

# Tutorial, Part 7

Model diagnostics, ctd.

## Example

### c) Jarque-Bera test

$H_0$ : Residuals (variable  $u$ ) are normally distributed.

$H_1$ :  $\text{---} \parallel \text{---}$  NOT  $\text{---} / \text{---} \text{---}$

$$S = -0.119$$

$$K = 3.034$$

$$n = 96$$

$$\begin{aligned} JB &= n \cdot \left( \frac{S^2}{6} + \frac{(K-3)^2}{24} \right) = 96 \cdot \left( \frac{(-0.119)^2}{6} + \frac{(3.034-3)^2}{24} \right) = \\ &= 0.231 \sim \chi^2_{(2)} \end{aligned}$$

$$\chi^2_{c(\alpha=0.05)} = 5.997$$

$JB < \chi^2_c \Rightarrow$  cannot reject  $H_0$  (at  $\alpha=0.05$ )

(State: jb 6  $\rightarrow p = 0.8909 > \alpha=0.05$ )

### d) Multicollinearity caused by variable PPR

$$\boxed{\text{MME:}} \quad \hat{PPR}_t = b_1 + b_2 RVP_t + b_3 RVV_t + b_4 CZP_t$$

$$\hookrightarrow R^2_{PPR} = 0.8847$$

$$VIF_{PPR} = \frac{1}{1-R^2_{PPR}} = \frac{1}{1-0.8847} = 8.676 \underset{?}{\approx} 10$$

### e) White test

$$\boxed{H_B:} \quad HM1_t = b_1 + b_2 PPR_t + b_3 RVP_t + b_4 RW_t + b_5 CZP_t + \epsilon_t$$

Assume:  $G_t^2 = \alpha_1 + \alpha_2 PPR_t + \alpha_3 RVP_t + \alpha_4 RW_t + \alpha_5 CZP_t + \alpha_6 PPR_t^2 + \alpha_7 RVP_t^2 + \alpha_8 RW_t^2 + \alpha_9 CZP_t^2 + \alpha_{10} PPR_t \cdot RVP_t + \alpha_{11} PPR_t \cdot RW_t + \alpha_{12} PPR_t \cdot CZP_t + \alpha_{13} RVP_t \cdot RW_t + \alpha_{14} RVP_t \cdot CZP_t + \alpha_{15} RW_t \cdot CZP_t$

$$H_0: \alpha_j = 0 \quad ; \quad H_j = 2, \dots, 15 \quad (\text{homoscedasticity})$$

$$H_1: \alpha_j \neq 0 \quad ; \quad \exists j = 2, \dots, 15 \quad (\text{heteroscedasticity})$$

$$\boxed{M_{AR}:} \quad \hat{\epsilon}_t^2 = \alpha_1 + \alpha_2 PPR_t + \alpha_3 RVP_t + \alpha_4 RW_t + \alpha_5 CZP_t + \alpha_6 PPR_t^2 + \alpha_7 RVP_t^2 + \alpha_8 RW_t^2 + \alpha_9 CZP_t^2 + \alpha_{10} PPR_t \cdot RVP_t + \alpha_{11} PPR_t \cdot RW_t + \alpha_{12} PPR_t \cdot CZP_t + \alpha_{13} RVP_t \cdot RW_t + \alpha_{14} RVP_t \cdot CZP_t + \alpha_{15} RW_t \cdot CZP_t$$

(LM)  $\hookrightarrow R_{AR}^2 = 0.5607$

$$\chi^2 = n \cdot R_{AR}^2 = 96 \cdot 0.5607 = 53.83 \sim \chi^2_{k-1} \quad \text{MAR}$$

$$\chi^2_C (m=k-1=14, \alpha=0.05) = 23.685$$

$\chi^2 > \chi^2_C \Rightarrow$  reject  $H_0$  at  $\alpha=0.05$  and conclude from  $H_1 \Rightarrow$  there is heteroscedasticity problem in the model  
 (state: p-value  $< 0.0000$ )

Solution: robust estimator

(HC - heteroscedasticity consistent estimator)  $\rightarrow$  BLUE

f) Breusch-Godfrey test  
(e.g. for AC of order up to  $p=4$ )

Assume:  $U_t = \beta_1 U_{t-1} + \beta_2 U_{t-2} + \beta_3 U_{t-3} + \beta_4 U_{t-4} + \varepsilon_t$

$$\begin{aligned} H_0: \quad \beta_j = 0 & \quad ; \quad \nexists j = 1, \dots, 4 \quad (\text{no AC of order up to } p) \\ H_1: \quad \beta_j \neq 0 & \quad ; \quad \exists j = 1, \dots, 4 \quad (\text{there is AC of order up to } 4) \end{aligned}$$

MAR:  $\hat{U}_t = b_1 + b_2 PPR_t + b_3 RW_t + b_4 RV_t + b_5 CZP_t +$   
 $\hat{\beta}_1 \hat{U}_{t-1} + \hat{\beta}_2 \hat{U}_{t-2} + \hat{\beta}_3 \hat{U}_{t-3} + \hat{\beta}_4 \hat{U}_{t-4}$

$$LM = (n-p) \cdot R_{AR}^2 = (96-4) \cdot 0.7446 = 68.50 \sim \chi_p^2$$

$$\chi_c^2 (m=p=4, \alpha=0.05) = 9.4877$$

$LM > \chi_c^2 \Rightarrow$  reject  $H_0$  at  $\alpha=0.05$  and conclude  
from  $H_1 \Rightarrow$  there is AC of order  
up to 4 present in the model  
(Stet:  $p\text{-value} = 0.0005$ )

Solution: robust estimator

(HAC - heteroscedasticity &  
Auto correlation consistent  
estimator)  $\rightarrow$  ~~BLUE~~

(Max lag:  $n-1-k = 96-1-5 = 90$ )

# Tutorial, Part 8

Dummy expl. vars

## Example 1

$$M_B : \hat{\text{PROFIT}}_t = b_1 + b_2 \text{SALES}_t$$

quarterly data  $\rightarrow$  likely presence of seasonal component

SEASON : quarter 1, quarter 2, quarter 3, quarter 4  
(q)

Generate DUMMY VARIABLES :

$m$  categories  $\rightarrow m-1$  dummies (one category  
is base/reference)

Let quarter 1 be BASE category :

$$D_2 = \begin{cases} 1 &; \text{quarter 2} \\ 0 &; \text{else} \end{cases}$$

$$D_3 = \begin{cases} 1 &; \text{quarter 3} \\ 0 &; \text{else} \end{cases}$$

$$D_4 = \begin{cases} 1 &; \text{quarter 4} \\ 0 &; \text{else} \end{cases}$$

$$M_S : \hat{\text{PROFIT}}_t = b_1 + b_2 \text{SALES}_t + b_3 D_{2t} + b_4 D_{3t} + b_5 D_{4t}$$

$\hookrightarrow$  model with seasonal component

$D_j$  : differential intercept coefficient

If  $D_j$  insignificant ( $t_j$ )  $\rightarrow M_S$  same as  $M_B$

Stata:  $D_3$  and  $D_4$  insignificant?

$$H_0: \beta_4 = \beta_5 = 0$$

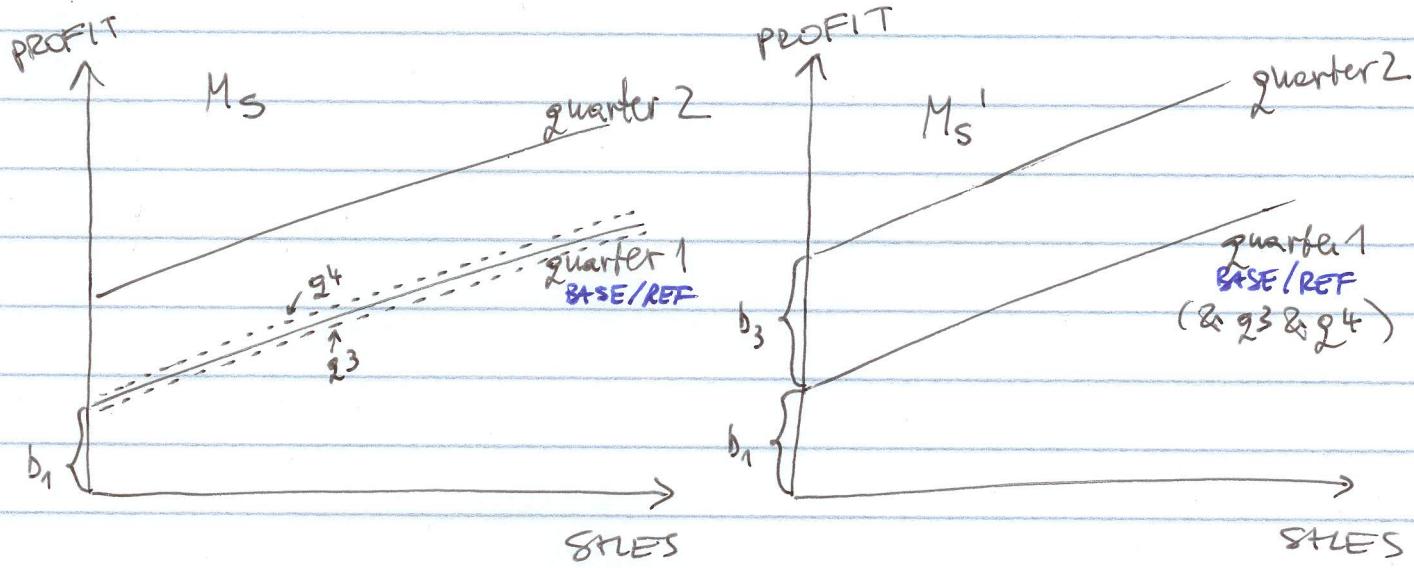
$$H_1: \beta_j \neq 0; j=4,5$$

$$F = 0.20$$

$$p\text{-value} = 0.8197 > \alpha = 0.05$$

→ cannot reject  $H_0$

MS<sup>1</sup>:  $\widehat{\text{PROFIT}}_t = b_1 + b_2 \text{SALES} + b_3 D_{2,t}$



Interpretation of partial regression coefficients:

$b_2 = 0.039$ : If sales increase by 1 billion  $(10^9)$  USD the profits increase on average by 39 million  $(10^6)$  USD, controlling for the seasonal effect.

$b_3 = 1.33$ : The profits in the 2<sup>nd</sup> quarter are on average by 1.33 billion USD higher compared to the 1<sup>st</sup> quarter <sup>BASE CATEGORY</sup> (as well as compared to the 3<sup>rd</sup> and 4<sup>th</sup>).  
↑ implicitly