

Problem Set 7

You are strongly encouraged to solve the following exercises before next week's tutorial:

- Starting on page 362 (end of Chapter 9): 12 (note that the test is a GLRT and not an LRT as stated in the question), 13 (a-c), 23, 26 (a-e, use “greater than or equal to” instead of “less than or equal to” in (a)) and 28.
- Starting on page 459 (end of Chapter 11): 1 (d-g), 11 and 21 (a, also test the assumption that the variances of the two groups are equal at the 5% level).

Additional Exercise:

Let $X_1, \dots, X_n \stackrel{\text{i.i.d.}}{\sim} \mathcal{N}(\mu, \sigma^2)$, and consider the problem of testing

$$\begin{cases} \mathcal{H}_0 : \sigma^2 = \sigma_0^2 \\ \mathcal{H}_1 : \sigma^2 \neq \sigma_0^2 \end{cases}$$

at level α .

- Calculate the generalized likelihood ratio statistic $\Lambda(\underline{X})$.
- Denote $\mathcal{X}^2 = \frac{1}{\sigma_0^2} \sum_{i=1}^n (X_i - \bar{X})^2$. Express $\Lambda(\underline{X})$ in terms of \mathcal{X}^2 .
- Show that Λ , as a function of \mathcal{X}^2 , has a single minimum. Conclude that $\Lambda \geq c \iff \mathcal{X}^2 \leq c_1$ or $\mathcal{X}^2 \geq c_2$ for some constants c_1 and c_2 .
- What is the distribution of \mathcal{X}^2 under \mathcal{H}_0 ? Find a test at level α .

Solution:

- The likelihood is of course

$$\mathcal{L}(\mu, \sigma^2) = (2\pi\sigma^2)^{-n/2} \exp \left\{ -\frac{1}{2\sigma^2} \sum_{i=1}^n (X_i - \mu)^2 \right\},$$

and under no constraints, the MLEs are $\hat{\mu} = \bar{X}$ and $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i - \bar{X})^2$, hence

$$\mathcal{L}(\hat{\mu}, \hat{\sigma}^2) = (2\pi e \hat{\sigma}^2)^{-n/2}.$$

Under \mathcal{H}_0 , the MLE for μ remains unchanged, and obviously $\hat{\sigma}_0^2 = \sigma_0^2$, hence

$$\mathcal{L}(\hat{\mu}_0, \hat{\sigma}_0^2) = (2\pi \sigma_0^2)^{-n/2} \exp \left\{ -\frac{1}{2\sigma_0^2} \sum_{i=1}^n (X_i - \bar{X})^2 \right\}.$$

Finally,

$$\Lambda(\underline{X}) = \frac{\mathcal{L}(\hat{\mu}, \hat{\sigma}^2)}{\mathcal{L}(\hat{\mu}_0, \hat{\sigma}_0^2)} = \left(\frac{\sigma_0^2}{e \hat{\sigma}^2} \right)^{n/2} \exp \left\{ \frac{1}{2\sigma_0^2} \sum_{i=1}^n (X_i - \bar{X})^2 \right\}.$$

(b) Since $\hat{\sigma}^2 = \frac{\sigma_0^2 \mathcal{X}^2}{n}$, we can rewrite

$$\Lambda = \left(\frac{n}{\mathcal{X}^2} \right)^{n/2} \exp \left\{ \frac{\mathcal{X}^2 - n}{2} \right\}.$$

(c) Write $\Lambda(t) = n^{n/2} t^{-n/2} e^{\frac{t-n}{2}}$, then it is easy to verify that

$$\lim_{t \rightarrow 0} \Lambda(t) = \lim_{t \rightarrow \infty} \Lambda(t) = \infty \quad (1)$$

and that

why is this: find the shape of distribution of Lambda

$$\Lambda'(t) = \frac{n^{n/2}}{2} t^{-\frac{n+1}{2}} (t - n) e^{\frac{t-n}{2}}.$$

The function then has a single extremum at $t = n$, and using (1) we know that its shape is as in Figure 1. From the figure it becomes clear that the rejection region of the GLRT

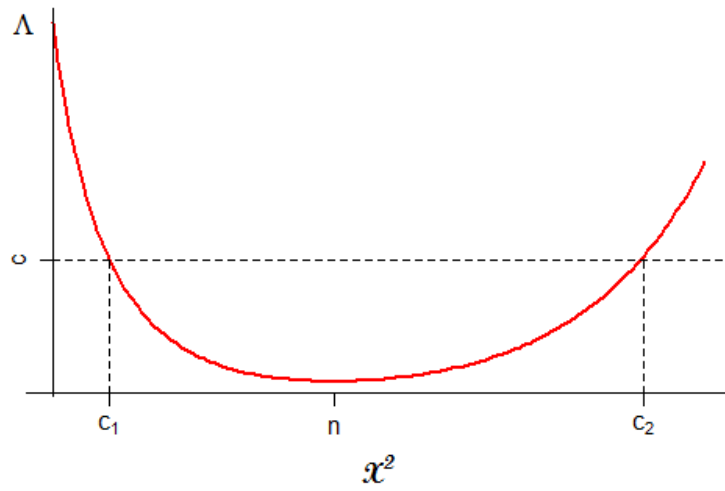


Figure 1: Λ as a function of χ^2 .

is

$$\mathcal{C} = \left\{ \Lambda(\underline{X}) \geq c \right\} = \left\{ \chi^2 \leq c_1 \right\} \cup \left\{ \chi^2 \geq c_2 \right\}.$$

(d) Clearly $\chi^2 \stackrel{\mathcal{H}_0}{\sim} \chi^2_{n-1}$, hence a rejection region of a test at level α would be –

$$\mathcal{C} = \left\{ \chi^2 \leq \chi^2_{n-1, \alpha/2} \right\} \cup \left\{ \chi^2 \geq \chi^2_{n-1, 1-\alpha/2} \right\}.$$