

St.JOSEPH'S COLLEGE OF ENGINEERING, CHENNAI-119
St.JOSEPH'S INSTITUTE OF TECHNOLOGY, CHENNAI-119
MA6453 – PROBABILITY AND QUEUEING THEORY
UNIT I RANDOM VARIABLES
FORMULAE SHEET

S.NO	DISCRETE RANDOM VARIABLE:	CONTINUOUS RANDOM VARIABLE:
1.	Discrete Random Variable: Let X be a discrete random variables with values x_1, x_2, x_3, \dots . The function $P(X = x_i)$ is said to be probability mass function if i) $P(X = x_i) \geq 0, \forall i$ ii) $\sum_{i=1}^{\infty} P(X = x_i) = 1$	Continuous Random Variable: Let X be a continuous random variable. A function $f(x)$ is said to be probability density function of X if i) $f(x) \geq 0, \forall x$ ii) $\int_{-\infty}^{\infty} f(x) dx = 1$
2.	Distribution Function (or) Cumulative Distribution Function (C.D.F): $F(x) = P(X \leq x) = \sum_{x_i} P(X \leq x_i)$	Distribution Function (or) Cumulative Distribution Function (C.D.F): $F(x) = P(X \leq x) = \int_{-\infty}^x f(x) dx$ Properties of C.D.F: i) $F(-\infty) = \lim_{x \rightarrow -\infty} F(x) = 0$ and $F(\infty) = \lim_{x \rightarrow \infty} F(x) = 1$ ii) $P(a < X < b) = P(a \leq X \leq b) = P(a < X \leq b) = P(a \leq X < b) = F(b) - F(a)$ iii) $P(X \leq a) = F(a)$ vi) $P(X > a) = 1 - P(X \leq a) = 1 - F(a)$ v) $P(X = a) = 0$ vi) $f(x) = \frac{d}{dx} [F(x)]$
3.	Mean (or) Expectation of $X = E(X) = \bar{X} = \mu_1'$ $E(X) = \sum_{x_i} x_i P(X = x_i)$	Mean (or) Expectation of $X = E(X) = \bar{X} = \mu_1'$ $E(X) = \int_{-\infty}^{\infty} x f(x) dx$
4.	Variance of X: $Var(X) = E(X^2) - [E(X)]^2$ (or) $Var(X) = \mu_2' - (\mu_1')^2$	Variance of X: $Var(X) = E(X^2) - [E(X)]^2$ (or) $Var(X) = \mu_2' - (\mu_1')^2$

	Where $E(X^2) = \mu_2' = \sum_{x_i} x_i^2 P(X = x_i)$	Where $E(X^2) = \mu_2' = \int_{-\infty}^{\infty} x^2 f(x) dx$
5.	r^{th} Order raw Moment : $\mu_r' = E(X^r) = \sum_{x_i} x_i^r P(X = x_i)$	r^{th} Order raw Moment : $\mu_r' = E(X^r) = \int_{-\infty}^{\infty} x^r f(x) dx$
6.	r^{th} Central Moment : $\mu_r = E(X - \bar{X})^r = \sum_{x_i} (x_i - \bar{X})^r P(X = x_i)$	r^{th} Central Moment : $\mu_r = E[X - \bar{X}]^r = \int_{-\infty}^{\infty} (x - \bar{X})^r f(x) dx$
7.	Moment generating function: $M_X(t) = E(e^{tX}) = \sum_{x_i} e^{tx} P(X = x_i)$	Moment generating function: $M_X(t) = E(e^{tX}) = \int_{-\infty}^{\infty} e^{tx} f(x) dx$
8.	First four central moments [Relationship between raw moment and central moment] $\mu_1 = 0$ (always) $\mu_2 = \mu_2' - (\mu_1')^2$ $\mu_3 = \mu_3' - 3\mu_2'\mu_1' + 2(\mu_1')^3$ $\mu_4 = \mu_4' - 4\mu_3'\mu_1' + 6\mu_2'(\mu_1')^2 - 3(\mu_1')^4$ In general $\mu_n = \mu_n' - nC_1\mu_{n-1}'\mu_1' + nC_2\mu_{n-2}'(\mu_1')^2 - nC_3\mu_{n-3}'(\mu_1')^3 + \dots + (-1)^{n+1}(n-1)(\mu_1')^n$	
<u>PROPERTIES OF MEAN & VARIANCE</u> 1) $E(c) = c$ where c is any constant. 2) $E(aX \pm b) = aE(X) \pm b$ 3) $\text{var}(c) = 0$ where c is any constant. 4) $\text{var}(aX \pm b) = a^2 \text{var}(X)$		
<u>PROPERTIES OF MOMENT GENERATING FUNCTION:</u> 1) $E(X^r) = \mu_r' = \text{coefficient of } \frac{t^r}{r!} \text{ in the expansion of } M_X(t)$ 2) $E(X^r) = \mu_r' = \left[\frac{d^r}{dt^r} (M_X(t)) \right]_{t=0}$ 3) $M_{cX}(t) = M_X(Ct)$, where C being a constant. 4) $M_{X_1+X_2}(t) = M_{X_1}(t)M_{X_2}(t)$ if X_1 & X_2 are independent. 5) $M_{aX+b}(t) = e^{bt} M_X(at)$		

STANDARD DISTRIBUTIONS:					
Discret Distributions	Probability Mass Function	MGF { $M_x(t)$ }	MEAN { $E(X)$ }	VARIANCE { $Var(X)$ }	Condition for applying (or) Remarks
Binomial Distribution $X \sim B(n, p)$	$P(X = x) = {}^nC_x p^x q^{n-x}, x = 0, 1, 2, 3, \dots, n$ Where n – number of trials p – probability of Success q – probability of failures X – Number of Success (out of 'n' trials) and $p + q = 1$	$(q + pe^t)^n$	np	npq	i) The n trials are independent ii) n is small ($n < 30$)
Poisson Distribution $X \sim P(\lambda)$	$P(X = x) = \frac{e^{-\lambda} \lambda^x}{x!}, x = 0, 1, 2, 3, \dots$ Where n – number of trials p – probability of Success X – Number of Success (out of 'n' trials) and $p + q = 1$	$e^{\lambda(e^t - 1)}$	λ	λ	i) n is infinitely large (i.e) $n \rightarrow \infty$ ii) p is very small (i.e) $p \rightarrow 0$ iii) $\lambda = np$
Geometric Distribution $X \sim G(p)$	Form I $P(X = x) = q^{x-1} p, x = 1, 2, 3, \dots, n$ X – number of trials required to get a first success	$\frac{pe^t}{1 - qe^t}$	$\frac{1}{p}$	$\frac{q}{p^2}$	Memoryless property: If X is a random variable such that for all $s, t > 0$ then X is said to have memory less property $P(X > s + t / X > s) = P(X > t), \forall s \& t > 0$
	Form II $P(X = x) = q^x p, x = 0, 1, 2, 3, \dots$ X – number of failures before the first success	$\frac{p}{1 - qe^t}$	$\frac{q}{p}$	$\frac{q}{p^2}$	

Continuous Distributions	Probability density Function	MGF { $M_X(t)$ }	MEAN { $E(X)$ }	VARIANCE { $Var(X)$ }	
Uniform Distribution $X \sim U(a, b)$	$f(x) = \begin{cases} \frac{1}{b-a}, & a < x < b \\ 0, & \text{otherwise} \end{cases}$	$\frac{e^{bt} - e^{at}}{(b-a)t}, t \neq 0$	$\frac{a+b}{2}$	$\frac{(b-a)^2}{12}$	-
Exponential Distribution $X \sim e(\lambda)$	$f(x) = \begin{cases} \lambda e^{-\lambda x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$	$\frac{\lambda}{\lambda - t}$	$\frac{1}{\lambda}$	$\frac{1}{\lambda^2}$	Memoryless property: If X is a random variable with exponential distribution, then X lacks memory, in the sense that $P(X > s + t / X > s) = P(X > t), \forall s, t > 0$
Gamma Distribution $X \sim G(\lambda, \alpha)$	$f(x) = \begin{cases} \lambda e^{-\lambda x} \frac{(\lambda x)^{\alpha-1}}{\Gamma \alpha}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases}$ The parameters λ and α are positive	$\left[\frac{\lambda}{\lambda - t} \right]^\alpha, \lambda > t$	$\frac{\alpha}{\lambda}$	$\frac{\alpha^2}{\lambda^2}$	If $\alpha=1$ Gamma distribution becomes exponential distribution
Normal Distribution $X \sim N(\mu, \sigma^2)$	$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}, -\infty < x < \infty$	$e^{\mu t + \frac{t^2 \sigma^2}{2}}$	μ	σ^2	-