

Definition: Random variables X_1, X_2, \dots, X_n are independent if $P(X_1 \in A_1, X_2 \in A_2, \dots, X_n \in A_n) = P(X_1 \in A_1)P(X_2 \in A_2) \cdots P(X_n \in A_n)$ for all subsets $A_1, \dots, A_n \subseteq \mathbb{R}$.

↪ If r.v.'s are continuous, the joint pdf factorizes:

$$f_{X_1, \dots, X_n}(x_1, \dots, x_n) = f_{X_1}(x_1)f_{X_2}(x_2) \cdots f_{X_n}(x_n),$$

Similar for discrete case, using pmf instead of pdf.

Def: We say X_1, \dots, X_n are iid (independent & identically distributed) if they're independent R.V.s and each has the same distribution.

↪ $P(X_1 \in A) = P(X_2 \in A) = \dots = P(X_n \in A)$ for every $A \subseteq \mathbb{R}$.

↪ If continuous, this means each X_i has same PDF.

Ex 1: Maximum of iid:

Let X_1, X_2, \dots, X_n be iid with common PDF $f(x)$,
continuous R.V.s

Let $Y = \max(X_1, \dots, X_n)$.

Goal: Find CDF/PDF of Y .

Start with CDF. Let $y \in \mathbb{R}$,

$$\begin{aligned} P(Y \leq y) &= P(\max(X_1, \dots, X_n) \leq y) && \text{same event!} \\ &= P(X_1 \leq y, X_2 \leq y, \dots, X_n \leq y) \\ &= P(X_1 \leq y)P(X_2 \leq y) \cdots P(X_n \leq y) && \text{by independence} \\ &= F_{X_1}(y)F_{X_2}(y) \cdots F_{X_n}(y). \end{aligned}$$

where F_{X_i} is the CDF of X_i . But these CDFs are all identical. Let $F = F_{X_1} = F_{X_2} = \dots = F_{X_n}$.

So

$$F_Y(y) = P(Y \leq y) = (F(y))^n \quad \text{is the CDF of } Y.$$

The PDF is

$$\begin{aligned} f_Y(y) &= \frac{d}{dy} F_Y(y) = \frac{d}{dy} [F(y)^n] \\ &= n F(y)^{n-1} \frac{d}{dy} F(y) \\ &= n F(y)^{n-1} f(y), \quad \text{where } f \text{ was the PDF} \\ &\quad \text{of the } X_i \text{'s.} \end{aligned}$$

Ex 2: Minimum of iid

Let X_1, \dots, X_n are iid continuous RVs with CDF F and PDF f . Let $Z = \min(X_1, \dots, X_n)$.

Goal: Find CDF & PDF of Z .

Let $r \in \mathbb{R}$. Start with CDF:

$$\begin{aligned} P(Z \leq r) &= P(\min(X_1, \dots, X_n) \leq r) \\ &= P(\text{at least one of the } X_i's \text{ is } \leq r) \\ &= \dots \text{ hard to use independence here...} \end{aligned}$$

Instead note $P(Z \leq r) = 1 - P(Z > r)$.

(calculate)

$$\begin{aligned} P(Z > r) &= P(\underline{\min(X_1, \dots, X_n)} > r) \\ &= P(\underline{X_1 > r, X_2 > r, \dots, X_n > r}) \quad \text{same event!} \\ &= P(X_1 > r)P(X_2 > r) \cdots P(X_n > r) \quad \text{by independence} \\ &= (1 - P(X_1 \leq r))(1 - P(X_2 \leq r)) \cdots (1 - P(X_n \leq r)) \\ &= (1 - F(r))(1 - F(r)) \cdots (1 - F(r)) \\ &= (1 - F(r))^n. \end{aligned}$$

So CDF of Z is

$$\begin{aligned} F_Z(r) &= P(Z \leq r) = 1 - P(Z > r) \\ &= 1 - (1 - F(r))^n. \end{aligned}$$

The PDF is then

$$\begin{aligned} f_Z(r) &= \frac{d}{dr} F_Z(r) = \frac{d}{dr} [1 - (1 - F(r))^n] \\ &= -n(1 - F(r))^{n-1} \frac{d}{dr} [1 - F(r)] \\ &= n f(r) (1 - F(r))^{n-1} \end{aligned}$$

$$F' = f$$

where f and F are the PDF/CDF of X_i .

Ex 3: Let X_1, \dots, X_n be iid $\text{Exp}(\lambda)$ where $\lambda > 0$.

The PDF is

$$f(x) = \begin{cases} \lambda e^{-\lambda x} & \text{if } x \geq 0 \\ 0 & \text{if } x < 0. \end{cases}$$

Let $Z = \min(X_1, \dots, X_n)$.

Think of X_i as time until bus # i arrives.

Then Z models the time until any bus arrives.

We know CDF of X_i is $F(x) = 1 - e^{-\lambda x}$ from before.

The PDF of Z is

$$\begin{aligned} f_Z(r) &= n f(r) (1 - F(r))^{n-1} = n \lambda e^{-\lambda r} (e^{-\lambda r})^{n-1} \\ &= n \lambda e^{-n\lambda r} \quad \text{for } r \geq 0 \end{aligned}$$

and $f_Z(r) = 0$ for $r < 0$.

Thus $Z \sim \text{Exp}(n\lambda)$.

Conditioning with Continuous RVs

Recap: If X and Y are discrete RVs with joint PMF $P_{X,Y}$ and marginals P_X , P_Y , then the conditional PMF of X given Y is the function

$$P_{X|Y}(x|y) = P(X=x | Y=y) = \frac{P(X=x, Y=y)}{P(Y=y)}$$

$$= \frac{P_{X,Y}(x,y)}{\dots} \quad \text{when } P_Y(y) > 0.$$

$$= \frac{P_{X,Y}(x,y)}{P_Y(y)}, \quad \text{when } P_Y(y) > 0.$$

Tricky to define conditional distribution for continuous RVs
because $P(Y=y) = 0$.

Def:

Let X and Y be continuous RVs with joint PDF $f_{X,Y}$
and marginals f_X and f_Y . The conditional PDF
of X given Y is defined as the function

$$f_{X|Y}(x|y) = \frac{f_{X,Y}(x,y)}{f_Y(y)} \quad \text{for } x, y \text{ such that } f_Y(y) > 0,$$

Note PDFs are not probabilities! This is not equal
to $\frac{P(X=x, Y=y)}{P(Y=y)}$ like in discrete case,

How to use this:

$$\textcircled{1} \quad P(a \leq X \leq b | Y=y) = \int_a^b f_{X|Y}(x|y) dx \quad \text{for } a < b.$$

Think of y as fixed.

The function $f_{X|Y}(x|y)$ is a PDF, for fixed y .

$$\textcircled{2} \quad E[X | Y=y] = \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx$$

and similarly

and similarly

$$\mathbb{E}[g(x) | Y=y] = \int_{-\infty}^{\infty} g(x) f_{x|y}(x|y) dx.$$

Same rules as in unconditional case apply, just with conditioning on $Y=y$ throughout,

(3) $f_{x|y}(x|y)$ is a PDF, for fixed y :

a) $f_{x|y} \geq 0$ ✓

b) Check if $\int_{-\infty}^{\infty} f_{x|y}(x|y) dx = 1$?

$$f_{x|y}(x|y) = \frac{f_{x,y}(x,y)}{f_y(y)}$$

Thus

$$\int_{-\infty}^{\infty} f_{x|y}(x|y) dx = \int_{-\infty}^{\infty} \frac{f_{x,y}(x,y)}{f_y(y)} dx$$

$$= \frac{1}{f_y(y)} \int_{-\infty}^{\infty} f_{x,y}(x,y) dx$$

$$= \frac{1}{f_y(y)} f_y(y)$$

$$= 1.$$

✓

(4) If X and Y are independent, then $f_{x|y}(x|y) = f_x(x)$.

Proof:

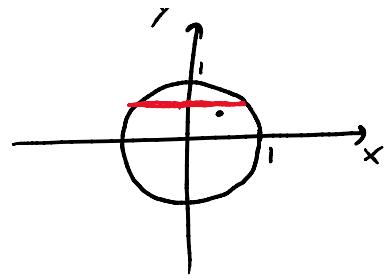
$$\begin{aligned} f_{x|y}(x|y) &= \frac{f_{x,y}(x,y)}{f_y(y)} = \frac{f_x(x)f_y(y)}{f_y(y)} \quad \text{by independence} \\ &= f_x(x). \end{aligned}$$

Recall:
A indep of B
if $P(A|B) = P(A)$

$\therefore 4 \cdot 1 \cdot 1 / (v \cdot v) := \dots - t - 1.$



Ex 4: Let (x, y) is a uniformly random point in the circle of radius 1 around origin. The joint PDF is



$$f_{x,y}(x,y) = \begin{cases} \text{const.} & \text{if } (x,y) \text{ lies in circle} \\ 0 & \text{otherwise} \end{cases}$$

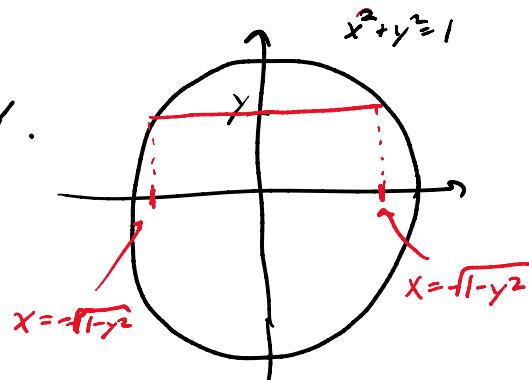
Here const. should be $\frac{1}{\text{area of circle}}$. So

$$f_{x,y}(x,y) = \begin{cases} \frac{1}{\pi} & \text{if } x^2+y^2 \leq 1 \\ 0 & \text{otherwise.} \end{cases}$$

Goal 1: Find conditional PDF of X given Y .

First need marginal of Y :

$$\begin{aligned} f_Y(y) &= \int_{-\infty}^{\infty} f_{x,y}(x,y) dx \\ &= \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} \frac{1}{\pi} dx \\ &= \frac{2}{\pi} \sqrt{1-y^2} \quad \text{for } -1 \leq y \leq 1 \end{aligned}$$



and $f_Y(y) = 0$ for $|y| > 1$.

Conditional PDF:

$$\begin{aligned} f_{X|Y}(x|y) &= \frac{f_{x,y}(x,y)}{f_Y(y)} = \frac{\frac{1}{\pi}}{\frac{2}{\pi} \sqrt{1-y^2}} \\ &= \frac{1}{2\sqrt{1-y^2}} \quad \text{for } (x,y) \text{ in the circle.} \end{aligned}$$

$$\text{So } f_{X|Y}(x|y) = \begin{cases} \frac{1}{2\sqrt{1-y^2}} & \text{for } x^2+y^2 \leq 1 \\ 0 & \text{otherwise.} \end{cases}$$

$$= \begin{cases} \frac{1}{2\sqrt{1-y^2}} & \text{for } -\sqrt{1-y^2} \leq x \leq \sqrt{1-y^2} \\ 0 & \text{otherwise} \end{cases}$$

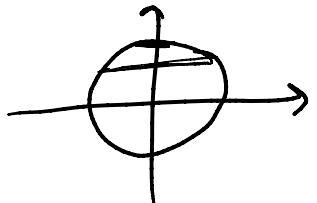
Since y is fixed, as a function of x , $f_{X|Y}(x|y)$ is the PDF of a $\text{Unif}[-\sqrt{1-y^2}, \sqrt{1-y^2}]$.

Goal 2: Find $E[X|Y=y]$ for $-1 < y < 1$.

This is 0, midpoint of $[-\sqrt{1-y^2}, \sqrt{1-y^2}]$.

Or directly:

$$\begin{aligned} E[X|Y=y] &= \int_{-\infty}^{\infty} x f_{X|Y}(x|y) dx \\ &= \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} x \cdot \frac{1}{2\sqrt{1-y^2}} dx \\ &= \frac{1}{2\sqrt{1-y^2}} \left[\frac{1}{2} x^2 \right]_{x=-\sqrt{1-y^2}}^{\sqrt{1-y^2}} = 0. \end{aligned}$$



Similarly:

$$\begin{aligned} E[X^2|Y=y] &= \int_{-\infty}^{\infty} x^2 f_{X|Y}(x|y) dx \\ &= \int_{-\sqrt{1-y^2}}^{\sqrt{1-y^2}} x^2 \cdot \frac{1}{2\sqrt{1-y^2}} dx \\ &= \dots = \frac{1-y^2}{3}. \end{aligned}$$

Def: The conditional variance of X given $Y=y$

$$\text{is } \text{Var}(X|Y=y) = E[(X - E[X|Y=y])^2 | Y=y]$$

$$\begin{aligned}\text{Var}(X|Y=y) &= \mathbb{E}[(X - \mathbb{E}[X|Y=y])^2 | Y=y] \\ &= \mathbb{E}[X^2 | Y=y] - (\mathbb{E}[X | Y=y])^2.\end{aligned}$$

In Ex 4 above:

$$\begin{aligned}\text{Var}(X|Y=y) &= \mathbb{E}[X^2 | Y=y] - (\mathbb{E}[X | Y=y])^2 \\ &= \frac{1-y^2}{3} - 0 \\ &= \frac{1-y^2}{3}.\end{aligned}$$

Note this approaches 0 as $y \rightarrow \pm 1$.

