# ECE 595: Homework 1

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### Exercise 2

(a) For a guassian distribution:

$$E[x] = \int_{-\infty}^{\infty} x \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}} dx$$

$$= \int_{-\infty}^{\infty} (x+\mu) \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{x^2}{2\sigma^2}} dx$$

$$= \int_{-\infty}^{\infty} x \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{x^2}{2\sigma^2}} dx + \int_{-\infty}^{\infty} \mu \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{x^2}{2\sigma^2}} dx$$

$$= 0 + \mu \frac{1}{\sqrt{2\pi\sigma^2}} \times \sigma \sqrt{2\pi}$$

$$= \mu$$
(1)

$$Var[x] = \int_{-\infty}^{\infty} (x - \mu)^2 \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x - \mu)^2}{2\sigma^2}} dx$$

$$= \frac{1}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} x^2 e^{-\frac{x^2}{2\sigma^2}} dx$$

$$let \ y = \frac{x}{\sigma}, \ then \ dy = \frac{1}{\sigma} dx$$

$$= \frac{\sigma^3}{\sqrt{2\pi\sigma^2}} \int_{-\infty}^{\infty} y^2 e^{-\frac{y^2}{2}} dy$$

$$= \sigma^2$$

$$(2)$$

(b) Data generated and plotted as follows.

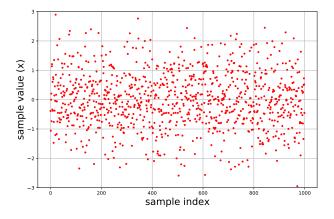
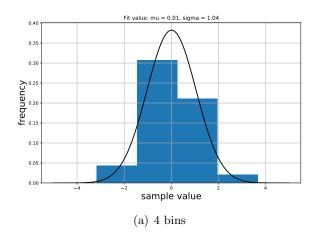
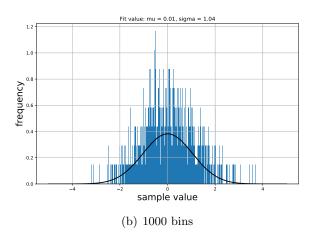


Figure 1: Gussian random data.

(c)

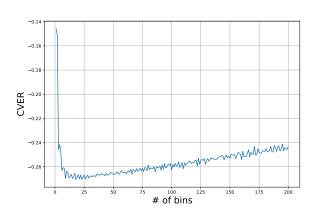
## (i)..(iv) plots shown below

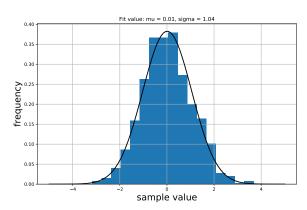




(v) TODO: fill this

(d) compare to part (c), the histogram fits a lot better with the PDF plots shown below





(c) Cross validation estimator of risk vs. # of bins

(d) Histogram and PDF overlayed with optimized number of bins

### Exercise 3

(a)

$$f_{\mathbf{X}}(\mathbf{x}) = \frac{1}{\sqrt{2\pi^2 |\mathbf{\Sigma}|}} exp\{-\frac{1}{2}(\mathbf{x} - \boldsymbol{\mu})^T \boldsymbol{\Sigma}^{-1}(\mathbf{x} - \boldsymbol{\mu})\}$$

(i) plug in

$$\mathbf{X} = \begin{bmatrix} X_1 \\ X_2 \end{bmatrix}, \, \mathbf{x} = \begin{bmatrix} x_1 \\ x_2 \end{bmatrix}, \, \boldsymbol{\mu} = \begin{bmatrix} 2 \\ 6 \end{bmatrix}, \, \mathrm{and} \, \boldsymbol{\Sigma} = \begin{bmatrix} 2 & 1 \\ 1 & 2 \end{bmatrix}$$

we get

$$f_{\begin{bmatrix} X_1 \\ X_2 \end{bmatrix}} \left( \begin{bmatrix} x_1 \\ x_2 \end{bmatrix} \right) = \frac{1}{\sqrt{2\pi^2 \begin{vmatrix} 2 & 1 \\ 1 & 2 \end{vmatrix}}} exp \left\{ -\frac{1}{2} \begin{bmatrix} x_1 - 2 \\ x_2 - 6 \end{bmatrix}^T \begin{bmatrix} 2 & 1 \\ 1 & 2 \end{bmatrix}^{-1} \begin{bmatrix} x_1 - 2 \\ x_2 - 6 \end{bmatrix} \right\}$$

$$= \frac{1}{\pi\sqrt{6}} exp \left\{ -\frac{1}{3} \left( (x_1 - 2)^2 - (x_1 - 2)(x_2 - 6) + (x_2 - 6)^2 \right) \right\}$$
(3)

(ii) plot shown below

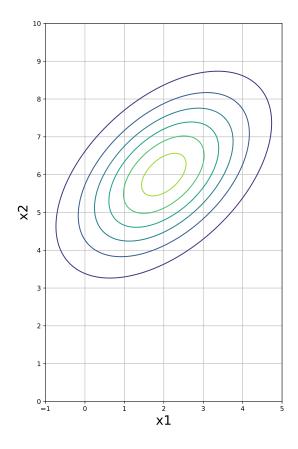


Figure 2: Gussian random data

(b)

(i)

To prove:  $\mu_{\mathbf{Y}} = b$ 

$$\mu_{\mathbf{Y}} = \mathbb{E}[\mathbf{Y}] = \mathbb{E}[\mathbf{A}\mathbf{X} + \mathbf{b}] = \mathbb{E}[\mathbf{A}\mathbf{X}] + \mathbf{b}$$

$$= \mathbf{A}\mathbb{E}[\mathbf{X}] + \mathbf{b}$$

$$since \ \mathbf{X} \in \mathcal{N}(\mathbf{0}, \mathbf{I}), \ \mathbb{E}[\mathbf{X}] = \mathbf{0}$$

$$= \mathbf{A}\mathbf{0} + \mathbf{b}$$

$$= \mathbf{b}$$
(4)

To prove:  $\Sigma_{\mathbf{Y}} = \mathbf{A}\mathbf{A}^T$ 

$$\Sigma_{\mathbf{Y}} = \mathbb{E}[(\mathbf{Y} - \boldsymbol{\mu}_{\mathbf{Y}})(\mathbf{Y} - \boldsymbol{\mu}_{\mathbf{Y}})^{T}]$$

$$= \mathbb{E}[\mathbf{A}\mathbf{X}\mathbf{X}^{T}\mathbf{A}^{T}]$$

$$= \mathbf{A}\mathbb{E}[\mathbf{X}\mathbf{X}^{T}]\mathbf{A}^{T}$$

$$= \mathbf{A}(\mathbf{I} + \mathbf{0})\mathbf{A}^{T}$$

$$= \mathbf{A}\mathbf{A}^{T}$$
(5)

(ii) To prove:  $\Sigma_{\mathbf{Y}}$  is symmetric and positive semi-definite

$$\Sigma_{\mathbf{Y}ij} = \Sigma_{x=1}^{n} \mathbf{A}_{ix} \mathbf{A}_{xj}^{T}$$
$$= \Sigma_{x=1}^{n} \mathbf{A}_{jx} \mathbf{A}_{xi}^{T}$$
$$= \Sigma_{\mathbf{Y}ji}$$

thus symmetric

$$\mathbf{x}^{T} \mathbf{\Sigma}_{\mathbf{Y}} \mathbf{x} = \mathbf{x}^{T} \mathbf{A} \mathbf{A}^{T} \mathbf{x}$$

$$let \ \mathbf{y} = \mathbf{x}^{T} \mathbf{A}$$
(6)

 $= \mathbf{y}\mathbf{y}^{T}$   $= \|\mathbf{y}\|^{2} \ge 0$ 

 $thus\ positive\ semi-definite$ 

(iii)  $Null(\mathbf{A}) = \mathbf{0}$ 

(iv) By inspection,

$$b = \boldsymbol{\mu}_{\mathbf{Y}} = \begin{bmatrix} 2 \\ 6 \end{bmatrix}$$

By Cholesky decomposition,

$$oldsymbol{\Sigma_Y} = \mathbf{A}\mathbf{A}^T$$
  $\mathbf{A} = egin{bmatrix} \sqrt{2} & 0 \ rac{\sqrt{2}}{2} & rac{\sqrt{6}}{2} \end{bmatrix}$ 

(c)

(i) Data points drawn from 2D standard guassian distribution are shown below.

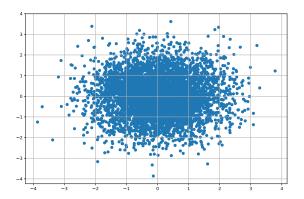


Figure 3: 2D standard Gussian random data

(ii) After applying the affine transformation, the data plot shown below.

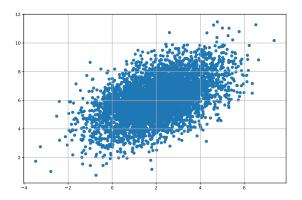


Figure 4: After applying affine transformation

(iii) From my favoriate python and numpy, for the data above:

$$\boldsymbol{\mu}_{\mathbf{Y}} = \begin{bmatrix} 2.003 \\ 6.004 \end{bmatrix}$$

$$\Sigma_{\mathbf{Y}} = \begin{bmatrix} 2.017 & 1.050 \\ 1.050 & 2.074 \end{bmatrix}$$

#### Exercise 4

(a)

$$|\mathbf{x}^{T}\mathbf{A}\mathbf{y}| = \left|\sum_{i}\sum_{j} a_{ij}x_{i}y_{j}\right|$$

$$= \left|\sum_{i}x_{i}\sum_{j} a_{ij}y_{j}\right|$$
(7)

(b)

(i) For a invertible  $n \times n$  matrix **A** 

$$\mathbf{A} = \mathbf{PDP}$$

$$\mathbf{A}^{-1} = (\mathbf{PDP}^{-1})^{-1}$$

$$= \mathbf{P}^{-1}\mathbf{D}^{-1}\mathbf{P}$$
(8)

Since A is positive definite,  $\lambda_i > 0$  for  $1 \le i \le n$ , **D** is invertible.

Thus **A** is also invertible.

(ii) 
$$f\left(\begin{bmatrix}x1\\x2\end{bmatrix}\right) = x_1^2 - x_2^2$$

the Hessian of this function is

$$\begin{bmatrix} 2 & 0 \\ 0 & -2 \end{bmatrix}$$

which is invertible but not positive definite

(iii) If a matrix is invertible while positive semi-definite, the matrix is positive definite. By definition of positive semi-definite,  $\mathbf{x}\mathbf{A}\mathbf{x}^T \geq 0$ , and for a eigenvalue of  $\mathbf{A}$ ,  $\lambda_i$ ,  $\lambda_i\mathbf{u}_i = \mathbf{A}\mathbf{u}_i$  So,

$$\mathbf{u}_i^T \mathbf{A} \mathbf{u}_i = \lambda_i$$

$$\lambda_i \ge 0$$

$$(9)$$

If the matrix is invertible, it cannot have  $\lambda_i = 0$ . So,  $\lambda_i > 0$ , thus  $\mathbf{x} \mathbf{A} \mathbf{x}^T > 0$ .

The matrix **A** is positive definite.

(c) To prove:  $(\exists \mathbf{A}^{\dagger}) \ \mathbf{A} \mathbf{A}^{\dagger} \mathbf{A} = \mathbf{A}$ 

$$\mathbf{A} = \mathbf{U}\boldsymbol{\Lambda}\mathbf{U}^{T}$$

$$\mathbf{A}\mathbf{A}^{\dagger}\mathbf{A} = \mathbf{U}\boldsymbol{\Lambda}\mathbf{U}^{T}\mathbf{A}^{\dagger}\mathbf{U}\boldsymbol{\Lambda}\mathbf{U}^{T}$$

$$= \mathbf{U}\boldsymbol{\Lambda}\mathbf{U}^{T}$$

$$simplify$$

$$\mathbf{U}^{T}\mathbf{A}^{\dagger}\mathbf{U}\boldsymbol{\Lambda} = \boldsymbol{\Lambda}\boldsymbol{\Lambda}^{-}$$

$$\mathbf{A}^{\dagger} = \mathbf{U}\boldsymbol{\Lambda}^{-}\mathbf{U}^{T}$$
(10)

Since  ${\bf A}$  is symmetric, it is guaranteed that can be decompose, thus  ${\bf A}^\dagger$  exist.