

POLS 904 Final Project

Monte Carlo Simulation on Causal Forest

Jiacheng He

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Introduction

In social science, researchers might be interested to estimate the effect of a binary treatment (either treated or not treated). In experimental setting, individuals are randomly assigned into a control group and a treated group. Formally, denote y_i as the outcome variable, W_i as the treatment assignment variable ($W_i = 1$ if in the treated group, $W_i = 0$ if in the control group), and X_i as a set of observed covariates (e.g. age, gender, race, education, etc). Then the researcher can estimate the classical linear model:

$$Y_i = \tau W_i + X_i \beta + \epsilon_i \quad (1)$$

Here $\tau = E[Y_i|W_i = 1] - E[Y_i|W_i = 0]$ is interpreted as the average treatment effect (ATE) across all individuals. X_i is included into the regression to make sure of unconfoundedness and to reduce the variance of the estimator $\hat{\tau}$.

But sometimes researchers might want to go beyond ATE, and try to further estimate heterogeneous treatment effect and identify the subgroup of the population who will benefit the most (or least) from the treatment. One approach is to estimate the conditional average treatment effect (CATE) $\tau(X_i) = E[Y_i|X_i, W_i = 1] - E[Y_i|X_i, W_i = 0]$, that is, express the treatment effect τ as a function of the observed covariates.

Causal forest developed by Wager and Athey (2017) aims to algorithmically search for the covariate space, identify the subspace where heterogeneity exists, and estimate the CATE in these subspaces. It is very similar to the popular random forest method. Wager and Athey (2017) also derived asymptotic distribution of the causal forest estimator so that statistical inference and hypothesis test become feasible when adopting this forest based method.

In this project, I run Monte Carlo simulation on the causal forest to examine its finite sample performance, such as the mean squared error (MSE) and the confidence interval coverage rate.

Model Framework

Consider a simple additive model. Define $\tau(X_i) = E[Y_i|X_i, W_i = 1] - E[Y_i|X_i, W_i = 0]$ We will have such relationship:

$$E[Y_i|X_i, W_i = 1] = E[Y_i|X_i, W_i = 0] + W_i \cdot \tau(X_i)$$

With some derivation, we will have such relationship:

$$Y_i = m(X_i) + \frac{W_i}{2}\tau(X_i) + \frac{1 - W_i}{2}\tau(X_i) + \epsilon_i$$

where

$$\begin{aligned} \tau(X_i) &= E[Y_i|X_i, W_i = 1] - E[Y_i|X_i, W_i = 0] \\ m(X_i) &= E[Y_i|X_i] \\ e(X_i) &= E[W_i|X_i] \end{aligned}$$

Both $\tau(\cdot)$, $m(\cdot)$, and $e(\cdot)$ are nonparametric functions of the observed covariates X_i .

There are several challenges in nonparametrically estimate the CATE function $\tau(X_i)$. First, in real world application, we never observe the true individual treatment effect τ_i . At each moment, an individual is either in the treated status or in the non-treated status, so we never know what would have happened to the individual if the individual would have shifted his/her status. This is the fundamental problem in causal inference. As a result of the absence of the true τ_i , we can not perform cross validation, which is the routine in predictive machine learning.

Second, the existence of non-constant $m(\cdot)$ and $e(\cdot)$ will tend to confound our estimation, as I showed in the final presentation.

Causal Forest

The estimation algorithm of the causal forest is very closed to the random forest. There are two major divergence:

1. When growing each tree in the causal forest, we place the split at the point \tilde{x}_i , which maximizes the difference of $\hat{E}[Y_i|X_i = x_i, W_i = 1] - \hat{E}[Y_i|X_i = x_i, W_i = 0]$ ($\hat{\tau}$) across the two sides of \tilde{x}_i . While in the case of random forest we place the split based on $\hat{E}[Y_i|X_i = x_i]$ (\hat{y}).
2. When growing each tree, we use half of the training sample to do Step 1 above (placing split, identifying heterogeneity covariate subspace), and use the other half of the training sample to calculate the $\hat{\tau}$ (estimation of the CATE in that subspace). Wager and Athey (2017) refer to this criterion as “honest splitting”.

Also, similar to random forest, in causal forest algorithm it is not necessary to implement regularization or pruning.

Simulation Setup

In the Monte Carlo simulation experiment, I am interested to see how the algorithm performs as sample size and number of covariate change, under two different scenarios: 1. constant treatment effect; 2. heterogeneous treatment effect. I set up two data generating processes (DGP) as:

DGP1 (constant τ)

$$\begin{aligned}\tau(X_i) &= 0 \\ e(X_i) &= (1 + f_{beta}^{2,4}(X_{1i}))/4 \\ m(X_i) &= 2X_{1i} - 1\end{aligned}$$

where $f_{beta}^{2,4}(\cdot)$ is the density function of Beta distribution with shape parameters 2 and 4.

DGP2 (heterogeneous τ)

$$\begin{aligned}\tau(X_i) &= 1 + \frac{1}{(1 + e^{-20(X_{1i}-1/3)})(1 + e^{-20(X_{2i}-1/3)})} \\ e(X_i) &= 0.5 \\ m(X_i) &= 0\end{aligned}$$

Training causal forests also requires setting up tuning parameters, the same as when we train random forests. In this project, I also try to vary different tuning parameters and evaluate the performance of these trained models. The tuning parameters I try are as follows: (1) sample fraction used in growing each tree; (2) covariates used in growing each tree; (3) Number of trees to build the forest; (4) Minimum number of observations in each terminal leaf; (5) Regularization parameter λ .

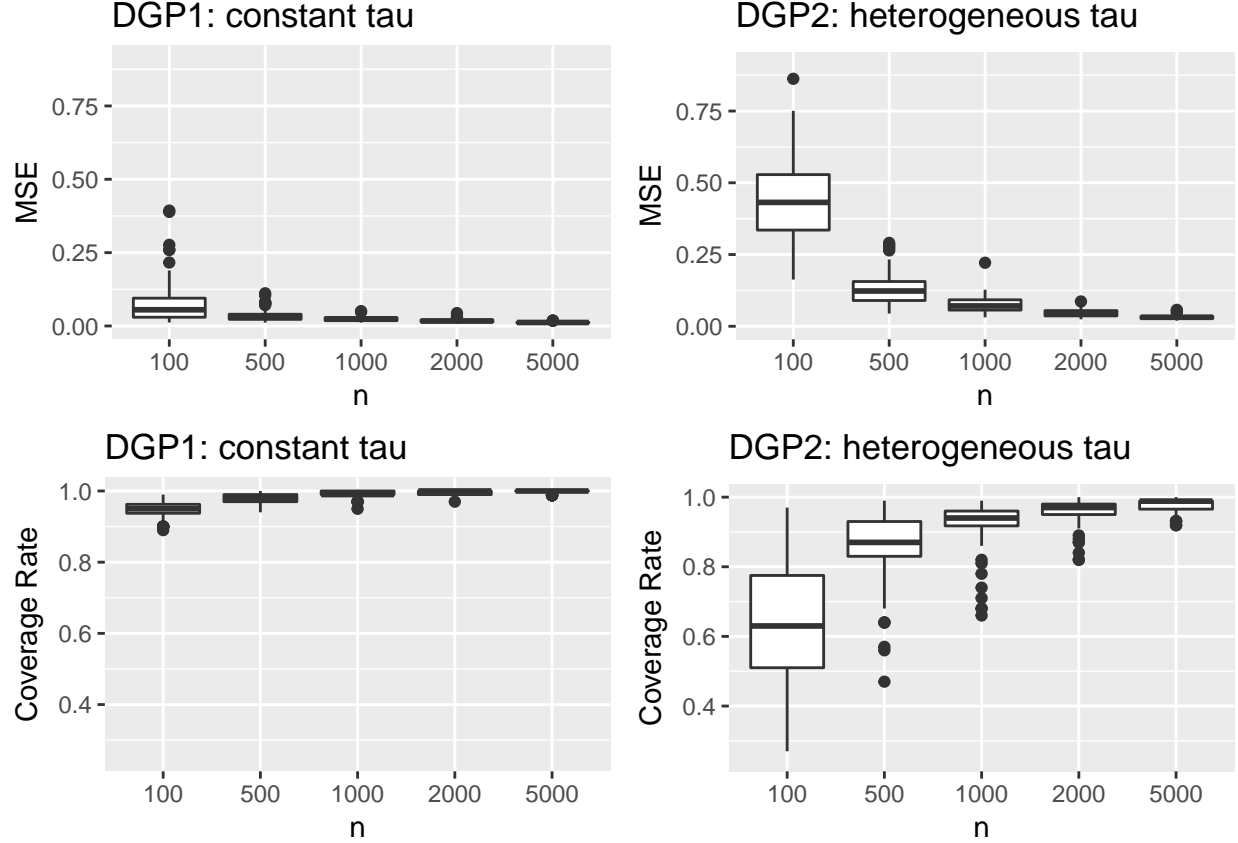


Figure 1: MSE and Coverage Rate with Different Sample Size n

Please keep in mind that cross validation is feasible in treatment effect estimation. So in practice there is no general guidance to select these tuning parameters in the training stage. Evaluation of choices of tuning parameters is only possible when we assume a data generating process and hence know the true τ_i in Monte Carlo simulation. And we can only implement the evaluation in the test set.

I draw $X_i \sim U(0, 1)^d$, $W_i \sim \text{binom}(1, e(X_i))$, $\epsilon_i \sim N(0, 1)$. (d is the number of covariates). Then I train the causal forest model on a training set, and evaluate the trained model on a test set with 100 data points. For each senario, I replicate it for 100 times. For each replication I generate a new training set, while the test set is invariant for all replications. Then I plot the box plot of the MSE and 95% confidence interval coverage rate.

Result

Sample Size and Number of Covariates

First, I will look at how the performance of the causal forest respond to the change of sample size n and number of covariates d :

1. Fix $d = 10$, try $n = 100, 500, 1000, 2000, 5000$;
2. Fix $n = 1000$, try $d = 2, 4, 10, 20, 40$;

Figure 1 displays the MSE and 95% confidence interval coverage rate. Under the data generating process of constant treatment effect, the MSE decays very fast as sample size n increases. Under the DGP of

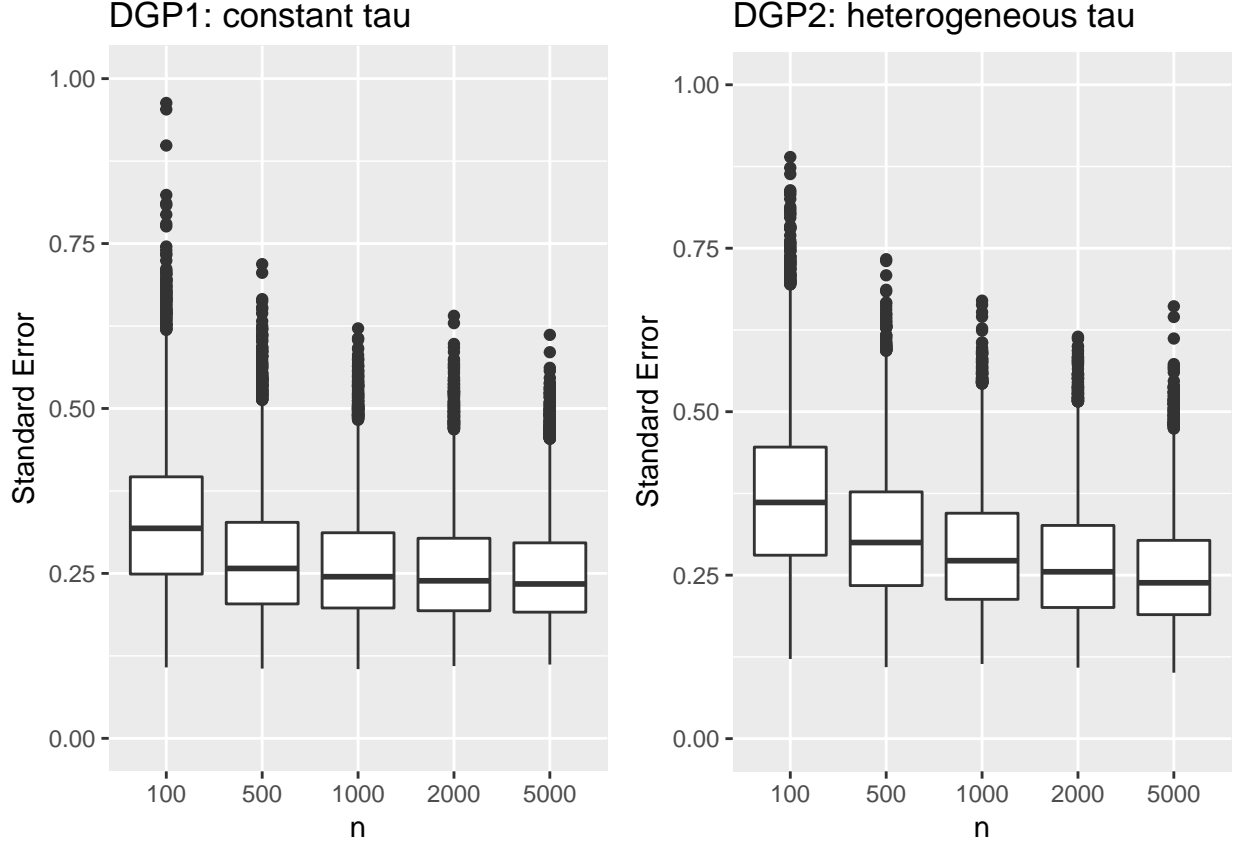


Figure 2: Standard Errors with Different Sample Size n

heterogeneous treatment effect, the MSE is slightly larger and it decays slightly slower. When n is as large as 5000, we achieve considerably small MSE.

Surprisingly, the confidence interval performs very well for the DGP of constant τ . Even in small sample $n = 100$, the simulation coverage rate achieves about 95%. What is interesting is that the confidence interval seems “over-accurate” for DGP1. In the case of $n = 5000$, we obtain almost 100% accuracy that the 95% confidence interval will always cover the true τ .

As for the DGP2, we need larger sample for the confidence coverage to converge. When $n = 100$, in median the 95% confidence interval successfully covers the true τ for only about 60% of the time. When $n = 1000$, the median coverage rate achieves 95%. However, when $n = 5000$, we have the “over-accurate” issue again.

In Figure 2 I plot the standard errors in the simulation. We can see that the standard errors are pretty stable and have not converged even when $n = 5000$. Theoretically, as sample size increases, we should have smaller standard error for our estimator. But it is clearly not the case here. Figure 1 has told us that the MSE is shrinking and basically converges when $n = 5000$. In other words, the point estimates are very close to the true value when $n = 5000$. Therefore, I conjecture that the standard errors are “too large” in large sample, and hence the 95% confidence intervals are “too wide”, which might lead us to the situation that we are less likely to reject than we should in hypothesis testing.

In both data generating processes, only the first two covariates X_1, X_2 contribute to the $\tau(\cdot)$ function. Therefore, adding extra covariates is purely adding “noise” to the causal forest algorithm. We should expect the variance of the heterogenous effect estimates increases as d increases.

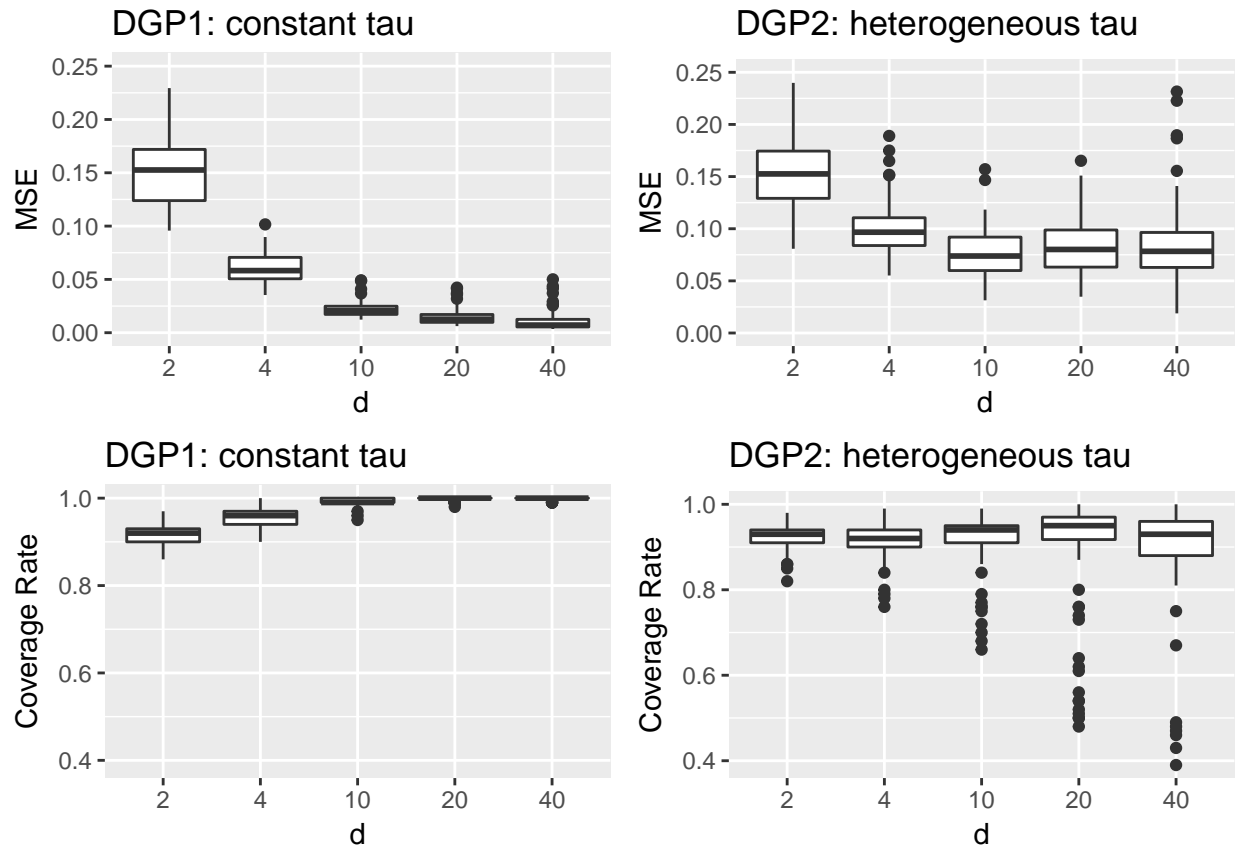


Figure 3: MSE and Coverage Rate with Different Numbers of Covariates d

Tuning Parameters

I try varying five tuning parameters, one at a time. I use DGP2 and fix $n = 1000, d = 10$

1. Sample fraction used in each tree training; (default 0.5)
2. Covariates used in each tree training; (default $\frac{2}{3}d$)
3. Number of trees; (default 2000)
4. Minimum # observations in each terminal node; (default NULL)
5. Regularization parameter λ ; (default 0)
6. Try sample fraction $s = 0.1, 0.2, 0.3, 0.4, 0.5$

s	MSE	coverage
0.1	0.242	0.535
0.2	0.123	0.87
0.3	0.082	0.92
0.4	0.073	0.94
0.5	0.075	0.94

2. Try # covariates in each tree training $t = 4, 5, 6, 7, 8$

t	MSE	coverage
4	0.102	0.865
5	0.082	0.92
6	0.079	0.92
7	0.07	0.94
8	0.074	0.94

3. Try # trees $b = 500, 1000, 2000, 4000, 6000$

b	MSE	coverage
500	0.088	0.97
1000	0.079	0.96
2000	0.076	0.95
4000	0.077	0.94
6000	0.075	0.915

4. Try minimum node size = 0, 10, 20, 40, 80

size	MSE	coverage
0	0.068	0.94
10	0.066	0.89
20	0.066	0.9
40	0.063	0.915
80	0.076	0.905

5. Try $\lambda = 0.1, 1, 5, 10, 100$

lambda	MSE	coverage
0	0.067	0.95
0.1	0.069	0.94
1	0.066	0.94
5	0.076	0.925
10	0.079	0.92