

Tutorial, Part 15

Panel Data (static)

Example 1

Panel data: cross-section (i) & time-series (t)
 POOLED (data: x + set i -var t -var)

Balanced panel: all units (i) have the same number of time observations (t)

With e.g. OLS estimation we neglect the panel structure of the data \rightarrow bias, inconsistency
Not independent

\rightarrow Fixed-effects estimator, Random-effects estimator

a) 10 firms, 15 years

Pattern of an unbalanced panel: the units usually drop out.

Bottom row: x - if at least one firm observed in that year

Source of variation: - overall: $x_{it} - \bar{x}$

(i - one unit, "group")

- between: $\bar{x}_i - \bar{x}$

- within: $x_{it} - \bar{x}_i$ \leftarrow unit (group) i average

b) LSDV (Least Squares Dummy Variable) Estimator

$$Q_{it} = \beta_0 F_{it}^{\beta_1} M_{it}^{\beta_2} E_{it}^{\beta_3} L_{it}^{\beta_4} P_{it}^{\beta_5} \cdot M_{it} \quad / \log$$

$$\log Q_{it} = \log \beta_0 + \beta_1 \log F_{it} + \beta_2 \log M_{it} + \beta_3 \log E_{it} + \beta_4 \log L_{it} + \beta_5 \log P_{it} + M_{it} + c_i$$

FEE: $c_i \rightarrow \alpha_i$ (fixed) $\Rightarrow E(\alpha_i | x_i) = \rho(x_i) \neq 0$ (no general intercept)

REE: $c_i \rightarrow v_i$ (variable, random) $\Rightarrow E(v_i | x_i) = 0$
 $\rightarrow M_{it} + v_i = \omega_{it}$ (compound disturbance)

Dummy coefficients: differential intercept coefficients
 $\hat{\alpha}_i = \hat{\alpha}_1 + \text{coeff}_i \Rightarrow \text{fixed} = \text{constant over time}$

LSDVE \rightarrow add dummies \rightarrow OK for small n

Example: partial elasticities (interpretation!)

FEE (Fixed Effects Estimator) (the "within" regression)

$$n = 125$$

$$\text{df total} = n - 1 = 124$$

$$m_1 = \# \text{ "true" explanatory vars} = 5$$

$$m_2 = (n-1) - m_1 - \# \text{ dummies} = 110$$

$$\text{! corr}(\mu_{-i}, x_b) \neq 0$$

\Rightarrow same coeffs for main expl. vars as with LSDVE

$$\Rightarrow -\text{cons} = \bar{\alpha}_i \text{ (mean FE)}$$

$$\text{F-test: } H_0: \alpha_i = 0 \text{ } \forall i$$

df: # dummies

$$H_1: \alpha_i \neq 0 \text{ } \exists i$$

\Rightarrow "manual" calculation of FE($\hat{\alpha}_i$)

$$\hat{\alpha}_i = \bar{y}_{i\cdot} - x_{i\cdot}^T \hat{\beta}_{FE} \quad ; \quad i = 1, 2, \dots, n_{(10)}$$

\uparrow \uparrow
 vectors of means

But: FEE

- neglects between variation

- loss of df due to too many dummies

- elimination of time-invariant expl. vars.

(due to the "within" transformation $y_{it} - \bar{y}_i = \tilde{y}_{it}$)

\Rightarrow REE

(preferable, if it passes the Hausman test)

c) REE (Random Effects Estimator)

(weighted average of within and between estimator)

→ individual specific constant terms considered as randomly distributed ($\alpha_i \rightarrow v_i$)

- the disturbance term thus not extrinsic
- the disturbance term uncorrelated with regressors ($E(v_i|x_i) = 0$)
- the model includes intercept

- REE:
- more efficient
 - better prediction
 - less parameters
 - time-invariant expl. vars not neglected

d) But REE can be used only if it "passes" the Hausman test:

H_0 : FEE & REE consistent, FEE inefficient → use REE

H_1 : REE inconsistent → use FEE

Estimator (model) \ True process	RE (H_0 not rejected)	FE (H_0 rejected)
REE	<ul style="list-style-type: none"> • consistent • efficient 	INCONSISTENT! (efficiency doesn't matter)
FEE	<ul style="list-style-type: none"> • consistent • inefficient 	<ul style="list-style-type: none"> • consistent • possibly efficient

if not
sure:
use FEE

e)

reject H_0 → use FEE

Constant returns to scale (Σ elasticities = 1)

$$H_0: \beta_1 + \beta_2 + \beta_3 + \beta_4 + \beta_5 = 1$$

$$H_1: \beta_1 + \beta_2 + \beta_3 + \beta_4 + \beta_5 \neq 1$$

→ cannot reject H_0

Example 2

a) elasticities

$$H_0: \alpha_i = 0 \quad \forall i \Rightarrow \text{reject } H_0$$

b) year-dummies not sig. anymore as wage vars. explain changes through time

$$H_0: \text{all these coeffs.} = 0$$

$$H_1: \text{at least one coeff.} \neq 0 \Rightarrow \text{reject } H_0$$

c) First-difference OLS (3rd option for FE)

$$\Delta y_{it} = \Delta x_{it}^T \beta + \Delta u_{it}$$

→ usually less efficient (higher standard errors)

d) Hausman test

$$H_0: \text{REE}$$

$$H_1: \text{FEE} \Rightarrow \text{reject } H_0 \rightarrow \text{use FEE}$$

e) H_0 : homoscedasticity (stats: $xttest3$)

$$H_1: \text{heteroscedasticity} \Rightarrow \text{reject } H_0$$

H_0 : no first-order autocorrelation (stats: $xtserial$)

$$H_1: \text{first-order autocorrelation present} \Rightarrow \text{reject } H_0$$

Solution: $xtreg$... , fe (in case of autocorrelation)
 $xtreg$... , fe vce(robust) (in case of heteroscedasticity)

Tutorial, Part 16

Panel data (dynamic)

Example 1

- a) Static panel data model: long-run effects
- Dynamic panel data model: short-run effects
 - past values of the expl. vars affect the current value of the dep. var (through y_{t-1})
- ! Dynamic panel data estimators are highly sensitive to particular specification of the model and the instruments.
 → robustness checks → experimentation

- b) Arellano-Bond estimator: (GMM* estimator)
- "internal" and "external" instruments → more efficient (compared to Anderson-Hsiao ← only "internal" i.v.)
 ↳ substantial loss of eff.

* GMM - another approach to estimation:

finite-dimensional parameters but unknown shape of distribution → MLE not possible

Generalised Method of Moments (1st Moment: mean, 2nd: variance, 3rd: skewness...)
 ↳ "powers" of derivations

stata: xtabond 2

(gmmstyle → lags of dependant variable
 ivstyle → exogenous variables)

- 1* → inclusion of time dummies to prevent contemporaneous ("cross-individual") correlation

twostep → VCE matrix estimated by levels and differences

$\hat{\rho}$ - the coeff. with lagged dep. var. ($y_{i,t-1}$)
 $\hat{\rho} < 1 \rightarrow$ dynamic stability

Diagnostics:

- autocorrelation:
 - of order 1 \rightarrow present by construction
 - of higher orders \rightarrow shouldn't be present
 - joint validity of moment conditions:
 (= identifying restrictions = joint validity of instruments)
 - Sargan test: not robust, not weakened by many instruments
 - Hansen J-test: robust, weakened by many instruments \rightarrow nevertheless superior to Sargan
- $\rightarrow H_0$: instruments are valid

- c) improvement: instruments are lagged levels & lagged differences (lagged levels poor instruments for Δ)
 \rightarrow system of two equations:
- first differences equation
 - levels equation
- } system GMM

(state: remove nolevel option)

- d) Collapse the columns of all available lags as instruments into a single column (1 instrument for var & lag, not 1 for var & lag & time observation) (Z notation)
 \rightarrow recommended for small samples (avoids bias due to too many instruments)

e) 3 ways of dealing with endogeneity in the model:

- "within" transformation: subtracts the average of all values from the current observation
- first-difference (FD, Δ) transformation (GMM default): subtracts the previous value from the current observation (\rightarrow drops 1st observation, magnifies gaps in an unbalanced p.)
- forward orthogonal deviations (FOD) transformation: subtracts the average of all available future observations from the current value (\rightarrow drops last observation)

f) Coeffs from dynamic panel data models are estimates of short-run effects.

Long-run effect : $b_{j, \text{long}} = b_{j, \text{short}} \cdot \frac{1}{1 - \hat{\beta}}$
 (Stata: nlcom)

\rightarrow Interpretation: as in an ordinary regression function