

Assessing Latent Level Associations Between PTSD and Dissociative Factors: Is Depersonalization and Derealization Related to PTSD Factors More So than Alternative Dissociative Factors?

Cherie Armour · Ateka A. Contractor ·
Patrick A. Palmieri · Jon. D. Elhai

Received: 1 April 2014 / Accepted: 7 May 2014 / Published online: 30 May 2014
© Springer Science+Business Media New York 2014

Abstract Posttraumatic stress disorder (PTSD) criteria in DSM-5 included a dissociative diagnostic subtype characterized by a depersonalization item and a derealization item. Researchers have queried whether this was too restrictive, as alternative dissociative symptomatology may also be characteristic of the subtype. The current study utilized data from 318 Northern Irish students, of which 165 were trauma exposed. Participants were assessed for PTSD symptomatology based on DSM-5 criteria via a modified version of the PTSD Symptom Scale-Self-Report (PSS-5) and dissociative experiences via the Dissociative Experiences Scale (DES). Confirmatory factor analysis of PTSD and DES models revealed an optimal four-factor DSM-5 PTSD model including reexperiencing, avoidance, negative alterations in mood and cognitions, and alterations in hyperarousal and reactivity factors, and an optimal three-factor DES model including absorption, amnesia, and depersonalization/derealization factors. When comparing the correlations between depersonalization/derealization and the four PTSD factors, significant Wald tests of parameter constraints revealed that depersonalization/derealization is more related to alterations in arousal and reactivity ($r=.432$)

compared to avoidance ($r=.289$), $\chi^2(1, N=165)=8.352$, $p=.004$. We discuss whether the mechanism for comorbid PTSD and dissociation may be related to PTSD's arousal factor.

Keywords Confirmatory factor analysis · Posttraumatic stress disorder · Dissociation · *DSM-5* · Dissociative Experiences Scale

Introduction

The fifth edition of the Diagnostic and Statistical Manual of Mental Disorders (DSM-5) released in May 2013 introduced a number of revisions to the Posttraumatic stress disorder (PTSD) diagnostic criteria. Revisions included an amendment to the definition of a traumatic exposure, revisions to existing symptom descriptions, the addition of symptoms, the grouping of symptoms into four rather than three symptom clusters, and the inclusion of a dissociative PTSD subtype characterized by the endorsement of depersonalization and derealization symptoms (cf. Friedman, Resick, Bryant, Strain, Horowitz et al., 2011). The latter two revisions are of central focus in the current study.

Trauma exposure, PTSD, and dissociation have long been associated. The theory which relates traumatic experience to the development of dissociative psychopathology dates back to the work of Janet, (1907) and Breuer and Freud, (1955). More recently, it has been suggested that dissociative symptomatology, both peritrauma and posttrauma, functions to minimize the negative emotional reactions to traumatic exposure (Putman, 1989; Spiegel, 1991; van der Kolk and van der Hart, 1989). Subsequent empirical research has shown that dissociative symptomatology develops in response to

C. Armour (✉)
School of Psychology, Coleraine Campus, University of Ulster,
Northern Ireland, UK
e-mail: armour.cherie@gmail.com

A. A. Contractor
Department of Psychology, University of Toledo, Toledo, OH 43606,
USA

P. A. Palmieri
Center for the Treatment and Study of Traumatic Stress, Department
of Psychiatry, Summa Health System, Akron, OH, USA

J. D. Elhai
Department of Psychology, and Department of Psychiatry, University
of Toledo, Toledo, OH 43606, USA

traumatic experience (cf. Gershuny and Thayer, 1999) and that peritraumatic dissociation is one of the strongest risk factors for the subsequent development of PTSD (Ozer, Best, Lipsey, and Weiss, 2003). Indeed, peritraumatic dissociation has been highlighted as a salient predictor for both posttraumatic symptoms and PTSD in two meta-analyses: one of studies conducted between 1980 and 2000 (Ozer et al., 2003) and the other of studies conducted between 1995 and 2004 (Breh and Seidler, 2007). Early traumatic responding in the form of acute stress disorder (ASD) requires participants to endorse three of four (numbing/detachment, reduced awareness of surroundings, derealization, and depersonalization) dissociative symptomatology. Additionally, PTSD symptoms in the form of flashbacks and amnesia are considered dissociative.

To date, much of the work assessing the role of dissociation in PTSD's symptomatology has focused on peritraumatic dissociative experiences. This differs from the dissociative experiences outlined in the DSM-5 for the dissociative PTSD subtype which is characterized by persistent dissociative experiences. Peritraumatic and persistent dissociative experiences have been shown to be correlated at .25 (Tichenor, Marmar, Weiss, Metzler, and Ronfeldt, 1996), and therefore, one must be careful when generalizing findings based on studies of peritraumatic dissociation to matters of persistent dissociation.

The DSM-5 dissociative PTSD subtype was proposed and has subsequently been supported based on both clinical and neurobiological evidence (Armour, Karstoft, and Richardson, 2014a, b; Cloitre, Petkova, Wang, and Lassell, 2012; Ginzburg, Koopman, Butler, et al., 2006; Lanius, Vermetten, Loewenstein, et al., 2010, 2012; Resick, Suvak, Johnides, Mitchell, Iverson, 2012; Spiegel, Loewenstein, Lewis-Fernandez, Sar, Simeon, Vermetten et al., 2011; Steuwe, Lanius, and Frewen, 2012; Wolf, Miller et al., 2012). Researchers initially proposed that the dissociative PTSD subtype be characterized by the presence of only depersonalization and derealization symptoms of dissociation (cf. Lanius et al., 2010). However, one study conducted by Armour, Elklit, Lauterbach, and Elhai (2014b) highlighted that an additional indicator of dissociation (reduced awareness) was the dissociative item most associated with overall PTSD ($r=.57$), in addition to being the item most likely to be endorsed by members of the dissociative PTSD class compared to the depersonalization and derealization items. This raised questions as to whether a wider range of dissociative symptomatology may be indicative of the subtype.

PTSD's Factor Structure (DSM-III to DSM-IV-TR)

The underlying dimensionality of PTSD has been contended within the literature. Since PTSD's initial inclusion, the DSM has categorized PTSD symptoms into three distinct symptom

groupings: reexperiencing, avoidance/numbing, and hyperarousal (APA 1980, 1987, 2000). Generally speaking, each of the symptom groupings is proposed to have distinct yet somewhat related functions. Reexperiencing symptoms are thought to occur as an attempt at integrating trauma memories with preexisting knowledge. Avoidance/numbing symptoms attempt to avoid trauma-related stimuli with the aim of avoiding the distressing reexperiencing symptoms. Avoidance of trauma-related stimuli is assisted by hyperarousal symptoms, as hyperarousal increases an individual's vigilance for exposure to such stimuli (reviewed in Bewin and Holmes, 2003).

Despite a comprehensive theoretical grounding for the tripartite model of PTSD, empirical support largely lies with alternative model conceptualizations (reviewed in Yufik and Simms, 2010; Elhai and Palmieri, 2011). Indeed, plentiful support has been found for the emotional numbing model (King, Leskin, King, and Weathers, 1998), the dysphoria model (Simms, Watson, and Doebbeling, 2002), and a more recently proposed five-factor model, the dysphoric arousal model (Elhai, Biehn, Armour, Klopfer, Frueh et al., 2011a, b). When comparing the emotional numbing model and the dysphoria model, the latter was deemed optimal in a meta-analysis of 40 PTSD studies ($N=14,827$ participants) (Yufik and Simms, 2010). Of note, the dysphoric arousal model (Elhai et al., 2011a, b) has been the recipient of support from the most recent of factor analytic studies (cf. Armour, Ghazali, and Elklit, 2013a, b, c; Harpaz-Rotem, Tsai, Pietrzak, and Hoff, 2014; Reddy et al. 2013; Semage et al., 2013; Wang, Cao, Wang, Zhanga, and Li, 2012, 2013).

PTSD's Factor Structure (DSM-5)

The DSM-5 (APA, 2013), published in May of 2013, now includes 20 rather than 17 PTSD symptoms, with some of the original 17 symptom descriptions being revised. Most pertinent to the focus of the current study, the 20 symptoms are now categorized across four rather than three symptom groupings: (1) reexperiencing, (2) avoidance, (3) negative alterations in mood and cognitions (predominately numbing symptoms with the addition of three symptoms: (a) persistent negative expectations about oneself, others, or the world; (b) persistent distorted blame of self or others about the cause or consequences of the trauma; and (c) pervasive negative emotional states), and (4) alterations in arousal and reactivity (predominately hyperarousal symptoms with irritability/anger and the addition of a symptom of reckless behavior) (cf. Friedman et al., 2011). This model conceptualization is closest to that of the DSM-IV-TR's emotional numbing model noted above. An alternative DSM-5 dysphoria model has been put forth to represent the DSM-IV-TR dysphoria model; this model includes an eight-item dysphoria factor (see Table 1 for item mappings of both models). The DSM-5 model has

Table 1 PSS-SR item distribution across PTSD models

| DSM-5 PTSD symptoms from the PSS-SR | PTSD models | |
|---|-------------|-------------------|
| | (DSM-5) | (DSM-5 dysphoria) |
| B1: intrusive thoughts | RE | RE |
| B2: nightmares | RE | RE |
| B3: flashbacks | RE | RE |
| B4: emotional cue reactivity | RE | RE |
| B5: physiological cue reactivity | RE | RE |
| C1: avoidance of thoughts | A | A |
| C2: avoidance of reminders | A | A |
| D1: memory impairment | NAMC | D |
| D2: negative beliefs | NAMC | D |
| D3: distorted blame | NAMC | D |
| D4: persistent negative emotional states | NAMC | D |
| D5: lack of interest | NAMC | D |
| D6: feeling detached | NAMC | D |
| D7: inability to experience positive emotions | NAMC | D |
| E1: irritable/angry | AR | D |
| E2: reckless behavior | AR | H |
| E3: hypervigilance | AR | H |
| E4: easily startled | AR | H |
| E5: difficulty concentrating | AR | D |
| E6: difficulty sleeping | AR | D |

RE reexperiencing, *A* avoidance, *NAMC* negative alterations in mood and cognitions, *AR* alterations in arousal and reactivity, *H* hyperarousal, *D* dysphoria

garnered the most empirical support to date, particularly in studies comparing it to the alternative DSM-5 dysphoria model (Biehn et al., 2013; Contractor et al., 2014; Elhai et al., 2012a, b).

Dissociative Experiences Scale Factor Structure

The Dissociative Experiences Scales (DES; Bernstein and Putman, 1986) is arguably one of the most widely used measures of dissociative phenomenon in both clinical and community-based populations (van Ijzendoorn and Schuengel, 1996). Notable criticisms of the measure have however been highlighted; one in particular pertains to the variability with which studies report the DES's underlying dimensionality. Bernstein and Putman (1986) originally reported three underlying item groupings: (1) absorption, (2) depersonalization/derealization, and (3) amnesia. Absorption is characterized by a state of focused attention in which individuals become fully immersed in their own thoughts, for example, DES absorption items query whether individuals have become completely absorbed in a story or found themselves staring off into space, thinking of nothing, and being

unaware of time. Depersonalization and derealization are characterized by a state of un-connectedness with both the individuals self and the world around them, for example, DES depersonalization and derealization items query whether individuals have looked in a mirror and felt that they did not recognize themselves and whether they feel that they are looking at the world through a fog. Amnesia represents memory impairment; DES amnesia items query if individuals report having no memory for important events or finding themselves in a certain place with no memory of how they got there. A number of studies have subsequently agreed that the three-factor model best represented their DES data; however, the items which load on each of the three factors have not always been consistent (Ross, Joshi, and Currie, 1991, 1995; Ruiz, Poythress, Lilienfeld, and Douglas, et al. 2008; Sanders and Green, 1994; Stockdale, Gridley, Ware and Holtgraves, 2002). Moreover, a number of studies have reported alternative model conceptualizations comprising 1, 2, 4, and 7 factors (cf. Dunn, Ryan, and Paolo, 1994; Fisher and Elnitsky, 1990; Holtgraves and Stockdale 1997; Ray and Faith 1995; Olsen, Clapp, Parra, and Beck, 2013). Stockdale et al. (2002) additionally assessed two bi-factor (hierarchical) models; however, these were shown to have poorer fit compared to two-, three-, and four-factor inter-correlated models.

Notably, much of the factor analytic work has been exploratory rather than confirmatory in nature with the exception of a few studies (Ruiz et al., 2008; Stockdale et al., 2002; Olsen et al., 2013). An issue which is thought to further compound the consistency of results is the variability across studies in regard to the characteristics of their samples, in particular with regard to differences related to trauma exposures. Studies to date have focused on combat veterans reporting PTSD, trauma-exposed offenders (Amdur and Liberzon, 1996; Ruiz et al., 2008) and samples of college students (e.g., Sanders and Green, 1994; Olsen et al., 2013). Moreover, the clinical profiles of samples have also greatly varied; some are comprised of PTSD-diagnosed participants (Amdur and Liberzon, 1996) whereas others are comprised of individuals diagnosed with multiple personality disorder (Ross et al., 1995). Studies to date generally support a model comprising three factors: (1) absorption, (2) depersonalization and derealization, and (3) amnesia (cf. Stockdale et al., 2002). Notably, however, the most recent of confirmatory factor analysis (CFA) studies on DES items supported a two-factor model: (1) absorption and (2) amnesia/depersonalization.

Current Study

When researchers wish to compare the relationships between underlying groupings of disorders, one analytic strategy is to first assess factor analytical models of the disorders and then assess the latent-level associations between groupings. Examples of such studies include those assessing the latent-level

associations between PTSD and depression factors with a view to clarify the comorbidity mechanism (Biehn et al., 2013; Elhai et al., 2011a; Contractor, Durham, Brennan, Armour, Wutrick, Frueh et al., 2014). To date, however, no study has utilized the same analytic techniques to assess the latent-level relationships between PTSD and dissociation; this is the focus of the current study.

The current study aimed to (1) assess the fit of competing DSM-5 PTSD models, (2) assess the fit of competing dissociation models (using DES items), and (3) assess the latent-level relationships between the factors of the optimal PTSD and DES models in a combined model CFA. In doing so, we hoped to provide clarification on which factors may best count for the relationship between PTSD and dissociation. This line of inquiry is pertinent given the recent inclusion of a dissociative PTSD subtype in the DSM-5 nosology.

Method

Participants and Procedure

The current study is based on a convenience sample which consisted of 318 Northern Irish university students. Participants were approached on multiple campuses either during lecture breaks or in the campus library. All data were collected during the first academic semester (September–December) of the 2013–2014 academic year. Participation was entirely voluntary, and informed consent was given by all participants. Ethical permission for data collection was approved by a university ethics board as part of a larger data collection project.

Measures

Stressful Life Events Screening Questionnaire (SLESQ; Goodman, Corcoran, Turner, Yuan, and Green, 1998). The SLESQ assesses 12 traumatic experiences which qualify as Criterion A1 traumatic stressors as outlined in the DSM-IV. To ensure that the traumas queried were representative of the new DSM-5 trauma definition, we queried whether the “witnessed exposure item” and the “extreme exposure to gruesome or horrific details of trauma” were witnessed exclusively through electronic media. For the latter, we also queried whether “extreme exposure to gruesome or horrific details of trauma” was part of the individual’s occupational role. By using these qualifiers, we ensured that only individuals reporting trauma as defined by the DSM-5 were used in subsequent analysis. Previous studies have utilized this version of the SLESQ (cf. Biehn et al., 2013; Elhai, Miller, et al., 2012a, b).

PTSD Symptom Scale-Self-Report (PSS-SR; Foa, Riggs, Dancu, and Rothbaum, 1993). The PSS is a self-report

measure of the 17 PTSD items found in the DSM-IV-TR nosology. Items are rated using a four-category response format (0 “not at all” to 3 “five times or more per week/very much/almost always”). The measure used in the current study is a modified version of the original PSS-SR which was first used in a study by Elhai, Contractor et al., (2012a, b). The PSS-SR was modified to reflect the DSM-5 criteria of PTSD; thus, items reflecting a pervasive negative emotional state, excessive trauma-related blame, reckless behavior, and a negative perception of either oneself, the future, or the world were added. Previous studies conducted on the original version of the PSS-SR have reported that internal consistency ranges from .66 to .71, with test-retest reliability of .66–.77 (Foa and Tolin 2000). Contractor et al., (2014) implemented the modified measure used herein and reported an alpha coefficient of .96, which was identical to the internal consistency of the modified PSS-SR in the current study.

Dissociative Experiences Scale (DES; Bernstein and Putnam, 1986). The DES is a self-report measure of 28 items. Items are rated using an 11-category response format (0 “0 %” to 10 “10 %”). The measure was originally developed as a screening measure for dissociative disorders; however, it has since been utilized within both non-clinical and clinical samples (van Ijzendoorn and Schuengel, 1996). Bernstein and Putnam, (1986) reported that the internal consistency of the DES was high (.83–.93). Test-retest reliability has also been evidenced (.78–.84) (Carlson and Putnam, 1993). The internal consistency of the DES in the current study was high (alpha=.93).

Missing Data

A total of 318 responders completed the survey; of those 202 endorsed, a DSM-5 defined Criterion A trauma. An investigation into the item level-missing data across the PTSD items revealed that 35 participants were missing over 70 % of the PSS items reducing the sample to 167 participants. An examination of the 167 cases revealed that 2 of these cases were missing over 70 % of DES items, so they were removed from the sample. Thus, the effective sample size consisted of 165 trauma-exposed participants. The remaining nominal amounts of missing data from the PSS and DES items were treated within the Mplus 6.1 software using maximum likelihood (ML) procedures (Graham, 2009).

Analyses

All analyses were conducted using Mplus 6 software. The analyses were conducted in three steps. Step one implemented testing the fit of PTSD’s competing models using the 20 PSS-SR items which reflect DSM-5 PTSD diagnostic criteria. Two four-factor models were specified and estimated: (1) a DSM-5 dysphoric model and (2) the DSM-5 model (see Table 1).

PSS-SR items have four response options therefore were treated as ordinal indicators within the CFA. When estimating ordinal indicators, CFA uses a polychoric covariance matrix, the robust-weighted least squares, with a mean- and variance-adjusted chi-squared (WLSMV) and probit regression coefficients (Flora and Curran, 2004). In addition, error covariances were fixed to zero, and the variances of the latent factors were fixed to one. Of note, these models were reestimated using the ML estimator to achieve Bayesian information criterion (BIC) values to allow comparison between the two non-nested models. Chi-square difference testing is only appropriate when models are nested.

Step two also implemented testing competing model CFA; this time utilizing 28 indicators from the DES. Four competing models were specified and estimated: (1) a single factor model, (2) a two-factor absorption and amnesia model, (3) a three-factor absorption, amnesia, and depersonalization/derealization model, and (4) a four-factor depersonalization/derealization, absorption, distractibility, and memory disturbance model (see Table 2). DES items have ten response options therefore were treated as continuous indicators within the CFA. When estimating continuous indicators, CFA uses the ML estimator. ML estimation has however been shown to inflate chi-square statistics and thus bias the fit of the model under conditions of non-normality (Curran, West, and Finch, 1996). Given extreme violations in normality on the DES indicators, the models were estimated using the robust maximum likelihood (MLR) estimator which includes a correction factor to adjust the chi-square to account for non-normality (Satorra and Bentler, 1994). In the DES models, error covariances were also fixed to zero and the variances of the latent factors were fixed to one.

Step three implemented Wald chi-squared tests to assess the latent-level relations between PTSD's optimal model and the optimal model of the DES indicators. Both optimal CFA models were specified and estimated in a single model which allowed for the assessment of correlations between all PTSD and DES factors. Wald chi-squared tests of parameter constraints assessed (1) whether a particular DES model factor was more or less related to the PTSD model factors and (2) whether a particular PTSD model factor was more or less related to the DES model factors. For example, we tested whether absorption was more or less related to reexperiencing compared to avoidance.

Optimal model selection in steps one and two rely on comparative fit index (CFI) and Tucker-Lewis index (TLI) values equal to or greater than .95 (.90–.94) and root mean square error of approximation (RMSEA) values equal or greater than .06 (.07–.08). Models which meet these specifications are generally deemed as adequate (Hu and Bentler, 1999). As noted above, the BIC also assists in model selection with lower values being preferential. Indeed a ten-point difference in BIC values is indicative of very strong support for

the model with the lowest value (Kass and Raftery, 1995). In step three, a significant Wald chi-squared test ($p < .05$) indicates that the two pairs of correlations differ from 0 and thus are significantly different.

Results

Demographics of Effective Sample

The majority of participants were female ($N=133$; 80.6 %) and ranged in age from 18 to 48 ($M=23.07$, $SD=6.88$; [29 were >25]) and were predominantly Caucasian ($N=164$; 99.4 %). Almost half of the Caucasian participants nominated themselves as white British ($N=80$; 48.5 %). The remaining students nominated themselves as either white Irish ($N=30$; 18.2 %) or white other ($N=54$; 32.7 %). All participants were current university students. Details pertaining to trauma exposures are presented in Table 3. The most frequently endorsed trauma within the current sample was “Has an immediate family member, romantic partner, or very close friend died as a result of accident, homicide, or suicide?” ($N=79$; 47.9 %) followed by “physical violence” ($N=55$; 33.3 %). Mean PTSD symptom cluster scores in the effective sample were 3.84 ($SD=7.32$), 1.81 ($SD=3.13$), 4.30 ($SD=5.35$), 4.01 ($SD=7.21$) (reexperiencing [range=0–73], avoidance [0–33], negative alterations in mood and cognitions [0–32], and alterations in arousal and reactivity [0–78], respectively). Probable PTSD based on meeting trauma endorsement criteria and positive endorsement of all four PTSD symptom clusters was met by 13.3 % ($N=21$) of the sample. Of note, a further 13.3 % ($N=21$) met three of four PTSD symptom clusters.

CFA Results: PTSD Model and Dissociation Model

The DSM-5 PTSD model fit the data well, as did the alternative DSM-5 dysphoria model. Notably, the fit of these two models was almost identical in the current sample. When rerunning the models using the MLR estimator to achieve BIC values, these differed by only 1 point. Therefore, given that the DSM-5 model is current and has been supported across previous studies, this model was chosen for our analysis. Of the four models of the DES, the three-factor model (see Table 2) provided the best fit to the data and thus was deemed the optimal model for use in subsequent analysis. Fit indices for competing PTSD models can be viewed in Table 4. Fit indices for competing dissociation models can be viewed in Table 5.

Combined CFA Model

The CFA model which simultaneously specified and estimated the optimal four-factor DSM-5 PTSD model and the

Table 2 DES item distribution across dissociative models

| DES items | DES models | | | |
|---|--------------------|--------------------|----------------------|---------------------|
| | (One-factor model) | (Two-factor model) | (Three-factor model) | (Four-factor model) |
| 1: Forgot bus or car journey | D | AB | AB | DST |
| 2: Did not hear a conversation | D | AB | AB | DST |
| 3: Found in a place with no idea how you got there | D | AMN | AMN | AMN |
| 4: Found dressed in clothes that you don't remember putting on | D | AMN | AMN | AMN |
| 5: Found new things that you did not remember buying | D | AMN | AMN | AMN |
| 6: Approached by strangers | D | AMN | AMN | AMN |
| 7: Standing next to yourself/watching yourself | D | AMN | DEP | DEP |
| 8: Not recognizing friends or family | D | AMN | AMN | AMN |
| 9: No memory for important events | D | AMN | AMN | AMN |
| 10: Accused of lying | D | AB | AB | DEP |
| 11: No self-recognition in mirror | D | AMN | DEP | DEP |
| 12: Feeling other people, objects, and the world are not real | D | AMN | DEP | DEP |
| 13: Feeling your body does not belong to you | D | AMN | DEP | DEP |
| 14: Sometimes remembering a past event so vividly | D | AB | AB | AB |
| 15: Not knowing whether things you remember happened | D | AB | AB | DST |
| 16: Finding a familiar place strange and unfamiliar. | D | AB | AB | DEP |
| 17: Completely absorbed in a story | D | AB | AB | AB |
| 18: So involved in a fantasy/daydream that it seems real | D | AB | AB | AB |
| 19: Able to ignore pain | D | AB | AB | AB |
| 20: Staring off into space, thinking of nothing, and unaware of time | D | AB | AB | AB |
| 21: Talking out loud to yourself | D | AB | AB | DST |
| 22: Acting differently across situations | D | AB | AB | ABS |
| 23: Able to do things with amazing ease and spontaneity | D | AB | AB | AB |
| 24: Cannot remember whether you have done something | D | AB | AB | DST |
| 25: Evidence that you have done things that you do not remember | D | AB | AB | DST |
| 26: Found writings, drawings, or notes that you cannot remember doing | D | AB | AB | DST |
| 27: Heard voices telling you to do things | D | AMN | DEP | DEP |
| 28: Looking at the world through a fog | D | AMN | DEP | DEP |

D dissociation, *AB* absorption, *AMN* amnesia, *DEP* depersonalization/derealization, *DST* distractibility

optimal three-factor DES model fit the data well, $\chi^2(1,059)=1,360.057$, $p<.0001$; RMSEA=.042; CFI=.926; TLI=.921. The inter-factor correlations between the four PTSD factors and the three DES factors were all positive and statistically significant ($p<.001$). Correlations are presented across Tables 5 and 6.

Wald Tests of Parameter Constraint

When comparing the relation of each PTSD factor to depersonalization/derealization (DEP) and alternative DES factors, most Wald tests of parameter constraints revealed non-significant ($p>.01$) findings (see Tables 6 and 7). Of note, when comparing the correlation between reexperiencing (RE) with absorption (AB) ($r=.432$) compared to RE with amnesia (AMN) ($r=.242$), the Wald test of parameter constraint $\chi^2(1,$

$N=165)=12.664$, $p<.001$ indicated that AB was more related to RE compared to AMN. Likewise, when comparing the correlation between alterations in arousal and reactivity (AR) with AB ($r=.462$) compared to AR with AMN ($r=.302$), the Wald test of parameter constraint $\chi^2(1, N=165)=7.776$, $p=.005$ indicated that AB was more related to AR compared to AMN (see Table 6).

When assessing the latent-level relationships between DEP and the four PTSD factors, significant Wald tests of parameter constraints revealed that DEP is more related to AR ($r=.432$) compared to avoidance (AV) ($r=.289$), $\chi^2(1, N=165)=8.352$, $p=.004$ (see Table 7). See Tables 5 and 6 for the results of all correlations and Wald's chi-squared parameter tests.

Given the reported relationship between DEP and AR, we conducted post hoc analysis to determine how each of the

Table 3 Trauma endorsement of the effective sample ($N=165$)

| Trauma experience | Frequency (%) |
|--|---------------|
| Death of a close friend or family member | 79 (47.9) |
| Physical violence | 55 (33.3) |
| Sexual assault | 45 (27.3) |
| Present when another was killed, injured, physically, or sexually assaulted (excluding when this was on TV) | 46 (27.9) |
| Life-threatening accident | 38 (23.0) |
| Extreme exposure to gruesome or horrifying details of another person's death, serious injury, or sexual violation (excluding if only on TV but allowing if in a work role) | 38 (23.0) |
| Physical maltreatment | 29 (17.6) |
| Life-threatening illness | 27 (16.4) |
| Attempted rape | 25 (15.2) |
| Threatened with a weapon | 23 (13.9) |
| Rape | 19 (11.5) |
| Serious life danger e.g., military combat/living in a war zone | 10 (6.1) |
| Physical force or a weapon used against you in a robbery or mugging | 8 (4.8) |

Trauma endorsements are not mutually exclusive

individual items of AR correlated with the DEP factor. All AR items significantly correlated with the DEP factor: irritable/angry ($r=.336$), reckless behavior ($r=.376$), hypervigilance ($r=.323$), easily startled ($r=.346$), difficulty concentrating ($r=.367$), and difficulty sleeping ($r=.231$). All correlations were significant at the $p<.001$ level. When comparing pairs of correlations, Wald tests of parameter constraints revealed that no item of AR was significantly more or less related to DEP than others ($p>.001$).

Discussion

The current study utilized a university sample of 167 trauma-exposed individuals. In general, trauma exposure was high; 107 (64.8 %) of participants endorsed two or more traumatic event exposures, and 25 (15.2 %) endorsed five or more

traumatic event exposures. This is however not uncommon within the Northern Ireland population where trauma exposure is prevalent (cf. Bunting, Ferry, Murphy, O'Neil and Bolton, 2013). Overall, results suggest that alternative dissociative symptom groupings are neither more or less related to PTSD groupings compared to DEP; however, differential relations exist in relation to how DEP associates with PTSD factors.

Neither of the two four-factor PTSD models was found to be superior to the other; therefore, the DSM-5 model was used in further analyses in the current study. This decision was based on the findings of previous studies. For example, Elhai et al., (2012a, b) reported that of competing models, the DSM-5 four-factor model provided the best fit to their data. Additional studies also found superior fit for the DSM-5 model compared to the alternative DSM-5 dysphoria model (cf. Biehn et al., 2013; Contractor et al., 2014), the former using a sample of undergraduate psychology students and the latter a primary care sample.

Of competing DES models, the three-factor model comprising an absorption, amnesia, and depersonalization/derealization factor provided the best fit to the data with some evidence for a four-factor solution (see Table 2). Superior fit of the three-factor model concurs with previous research, for example, with Stockdale et al., (2002) who investigated the factor structure of the DES in two samples of midwestern university students ($N=971$ [69 % female] and $N=400$ [69 % female]). Utilizing exploratory factor analysis, on the sample comprising 971 students, Stockdale et al., (2002) reported three defensible solutions comprising one-, two-, and three-factors. A four-factor solution was also deemed defensible but less, so given it produced a fourth factor comprising only two items and resulted in an Eigen value of 1.05 (close to the cutoff value of 1). Utilizing CFA, the four inter-correlated models were fit to the same data ($N=971$). In addition, two hierarchical models were specified and estimated. Of these competing models, the three-factor inter-correlated model was deemed optimal. To further validate the models, Stockdale et al. utilized the second sample of 400 undergraduate students and conducted a competing model CFA, again concluding that optimal fit was provided by the three-factor model which is supported in the current study. Of note, the poor fit demonstrated by the DES in this trauma-exposed sample may be attributable to criticisms of the DES in that it is not particularly representative of trauma-related dissociation rather a number of items are said to represent either relatively common and minor experiences of dissociation found even in the absence of psychological disorder whereas others, admittedly, represent more severe experiences, however, are primarily associated with dissociative disorders (Carlson et al., submitted; Dalenberg and Paulson, 2009).

The combined CFA models indicated that all correlations between PTSD and dissociative factors were positive and

Table 4 Fit indices for competing PSS models

| Fit statistics | DSM-5 dysphoria | DSM-5 numbing |
|----------------|------------------------|------------------------|
| Chi-squared | $\chi^2 (164)=323.900$ | $\chi^2 (164)=317.964$ |
| RMSEA | .077 | .075 |
| CFI | .969 | .970 |
| TLI | .964 | .965 |
| BIC | 7,200.522* | 7,201.190* |

RMSEA root mean square error of approximation, CFI comparative fit index, TLI Tucker-Lewis index, BIC Bayesian information criteria

Table 5 Fit indices for competing DES models

| | DES—1 factor | DES—2 factor | DES—3 factor | DES—4 factor |
|----------------|------------------------|------------------------|------------------------|------------------------|
| Fit statistics | χ^2 (350)=797.310 | χ^2 (349)=760.874 | χ^2 (347)=727.020 | χ^2 (344)=733.587 |
| Chi-square | .088 (.080–.096) | .085 (.076–.093) | .081 (.073–.090) | .083 (.075–.091) |
| RMSEA | .697 | .721 | .743 | .736 |
| CFI | .673 | .698 | .720 | .710 |
| TLI | 19,001.691 | 18,923.890 | 18,863.281 | 18,894.181 |
| BIC | | | | |

RMSEA root mean square error of approximation, CFI comparative fit index, TLI Tucker-Lewis index, BIC Bayesian information criteria

statistically significant. Correlations ranged from .242 (RE with AMN) to .462 (AR with AB). When focusing on the correlations between the DEP factor with PTSD factors, the correlations ranged from .289 (with AV) to .432 (with AR). All correlations are provided in Tables 6 and 7. Given that PTSD and dissociation have long been associated, we would expect that the latent factors of the two disorders would correlate. Furthermore, although we expect these to be related, we would not expect them to be highly correlated given they are essentially two distinct disorders. This is confirmed in the current study.

In assessing if alternative dissociative factors were more or less related to PTSD factors compared to DEP, results revealed that no alternative dissociative factor (AB or AMN) was more or less related to any of the four PTSD factors as compared to the DEP factor. This is contrary to what we would have expected given the emphasis on depersonalization and derealization items in the newly proposed dissociative PTSD subtype in the DSM-5 (cf. Lanius et al. 2010). Although not

of central focus to the current study, it is notable that Wald chi-squared tests of parameter constraints revealed that absorption was more related to AR and RE compared to AMN, despite there being an amnesia item included within PTSD's negative alterations in mood and cognitions (NAMC) factor. Notably, however, CFA studies have often reported that the lowest factor loading across all items was the factor loading of memory impairment (cf. Armour, Carragher, and Elhai, 2013a, b, c, 2012).

In comparing correlations of DEP with PTSD factors, the current results indicated only one significant finding that DEP was more related to AR ($r=.432$) compared to AV ($r=.289$). This is in part contradictory to results reported by Stein et al., (2013) who reported that among respondents meeting PTSD criteria, the presence of dissociation was positively associated with the number of reexperiencing symptoms but not with the number of hyperarousal symptoms. This discrepancy across studies may in part be accountable to both reexperiencing and hyperarousal symptoms belonging to the fear/threat processing dimension of the PTSD phenotype rather than the distress dimension. Thus, both have commonalities, and therefore, it is reasonable that both may in some way be related to dissociation (Zoellner, Pruitt, Farach and Jun, 2014). Furthermore, alternative differences across study samples may increase or decrease the association of each with symptoms of

Table 6 Wald's chi-squared test of parameter constraints comparing the relation between each PTSD factor with factors of dissociation

| Correlated factors | <i>r</i> | Correlated factors | <i>r</i> | Wald test (<i>p</i>) |
|--------------------|----------|--------------------|----------|------------------------|
| RE with AB | .432 | RE with AMN | .242 | 12.664 (.000)* |
| RE with AB | .432 | RE with DEP | .327 | 2.703 (.100) |
| RE with AMN | .242 | RE with DEP | .327 | 1.58 (.209) |
| AV with AB | .397 | AV with AMN | .305 | 2.257 (.133) |
| AV with AB | .397 | AV with DEP | .289 | 2.514 (.113) |
| AV with AMN | .305 | AV with DEP | .289 | 0.045 (.833) |
| NAMC with AB | .386 | NAMC with AMN | .267 | 4.656 (.031) |
| NAMC with AB | .386 | NAMC with DEP | .377 | 0.028 (.867) |
| NAMC with AMN | .267 | NAMC with DEP | .377 | 2.793 (.095) |
| AR with AB | .462 | AR with AMN | .302 | 7.776 (.005)* |
| AR with AB | .462 | AR with DEP | .432 | .270 (.604) |
| AR with AMN | .302 | AR with DEP | .432 | 3.306 (.069) |

RE reexperiencing, AV avoidance, NAMC negative alterations in mood and cognitions, AR alterations in arousal and reactivity, AB absorption, AMN amnesia, DEP depersonalization/derealization

All correlations were significant ($p<.001$). Significantly different correlated pairs ($p<.01$) are highlighted by *

Table 7 Wald's chi-squared test of parameter constraints comparing depersonalization/derealization across PTSD factors

| Correlated factors | <i>r</i> | Correlated factors | <i>r</i> | Wald test (<i>p</i>) |
|--------------------|----------|--------------------|----------|------------------------|
| DEP with RE | .327 | DEP with AV | .289 | .856 (.355) |
| DEP with RE | .327 | DEP with NAMC | .377 | 1.838 (.175) |
| DEP with RE | .327 | DEP with AR | .432 | 5.853 (.016) |
| DEP with AV | .289 | DEP with NAMC | .377 | 4.458 (.035) |
| DEP with AV | .289 | DEP with AR | .432 | 8.352 (.004)* |
| DEP with NAMC | .377 | DEP with AR | .432 | 2.480 (.115) |

RE reexperiencing, AV avoidance, NAMC negative alterations in mood and cognitions, AR alterations in arousal and reactivity, DEP depersonalization/derealization

All correlations were significant ($p<.001$). Significantly different correlated pairs ($p<.01$) are highlighted by *

dissociation. Notably, however, previous research has reported that levels of arousal and dissociation are highly correlated (Moleman et al., 1992). Moreover, McFarlane, (2013) states “The accumulated evidence is that dissociation is a regulatory strategy to deal with the intense hyperarousal that PTSD sufferers experience” (p. 296). This is therefore supportive of the finding that DEP was more related to AR compared to AV. It is however important to acknowledge that although the results suggest a potential leading role for AR as it relates to DEP, we must stress that this is only based on an examination of inter-correlations between latent factors and only in comparison to AV. Indeed, AR was not significantly different in its relationship with DEP compared to both NAMC and RE (albeit the latter would have been significant at the $p \leq .05$ level). It is also important to acknowledge that we are comparing the depersonalization and derealization dissociative indicators specifically to DSM-5 PTSD groupings whereas Stein et al. focused their comparisons on the three PTSD symptom groupings comprising the 17 PTSD items from the DSM-IV-TR. Of note, given the aim of their study, dissociation represented only two depersonalization items and one derealization item rather than depersonalization/derealization (or indeed dissociation) more generally.

In conclusion, by assessing the inter-relations between PTSD and dissociation factors, the current study raises questions as to which dissociative items should be included in the diagnostic symptom set for dissociative PTSD. Indeed, Wald chi-squared tests of parameter constraints revealed that the DEP factor was no more or less related to any PTSD factor compared to alternative dissociative factors. The results did however highlight that AB was more related to PTSD factors compared to AMN which warrants further exploration. Interestingly, few differential relations between the DEP factor of dissociation and the four factors of PTSD were evident (i.e., a single significant finding that DEP was more related to AR ($r = .432$) compared to AV ($r = .289$), Wald χ^2 (1, $N = 165$) 8.352, $p = .004$). The current results therefore provide tentative support that the mechanism for comorbid PTSD and dissociation may be related to PTSD’s arousal factor.

Forensic and Clinical Implications

In forensic mental health, an important aspect related to PTSD is the identification of valid from non-valid cases (malingering). Historically, forensic cases related to PTSD and dissociation have been successfully concluded as acts of self-defense and related to diminished capacity; in fact, some of the PTSD symptoms most related to criminal cases are dissociation and hyperarousal (Berger et al., 2012). Thus, given the addition of a dissociative PTSD subtype in the DSM-5 nosology, the current study’s overarching aim, to assess PTSD and dissociation latent-level relations is timely and important. In doing so, we have highlighted based on the

inter-correlations between latent factors, that DEP items, as would be expected from the DSM-5 dissociative PTSD criteria, are perhaps not the only dissociative symptoms indicative of dissociative PTSD. Thus, patients presenting with a wide range of dissociative experiences should also be queried in relation to PTSD. In forensic cases, it would be also be helpful to assess for diversity of dissociative phenomena given its relation to different PTSD symptom clusters in the current study. We highlighted that the mechanism for comorbid PTSD and dissociation may be related to PTSD’s AR factor. If this proves to be the case, clinicians must be aware that those presenting with high levels of AR may also be experiencing dissociative symptomatology, which is known to associate with self-harm and suicidality (Foot, Smolin, Neft, and Lipschitz, 2008), and that this may be particularly prevalent in forensic populations (Hawton Linsell Adeniji Sariaslan, and Fazel, 2014). In fact, the current study results add to the idea that self-defense in forensic cases could be related to hyperarousal symptoms (e.g., poor impulse control and overestimation of danger) and dissociative flashbacks (Berger et al., 2012). The current study is also pertinent in the psychological injury context, e.g., disability claims, given that assessment using DSM-5 criteria will quickly become more prevalent.

Study Limitations

The current study is the first, to our knowledge, to assess the latent-level relationships between groupings of PTSD and dissociative items. However, several limitations are present. First, we employed the use of data from university students and so are unable to determine if the current findings will generalize to alternative trauma-exposed samples. In line with this, the current study focused on only those who were trauma exposed resulting in a smaller sample size. It is important to note that larger samples with a more diverse trauma profile may reveal further associations between PTSD and dissociation factors. It is therefore important that the current study is replicated in alternative samples. In addition, although our sample is comprised of individuals who have endorsed Criterion A traumas as per the DSM-5 stipulations, different result may be found in samples which focus on specific trauma exposures, for example rape and sexual assault victims, rather than heterogeneous trauma exposures. Moreover, 13.3 % of the current sample met the criteria for probable PTSD; results may differ in samples reporting a greater prevalence of PTSD. Indeed, Stein et al., (2013) reported that the prevalence of dissociative symptoms was increased in those with PTSD who reported sexual violence (19.9 %) compared to those reporting alternative traumas such as the death of a loved one (13.0 %). Second, we assessed trauma exposure, PTSD symptoms, and dissociative symptoms using self-report measures rather than via a clinical interview. Thus, the results may be biased to issues of socially desirable answering and issues with memory

recall. Notably, however, diagnostic concordance rates between the use of self-report measures and clinical interviews have been reported as high (Coffey et al., 1998; Harrington and Newman, 2007). Similarly, we acknowledge the potential for sampling bias in this study given that interest in participation may be linked to students who identify with the area of research. We also acknowledge that we used a modified version of the PSS to encapsulate the DSM-5 criteria; as such, we can evidence reliability of this measure but not the validity. Third, the DES model used for latent-level comparisons provided poor fit to the data, albeit superior fit to alternative DES models. However, the model used has gathered previous support in two samples of university students (Stockdale et al., 2002) and thus was deemed the best option available to us.

To truly determine the latent-level associations between PTSD and dissociative symptom groupings, in particular the DEP grouping, this study should be replicated using both clinical samples and samples who present with alternative trauma profiles. Furthermore, researchers should endeavor to assess the temporal relationships between these latent factors in an attempt to clarify whether certain factors drive the course of alternative factors. One option would be to determine if alterations in arousal and reactivity covary or not with depersonalization/derealization over time and during treatment. Given that we only assessed the inter-relationships between latent factors of PTSD and dissociation, an additional line of inquiry would be to ascertain the predictive value of individual dissociative indicators as they relate to the dissociative PTSD subtype.

References

- Amdur, R., & Liberzon, I. (1996). Dimensionality of dissociation in subjects with PTSD. *Dissociation: Progress in the Dissociative Disorders*, 9, 118–124.
- American Psychiatric Association. (1980). *Diagnostic and statistical manual of mental disorders* (3rd ed.). Washington, DC: Author.
- American Psychiatric Association. (1987). *Diagnostic and statistical manual of mental disorders* (3rd ed., text rev.). Washington, DC: Author.
- American Psychiatric Association. (2000). *Diagnostic and statistical manual of mental disorders* (4th ed., text rev.). Washington, DC: Author.
- American Psychiatric Association. (2013). *Diagnostic and statistical manual of mental disorders* (5th ed.). Washington, DC: Author.
- Armour, C., Elhai, J. D., Richardson, D., Ractliffe, K., Wang, L., & Elklit, A. (2012). Assessing a five factor model of PTSD: is dysphoric arousal a unique PTSD construct showing differential relationships with anxiety and depression? *Journal of Anxiety Disorders*, 26(2), 368–376. doi:10.1016/j.janxdis.2011.12.002
- Armour, C., Carragher, N., & Elhai, J. D. (2013a). Assessing the fit of the dysphoric arousal model across two nationally representative epidemiological surveys; the Australian NSMHWB and the United States NESARC. *Journal of Anxiety Disorders*, 27, 109–115.
- Armour, C., Ghazali, S. R., & Elklit, A. (2013b). PTSD's latent structure in Malayan tsunami victims: assessing the newly proposed dysphoric arousal model. *Psychiatry Research*, 1, 26–32.
- Armour, C., O'Conner, M., Elklit, A., & Elhai, J. D. (2013c). Assessing PTSD's latent structure in European trauma victims: evidence for a five factor dysphoric and anxious arousal model. *Journal of Nervous and Mental Disease*, 201, 901–906.
- Armour, C., Elklit, A., Lauterbach, D., Elhai, J. D. (2014a). The DSM-5 dissociative-PTSD subtype: can levels of depression, anxiety, hostility, and sleeping difficulties differentiate between dissociative-PTSD and PTSD in rape victims? *Journal of Anxiety Disorders*. (in press).
- Armour, C., Karstoft, K.I., & Richardson, J. D. (2014b). The co-occurrence of PTSD and dissociation: differentiating severe PTSD from dissociative-PTSD. *Social Psychiatry & Psychiatric Epidemiology*. (in press).
- Berger, O., McNiel, D. E., & Binder, R. L. (2012). PTSD as a criminal defense: a review of case law. *Journal of the American Academy of Psychiatry & Law*, 4, 509–521.
- Bernstein, E. M., & Putnam, F. W. (1986). Development, reliability, and validity of a dissociation scale. *Journal of Nervous and Mental Disease*, 174(12), 727–735.
- Bewin, C. R., & Holmes, E. A. (2003). Psychological theories of post-traumatic stress disorder. *Clinical Psychology Review*, 23, 339–376.
- Biehn, T. L., Contractor, A., Elhai, J. D., Tamburrino, M., Fine, T. H., Prescott, M. R., et al. (2013). Relations between the underlying dimensions of PTSD and major depression using an epidemiological survey of deployed Ohio National Guard soldiers. *Journal of Affective Disorders*, 144, 106–111.
- Breh, D. C., & Seidler, G. H. (2007). Is peritraumatic dissociation a risk factor for PTSD? *Journal of Trauma and Dissociation*, 8, 53–69.
- Breuer, J., & Freud, S. (1955). *Studies on hysteria. The Standard Edition of the Complete Psychological Works of Sigmund Freud*. London: The Hogarth Press (Original Publication, 1895).
- Bunting, B., Ferry, F., Murphy, S. D., O'Neill, S., & Bolton, D. (2013). Trauma associated with civil conflict and posttraumatic stress disorder: evidence from the Northern Ireland Study of Health and Stress. *Journal of Traumatic Stress*, 26, 134–141.
- Carlson, E. B., Waelde, L., Palmieri, P. A., McDade-Montez, E. Macia, K. S. (submitted). Development and validation of the Dissociative Symptoms Scale.
- Carlson, E. B., & Putnam, F. W. (1993). An update on the Dissociative Experiences Scale. *Dissociation*, 6, 16–27.
- Cloitre M, Petkova E, Wang J, Lu Lassell F (2012). An examination of the influence of a sequential treatment on the course and impact of dissociation among women with PTSD related to childhood abuse [published online ahead of print May 1]. *Depress Anxiety*. <http://dx.doi.org/10.1002/da.21920>
- Coffey, S. F., Dansky, B. S., Falsetti, S. A., Saladin, M. E., & Brady, K. T. (1998). Screening for PTSD in a substance abuse sample: psychometric properties of a modified version of the PTSD Symptom Scale Self-Report. *Journal of Traumatic Stress*, 11, 393–399.
- Contractor, A. A., Durham, T. A., Brennan, J. A., Armour, C., Wutrick, H. R., Frueh, B. C., & Elhai, J. D. (2014). DSM-5 PTSD's symptom dimensions and relations with major depression's symptom dimensions in a primary care sample. *Psychiatry Research*, 215, 146–153.
- Curran, P. J., West, S. G., & Finch, J. F. (1996). The robustness of test statistics to nonnormality and specification error in confirmatory factor analysis. *Psychological Methods*, 1, 16–29.
- Dalenberg, C., & Paulson, K. (2009). The case for the study of "normal" dissociation processes. In P. F. Dell, J. O'Neil, & E. Somer (Eds.), *Dissociation and the dissociative disorders: DSM-V and beyond* (pp. 145–154). New York: Routledge.

- Dunn, G. E., Ryan, J. J., & Paolo, A. M. (1994). A principal components analysis of the dissociative experiences scale in a substance abuse population. *Journal of Clinical Psychology*, 50, 936–940.
- Elhai, J. D., & Palmieri, P. A. (2011). The factor structure of posttraumatic stress disorder: a literature update, critique of methodology, and agenda for future research. *Journal of Anxiety Disorders*, 25, 849–854. doi:10.1016/j.janxdis.2011.04.007
- Elhai, J. D., Biehn, T. L., Armour, C., Klopfer, J. J., Frueh, B. C., & Palmieri, P. A. (2011a). Evidence for a unique PTSD construct represented by PTSD's D1-D3 symptoms. *Journal of Anxiety Disorders*, 25, 340–345. doi:10.1016/j.janxdis.2010.10.007
- Elhai, J. D., Contractor, A. A., Palmieri, P. A., Forbes, D., & Richardson, J. D. (2011b). Exploring the relationship between underlying dimensions of posttraumatic stress disorder and depression in a national, trauma-exposed military sample. *Journal of Affective Disorders*, 133, 477–480. doi:10.1016/j.jad.2011.04.035
- Elhai, J. D., Contractor, A., Tamburrino, M., Fine, T. H., Prescott, M. R., Shirley, E., et al. (2012a). The factor structure of major depression symptoms: a test of four competing models using the Patient Health Questionnaire-9. *Psychiatry Research*, 199, 169–173. doi:10.1016/j.psychres.2012.05.018
- Elhai, J. D., Miller, M. E., Ford, J. D., Biehn, T. L., Palmieri, P. A., & Frueh, B. C. (2012b). Posttraumatic stress disorder in DSM-5: estimates of prevalence and symptom structure in a nonclinical sample of college students. *Journal of Anxiety Disorders*, 26, 58–64. doi:10.1016/j.janxdis.2011.08.013
- Fisher, D. G., & Elnitsky, S. (1990). A factor analytic study of two scales measuring dissociation. *American Journal of Clinical Hypnosis*, 32, 201–207.
- Flora, D. B., & Curran, P. J. (2004). An empirical evaluation of alternative methods of estimation for confirmatory factor analysis with ordinal data. *Psychological Methods*, 9(4), 466–491. doi:10.1037/1082-989x.9.4.466
- Foa, E. B., Riggs, D. S., & Gershuny, B. S. (1995). Arousal, numbing, and intrusion: symptom structure of PTSD following assault. *The American Journal of Psychiatry*, 152(1), 116–120.
- Foa, E. B., & Tolin, D. F. (2000). Comparison of the PTSD Symptom Scale-Interview version and the Clinician-Administered PTSD Scale. *Journal of Traumatic Stress*, 13, 181–191. doi:10.1023/A:1007781909213
- Foote, B., Smolin, Y., Neft, D. I., & Lipschitz, D. J. (2008). Dissociative disorders and suicidality in psychiatric outpatients. *Journal of Nervous and Mental Disease*, 1, 29–36.
- Friedman, M. J., Resick, P. A., Bryant, R. A., Strain, J., Horowitz, M., & Spiegel, D. (2011). Classification of trauma and stressor-related disorders in DSM-5. *Depression and Anxiety*, 28, 737–749. doi:10.1002/da.20845
- Gershuny, B. S., & Thayer, J. F. (1999). Relations among psychological trauma, dissociative phenomena, and trauma-related distress: a review and integration. *Clinical Psychology Review*, 19, 631–637.
- Ginzburg, K., Koopman, C., Butler, L. D., Palesh, O., Kraemer, H. C., Classen, C. C., & Spiegel, D. (2006). Evidence for a dissociative subtype of post-traumatic stress disorder among help-seeking childhood sexual abuse survivors. *Journal of Trauma & Dissociation*, 7, 7–27.
- Goodman, L., Corcoran, C., Turner, K., Yuan, N., & Green, B. L. (1998). Assessing traumatic event exposure: general issues and preliminary findings for the Stressful Life Events Screening Questionnaire. *Journal of Traumatic Stress*, 11, 521–542.
- Graham, J. W. (2009). Missing data analysis: making it work in the real world. *Annual Review of Psychology*, 60(1), 549–576. doi:10.1146/annurev.psych.58.110405.085530
- Harpaz-Rotem, I., Tsai, J., Pietrzak, R. H., & Hoff, R. (2014). The dimensional structure of posttraumatic stress symptomatology in 323,903 U.S. veterans. *Journal of Psychiatric Research*, 49, 31–36. doi:10.1016/j.jpsychores.2013.10.020
- Harrington, T., & Newman, E. (2007). The psychometric utility of two self-report measures of PTSD among women substance users. *Addictive Behaviours*, 32, 2788–2798.
- Hawton, K., Linsell, L., Adeniji, T., Sariaslan, A., & Fazel, S. (2014). Self-harm in prisons in England and Wales: an epidemiological study of prevalence, risk factors, clustering, and subsequent suicide. *Lancet*, 383, 1147–54. doi:10.1016/S0140-6736(13)62118-2
- Holtgraves, T., & Stockdale, G. (1997). The assessment of dissociative experiences in a non-clinical population: reliability, validity, and factor structure of the dissociative experiences scale. *Personality and Individual Differences*, 22, 699–706.
- Hu, L., & Bentler, P. M. (1999). Cutoff criteria for fit indexes in covariance structure analysis: conventional criteria versus new alternatives. *Structural Equation Modeling*, 6, 1–55.
- Janet, P. (1907). The major symptoms of hysteria, Macmillan, London/New York. Second edition with new matter: 1920. Reprint of 1920-edition: Hafner, New York, 1965. Janet, P. (1909a). Les névroses, Flammarion, Paris.
- Kass, R. E., & Rafferty, A. E. (1995). Bayes factor. *Journal of the American Statistical Association*, 90, 430. doi:10.1080/01621459.1995.10476572
- King, D. W., Leskin, G. A., King, L. A., & Weathers, F. W. (1998). Confirmatory factor analysis of the Clinician-Administered PTSD Scale. *Psychological Assessment*, 10(2), 90–96. doi:10.1037/1040-3590.10.2.90
- Lanius, R. A., Vermetten, E., Loewenstein, R. J., Brand, B., Schmahl, C., Bremner, J. D., & Spiegel, D. (2010). Emotion modulation in PTSD: clinical and neurobiological evidence for a dissociative subtype. *American Journal of Psychiatry*, 167, 640–647.
- Lanius, R. A., Brand, B., Vermetten, E., Frewen, P. A., & Spiegel, D. (2012). The dissociative subtype of posttraumatic stress disorder: rationale, clinical and neurobiological evidence, and implications. *Depress & Anxiety*, 29, 701–708.
- McFarlane, A. C. (2013). Biology not culture explains dissociation in posttraumatic stress disorder. *Biological Psychiatry*, 73, 296–297.
- Moleman, N., van der Hart, O., & van der Kolk, B. A. (1992). Dissociation and hypnotizability in posttraumatic stress disorder. *Journal of Nervous and Mental Disease*, 180, 271–272.
- Olsen, S. A., Clapp, J. D., Parra, G. R., & Beck, J. B. (2013). Factor structure of the Dissociative Experiences Scale: an examination across sexual assault status. *Journal of Psychopathology & Behavioral Assessment*, 35, 394–403. doi:10.1007/s10862-013-9347-4
- Ozer, E. J., Best, S. R., Lipsey, T. L., & Weiss, D. S. (2003). Predictors of posttraumatic stress disorder and symptoms in adults: a meta-analysis. *Psychological Bulletin*, 129, 52–73.
- Putman, F. W. (1989). Pierre Janet and modern views of dissociation. *Journal of Traumatic Stress*, 2, 413–429. doi:10.1002/jts.2490020406
- Ray, W. J., & Faith, M. (1995). Dissociative experiences in a college age population: follow-up with 1190 subjects. *Personality and Individual Differences*, 18, 223–230.
- Reddy, M. K., Anderson, B. J., Liebschutze, J., & Steind, M. D. (2013). Factor structure of PTSD symptoms in opioid-dependent patients rating their overall trauma history. *Drug and Alcohol Dependence*, 132, 597–602.
- Resick, P. A., Suvak, M. K., Johnides, B. D., Mitchell, K. S., Iverson, K. M. (2012). The impact of dissociation on PTSD treatment with cognitive processing therapy [published online ahead of print March 30]. *Depress Anxiety*. http://dx.doi.org/10.1002/da.21938
- Ross, C. A., Joshi, S., & Currie, R. (1991). Dissociative experiences in the general population: identification of three factors. *Hospital & Community Psychiatry*, 42, 297–300.
- Ross, C. A., Ellason, J. W., & Anderson, G. (1995). A factor analysis of the Dissociative Experiences Scale in dissociative identity disorder. *Dissociation*, 8, 229–235.

- Ruiz, M., Poythress, N. G., Lilienfeld, S. O., & Douglass, K. S. (2008). Factor structure and correlates of the Dissociative Experiences Scale in a large offender sample. *Assessment*, 15, 511–521.
- Sanders, A., & Green, G. A. (1994). The factor structure of the dissociative experiences scale in college students. *Dissociation*, 7, 23–27.
- Satorra, A., & Bentler, E. M. (1994). Corrections to test statistics and standard errors in covariance structure analysis. In A. von Eye & C. C. Clogg (Eds.), *Latent variables analysis: Applications for developmental research* (pp. 399–419). Thousand Oaks: Sage.
- Semage, S. N., Sivayogan, S., Forbes, D., O'Donnell, M., Monaragala, R. M., Lockwood, E., & Dunt, D. (2013). Cross-cultural and factorial validity of PTSD check list-military version (PCL-M) in Sinhalese language. *European Journal of Psychotraumatology*, 4. doi:10.3402/ejpt.v4i0.19707.
- Simms, L. J., Watson, D., & Doebbeling, B. N. (2002). Confirmatory factor analyses of posttraumatic stress symptoms in deployed and nondeployed veterans of the Gulf War. *Journal of Abnormal Psychology*, 111(4), 637–647. doi:10.1037//0021-843x.111.4.637
- Spiegel, D. (1991). Neurophysiological correlates of hypnosis and dissociation. *Journal of Neuropsychiatry*, 3, 440–445.
- Spiegel, D., Loewenstein, R. J., Lewis-Fernandez, R., Sar, V., Simeon, D., Vermetten, E., et al. (2011). Dissociative disorders in DSM-5. *Depression and Anxiety*, 28, E17–E45.
- Stein, D. J., Koenen, K. C., Friedman, M. J., Hill, E., McLaughlin, K. A., Petukhova, M., et al. (2013). Dissociation in posttraumatic stress disorder: evidence from the world mental health surveys. *Biological Psychiatry*, 15, 302–312. doi:10.1016/j.biopsych.2012.08.022
- Steuwe, C., Lanius, R. A., & Frewen, P. A. (2012). Evidence for a dissociative subtype of PTSD by latent profile and confirmatory factor analyses in a civilian sample. *Depression and Anxiety*, 29, 689–700. doi:10.1002/da.21944
- Stockdale, G. D., Gridley, B. E., Balogh, D. W., & Holtgraves, T. (2002). Confirmatory factor analysis of single-and multiple-factor competing models of the dissociative experiences scale in a non-clinical sample. *Assessment*, 9, 94–106.
- Tichenor, V., Marmar, C. R., Weiss, D. S., Metzler, T. J., & Ronfeldt, H. M. (1996). The relationship of peritraumatic dissociation and post-traumatic stress: findings in female Vietnam theater veterans. *Journal of Consulting and Clinical Psychology*, 64, 1054–1059.
- van der Kolk, B. A., & van der Hart, O. (1989). Pierre Janet and the breakdown of adaptation in psychological trauma. *American Journal of Psychiatry*, 146, 1530–1540.
- van IJzendoorn, M. H., & Schuengel, C. (1996). The measurement of dissociation in normal and clinical populations: meta-analytic validation of the Dissociative Experiences scale (DES). *Clinical Psychology Review*, 16, 365–382.
- Wang, L., Cao, C., Wang, R., Zhanga, J., & Li, Z. (2012). The dimensionality of PTSD symptoms and their relationship to health-related quality of life in Chinese earthquake survivors. *Journal of Anxiety Disorders*, 26, 711–718.
- Wang, M., Armour, C., Li, X., Dai, X., Zu, X., & Yao, S. (2013). Further evidence for a five-factor model of PTSD: factorial invariance across gender in Chinese earthquake survivors. *Journal of Nervous and Mental Disease*, 201, 145–152.
- Wolf, E. J., Miller, M. W., Reardon, A. F., Ryabchenko, K. A., Castillo, D., & Freund, R. (2012). A latent class analysis of dissociation and posttraumatic stress disorder evidence for a dissociative subtype. *Archives of General Psychiatry*, 69, 698–705.
- Wolf, E. J., Miller, M. W., Reardon, A. F., Ryabchenko, K. A., Castillo, D., Freund, R. (2014). A latent class analysis of dissociation and PTSD: evidence for a dissociative subtype. *Archives of General Psychiatry* (in press)
- Yufik, T., & Simms, L. J. (2010). A meta-analytic investigation of the structure of posttraumatic stress disorder symptoms. *Journal of Abnormal Psychology*, 119(4), 764–776. doi:10.1037/a0020981
- Zoellner, L. A., Pruitt, L. D., Farach, F. J., & Jun, J. J. (2014). Understanding heterogeneity in PTSD: fear, dysphoria, and distress. *Depression & Anxiety*, 31, 97–106.