

Concentration Review

(Ref: Chap 5, Bandit Algorithms).

$\{x_1, x_2, \dots, x_n, \dots\}$ sequence of

iid r.v.s \rightarrow random variables

\hookrightarrow independent, identically distributed

Law of large numbers:

$$\hat{\mu}_n = \frac{1}{n} \sum_{i=1}^n x_i \longrightarrow E[x] \quad \text{a.s.}$$

Notation: $E[x] = \mu, \quad \text{Var}(x) = \sigma^2.$

Question: With n samples, $\{x_1, \dots, x_n\}$,

how "close" is the empirical estimate of
the mean to the true mean?

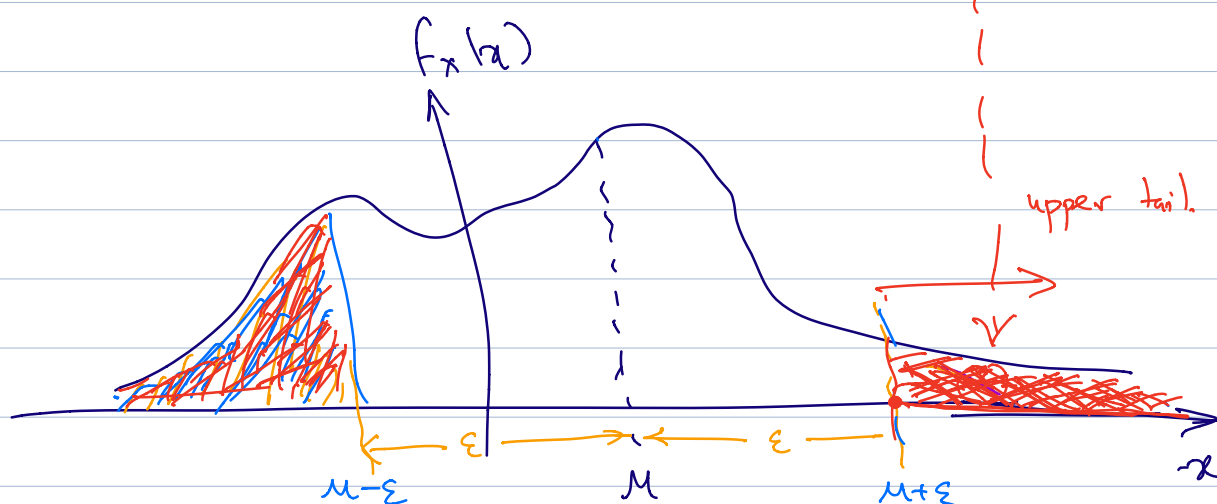
i.e., $\hat{\mu}_n = \frac{1}{n} \sum_{i=1}^n x_i$ a r.v.
(empirical estimate)

$\mu = E[x]$ a fixed number.

$$P(\hat{\mu}_n - \mu > \varepsilon) \leq \boxed{?}$$

$$P(|\hat{\mu}_n - \mu| > \varepsilon) \leq \boxed{?}$$

tail
probability



Why do we care: In bandit algorithms, past sample of rewards allows us to estimate "goodness" of actions. Concentrations permit us to quantify how much we trust these estimates.

Basic inequalities

Markov Inequality: $X \geq 0$ a non-negative r.v.
and $\varepsilon > 0$. Then,

$$P(X > \varepsilon) \leq \frac{E[X]}{\varepsilon}$$

Chebyshev's Inequality: X a r.v. with $E[X] = \mu$,
 $\text{Var}(X) = \sigma^2$,

$$P(|X - \mu| > \varepsilon) \leq \frac{\sigma^2}{\varepsilon^2}.$$

Chernoff Bound: X a r.v. with MGF $M_X(\theta)$,
Moment Generating function

$$M_X(\theta) = E[e^{\theta X}]$$

exists atleast for
"small" θ

Then,
$$P(X > \varepsilon) \leq e^{-I(\varepsilon)}$$

$$\boxed{\varepsilon > E[X]}$$

$$I(\varepsilon) = \max_{\theta} \{ \varepsilon \theta - \ln M_X(\theta) \}.$$

↪ rate function. = Legendre-Fenchel Transform

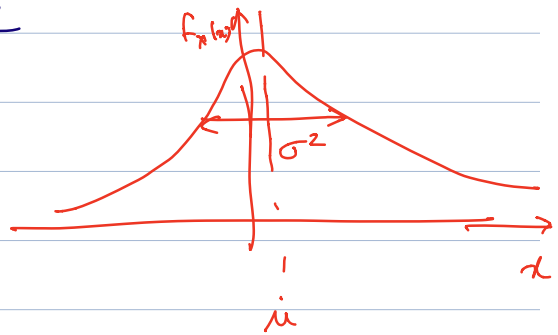
of "the log-MGF

Most often, we will simplify and assume noise terms are Subgaussian

Recall: $X \sim N(\mu, \sigma^2)$ \rightarrow Gaussian r.v.
aka Normal r.v.
 \hookrightarrow "distributed as"
"has the pdf of"

if

$$f_X(x) = \frac{1}{\sqrt{2\pi\sigma^2}} e^{-\frac{(x-\mu)^2}{2\sigma^2}}$$



For $Z \sim N(0, \sigma^2)$

$$\underbrace{E[e^{\theta Z}]}_{\text{MGF}} = e^{\theta^2 \sigma^2 / 2}$$

$$\boxed{\begin{aligned} P(X > \mu + \varepsilon) \\ \leq e^{-\varepsilon^2 / 2\sigma^2} \end{aligned}}$$

MGF = Moment Generating Function

$\gamma \sim \text{SubGaussian}(\sigma)$ (also σ -subgaussian) if

$$(a) \quad E[\gamma] = 0$$

$$(b) \quad E[e^{\theta\gamma}] \leq e^{\theta^2\sigma^2/2}$$

Properties $\gamma \sim \text{subgaussian}(\sigma)$

$$(1) \quad P(\gamma \geq \varepsilon) \leq e^{-\varepsilon^2/2\sigma^2}$$

$$(2) \quad P\left(\gamma \geq \sqrt{2\sigma^2 \ln(1/\delta)}\right) \leq \delta,$$

for any $\delta > 0$

$$(3) \quad P\left(|\gamma| \geq \sqrt{2\sigma^2 \ln(2/\delta)}\right) \leq \delta \quad \text{for}$$

any $\delta > 0$.

$$(4) \quad \text{Var}(\gamma) \leq \sigma^2$$

$$(5) \quad Z = c\gamma \sim \text{subgaussian}(\underbrace{|c|}_{\text{std-dev}}\sigma)$$

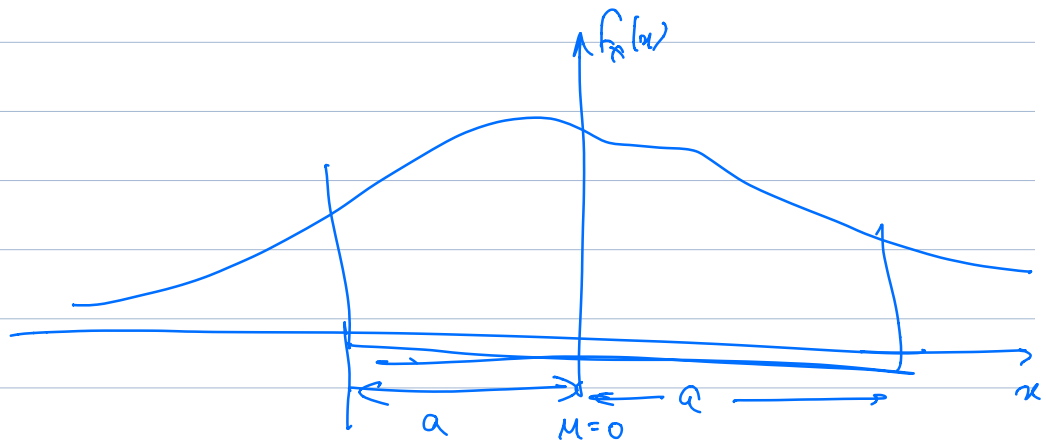
(Proof: See chapter 5, Bandit Algorithms).

⑥ $Z_1 \sim \sigma_1$ -subgaussian, $Z_2 \sim \sigma_2$ -subgaussian

$(Z_1 + Z_2) \sim \sqrt{\sigma_1^2 + \sigma_2^2}$ -subgaussian.

Corollary: $\{X_i\} \sim 1$ -subgaussian, iid

$\frac{1}{m} \sum_{i=1}^m X_i \sim \frac{1}{\sqrt{m}}$ -subgaussian



$$P\left(\underbrace{\frac{1}{m} \sum_{i=1}^m X_i}_{\text{sample mean}} > \varepsilon\right) \leq e^{-\frac{\varepsilon^2}{2\left(\frac{1}{\sqrt{m}}\right)^2}}$$

$$= e^{-\varepsilon^2/2/m}$$

$$= \underline{e^{-(m\varepsilon^2/2)}}.$$