INTRODUCTION TO POMP BY EXAMPLE

AARON A. KING

Contents

1.	A first example: the two-dimensional Ornstein-Uhlenbeck process.	1
1.1	Building the pomp object	1
1.2	The low-level interface	4
2.	Particle filter.	6
3.	The maximum likelihood by iterated filtering (MIF) algorithm	7

1. A first example: the two-dimensional Ornstein-Uhlenbeck process.

To keep things simple, we will study a discrete-time process. The pomp package is designed with continuous-time processes in mind, but the associated computational effort is typically greater, and the additional complexities are best postponed until the structure and usage of the package is understood. The unobserved Ornstein-Uhlenbeck (OU) process $X_t \in \mathbb{R}^2$ satisfies

$$X_t = A X_{t-1} + \xi_t.$$

The observation process is

$$Y_t = B X_t + \varepsilon_t$$
.

In these equations, A and and B are 2×2 constant matrices; ξ_t and ε_t are mutually-independent families of i.i.d. bivariate normal random variables. We let $\sigma\sigma^T$ be the variance-covariance matrix of ξ_t , where σ is lower-triangular; likewise, we let $\tau\tau^T$ be that of ε_t .

In order to specify this partially-observed Markov process, we must implement both the process model (i.e., the unobserved process) and the measurement model (the observation process). In particular, we will need to be able to both simulate from and compute the p.d.f. of each of these models. In pomp, one constructs an object of class pomp to hold functions that do all of these things, together with data and other information. The documentation (?pomp) spells out the usage of the pomp constructor, including detailed specifications for all its arguments and a worked example.

- 1.1. **Building the pomp object.** We build a pomp object by specifying the four basic elements mentioned above. First, we write a function that implements the process model simulator. In this function, we assume that:
 - (1) **xstart** will be a matrix, each column of which is a vector of initial values of the state process;
 - (2) params will be a matrix, the columns of which are parameter vectors;
 - (3) times will be a vector of times at which realizations of the state process are required. In particular, times[1] is the initial time (corresponding to xstart).

```
> ou2.rprocess <- function(xstart, times, params,
      ...) {
      nsims <- ncol(xstart)</pre>
      ntimes <- length(times)</pre>
      alpha <- array(params[c("alpha.1", "alpha.2",</pre>
           +
      sigma <- array(params[c("sigma.1", "sigma.2",</pre>
          "sigma.2", "sigma.3"), ], dim = c(2, 2, 3)
+
          nsims))
      sigma[1, 2, ] <- 0
      x \leftarrow array(0, dim = c(2, nsims, ntimes))
      rownames(x) <- rownames(xstart)</pre>
      x[, , 1] <- xstart
+
      for (k in 1:nsims) {
          for (j in 2:ntimes) {
              x[, k, j] \leftarrow alpha[, k] %*% x[,
                   k, j - 1] + sigma[, , k] %*% rnorm(2)
          }
      }
      \boldsymbol{X}
+ }
```

Notice that this function returns a rank-3 array (x), which has the realized values of the state process at the requested times. Notice too that x has rownames. When this function is called, in the course of any algorithm that uses it, some basic error checks will be performed.

Next, we write a function that computes the likelihoods of a set of process model state transitions. Again, pay special attention to the structure of the input arguments and return value.

```
> ou2.dprocess <- function(x, times, params, log = FALSE,
      ...) {
+
      nsims \leftarrow ncol(x)
      ntimes <- length(times)</pre>
      alpha <- array(params[c("alpha.1", "alpha.2",</pre>
          nsims))
      sigma <- array(params[c("sigma.1", "sigma.2",</pre>
          "sigma.2", "sigma.3"), ], \dim = c(2, 2, 2, 1)
          nsims))
      sigma[1, 2, ] <- 0
      d \leftarrow array(0, dim = c(nsims, ntimes - 1))
      for (k in 1:nsims) {
          for (j in 2:ntimes) {
              z <- forwardsolve(sigma[, , k], x[,</pre>
                  k, j] - alpha[, , k] %*% x[, k,
                   j - 1])
              if (log) {
                   d[k, j-1] \leftarrow sum(dnorm(z, mean = 0,
                     sd = 1, log = TRUE))
              }
              else {
                   d[k, j-1] \leftarrow \exp(sum(dnorm(z,
```

In this function, times and params are as before, and x is a rank-3 array. In practice, you can think of x as an array that might have been generated by a call to the rprocess function above.

Third, we write a measurement model simulator. In this function, x, times, and params are as they were above. Notice that we give the returned array, y, rownames.

Finally, we specify how to evaluate the likelihood of an observation given the underlying state.

With these four functions in hand, we construct the pomp object:

```
> ou2 <- pomp(times = seq(1, 100), data = rbind(y1 = rep(0,
+ 100), y2 = rep(0, 100)), t0 = 0, rprocess = ou2.rprocess,
+ dprocess = ou2.dprocess, rmeasure = bvnorm.rmeasure,
+ dmeasure = bvnorm.dmeasure)
```

In the above, times are the times at which the observations (given by data) were observed. The scalar t0 is the time at which the process model is initialized: t0 should not be any later than the first observation

time times[1]. In the present case, it was easy to specify all four of the basic functions. That won't always be the case and it's not necessary to specify all of them to construct a pomp object. If any statistical method using the pomp object wants access to a function that hasn't been provided, however, an error will be generated.

We'll now specify some "true" parameters and initial states in the form of a named numeric vector:

```
> true.p <- c(alpha.1 = 0.9, alpha.2 = 0, alpha.3 = 0,
+ alpha.4 = 0.99, sigma.1 = 1, sigma.2 = 0,
+ sigma.3 = 2, tau = 1, x1.0 = 50, x2.0 = -50)</pre>
```

Note that the initial states are specified by parameters that have names ending in '.0'. This is important, since this identifies them as initial-value parameters. By default, the unobserved (state) process will be initialized with these values, and the names of the state variables will be obtained by dropping the '.0'. In applications, one will frequently want more flexibility in parameterizing the initial state. This is available: one can optionally specify an alternative initializer. See ?pomp for details.

The pomp object ou2 we just constructed has no data. If we simulate the model, we'll obtain another pomp object just like ou2, but with the data slot filled with simulated data:

```
> tic <- Sys.time()
> ou2 <- simulate(ou2, params = true.p, nsim = 1000)
> toc <- Sys.time()
> print(toc - tic)

Time difference of 9.960198 secs
> ou2 <- ou2[[1]]</pre>
```

Here, we actually ran 1000 simulations: the default behavior of simulate is to return a list of pomp objects.

There are a number of *methods* that perform operations on pomp objects. One can read the documentation on all of these by doing class?pomp. For example, one can coerce a pomp object to a data frame:

```
> as(ou2, "data.frame")
```

and if we print a pomp object, the resulting data frame is what is shown. One can access the data and the observation times using

```
> data.array(ou2)
> time(ou2)
```

One can also plot a pomp object (Fig. 1).

1.2. **The low-level interface.** There is a low-level interface to pomp objects, primarily designed for package developers. Ordinary users should have little reason to use this interface. Here, I'll just introduce each of the methods that make up this interface.

The init.state method is called to initialize the state (unobserved) process. It takes a vector or matrix of parameters and returns a matrix of initial states.

```
> x0 <- init.state(ou2, params = true.p)
> x0
```



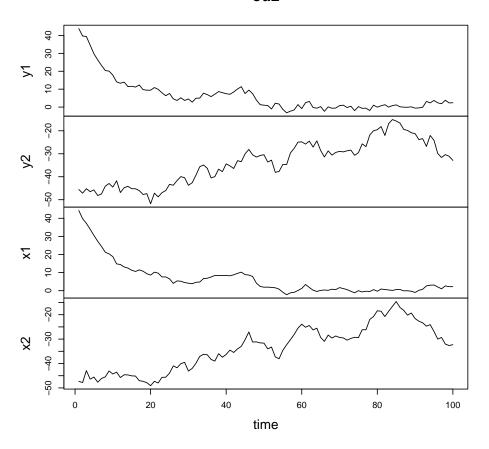


FIGURE 1. One can plot a pomp object. This shows the result of plot(ou2).

The rprocess method gives access to the process model simulator. It takes initial conditions (which need not correspond to the zero-time t0 specified when the pomp object was constructed), a set of times, and a set of parameters. The initial states and parameters must be matrices, and they are checked for commensurability. The method returns a rank-3 array containing simulated state trajectories, sampled at the times specified.

```
> x <- rprocess(ou2, xstart = as.matrix(x0), times = c(0,
+ time(ou2)), params = as.matrix(true.p))
> dim(x)

[1] 2 1 101
```

Note that the dimensions of x are nvars x nsims x ntimes, where nvars is the number of state variables, nsims is the number of simulated trajectories (which is the number of columns in the params and

xstart matrices), and ntimes is the length of the times argument. Note also that x[,,1] is identical to xstart.

The rmeasure method gives access to the measurement model simulator:

The dmeasure and dprocess methods give access to the measurement and process model densities, respectively.

All of these are to be preferred to direct access to the slots of the pomp object, because they do sanity checks on the inputs and outputs.

Since the codes above that implement the model are written in R, they're not particularly fast. To maximize computational efficiency, you'll frequently want to instead use compiled codes. For more information on this topic, see the vignette, *Using compiled code in pomp*. For the remainder of this vignette, we'll use a version of ou2 in which each of the four functions — dprocess, rprocess, dmeasure, rmeasure — is written in C. This pomp object is provided in the package; just do

> data(ou2)

to retrieve it. The C codes that make it up are included in the source distribution of the package (look in the file 'ou2.c'). Another example, an SIR model of disease transmission, is included in the package 'examples' directory.

2. Particle filter.

We can run a particle filter as follows:

```
> fit1 <- pfilter(ou2, params = true.p, Np = 1000,
+ filter.mean = T, pred.mean = T, pred.var = T)</pre>
```

Since ou2 already contained the parameters p, it wasn't necessary to specify them; we could have done

```
> fit1 <- pfilter(ou2, Np = 1000)
```

with much the same result, for example.

We can compare the results against those of the Kalman filter, which is exact in the case of a linear, Gaussian model such as the one implemented in ou2. First, we need to implement the Kalman filter.

```
tau) {
      n \leftarrow nrow(y)
      ntimes <- ncol(y)</pre>
      sigma.sq <- sigma %*% t(sigma)
      tau.sq <- tau %*% t(tau)
      inv.tau.sq <- solve(tau.sq)</pre>
      cond.dev <- numeric(ntimes)</pre>
      filter.mean <- matrix(0, n, ntimes)
      pred.mean <- matrix(0, n, ntimes)</pre>
      pred.var \leftarrow array(0, dim = c(n, n, ntimes))
      dev <- 0
      m <- x0
      v <- diag(0, n)</pre>
      for (k in seq(length = ntimes)) {
           pred.mean[, k] <- M <- a %*% m</pre>
           pred.var[, , k] <- V <- a %*% v %*% t(a) +
               sigma.sq
           q <- b %*% V %*% t(b) + tau.sq
           r \leftarrow y[, k] - b \% *\% M
           cond.dev[k] \leftarrow n * log(2 * pi) + log(det(q)) +
               t(r) \%*\% solve(q, r)
           dev <- dev + cond.dev[k]</pre>
           q \leftarrow t(b) \% \% inv.tau.sq \% \% b + solve(V)
           v \leftarrow solve(q)
           filter.mean[, k] <- m <- v \% \% (t(b) \% \%
               inv.tau.sq %*% y[, k] + solve(V, M))
      list(pred.mean = pred.mean, pred.var = pred.var,
           filter.mean = filter.mean, cond.loglik = -0.5 *
               cond.dev, loglik = -0.5 * dev)
+ }
Now we can run it on the example data we generated above.
> y <- data.array(ou2)</pre>
> a <- matrix(true.p[c("alpha.1", "alpha.2", "alpha.3",</pre>
       "alpha.4")], 2, 2)
> b < - diag(1, 2)
> sigma <- matrix(c(true.p["sigma.1"], p["sigma.2"],</pre>
       0, true.p["sigma.3"]), 2, 2)
> tau <- diag(true.p["tau"], 2, 2)
> fit2 <- kalman.filter(y, x0, a, b, sigma, tau)
```

> kalman.filter <- function(y, x0, a, b, sigma,

In this case, the Kalman filter gives us a log likelihood of fit2\$loglik=-422.4, while the particle filter gives us fit1\$loglik=-421.9.

3. The maximum likelihood by iterated filtering (MIF) algorithm

The MIF algorithm works by modifying the model slightly. It replaces the model we are interested in fitting — which has time-invariant parameters — with a model that is just the same except that its parameters take a random walk in time. As the intensity of this random walk approaches zero, the modified model approaches the fixed-parameter model. MIF works by iterating a particle filter on this

model. The extra variability in the parameters combats the particle depletion that typically plagues simple particle filters.

At the beginning of each iteration, MIF must create an initial distribution of particles in the state-parameter space. For this purpose, MIF uses a function, particles, which can be optionally specified by the user. By default, MIF uses a multivariate normal particle distribution. The particles function takes an argument, sd, that scales the width of the distribution of particles in each of the directions of the state-parameter space. In particular, this distribution must be such that, when sd=0, all the particles are identical. In this vignette, we'll use the default (multivariate normal) particle distribution.

Let's jump right in and run MIF to maximize the likelihood over two of the parameters and both initial conditions. We'll use 1000 particles, an exponential cooling factor of 0.95, and a fixed-lag smoother with lag 10 for the initial conditions:

```
> alg.pars <- list(Np = 1000, var.factor = 1, ic.lag = 10,
+ cooling.factor = 0.95)</pre>
```

Just to make it interesting, we'll start far from the true parameter values:

```
> start.p <- true.p
> start.p[c("x1.0", "x2.0", "alpha.1", "alpha.4")] <- c(45,
      -60, 0.8, 0.9)
> tic <- Sys.time()</pre>
> fit <- mif(ou2, Nmif = 1, start = start.p, pars = c("alpha.1",
      "alpha.4"), ivps = c("x1.0", "x2.0"), rw.sd = c(x1.0 = 5,
      x2.0 = 5, alpha.1 = 0.1, alpha.2 = 0, alpha.3 = 0,
      alpha.4 = 0.1, sigma.1 = 0, sigma.2 = 0, sigma.3 = 0,
      tau = 0), alg.pars = alg.pars, max.fail = 100)
> fit <- continue(fit, Nmif = 79, max.fail = 100)</pre>
> toc <- Sys.time()</pre>
> print(toc - tic)
Time difference of 1.63124 mins
> fitted.pars <- c("alpha.1", "alpha.4", "x1.0",
      "x2.0")
> cbind(start = start.p[fitted.pars], mle = signif(coef(fit,
      fitted.pars), 3), truth = true.p[fitted.pars])
        start
                  mle truth
alpha.1
          0.8
                0.907
                        0.90
alpha.4
          0.9
                0.973
                        0.99
x1.0
         45.0
               50.400 50.00
x2.0
        -60.0 -51.400 -50.00
```

One can plot various diagnostics for the fitted mif object using

```
> plot(fit)
```

Here, we'll just plot the convergence records for the log likelihood and the two α parameters (Fig. 2). In applications, a good strategy is to start several MIFs from different starting points. A good diagnostic for convergence is obtained by plotting the *convergence records* (see the documentation for conv.rec) and verifying that all the MIF iterations converge to the same parameters. One plots these—and other—diagnostics using compare.mif applied to a list of mif objects.

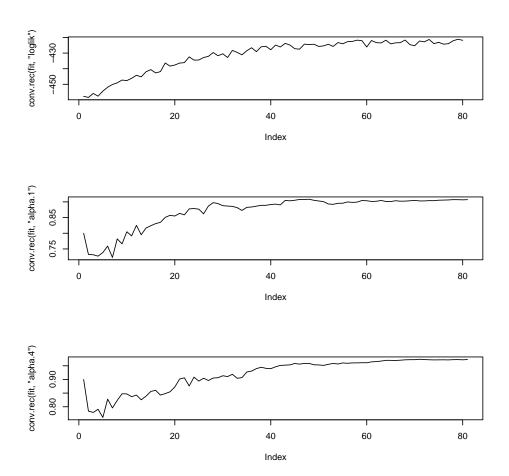


FIGURE 2. Convergence plots can be used to help diagnose convergence of the MIF algorithm.

The log likelihood of the random-parameter model at the end of the mif iterations—which should be a rough approximation of that of the fixed-parameter model—is logLik(fit)=-421.8. To get the log likelihood of the fixed-parameter model (up to Monte Carlo error) we can use pfilter:

> round(pfilter(fit)\$loglik, 1)

[1] -421.8

Like pomp objects, one can simulate from a fitted mif object (Fig. 3). In this case, the pomp is simulated at the MLE.

A. A. King, Departments of Ecology & Evolutionary Biology and Mathematics, University of Michigan, Ann Arbor, Michigan 48109-1048 USA

 $E ext{-}mail\ address:$ kingaa at umich dot edu

 URL : http://www.umich.edu/~kingaa

simulate(fit)[[1]]

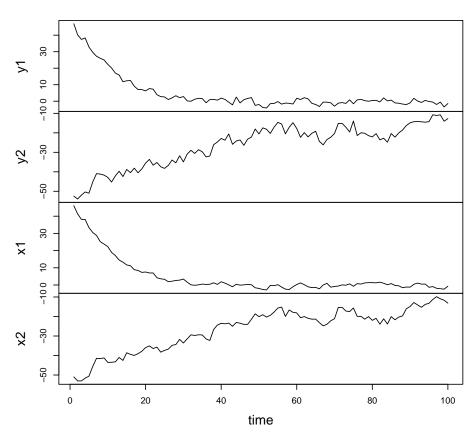


FIGURE 3. mif objects can be simulated.