

Lecture 19: Mixture Models

20/10/2022

Lecturer: Abir De

Scribe: Course Team

1 Introduction

We observe a data set $D = \{X_i\}_{i=1}^N$, where each $X_i = x_i$ is being sampled from one of the K mixture components.

Each of the mixture component is a multivariate Gaussian density with its own parameters $\theta_k = \{\mu_k, \Sigma_k\}$

$$p_k(x_i | \theta_k) = \frac{1}{(2\pi)^{d/2} |\Sigma_k|^{1/2}} e^{-\frac{1}{2}(x_i - \mu_k)^t \Sigma_k^{-1} (x_i - \mu_k)}$$

We now have to estimate the parameters of the K mixture components, θ_k and the mixture weights, which represent the probability that a randomly selected \bar{x} was generated by k^{th} component, $\pi_k = P(c(\bar{x}) = k)$, where $\sum_{k=1}^K \pi_k = 1$.

2 Computing posterior distribution $P(c(X_i) = k | X_i)$

Using initial estimates for ω , we obtain the posterior in the following way -

$$P(c(X_i) = k | X_i, \omega) = \frac{P(\mathbf{X}_i | c(X_i) = k, \theta_k) \cdot P(c = k)}{\sum_m P(\mathbf{X}_i | c(X_i) = m, \theta_m) P(c = m)} = \frac{N(X_i; \theta_k) \pi_k}{\sum_m N(X_i; \theta_m) \pi_m}$$

This follows a direct application of Bayes rule. These membership weights reflect the uncertainty, given $X_i = x_i$ and ω , about which of the K components generated vector $X_i = x_i$.

3 Maximum Likelihood Estimation

The complete set of parameters for a mixture model with K components is -

$$\omega = \{\pi_1, \pi_2, \dots, \pi_K, \theta_1, \dots, \theta_K\}$$

We now maximize the likelihood of data, $P(D) = P(X_1 = x_1, X_2 = x_2, \dots, X_N = x_N)$ w.r.t ω .

$$\begin{aligned} P(D) &= \prod_{i=1}^N P(\mathbf{X}_i = \mathbf{x}_i) \\ \implies \log(P(D)) &= \sum_{i=1}^N \log(P(\mathbf{X}_i = \mathbf{x}_i)) \end{aligned}$$

We know that marginal probability of X_i is,

$$P(\mathbf{X}_i = \mathbf{x}_i) = \sum_{k=1}^K P(\mathbf{X}_i = \mathbf{x}_i \mid c(X_i) = k)P(c = k)$$

$$\implies P(\mathbf{X}_i = \mathbf{x}_i) = \sum_{k=1}^K P(\mathbf{X}_i = \mathbf{x}_i \mid c(X_i) = k)\pi_k$$

Using the above,

$$\log(P(D)) = \sum_{i=1}^N \log\left(\sum_{k=1}^K P(\mathbf{X}_i \mid c(X_i) = k)\pi_k\right)$$

Differentiating the above w.r.t π_k , μ_k and \sum_k , we obtain the new parameters (and using the equation presented in Section 2)-

$$\text{Let } N_k = \sum_{i=1}^N P(c(X_i) = k \mid X_i, \omega)$$

$$\pi_k^{new} = \frac{N_k}{N}$$

$$\mu_k^{new} = \left(\frac{1}{N_k}\right) \sum_{i=1}^N X_i \cdot P(c(X_i) = k \mid X_i, \omega)$$

$$\sum_k^{new} = \left(\frac{1}{N_k}\right) \sum_{i=1}^N P(c(X_i) = k \mid X_i, \omega) \cdot (X_i - \mu_k^{new}) \cdot (X_i - \mu_k^{new})^t$$

4 Iterative Procedure for Parameter Estimation

We now work on choosing a suitable initial prior for π_k . Entropy of a distribution is defined as $-\sum_{i=1}^N P(\mathbf{X}_i) * \log(P(\mathbf{X}_i))$ where \mathbf{X}_i are random variables of the distribution. In a K-means cluster distribution, we have $\pi_1, \pi_2, \dots, \pi_K$. In order to maximize the randomness, we assign each one of the random variables probability $1/K$.

Now, using the above initial prior for π_k , and some initial parameter estimates θ_k , we derive the posterior $P(c(X_i) = k \mid X_i)$ (membership weights) as presented in Section 2.

Using these new membership weights, we calculate the new π_k , μ_k and \sum_k using the equations given at the end of Section 3 (derived by differentiating the log likelihood).

Using these new parameter estimates, we calculate the new membership weights and repeat the steps again until the value of likelihood of data converges.

$$\text{Log likelihood of data} - \log \prod_{i=1}^N P(\mathbf{X}_i) = \sum_{i=1}^N \log \left(\sum_{k=1}^K P(\mathbf{X}_i | c(\mathbf{X}_i) = k) P(c = k) \right)$$

$$\text{Let } P_\omega = P(\mathbf{X}_i | c(\mathbf{X}_i) = k), P_c = P(c = k | \mathbf{X}_i)$$

$$\omega = \omega^{t-1}$$

$$\text{At time } t, \max_{\omega} \sum_{i=1}^N \log \left(\sum_{k=1}^K P_\omega P_c(\omega^{t-1}) \right) \text{ will give us the new parameter estimates } \omega$$

5 Representation in terms of Expectation

We can also represent the likelihood of data $\{\prod_{i=1}^N P(\mathbf{X}_i)\}$ as below.

$$\text{Now, } P(\mathbf{X}) = \sum_Z P(\mathbf{X}|Z)P(Z)$$

$$\text{implies } P(\mathbf{X}) = \mathbf{E}_Z[P(\mathbf{X}|Z)]$$

$$\text{Hence, } P(\mathbf{X}_i) = \mathbf{E}_c[P(\mathbf{X}_i | c)]$$

$$\prod_{i=1}^N P(\mathbf{X}_i) = \prod_{i=1}^N \mathbf{E}_c[P(\mathbf{X}_i | c)]$$

$$\text{Now, } \prod_{i=1}^N \sum_{k=1}^K P(\mathbf{X}_i | c = k) P(c = k)$$

$$\text{is equal to, } \sum_{k_1=1}^K \sum_{k_2=1}^K \sum_{k_3=1}^K \dots \sum_{k_N=1}^K \left(\prod_{i=1}^N P(\mathbf{X}_i | c = k_i) P(c = k_i) \right)$$

$$\prod_{i=1}^N P(\mathbf{X}_i) = \mathbf{E}_{(k_1, k_2, k_3, \dots, k_N)} \left[\prod_{i=1}^N P(\mathbf{X}_i | c = k_i) \right]$$

6 Mixture Model to K-Means iterative algorithm

Entropy of a distribution is defined as $S(X) = -\sum_{i=1}^N P(X_i) \log(P(X_i))$ where X_i are random variables of the distribution. Entropy is maximised when all of these probabilities are equal (easily proved with differentiation).

So we set the prior $w_k = 1/K$ for all k initially to maximise entropy in K-Means. We set random initial parameter estimates θ . In addition, for later iterations we set $P(c = k) = \mathbf{I}(c = k)$ where \mathbf{I} is the delta function.

Using maximum likelihood estimation of data, we calculate the new parameters and weights and stop when the likelihood converges.