

2025 SIAM Annual Meeting

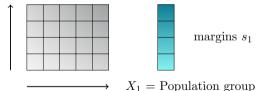
Raking Methods and Applications to Health Metrics

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What is raking?



$$X_2 = \text{County}$$



Two categorical variables X_1 and X_2 taking I and J possible values.

When summing the rows and columns of the table y, the observations y_{ij} do not add up to the values in the margins s_1 and s_2 .

$$\sum_{i=1}^{I} y_{ij} \neq s_{1j} \quad j = 1, \cdots, J$$

$$\sum_{i=1}^{J} y_{ij} \neq s_{2i} \quad i = 1, \cdots, I$$

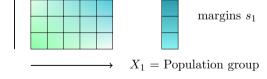


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What is raking?







After raking, the raked values β_{ij} in the updated table β sum correctly to the values in the margins s_1 and s_2 .

$$\sum_{i=1}^{I} \beta_{ij} = s_{1j} \quad j = 1, \cdots, J$$

$$\sum_{j=1}^{J} \beta_{ij} = s_{2i} \quad i = 1, \cdots, I$$

Note: For the problem to have a solution, we need the margins to be consistent:

$$\sum_{j=1}^{J} s_{1j} = \sum_{i=1}^{I} s_{2i}$$



Global health example

- The observation table may be the number of obesity cases for each population group *i* and each county *j*. The margins are the number of obesity cases for the entire population for each county *j* and the number of cases for each population group *i* for the entire state.
- For some reason (e.g. errors in data collection, the table is the output of a model that does not
 include the constraints on the margins), the partial sums on the observations do not match the
 margins.
- We trust more the margins than the observations.



Raking as an optimization problem

 $y \in \mathbb{R}^p$ is the vectorized observation table.

 $s \in \mathbb{R}^k$ are the known margins, i.e. the known partial sums on the table.

 $A \in \mathbb{R}^{k \times p}$ summarizes how to compute the partial sums.

 $\beta \in \mathbb{R}^p$ are the unknown raked values.

 $w \in \mathbb{R}^p$ are raking weights chosen by the user.

 f^w is a separable, derivable, positive, strictly convex function chosen by the user.

$$\min_{\beta \in \mathbb{R}^p} f^w\left(\beta;y\right) \quad \text{s.t.} \quad A\beta = s \quad \text{with} \quad f^w\left(\beta;y\right) = \sum_{i=1}^p w_i f_i\left(\beta_i,y_i\right) \quad \text{and e.g.} \quad A = \begin{pmatrix} I_J \otimes \mathbb{1}_I^T \\ \mathbb{1}_J^T \otimes I_I \end{pmatrix}$$

Note: We need to ensure that all the constraints are consistent and we trim the redundant constraints such that rank $(A \in \mathbb{R}^{k \times p}) = k \leq p$.



Dual formulation

$$\mathcal{P}:\quad \min_{\beta\in\mathbb{R}^{p}}f^{w}\left(\beta,y\right)\quad \text{s.t}\quad A\beta=s$$

$$\mathcal{L}:\quad f^{w}\left(\beta,y\right)+\lambda^{T}(A\beta-s)$$

$$\mathcal{D}: \quad \min_{\lambda \in \mathbb{R}^k} f^{w*} \left(-A^T \lambda \right) + \lambda^T s$$

As $k \le p$, we decrease the dimension of the problem by using the dual formulation instead of the primal formulation.

We solve the dual problem using Newton's method.

Common distance functions

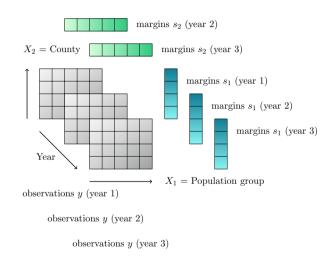
	Distance $f_{i}\left(\beta_{i};y_{i}\right)$	Solution	Note
χ^2	$rac{1}{2y_i}\left(eta_i-y_i ight)^2$	$eta^* = y \odot \left(1 - rac{1}{w} \odot A^T \lambda^* ight)$	Solved in 1 iteration.
			The raked values have
Entropic	$\boldsymbol{\beta}_i \log \left(\frac{\beta_i}{y_i}\right) - \boldsymbol{\beta}_i + \boldsymbol{y}_i$	$\beta^* = y \odot \exp\left(-\frac{1}{w} \odot A^T \lambda^*\right)$	the same sign as
			the initial observations.
Logit	$(\beta_i - l_i) \log \tfrac{\beta_i - l_i}{y_i - l_i} + (h_i - \beta_i) \log \tfrac{h_i - \beta_i}{h_i - y_i}$	$\beta^* = \frac{l \odot (h-y) + h \odot (y-l) \odot e^{-\frac{1}{w} \odot A^T \lambda^*}}{(h-y) + (y-l) \odot e^{-\frac{1}{w} \odot A^T \lambda^*}}$	The raked values stay
			between l_i and h_i
			when we rake
			prevalence observations.



Prior ordinal constraints

We are given observations and margins for n different years \rightarrow We can solve n independent raking problems:

$$\begin{split} & \min_{\beta_1 \in \mathbb{R}^p} f^w \left(\beta_1, y_1\right) & \text{ s.t. } & A_1 \beta_1 = s_1 \\ & \vdots \\ & \min_{\beta_n \in \mathbb{R}^p} f^w \left(\beta_n, y_n\right) & \text{ s.t. } & A_n \beta_n = s_n \end{split}$$



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Prior ordinal constraints

We want the raking process to preserve the sign of the trend observed between year 1 and year 2, year 2 and year 3, and so on and so forth until year n.

 $(n-1)\,p$ additional constraints must then be added:

$$\begin{split} \left(\beta_{1,i}-\beta_{2,i}\right)\left(y_{1,i}-y_{2,i}\right) &\geq 0 \quad \forall i=1,\cdots,p \\ \vdots \\ \left(\beta_{n-1,i}-\beta_{n,i}\right)\left(y_{n-1,i}-y_{n,i}\right) &\geq 0 \quad \forall i=1,\cdots,p \end{split}$$

The raking problem becomes:

$$\begin{split} & \min_{\beta_1,\cdots,\beta_n} f^w\left(\beta_1,y_1\right) + \cdots + f^w\left(\beta_n,y_n\right) \\ & \text{s.t.} \quad \begin{cases} A_1\beta_1 = s_1, \\ \cdots \\ A_n\beta_n = s_n, \\ -\left(\beta_1 - \beta_2\right) \odot \left(y_1 - y_2\right) \leq 0 \\ \cdots \\ -\left(\beta_{n-1} - \beta_n\right) \odot \left(y_{n-1} - y_n\right) \leq 0 \end{cases} \end{split}$$

Prior ordinal constraints

We end up with a minimization problem with the same form as before:

Inequality constraints

$$\min_{\beta \in \mathbb{R}^{n_p}} f^w \left(\beta, y\right) \quad \text{s.t.} \quad \begin{cases} A\beta = s, \\ C\beta \leq c \end{cases}$$

The feasible set may be empty.

Penalty

$$\min_{\beta \in \mathbb{R}^{n_p}} f^w\left(\beta,y\right) + L\left(c - C\beta,\alpha\right) \quad \text{s.t.} \quad A\beta = s$$

 α is a penalty parameter and L can be the logistic loss:

$$L^{\text{logit}}\left(x\right) = \sum_{i=1}^{m} \log\left(1 + \exp\left(-x_{i}\right)\right)$$

Given:

- $\Sigma_y \in \mathbb{R}^{p \times p}$, the covariance matrix of the observations vector y,
- $\Sigma_s \in \mathbb{R}^{k \times k}$, the covariance matrix of the margins vector s and
- $\Sigma_{ys} \in \mathbb{R}^{p \times k}$, the covariance matrix of y and s,

find:

• $\Sigma_{\beta^*} \in \mathbb{R}^{p \times p}$, the covariance matrix of the estimated raked values β^* .



The primal problem:

$$\min_{\beta \in \mathbb{R}^{p}} \max_{\lambda \in \mathbb{R}^{k}} f^{w}\left(\beta, y\right) + L\left(c - C\beta, \alpha\right) + \lambda^{T}\left(A\beta - s\right)$$

can also be written:

$$F(\beta, \lambda; y, s) = \begin{bmatrix} \nabla_{\beta} f^{w}\left(\beta, y\right) - C^{T} \nabla_{x} L\left(c - C\beta, \alpha\right) + A^{T} \lambda \\ A\beta - s \end{bmatrix} = 0$$

and has solution:

$$\beta^* = \phi(y, s) \text{ with } \phi: \mathbb{R}^{p+k} \to \mathbb{R}^p$$

We get:

$$\Sigma_{\beta^*} = \phi'_{ys}\left(y, s\right) \Sigma \phi'^{T}_{ys}\left(y, s\right)$$

with:

$$\phi_{ys}^{\prime}\left(y,s\right) = \begin{pmatrix} \frac{\partial\beta^{*}}{\partial y} & \frac{\partial\beta^{*}}{\partial s} \end{pmatrix} = \begin{pmatrix} \frac{\partial\phi_{1}}{\partial y_{1}}\left(y,s\right) & \dots & \frac{\partial\phi_{1}}{\partial y_{p}}\left(y,s\right) & \frac{\partial\phi_{1}}{\partial s_{1}}\left(y,s\right) & \dots & \frac{\partial\phi_{1}}{\partial s_{k}}\left(y;s\right) \\ \vdots & & \vdots & & \vdots \\ \frac{\partial\phi_{p}}{\partial y_{1}}\left(y,s\right) & \dots & \frac{\partial\phi_{p}}{\partial y_{p}}\left(y,s\right) & \frac{\partial\phi_{p}}{\partial s_{1}}\left(y,s\right) & \dots & \frac{\partial\phi_{p}}{\partial s_{k}}\left(y,s\right) \end{pmatrix}$$

and:

$$\Sigma = \begin{pmatrix} \Sigma_y & \Sigma_{ys} \\ \Sigma_{ys}^T & \Sigma_s \end{pmatrix}$$

Implicit Function Theorem: When differentiating the primal problem $F(y,s;\phi(y,s))=0$ at the solution (β^*,λ^*) , we get:

$$\left[D_{\beta,\lambda}F(y,s;\beta^*,\lambda^*)\right]\left[D_{y,s}\phi\left(y,s\right)\right]+\left[D_{y,s}F(y,s;\beta^*,\lambda^*)\right]=0$$

Knowing $D_{\beta,\lambda}F$ and $D_{u,s}\phi$, we can compute:

$$D_{y,s}\phi = \begin{pmatrix} \frac{\partial \beta^*}{\partial y} & \frac{\partial \beta^*}{\partial s} \\ \frac{\partial \lambda^*}{\partial y} & \frac{\partial \lambda^*}{\partial s} \end{pmatrix} \text{ and } \phi'_{ys}\left(y,s\right) = \begin{pmatrix} \frac{\partial \beta^*}{\partial y} & \frac{\partial \beta^*}{\partial s} \end{pmatrix}$$

We have:

$$D_{\beta,\lambda}F = \begin{pmatrix} \nabla_{\beta}^{2}f^{w}\left(\beta^{*},y\right) + C^{T}\nabla_{x^{2}}L\left(c - C\beta^{*},\alpha\right)C & A^{T} \\ A & 0_{k\times k} \end{pmatrix}$$



We denote:

$$\left[A*b\right]_{ij} = \sum_{j=1}^{n} A_{ijk}b_{j} \quad \text{for} \quad A \in \mathbb{R}^{m \times n \times p} \quad , \quad \beta \in \mathbb{R}^{n} \quad \text{and} \quad A*b \in \mathbb{R}^{m \times p}$$

$$\left[\nabla_{y}A\right]_{ijk} = \frac{\partial A_{ij}\left(y\right)}{\partial y_{k}}$$

We get:

$$D_{y,s}F = \begin{pmatrix} \nabla_{\beta y}^2 f^w \left(\beta^*;y\right) - \left[\nabla_y C^T\right] \nabla_x L\left(c - C\beta^*,\alpha\right) + C^T \nabla_x^2 L\left(c - C\beta^*,\alpha\right) \left[\nabla_y C * \beta\right] & 0_{p \times k} \\ 0_{k \times p} & -I_{k \times k} \end{pmatrix}$$

Some figures



Questions?

PyPI: https://pypi.org/project/raking/

GitHub: https://github.com/ihmeuw-msca/raking

