Chapter 1 Linear Regression with One Predictor Variable

1.3 Simple Linear Regression Model with Distribution of Error Unspecified

Definition. Simple Linear Model A model that is linear in simple (1 predictor variable) and linear in parameters and linear in predictor variable. A model linear in parameter and predictor variable is called **first-order model**

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$$

where

- 1. Y_i is value of response variable and X_i is the value of predictor variable in the i-th trial
- 2. ϵ_i is a random error term with mean $\mathbb{E}[\epsilon_i] = 0$ and $\sigma^2[\epsilon_i] = \epsilon^2$. Also ϵ_i and ϵ_j are uncorrelated $(\sigma^2[\epsilon_1], \epsilon_j = 0$ for all $i \neq j$)
- 3. regression coefficients β_1 is slope of the regression line, indicating change in mean of probability distribution of Y per unit increase in X. β_0 is the Y intercept of the regression line

Properties

- 1. Y_i is a random variable, a summation of a constant $\beta_0 + \beta_1 X_i$ and the random error ϵ_i .
- 2. Distribution of Y_i

(a) By
$$\mathbb{E}[\epsilon_i] = 0$$

$$\mathbb{E}[Y_i] = \mathbb{E}[\beta_0 + \beta_1 X_i + \epsilon_i] = \beta_0 + \beta_1 X_i$$

(b) By
$$\sigma^2\{\epsilon_i\} = \sigma^2$$

$$\sigma^2\{Y_i\} = \sigma^2\{\epsilon_i\} = \sigma^2$$

3. The **regression function** relates mean of probability distribution of Y for given S to level of X

$$\mathbb{E}[Y] = \beta_0 + \beta_1 X$$

4. Y_i and Y_j are uncorrelated since errors are uncorrelated

In summary this regression model implie reseponse Y_i come from probability distribution whose means are $\mathbb{E}\{Y\} = \beta_0 + \beta_1 X_i$ whose variances are σ^2 (same for all levels of X), furthermore, two responses Y_i and Y_j are uncorrelated

1.6 Estimation of Regression Function

Definition. Method of Least Squares is used to estimate regression parameters β_0 and β_1 . The MLS considers sum of n squared deviations of Y_i from its expected value

$$RSS = \sum_{i=1}^{n} (y_i - \beta_0 - \beta_1 y_i)^2$$

estimators of β_0 and β_1 are values b_0 and b_1 that minimize Q given sample observations $(x_1, y_1), \cdots$. By taking partials of RSS and set it to zero, we derive a pair of **Normal Equation**

$$\sum y_i = nb_0 + b_1 \sum x_i$$
$$\sum x_i y_i = b_0 \sum x_i + b_1 \sum x_i^2$$

Solving it we get LS estimator

$$\hat{\beta}_1 = \frac{\sum (x_i - \overline{x})(y_i - \overline{y})}{\sum (x_i - \overline{x})^2} = \frac{S_{XY}}{S_{XX}}$$
$$\hat{\beta}_0 = \overline{y} - \hat{\beta}_1 \overline{x}$$

Properties of LS regression coefficient mostly from derivation of LS estimators

$$\sum_{i=1}^{n} e_{i} = 0 (from \frac{\partial RSS}{\beta_{0}}) \qquad \sum_{i=1}^{n} e_{i}^{2} \quad minimized (from LSE)$$

$$\sum y_{i} = \sum \hat{y}_{i} (since \sum_{i=1}^{n} e_{i} = 0) \qquad \sum x_{i}e_{i} = 0 (from \frac{\partial RSS}{\beta_{1}}) \qquad \sum \hat{y}_{i}e_{i} = 0$$

$$\overline{y} = \hat{\beta}_{0} + \hat{\beta}_{1}\overline{x}$$

Definition. Gauss Markov Theorem Under conditions of regression model, the least squares estimator $\hat{\beta}_0$ and $\hat{\beta}_1$ are unbiased and have minimum variance among all unbiased linear estimators

Proof. 1. For $\hat{\beta}_1$

$$\hat{\beta}_1 = \sum c_i y_i$$
 where c_i is arbitrary

Now we prove its unbiased

$$\mathbb{E}[\hat{\beta}_1] = \mathbb{E}\{\sum c_i y_i\} = \sum c_i \mathbb{E}[Y_i] = \sum c_i (\beta_0 + \beta_1 x_i) = \beta_0 \sum c_i + \beta_1 \sum c_i x_i = \beta_1$$

given the restriction that

$$\sum c_i = 0 \quad \sum c_i x_i = 1$$

which holds for both $\hat{\beta}_1$ and $\hat{\beta}_0$...

Definition. Point estimation of mean response Given regression function

$$\mathbb{E}[Y] = \beta_0 + \beta_1 X$$

so we have a estimated regression line

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$$

where \hat{y}_i is value of estimated regression level at level x, it is a point estimate of the mean response when the level of the predictor variable is X. By the previous theorem, \hat{y} is an unbiased estimator of $\mathbb{E}[Y]$ with minimum variance in the class of unbiased linear estimators.

Definition. residual the i-th residual is the difference between the observed value y_i and the corresponding fitted value \hat{y}_i .

$$e_i = y_i - \hat{y}_i$$

In the case of simple linear model, we have

$$e_i = y_i - \hat{\beta}_0 - \hat{\beta}_1 x_i$$

Informally, it is the vertical distance of y_i from the fitted value \hat{y}_i on the estimated regression line, which is known.

1.7 Estimation of Error terms variance σ^2

Definition. Point Estimator of σ^2

1. Single Population We use mean square s^2 to estimate population variance σ^2

$$s^2 = \frac{1}{n-1} \sum_{i=1}^n (y_i - \overline{y})^2$$

The lost degree of freedom comes with using \overline{y} to estimate mean

- 2. Regression Model Note y_i comes with difference probability distribution based on levels of x_i . So to calculate sum of squared deviation, we have to calculate deviation around its own estimated mean \hat{y}_i .
 - (a) Residual/Error sum of square (RSS, SSE)

$$RSS/SSE = \sum_{i} (y_i - \hat{y}_i)^2 = \sum_{i} e_i^2$$

(b) Residual/Error Mean Square (MSE)

$$s^2 = \frac{RSS}{n-2} = \frac{\sum e_i^2}{n-2}$$

The loss of 2 degree of freedom comes from using $\hat{\beta}_0$ and $\hat{\beta}_1$ to estimate regression coefficient to get the \hat{y}_i

(c) MSE is an unbiased estimator of σ^2

1.8 Normal Error Regression Model

Note. Motivation Least squared method provides unbiased point estimator for β_0 and β_1 regardless of the distribution of ϵ_i (and hence of Y_i). However, need to make assumption about form of distribution of ϵ_i to set up **interval estimate** and make tests.

Definition. Normal Error Regression Model

$$Y_i = \beta_0 + \beta_1 X_i + \epsilon_i$$

Additionally, $\epsilon_i \stackrel{i.i.d.}{\sim} \mathcal{N}(0, \sigma^2)$ for $i = 1, \dots, n$.

Properties

- 1. Note independence of ϵ_i comes from the uncorrelatedness of ϵ_i and properties of normal distribution
- 2. Y_i are independent normal random variable
- 3. ϵ_i being normal is somewhat justified as it represent all factors which tend to comply with CLT and cause error distribution approach normality as number of factor effects becomes large

Definition. Parameter (β_0 , β_1 , σ^2) estimation by Method of maximum likelihood Turns out $\hat{\beta}_0$ and $\hat{\beta}_1$ are same as least squared estimator. The estimator for σ^2 is different however

$$\hat{\sigma}^2 = \frac{1}{n} \sum_{i} (y_i - \hat{y}_i)^2$$

estimator $\hat{\sigma}^2$ is biased with following relationship to mean squre error

$$s^2 = MSE = \frac{n}{n-2}\hat{\sigma}^2$$