The problem has existed over endless years: Racialized difference in commuting, 1980–2019

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"The problem has existed over endless years..."

- Quote from Dr. Martin Luther King Jr. about discrimination faced by Black bus riders, made during the Montgomery Bus Boycott (1955).
- Plessy v. Ferguson (1896), which legitimized 'separate but equal', was about segregation on trains.
- ⇒ Racialized difference in transportation is pervasive throughout US history

This paper: Provide a comprehensive accounting of racialized differences in commuting from 1980 to 2019

- Update prior literature in economics and sociology and study trend
- Decompose differences: mode, location, observables
- Study aggregate (city-level) determinants
- Suggest an interpretation

Summary of Findings

- 1. Black commuters face longer commutes than White commuters
 - Difference declining since 1980, but still persists (even conditionally)
- 2. Amongst transit users, the difference has not decreased or has increased
 - Suggests partial convergence due largely to car adoption by Black commuters
 - But there is still a small but significant difference for car commuters
- 3. Within-city res. location (PUMA) does not account for much of the difference
- 4. Difference largest at lower incomes, but are present at high incomes too
- 5. While differences have mostly shrunk since 1980, they persist in:
 - Large, congested cities that have experienced rapid house price growth.
 - IV to show high housing costs drive at least part of the residual difference
 - Evidence supports increasing spatial stratification on commute times in \$\$\$ cities

Data

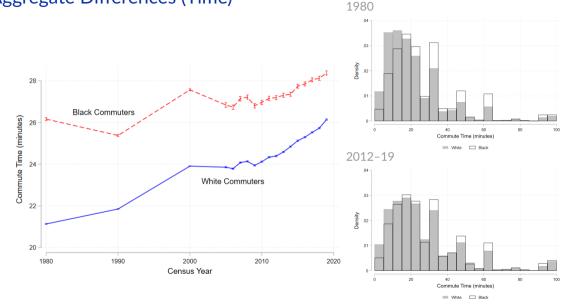
Census/ACS, 1980–2019, for all commuters

- Journey to Work questions ask about race and commute time/mode
- Assign to consistent commuting zones (CZs) (Autor & Dorn '13)
 - Lightly modify to bring together large markets, e.g., DFW, NYC/Newark
- ▶ Often focus on year bins: 1980, 1990, 2000, 2005–11, 2012–19
- Extend back to 1960 for aggregate mode share

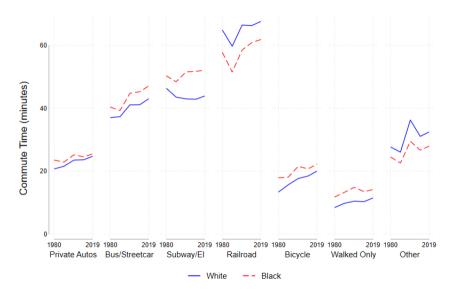
Additional sources

- ► NHGIS for finer geographic aggregates
- ▶ Zip Code Business Patterns for spatial dist. of work locations

Aggregate Differences (Time)



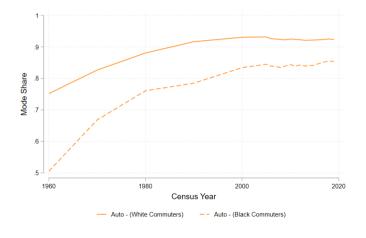
Aggregate Differences (Time by Mode)



Aggregate Differences (Mode)

Major increases in auto commuting

- Primarily at the expense of Bus/Streetcar use by Black commuters
- Also substantial reduction of Walking for all commuters



Analytic Framework

Use regression compatible decomposition approach (Fortin '08, Fortin et al. '11)

$$ln(\tau_{ict}) = \beta_t 1[Black_{ict}] + x'_{ict}\mu_t + \lambda_{ct} + u_{ict}$$

Group obs. covariates into categories:

- ► CZ
- Demographics/Educ
- Transportation Mode
- Job/Income

Permit time-varying coeffs

Can estimate β_{ct} (at city-level)!

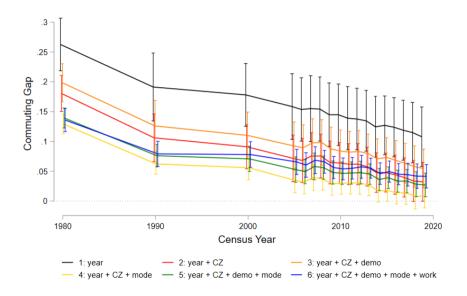
CAUTION: Intepretation

- What does it mean to control for e.g., income?
- Discrimination or structural racism could drive income differences
- Interpret as potential mechanisms

CAUTION: Selection

Many margins of selection, esp: LFP, mode, and job

Baseline Results



Decomposition

Decompose along observable dimensions (non-sequentially a la Gelbach '16)

	$\Delta_{\{t\}}$	$\Delta^{Unexplained}_{\{t\}}$	$\Delta^{Explained}_{\{t\}}$			
			$\Delta^{Demog.}_{\{t\}}$	$\Delta^{Tr.Mode}_{\{t\}}$	$\Delta^{Work/Inc.}_{\{t\}}$	$\Delta^{CZ}_{\{t\}}$
$1[Black] imes t_{1980}$	0.263*** (0.022)	0.136*** (0.010)	-0.008*** (0.000)	0.073*** (0.016)	-0.001 (0.002)	0.062***
	(0.022)	51.7%	-3.0%	27.8%	-0.2%	23.7%
$1[Black] imes t_{1990}$	0.191***	0.079***	-0.009***	0.063***	-0.007***	0.065***
2000	(0.029)	(0.011)	(0.000)	(0.018)	(0.002)	(0.009)
		41.4%	-5.0%	32.9%	-3.4%	34.0%
$1[Black] imes t_{2000}$	0.178***	0.078***	-0.008***	0.050***	-0.011***	0.069***
	(0.027)	(0.011)	(0.000)	(0.013)	(0.002)	(0.009)
		43.9%	-4.6%	28.1%	-6.3%	39.0%
$1[Black] imes t_{2005-11}$	0.150***	0.061***	-0.009***	0.049***	-0.014***	0.063***
	(0.027)	(0.010)	(0.000)	(0.014)	(0.002)	(0.009)
		40.5%	-6.1%	33.0%	-9.5%	42.1%
$1[Black] imes t_{2012-19}$	0.124***	0.049***	-0.008***	0.040***	-0.019***	0.063***
	(0.025)	(0.009)	(0.000)	(0.012)	(0.001)	(0.010)
		39.1%	-6.6%	32.5%	-15.4%	50.4%

N = 46.8 mil.

Decomposition

Contribution of **location (CZ)** to difference is ~ constant (6–7 log points)

Makes up a larger share as unconditional difference shrinks

Demographics and **job/income** do not play a huge role

- Typically increases difference!
- Job/income playing a larger role over time

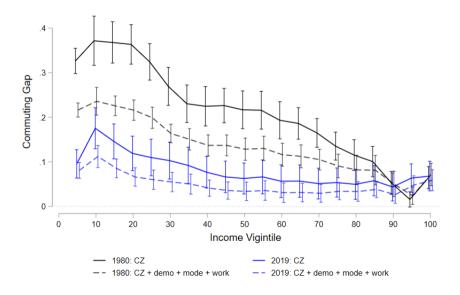
Transportation mode makes up a roughly constant share of difference

One-quarter of decline is explained by partial convergence in mode share

Share of difference unexplained by our observables falling

► Still about 39% of unconditional difference

Heterogeneity by Income



Additional Aggregate Specs

Differences conditional on mode

- Car: Strongest evidence of partial convergence Car
- ▶ Bus & Subway: Difference flat or growing (~ 10 log points) Bus Subway

With PUMA FEs (2000 and later, might extend to 1990)

- Makes surprisingly little difference (except for subway)
- Meso-scale residential sorting doesn't explain difference

By city type

- Differences larger in bigger cities w/ heavy rail (even for auto)
- ► An exception: walkers, much larger differences in non-large non-transit CZs

City-level Heterogeneity

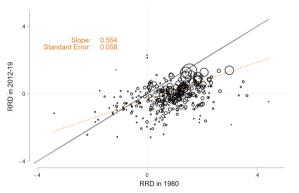
What drives city-level variation in this difference:

$$\begin{split} \ln(\tau_{ict}) &= \beta_{ct} 1 [\mathsf{Black}_{ict}] + x_{ict}' \mu_{ct} + \lambda_{ct} + u_{ict} \\ \hat{\beta}_{ct} &= z_{ct}' \gamma + D_c + T_t + e_{ct} \end{split}$$

We term $\hat{\beta}_{ct}$ the residual racialized difference (RRD) in commute time

- ▶ RRD can be shown to contribute to $\Delta_{\{t\}}^{\text{Unexplained}}$
- ightharpoonup Dealing with generated β and heteroskedasticity
 - Drop CZs with <1k commuters, CZs with <50 unique Black commuter Census respondents
 - Weight second stage by number unique Black commuter Census respondents (could also weight by robust variance estimate of $\hat{\beta}_{ct}$
 - Cluster SEs by CZ

Summary and Persistence

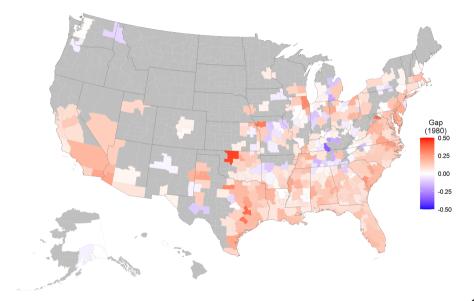


	Mean	SD	Min	Max
1980	0.131	0.072	-0.339	0.442
1990	0.070	0.072	-0.326	0.246
2000	0.068	0.077	-0.412	0.247
2005-11	0.053	0.073	-0.384	0.220
2012-19	0.036	0.070	-0.257	0.230
$N_t = 339$				

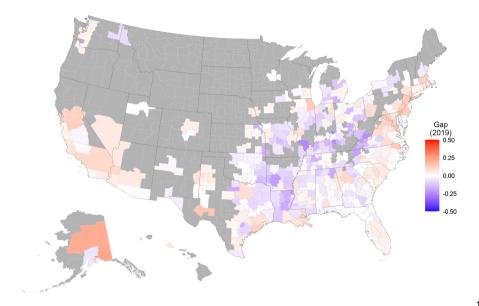
Declining mean, but not much decline in SD

▶ Relatively high but not uniform persistence over 40-year interval

Map (1980)



Map (2012-19)



Correlates

	1980		20	000	2012-19		
	(1)	(2)	(3)	(4)	(5)	(6)	
Ln(Population)	0.015*** (0.004)	0.022*** (0.005)	0.028*** (0.003)	0.034*** (0.003)	0.029*** (0.002)	0.032*** (0.002)	
% Black		0.318*** (0.050)		0.253*** (0.038)		0.123*** (0.026)	
Obs. R-sq	339 0.131	339 0.264	339 0.380	339 0.475	339 0.545	339 0.581	

- Population playing an increasingly large role in RRD
- ▶ Black share of population playing a smaller role
- ▶ Large (& often coastal cities) cities see much of the RRD ⇒ Housing Costs?

Housing Prices

- 1. Access is a somewhat persistent 'second-nature' n'hood amenity (e.g., Cronon '91)
- 2. Big, expensive cities features lots of variation in job access
 - Geography lurking in the background (Saiz '10; Lee & Lin '18; Saiz & Wang '21)
- 3. Inelastic supply likely binds first in \$\$\$ places → increased stratification (Van Nieuwerburgh & Weill '10; Guerrieri, Hartley, Hurst '13; Gyourko, Mayer, Sinai '13)

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Test link between rising housing prices and stratification

- Estimate panel (TWFE) model of RRD on housing prices
- ► Use Guren et al. '21-type IV (like Saiz elasticity) Description
- → High housing costs play a role at preserving RRD

	0	LS	ľ	V	
Ln(Hous. Pr.)	0.057** (0.013)	0.058** (0.013)	0.043+ (0.025)	0.044+ (0.024)	
Perc. Black		0.059 (0.125)		0.046 (0.123)	
First Stage			0.583** (0.096)	0.573** (0.099)	
F-stat (CD) F-stat (KP) Obs.	1695	1695	1070 37 1695	1039 34 1695	

Other Aggregate Measures

Estimate additional correlations

- Cross section: with and without population control
- ► Panel: with TWFEs (CZ & year-bin), with and without population control

Suggestive evidence that

- Larger difference persists in transit-heavy, longer-commute (slower) CZs
- Urban form, segregation, daytime vs. nighttime population variables only matter insomuch as they are correlates of population
- Places with increasingly correlated (Expensive Housing, Short Commute) see higher RRD (supports stratification)

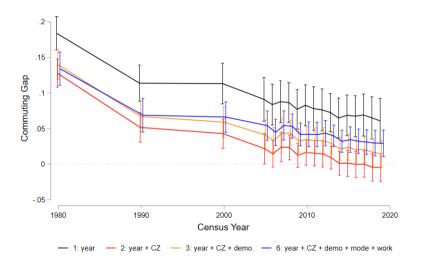


To Do - Thoughts and Comments?!

- 1. Additional aggregate city correlates of RRD
 - Race-specific urban form measures
- 2. Stratify analysis by city size/history
 - E.g., Large vs. Medium CZs, Old vs. New CZs.
- 3. How important is selection?
 - Selection into employment (not commute times if not)
 - Selection in mode, etc.

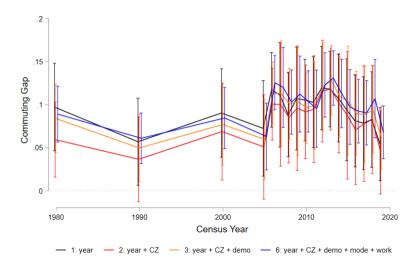
Thank you!

Baseline Results - Car



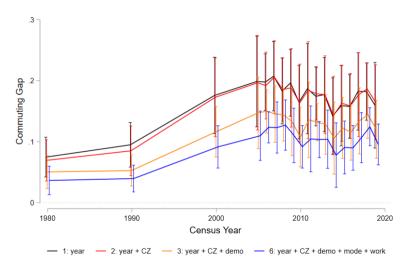


Baseline Results - Bus





Baseline Results - Subway





Details of Housing IV

Confounders, e.g.: localized land use regs, prod. shocks to clustered industries

Estimate division-level variant of IV in Guren et al. '21

$$P_{cdt} = \delta_c \bar{P}_{(-c)dt} + \psi m_{cdt} + \phi_c t + D_c + \epsilon_{cdt}$$

- $ightharpoonup P_{cdt}$ is log mean housing price in CZ c in Census division d in year-bin t
- $ightharpoonup ar{P}_{(-c)dt}$ is the leave-c-out log mean housing price in the Census division
- lacktriangledown $\hat{\delta}_car{P}_{(-c)dt}$ measures local response to reg. price movements o time-varying IV

Identification excludes the presence of any unobserved factor that

- i. is correlated with regional house price movements, and
- ii. differentially impacts the RRD according to housing price sensitivity



Other Aggregate Measures - Results

Dependent Variable: RRD

	Aggre	Aggregates		Commuting Chars.			Urban Form/Sorting			Comm. Strat.	
	Ln(Pop.)	% Black	% Transit	Ave. Car Time	Ave. Transit Time	Centr'	Dis- simi- larity	Hut- chens	Zip Empl. Gini	Corr(Time, Hous.)	%Hous. -%Time
Unconditio	nal										
Measure	0.070*** (0.018)	0.005 (0.128)	0.289+ (0.150)	0.006** (0.002)	0.001*** (0.000)	0.065 (0.142)	-0.069 (0.048)	-0.043 (0.067)	-0.001 (0.045)	-0.025* (0.010)	-0.094* (0.047)
Controlling	for Log Popul	ation									
Measure	,	0.037 (0.099)	0.287+ (0.156)	0.004* (0.002)	0.001** (0.000)	0.181 (0.121)	-0.021 (0.040)	0.012 (0.057)	0.009 (0.040)	-0.020* (0.009)	-0.079+ (0.046)

Return