## An example where the MLE does not follow an asymptotically normal distribution

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Let  $X_1, \ldots, X_n$  be a sample from a truncated exponential distribution with a density

$$f(x|\theta) = \begin{cases} e^{-(x-\theta)} & x > \theta \\ 0 & \text{otherwise} \end{cases}$$

- 1. find the MLE for  $\theta$
- 2. show that  $n(\hat{\theta} \theta)$  is exponentially distributed
- 3. find the asymptotic distribution of  $\sqrt{n}(\hat{\theta} \theta)$
- 4. why does the asymptotic normality of the MLEs not hold in this case?

## **Solution:**

1. Writing

$$\mathcal{L}(\theta) = \prod_{i=1}^{n} e^{-(x_i - \theta)} I_{[\theta, \infty]}(x_i) = \begin{cases} e^{-n\overline{x}} e^{n\theta} & x_{\min} > \theta \\ 0 & \text{otherwise} \end{cases}$$

we see that for a given sample (fixed  $\overline{x}$ ) the likelihood is monotonically increasing in  $\theta$  before being truncated at  $x_{\min}$ . Figure 1 makes it clear that the MLE is  $\widehat{\theta} = X_{\min}$ .

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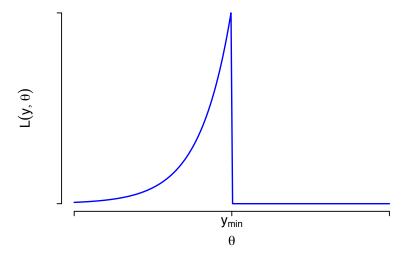


Figure 1: A sketch of the likelihood

2. First, for  $x > \theta$ 

$$F(x|\theta) = \mathbb{P}(X \le x) = \int_{\theta}^{x} e^{-(t-\theta)} dt = 1 - e^{-(t-\theta)}.$$

Now,

$$\mathbb{P}\left(n(\widehat{\theta} - \theta) \le t\right) = \mathbb{P}\left(X_{\min} \le \frac{t}{n} + \theta\right)$$

$$= 1 - \mathbb{P}\left(X_{\min} > \frac{t}{n} + \theta\right) = 1 - \prod_{i=1}^{n} \left[1 - F\left(\theta + \frac{t}{n}\right)\right]$$

$$= 1 - \left\{\begin{array}{c} \exp\left\{-\sum_{i=1}^{n} \left(\theta + \frac{t}{n} - \theta\right)\right\} & \theta + \frac{t}{n} > \theta\\ 0 & \text{otherwise} \end{array}\right.$$

$$= \left\{\begin{array}{c} 1 - e^{-t} & t > 0\\ 0 & \text{otherwise} \end{array}\right.$$

and we see that  $n(\widehat{\theta}-\theta) \sim \operatorname{Exp}(1)$  .

3. To derive the desired distribution, note that

$$\mathbb{P}\left(\sqrt{n}(\hat{\theta} - \theta) \le t\right) = \mathbb{P}\left(n(\hat{\theta} - \theta) \le \sqrt{n}t\right)$$

$$\stackrel{(2)}{=} \begin{cases} 1 - e^{-\sqrt{n}t} & \sqrt{n}t > 0 \\ 0 & \text{otherwise} \end{cases}$$

$$= \begin{cases} 1 - e^{-\sqrt{n}t} & t > 0 \\ 0 & \text{otherwise} \end{cases}$$

$$\xrightarrow{n \to \infty} \begin{cases} 1 & t > 0 \\ 0 & \text{otherwise} \end{cases},$$

therefore the cdf of  $\sqrt{n}(\hat{\theta}-\theta)$  approaches the cdf of the constant random variable 0 instead of that of a normal distribution.

4. Note that the likelihood is not even differentiable once at the MLE, deeming the proof of the asymptotic normality invalid.