Example 1: We will analyse monthly UK house price series for the period from January 1991 to May 2013, already prepared in Stata data file ukhp.dta. The programming code is given in Stata Do file ukhp-commands.do.

First, we will inspect, test and ensure stationarity of our monthly UK house price series.

- a) Inspect the UK house price (time series *HP*) visually. Does it look stationary? Perform the Augmented Dickey-Fuller (ADF) test on the time series. How many lags should you use? Should you include a trend into the DF regression? What do you find?
- b) Also perform the Dickey-Fuller GLS test (DFGLS), which has in general more statistical power than the ADF test. What do you find? You can use information on the number of lags from the DFGLS test in the ADF test. Has the conclusion changed?
- c) How can you make a time series stationary? Apply first differences on the time series *HP* and inspect the new time series visually. Does it look stationary? Repeat the ADF test on the new time series. What do you find?
- d) Another way of making a time series stationary, often used in finance, is to generate continuously compounded percentage returns (percentage changes). Do this and inspect the new time series visually. Repeat the ADF test on this time series. What do you find?

Now, we will build an ARMA model for the house price changes. Recall that there are three stages involved: identification, estimation and diagnostic checking.

- e) Identification is carried out by looking at the autocorrelation and partial autocorrelation coefficients to identify any structure in the data. Generate a table of autocorrelations, partial correlations and related test statistics. Draw the autocorrelation function (ACF) and the partial autocorrelation function (PACF). Does the time series appear to be an autoregressive (AR) process, a moving average (MA) process or a mixed ARMA process?
- f) Use information criteria to decide on model orders p and q. Consider at least the first four lags. Which information criterion will you focus on? What do you find?
- g) Estimate the model, proposed by the information criteria analysis, by employing the maximum likelihood estimator. Interpret the Stata output, but keep in mind that because the construction of these models is not based on economic or financial theory, it is often best not to interpret the individual parameter estimates, but rather to examine the plausibility of the model as a whole, and to determine whether it describes the data well and produces accurate forecasts. Also obtain the impulse responses (IRs).
- h) Check for stability of the chosen model by examining inverse AR and MA roots of the characteristic equation. What do you find?
- i) Now, suppose that the chosen time series model was estimated using observations from the period February 1991 December 2010, leaving 29 remaining observations for constructing forecasts (for the period January 2011 May 2013). Generate static and dynamic forecasts and present them graphically, together with the actual data, which are in fact available for the forecasting period. What do you find?

Static forecasts imply a sequence of one-step-ahead forecasts, rolling the sample forwards one observation after each forecast. Dynamic forecast calculates multi-step forecasts starting from the first period in the forecast sample. The dynamic forecasts quickly converge upon the long-term unconditional mean value as the horizon increases. This does not occur with the series of one-step-ahead forecasts, which seem to more closely resemble the actual series.

# Computer printout of the results in Stata:

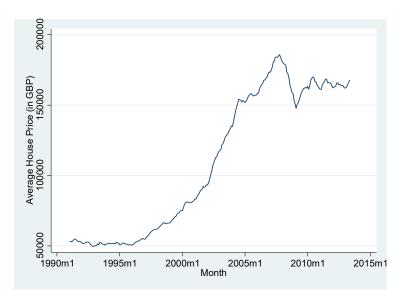
Data exploration

#### . tsset Month

time variable: Month, 1991m1 to 2013m5

delta: 1 month

### . twoway line HP Month



Performing the Dickey-Fuller tests

# . dfuller HP, lags(12)

Augmented Dickey-Fuller test for unit root Number of obs = 256

		Interpolated Dickey-Fuller					
	Test	1% Critical	5% Critical	10% Critical			
	Statistic	Value	Value	Value			
Z(t)	-0.960	-3.460	-2.880	-2.570			

MacKinnon approximate p-value for Z(t) = 0.7676

### . dfuller HP, regress lags(12)

Augmented Dickey-Fuller test for unit root Number of obs = 256

	Test	Interpolated Dickey-Fuller 1% Critical 5% Critical 10% Critica					
	Statistic	Value	Value	Value			
Z(t)	-0.960	-3.460	-2.880	-2.570			

MacKinnon approximate p-value for Z(t) = 0.7676

D.HP	Coef.	Std. Err.	t	P> t	 [95% Conf.	Interval]
	0014392 .2859919			0.338	0043926 .1612334	.0015143

L2D.		.3226638	.0656779	4.91	0.000	.1932906	.4520371
L3D.		.0217977	.0687253	0.32	0.751	1135784	.1571738
L4D.		.0249136	.0685106	0.36	0.716	1100396	.1598668
L5D.		0288369	.0684909	-0.42	0.674	1637514	.1060775
L6D.	1	.0128432	.0683735	0.19	0.851	1218399	.1475263
L7D.	i	1132738	.0683326	-1.66	0.099	2478763	.0213288
L8D.	i	.0238635	.0687485	0.35	0.729	1115583	.1592852
L9D.		.0884816	.0686746	1.29	0.199	0467948	.223758
L10D.		1151423	.0689386	-1.67	0.096	2509386	.020654
L11D.		.0911545	.0662825	1.38	0.170	0394098	.2217187
L12D.	1	.1653165	.0640052	2.58	0.010	.039238	.2913949
	i						
_cons	1	271.1407	183.6549	1.48	0.141	-90.62553	632.9069

# . dfuller HP, trend lags(12)

Augmented Dickey-Fuller test for unit root Number of obs = 256

		Interpolated Dickey-Fuller					
	Test	1% Critical	5% Critical	10% Critical			
	Statistic	Value	Value	Value			
Z(t)	-2.364	-3.990	-3.430	-3.130			

MacKinnon approximate p-value for Z(t) = 0.3987

# . dfuller HP, trend regress lags(12)

Augmented Dickey-Fuller test for unit root Number of obs = 256

Test 1% Critical 5% Critical 10% Critical Statistic Value Value Value  Z(t) -2.364 -3.990 -3.430 -3.130			Interpolated Dickey-Fuller					
		Test	1% Critical	5% Critical	10% Critical			
Z(t) -2.364 -3.990 -3.430 -3.130		Statistic	Value	Value	Value			
	Z(t)	-2.364	-3.990	-3.430	-3.130			

MacKinnon approximate p-value for Z(t) = 0.3987

D.HP	Coef.	Std. Err.	t	P> t	[95% Conf.	<pre>Interval]</pre>
HP						
L1.	0109715	.0046402	-2.36	0.019	0201121	0018309
LD.	.2790635	.0629371	4.43	0.000	.1550865	.4030405
L2D.	.3231062	.0651813	4.96	0.000	.1947084	.4515039
L3D.	.0278044	.0682616	0.41	0.684	1066611	.1622698
L4D.	.0331733	.0680988	0.49	0.627	1009716	.1673182
L5D.	0203513	.0680853	-0.30	0.765	1544694	.1137669
L6D.	.0162966	.0678748	0.24	0.810	1174071	.1500003
L7D.	1099157	.0678333	-1.62	0.106	2435375	.0237061
L8D.	.0258148	.0682343	0.38	0.706	1085969	.1602265
L9D.	.0919464	.0681738	1.35	0.179	0423462	.226239
L10D.	1083164	.0684894	-1.58	0.115	2432307	.0265979
L11D.	.0991286	.0658837	1.50	0.134	0306528	.22891
L12D.	.1798951	.0638756	2.82	0.005	.0540692	.3057209
trend	6.705113	3.091602	2.17	0.031	.6151001	12.79512
_cons	369.3101	187.8018	1.97	0.050	6325458	739.2527

#### . dfgls HP, maxlag(12)

256 DF-GLS for HP Number of obs =

[lags]	DF-GLS tau Test Statistic	1% Critical Value	5% Critical Value	10% Critical Value
12	-2.019	-3.480	-2.836	-2.554
11	-1.706	-3.480	-2.844	-2.561
10	-1.478	-3.480	-2.851	-2.568
9	-1.510	-3.480	-2.859	-2.575
8	-1.317	-3.480	-2.866	-2.582
7	-1.223	-3.480	-2.873	-2.589
6	-1.296	-3.480	-2.880	-2.595
5	-1.269	-3.480	-2.887	-2.601
4	-1.344	-3.480	-2.893	-2.606
3	-1.349	-3.480	-2.899	-2.612
2	-1.310	-3.480	-2.905	-2.617
1	-0.897	-3.480	-2.910	-2.622

Opt Lag (Ng-Perron seq t) = 12 with RMSE 1139.738Min SC = 14.22714 at lag 2 with RMSE 1189.249 Min MAIC = 14.19142 at lag 2 with RMSE 1189.249

### . dfuller HP, trend lags(2)

Augmented Dickey-Fuller test for unit root

Number of obs =

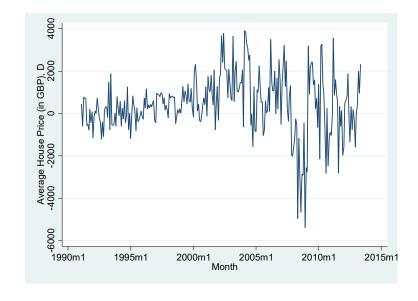
266

		Interpolated Dickey-Fuller					
	Test	1% Critical	5% Critical	10% Critical			
	Statistic	Value	Value	Value			
Z(t)	-1.674	-3.989	-3.429	-3.130			

MacKinnon approximate p-value for Z(t) = 0.7623

Applying first differences and repeating the ADF test

### . twoway line d.HP Month



#### . dfuller d.HP, lags(2)

Augmented Dickey-Fuller test for unit root Number of obs =

		Interpolated Dickey-Fuller					
	Test	1% Critical	5% Critical	10% Critical			
	Statistic	Value	Value	Value			
Z(t)	-5.344	-3.459	-2.879	-2.570			

265

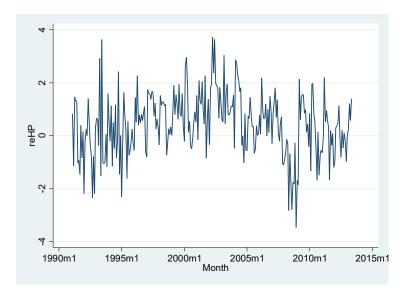
MacKinnon approximate p-value for Z(t) = 0.0000

Generating continuously compounded percentage returns (percentage changes) and repeating the ADF test  $\,$ 

# . generate reHP=100\*(ln(HP/L.HP))

(1 missing value generated)

#### . twoway line reHP Month



### . dfuller reHP, lags(2)

Augmented Dickey-Fuller test for unit root Number of obs = 265

		Interpolated Dickey-Fuller					
	Test	1% Critical	5% Critical	10% Critical			
	Statistic	Value	Value	Value			
Z(t)	 -5.771	-3.459	-2.879	-2.570			

MacKinnon approximate p-value for Z(t) = 0.0000

Constructing an ARMA model

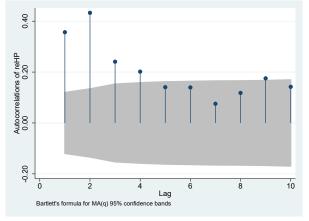
# . corrgram reHP, lags(10)

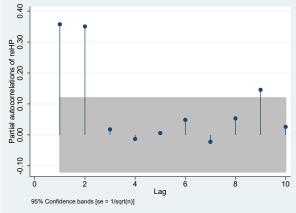
					-1 0	1 -1	0 1
LAG	AC	PAC	Q	Prob>Q	[Autocorrelat	ion] [Pa:	rtial Autocor]
1	0.3566	0.3575	34.468	0.0000			
2	0.4329	0.3504	85.444	0.0000			
3	0.2407	0.0179	101.27	0.0000	-		1

```
112.43 0.0000
117.86 0.0000
         0.2018 -0.0132
                 0.0058
5
         0.1406
6
         0.1396
                  0.0486
                            123.24 0.0000
7
         0.0755
                  -0.0225
                             124.82 0.0000
8
         0.1182
                  0.0533
                            128.71 0.0000
                            137.31 0.0000
142.97 0.0000
9
         0.1754
                   0.1457
10
         0.1421
                   0.0260
```

#### . ac reHP, lags(10)

#### . pac reHP, lags(10)





### . arima reHP, arima(0,0,0)

(setting optimization to BHHH)

Iteration 0: log likelihood = -427.80328
Iteration 1: log likelihood = -427.80328

ARIMA regression

| OPG
reHP | Coef. Std. Err. z P>|z| [95% Conf. Interval]

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

#### . estat ic

Akaike's information criterion and Bayesian information criterion

Model	Obs	ll(null)	ll(model)	df	AIC	BIC
	268		-427.8033	2	859.6066	866.7885

Note: N=Obs used in calculating BIC; see [R] BIC note.

// We continue to estimate all possible ARMA models manually. // // Alternatively, we can use the following program: //

```
. generate variable=reHP
(1 missing value generated)
. scalar pset=4
. scalar qset=4
. forvalues p=0/`=pset' {
  2. forvalues q=0/`=qset' {

 qui arima variable, arima(`p',0,`q')

    scalar aic_`p'_`q'=-2*e(11)+2*e(rank)
    scalar bic_`p'_`q'=-2*e(11)+e(rank)*ln(e(N))

  6. }
  7. }
. forvalues p=0/`=pset' {
  2. forvalues q=0/`=qset' {
  3. display "AIC(`p', `q') = "aic_`p'_`q'
  4. }
  5. }
AIC(0,0) = 859.60656
AIC(0,1) = 840.82476
AIC(0,2) = 804.35416
AIC(0,3) = 801.8014
AIC(0,4) = 800.01893
AIC(1,0) = 825.17453
AIC(1,1) = 802.16762
AIC(1,2) = 795.50644
AIC(1,3) = 794.0435
AIC(1,4) = 795.2866
AIC(2,0) = 792.09006
AIC(2,1) = 794.01043
AIC(2,2) = 795.95615
AIC(2,3) = 794.75456
AIC(2,4) = 796.10834
AIC(3,0) = 794.00444
AIC(3,1) = 793.38271
AIC(3,2) = 797.94062
AIC(3,3) = 796.4548
AIC(3,4) = 795.86876
AIC(4,0) = 795.95778
AIC(4,1) = 795.26527
AIC(4,2) = 791.60913
AIC(4,3) = 791.43349
AIC(4,4) = 795.39434
. forvalues p=0/`=pset' {
  2. forvalues q=0/`=qset' {
  3. display "SBIC(`p', `q') = "bic_`p'_`q'
  4. }
  5. }
SBIC(0,0) = 866.78854
SBIC(0,1) = 851.59773
SBIC(0,2) = 818.71811
SBIC(0,3) = 819.75633
SBIC(0,4) = 821.56485
SBIC(1,0) = 835.94749
SBIC(1,1) = 816.53157
SBIC(1,2) = 813.46137
SBIC(1,3) = 815.58942
SBIC(1,4) = 820.42351
SBIC(2,0) = 806.45401
SBIC(2,1) = 811.96537
SBIC(2,2) = 817.50207
```

```
SBIC(2,3) = 819.89146
SBIC(2,4) = 824.83624
SBIC(3,0) = 811.95937
SBIC(3,1) = 814.92863
SBIC(3,2) = 823.07753
SBIC(3,3) = 825.1827
SBIC(3,4) = 828.18765
SBIC(4,0) = 817.5037
SBIC(4,1) = 820.40218
SBIC(4,2) = 816.74604
SBIC(4,3) = 823.75237
SBIC(4,4) = 831.30421
// SBIC is consistent, but requires a large sample, which we seem to have (T = 268) //
. arima reHP, ar(1/2)
(setting optimization to BHHH)
Iteration 0: \log likelihood = -392.04832
Iteration 1: \log likelihood = -392.04526
Iteration 2: \log likelihood = -392.04505
Iteration 3: \log likelihood = -392.04503
ARIMA regression
                                            Number of obs = Wald chi2(2) = Prob > chi2 =
Sample: 1991m2 - 2013m5
                                                                 79.64
                                                                  0.0000
Log likelihood = -392.045
                                             Prob > chi2
______
                            OPG
      reHP | Coef. Std. Err. z P>|z| [95% Conf. Interval]
_cons |
                                                      .1301183
                                                                 .7347255
               .4324219 .1542394 2.80 0.005
ARMA
        ar |
        L1. | .2299696 .0508839 4.52 0.000 .130239 .3297002
L2. | .3508321 .0488959 7.18 0.000 .2549978 .4466663
    /sigma | 1.044089 .0415722 25.12 0.000 .9626087 1.125569
```

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

#### . arima reHP, arima(2,0,0)

(setting optimization to BHHH)

ARIMA regression

	reHP	   Coef.	OPG Std. Err.	Z	P> z	[95% Conf.	Interval]
reHP	_cons	.4324219	.1542394	2.80	0.005	.1301183	.7347255
ARMA	ar L1. L2.	.2299696 .3508321	.0508839	4.52 7.18	0.000	.130239 .2549978	.3297002
	/sigma	1.044089	.0415722	25.12	0.000	.9626087	1.125569

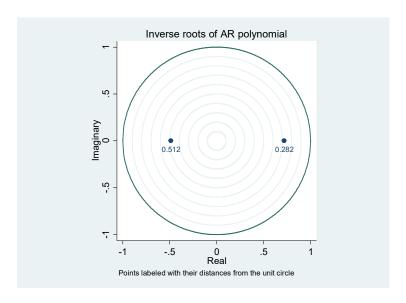
Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

#### . estat aroots, dlabel

Eigenvalue stability condition

	Eigenvalue	   +	Modulus	+   
	.7183533 .4883837	i I	.718353 .488384	

All the eigenvalues lie inside the unit circle. AR parameters satisfy stability condition.



#### . irf create AR2, set(myirf, replace) step(12)

(file myirf.irf created)
(file myirf.irf now active)
(file myirf.irf updated)

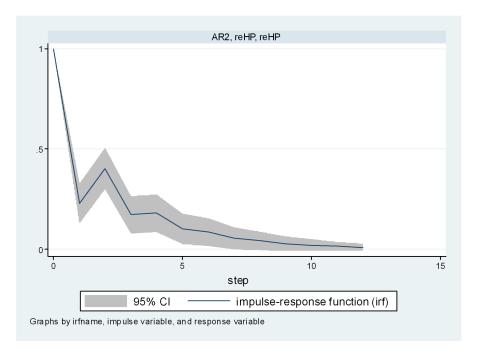
### . irf describe

Contains irf results from myirf.irf

irfname	model +	endogenous	variables 	and	order	(*)	_
AR2	arima	reHP					_

<sup>(\*)</sup> order is relevant only when model is var

#### . irf graph irf



#### . irf table irf

Results from AR2

1					
	step	   (1)   irf	(1) Lower	(1) Upper	
10		+   1	1	1	<sub> </sub>
1		.22997	.130239	.3297	
12		.403718	.301934	.505502	
3		.173524	.081296	.265751	
4		.181542	.088293	.274792	
5		.102627	.027752	.177502	
16		.087292	.018908	.155676	
7		.056079	.002009	.11015	
8		.043521	002335	.089378	
19		.029683	006318	.065684	
10	)	.022095	007103	.051293	
11		.015495	007249	.038239	
12		.011315	00664	.02927	

95% lower and upper bounds reported

(1) irfname = AR2, impulse = reHP, and response = reHP

Forecasting using an ARMA model

# . arima reHP if Month<=tm(2010m12), arima(2,0,0)

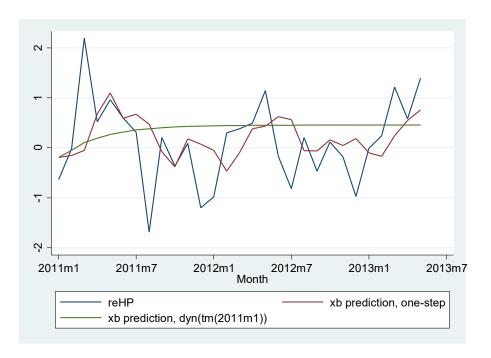
```
(setting optimization to BHHH)
Iteration 0: log likelihood = -355.03462
Iteration 1: log likelihood = -355.02874
Iteration 2: log likelihood = -355.02815
Iteration 3: log likelihood = -355.02809
Iteration 4: log likelihood = -355.02808
```

#### ARIMA regression

Sampl	.e: 1991r	m2 - 2010m12			Number		=	239
Log l	ikelihood	d = -355.0281			Wald ch	, ,	=	73.84
	reHP	   Coef.	OPG Std. Err.	z	P>   z	 [95% C	Conf.	Interval]
reHP	_cons	.4540309	.1702274	2.67	0.008	.12039	913	.7876704
ARMA	ar L1. L2.	.2255672	.0535307	4.21 7.04		.1206		.3304854
	/sigma	1.06785	.0460784	23.17	0.000	.97753	885	1.158163

Note: The test of the variance against zero is one sided, and the two-sided confidence interval is truncated at zero.

- . predict reHP\_fstatic, xb
- . predict reHP\_fdynamic, xb dynamic(tm(2011m1))
- . twoway (tsline reHP) (tsline reHP\_fstatic) (tsline reHP\_fdynamic) if Month>tm(2010m12)



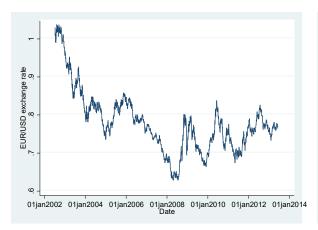
Example 2: We will examine whether there are lead-lag relationships between the returns to three exchange rates against the US dollar: the euro, the British pound and the Japanese yen. The data are daily and run from 7 July 2002 to 6 June 2013, giving a total of 3,988 observations. The data are contained in Stata data file currencies.dta. The programming code is given in Stata Do file currencies-commands-118.do.

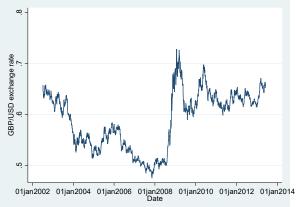
- a) Inspect the three time series visually. Do they look stationary? Perform the Dickey-Fuller GLS test (DFGLS) on each of the three time series. What do you find?
- b) Construct a set of continuously compounded percentage returns (percentage changes). Check for stationarity of the newly generated time series. What do you find?
- c) Now, we will focus on a vector autoregressive (VAR) model based on the returns to three exchange rates. Choose the optimal lag length based on multivariate information criteria and estimate the suggested VAR model. Interpret the Stata output, but keep in mind that because the construction of these models is usually not based on economic or financial theory, it is often best not to interpret the individual parameter estimates, but rather to examine the plausibility of the model as a whole, and to determine whether it describes the data well and produces accurate forecasts.
- d) After fitting a VAR model, one hypothesis of interest is that all the variables at a given lag are jointly zero. In order to test this hypothesis, perform the Wald lag-exclusion tests for different lag orders. What do you find?
- e) Check for stability of the chosen model by examining roots of the companion matrix (whether the eigenvalues lie inside the unit circle). What do you find?
- f) Check for residual autocorrelation of at least the first two lag orders with an appropriate Lagrange-multiplier (LM) test. What do you find?
- g) Run Granger causality Wald tests and figure out whether the "excluded" variable Granger-causes the "equation" variable. Do you find evidence of lead-lag interactions between the series? Which exchange rate reacts the fastest to market information?
- h) Obtain the impulse responses (IRs), orthogonalised impulse responses (OIRs) and forecasterror variance decompositions (FEVDs) for the estimated model. Interpret the results. Could these results have been expected based on previous results?
- i) Remember that the ordering of variables has an effect on the impulse responses and variance decompositions, and when, as in this case, theory does not suggest an obvious ordering of the series, sensitivity analysis should be undertaken. Reverse the ordering of variables and compare the new results with the initial ones. What do you find?
- j) Finally, employ the estimated VAR model in order to compute dynamic forecasts of the returns to three exchange rates for the following 20 days. Present the forecasts graphically, together with 95% confidence intervals.

# Computer printout of the results in Stata:

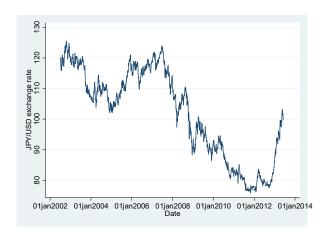
#### . twoway line EUR Date

#### . twoway line GBP Date





#### . twoway line JPY Date



Performing the Dickey-Fuller GLS test

# . dfgls EUR, maxlag(5)

DF-GLS for EUR

[lags]	DF-GLS tau Test Statistic	1% Critical Value	5% Critical Value	10% Critical Value
 5	-0.985	-3.480	-2.841	-2.554
4	-0.960	-3.480	-2.842	-2.554
3	-1.000	-3.480	-2.842	-2.554
2	-1.016	-3.480	-2.842	-2.554
1	-1.026	-3.480	-2.843	-2.555

Opt Lag (Ng-Perron seq t) = 1 with RMSE .0037332 Min SC = -11.17684 at lag 1 with RMSE .0037332Min MAIC = -11.17997 at lag 1 with RMSE .0037332

#### . dfgls GBP, maxlag(5)

DF-GLS for GBP

Number	οf	obs	=	3982
Number	$O_{\perp}$			3302

Number of obs = 3982

[lags]	DF-GLS tau	1% Critical	5% Critical	10% Critical
	Test Statistic	Value	Value	Value
5	-1.185	-3.480	-2.841	-2.554
4	-1.204	-3.480	-2.842	-2.554

3	-1.237	-3.480	-2.842	-2.554
2	-1.223	-3.480	-2.842	-2.554
1	-1.306	-3.480	-2.843	-2.555

Opt Lag (Ng-Perron seq t) = 2 with RMSE .002617 Min SC = -11.88519 at lag 2 with RMSE .002617 Min MAIC = -11.88968 at lag 2 with RMSE .002617

#### . dfgls JPY, maxlag(5)

DF-GLS for JPY

Number of obs = 3982

[lags]	DF-GLS tau Test Statistic	1% Critical Value	5% Critical Value	10% Critical Value
5	-1.625	-3.480	-2.841	-2.554
4	-1.644	-3.480	-2.842	-2.554
3	-1.771	-3.480	-2.842	-2.554
2	-1.788	-3.480	-2.842	-2.554
1	-1.764	-3.480	-2.843	-2.555

Opt Lag (Ng-Perron seq t) = 4 with RMSE .4787147 Min SC = -1.46636 at lag 1 with RMSE .47938 Min MAIC = -1.469925 at lag 4 with RMSE .4787147

Generating continuously compounded percentage returns (percentage changes) and repeating the Dickey-Fuller GLS test

#### . generate reur=100\*(ln(EUR/L.EUR))

(1 missing value generated)

# . generate rgbp=100\*(ln(GBP/L.GBP))

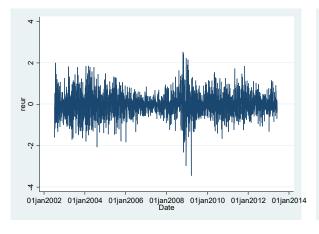
(1 missing value generated)

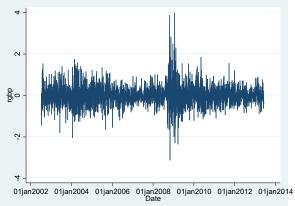
### . generate rjpy=100\*(ln(JPY/L.JPY))

(1 missing value generated)

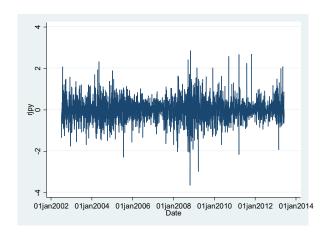
#### . twoway line reur Date

#### . twoway line rgbp Date





#### . twoway line rjpy Date



#### . dfgls reur, maxlag(5)

DF-GLS for reur

Number of obs = 3981

[lags]	DF-GLS tau Test Statistic	1% Critical Value	5% Critical Value	10% Critical Value
5	-24.487	-3.480	-2.841	-2.554
4	-26.944	-3.480	-2.842	-2.554
3	-30.815	-3.480	-2.842	-2.554
2	-34.634	-3.480	-2.842	-2.554
1	-41.158	-3.480	-2.843	-2.555

Opt Lag (Ng-Perron seq t) = 4 with RMSE .4722431 Min SC = -1.495368 at lag 1 with RMSE .4724769 Min MAIC = -.0863627 at lag 5 with RMSE .4722112

# . dfgls rgbp, maxlag(5)

DF-GLS for rgbp

Number of obs = 3981

[lags]	DF-GLS tau Test Statistic	1% Critical Value	5% Critical Value	10% Critical Value
5	-25.393	-3.480	-2.841	-2.554
4	-27.246	-3.480	-2.842	-2.554
3	-30.256	-3.480	-2.842	-2.554
2	-33.928	-3.480	-2.842	-2.554
1	-40.914	-3.480	-2.843	-2.555

Opt Lag (Ng-Perron seq t) = 1 with RMSE .4333588 Min SC = -1.668214 at lag 1 with RMSE .4333588 Min MAIC = -.3197095 at lag 2 with RMSE .4333307

### . dfgls rjpy, maxlag(5)

DF-GLS for rjpy

Number of obs = 3981

[lags]	DF-GLS tau Test Statistic	1% Critical Value	5% Critical Value	10% Critical Value
5	-24.263	-3.480	-2.841	-2.554
4	-27.220	-3.480	-2.842	-2.554
3	-30.605	-3.480	-2.842	-2.554
2	-33.610	-3.480	-2.842	-2.554
1	-40.031	-3.480	-2.843	-2.555

```
Opt Lag (Ng-Perron seq t) = 5 with RMSE .4701428 Min SC = -1.503243 at lag 1 with RMSE .4706203 Min MAIC = -.1675009 at lag 5 with RMSE .4701428
```

Estimating a VAR model

#### . varsoc reur rgbp rjpy, maxlag(10)

Endogenous: reur rgbp rjpy

Exogenous: \_cons

#### . var reur rgbp rjpy, lags(1/2)

Vector autoregression

Equation	Parms	RMSE	R-sq	chi2	P>chi2
reur rgbp rjpy	7 7 7	.470301 .430566 .466151		104.1884 219.6704 99.23658	0.0000 0.0000 0.0000

	 	Coef.	Std. Err.	z	P> z	[95% Conf.	Interval]
reur		 					
	reur						
	L1.	.2001552	.0226876	8.82	0.000	.1556883	.2446221
	L2.	0334134	.0225992	-1.48	0.139	077707	.0108802
	rgbp						
	L1.	0615658	.0240862	-2.56	0.011	1087738	0143578
	L2.	.024656	.0240581	1.02	0.305	0224971	.0718091
	rjpy						
	L1.	0201509	.0166431	-1.21	0.226	0527707	.0124689
	L2.	.002628	.0166676	0.16	0.875	0300398	.0352958
	_cons	0058355	.0074464	-0.78	0.433	0204301	.0087591

rgbp							
	reur						
	L1.	0427769	.0207708	-2.06	0.039	0834868	0020669
	L2.	.0567707	.0206898	2.74	0.006	.0162194	.097322
	rahn	 					
	rgbp L1.	ı I .2616429	.0220512	11.87	0.000	.2184234	.3048623
	L2.	10920986	.0220312	-4.18	0.000	1352678	0489294
	шг.	1	.0220255	4.10	0.000	.1332070	.04002
	rjpy	! 					
	L1.	0566386	.0152369	-3.72	0.000	0865024	0267747
	L2.	.0029643	.0152593	0.19	0.846	0269435	.0328721
	_cons	.0000454	.0068172	0.01	0.995	0133161	.0134069
		+ '					
rjpy	reur	 					
	L1.	.0241862	.0224874	1.08	0.282	0198883	.0682607
	L2.	0313338	.0223997	-1.40	0.162	0752365	.0125689
	ше.	1	.0223337	1.10	0.102	.0732303	.0123003
	rgbp	! 					
	Ĺ1.	0679786	.0238736	-2.85	0.004	1147701	0211872
	L2.	.0324034	.0238458	1.36	0.174	0143336	.0791404
	rjpy						
	L1.	.1508446	.0164962	9.14	0.000	.1185127	.1831766
	L2.	.0007184	.0165205	0.04	0.965	0316611	.0330979
	_cons	0036822	.0073807	-0.50	0.618	0181481	.0107836

VAR model diagnostics

# . varwle

Equation: reur

la	.g			Prob > c	
		104.0216	3	0.000 0.528	     

Equation: rgbp

+-						
	lag				Prob > chi2	
-		-+-				-
	1		214.893	3	0.000	
	2		17.60756	3	0.001	

Equation: rjpy

	lag				Prob > chi	
- 1						
	1		95.98737	3	0.000	- 1
	2		2.298777	3	0.513	
- 1						

Equation: All

· :	lag	1	chi2	df	Prob	> ch	i2
		+-					
1	1	-	592.6466	9	0 .	.000	

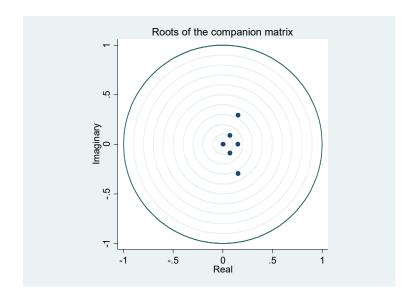
```
| 2 | 51.14752 9 0.000 |
```

### . varstable, graph

Eigenvalue stability condition

+-	Eige		Modulus	-+   -!		
	.1556644	+	.2942196i	į	.332861	
	.1531685	_			.153168	
	.07266967	+	.08855483i .08855483i		.114555 .114555	
	.00280605				.002806	

All the eigenvalues lie inside the unit circle.  $\ensuremath{\text{VAR}}$  satisfies stability condition.



### . varlmar

Lagrange-multiplier test

+						-+
	lag		chi2	df	Prob > chi2	
		-+-				-
	1	- 1	6.8459	9	0.65316	
-	2	- 1	9.7778	9	0.36877	
+						

HO: no autocorrelation at lag order

Performing Granger causality tests

## . vargranger

Granger causality Wald tests

į	Equation	Excluded		chi2	df Pi	rob > chi2	2   
	reur reur reur	rgbp   rjpy   ALL		6.6529 1.4668 7.9253	2 2 4	0.036 0.480 0.094	
 	   rgbp	reur	+- 	9.8352	2	0.007	 

	rgbp rgbp	211.	13.963 28.095	2 4	0.001	 
       	rjpy rjpy rjpy	reur   rgbp   ALL		2 2 2 4		

Impulse responses and variance decompositions for 20 time periods (days)

# . varbasic reur rgbp rjpy, lags(1/2) step(20) irf // Impulse response functions, IRFs //

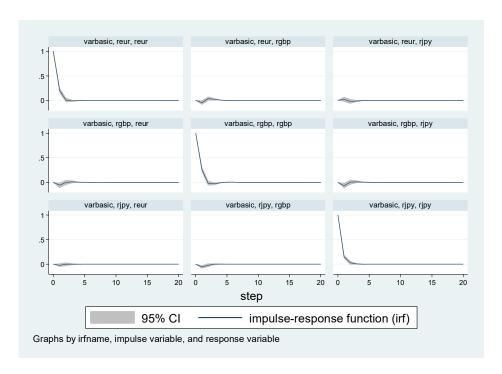
Vector autoregression

Sample:	10jul20	002	- 06jun2013	Number of obs	=	3 <b>,</b> 985
Log like	lihood =	=	-6043.54	AIC	=	3.043684
FPE	=	=	.0042115	HQIC	=	3.055437
Det(Sigma	a ml) =	=	.0041673	SBIC	=	3.076832

Equation	Parms	RMSE	R-sq	chi2	P>chi2
reur	7	.470301	0.0255	104.1884	0.0000
rgbp	7	.430566	0.0522	219.6704	0.0000
rjpy	7	.466151	0.0243	99.23658	0.0000

		Coef.	Std. Err.	Z	P> z	[95% Conf.	Interval]
reur							
	reur L1. L2.	.2001552  0334134	.0226876	8.82 -1.48	0.000 0.139	.1556883 077707	.2446221
	rgbp L1. L2.	0615658 .024656	.0240862	-2.56 1.02	0.011 0.305	1087738 0224971	0143578 .0718091
	rjpy L1. L2.	0201509 .002628	.0166431	-1.21 0.16	0.226 0.875	0527707 0300398	.0124689
	_cons	0058355	.0074464	-0.78	0.433	0204301	.0087591
rgbp							
	reur L1. L2.	0427769 .0567707	.0207708	-2.06 2.74	0.039	0834868 .0162194	0020669 .097322
	rgbp L1. L2.	.2616429  0920986	.0220512	11.87 -4.18	0.000	.2184234 1352678	.3048623 0489294
	rjpy L1. L2.	0566386 .0029643	.0152369	-3.72 0.19	0.000 0.846	0865024 0269435	0267747 .0328721
	_cons	.0000454	.0068172	0.01	0.995	0133161	.0134069
rjpy	reur L1.	.0241862	.0224874	1.08	0.282	0198883	.0682607
	L2.	0313338	.0223997	-1.40	0.162	0752365	.0125689

rgbp						
L1.	0679786	.0238736	-2.85	0.004	1147701	0211872
L2.	.0324034	.0238458	1.36	0.174	0143336	.0791404
rjpy						
L1.	.1508446	.0164962	9.14	0.000	.1185127	.1831766
L2.	.0007184	.0165205	0.04	0.965	0316611	.0330979
_cons	0036822	.0073807	-0.50	0.618	0181481	.0107836
<del>_</del>						



. varbasic reur rgbp rjpy, lags(1/2) step(20)
 // Orthogonalised impulse response functions, OIRFs //

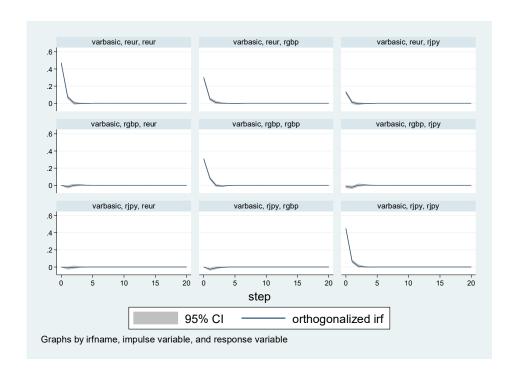
Vector autoregression

<pre>Sample: 10jul20 Log likelihood = FPE</pre>	-6043.54 .0042115	<u>l</u> 5		Number of AIC HQIC SBIC	of obs	= =	3,985 3.043684 3.055437 3.076832
Equation	Parms	RMSE	R-sq	chi2	P>chi2		3.070032

Equation	Parms	RMSE	R-sq	chi2	P>chi2
reur rgbp	7 7	.470301 .430566	0.0255 0.0522	104.1884 219.6704	0.0000
rjpy	7	.466151	0.0243	99.23658	0.0000

		 Coef.	Std. Err.		 P> z		Intervall
	ا +		5ta. EII.	Z 		. [93% CONI.	
reur	1						
	reur						
	L1.	.2001552	.0226876	8.82	0.000	.1556883	.2446221
	L2.	0334134	.0225992	-1.48	0.139	077707	.0108802
	- 1						
	rgbp						
	L1.	0615658	.0240862	-2.56	0.011	1087738	0143578
	L2.	.024656	.0240581	1.02	0.305	0224971	.0718091

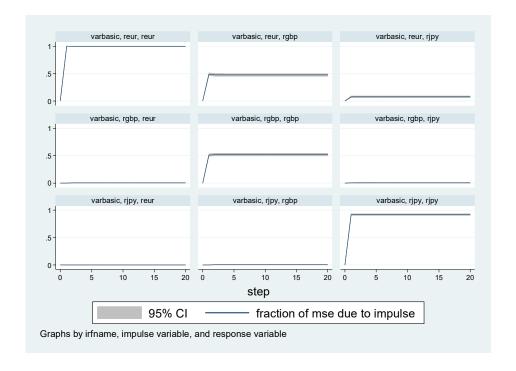
		l					
	rjpy L1.	  0201509	.0166431	-1.21	0.226	0527707	.0124689
	L2.	.002628	.0166676	0.16	0.875	0300398	.0352958
			0074464	0.70	0 400	0004001	0007501
	_cons	0058355 +	.0074464 	-0.78	0.433	0204301 	.0087591
rgbp							
	reur		0007700	0.06	0 000	0004060	0000660
	L1.	0427769	.0207708	-2.06	0.039	0834868	0020669
	L2.	.0567707	.0206898	2.74	0.006	.0162194	.097322
	rgbp	! 					
	L1.	.2616429	.0220512	11.87	0.000	.2184234	.3048623
	L2.	0920986	.0220255	-4.18	0.000	1352678	0489294
	rjpy	 					
	L1.	0566386	.0152369	-3.72	0.000	0865024	0267747
	L2.	.0029643	.0152593	0.19	0.846	0269435	.0328721
	cons	l .0000454	.0068172	0.01	0.995	0133161	.0134069
		+					
rjpy		[					
	reur	0041060	0004074	1 00	0 000	010000	0.600.607
	L1. L2.	.0241862	.0224874	1.08 -1.40	0.282 0.162	0198883 0752365	.0682607
	⊥∠.	0313338	.0223997	-1.40	0.162	0/52365	.0123689
	rgbp						
	L1.	0679786	.0238736	-2.85	0.004	1147701	0211872
	L2.	.0324034	.0238458	1.36	0.174	0143336	.0791404
	rjpy	 					
	L1.	.1508446	.0164962	9.14	0.000	.1185127	.1831766
	L2.	.0007184	.0165205	0.04	0.965	0316611	.0330979
	cons	  0036822	.0073807	-0.50	0.618	0181481	.0107836
			.00/000/	0.00	3.010	•0101101	•010,000



# . varbasic reur rgbp rjpy, lags(1/2) step(20) fevd // Forecast-error variance decompositions, FEVDs //

Vector autoregression

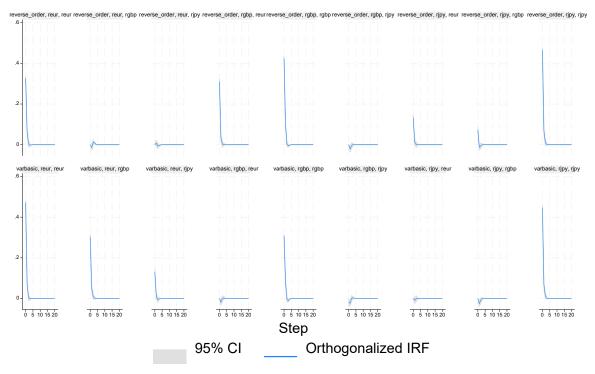
Vector autoreo	gression					
Sample: 10jul Log likelihood FPE Det(Sigma_ml)	12002 - 06jun 1 = -6043.54 = .0042115 = .0041673	4 5		Number of AIC HQIC SBIC	of obs	= 3,985 = 3.043684 = 3.055437 = 3.076832
Equation	Parms	RMSE	R-sq	chi2	P>chi2	
reur rgbp rjpy	7 7 7	.470301 .430566 .466151	0.0255 0.0522 0.0243	104.1884 219.6704 99.23658	0.0000 0.0000 0.0000	
	Coef.	Std. Err.	z	P>   z	[95% Cor	nf. Interval]
reur						
reur   L1.   L2.	.2001552 0334134	.0226876 .0225992	8.82 -1.48	0.000 0.139	.1556883 077707	
rgbp   L1.   L2.		.0240862	-2.56 1.02	0.011 0.305	1087738 0224971	
rjpy   L1.   L2.		.0166431	-1.21 0.16	0.226 0.875	0527707 0300398	
_cons	0058355	.0074464	-0.78	0.433	0204301	.0087591
rgbp	 					
reur   L1.   L2.	0427769	.0207708	-2.06 2.74	0.039	0834868 .0162194	
rgbp   L1.   L2.	.2616429 0920986	.0220512	11.87 -4.18	0.000	.2184234	
rjpy   L1.   L2.		.0152369	-3.72 0.19	0.000	0865024 0269435	
_cons	.0000454	.0068172	0.01	0.995	0133161	.0134069
rjpy reur	.0241862	.0224874	1 00	0 202	0100003	.0682607
L2.	0313338	.0223997	1.08 -1.40	0.282 0.162	0198883 0752365	
rgbp   L1.   L2.		.0238736	-2.85 1.36	0.004 0.174	1147701 0143336	
rjpy   L1.   L2.		.0164962	9.14	0.000 0.965	.1185127	
_cons	0036822	.0073807	-0.50	0.618	0181481	.0107836



Changing the ordering of variables

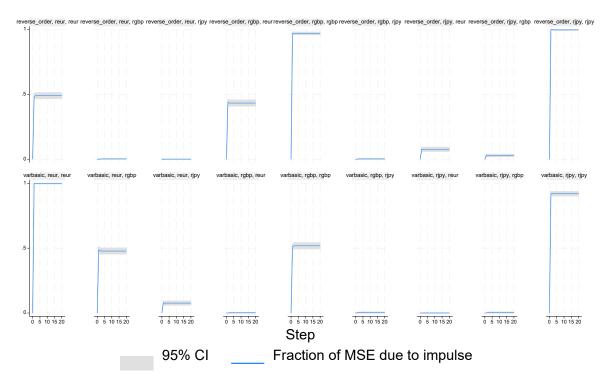
. irf create reverse\_order, replace step(20) order(rjpy rgbp reur)
irfname reverse\_order not found in \_varbasic.irf
(file \_varbasic.irf updated)

### . irf graph oirf, byopts(c(9))



Graphs by irfname, impulse variable, and response variable

### . irf graph fevd, byopts(c(9))



Graphs by irfname, impulse variable, and response variable

Forecasting based on VAR(2) for the following 20 time periods (days)

- . qui var reur rgbp rjpy, lags(1/2)
- . fcast compute fc\_, step(20)
   // generates level, lower bound, upper bound, and std error //
- . fcast graph fc\_reur fc\_rgbp fc\_rjpy

