

432 Homework 3 Answer Sketch

Due 2019-02-22. Version: 2019-02-22

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Setup and Data Ingest

```
library(skimr); library(broom); library(janitor)
library(leaps); library(modelr); library(tidyverse)

skim_with(numeric = list(hist = NULL),
          integer = list(hist = NULL))

hbp330 <- read_csv("data/hbp330.csv") %>%
  clean_names()
```

Question 1 (20 points)

Again, consider the `hbp330` data used in Homeworks 1 and 2. Build your best model for the prediction of body-mass index, considering the following 14 predictors: `practice`, `age`, `race`, `eth_hisp`, `sex`, `insurance`, `income`, `hsgrad`, `tobacco`, `depdiag`, `sbp`, `dbp`, `statin` and `bpmed`. Use an appropriate best subsets procedure to aid in your search, and use a cross-validation strategy to assess and compare potential models. Be sure to provide a written explanation of your conclusions and specify the variables in your final model, in complete sentences.

Preparing the data for `regsubsets`

As mentioned, in the answer sketch, we will use a complete cases analysis to deal with missing data.

To get the `regsubsets` function in the `leaps` package to do what we want, we will have to make sure that all of the multi-categorical predictors are expressed as factors (we can do this in a batch by changing all of the character variables to factor variables with `mutate_if`), and we'll need to drop all missing values (we could have first imputed them.)

```
hw3q1 <- hbp330 %>%
  mutate( bmi = weight / (height*height) ) %>%
  mutate_if(is.character, as.factor) %>%
  select(bmi, practice, age, race, eth_hisp,
         sex, insurance, income, hsgrad, tobacco,
         depdiag, sbp, dbp, statin, bpmed) %>%
  drop_na
```

Let's check to be sure everything is either a factor or numeric, and that we now have no missing values.

```
skim(hw3q1)
```

Skim summary statistics

```
n obs: 325
n variables: 15
```

```
-- Variable type:factor -----
variable missing complete  n n_unique
depdiag      0      325 325      2
eth_hisp     0      325 325      2
insurance    0      325 325      4
practice     0      325 325      2
race         0      325 325      4
sex          0      325 325      2
tobacco      0      325 325      3
      top_counts ordered
No: 211, Yes: 114, NA: 0  FALSE
No: 261, Yes: 64, NA: 0  FALSE
Med: 131, Med: 128, Com: 53, Uni: 13  FALSE
A: 176, B: 149, NA: 0  FALSE
Bla: 178, Whi: 131, Asi: 10, Mul: 6  FALSE
F: 201, M: 124, NA: 0  FALSE
nev: 138, for: 115, cur: 72, NA: 0  FALSE
```

```
-- Variable type:numeric -----
variable missing complete  n    mean    sd    p0    p25    p50
age          0      325 325    55.5    11.53   23     48     57
bmi          0      325 325    34.83    8.05   16.73   29.73   33.91
bpmed        0      325 325     0.66    0.48    0        0        1
dbp          0      325 325    74.73   10.24   41        68       74
hsgrad       0      325 325    81.92    8.55   57        75       81
income       0      325 325 35480.92 15901.56 6800   25600   30600
sbp          0      325 325   128.28   17.39   84       116     128
statin       0      325 325     0.7     0.46    0        0        1
p75          65
p100         77
39.22        64.04
1            1
82          106
89          100
```

```

42600    147400
  138      194
    1        1

```

Performing an exhaustive search with regsubsets

```

q1_preds <- with(hw3q1,
  cbind(practice, age, race, eth_hisp, sex,
        insurance, income, hsgrad, tobacco,
        depdiag, sbp, dbp, statin, bpmed))

q1_subs <- regsubsets(q1_preds, y = hw3q1$bmi, nvmax = 7)

q1_rs <- summary(q1_subs)

```

The `outmat` section of the summary output has the listing of fitted models that we want. Note that the multi-categorical `race` variables are either in or out, as a group, this way.

```

q1_rs$outmat

      practice age race eth_hisp sex insurance income hsgrad tobacco
1 ( 1 ) " "      " " " " " "      "*" " "      " "      " "      " "
2 ( 1 ) " "      "*" " " " "      "*" " "      " "      " "      " "
3 ( 1 ) " "      "*" " " " "      "*" " "      " "      " "      " "
4 ( 1 ) " "      "*" " " " "      "*" " "      " "      " "      "*"
5 ( 1 ) "*"      "*" "*" " "      "*" " "      " "      " "      " "
6 ( 1 ) "*"      "*" "*" " "      "*" " "      " "      " "      "*"
7 ( 1 ) "*"      "*" "*" " "      "*" " "      " "      " "      "*"

      depdiag sbp dbp statin bpmed
1 ( 1 ) " "      " " " " " "      " "
2 ( 1 ) " "      " " " " " "      " "
3 ( 1 ) " "      " " " " " "      "*"
4 ( 1 ) " "      " " " " " "      "*"
5 ( 1 ) " "      " " " " " "      "*"
6 ( 1 ) " "      " " " " " "      "*"
7 ( 1 ) "*"      " " " " " "      "*"

```

So, here are our “best subsets” models:

Inputs	Predictors Included (in addition to Intercept)
2	sex
3	sex, age
4	sex, age, bpmed
5	sex, age, bpmed, tobacco
6	sex, age, bpmed, practice, race
7	sex, age, bpmed, practice, race, tobacco
8	sex, age, bpmed, practice, race, tobacco, depdiag

```

round(q1_rs$adjr2, 4)

```

```

[1] 0.0373 0.0476 0.0634 0.0750 0.0889 0.1006 0.1008

```

```
round(q1_rs$cp, 1)
```

```
[1] 19.2 16.5 11.9 8.8 5.0 1.9 2.8
```

```
round(q1_rs$bic, 1)
```

```
[1] -1.8 -0.5 -1.2 -0.5 -0.6 0.0 4.6
```

since n for hw3q1 is 325, and we are looking at 2-8 inputs

```
q1_rs$aic.c <- 325*log(q1_rs$rss / 325) + 2*(2:8) +
  (2 * (2:8) * ((2:8)+1) / (325 - (2:8) - 1))
```

```
round(q1_rs$aic.c, 1)
```

```
[1] 1345.3 1342.8 1338.4 1335.4 1331.6 1328.5 1329.4
```

```
best_mods_1 <- tibble(
  k = 2:8,
  r2 = q1_rs$rsq,
  adjr2 = q1_rs$adjr2,
  cp = q1_rs$cp,
  aic.c = q1_rs$aic.c,
  bic = q1_rs$bic
)
```

```
best_mods <- cbind(best_mods_1, q1_rs$which)
```

```
best_mods
```

	k	r2	adjr2	cp	aic.c	bic (Intercept)
1	2	0.04024985	0.03727849	19.201108	1345.312	-1.78409378
2	3	0.05352477	0.04764604	16.495568	1342.823	-0.52693024
3	4	0.07206846	0.06339620	11.922415	1338.442	-1.17382911
4	5	0.08646901	0.07504987	8.817876	1335.422	-0.47321713
5	6	0.10295113	0.08889080	4.975487	1331.581	-0.60665159
6	7	0.11721042	0.10055401	1.921019	1328.462	-0.03045461
7	8	0.12027149	0.10084531	2.835965	1329.436	4.62447441

	practice	age	race	eth_hisp	sex	insurance	income	hsgrad	tobacco
1	FALSE	FALSE	FALSE	FALSE	TRUE	FALSE	FALSE	FALSE	FALSE
2	FALSE	TRUE	FALSE	FALSE	TRUE	FALSE	FALSE	FALSE	FALSE
3	FALSE	TRUE	FALSE	FALSE	TRUE	FALSE	FALSE	FALSE	FALSE
4	FALSE	TRUE	FALSE	FALSE	TRUE	FALSE	FALSE	FALSE	TRUE
5	TRUE	TRUE	TRUE	FALSE	TRUE	FALSE	FALSE	FALSE	FALSE
6	TRUE	TRUE	TRUE	FALSE	TRUE	FALSE	FALSE	FALSE	TRUE
7	TRUE	TRUE	TRUE	FALSE	TRUE	FALSE	FALSE	FALSE	TRUE

	depdiag	sbp	dbp	statin	bpm
1	FALSE	FALSE	FALSE	FALSE	FALSE
2	FALSE	FALSE	FALSE	FALSE	FALSE
3	FALSE	FALSE	FALSE	FALSE	TRUE
4	FALSE	FALSE	FALSE	FALSE	TRUE
5	FALSE	FALSE	FALSE	FALSE	TRUE
6	FALSE	FALSE	FALSE	FALSE	TRUE
7	TRUE	FALSE	FALSE	FALSE	TRUE

So, now, which of these models shows the best results?

- By Adjusted R²

- By C_p
- By corrected AIC
- By BIC

Building our 4 Plots

```
p1 <- ggplot(best_mods, aes(x = k, y = adjr2,
                           label = round(adjr2,2))) +
  geom_line() +
  geom_label() +
  geom_label(data = subset(best_mods,
                           adjr2 == max(adjr2)),
            aes(x = k, y = adjr2, label = round(adjr2,2)),
            fill = "yellow", col = "blue") +
  theme_bw() +
  scale_x_continuous(breaks = 2:9) +
  labs(x = "# of predictors (including intercept)",
       y = "Adjusted R-squared")

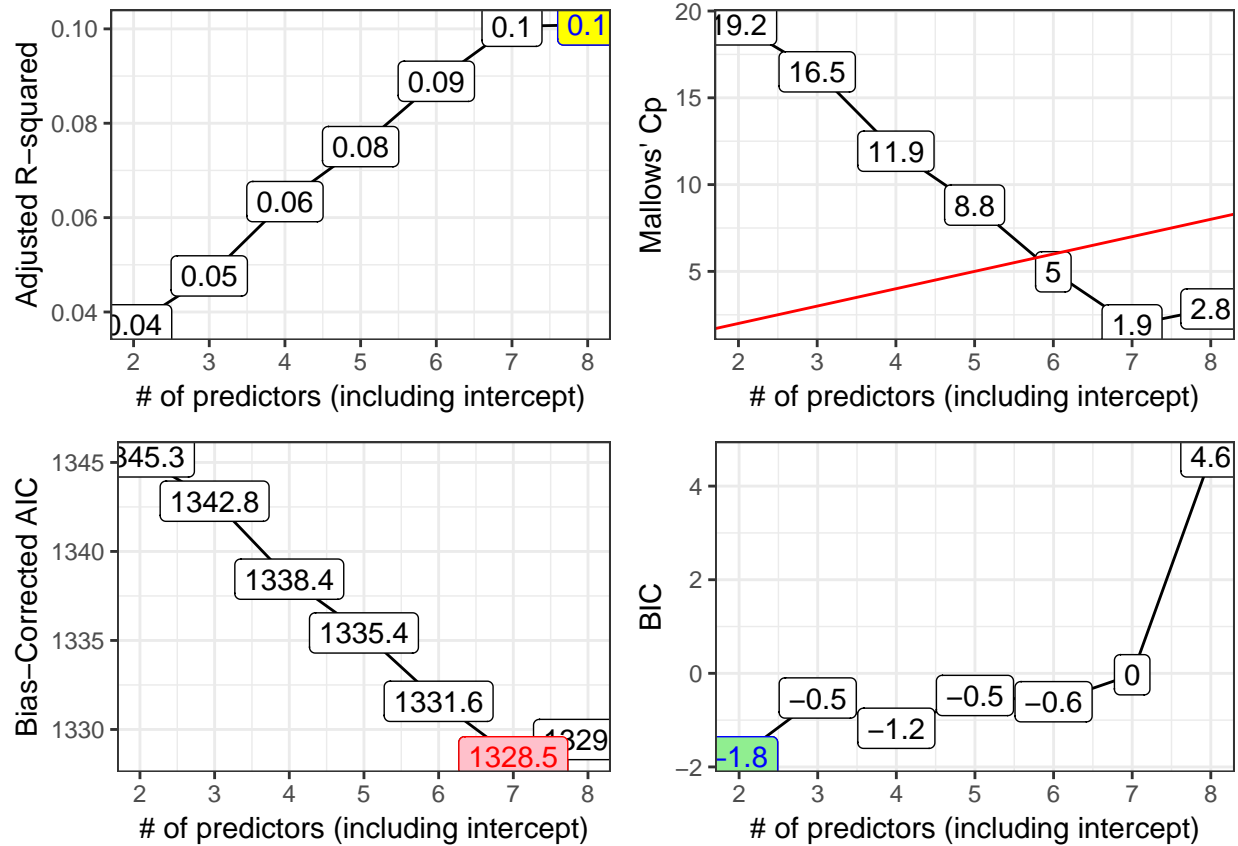
p2 <- ggplot(best_mods, aes(x = k, y = cp,
                           label = round(cp,1))) +
  geom_line() +
  geom_label() +
  geom_abline(intercept = 0, slope = 1,
             col = "red") +
  theme_bw() +
  scale_x_continuous(breaks = 2:9) +
  labs(x = "# of predictors (including intercept)",
       y = "Mallows' Cp")

p3 <- ggplot(best_mods, aes(x = k, y = aic.c,
                           label = round(aic.c,1))) +
  geom_line() +
  geom_label() +
  geom_label(data = subset(best_mods, aic.c == min(aic.c)),
            aes(x = k, y = aic.c), fill = "pink",
            col = "red") +
  theme_bw() +
  scale_x_continuous(breaks = 2:9) +
  labs(x = "# of predictors (including intercept)",
       y = "Bias-Corrected AIC")

p4 <- ggplot(best_mods, aes(x = k, y = bic,
                           label = round(bic,1))) +
  geom_line() +
  geom_label() +
  geom_label(data = subset(best_mods, bic == min(bic)),
            aes(x = k, y = bic),
            fill = "lightgreen", col = "blue") +
  theme_bw() +
  scale_x_continuous(breaks = 2:9) +
  labs(x = "# of predictors (including intercept)",
```

```
y = "BIC")
```

```
gridExtra::grid.arrange(p1, p2, p3, p4, nrow = 2)
```



Selecting a Winner

The models we'll consider are:

Inputs	Predictors Included	Reason
2	Intercept, sex	lowest BIC
5	Add age, bpmmed, tobacco to 2	(sort of) suggested by C_p
6	Add age, bpmmed, practice, race to 2	suggested by C_p
7	Add tobacco to 6	lowest AIC(corr.)
8	Add depdiag to 7	highest adj. R^2

We'll fit each of these models (and, in fact, the 5 predictor one, too) in turn, and then perform a 5-fold cross validation for each, then compare results. In each case, we'll calculate the root mean squared error of the predictions, and the mean absolute prediction error across the complete samples.

Model 2 cross-validation

There's not much point to this - though BIC likes it, we're not likely to be fond of a model that uses only a single binary predictor.

```

set.seed(4320142)

q1m2 <- hw3q1 %>%
  crossv_kfold(k = 5) %>%
  mutate(model = map(train, ~ lm(bmi ~ sex, data = .)))

q1m2_pred <- q1m2 %>%
  unnest(map2(model, test, ~ augment(.x, newdata = .y)))

res2 <- q1m2_pred %>%
  summarise(Model = "2",
            RMSE = sqrt(mean((bmi - .fitted) ^2)),
            MAE = mean(abs(bmi - .fitted)))

```

Note: Sometimes you'll get an error here if R thinks you want a different version of `summarize` than the one in `dplyr`. A way around this is to use `summarise` rather than `summarize`.

Model 5 cross-validation

```

set.seed(4320145)

q1m5 <- hw3q1 %>%
  crossv_kfold(k = 5) %>%
  mutate(model = map(train,
                    ~ lm(bmi ~ sex + age + bpmed + tobacco, data = .)))

q1m5_pred <- q1m5 %>%
  unnest(map2(model, test, ~ augment(.x, newdata = .y)))

res5 <- q1m5_pred %>%
  summarize(Model = "5",
            RMSE = sqrt(mean((bmi - .fitted) ^2)),
            MAE = mean(abs(bmi - .fitted)))

```

Model 6 cross-validation

```

set.seed(4320146)

q1m6 <- hw3q1 %>%
  crossv_kfold(k = 5) %>%
  mutate(model = map(train,
                    ~ lm(bmi ~ sex + age + bpmed + practice + race, data = .)))

q1m6_pred <- q1m6 %>%
  unnest(map2(model, test, ~ augment(.x, newdata = .y)))

res6 <- q1m6_pred %>%
  summarize(Model = "6",
            RMSE = sqrt(mean((bmi - .fitted) ^2)),
            MAE = mean(abs(bmi - .fitted)))

```

Model 7 cross-validation

```
set.seed(4320147)

q1m7 <- hw3q1 %>%
  crossv_kfold(k = 5) %>%
  mutate(model = map(train,
    ~ lm(bmi ~ sex + age + bpmed +
          practice + race + tobacco,
          data = .)))

q1m7_pred <- q1m7 %>%
  unnest(map2(model, test, ~ augment(.x, newdata = .y)))

res7 <- q1m7_pred %>%
  summarize(Model = "7",
    RMSE = sqrt(mean((bmi - .fitted) ^2)),
    MAE = mean(abs(bmi - .fitted)))
```

Model 8 cross-validation

```
set.seed(4320148)

q1m8 <- hw3q1 %>%
  crossv_kfold(k = 5) %>%
  mutate(model = map(train,
    ~ lm(bmi ~ sex + age + bpmed +
          practice + race + tobacco +
          depdiag, data = .)))

q1m8_pred <- q1m8 %>%
  unnest(map2(model, test, ~ augment(.x, newdata = .y)))

res8 <- q1m8_pred %>%
  summarize(Model = "8",
    RMSE = sqrt(mean((bmi - .fitted) ^2)),
    MAE = mean(abs(bmi - .fitted)))
```

Summary Table

```
bind_rows(res2, res5, res6, res7, res8)
```

```
# A tibble: 5 x 3
  Model  RMSE  MAE
  <chr> <dbl> <dbl>
1 2      7.99  6.17
2 5      7.64  5.90
3 6      7.78  5.96
4 7      7.73  5.96
5 8      7.66  5.90
```


Model 5 looks best by RMSE, and model 8 looks best by MAE. So we need to make a choice. I will pick the 5-input model, mainly because it has fewer predictors.

Moving forward with the sex, age, bpmmed and tobacco model

Refitting this model to the complete case sample of people without missing values on the variables we decided to use at the beginning, we have the following summary results.

```
summary(lm(bmi ~ sex + age + bpmmed + tobacco, data = hw3q1))
```

Call:

```
lm(formula = bmi ~ sex + age + bpmmed + tobacco, data = hw3q1)
```

Residuals:

Min	1Q	Median	3Q	Max
-18.6230	-5.1338	-0.8431	4.2077	27.4879

Coefficients:

	Estimate	Std. Error	t value	Pr(> t)
(Intercept)	38.57464	2.25912	17.075	< 2e-16 ***
sexM	-3.52607	0.88621	-3.979	8.58e-05 ***
age	-0.12415	0.03852	-3.223	0.001398 **
bpmmed	2.34229	0.91966	2.547	0.011338 *
tobaccoformer	4.50459	1.16070	3.881	0.000126 ***
tobacconeever	3.20177	1.11980	2.859	0.004526 **

Signif. codes: 0 '***' 0.001 '**' 0.01 '*' 0.05 '.' 0.1 ' ' 1

Residual standard error: 7.633 on 319 degrees of freedom

Multiple R-squared: 0.1146, Adjusted R-squared: 0.1007

F-statistic: 8.256 on 5 and 319 DF, p-value: 2.377e-07

Question 2 (10 points)

Refer to the modeling task you accomplished in Question 1. Now, your job is to fit a Spearman rho-squared plot to identify the candidate variables (out of the 14 you studied) on which you might most reasonably try to address non-linearity in a model predicting body-mass index, now making use of as much of the data set that missing data allow (without imputation.) Show the plot, and provide a written explanation of your conclusions about it, and specify the variables that are most appealing for non-linear augmentations, all in complete sentences. Which variables are most appealing candidates to add non-linear evaluations to a linear fit to the complete set of 14 predictors, and why?

Full hbp330 data (including Missing Values)

```
hbp330_full <- hbp330 %>%  
  mutate( bmi = weight / (height*height)) %>%  
  select(subject, bmi, practice, age, race, eth_hisp, sex,  
         insurance, income, hsgrad, tobacco,  
         depdiag, sbp, dbp, statin, bpmmed)
```

```
hbp330_full %>% skim(-subject)
```

Skim summary statistics

n obs: 330

n variables: 16

-- Variable type:character -----

variable	missing	complete	n	min	max	empty	n_unique
depdiag	0	330	330	2	3	0	2
eth_hisp	5	325	330	2	3	0	2
insurance	0	330	330	8	10	0	4
practice	0	330	330	1	1	0	2
race	2	328	330	5	12	0	4
sex	0	330	330	1	1	0	2
tobacco	0	330	330	5	7	0	3

-- Variable type:numeric -----

variable	missing	complete	n	mean	sd	p0	p25	p50
age	0	330	330	55.35	11.53	23	48	57
bmi	0	330	330	34.83	8.03	16.73	29.73	33.92
bpmed	0	330	330	0.66	0.48	0	0	1
dbp	0	330	330	74.75	10.2	41	68	74
hsgrad	0	330	330	81.93	8.54	57	75	81
income	0	330	330	35342.73	15888.27	6800	25600	30600
sbp	0	330	330	128.37	17.3	84	116	128.5
statin	0	330	330	0.71	0.46	0	0	1
p75	p100							
65	77							
39.18	64.04							
1	1							
82	106							
89	100							
42475	147400							
138	194							
1	1							

There are 5 missing `eth_hisp` values and 2 missing `race` values.

Complete Cases: hbp330 data after deleting cases with NAs

```
hbp330_noNA <- hbp330 %>%
  mutate( bmi = weight / (height*height)) %>%
  select(subject, bmi, practice, age, race, eth_hisp, sex,
         insurance, income, hsgrad, tobacco,
         depdiag, sbp, dbp, statin, bpmed) %>%
  drop_na
dim(hbp330_noNA)
```

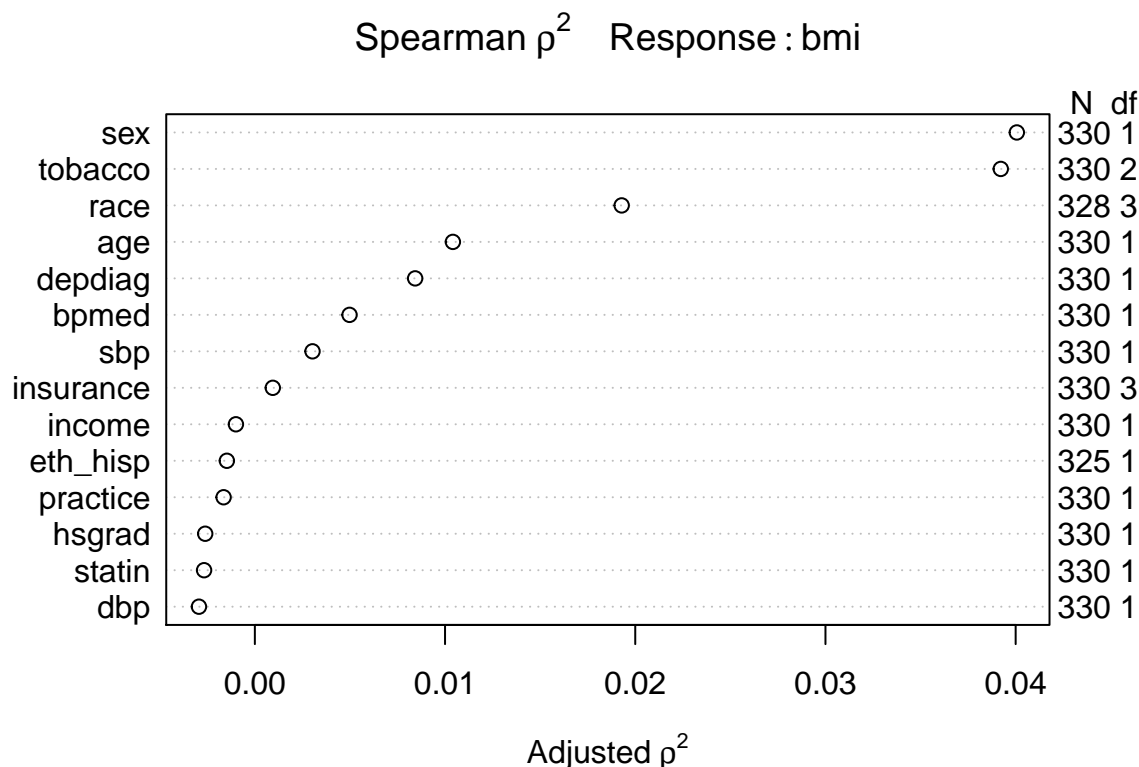
```
[1] 325 16
```

Again, we lose a total of five observations (dropping from 330 to 325 subjects) by dropping missing values.

Spearman rho-squared plot (applied to full data)

You might have chosen to include all observations, and simply allow the Spearman ρ^2 plot to reduce the sample size for the specific variables (race and eth_hisp) that had missing values. That produces this result.

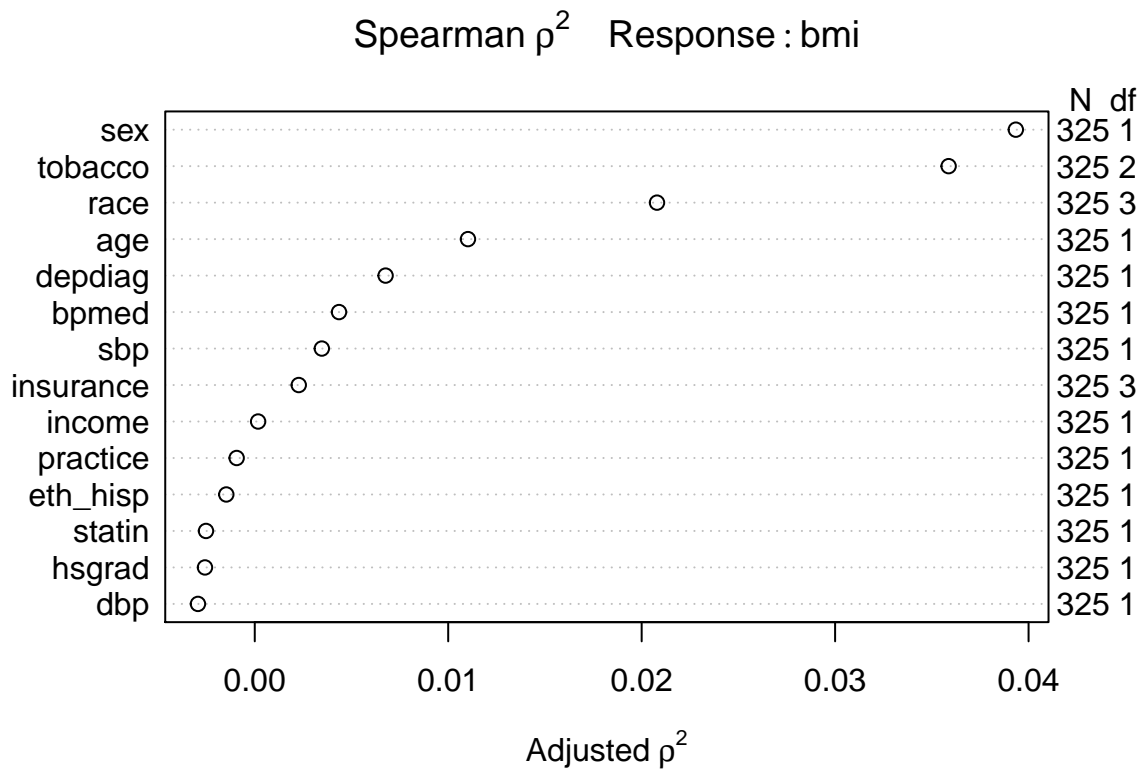
```
plot(Hmisc::spearman2(bmi ~ practice + age + race +
  eth_hisp + sex + insurance +
  income + hsgrad + tobacco +
  depdiag + sbp + dbp + statin +
  bpmed, data = hbp330_full))
```



Spearman rho-squared plot (applied to 325 complete cases)

Or, you might have chosen to include only the 325 complete cases, and so the Spearman ρ^2 plot would address only those subjects without missing eth_hisp or race. That produces this result.

```
plot(Hmisc::spearman2(bmi ~ practice + age + race +
  eth_hisp + sex + insurance +
  income + hsgrad + tobacco +
  depdiag + sbp + dbp + statin +
  bpmed, data = hbp330_noNA))
```



In either case, the variables which pack the largest “potential predictive punch” in this setting are, in order, (1) **sex** and (2) **tobacco**. Certainly, those are the most appealing variables for which we should consider non-linear augmentations. Since these are categorical variables, the inclusion of interaction terms seems appealing. We might, for instance, include the **sex-tobacco** interaction or an interaction of **sex** or **tobacco** or both with the next two highest variables on the list: **race** and **age**.

Question 3 (20 points)

First, in 2-4 complete English sentences, please specify, using your own words and complete English sentences, the most useful and relevant piece of advice you took away from reading Jeff Leek’s *How To Be A Modern Scientist*.

Second, in an essay of 4-8 additional sentences, describe why this particular piece of advice was meaningful or useful for you, personally, and how it will affect the way you move forward.

We don’t write sketches for essay questions. We hope to share a few of the more interesting responses with you after they’ve been graded.

Session Information

```
sessioninfo::session_info()
```

```
- Session info -----
setting  value
```

```

version R version 3.5.2 (2018-12-20)
os      Windows 10 x64
system  x86_64, mingw32
ui       RTerm
language (EN)
collate English_United States.1252
ctype   English_United States.1252
tz       America/New_York
date     2019-02-22

```

```

- Packages -----
package      * version      date      lib source
acepack      1.4.1        2016-10-29 [1] CRAN (R 3.5.0)
assertthat   0.2.0        2017-04-11 [1] CRAN (R 3.5.0)
backports    1.1.3        2018-12-14 [1] CRAN (R 3.5.2)
base64enc    0.1-3        2015-07-28 [1] CRAN (R 3.5.0)
bindr        0.1.1        2018-03-13 [1] CRAN (R 3.5.0)
bindrcpp     * 0.2.2        2018-03-29 [1] CRAN (R 3.5.0)
broom        * 0.5.1        2018-12-05 [1] CRAN (R 3.5.2)
cellranger   1.1.0        2016-07-27 [1] CRAN (R 3.5.0)
checkmate    1.9.1        2019-01-15 [1] CRAN (R 3.5.2)
cli          1.0.1        2018-09-25 [1] CRAN (R 3.5.1)
cluster      2.0.7-1      2018-04-13 [2] CRAN (R 3.5.2)
colorspace   1.4-0        2019-01-13 [1] CRAN (R 3.5.2)
crayon       1.3.4        2017-09-16 [1] CRAN (R 3.5.0)
data.table   1.12.0       2019-01-13 [1] CRAN (R 3.5.2)
digest       0.6.18       2018-10-10 [1] CRAN (R 3.5.1)
dplyr        * 0.7.8        2018-11-10 [1] CRAN (R 3.5.1)
evaluate     0.12         2018-10-09 [1] CRAN (R 3.5.1)
fanside      0.4.0        2018-10-05 [1] CRAN (R 3.5.1)
forcats      * 0.3.0        2018-02-19 [1] CRAN (R 3.5.0)
foreign      0.8-71       2018-07-20 [1] CRAN (R 3.5.2)
Formula      1.2-3        2018-05-03 [1] CRAN (R 3.5.0)
generics     0.0.2        2018-11-29 [1] CRAN (R 3.5.1)
ggplot2      * 3.1.0        2018-10-25 [1] CRAN (R 3.5.1)
glue         1.3.0        2018-07-17 [1] CRAN (R 3.5.2)
gridExtra    2.3          2017-09-09 [1] CRAN (R 3.5.2)
gtable       0.2.0        2016-02-26 [1] CRAN (R 3.5.0)
haven        2.0.0.9000   2018-12-11 [1] Github (tidyverse/haven@f0dc4e5)
Hmisc        4.2-0        2019-01-26 [1] CRAN (R 3.5.2)
hms          0.4.2        2018-03-10 [1] CRAN (R 3.5.0)
htmlTable    1.13.1       2019-01-07 [1] CRAN (R 3.5.2)
htmltools    0.3.6        2017-04-28 [1] CRAN (R 3.5.0)
htmlwidgets  1.3          2018-09-30 [1] CRAN (R 3.5.1)
httr         1.4.0        2018-12-11 [1] CRAN (R 3.5.2)
janitor      * 1.1.1        2018-07-31 [1] CRAN (R 3.5.1)
jsonlite     1.6          2018-12-07 [1] CRAN (R 3.5.2)
knitr        1.21         2018-12-10 [1] CRAN (R 3.5.2)
labeling     0.3          2014-08-23 [1] CRAN (R 3.5.0)
lattice      0.20-38      2018-11-04 [1] CRAN (R 3.5.2)
latticeExtra 0.6-28       2016-02-09 [1] CRAN (R 3.5.0)
lazyeval     0.2.1        2017-10-29 [1] CRAN (R 3.5.0)
leaps        * 3.0          2017-01-10 [1] CRAN (R 3.5.2)
lubridate    1.7.4        2018-04-11 [1] CRAN (R 3.5.0)

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magrittr	1.5	2014-11-22	[1]	CRAN	(R 3.5.2)
Matrix	1.2-15	2018-11-01	[2]	CRAN	(R 3.5.2)
modelr	* 0.1.2	2018-05-11	[1]	CRAN	(R 3.5.0)
munsell	0.5.0	2018-06-12	[1]	CRAN	(R 3.5.0)
nlme	3.1-137	2018-04-07	[2]	CRAN	(R 3.5.2)
nnet	7.3-12	2016-02-02	[1]	CRAN	(R 3.5.2)
pillar	1.3.1	2018-12-15	[1]	CRAN	(R 3.5.2)
pkgconfig	2.0.2	2018-08-16	[1]	CRAN	(R 3.5.1)
plyr	1.8.4	2016-06-08	[1]	CRAN	(R 3.5.0)
purrr	* 0.3.0	2019-01-27	[1]	CRAN	(R 3.5.2)
R6	2.3.0	2018-10-04	[1]	CRAN	(R 3.5.1)
RColorBrewer	1.1-2	2014-12-07	[1]	CRAN	(R 3.5.0)
Rcpp	1.0.0	2018-11-07	[1]	CRAN	(R 3.5.1)
readr	* 1.3.1	2018-12-21	[1]	CRAN	(R 3.5.2)
readxl	1.3.0	2019-02-15	[1]	CRAN	(R 3.5.2)
rlang	0.3.1	2019-01-08	[1]	CRAN	(R 3.5.2)
rmarkdown	1.11	2018-12-08	[1]	CRAN	(R 3.5.2)
rpart	4.1-13	2018-02-23	[2]	CRAN	(R 3.5.2)
rstudioapi	0.9.0	2019-01-09	[1]	CRAN	(R 3.5.2)
rvest	0.3.2	2016-06-17	[1]	CRAN	(R 3.5.0)
scales	1.0.0	2018-08-09	[1]	CRAN	(R 3.5.1)
sessioninfo	1.1.1	2018-11-05	[1]	CRAN	(R 3.5.1)
skimr	* 1.0.4	2019-01-13	[1]	CRAN	(R 3.5.2)
snakecase	0.9.2	2018-08-14	[1]	CRAN	(R 3.5.1)
stringi	1.2.4	2018-07-20	[1]	CRAN	(R 3.5.1)
stringr	* 1.3.1	2018-05-10	[1]	CRAN	(R 3.5.0)
survival	2.43-3	2018-11-26	[1]	CRAN	(R 3.5.2)
tibble	* 2.0.1	2019-01-12	[1]	CRAN	(R 3.5.2)
tidyr	* 0.8.2	2018-10-28	[1]	CRAN	(R 3.5.1)
tidyselect	0.2.5	2018-10-11	[1]	CRAN	(R 3.5.1)
tidyverse	* 1.2.1	2017-11-14	[1]	CRAN	(R 3.5.2)
utf8	1.1.4	2018-05-24	[1]	CRAN	(R 3.5.0)
withr	2.1.2	2018-03-15	[1]	CRAN	(R 3.5.0)
xfun	0.4	2018-10-23	[1]	CRAN	(R 3.5.1)
xml2	1.2.0	2018-01-24	[1]	CRAN	(R 3.5.0)
yaml	2.2.0	2018-07-25	[1]	CRAN	(R 3.5.1)

[1] C:/Users/Thomas/Documents/R/win-library/3.5

[2] C:/Program Files/R/R-3.5.2/library