

Homework 8 Solutions

Chapter 3:

68. Length of time to complete a job of a server has exponential pdf and cdf $F_X(x) = 1 - e^{-\lambda x}$. Call Y the waiting time until service of the first job. Its distribution is distribution of the 1st order statistics.

$$f_Y(y) = f_1(y) = n f_X(y) (1 - F_X(y))^{n-1} = n \lambda e^{-\lambda y} (1 - (1 - e^{-\lambda y}))^{n-1} = n \lambda e^{-\lambda y} e^{-\lambda y(n-1)} = n \lambda e^{-n\lambda y}, \quad y \geq 0$$

Call Z the waiting time until service of the second job. Its distribution is distribution of the 2nd order statistics.

$$\begin{aligned} f_Z(z) &= f_2(z) = n(n-1) f_X(z) (F_X(z))^{2-1} (1 - F_X(z))^{n-2} = n(n-1) \lambda e^{-\lambda z} (1 - e^{-\lambda z}) (1 - (1 - e^{-\lambda z}))^{n-2} \\ &= n(n-1) \lambda e^{-(n-1)\lambda z} (1 - e^{-\lambda z}), \quad y \geq 0 \end{aligned}$$

Chapter 4:

6. Let X continuous random variable with pdf $f_X(x) = 2x, 0 \leq x \leq 1$

a. Find $E(X)$:

$$\text{Ans: } E(X) = \int_{-\infty}^{\infty} x f_X(x) dx = \int_0^1 x 2x dx = \int_0^1 2x^2 dx = \left[\frac{2x^3}{3} \right]_{x=0}^1 = \frac{2}{3}$$

b. Let $Y = X^2$. Find pdf of Y and use it to find $E(Y)$.

$$\text{Ans: } Y = X^2 \text{ so } x = g^{-1}(y) = \sqrt{y}, \quad \frac{d}{dy} g^{-1}(y) = \frac{d}{dy} \sqrt{y} = \frac{1}{2\sqrt{y}}$$

$$\text{Hence } f_Y(y) = f_X(g^{-1}(y)) \left| \frac{d}{dy} g^{-1}(y) \right| = 2\sqrt{y} \frac{1}{2\sqrt{y}} = 1, 0 \leq y \leq 1 \text{ So}$$

$$E(Y) = \int_{-\infty}^{\infty} x f_Y(x) dx = \int_0^1 x dx = \left[\frac{x^2}{2} \right]_{x=0}^1 = \frac{1}{2}$$

$$\text{c. } E(X^2) = \int_{-\infty}^{\infty} x^2 f_X(x) dx = \int_0^1 x^2 2x dx = \int_0^1 2x^3 dx = \left[\frac{2x^4}{4} \right]_{x=0}^1 = \frac{1}{2}$$

b. and c. give the same answer.

$$\text{d. } \text{Var}(X) = E(X^2) - (E(X))^2 = \frac{1}{2} - \left(\frac{2}{3}\right)^2 = \frac{1}{2} - \frac{4}{9} = \frac{1}{18}$$

29. Prove that $E(g(X)h(Y)) = E(g(X))E(h(Y))$

Ans: Continuous RVs

$$\begin{aligned} E(g(X)h(Y)) &= \iint_{-\infty}^{\infty} g(x)h(y) f_{X,Y}(x,y) dx dy = \iint_{-\infty}^{\infty} g(x)h(y) f_X(x) f_Y(y) dx dy = \\ &= \int_{-\infty}^{\infty} h(y) f_Y(y) \int_{-\infty}^{\infty} g(x) f_X(x) dx dy = \int_{-\infty}^{\infty} h(y) f_Y(y) E(g(X)) dy = E(g(X)) \int_{-\infty}^{\infty} h(y) f_Y(y) dy = \\ &= E(g(X))E(h(Y)) \end{aligned}$$

The second equality come from X,Y independent

30. Find $E\left(\frac{X}{X+1}\right)$ where X is a Poisson random variable.

Ans: X is Poisson random variable so pmf of X is: $p_X(k) = \frac{\lambda^k}{k!} e^{-\lambda}, \quad k=0,1,2,\dots$

$$\begin{aligned} E\left(\frac{1}{X+1}\right) &= \sum_{k=0}^{\infty} \frac{1}{k+1} p_X(k) = \sum_{k=0}^{\infty} \frac{1}{k+1} \frac{\lambda^k}{k!} e^{-\lambda} = \sum_{k=0}^{\infty} \frac{\lambda^k}{(k+1)!} e^{-\lambda} = \frac{1}{\lambda} \sum_{k=0}^{\infty} \frac{\lambda^{k+1}}{(k+1)!} e^{-\lambda} = \frac{1}{\lambda} \sum_{l=1}^{\infty} \frac{\lambda^l}{l!} e^{-\lambda} = \\ &= \frac{1}{\lambda} (\sum_{l=0}^{\infty} \frac{\lambda^l}{l!} e^{-\lambda} - e^{-\lambda}) = \frac{1}{\lambda} (1 - e^{-\lambda}) \end{aligned}$$

The last equality come from “total sum of any pdf (here Poisson) is 1”.