# AQM Assignment One

### Ada Ziemyte, Shalva, Anders Riis

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### a. ARIMA Models

Identify, estimate and diagnostic check ARIMA models for each of the following variables: ln inv, ln inc, and ln consump.

To identity an ARIMA model we must check for stationarity, correct for stationarity and examine the graphs for the autocorrelations and partial autocorrelations.

## a.i. ARIMA model for $ln_inv$

As can be seen in figure 1 on the following page, the ACF decays monotonically. That the ACF decays slowly indicates that the data is non-stationary, i.e. that we have a non-stationary mean.<sup>2</sup> Stationarity is when we have a flat looking series without trend.<sup>3</sup> The scatterplot in figure 1 on the next page also indicates

 $<sup>^{1}</sup> http://www.polsci.wvu.edu/duval/ps791c/Notes/Stata/arimaident.html \\ ^{2} http://www.edu/duval/ps791c/Notes/Stata/arimaident.html \\ ^{2} http://www.edu/duval/ps791c/Notes/Stata/arimaident.html \\ ^{2} http://www.edu/duval/ps791c/Notes/Stata/arimaident.html \\ ^{2} http://www.edu/duval/ps791c/Notes/Stata/arimaident.html \\ ^{2} http://www.$ 

 $<sup>^2</sup> http://www.polsci.wvu.edu/duval/ps791c/Notes/Stata/arimaident.html\\$ 

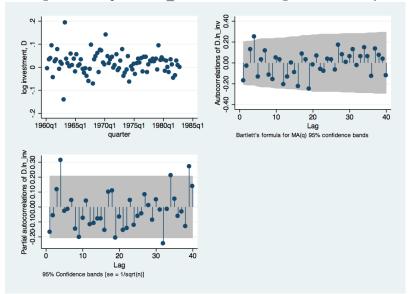
 $<sup>^3</sup> http://www.itl.nist.gov/div898/handbook/pmc/section4/pmc442.htm$ 

that the series is non-stationary as we can see a clear trend. The PACF cuts off after one lag (see figure 1), so we probably have an AR(1) process. However, because of non-stationarity we first attempt to correct for this by taking the first difference.

Now in figure 2 on the next page we see that the scatterplot is flatter, portraying stationarity. The ACF decays in an alternating or oscillating fashion, probably portraying some sort of MA as it seems to cut off (becomes zero) after four values or after zero values (q=0). The PACF never seems to cut off (p=0), although it can be seen as decaying. Thus, we have evidence of an ARIMA(0,1,0) model:

```
. arima ln_inv, arima(0,1,0)
(setting optimization to BHHH)
Iteration 0: log likelihood = 154.00497
              log likelihood = 154.00497
Iteration 1:
ARIMA regression
Sample: 1960q2 - 1982q4
                                               Number of obs
                                                                           91
                                               Wald chi2(.)
Log likelihood =
                  154.005
                                               Prob > chi2
            OPG
                           Std. Err.
                                                         [95% Conf. Interval]
   D.ln_inv |
                   Coef.
                                               P>|z|
ln_inv
```

Figure 2: Graphs for *ln inv* after correcting for stationarity



_cons	.0167964	.0047276	3.55	0.000	.0075304	.0260623
+						
/sigma	.044543	.0020718	21.50	0.000	.0404823	.0486037

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at z

It does not make sense to do an armaroot test as we have no AR or MA parameters in the model.

Another way of interpreting the graphs in figure 2 is that we have an ARIMA(4,1,4) model, the results of which are:

. arima  $ln_inv$ , arima(4,1,4)

```
(setting optimization to BHHH)
Iteration 0:
              log likelihood =
                                151.19416
              log likelihood =
                                153.8349
Iteration 1:
Iteration 2:
              log likelihood = 155.72441
              log likelihood = 158.06115
Iteration 3:
              log likelihood = 158.78522
Iteration 4:
(switching optimization to BFGS)
              log likelihood = 160.42741
Iteration 5:
Iteration 6:
              log likelihood = 161.31532
Iteration 7:
              log likelihood = 161.57824
Iteration 8:
              log likelihood = 162.63005
Iteration 9:
              log likelihood = 162.99665
Iteration 10: log likelihood = 163.47528
Iteration 11: log likelihood = 163.81731
Iteration 12: log likelihood = 164.09801
Iteration 13: log likelihood = 164.17117
```

```
Iteration 14: log likelihood = 164.23693
(switching optimization to BHHH)
Iteration 15: log likelihood = 164.28633
Iteration 16: log likelihood = 164.28633
                                           (backed up)
Iteration 17: log likelihood = 164.28633
                                           (not concave)
Iteration 18: log likelihood = 164.29517
Iteration 19: log likelihood = 164.29517
                                           (backed up)
(switching optimization to BFGS)
Iteration 20: log likelihood = 164.29519
                                           (backed up)
Iteration 21: log likelihood = 164.29682
Iteration 22: log likelihood = 164.30153
Iteration 23: log likelihood = 164.30185
Iteration 24: log likelihood = 164.3047
Iteration 25: log likelihood = 164.30508
Iteration 26: log likelihood = 164.30518
Iteration 27: log likelihood = 164.30518
ARIMA regression
Sample: 1960q2 - 1982q4
                                               Number of obs
                                                                           91
                                               Wald chi2(7)
                                                                       361.32
Log likelihood = 164.3052
                                               Prob > chi2
                                                                       0.0000
______
                            OPG
                 Coef. Std. Err.
   D.ln_inv |
                                       z P>|z| [95% Conf. Interval]
_cons | .0170483 .0066675 2.56 0.011 .0039803
                                                                    .0301164
ARMA
         ar |
        L1. |
                .250352 .3175816 0.79 0.431
                                                        -.3720964
                                                                    .8728004

    L2. | -.6410894
    .3055835
    -2.10
    0.036
    -1.240022

    L3. | .6959524
    .2326489
    2.99
    0.003
    .2399689

    L4. | -.1234277
    .371418
    -0.33
    0.740
    -.8513935

                                                                   -.0421568
                                                                    1.151936
                                                                   .6045381
        L1. | -.4332859 .3581174 -1.21 0.226 -1.135183

L2. | .7828925 .3674864 2.13 0.033 .0626323

L3. | -.7487543 . . . . . . . . . .
                                                                    .2686114
                                                                    1.503153
        L4. | .5336774 .347815 1.53 0.125
______
     /sigma | .0387658 .0050688 7.65 0.000 .0288311 .0487004
```

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at z

To test for stationarity we use the -armaroots- command to know whether the absolute value of the roots of the characteristic equation are smaller than one, which is required for stationarity. As we can see in the following table all the roots are smaller than one in absolute value:

#### . armaroots

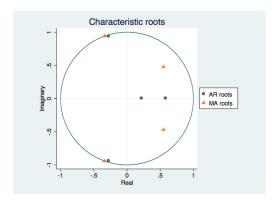
Characteristic roots of AR-polynomial

1	Characteri	stic roots	•				
•					T		!
-	2762973 +	.93//0/51	- 1	.977566	ı	3.38	ı
	2762973 -	.9377075i	-	.977566	-	3.38	- 1
-	.5804232		-1	.580423	-		- 1
-	.2225233		-1	.222523	- 1		- 1
_							_

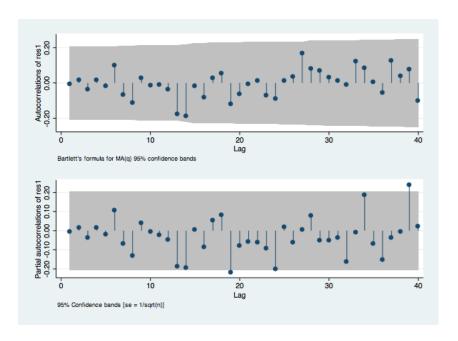
### ${\tt Characteristic\ roots\ of\ MA-polynomial}$

+-							+
 	Characteris			Modulus		Period	 
i	3382539 +	.941054i	Ī	.999999	İ	3.28	i
	3382539 -	.941054i		.999999	- 1	3.28	- 1
1	.5548968 +	.4751503i	-	.730533	- 1	8.87	-
1	.5548968 -	.4751503i	-	.730533	- 1	8.87	-
+-							+

Graphically we can also confirm that none of the roots are larger than one in absolute value:



We now turn to the ACF and the PACF for the residuals to confirm that these have no distinct pattern:



The graph above indeed seems to confirm that the residuals are white noise. The Portmanteau statistics confirms this by yielding:

. wntestq res1

```
Portmanteau test for white noise
-----

Portmanteau (Q) statistic = 30.0852

Prob > chi2(40) = 0.8728
```

## a.ii. ARIMA model for $ln\_inc$

We now turn to the variable  $ln\_inc$ . The graphs shown in 3 on the next page again demonstrate that we probably have non-stationary data, so we do a first difference transformation again.

The results of this first difference transformation can be seen in figure 4 on the following page. Now the data does not portray a non-stationary mean, so we decide that d = 1. Again the case is more ambiguous for the ACF and PACF. We could again conclude that the ACF never cuts off, so we have a MA(0), where q = 0, and the PACF never cuts off (becomes zero), so p = 0.

Estimating an ARIMA(0,1,0) model gives the results shown in figure 5 on page 8.

The Portmanteau statistics yield...

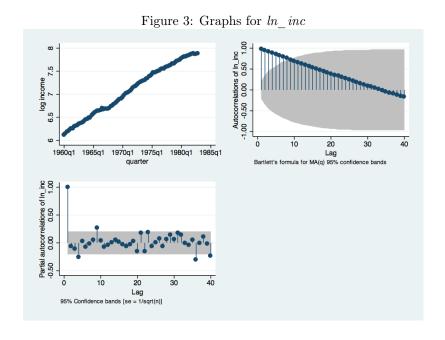


Figure 4: Graphs of  $ln_i$  inc after first differencing  $\frac{8}{8}$   $\frac{1}{9}$   $\frac{1}{9}$ 

Figure 5: Results from ARIMA(0,1,0) of $ln_i$						
		(1)	(2)	_		
	VARIABLES	$\ln_{\mathrm{inc}}$	$_{ m sigma}$	_		
	Constant	0.0195*** (0.00129)	0.0119*** (0.000645)			
	Observations	91	91			
	Standard errors in parentheses					
	*** p<0.01	, ** p<0.05,	* p<0.1			

### a.iii. ARIMA model for ln consump

Identifying an ARIMA model for the variable  $ln\_consump$  we again first look at some graphs in figure 6 on the following page. Again we have to take the first difference to correct for stationarity. After this is done we see in the lower pane that we have a stationary mean.

Estimating an ARIMA(?,1,?) gives...

## b. Dynamic Linear Model

Formulate and estimate a dynamic linear model with four lags for  $ln\_consump$  with  $ln\_inc$  and  $ln\_inv$  as explanatory variables.

The model can be written as

$$y_t = a_0 + a_1 y_{t-1} + a_2 y_{t-2} + a_3 y_{t-3} + a_4 y_{t-4} + b_0 \ln inc_t + c_0 \ln inv_t + u_t, \quad (1)$$

where y is  $ln\_consump.^4$ 

In figure 7 on page 11 we see the results of a dynamic linear model with four lags for  $ln\_consump$ . Mis-specification tests are:

```
Ramsey RESET test using powers of the fitted values of ln_consump Ho: model has no omitted variables F(3,\ 78) = 6.02 Prob > F = 0.0010
```

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity
Ho: Constant variance

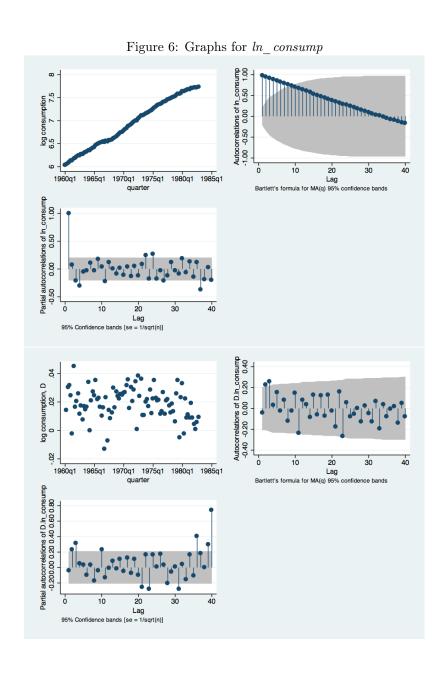
Variables: fitted values of ln\_consump

$$chi2(1) = 1.10$$
  
Prob > chi2 = 0.2944

White's test for Ho: homoskedasticity

against Ha: unrestricted heteroskedasticity

 $<sup>^4{\</sup>rm See}$  AQM lecture notes chapter 3, p. 18.



chi2(27) = 30.03Prob > chi2 = 0.3130

Cameron & Trivedi's decomposition of IM-test

				_		
Source	chi2					
Heteroskedasticity						
	6.37					
	1 4.75					
Nui 00515	•					
Total	41.14	34	0.1863			
				-		
(n = 88)   D	 -H P-value					
Residuals   6.15		3.60	067	0.1647		
Breusch-Godfrey LM t						
lags(p)	chi2	C	lf		Prob > chi2	
·	2.096		1		0.1476	
HO: no serial correlation LM test for autoregressive conditional heteroskedasticity (ARCH)						
lags(p)			lf		Prob > chi2	

HO: no ARCH effects vs. H1: ARCH(p) disturbance

## c. Error Correction Model

1 | 1.558

Reformulate the model in b. as an error correction model; store the log-likelihood value using the command—estimates store llu-; and subject the model to mis-specification testing.

1

Writing equation 1 on page 8 as an error correction model yields:

 $\Delta y_t = a_0 + \alpha_1 \Delta y_{t-1} + \alpha_2 \Delta y_{t-2} + \alpha \Delta_3 y_{t-3} + \beta_0 \Delta ln\_inc_t + \gamma_0 \Delta ln\_inv_t + ay_{t-1} + u_t,$  again where y is  $ln\_consump$ .<sup>5</sup> Running the regression gives:

 $<sup>^5{\</sup>rm See}$  answer to assignment 4, IQM, p. 13.

 $\begin{tabular}{ll} Figur \underline{e} \begin{tabular}{ll} Figur \underline{e} \begin{tabular}{ll} F: Dynamic linear model with four lags \\ \hline (1) \end{tabular}$ 

VARIABLES	ln consump
L.ln_consump	0.405***
	(0.106)
$L2.ln\_consump$	0.278**
	(0.113)
$L3.ln\_consump$	0.118
	(0.111)
L4.ln_consump	-0.154*
	(0.0871)
ln inc	0.333***
_	(0.0490)
ln inv	0.00970
_	(0.0121)
Constant	0.0493***
	(0.0135)
Observations	88
R-squared	1.000
Standard errors	in parentheses
*** p<0.01, ** p	<0.05, * p<0.1
- / -	

Source	SS	df			Number of obs	
+-					F( 6, 81)	
	.004719858				Prob > F	
Residual	.005759951	81 .00	0071111		R-squared	= 0.4504
+-					Adj R-squared	= 0.4097
Total	.010479809	87 .000	0120458		Root MSE	= .00843
dln_consump	Coef.	Std. Err.	t	P> t	[95% Conf.	Interval]
+-						
${\tt dln\_consump}$						
L1.	2673194	.087738	-3.05	0.003	4418905	0927483
L2.	.1660194	.0843342	1.97	0.052	0017792	.333818
L3.	.2591583	.0865452	2.99	0.004	.0869605	.4313562
1						
dln_inc	.4554999	.0815985	5.58	0.000	.2931446	.6178553
dln_inv	.062451	.0202667	3.08	0.003	.0221265	.1027754
1						
ln_consump						
L1.	0004039	.0018446	-0.22	0.827	0040741	.0032664
1						
_cons	.008395	.0138116	0.61	0.545	0190858	.0358757

The results of mis-specification tests are:

```
Ramsey RESET test using powers of the fitted values of dln_consump
```

Ho: model has no omitted variables F(3, 78) = 1.19Prob > F = 0.3193

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity

Ho: Constant variance

Variables: fitted values of dln\_consump

chi2(1) = 0.47Prob > chi2 = 0.4911

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity

Ho: Constant variance

Variables: L.dln\_consump L2.dln\_consump L3.dln\_consump dln\_inc dln\_inv L.ln\_consump

chi2(6) = 7.60Prob > chi2 = 0.2686

White's test for Ho: homoskedasticity

against Ha: unrestricted heteroskedasticity

chi2(27) = 38.39Prob > chi2 = 0.0719

Cameron & Trivedi's decomposition of IM-test  $\,$ 

Source | chi2 df p

	-+			
Heteroskedasticity	1	38.39	27	0.0719
Skewness	1	9.38	6	0.1532
Kurtosis	1	0.65	1	0.4218
	+			
Total	1	48.42	34	0.0519

(n = 88)	 P-value	•	P-value
Residuals	0.2732	2.4639	0.2917

Breusch-Godfrey LM test for autocorrelation

lags(p		chi2	df	Prob > chi2
1 2	l I	4.266 8.164	1 2	0.0389 0.0169
3	i	9.789	3	0.0205
4	I	10.731	4	0.0298

HO: no serial correlation

LM test for autoregressive conditional heteroskedasticity (ARCH)

lags(p)		chi2	df	Prob > chi2
1 2 3		3.184 8.175 1.331	1 2 3	0.0743 0.0168 0.7217
4	ı	0.949	4	0.9175

HO: no ARCH effects vs. H1: ARCH(p) disturbance

The Ramsey RESET test and the two Breusch-Pagan tests for heterosked asticity do not reveal any problems as we have p-values comfortably above, say, a 10% significance level. However, the White's test for homosked asticity is rejected at a 10% significance level. Moreover, the Breusch-Godfrey LM test for autocorrelation shows some problems as the null hypothesis of no serial correlation is rejected for all the lags at the 5% significance level.

## d. Model Specification

Reduce the model in successive steps by dropping insignificant variables (subjecting each step to mis-specification testing) in order to obtain a parsimonious model.

We now drop the first lag of  $ln\_consump$  in the error correction model above. This gives us the following result:

Source	SS	df		MS		Number of obs	=	88
+						F( 5, 82)	=	13.42
Model	.004716449	5	.00	0094329		Prob > F	=	0.0000
Residual	.005763359	82	.000	0070285		R-squared	=	0.4501
+						Adj R-squared	=	0.4165
Total	.010479809	87	.000	120458		Root MSE	=	.00838
-						[95% Conf.		
dln_consump								
L1.	2657825	.0869	475	-3.06	0.003	4387488		0928161
L2.	.1683854	.083	152	2.03	0.046	.0029696		3338012
L3.	.2603634	.0858	671	3.03	0.003	.0895463		4311805
1								
dln_inc	.4581345	.0802	364	5.71	0.000	.2985187		6177503
dln_inv	.0626535	.0201	277	3.11	0.003	.022613		.102694
_cons	.0054381	.0028	747	1.89	0.062	0002807		0111569

Now we see that we do not have insignificant variables.

The mis-specification tests of this model are:

Ramsey RESET test using powers of the fitted values of dln\_consump

Ho: model has no omitted variables F(3, 79) = 1.20

Prob > F = 0.3151

Breusch-Pagan / Cook-Weisberg test for heteroskedasticity

Ho: Constant variance

Variables: fitted values of dln\_consump

chi2(1) = 0.57Prob > chi2 = 0.4492

 ${\tt Breusch-Pagan \ / \ Cook-Weisberg \ test \ for \ heterosked a sticity}$ 

Ho: Constant variance

Variables: L.dln\_consump L2.dln\_consump L3.dln\_consump dln\_inc dln\_inv

chi2(5) = 4.26Prob > chi2 = 0.5122

White's test for Ho: homoskedasticity

against Ha: unrestricted heteroskedasticity

chi2(20) = 25.95Prob > chi2 = 0.1675

Cameron & Trivedi's decomposition of IM-test

Source	chi2	df	p
Heteroskedasticity	25.95	20	0.1675
Skewness	9.69	5	0.0846

_	Kurtosis		_	
		36.25		

Residuals   2.7414 0.2539 2.6979 0.2595	(n = 88)	D-H	P-value	asy.	P-value

 ${\tt Breusch\text{-}Godfrey}\ {\tt LM}\ {\tt test}\ {\tt for}\ {\tt autocorrelation}$ 

lags(p)	   +	chi2	df	Prob > chi2
1 2 3 4	       	4.309 8.027 9.487 10.267	1 2 3 4	0.0379 0.0181 0.0235 0.0362

HO: no serial correlation

LM test for autoregressive conditional heteroskedasticity (ARCH)

lags(p)		chi2	df	Prob > chi2
1		3.173	1	0.0748
2		8.125	2	0.0172
3		1.381	3	0.7100
4		0.998	4	0.9101

HO: no ARCH effects vs. H1: ARCH(p) disturbance

We now test whether we can exclude the last lag, but the -testparm- command does lead to rejection, so we keep the third lag.

(1) L3.dln\_consump = 0

F(1, 82) = 9.19Prob > F = 0.0032

## e. Comparison of Models

Save the log-likelihood value of the parsimonious model (-estimates store llr-)and perform a likelihood ratio test comparing the model in c. with the model in d. (-lrtest llu llr). You may wish to save the log-likelihood value at each step of the reduction using different names with the estimates command at each step.

. lrtest llu1 llu2

LR chi2(1) = 0.05

# f. VAR(4) Model

(Assumption: 11u2 nested in 11u1)

Estimate a VAR(4) for the three variables dln\_consump, dln\_inc, and dln\_inv, that is, for the first difference of ln\_inv, ln\_inc, and ln\_consump. Then determine the minimum number of lags required for the VAR, and subject it to mis-specification testing.

We determine the lag structure by starting with three lags:

. var dln\_consump dln\_inc dln\_inv, lags(1/3) dfk

### Vector autoregression

Sample: 1961	q1 -	1982q4	No. of obs	=	88
Log likelihoo	d =	737.5558	AIC	=	-16.08081
FPE	=	2.09e-11	HQIC	=	-15.74057
Det(Sigma_ml)	=	1.05e-11	SBIC	=	-15.23627

Equation	Parms	RMSE	R-sq	chi2	P>chi2
<pre>dln_consump dln_inc dln_inv</pre>	10 10 10	.00973 .011309 .045251	0.2953 0.1800 0.1096	32.68978 17.12733 9.599902	0.0002 0.0468 0.3838

		Std. Err.			[95% Conf	. Interval]
dln_consump						
dln_consump						
L1.	3930788	.1385722	-2.84	0.005	6646753	1214823
L2.	1281016	.1441055	-0.89	0.374	4105431	. 15434
L3.	.1015717	.1278624	0.79	0.427	149034	.3521773
dln_inc						
L1.	.2940007	.1174905	2.50	0.012	.0637236	.5242778
L2.	.3438497	.1205479	2.85	0.004	.1075802	.5801192
L3.	.1810283	.1187276	1.52	0.127	0516734	.4137301
dln_inv						
L1.	.0034602	.0256856	0.13	0.893	0468826	.0538031
L2.	.0437956	.0264529	1.66	0.098	0080512	.0956423
L3.		.0263789	0.64	0.523	034872	.0685314
_cons	.0091472	.0034369	2.66	0.008	.0024111	.0158833
dln_inc						
dln_consump						
L1.	.1923453	.161054	1.19	0.232	1233148	.5080055
L2.	0339177	.167485	-0.20	0.840	3621824	.294347
L3.	0434862	.1486067	-0.29	0.770	33475	. 2477776

dln_inc	l					
L1.	0815851	.136552	-0.60	0.550	3492222	.186052
L2.	.0624807	.1401055	0.45	0.656	212121	.3370824
L3.	.1984851	.1379898	1.44	0.150	07197	.4689402
	l					
dln_inv	I					
L1.	.0476297	.0298528	1.60	0.111	0108808	.1061401
L2.	.0598946	.0307446	1.95	0.051	0003637	.1201529
L3.	.0155151	.0306586	0.51	0.613	0445747	.0756048
	l					
_cons	.011234	.0039945	2.81	0.005	.003405	.019063
	+					
dln_inv	l					
dln_consump	l					
L1.	.4585587	.644436	0.71	0.477	8045127	1.72163
L2.	.4443132	.6701688	0.66	0.507	8691936	1.75782
L3.	199181	.5946297	-0.33	0.738	-1.364634	.9662718
	l					
dln_inc	I					
L1.	.3711541	.5463946	0.68	0.497	6997597	1.442068
L2.	.2534399	.5606132	0.45	0.651	8453418	1.352222
L3.	.3569366	.5521478	0.65	0.518	7252532	1.439126
	I					
dln_inv	I					
L1.	2598316	.1194521	-2.18	0.030	4939533	0257098
L2.	1253296	.1230204	-1.02	0.308	3664451	.115786
L3.	.0466697	.1226763	0.38	0.704	1937714	.2871107
	I					
_cons	0100135	.0159833	-0.63	0.531	0413401	.0213131

. varsoc

Selection-order criteria Sample: 1961q1 - 1982q4

Number of obs = 88

+										-+
١	lag	LL	LR	df	p	FPE	AIC	HQIC	SBIC	1
									-15.8428*	•
- 1	1	719.587	31.572	9	0.000	2.1e-11	-16.0815	-15.9454*	-15.7437	1
- 1	2	732.981	26.787*	9	0.002	1.9e-11*	-16.1814*	-15.9432	-15.5902	1
- 1	3	737.556	9.1506	9	0.423	2.1e-11	-16.0808	-15.7406	-15.2363	-
+										-+

Endogenous: dln\_consump dln\_inc dln\_inv

Exogenous: \_cons

The -varsoc- command shows us that we can drop the third lag, so we try with two lags:

. dln\_consump dln\_inc dln\_inv, lags(1/2) dfk

### Vector autoregression

Sample: 1960q4 - 1982q4 No. of obs = 89
Log likelihood = 742.2131 AIC = -16.20704
FPE = 1.84e-11 HQIC = -15.97035
Det(Sigma\_ml) = 1.15e-11 SBIC = -15.61983

Equation	Parms	RMSE	R-sq	chi2	P>chi2	
dln_consump	7	.009938	0.2400	25.88962	0.0002	
dln_inc	7	.011224	0.1514	14.62996	0.0233	
dln_inv	7	.044295	0.1051	9.633771	0.1409	
	Coef.	Std. Err.		P> z	[95% Conf	. Interval]
dln_consump						
dln_consump						
L1.	2845172	.1274068	-2.23	0.026	5342299	0348044
L2.	1159776	.1270233	-0.91	0.361	3649386	.1329834
43 d.u. a						
dln_inc		110010	0.50	0.010	060101	E004407
L1.		.112313	2.58	0.010	.069191	.5094497
L2.	.3664341 	.1089872	3.36	0.001	.1528231	.5800451
dln_inv						
L1.	.0027381	.0255556	0.11	0.915	0473499	.0528261
L2.	.0497402	.025462	1.95	0.051	0001644	.0996447
		0000000	4 40	0.000	0005770	0404040
_cons	.0123795 	.0029602	4.18	0.000	.0065776	.0181813
${\tt dln\_inc}$	1					
dln_consump						
L1.		.1438994	2.12	0.034	.0230195	.5870947
L2.	.0490208	.1434662	0.34	0.733	2321677	.3302094
	l					
dln_inc						
L1.		.1268517	-0.97	0.331	371879	.1253704
L2.	.0209769	.1230954	0.17	0.865	2202857	. 2622394
dln_inv	! 					
_ L1.		.0288637	1.50	0.133	0132244	.0999191
L2.	.0616319	.028758	2.14	0.032	.0052673	.1179965
	l					
_cons	.0125949 	.0033434	3.77	0.000	.006042	.0191478
dln_inv	l					
dln_consump						
L1.	.6520473	.5678885	1.15	0.251	4609936	1.765088
L2.	.5980687	.566179	1.06	0.291	5116217	1.707759
dln_inc						
L1.	.3374819	.5006109	0.67	0.500	6436975	1.318661
L2.	.1827302	.4857871	0.38	0.707	7693951	1.134855
dln_inv	] 					
L1.		.1139085	-2.39	0.017	4958218	0493089
L1. L2.		.1134912	-2.39	0.017	3564891	.0883884
ьг.	1340503	.1104312	-1.10	0.230	3504031	.0000004
_cons	0099191	.0131944	-0.75	0.452	0357798	.0159415

<sup>.</sup> varsoc

```
Selection-order criteria
Sample: 1960q4 - 1982q4
                              Number of obs =
 | \texttt{lag} \ | \ \texttt{LL} \qquad \texttt{LR} \qquad \texttt{df} \quad p \qquad \texttt{FPE} \qquad \texttt{AIC} \qquad \texttt{HQIC} \qquad \texttt{SBIC} \quad | 
0 | 712.092
                   2.4e-11 -15.9346 -15.9008 -15.8508* |
| 1 | 728.167 32.15 9 0.000 2.1e-11 -16.0936 -15.9584 -15.7581 |
+-----+
```

Endogenous: dln\_consump dln\_inc dln\_inv

Exogenous: \_cons

#### Final Model g.

Finally, drop individual variables from each of the three equations until you end up with a parsimonious model.

We now drop individual variables from each equation until we end up with the following parsimonious model with eight constraints:

. var dln\_consump dln\_inc dln\_inv, lags(1/2) dfk constraint (1 2 3 4 5 6 7 8) Estimating VAR coefficients

```
Iteration 1:
            tolerance = .08076175
Iteration 2: tolerance = .00197295
Iteration 3: tolerance = .0000493
Iteration 4: tolerance = 1.529e-06
Iteration 5: tolerance = 5.099e-08
```

### Vector autoregression

Sample: 19	60q4 -	1982q4	No. of obs	=	89
Log likelih	ood =	739.5869	AIC	=	-16.14802
FPE	=	2.18e-15	HQIC	=	-15.91133
Det(Sigma_m	1) =	1.63e-15	SBIC	=	-15.56081

Equation	Parms	RMSE	R-sq	chi2	P>chi2
dln_consump dln_inc dln inv	5	.009606	0.2292	37.83738	0.0000
	4	.01085	0.1394	15.67358	0.0013
	4	.042994	0.0849	9.673415	0.0216

- ( 1) [dln\_consump]L.dln\_inv = 0
- $(2) [dln_inc]L2.dln_inc = 0$
- $(3) [dln_inv]L2.dln_inc = 0$
- (4) [dln\_inc]L2.dln\_consump = 0
- (5) [dln\_inv]L.dln\_inc = 0
- (6) [dln\_inc]L.dln\_inc = 0
- $(7) [dln_inv]L2.dln_inv = 0$
- (8) [dln\_consump]L2.dln\_consump = 0

	Coef.	Std. Err.	z	P> z	[95% Conf.	Interval]
dln_consump dln_consump	   					
L1.	2560975	.1075375	-2.38	0.017	4668671	0453279
L2.	2.70e-18	1.79e-17	0.15	0.880	-3.24e-17	3.78e-17
	l					
dln_inc						
L1.		.0825004	3.52	0.000	.1282945	.4516901
L2.	. 286996	.0723422	3.97	0.000	.1452079	.4287841
dln_inv		F	0.00	0.004	1 00 17	4 40 47
L1.		5.57e-18	0.02	0.984	-1.08e-17	1.10e-17
L2.	.0515305	.0226578	2.27	0.023	.0071221	.095939
_cons	.0112223	.0024842	4.52	0.000	.0063534	.0160913
dln_inc	r I					
dln_consump						
L1.		.1110663	2.12	0.034	.0177771	.4531489
L2.		2.73e-17	-1.22	0.224	-8.67e-17	2.03e-17
	1	21.00 1.		0.221	0.0.0 1.	27000 1.
dln_inc	ĺ					
_ L1.	-3.35e-17	2.71e-17	-1.24	0.215	-8.66e-17	1.95e-17
L2.	-1.73e-17	1.61e-17	-1.07	0.282	-4.88e-17	1.42e-17
	l					
dln_inv						
L1.	.0472669	.0230169	2.05	0.040	.0021546	.0923791
L2.	.0629894	.0263325	2.39	0.017	.0113786	.1146003
_cons	.0127585	.0024014	5.31	0.000	.0080519	.0174651
	+ ,					
dln_inv						
dln_consump		4462472	1 00	0 046	0146450	1 764005
L1.		.4463473	1.99	0.046	.0146459	1.764295
L2.	7532321	.4187462	1.80	0.072	0674953	1.573959
dln_inc						
L1.		1.10e-16	-0.80	0.421	-3.04e-16	1.27e-16
L2.		2.14e-16	-0.40	0.692	-5.05e-16	3.35e-16
ш.	-0.40C-17 	2.140-10	-0.40	0.002	-0.000-10	0.000-10
${\tt dln\_inv}$						
_ L1.		.1045915	-2.45	0.014	4612356	0512443
L2.	-2.63e-17	2.54e-17	-1.04	0.300	-7.62e-17	2.35e-17
_cons	0097218	.0125614	-0.77	0.439	0343416	.014898

### . varlmar

### Lagrange-multiplier test

+						-+
	0		chi2		Prob > chi2	
•			-7.1e+02	9		i
1	2	-	-7.2e+02	9	1.00000	-

HO: no autocorrelation at lag order

### . varnorm

### Jarque-Bera test

4	+				-+
	Equation	chi2		Prob > chi2	: [
	+   dln_consump	11.036		0.00401	-
i	dln_consump   dln_inc	0.869	2	0.64764	i
i	dln_inv	12.553	2	0.00188	i
١	ALL	24.457	6	0.00043	-1

### Skewness test

+						+
E	quation   Sk				Prob > chi	
dln_	consump	.74579	8.250	1	0.00407	i
1	dln_inc	. 17547	0.457	1	0.49915	- 1
1	dln_inv   .	.42156	2.636	1	0.10446	- 1
1	ALL		11.343	3	0.01001	
+						+

### Kurtosis test

+					+
Equation	1	Kurtosis	chi2	df	Prob > chi2
	-+-				
dln_consump	1	3.8666	2.785	1	0.09514
dln_inc	$\perp$	2.6666	0.412	1	0.52091
dln_inv	$\perp$	4.6353	9.917	1	0.00164
ALL	$\perp$		13.114	3	0.00440
+					+

dfk estimator used in computations