# Introduction to Financial Econometrics Chapter 1: Describing Financial Series

Christophe Hurlin

Univ Orléans

May, 2019

### 1. Introduction

The outline of this chapter is the following:

Section 2: Financial Econometrics

Section 3: Financial Data

Section 4: Statistical Properties of Financial Data

**Section 5:** Stylized Facts

## 1. Introduction

Notations: In this chapter, I will (try to...) follow some conventions of notation.

Y	random variable
У	realization
$f_{Y}(y)$	probability density or mass function
$F_{Y}(y)$	cumulative distribution function
Pr ()	probability
у	vector
Υ	matrix

**Problem:** this system of notations does not allow to discriminate between a vector (matrix) of random elements and a vector (matrix) of non-stochastic elements (realization).



Abadir and Magnus (2002), Notation in econometrics: a proposal for a standard, Econometrics Journal.

# Section 2

# Financial Econometrics

## **Objectives**

- Define financial econometrics
- Define the notions of sample and data generating process
- Oistinguish parametric and non-parametric model
- Define the cross-sectional data
- Define the time series data
- Opening the panel data
- Summarize the general approach of financial econometrics

What is financial econometrics? Not an easy question!

The main objective of this course is to define this term!

# Definition (financial econometrics)

**Financial econometrics** is the application of statistical methods to financial market data. [..] It differs from other forms of econometrics because the emphasis is usually on analyzing the prices of financial assets traded at competitive, liquid markets.

Source: Wikipedia.

6 / 136

## **Objectives of Financial Econometrics**

- Combining finance theory with statistical theory.
- Transferring a theoretical (financial) model into a (financial) econometric model.
- Modelling financial data.
- Predicting financial variables as well as relations thereof.
- Developing specific model well-adapted to the stylized facts of financial data.
   Example: GARCH models

#### Is it useful?

"People working in the finance industry or research in the finance sector often use econometric techniques in a range of activities – for example, in support of portfolio management and in the valuation of securities. Financial econometrics is essential for risk management when it is important to know how often 'bad' investment outcomes are expected to occur over future days, weeks, months and years".

Source: Wikipedia.

## **Academic development**

Financial econometrics attracted substantial attention in recent years in both academia as well as financial practice.

- 2003: Nobel prize for Robert F. Engle for the GARCH models (cf chapter 6).
- 2003: Foundation of the Journal of Financial Econometrics.
- 2007: Foundation of the Society for Financial Econometrics (SoFiE).



#### Remark

There is sometimes a confusion between the following terms

- Financial econometrics bridges the gap between financial economics, statistics and mathematical finance.
- Empirical finance covers all the finance studies based on data analysis, e.g. Journal of Empirical Finance.
- Financial economics is a "highly empirical discipline, perhaps the most empirical among the branches of economics and even among the social sciences in general" (Campbell, Lo, and MacKinlay, 1997)

"Econometrics is the quantitative analysis of actual economic phenomena based on the concurrent development of theory and observation, related by appropriate methods of inference", P. A. Samuelson, T. C. Koopmans, and J. R. N. Stone (1954)

(Financial) econometrics is fundamentally based on four elements:

- A sample of data
- An econometric model
- An estimation method
- Some inference methods

## Question: Why using a sample?

- Let us assume that we want to study a characteristic / property x of the individuals of a population.
- The individuals (unit) of the population are not necessarily some persons: it can be firms, assets, countries, time index etc..
- The characteristic x may be quantitative (price, returns, total asset, etc.) or qualitative (default, sector, etc.)
- The characteristic x may be **stochastic** or **deterministic**.

## Definition (Population)

A **population** can be defined as including all people or items with the characteristic one wishes to understand.

- In most of cases, it is impossible to observe the entire statistical population, due to cost constraints, time constraints, constraints of geographical accessibility.
- A researcher would instead observe a statistical sample from the population in order to attempt to learn something about the population as a whole.

In most of cases, the sample is random:

# Definition (Probability sampling)

A probability sampling is a **sampling method** in which every unit in the population has a chance (greater than zero) of being selected in the sample.

**Consequence:** a sample is a collection of **random variables** even the characteristic x is deterministic.

sample: 
$$\{X_1, X_2, ..., X_N\}$$

## Example (random sample)

Let us consider a population of four firms and denote by  $\widetilde{x}$  the characteristic (assumed to be non stochastic) of the firm with:

$$\widetilde{x}_A = 80 \quad \widetilde{x}_B = 50 \quad \widetilde{x}_C = 40 \quad \widetilde{x}_D = 90$$

Consider a random sample of N = 2 firms denoted by

$$\left\{ \underbrace{X_1}_{\text{characteristic. of the first firm selected in the sample}}, X_2 \right\}$$

So we can obtain a realization

$$\{x_1, x_2\} = \{50, 80\}$$
 or  $\{x_1, x_2\} = \{90, 40\}$  or  $\{x_1, x_2\} = \{90, 90\}$  etc.

15 / 136

## Fact (random sample - data set)

The result of the probability sampling is a random sample, i.e. a collection of random variables  $X_1, X_2, ..., X_N$ . In general, only one realization of the sample is available: this is your data set!

$$\{x_1, x_2, ..., x_N\}$$



Remark When ones consider two or more variables, the notion of population is replaced by the concept of Data Generating Process (DGP)

## Definition (Data Generating Process)

A **Data Generating Process (DGP)** is the joint probability distribution that is supposed to characterize the entire population from which the data set has been drawn.

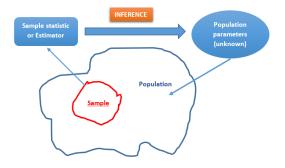
## Example (Data Generating Process)

Let us assume that there is a linear relationship between the returns  $R_t$  and  $R_{m,t}$  in the population, such that

$$\mathbb{E}\left(\left.R_{t}\right|R_{m,t}=r_{m}\right)=\alpha+\beta r_{m}$$

This relationship is the Data Generating Process for  $R_t$ .

The **challenge** of econometrics is to draw conclusions about a population (or DGP) after observing only **one** realization  $\{x_1, ... x_N\}$  of a random sample (your data set..).



In econometrics, data sets can be mainly distinguished in three types:

- Cross-sectional data
- 2 Time series data
- Panel data

#### Cross-sectional data

- Data for different entities: assets, portfolios, firms, and so forth.
- No time dimension (even if date of data collection varies somewhat across units, it is ignored).
- Order of data does not matter!

#### Time series data

- Data for a single entity (asset, firm, etc.) collected at multiple time periods.
   Repeated observations of the same variables (price, volume, etc.).
- Order of data is important!
- Observations are typically not independent over time;

## Panel data or longitudinal data

- Data for multiple entities (asset, firm, etc.) in which outcomes and characteristics of each entity are observed at multiple points in time.
- Combine cross-sectional and time series issues.
- Present several advantages with respect to cross-sectional and time series data (depending on the question of interest!).

# Definition (Econometric model)

A **model** specifies the statistical relationship that is believed to hold between the various economic quantities pertaining to a particular economic phenomenon under study.

## We can distinguish:

**Parametric model:** the relationship (joint probability distribution) between the dependent variable /vector Y and the explicative variables X is fully characterized by a set of parameters  $\theta$ 

$$Y = f(X; \theta) + \varepsilon$$

where link function f(.) is assumed to be known.

Non parametric and semi-parametric models: the link function can not be described using a finite number of parameters. The link function is assumed to be unknown and has to be estimated.

The general approach of (financial) econometrics is the following:

- Step 1: Specification of the model (Chapter 2)
- Step 2: Estimation of the parameters (Chapter 2)
- Step 3: Diagnostic tests (Chapter 3)
  - Significance tests;
  - Specification tests;
  - Backtesting tests;
  - etc.
- Step 4: Interpretation and use of the model (forecasting, historical study, etc.)
   (Chapter 4)

## **Key Concepts**

- Financial econometrics
- Sample, population and data generating process
- Parametric and non-parametric model
- Cross-sectional data
- Time series data
- Panel data
- General approach of financial econometrics

# Section 3

Financial Data

## **Objectives**

- 1 Introduce the various notions of asset prices
- Define the notion of sampling frequency
- 1 Distinguish high frequency data (intraday) vs. low frequency data
- **1** Define the **closing** and **opening** prices
- Define the adjusted (closing) prices
- Opening the simple (net) return
- Define the continuously compounded or log-return

27 / 136

## Liquid markets

Quantitative financial research mainly focuses on "liquid" financial markets, i.e. organized markets where transactions are frequent and the number of actors is large.

# Example ("liquid" markets)

Some example of "liquid" markets: foreign exchange market, organized futures markets, stock index markets and the market for large stocks

On these markets, prices are recorded several times a minutes or after each event (tick-by-tick data).

# Definition (prices)

Let  $P_t$  denotes the **price** of an asset at time t.

The prices may be observed (and recorded) at

- An irregular (sampling) frequency: tick-by-tick price observations, volume-event observations (prices observed when the volume exceeds a given threshold), price-events (transactions associated with significant price changes), etc.
- **4** A **regular (sampling) frequency**: the prices are observed every *m* periods of time.
- Banulescu D., Colletaz G., Hurlin C. and Tokpavi S. (2016), High Frequency Risk Measures, Journal of Forecasting, 35(3), 224-249.

29 / 136

## **Regular frequencies**

- Most of the financial econometric studies are based on regular frequencies.
- In general, ones make a distinction between intraday frequencies (called high frequencies) and lower frequencies because the statistical properties of the corresponding data are different and the econometric tools used to model them, too.

Table 1: Some examples of sampling frequencies used for financial data

High frequency (intra-day) data	Low frequency data
10 seconds	daily
1 minute	weekly
5 minutes	monthly
10 minutes	quarterly
30 minutes	semi-annual
1 hour	annual
etc.	etc.

# Example (sampling frequency for prices)

Plot the (closing) prices for Intel Corporation (Ticker: INTC) from January 2015 to September 2018, sampled at the daily, monthly, yearly frequencies, respectively.

Note: The data are available at https://finance.yahoo.com

Figure: Closing prices of Intel Corp, daily frequency (Jan 2013 - Sept 2018)



Source: Yahoo Finance

Figure: Closing prices of Intel Corp, monthly frequency (Jan 2013 - Sept 2018)



Source: Yahoo Finance

Figure: Closing prices of Intel Corp, yearly frequency (2010 - 2018)



Source: Yahoo Finance

#### Remark

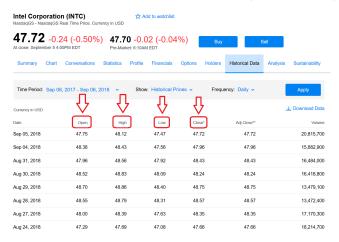
For a given sampling frequency (e.g., daily), the prices can be recorded at different time within the observational period.

For instance, for a daily sampling frequency, we can consider:

- The opening price is the price at which a security first trades upon the opening of an exchange on a trading day.
- The closing price is the final price at which a security is traded on a given trading day, representing the most up-to-date valuation of a security
- The **highest** (or lowest) price of the trading day.
- The adjusted closing prices have been treated to correct for splits and dividends.

In general, we consider the **end-of-period price**, i.e. the closing price in the case of a daily frequency.

Figure: Historical data available with Yahoo Finance (ticker: Intel Corp).



Source: Yahoo Finance

### Fact (prices vs returns)

Although **prices** are what we observe in financial markets, most empirical studies are based on **returns**.

- Returns are scale-free and have more attractive statistical properties than prices.
- Prices are, in general, non-stationary whereas returns are stationary.

We distinguish two types of return

- The simple (net) return
- The continuously compounded or log-return

### Definition (proportional return)

Holding a given asset for one period (from t-1 to t) yields the **one-period** (simple) return or proportional return:

$$\widetilde{R}_t \equiv \widetilde{R}_t[1] = \frac{P_t - P_{t-1}}{P_{t-1}}$$

where  $P_t$  is the (end-of-period) asset price at time t.

### Definition (log-return)

Holding a given asset for one period (from t-1 to t) yields the **continuously compounded** or **log-return**:

$$R_t \equiv R_t[1] = \ln\left(rac{P_t}{P_{t-1}}
ight) = p_t - p_{t-1}$$

where  $p_t$  is the log-price of the asset at time t.

#### Remark

Simple return and log-return are very similar if  $\tilde{R}_t$  is close to zero;

$$\begin{array}{lcl} R_t & = & \ln\left(\frac{P_t}{P_{t-1}}\right) \\ & = & \ln\left(1 + \left(\frac{P_t}{P_{t-1}} - 1\right)\right) \\ & = & \ln\left(1 + \widetilde{R}_t\right) \simeq \widetilde{R}_t \quad \text{if } \widetilde{R}_t \simeq 0 \end{array}$$

### Example (simple and log-returns)

The closing price of the equity Intel Corporation (INTC) was equal to \$47.17 the August 16, 2018 and \$46.9 the August 17, then the daily return is equal to

$$\widetilde{R}_{17/08} = \frac{46.9 - 47.7}{47.7} = -0.0058 = -0.58\%$$

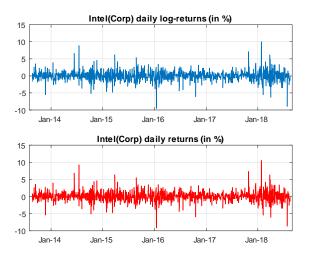
$$R_{17/08} = \ln\left(\frac{46.9}{47.7}\right) = -0.0057 = -0.57\%$$

## Example (Sample)

Consider a sample of **daily closing prices** for the equity Intel Corp. (ticker: INTC) from August 19, 2013 to August 17, 2018 (5 years). **Question:** compare the (simple) daily returns and the log-(daily) returns.

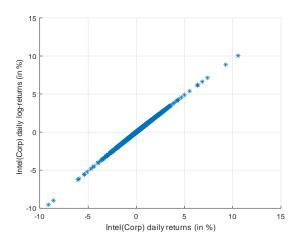
**Note:** The sample size T is equal to 1,259. Notice that each year corresponds to approximately 250 quotations, which implies about  $250 \times 5 = 1,250$  observations for 5 years.

Figure: Simple and log-returns for Intel(Corp), August 19, 2013 to August 17, 2018



44 / 136

Figure: Scatter-plot of simple and log-returns for Intel(Corp), August 19, 2013 to August 17, 2018



#### **Notations**

For the rest of this chapter and the rest of the course, we will keep the following set-up:

- We observe the end-of-period asset price  $P_t$  at time t.
- We consider the corresponding one-period (log-) returns  $R_t$ .
- The **frequency** (periodicity) of the observations can be daily, weekly, monthly, etc.
- We collect a T-sample of prices  $\{P_1, \ldots, P_T\}$ , where T is the sample size.

#### Definition

The **k-period log-return** is given by:

$$R_t[k] = \sum_{j=0}^{k-1} R_{t-j}[1].$$

**Proof:** Log-returns are quite convenient since multiplication becomes addition!

$$R_{t}[k] = \ln\left(\frac{P_{t}}{P_{t-k}}\right)$$

$$= \ln\left(\frac{P_{t}}{P_{t-1}} \times \frac{P_{t-1}}{P_{t-2}} \times \dots \times \frac{P_{t-k+2}}{P_{t-k+1}} \times \frac{P_{t-k+1}}{P_{t-k}}\right)$$

$$= \sum_{j=0}^{k-1} \ln\left(\frac{P_{t-j}}{P_{t-j-1}}\right) = \sum_{j=0}^{k-1} R_{t-j}[1]$$

## Definition (Excess return)

In many applications, it is convenient to use an excess return defined as

excess return = 
$$R_t - R_{f,t}$$

where  $R_{f,t}$  is a reference rate.

#### **Key Concepts**

- Prices and sampling frequency
- 4 High frequency data (intraday) vs low frequency data
- Olosing, opening, highest, lowest prices
- Adjusted prices
- Simple vs log-returns
- k-periods (log-) return
- Excess return

# Section 4

Statistical Properties of Financial Data

#### **Objectives**

- Introduce the concept of return distribution
- 2 Define the central and non-central moments of the returns distribution
- Oefine the mean, variance, skewness, and kurtosis
- Define the notions of sample, estimator, and estimate
- Study the sampling distribution of an estimator or a descriptive statistic

#### Notations

### Definition (distribution of returns)

We represent the (log-) returns by the **real-valued random variable**  $R_t$  (or  $R_t$  [k]). The distribution of  $R_t$  is called the **distribution of returns**, or returns distribution

- The **probability density function (pdf)** of  $R_t$  is denoted  $f_R(r)$ .
- The cumulative density function (cdf) of  $R_t$  is denoted  $F_R(r)$  such that

$$\Pr\left(R_{t} \leq r\right) \equiv F_{R}\left(r\right) = \int_{-\infty}^{r} f_{R}\left(x\right) dx$$

### Example (Normality assumption)

Let us assume that the return  $R_t$  has a normal distribution

$$R_t \sim \mathcal{N}\left(\mu, \sigma^2\right)$$

with  $\mu$  the expected (mean of) return and  $\sigma^2$  the variance of return. Then, the pdf and cdf of  $R_t$  are respectively defined by

$$\mathit{f}_{R}\left(r\right) = \frac{1}{\sigma}\phi\left(\frac{r-\mu}{\sigma}\right) = \frac{1}{\sigma\sqrt{2\pi}}\exp\left(-\frac{(r-\mu)^{2}}{2\sigma^{2}}\right)$$

$$F_R(r) = \Phi\left(\frac{r-\mu}{\sigma}\right) = \int_{-\infty}^{r} f_R(u) du$$

where  $\phi\left(.\right)$  and  $\Phi\left(.\right)$  respectively denote the pdf and cdf of the standard normal distribution.

→□▶→□▶→□▶→□ ● の○○

The distribution of returns can be characterized either by its pdf or by its moments.

### Definition (non-central and central moments)

The  $k^{th}$  non-central moment of  $R_t$  is given by

$$m_k = \mathbb{E}\left(R_t^k\right) = \int_{-\infty}^{\infty} x^k f_R(x) dx$$

The  $k^{th}$  central moment of  $R_t$  is given by

$$\mu_{k} = \mathbb{E}\left(\left(R_{t} - m_{1}\right)^{k}\right) = \int_{-\infty}^{\infty} (x - m_{1})^{k} f_{R}(x) dx$$

with  $m_1 = \mathbb{E}(R_t)$  the mean of the return distribution.

In financial econometrics, we are mainly interested by the fourth first moments:

#### Mean Variance Skewness Kurtosis

	Mean	Std. dev	Skew	Kurt	Max	Min	Q(15)	$Q_2(15)$
Durations								
RY	0.99	1.14	2.50	12.36	16.68	0.01	1325	387
PDG	0.93	1.50	7.68	170.59	63.33	0.01	5015	185
Returns								
RY	0.000	0.815	0.005	7.560	5.960	-6.422	8196	15417
PDG	0.001	0.890	-0.009	6.735	6.404	-6.421	2355	3118

The sample period runs from April 1st to June 30, 2001. It consists of 51,660 observations for the RY stock and 27,956 observations for the PDG stock. Mean is the sample mean, Std. dev is the sample standard deviation, Skew is the sample skewness coefficient, Kurt is the sample kurtosis, Max is the sample maximum, Min is the sample minimum. Q(15) is the Ljung-Box test statistics with 15 lags,  $Q_2(15)$  is the Ljung-Box test statistics applied to squared returns using 15 lags. The associated 95% critical value is 24.996.

Table 2: Descriptive Statistics of Deseasonalised Data for Royal Bank (RY) and Placer Dome (PDG) Stocks

Source: Dionne G., Duchesne, P., and Pacurar, M. (2005), Intraday Value at Risk (IVaR) using tick-by-tick data with application to the Toronto Stock

Exchange, Journal of Empirical Finance, 16(5), 777-792.

### Definition (mean and variance of return)

The **mean** (expected value) of the return  $R_t$ 

$$\mu \equiv m_1 = \mathbb{E}(R_t) = \int_{-\infty}^{\infty} x f_R(x) dx$$

The **variance** of the return  $R_t$  is

$$\sigma^{2} \equiv \mathbb{V}(R_{t}) = \mathbb{E}\left((R_{t} - \mu)^{2}\right) = \int_{-\infty}^{\infty} (x - \mu)^{2} f_{R}(x) dx$$

### Definition (volatility)

The volatility corresponds to the standard deviation of the return distribution

Volatility 
$$\equiv \sigma = \sqrt{\mathbb{V}\left(R_t\right)}$$

### Definition (skewness coefficient)

The third central moment measures the skewness of the distribution

$$\mu_3 = \mathbb{E}\left(\left(R_t - \mu\right)^3\right) = \int\limits_{-\infty}^{\infty} \left(x - \mu\right)^3 f_R(x) dx$$

The (standardized) skewness coefficient is defined as

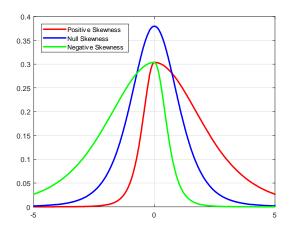
$$\mathbb{S}(R_t) = \mathbb{E}\left(\left(\frac{R_t - \mu}{\sigma}\right)^3\right) = \frac{\mu_3}{\sigma^3}$$

#### Skewness

- When  $S(R_t) < 0$ , large realizations of  $R_t$  are more often negative than positive (i.e., crashes are more likely than booms).
- The distribution is said to be left-skewed, left-tailed, or skewed to the left
- For **symmetric distributions** (e.g., normal, Student, etc.), the skewness coefficient is null

$$\mathbb{S}\left(R_{t}\right)=0$$

Figure: Illustration of the Skewness



### Definition (kurtosis coefficient)

The fourth central moment measures the tail heaviness of the distribution

$$\mu_4 = \mathbb{E}\left(\left(R_t - \mu\right)^4\right) = \int_{-\infty}^{\infty} \left(x - \mu\right)^4 f_R(x) \, \mathrm{d}x$$

The (standardized) kurtosis coefficient is defined as

$$\mathbb{K}\left(R_{t}\right) = \mathbb{E}\left(\left(\frac{R_{t} - \mu}{\sigma}\right)^{4}\right) = \frac{\mu_{4}}{\sigma^{4}}$$

### Kurtosis: properties

### Fact (normal distribution)

For a normal distribution  $R_t \sim \mathcal{N}\left(\mu, \sigma^2\right)$ , the kurtosis coefficient is equal to 3

$$\mathbb{K}(R_t)=3$$

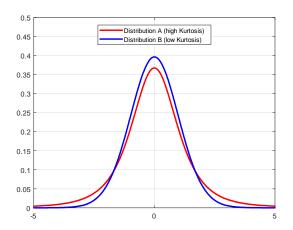
**Note:** the Gaussian distribution is not the only one to have a kurtosis equal to 3.

62 / 136

### Kurtosis: properties (cont'd)

- Large  $\mathbb{K}(R_t)$  implies that large realizations (positive or negative) are more likely to occur (compared to a normal distribution)
- If  $\mathbb{K}(R_t) > 3$ , the distribution of  $R_t$  is said to be **leptokurtic**.
- If  $\mathbb{K}(R_t) = 3$ , the distribution of  $R_t$  is said to be **mesokurtic**.
- If  $\mathbb{K}(R_t) < 3$ , the distribution of  $R_t$  is said to be platykurtic.
- The excess kurtosis coefficient is equal to  $\mathbb{K}(R_t) 3$ .

Figure: Illustration of the Kurtosis



### Summary

Each (central or non-central) moment gives an information on the pdf of returns

	Formula	Interpretation
Mean	$\mathbb{E}\left(R_{t}\right)=\mu$	Indicator of central tendency
Variance	$\mathbb{V}\left(R_{t}\right) = \mathbb{E}\left(\left(R_{t} - \mu\right)^{2}\right) = \sigma^{2}$	Indicator of dispersion around $\mu$
Skewness	$S(R_t) = \mathbb{E}\left(\left(R_t - \mu\right)^3\right) / \sigma^3$	Indicator of symmetry
Kurtosis	$\mathbb{K}\left(R_{t}\right) = \mathbb{E}\left(\left(R_{t} - \mu\right)^{4}\right) / \sigma^{4}$	Indicator of tail heaviness

### Example (normal distribution)

If we assume that the returns are **normally** distributed with  $R_t \sim \mathcal{N}\left(\mu, \sigma^2\right)$ , then

$$\mathbb{E}\left(R_{t}\right)=\mu$$

$$\mathbb{V}\left(R_{t}\right)=\sigma^{2}$$

$$S(R_t) = 0$$

$$\mathbb{K}(R_t)=3$$

The normal distribution is symmetric and by definition, mesokurtic.

### Example (Student distribution)

If we assume that the return  $R_t$  has a **Student distribution** with v degrees of freedom, i.e.  $R_t \sim t\left(v\right)$ , then we have

$$\mathbb{E}(R_t) = 0$$

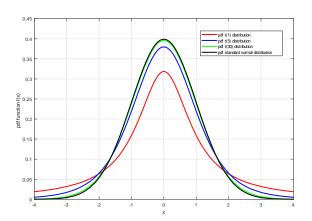
$$\mathbb{V}(R_t) = \frac{v}{v-2} \text{ if } v > 2$$

$$\mathbb{S}(R_t) = 0$$

$$\mathbb{K}(R_t) = 3 + \frac{6}{v-4} \text{ if } v > 4$$

The Student distribution is **symmetric** and **leptokurtic** (if v is "small"). Its excess of Kurtosis decreases with the number of degrees of freedom. When v tends to  $\infty$ ,  $t(v) \approx \mathcal{N}(0,1)$ , and the distribution is mesokurtic.

Figure: Density function for a Student distribution



#### Remarks

- The moments allow to characterize the shape of the returns distribution.
- However, the (theoretical) moments are unobservable and we need to estimate them.
- Denote by  $\{R_1, \dots, R_T\}$  a sample of i.i.d. variables that have the same distribution as  $R_t$
- Denote by  $\{r_1, \ldots, r_T\}$  the realization of this sequence => your **dataset**.

### Definition (estimator)

An **estimator** is any function  $T(R_1, R_2, ..., R_T)$  of a sample. Any **descriptive statistic** is an estimator.

### Example (sample mean)

Assume that  $R_1, R_2, ..., R_{\mathcal{T}}$  are i.i.d. random variables. The sample mean (or average)

$$\widehat{\mu}_T \equiv \overline{R}_T = \frac{1}{T} \sum_{t=1}^T R_t$$

is an estimator of the (theoretical) mean  $\mu=\mathbb{E}\left(R_{t}\right)$ .

### Example (sample variance)

Assume that  $R_1, R_2, ..., R_T$  are i.i.d. random variables. The **sample variance** 

$$\widehat{\sigma}_{T}^{2} = rac{1}{T-1}\sum_{t=1}^{T}\left(R_{t}-\overline{R}_{T}
ight)^{2}$$

is an estimator of the (theoretical) variance  $\sigma^2=\mathbb{V}\left(R_t
ight)$ .

**Note:** the denominator is equal to T-1 as to define a sample variance corrected for the small sample bias (cf. chapter 2).

### Example (sample skewness)

Assume that  $R_1, R_2, ..., R_T$  are i.i.d. random variables. The sample skewness coefficient

$$\widehat{S}_{T} = \frac{1}{T} \sum_{t=1}^{T} \left( \frac{R_{t} - \overline{R}_{T}}{\widehat{\sigma}} \right)^{3} = \frac{T^{-1} \sum_{t=1}^{T} \left( R_{t} - \overline{R}_{T} \right)^{3}}{\left( \left( T - 1 \right)^{-1} \sum_{t=1}^{T} \left( R_{t} - \overline{R}_{T} \right)^{2} \right)^{3/2}}$$

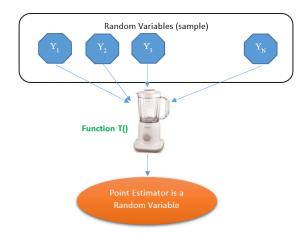
is an estimator of the (theoretical) skweness  $\mathbb{S}(R_t)$ .

#### Example (sample kurtosis)

Assume that  $R_1, R_2, ..., R_T$  are i.i.d. random variables. The **sample kurtosis** coefficient

$$\widehat{K}_{T} = \frac{1}{T} \sum_{t=1}^{T} \left( \frac{R_{t} - \overline{R}_{T}}{\widehat{\sigma}} \right)^{4} = \frac{T^{-1} \sum_{t=1}^{T} \left( R_{t} - \overline{R}_{T} \right)^{4}}{\left( (T - 1)^{-1} \sum_{t=1}^{T} \left( R_{t} - \overline{R}_{T} \right)^{2} \right)^{2}}$$

is an estimator of the (theoretical) kurtosis  $\mathbb{K}(R_t)$ .



#### Fact

An estimator  $\widehat{\theta}$  is a random variable.

**Consequence:**  $\widehat{\theta}$  has a (marginal or conditional) **probability distribution**. This **sampling distribution** is characterized by a pdf  $f_{\widehat{\theta}}(u)$ 

## Definition (Sampling Distribution)

The probability distribution of an estimator (or a statistic) is called the **sampling distribution**.

**Consequence:** The sampling distribution of  $\widehat{\theta}$  is characterized by **moments** such that the expectation  $\mathbb{E}\left(\widehat{\theta}\right)$ , the variance  $\mathbb{V}\left(\widehat{\theta}\right)$ , etc.

#### Definition (estimate)

A (point) **estimate** is the realized value of an estimator (i.e. a number) that is obtained when a sample is actually taken. For an estimator  $\widehat{\theta}$  it can be denoted by  $\widehat{\theta}\left(y\right)$ .

### Example (Point estimate)

For instance a realization of the estimator  $\overline{R}_T$ , denoted  $\overline{r}_T$ , is defined as

$$\overline{r}_T = \frac{1}{T} \sum_{t=1}^{T} r_t$$

- If T = 3 and  $\{r_1, r_2, r_3\} = \{3, -1, 2\}$  then  $\overline{r}_T = 1.333$ .
- If T=3 and  $\{r_1,r_2,r_3\}=\{4,-8,1\}$  then  $\overline{r}_T=-1.$
- etc..

#### Summary

Population moments (real nbs.)		Sample moments (random variables)	
Mean	$\mathbb{E}\left(R_{t} ight)=nb$	Sample mean	$\widehat{\mu}_T = \frac{1}{T} \sum_{t=1}^T R_t$
Variance	$\mathbb{V}\left(R_{t} ight)=nb$	Sample variance	$\widehat{\sigma}_{T}^{2} = \frac{1}{T-1} \sum_{t=1}^{T} \left( R_{t} - \overline{R}_{T} \right)^{2}$
Skewness	$\mathbb{S}\left(R_{t} ight)=nb$	Sample Skewness	$\widehat{S}_T = rac{1}{T} \sum_{t=1}^T \left( rac{R_t - \overline{R}_T}{\widehat{\sigma}}  ight)^3$
Kurtosis	$\mathbb{K}\left(R_{t} ight)=nb$	Sample Kurtosis	$\widehat{\mathcal{K}}_{\mathcal{T}} = rac{1}{\mathcal{T}} \sum_{t=1}^{\mathcal{T}} \left( rac{R_t - \overline{R}_{\mathcal{T}}}{\widehat{\sigma}}  ight)^4$

Question: What constitutes a good estimator?

The idea is to study the properties of the sampling distribution and especially its moments such as

- ullet  $\mathbb{E}\left(\widehat{ heta}
  ight)$  for the bias,
- ullet  $\mathbb{V}\left(\widehat{ heta}
  ight)$  for the precision,
- $S\left(\widehat{\theta}\right)$  for the symmetry,
- etc..

Estimators are compared on the basis of a variety of attributes.

- Finite sample properties (or finite sample distribution) of estimators are those attributes that can be compared regardless of the sample size.
- ullet However, the finite sample distribution is known only in the case of specific distributional assumption on  $R_t$  (typically, normality).
- When the normality assumption is no longer valid (and the finite sample distribution is unknown), estimators are evaluated on the basis on their large sample, or asymptotic properties

### Theorem (finite sample distributions)

If we assume that the returns  $R_1$ ,  $R_2$ , ...,  $R_T$  are i.i.d. with  $R_t \sim \mathcal{N}\left(\mu, \sigma^2\right)$ , then  $\widehat{\mu}_T$  and  $(T-1)\widehat{\sigma}_T^2/\sigma^2$  have a finite sample distribution

$$\widehat{\mu}_{T} \sim \mathcal{N}\left(\mu, \frac{\sigma^{2}}{T}\right) \quad \forall T \in \mathbb{N}$$

$$\frac{(T-1)}{\sigma^{2}}\widehat{\sigma}_{T}^{2} \sim \chi^{2}(T-1) \quad \forall T \geq 2$$

### Example (finite sample distribution)

If we assume that the returns  $R_1$ ,  $R_2$ , ...,  $R_{\mathcal{T}}$  are  $\mathcal{N}.i.d.(\mu,\sigma^2)$  , if  $\mathcal{T}=10$ , then

$$\widehat{\mu}_{10} \sim \mathcal{N}\left(\mu, \frac{\sigma^2}{10}\right)$$
  $9 \times \frac{\widehat{\sigma}_T^2}{\sigma^2} \sim \chi^2(9)$ 

In most of cases, it is impossible to derive the **exact/finite sample distribution** for the estimator (or a transformed variable).

#### Two reasons:

• In some cases, the exact distribution of  $R_1, R_2, \ldots, R_T$  is known, but the function T(.) is too complicated to derive the distribution of  $\widehat{\theta}$ :

$$\widehat{\theta} = T(R_1, R_2, \dots, R_T) \sim ??? \quad \forall T \in \mathbb{N}$$

**1** In most of cases, the distribution of the returns  $R_1, R_2, \ldots, R_T$  is **unknown**...

$$\widehat{\theta} = T(R_1, R_2, \dots, R_T) \sim ??? \quad \forall T \in \mathbb{N}$$

**Question:** what is the behavior of the estimator  $\widehat{\theta}_T$  when the sample size T tends to infinity?

## Definition (Asymptotic theory)

**Asymptotic** or **large sample theory** consists in the study of the distribution of the estimator when the sample size is sufficiently large.

The asymptotic theory is fundamentally based on the notion of convergence.

We distinguish two types of convergence

**Onvergence in probability** (or in mean squared, or almost sure):

$$\widehat{ heta}_{\mathcal{T}}$$
 converges to a real number

This convergence types are used to derive the **consistency** property of the estimators.

Convergence in distribution:

$$\sqrt{T}\left(\widehat{\theta}_{T}-\theta_{0}
ight)$$
 converges to a given distribution (say, normal)

This convergence is used to derive the asymptotic distribution of the estimators and to make inference (tests) about the true value of the parameters.

### Definition (Convergence in probability)

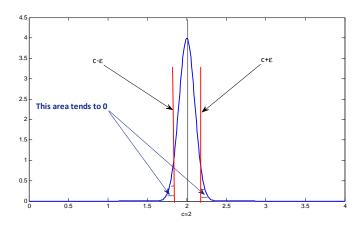
Let  $X_T$  be a sequence random variable indexed by the sample size.  $X_T$  converges in probability to a constant c, if, for any  $\varepsilon > 0$ ,

$$\lim_{T\to\infty} \Pr\left(|X_T - c| > \varepsilon\right) = 0$$

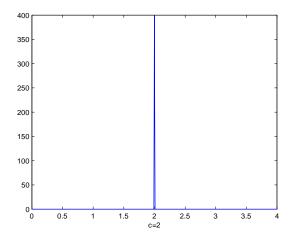
It is written

$$X_T \stackrel{p}{\rightarrow} c$$
 or plim  $X_T = c$ 

$$X_T \stackrel{p}{\to} c$$
 if  $\lim_{t \to \infty} \Pr(|X_T - c| > \varepsilon) = 0$ 



$$X_t \stackrel{p}{ o} c \quad ext{if} \quad \lim_{T o \infty} \Pr\left(|X_T - c| > \epsilon\right) = 0 \quad ext{for a very small } \epsilon...$$



#### Theorem (consistency)

If the returns  $R_1$ ,  $R_2$ , ...,  $R_T$  are i.i.d. with finite moments,  $\mathbb{E}\left(R_t\right) = \mu$ ,  $\mathbb{V}\left(R_t\right) = \sigma^2$ ,  $\mathbb{S}\left(R_t\right)$  and  $\mathbb{K}\left(R_t\right)$ , then the sample first four moments are **weakly consistent** 

$$\widehat{\mu}_{T} \stackrel{p}{\to} \mathbb{E}(R_{t}) = \mu$$

$$\widehat{\sigma}_{T}^{2} \stackrel{p}{\to} \mathbb{V}(R_{t}) = \sigma^{2}$$

$$\widehat{S}_{T} \stackrel{p}{\to} \mathbb{S}(R_{t})$$

$$\widehat{K}_{T} \stackrel{p}{\to} \mathbb{K}(R_{t})$$

#### Interpretation

- The distribution of the sample moments is highly concentrated around the true value (unknown) of the population moments of the returns when the sample size T tends to infinity.
- The realization of the sample moments (mean, variance, skewness, etc.) are then "close" to the value of the population moment when the sample size T tends to infinity.

#### Definition (Convergence in distribution)

Let  $X_T$  be a sequence random variable indexed by the sample size with a cdf  $F_T$  (.).  $X_T$  converges in distribution to a random variable X with cdf F (.) if

$$\lim_{T\to\infty} F_T(x) = F(x) \quad \forall x$$

It is written:

$$X_t \stackrel{d}{\longrightarrow} X$$

## Theorem (asymptotic distributions for sample mean and variance)

If the returns  $R_1, R_2, ..., R_T$  are i.i.d. with finite mean  $\mathbb{E}(R_t) = \mu$  and finite variance  $\mathbb{V}(R_t) = \sigma^2$ , then we can derive the following asymptotic distributions when T tends to infinity

$$\sqrt{T}\left(\widehat{\mu}_{T}-\mu\right) \stackrel{d}{\to} \mathcal{N}\left(0,\sigma^{2}\right)$$

$$\sqrt{T}\left(\widehat{\sigma}_{T}^{2}-\sigma^{2}\right)\overset{d}{\rightarrow}\mathcal{N}\left(0,2\sigma^{4}\right)$$

Remark: The asymptotic results

$$\sqrt{T} \left( \widehat{\mu}_T - \mu \right) \stackrel{d}{\to} \mathcal{N} \left( 0, \sigma^2 \right)$$

$$\sqrt{T} \left( \widehat{\sigma}_T^2 - \sigma^2 \right) \stackrel{d}{\to} \mathcal{N} \left( 0, 2\sigma^4 \right)$$

can be interpreted as

$$\widehat{\boldsymbol{\mu}}_{T} \overset{\text{asy}}{\approx} \mathcal{N}\left(\boldsymbol{\mu}, \frac{\sigma^{2}}{T}\right) \qquad \widehat{\boldsymbol{\sigma}}_{T} \overset{\text{asy}}{\approx} \mathcal{N}\left(\sigma^{2}, \frac{2\sigma^{4}}{T}\right)$$

where the symbol  $\stackrel{\textit{asy}}{pprox}$  means "asymptotically distributed as".

**Note:** These asymptotic results for the sample moments can be used to perform **statistical tests** about the (population) moments.

## Example (confidence interval)

Consider a sample of T=500 i.i.d. returns  $R_1,...,R_T$ , and a realization (dataset)  $\{r_1,...,r_T\}$  with

$$\widehat{\mu}_T = \overline{r}_T = -0.057$$
  $\widehat{\sigma}_T^2 = 0.0009$ 

**Question:** compute a 95% confidence interval on the true value of the expected return  $\mathbb{E}(R_t)$ .

Answer

We known that  $\widehat{\mu}_{T}\stackrel{\mathit{asy}}{\approx}\mathcal{N}\left(\mu,\sigma^{2}/T\right)$  , so we have

$$IC_{95\%} = \left[\mu \pm \frac{\sigma}{\sqrt{T}}\Phi^{-1}\left(0.975\right)\right]$$

Since  $\mu$  and  $\sigma^2$  are unobservable, we can use a consistent estimator (plug-in approach) and we get

$$IC_{95\%} = \left[ \widehat{\mu} \pm \frac{\widehat{\sigma}}{\sqrt{T}} \Phi^{-1} (0.975) \right]$$

$$= \left[ -0.057 \pm \sqrt{\frac{0.0009}{500}} \Phi^{-1} (0.975) \right]$$

$$= \left[ -0.0596; -0.0544 \right]$$

## Theorem (asymptotic distributions for sample skewness and kurtosis)

If we assume that the returns  $R_1$ ,  $R_2$ , ...,  $R_T$  are  $\mathcal{N}.i.d.(\mu, \sigma^2)$ , then we have

$$\sqrt{T}\left(\widehat{S}_{\mathcal{T}}-0\right)\overset{d}{\rightarrow}\mathcal{N}\left(0,6\right)$$

$$\sqrt{T}\left(\widehat{K}_{\mathcal{T}}-3\right)\overset{d}{\rightarrow}\mathcal{N}\left(0,24\right)$$

**Note:** These asymptotic results for the sample moments can be used to perform **statistical tests** about the returns distribution (Jarque-Bera test for instance).

#### **Key Concepts**

- Returns distribution
- 2 Central and non-central moments
- Mean, variance, skewness, and kurtosis
- Leptokurtic or mesokurtic returns distribution
- Symmetric or asymmetric returns distribution
- Sample, estimator, estimate, and (descriptive) statistics
- Independent and identically distributed (i.i.d.) random variables
- Sampling distribution of an estimator
- Finite sample vs asymptotic distribution

# Section 5

Stylized Facts

#### **Objectives**

- Present the main stylized facts of financial series
- Define the stationarity property of asset returns
- Show the absence of autocorrelations of asset returns
- Show the heavy tails and the asymmetry of the distribution of returns
- Show the volatility clustering
- Opening the aggregational Gaussianity
- Introduce the long range dependence
- Opening the leverage effect

- The statistical properties of financial data revealed a wealth of stylized facts
- These stylized facts are common in a wide range of financial time series.



Cont, R. (2001), Empirical Properties of Asset Returns: Stylized Facts and Statistical Issues, *Quantitative Finance*, 1, 223–236

## Example (Intel Corp)

In order to illustrate some of these stylized facts, we consider a sample of 1, 259 daily prices and (log-) returns for the S&P500 index (ticker: SPY) from August 19, 2013 to August 17, 2018 (5 years). The data are available in Data\_SP500.xlsx.

Fan and Yao (2015) identify 8 main stylized facts

- Stationarity
- Absence of autocorrelations
- Heavy tails
- Asymmetry
- Volatility clustering
- Aggregational Gaussianity
- Long range dependence
- Leverage effect



Fan, J. and Yao, Q. The Elements of Financial Econometrics, Science Press, Beijing, 2015

99 / 136

Fan and Yao (2015) identify 8 main stylized facts

- Stationarity
- Absence of autocorrelations
- Heavy tails
- Asymmetry
- Volatility clustering
- Aggregational Gaussianity
- Long range dependence
- Leverage effect

Fan, J. and Yao, Q. The Elements of Financial Econometrics, Science Press, Beijing, 2015

#### Stylized Fact 1: Stationarity

#### Fact (stationarity)

In general, the prices are non-stationary whereas the returns are stationary.

- The prices of an asset recorded over times are often not stationary due to, for example, the steady expansion of economy, the increase of productivity, and financial crisis.
- However their returns, typically fluctuates around a constant level, suggesting a constant mean over time.
- Most return sequences can be modeled as a stochastic processes with at least time-invariant two first moments: weak stationarity (cf. chapter 5)

#### Definition (weak or second-order stationarity)

A time-series process  $(x_t, t \in \mathbb{Z})$  is **weakly stationary** (second-order stationarity) if and only if:

- $\forall t \in \mathbb{Z}$ ,  $\mathbb{E}\left(x_t^2\right) < \infty$
- $\forall t \in \mathbb{Z}$ ,  $\mathbb{E}(x_t) = \mu$  does not depend on t
- ullet  $\forall (t,h) \in \mathbb{Z}^2$ ,  $\mathbb{C}$  ov  $(x_t,x_{t+h}) = \mathbb{E}\left(\left(x_{t+h}-m
  ight)(x_t-m)
  ight) = \gamma_h$  does not depend on t

Figure: Daily closing prices for the S&P500 index are non stationary

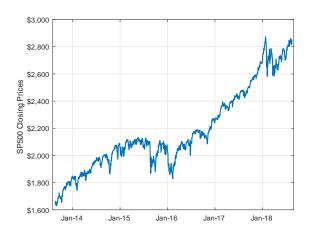
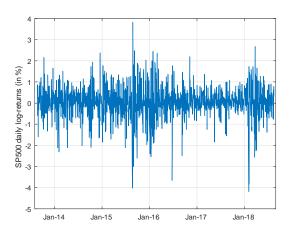


Figure: Daily returns for the S&P500 index are stationary



Fan and Yao (2015) identify 8 main stylized facts

- Stationarity
- Absence of autocorrelations
- Heavy tails
- Asymmetry
- Volatility clustering
- Aggregational Gaussianity
- Long range dependence
- Leverage effect



Fan, J. and Yao, Q. The Elements of Financial Econometrics, Science Press, Beijing, 2015

#### **Stylized Fact 2: Absence of autocorrelations**

#### Fact (absence of autocorrelations )

The autocorrelations of asset returns  $R_t$  are often insignificant, except for very small intraday time scales ( $\approx 20$  minutes) for which microstructure effects come into play.

**Note:** The fact that returns hardly show any serial correlation, does not mean that they are independent.

#### Definition (autocorrelation)

The **autocorrelation**, denoted  $\rho(k)$ , of a weak stationary process  $R_t$  is the correlation between values of the process at different times, defined as:

$$\rho_{k} = \operatorname{Corr}\left(R_{t}, R_{t-k}\right) = \frac{\mathbb{E}\left(\left(R_{t} - \mu\right)\left(R_{t-k} - \mu\right)\right)}{\mathbb{V}\left(R_{t}\right)} = \frac{\gamma_{k}}{\sigma^{2}}$$

with  $\mu = \mathbb{E}(R_t)$ ,  $\sigma^2 = \mathbb{V}(R_t)$ ,  $\forall t$  and  $\gamma_k$  the autocovariance of order k.

#### Definition (sample autocorrelation)

The **sample autocorrelation**, denoted  $\widehat{\rho}\left(k\right)$ , of a weak stationary process  $R_{t}$ , is an estimator of  $\rho\left(k\right)$  defined as

$$\widehat{\rho}_{k} = corr\left(R_{t}, R_{t-k}\right) = \frac{1}{\left(T - k\right)\widehat{\sigma}^{2}} \sum_{t=k+1}^{I} \left(R_{t} - \widehat{\mu}\right) \left(R_{t-k} - \widehat{\mu}\right)$$

where  $\widehat{\sigma}^2$  and  $\widehat{\mu}$  are consistent estimators of  $\mu = \mathbb{E}\left(R_t\right)$  and  $\sigma^2 = \mathbb{V}\left(R_t\right)$ ,  $\forall t$ .

The AutoCorrelation function (ACF) (or correlogram) represents the sample autocorrelation for different lags from k=1 to a maximum lag order, say k=15.

Figure: ACF for the S&P500 daily returns (Aug 2013 - Aug 2018). Source: Eviews 3

Correlogram of SP500_RETURNS								
Date: 09/09/18 Time: 11:25 Sample: 8/19/2013 8/17/2018 Included observations: 1259								
Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob			
4 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0 0		8 -0.037 9 -0.054 10 -0.019	-0.054 0.022 -0.045 -0.027 -0.014 0.011 -0.039 -0.057 -0.028 -0.003 -0.004 -0.022 -0.027	7.1488 8.1122 8.1869 8.3965 10.165 13.887 14.324 14.330 14.617 15.102	0.225 0.299 0.254 0.126 0.159 0.215 0.280 0.332			

#### Testing the nullity of autocorrelations

- It is possible to test the nullity of the autocorrelations through a Q-test (Box-Pierce test or Ljung-Box test).
- For more details on statistical tests: cf. Chapter 3.
- For a given lag-order K, the null hypothesis of the test is

$$H_0: \rho_1 = \ldots = \rho_K = 0$$

• The alternative hypothesis is

$$\mathsf{H}_1:\exists j\in\{1,\ldots,K\}$$
 /  $\widehat{
ho}_j
eq 0$ 

#### Definition (Box-Pierce test)

The **Box-Pierce** test statistic associated to the null  $H_0: \rho_1 = \ldots = \rho_K = 0$ , is defined as

$$Q_{BP} = T \sum_{k=1}^{K} \widehat{\rho}_{k}^{2} \xrightarrow[H_{0}]{d} \mathcal{X}^{2}\left(K\right)$$

#### Decision rule (cf. Chapter 3):

- If the realization of the test statistic  $Q_{BP}$  is larger than the quantile of a chi-squared distribution with K degrees of freedom at the probability  $1-\alpha\%$  (say, 95%), for a risk level  $\alpha\%$  (say, 5%) we reject the null hypothesis  $H_0$ .
- If the p-value associated to the test statistic  $Q_{BP}$  is smaller the risk level  $\alpha\%$  (say, 5%), we reject the null hypothesis  $H_0$ .
- If the null hypothesis  $H_0$  is rejected, there is at least one autocorrelation (between lag 1 and K) which is non-null, i.e. the time series  $R_t$  is autocorrelated.

**イロトイ団ト イミト イミト ミー り**900

Figure: Q-statistics for the S&P500 daily returns (Aug 2013 - Aug 2018). Source: Eviews 3

# Correlogram of SP500\_RETURNS Date: 09/09/18 Time: 11.25 Sample: 8/19/2013 8/17/2018 Included observations: 1259

A	Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
	utocorrelation  (i  ii  ii  ii  ii  ii  ii  ii  ii  i	Partial Correlation	13	0.019 0.054 0.024 0.028 0.008 0.013 0.037 0.054 0.019 0.002 0.001	-0.019 -0.054 0.022 -0.045 -0.027 -0.014 0.011 -0.039 -0.057 -0.028 -0.003 -0.004 -0.002	0.4455 4.1067 4.8464 7.1488 8.1122 8.1869 8.3965 10.165 13.887 14.324 14.330 14.330 14.617	0.504 0.128 0.183 0.128 0.150 0.225 0.299 0.254 0.126 0.159 0.215 0.280 0.332
	1				-0.027 -0.079	15.102 21.926	0.371 0.110

Fan and Yao (2015) identify 8 main stylized facts

- Stationarity
- Absence of autocorrelations
- Heavy tails
- Asymmetry
- Volatility clustering
- Aggregational Gaussianity
- Long range dependence
- Leverage effect



#### Stylized Fact 3: Heavy tails

# Fact (heavy tails)

The probability distribution of return often exhibits **heavier tails** than those of a normal distribution.

Figure: Kernel density estimate of the daily returns distribution (pdf) for the S&P500 index

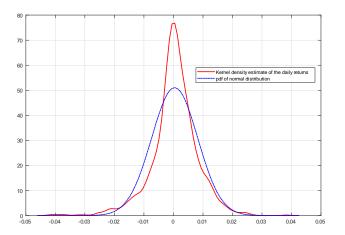
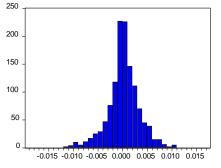


Figure: Descriptive statistics for the daily returns of the S&P500 index (Eviews 3)



Series: SP500_ Sample 8/20/20				
Observations 1259				
Mean	0.000189			
Median	0.000231			
Maximum	0.016630			
Minimum	-0.018172			
Std. Dev.	0.003390			
Skew ness	-0.566934			
Kurtosis	6.405986			
_				
Jarque-Bera	675.9989			
Probability	0.000000			

Fan and Yao (2015) identify 8 main stylized facts

- Stationarity
- Absence of autocorrelations
- Heavy tails
- Asymmetry
- Volatility clustering
- Aggregational Gaussianity
- Long range dependence
- Leverage effect

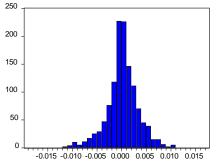


#### Stylized Fact 4: Asymmetry

#### Fact (asymmetry)

The distribution of return is **asymmetric** and often **negatively skewed**, reflecting the fact that the downturns of financial markets are often much steeper than the recoveries. Investors tend to react more strongly to negative news than to positive news

Figure: Descriptive statistics for the daily returns of the S&P500 index (Eviews 3)



Series: SP500_RETURNS Sample 8/20/2013 6/15/2018				
Observations 1	259			
Mean	0.000189			
Median	0.000231			
Maximum	0.016630			
Minimum	-0.018172			
Std. Dev.	0.003390			
Skew ness	-0.566934			
Kurtosis	6.405986			
Jarque-Bera	675.9989			
Probability	0.000000			

Fan and Yao (2015) identify 8 main stylized facts

- Stationarity
- Absence of autocorrelations
- Heavy tails
- Asymmetry
- Volatility clustering
- Aggregational Gaussianity
- Long range dependence
- Leverage effect



#### Stylized Fact 5: Volatility clustering

#### Fact (volatility clustering)

The volatility clustering means that large price changes (i.e. returns with large absolute values or large squares) occur in clusters. Indeed, large price changes tend to be followed by large price changes, and periods of tranquility alternate with periods of high volatility.

Note: the volatility clustering is the consequence of the autocorrelation of the squared returns (cf. Stylized fact 7)

Figure: Daily returns for the S&P500 index

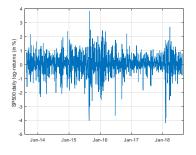
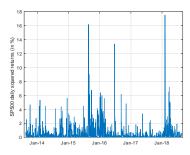


Figure: Daily squared-returns for the S&P500 index



Fan and Yao (2015) identify 8 main stylized facts

- Stationarity
- Absence of autocorrelations
- Heavy tails
- Asymmetry
- Volatility clustering
- Aggregational Gaussianity
- Long range dependence
- Leverage effect



#### Stylized Fact 6: Aggregational Gaussianity

# Fact (aggregational Gaussianity)

A return over k days is simply the aggregation of k daily returns. When the time horizon k increases, the central limit law sets in and the distribution of the returns over a long time-horizon (such as a month) tends toward a **normal distribution**.

#### Definition (Jarque-Bera test)

The Jarque-Bera test is

$$H_0: \mathbb{S}(R_t) = 0$$
 and  $\mathbb{K}(R_t) = 3$ 

$$\mathsf{H}_1:\mathbb{S}\left(R_t
ight)
eq 0 \ \ \mathsf{and/or} \ \ \mathbb{K}\left(R_t
ight)
eq 3$$

The Jarque-Bera test statistic is defined as

$$\mathsf{JB} = \frac{T}{6} \widehat{\mathsf{S}}_{T}^{2} + \frac{T}{24} \left( \widehat{\mathsf{K}}_{T} - 3 \right)^{2} \xrightarrow[\mathsf{H}_{0}]{d} \mathcal{X}^{2} \left( 2 \right)$$

where  $\widehat{S}_{\mathcal{T}}$  and  $\widehat{K}_{\mathcal{T}}$  denote the sample Skewness and Kurtosis coefficients, respectively.

#### Jarque-Bera test

#### Interpretation:

- The Jarque-Bera is not a normality test. However, if the null hypothesis is rejected the distribution of  $R_t$  cannot be normal.
- If one fails to reject the null hypothesis, it only implies that the distribution is symmetric and mesokurtic (the normal distribution is not the only one in this case).

#### Decision rule (cf. Chapter 3):

- If the JB test statistic is larger than the quantile of a chi-squared distribution with 2 degrees of freedom at the probability  $1-\alpha\%$  (say, 95%), for a risk level  $\alpha\%$  (say, 5%) we reject the null hypothesis H<sub>0</sub>.
- If the p-value associated to the JB test statistic is smaller the risk level  $\alpha\%$  (say, 5%), we reject the null hypothesis H<sub>0</sub>.

#### Example (Jarque-Bera test statistic)

For the S&P500 index, the sample Skewness and Kurtosis coefficient are equal to -0.5669 and 6.4059, respectively for a sample size T equal to 1259. **Question:** compute the realization of the Jarque-Bera test statistic and conclude for a 5% significance level.

Answer: The realization of the Jarque-Bera test statistic equal to

$$\mathsf{JB} = \frac{T}{6}\widehat{S}_{T}^{2} + \frac{T}{24}\left(\widehat{K}_{T} - 3\right)^{2} = \frac{1259}{6}\left(-0.5669\right)^{2} + \frac{1259}{24}\left(6.4059 - 3\right)^{2} = 675.9989$$

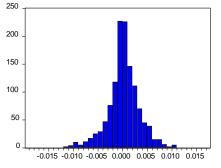
For a risk level  $\alpha=5\%$ , the critical value is equal to

critical value = 
$$F_2^{-1}(\alpha) = F_2^{-1}(0.95) = 5.9915$$

where F(.) is the cdf of a chi-squared distribution with 2 degrees of freedom.

**Conclusion:** For a risk level  $\alpha=5\%$ , test statistic is larger than the critical value, so we reject the null hypothesis (symmetric and mesokurtic distribution). The distribution of the daily returns **cannot** be normal.

Figure: Descriptive statistics for the daily returns of the S&P500 index (Eviews 3)



Series: SP500_RETURNS Sample 8/20/2013 6/15/2018 Observations 1259				
Mean	0.000189			
Median	0.000231			
Maximum	0.016630			
Minimum	-0.018172			
Std. Dev.	0.003390			
Skew ness	-0.566934			
Kurtosis	6.405986			
Jarque-Bera	675.9989			
Probability	0.000000			

Fan and Yao (2015) identify 8 main stylized facts

- Stationarity
- Absence of autocorrelations
- Heavy tails
- Asymmetry
- Volatility clustering
- Aggregational Gaussianity
- Long range dependence
- Leverage effect



#### **Stylized Fact 7: Long range dependence**

#### Fact (long range dependence)

Both daily squared and absolute returns often exhibit significant autocorrelations. Those autocorrelations are persistent, indicating possible long-memory properties.

**Note:** Those autocorrelations become weaker and less persistent when the sampling interval is increased from a day, to a week to a month.

Figure: ACF for the S&P500 returns

Date: 09/09/18 Time: 15:19 Sample: 8/19/2013 8/17/2018 Included observations: 1259

Autocorrelation	Partial Correlation	AC	PAC	Q-Stat	Prob
() () ()	(1 (1) (1) (1)	2 -0.05 3 0.02 4 -0.04 5 -0.02 6 -0.00	4 -0.054 4 0.022 3 -0.045 8 -0.027 8 -0.014	8.1869	0.183 0.128 0.150 0.225
0	1) () () () () () ()	8 -0.03	4 -0.057 9 -0.028 2 -0.003 1 -0.004 5 -0.022 0 -0.027	10.165 13.887 14.324 14.330 14.330 14.617 15.102	0.254 0.126 0.159 0.215 0.280 0.332 0.371

# Figure: ACF for the S&P500 squared returns

Date: 09/09/18 Time: 11:25 Sample: 8/19/2013 8/17/2018 Included observations: 1259

Autocorrelation	Partial Correlation		AC	PAC	Q-Stat	Prob
-		1	0.302	0.302	114.94	0.000
	<u> </u>	2	0.251	0.176	194.81	0.000
	<u> </u>	3	0.288	0.196	299.82	0.000
		4	0.251	0.116	379.50	0.00
•	1 0	5	0.115	-0.052	396.30	0.00
•	1 1	6	0.135	0.018	419.45	0.00
•	(	7	0.095	-0.020	430.83	0.00
•	1 1	8	0.108	0.044	445.63	0.00
•	1 1	9	0.102	0.043	458.93	0.00
•		10	0.123	0.060	478.14	0.00
•	1 1	11	0.084	0.003	487.21	0.00
4		12	0.050	-0.041	490.35	0.00
•	1 (	13	0.048	-0.020	493.24	0.00
•	1 (	14	0.040	-0.012	495.29	0.00
•	1 0	15	0.031	0.010	496.51	0.00

#### Remarks

- The autocorrelation of the squared returns is called ARCH effect (cf. chapter 6).
- The ARCH effect depends on the sampling frequency.
- It is most important with daily returns, and less important with low frequency returns (monthly, quarterly, etc.).

Fan and Yao (2015) identify 8 main stylized facts

- Stationarity
- Absence of autocorrelations
- Heavy tails
- Asymmetry
- Volatility clustering
- Aggregational Gaussianity
- Long range dependence
- Leverage effect



#### Stylized Fact 8: Leverage effect

#### Fact (leverage effect)

Asset returns are negatively correlated with the changes of their volatilities: this negative correlation is called leverage effect.

- As asset prices decline, companies become more leveraged (debt to equity ratios increase) and riskier, and hence their stock prices become more volatile.
- On the other hand, when stock prices become more volatile, investors demand high returns and hence stock prices go down.
- Volatilities caused by price decline are typically larger than the appreciations due to declined volatilities.

#### **Key Concepts**

- Stationarity
- Absence of autocorrelation
- ARCH effect
- Heavy tails and leptokurtic distribution
- 4 Asymmetric distribution
- Volatility clustering
- Aggregational Gaussianity
- Long range dependence
- Leverage effect

# End of the Chapter 1

Christophe Hurlin