Bayesian reasoning

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Bayesian reasoning: problem set

The problem set is available here:

tiny.cc/BR_problems

Problem 1: basic Bayesian reasoning

Setting

The Dark Mark

- stays with 20% probability if the maker dies
- stays with 100% probability if the maker is still alive

The Dark Lord survived his attack on Harry Potter with the small chance of 1/101

Question

Given that Severus Snapes Dark Mark has not faded, find the probability of the Dark Lord being alive.

Problem 1: basic Bayesian reasoning

- $x \in \{0, 1\}$ the Dark Lord (1 means that he is alive)
- $y \in \{0, 1\}$ the Dark Mark (1 means that it is still visible)

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Setting: p(y=1 \mid x=1) = 1, p(y=1 \mid x=0) = 0.2, p(x=1) = 1/101 Question: p(x=1 \mid y=1) = ?
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Problem 1: basic Bayesian reasoning

- $x \in \{0, 1\}$ the Dark Lord (1 means that he is alive)
- $y \in \{0,1\}$ the Dark Mark (1 means that it is still visible)

Setting:
$$p(y=1 \mid x=1) = 1$$
, $p(y=1 \mid x=0) = 0.2$, $p(x=1) = 1/101$ Question: $p(x=1 \mid y=1) = ?$

$$p(x=1 \mid y=1) = \frac{p(y=1 \mid x=1)p(x=1)}{\sum_{j} p(y=1 \mid x=j)p(x=j)} = \frac{1 \cdot \frac{1}{101}}{1 \cdot \frac{1}{101} + \frac{1}{5} \cdot \frac{100}{101}} = \frac{1}{1 + \frac{100}{5}} = \frac{1}{21}$$

Problem 2: frequentist framework

Setting

- $X = \{x_1, \dots, x_N\}$ independent dice rolls
- $N_k = \sum_{n=1}^{N} \mathbb{I}(x_n = k)$ counts
- $p(X \mid \theta) = \prod_{k=1}^K \theta_k^{N_k}$ multinomial likelihood, $\theta \in S_K$

Question

Maximum likelihood estimate for $\theta_{ML} = \arg \max_{\theta \in S_K} \log p(X \mid \theta)$

Problem 2: frequentist framework

 θ is restricted to a simplex. To omit the inequality restrictions change parameterization to $\mu_k = \log \theta_k, \quad \mu_k \in \mathbb{R}$

The Lagrangian has the form:

$$\mathcal{L}(\mu, \lambda) = \log p(X \mid \exp \mu) - \lambda (\sum_{k=1}^{K} \exp \mu_k - 1) =$$

$$= \sum_{k=1}^{K} (N_k \mu_k - \lambda \exp \mu_k) + \lambda$$

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Differentiation:

$$0 = \frac{\partial \mathcal{L}(\boldsymbol{\mu}, \lambda)}{\partial \mu_k} = N_k - \lambda \exp \mu_k \Rightarrow \theta_k = \exp \mu_k = \frac{N_k}{\lambda}$$

$$0 = \frac{\partial \mathcal{L}(\boldsymbol{\mu}, \lambda)}{\partial \lambda} = -\sum_{k=1}^K \exp \mu_k + 1 \Rightarrow \lambda = \sum_{k=1}^K N_k$$

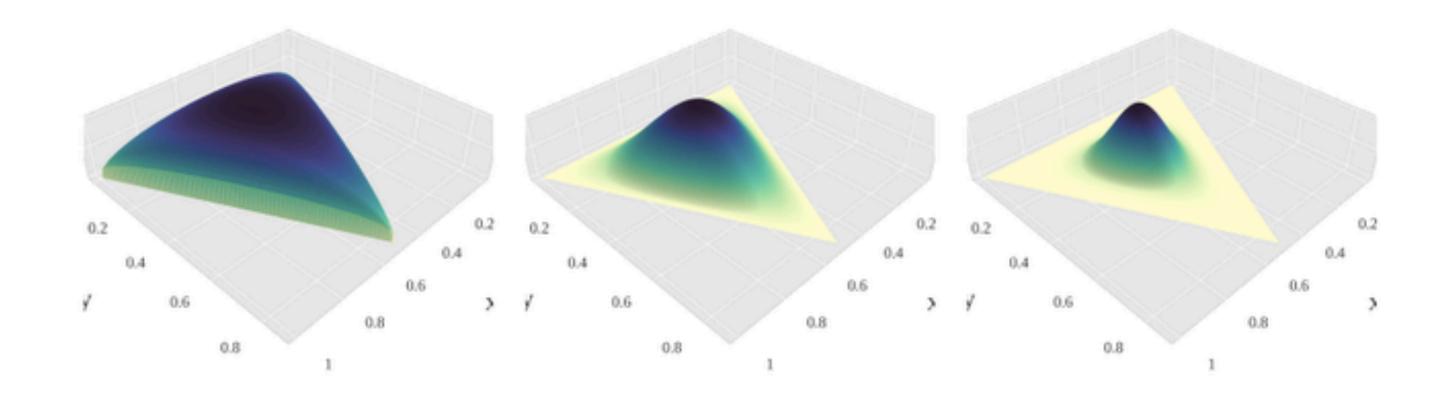
$$\theta_k = \frac{N_k}{\sum_{l=1}^K N_l}$$

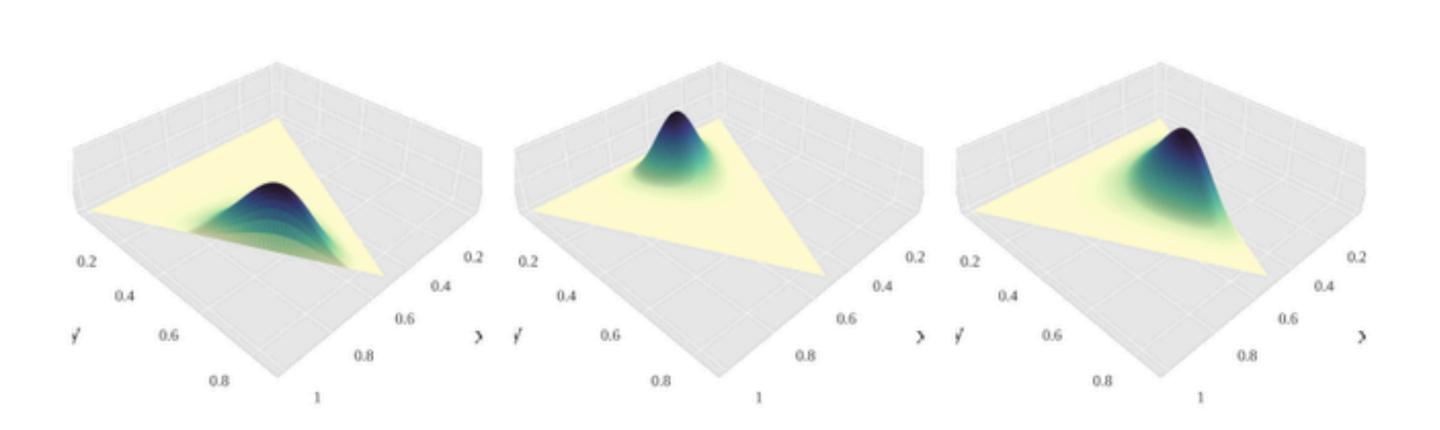
Setting

- $p(X \mid \theta) = \prod_{k=1}^K \theta_k^{N_k}$ multinomial likelihood, $\theta \in S_K$
- Dirichlet prior: $\operatorname{Dir}(\theta \mid \alpha) = \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K \theta_k^{\alpha_k 1}$

- Check that likelihood and prior are conjugate
- Compute the posterior $p(\theta|X,\alpha)$
- ullet Compare $\mathbb{E}_{p(heta|X,lpha)} heta$ and $heta_{ML}$
- Compute the predictive posterior $p(x_{N+1} = k | X, \alpha)$

Dirichlet distribution





Beta distribution is a special case of Dirichlet distribution:

$$\operatorname{Dir}(\theta \mid \alpha) \propto \prod_{k=1}^{K} \theta_k^{\alpha_k - 1}$$

$$Beta(\theta \mid a, b) \propto \theta^{a-1} (1 - \theta)^{b-1}$$

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Probabilistic model: $p(X, \theta) = p(X \mid \theta)p(\theta) = Dir(\theta \mid \alpha) \prod_{k=1}^{\infty} p(x_k \mid \theta)$

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Prior:
$$p(\theta) = \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K \theta_k^{\alpha_k - 1} = C \prod_{k=1}^K \theta_k^C$$

Here different constants are denoted with the same letter C for demonstration reasons.

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Posterior:
$$p(\theta \mid X) \propto p(X \mid \theta) p(\theta) = \prod_{k=1}^K \theta_k^{N_k} \cdot \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K \theta_k^{\alpha_k - 1} = \prod_{k=1}^K \theta_k^{N_k} \cdot \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K \theta_k^{\alpha_k - 1} = \prod_{k=1}^K \theta_k^{N_k} \cdot \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K \theta_k^{\alpha_k - 1} = \prod_{k=1}^K \theta_k^{N_k} \cdot \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K \theta_k^{\alpha_k - 1} = \prod_{k=1}^K \theta_k^{N_k} \cdot \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K \theta_k^{\alpha_k - 1} = \prod_{k=1}^K \theta_k^{N_k} \cdot \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K \theta_k^{\alpha_k - 1} = \prod_{k=1}^K \theta_k^{N_k} \cdot \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K \theta_k^{\alpha_k - 1} = \prod_{k=1}^K \theta_k^{N_k} \cdot \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K \theta_k^{\alpha_k - 1} = \prod_{k=1}^K \theta_k^{N_k} \cdot \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K \theta_k^{\alpha_k - 1} = \prod_{k=1}^K \theta_k^{N_k} \cdot \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^K$$

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$$= C \prod_{k=1}^{K} \theta_k^C$$

conjugate

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$$p(\theta \mid X) \propto p(X \mid \theta)p(\theta) = \prod_{k=1}^{K} \theta_k^{N_k} \cdot \frac{1}{B(\alpha_1, \dots, \alpha_K)} \prod_{k=1}^{K} \theta_k^{\alpha_k - 1} \propto \prod_{k=1}^{K} \theta_k^{N_k + \alpha_k - 1}$$

$$p(\theta \mid X) = Dir(\theta \mid \alpha'), \quad \alpha' = (\alpha_1 + N_1, \dots, \alpha_K + N_K)$$

Setting

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Maximum likelihood estimate:

$$\theta_k = \frac{N_k}{\sum_{l=1}^K N_l}$$

Expectation of the posterior:

$$\mathbb{E}_{p(\theta|X)}\theta_k = \frac{\alpha_k + N_k}{\sum_{l=1}^K \alpha_l + N_l}$$

Small $K \longrightarrow$ Bayesian estimate is mostly based on prior Large $K \longrightarrow$ Bayesian estimate is very similar to ML estimate

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$$p(x_{N+1} = k \mid X, \alpha) = \int_{S_K} p(x_{N+1} = k \mid \theta) p(\theta \mid X, \alpha) d\theta =$$

$$= \frac{\int_{S_K} \theta_k \prod_{k=1}^K \theta_k^{N_k + \alpha_k - 1} d\theta}{B(\alpha_1 + N_1, \dots, \alpha_K + N_K)} = \frac{B(\alpha_1 + N_1, \dots, \alpha_k + N_k + 1, \dots, \alpha_K + N_K)}{B(\alpha_1 + N_1, \dots, \alpha_k + N_k, \dots, \alpha_K + N_K)} =$$

$$= \frac{\Gamma(\alpha_1 + N_1) \dots \Gamma(\alpha_k + N_k + 1) \dots \Gamma(\alpha_K + N_K)}{\Gamma(\alpha_1 + N_1) \dots \Gamma(\alpha_k + N_k) \dots \Gamma(\alpha_K + N_K)} \cdot \frac{\Gamma(\sum_l (\alpha_l + N_l))}{\Gamma(\sum_l (\alpha_l + N_l) + 1)} =$$

$$= \frac{\alpha_k + N_k}{\sum_k \alpha_k + N_k}$$