# Forecasting Trade Balance in Goods and Services for the United States

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#### 1 Introduction

The economic variable that we chose to forecast is the U.S. International Trade Balance in Goods and Services. International Trade Balance in Goods and Services attempts to accurately measure the trade balance of the United States, which is the difference between imports and exports measured in the millions of US dollars (\$).

This economic time series also captures what the top exports and imports in the United States were in a given month. However, for the purpose of the forecast, we will be looking at the value of the Trade Balance for the coming 12 months. To provide a more complete picture of the forecasts, we look at both point and interval forecasts.

#### 2 Data and Metrics

The original publisher of the data we use is the Bureau of Economic Analysis (BEA). However, we obtained the data from the Federal Reserve Economic Data (FRED)[5] using the FRED code BOPGSTB.

For the purpose of comparing different models we use for our data, the metric we use is the Akaike Information Criterion (AIC) to choose the best forecasting model. Since we are just beginning to work out a model for the Trade Balance time series, we believe that all models we come up with are just approximations of the true model, and our goal here is to find the best approximation, rather than finding the true model[2]. Moreover, we do not have access to out-of-sample data. These are the reasons for us to choose the AIC criterion over others like the Bayesian Information Criterion (BIC).

#### 3 Methods and Models<sup>1</sup>

To begin exploring the nature of the time series data, we look at the auto-correlation function and the partial auto-correlation function and plot correlograms for the same. We see that the auto-correlation correlogram shows a slow decay trend and the partial auto-correlation correlogram seems to show no fixed trend (see Appendix A), except for indicating 7 significant lags after which the subsequent lags become insignificant. These trends suggest that we should look at an Autoregressive (AR) model for our forecast[4].

Since the partial auto-correlation correlogram shows us that the first 7 lags would be important to look at for our AR model, we went ahead and ran the ARIMA(7, 0, 0)

<sup>&</sup>lt;sup>1</sup>All code used to generate the results presented in the paper as well as the plots in the appendix can be found on https://github.com/AGoyal0512/forecasting\_project

model on the time series data. On doing so, we saw that only the p-values for the first 4 lags were significant, and the AIC for this model was 6838.614 (See Appendix B). Updating the forecast model to just an autoregressive model of order 4 or ARIMA(4, 0, 0), we got that all the four lags are significant, and more importantly, our AIC value was also lower at 6835.714, which was the motivating factor for us to move forward with this model.

Before proceeding, we wanted to check for stationarity in our model, and see if we need any fixes to make the series stationary if it is non-stationary. For this, we used the Augmented Dickey-Fuller (ADF) test[1]. In this test, our null-hypothesis was that of assuming non-stationarity in the time series, while the alternative was the opposite, assuming stationary. We decided to use a 0.05 size-of-test for this hypothesis test.

When we ran the ADF test on our original time series, we got a p-value of 0.922304, which means that we clearly fail to reject the null hypothesis for non-stationarity, which implies that there is a good chance that our series possesses non-stationarity, and that our series has a unit root. To fix this, we decided to use the method of differencing, and took a one-differenced time series of our original data, and then re-ran the ADF test on this series (See Appendix C). This time we obtained a p-value of almost 0.000, which indicates that we have a strong evidence to reject the null hypothesis and conclude that this new series is stationary. So we now decided to integrate the one-differenced component in our ARIMA model, to get the ARIMA(4, 1, 0). The lower AIC value of 6807.724 confirmed that we were thinking in the right direction (See Appendix D).

Laslty, we saw that on running the ARIMA(4, 1, 0) model, the  $4^{th}$  lag became insignificant with a p-value of greater than 0.05. So we decided to drop that lag and settle on an ARIMA(3, 1, 0) model, which had roughly a similar AIC value, but a lower BIC value<sup>2</sup>, which was an indication that this model might be closer to what the true forecasting model for this time series looks like. See Appendix D for a detailed description of the model and its parameter estimates.

#### 4 Forecasts and Discussions

Having settled on the ARIMA(3, 1, 0) model, our next step was to check if the model actually works well on the data and whether it would be suitable for our out-of-sample forecasts. For this, we decided to split our time series into a testing set and a training set with a 2:1 train-test split. We trained our model on the training set, and then used this model to predict values of the in-sample dates in the testing set.

On performing predictions on the test dataset, our Root Mean Squared Error[3] was 3851.806, which is a very reasonable error given that our observations are in the order of tens of thousands to almost a hundred thousand. Moreover, on plotting our predictions on top of the test data, we see that the forecast model performs a good job (See Appendix F).

Our first out-of-sample point forecast for the month of March is a trade balance in

<sup>&</sup>lt;sup>2</sup>The reason we used BIC here is just for it to act as a sort of a tie-breaker, since the AIC values are very similar for being able to draw conclusions

goods and services of -\$ 89756.094707 million for the United States.

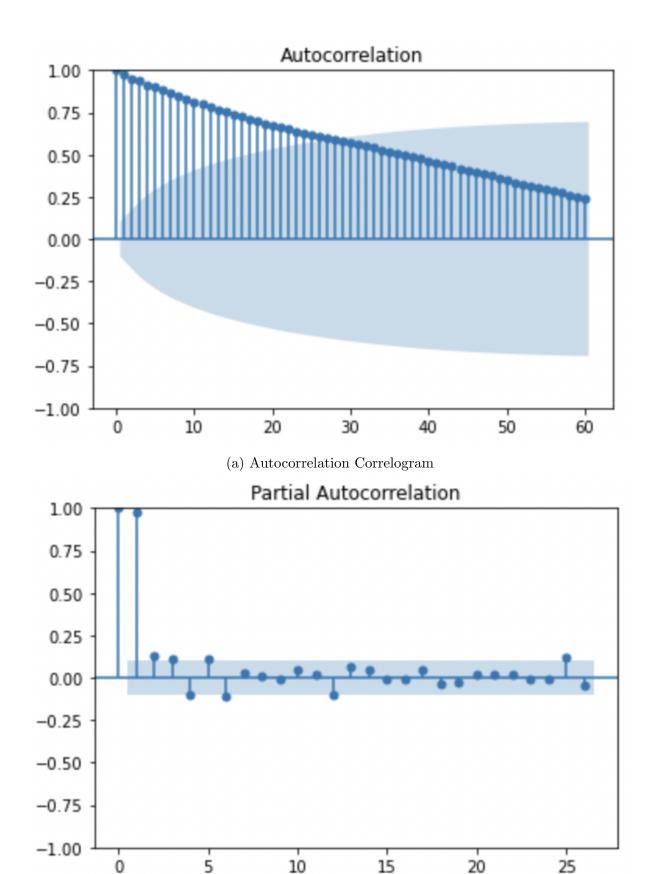
The point forecasts for the one upcoming year out-of-sample, for each of the 12 months lie at a trade balance in goods and services of -\$ 89756.094707 million, -\$ 89427.186116 million, -\$ 91032.214411 million, -\$ 90702.505979 million, -\$ 90600.974737 million, -\$ 90999.122135 million, -\$ 90820.712238 million, -\$ 90819.462587 million, -\$ 90918.924467 million, -\$ 90851.954543 million, -\$ 90863.320926 million, and -\$ -90886.540965 million.

The 95% interval forecasts along with the point forecasts for the 12 months are reported in a table in Appendix E. The plots showing these forecasts with an overall plot of a part of the original time series with the 12 out-of-sample predictions can also be found in Appendix F. All observations are in the units of millions of United States \$.

#### References

- [1] How to Check if Time Series Data is Stationary with Python: Machine Learning Mastery. URL: https://machinelearningmastery.com/time-series-data-stationary-python/. (accessed: 05.02.2022).
- [2] Introduction to AIC Akaike Information Criterion: Towards Data Science. URL: https://towardsdatascience.com/introduction-to-aic-akaike-information-criterion-9c9ba1c96ced. (accessed: 05.02.2022).
- [3] RMSE: Root Mean Square Error Statistics How To. URL: https://www.statisticshowto.com/probability-and-statistics/regression-analysis/rmse-root-mean-square-error/. (accessed: 05.02.2022).
- [4] Time Series Analysis: Identifying AR and MA using ACF and PACF Plots: Towards Data Science. URL: https://towardsdatascience.com/identifying-arand-ma-terms-using-acf-and-pacf-plots-in-time-series-forecasting-ccb9fd073db8. (accessed: 05.02.2022).
- [5] U.S. Census Bureau and U.S. Bureau of Economic Analysis, Trade Balance: Goods and Services, Balance of Payments Basis [BOPGSTB]. URL: https://fred.stlouisfed.org/series/BOPGSTB. (accessed: 05.02.2022).

## Appendix A



(b) Partial Autocorrelation Correlogram

# Appendix B

#### SARIMAX Results

Dep. Vari	lable:	BOP	GSTB No.	Observations	:	362
Model:		ARIMA(7, 0	, 0) Log	Likelihood		-3410.307
Date:	$\mathbf{T}$	ue, 03 May	2022 AIC			6838.614
Time:		15:2	0:05 BIC			6873.639
Sample:		01-31-	1992 HQIC	:		6852.538
		- 02-28-	2022			
Covarianc	ce Type:		opg			
=======		std err		P>   z	-	0.975]
const	-3.693e+04					7311.598
ar.L1	0.6862	0.040	17.291	0.000	0.608	0.764
ar.L2	0.2264	0.052	4.354	0.000	0.124	0.328
ar.L3	0.3215	0.053	6.081	0.000	0.218	0.425
ar.L4	-0.1406	0.064	-2.207	0.027	-0.265	-0.016
ar.L5	-0.0413	0.064	-0.640	0.522	-0.168	0.085
	-0.0397			0.564	-0.175	0.095
ar.L7	-0.0171	0.057	-0.301	0.763	-0.129	0.094
sigma2	8.916e+06	55.706	1.6e+05	0.000	8.92e+06	8.92e+06
Ljung-Box (L1) (Q): 0.01 Jarque-Bera (JB): 164						
			0.91	Prob(JB):	, ,	(
, – ,			6.64	Skew:		(
<pre>Prob(H) (two-sided):</pre>			0.00	Kurtosis:		6

#### (a) ARIMA(7, 0, 0) model

#### SARIMAX Results

Dep. Varia	ble:	BOP	GSTB N	lo.	Observations:	:	362	
Model:		ARIMA(4, 0	, 0) L	og :	Likelihood		-3411.857	
Date:	r	ue, 03 May	2022 A	IC			6835.714	
Time:		15:2	0:05 B	BIC			6859.064	
Sample:		01-31-	1992 н	QIC			6844.997	
		- 02-28-	2022					
Covariance	Type:		opg					
	coef	std err		z	P>   z	[0.025	0.975]	
const	-3.693e+04	3.34e+04	-1.1	.05	0.269	-1.02e+05	2.86e+04	
ar.L1	0.7072	0.036	19.8	04	0.000	0.637	0.777	
ar.L2	0.2104	0.047	4.4	41	0.000	0.118	0.303	
ar.L3	0.2908	0.043	6.7	37	0.000	0.206	0.375	
ar.L4	-0.2114	0.036	-5.8	81	0.000	-0.282	-0.141	
sigma2	8.93e+06	154.266	5.79e+	04	0.000	8.93e+06	8.93e+06	
Ljung-Box (L1) (Q): 0.22 Jarque-Bera (JB): 158.86								
Prob(Q):	(11) (2).		0.6		Prob(JB):	(02).		0.00
, ,	asticity (H)	•	6.5		Skew:			0.12
Prob(H) (t		•	0.0		Kurtosis:			6.24
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(b) ARIMA(4, 0, 0) model

### Appendix C

ADF Statistic: -0.322179

p-value: 0.922304

Critical Values:

1%: -3.449

5%: -2.870

10%: -2.571

(a) Results of the Augmented Dickey-Fuller Test BEFORE Differencing

ADF Statistic: -10.442180

p-value: 0.000000

Critical Values:

1%: -3.449

5%: -2.870

10%: -2.571

(b) Results of the Augmented Dickey-Fuller Test AFTER Differencing

# Appendix D

#### SARIMAX Results

Dep. Variable:	BOPGS	TB No.	Observations	:	362	
Model:	ARIMA(4, 1,	0) Log	Likelihood		-3398.862	
Date:	Tue, 03 May 20	22 AIC			6807.724	
Time:	15:20:	05 BIC			6827.168	
Sample:	01-31-19	92 HQIC	:		6815.455	
	- 02-28-20	22				
Covariance Type:	0	pg				
					========	
coef	std err	z	P>   z	[0.025	0.975]	
ar.L1 -0.3094	0.038	-8.042	0.000	-0.385	-0.234	
ar.L2 -0.0711	0.041	-1.719	0.086	-0.152	0.010	
ar.L3 0.2366	0.040	5.908	0.000	0.158	0.315	
ar.L4 0.0808	0.052	1.559	0.119	-0.021	0.182	
sigma2 8.9e+06	4.41e+05	20.201	0.000	8.04e+06	9.76e+06	
Ljung-Box (L1) (Q):		0.05	Jarque-Bera	 (JB):	173	3.90
Prob(Q):		0.82	Prob(JB):	(02):		0.00
Heteroskedasticity (H	):	6.59	Skew:			0.21
Prob(H) (two-sided):			Kurtosis:			5.37
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#### (a) ARIMA(4, 1, 0) model

#### SARIMAX Results

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Dep. Varia	able:	BOPGS	TB No.	Observation	s:	362	
Model:		ARIMA(3, 1,	0) Log	Likelihood		-3399.948	
Date:	Tu	ie, 03 May 20	22 AIC			6807.897	
Time:		15:20:	05 BIC			6823.452	
Sample:		01-31-19	92 HQIC			6814.082	
		- 02-28-20	22				
Covariance	Type:	o	pg				
========							
	coef	std err	Z	P>   z	[0.025	0.975]	
ar.L1		0.036				-0.223	
ar.L2	-0.0810	0.041	-1.990	0.047	-0.161	-0.001	
ar.L3	0.2107	0.036	5.798	0.000	0.139	0.282	
sigma2	8.933e+06	4.43e+05	20.149	0.000	8.06e+06	9.8e+06	
Ljung-Box	(L1) (Q):		0.23	Jarque-Ber	a (JB):	 16	2.88
Prob(Q):	, , , , , , ,		0.63	Prob(JB):	, ,		0.00
, ,	lasticity (H):		6.47	Skew:			0.16
	wo-sided):		0.00	Kurtosis:			6.27

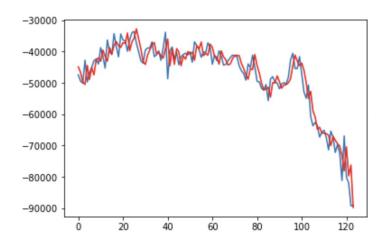
(b) ARIMA(3, 1, 0) model

# Appendix E

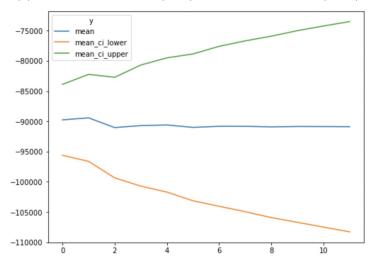
У	mean	mean_ci_lower	mean_ci_upper
0	-89756.094707	-95622.056400	-83890.133013
1	-89427.186116	-96615.360567	-82239.011665
2	-91032.214411	-99351.224045	-82713.204778
3	-90702.505979	-100711.305326	-80693.706632
4	-90600.974737	-101699.332363	-79502.617112
5	-90999.122135	-103138.178740	-78860.065531
6	-90820.712238	-104041.248931	-77600.175544
7	-90819.462587	-104943.319082	-76695.606092
8	-90918.924467	-105917.957144	-75919.891790
9	-90851.954543	-106699.150786	-75004.758299
10	-90863.320926	-107490.361646	-74236.280206
11	-90886.540965	-108268.788247	-73504.293683

Table containing Point and Interval Forecasts for 12 months

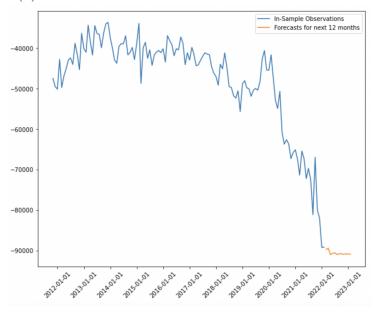
## Appendix F



(a) Model Predictions (Red) with Testing Data (Blue)



(b) Plot of Point and Interval Forecasts for 12 months



(c) Overall plot for Last 12 Years with Forecasts