Population genetics models with selection for phylogenetic inference

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Abstract

We present a phylogenetic approach rooted in the field of population genetics that more realistic models the evolution of protein-coding DNA under the assumption of stabilizing selection. The new set of models, which we collectively call selac models, fit phylogenetic data substantially better than current models, suggesting more accurate inference of phylogenies. Moreover, these models allow inference of population genetics parameters from data used for interspecific phylogenies.

Introduction

Phylogenetic analysis now plays a critical role in the fields of ecology, evolution, paleontology, medicine, conservation, and others. While the scale and impact of phylogenetic studies has increased substantially over the past two decades, by comparison the realism of the models used to infer the trees has changed relatively little. The simplest models assume neutrality between the different amino acid substitutions and may or may not include mutation bias (e.g. F81, F84, HYK85, TN93, and GTR for the former and JC69 and K80 for the latter, see Yang (2014) for an overview). The next set of models attempt to include a 'selection' term ω which is interpreted as describing stabilizing or diversifying selection depending on whether ω is less than or greater than 1, respectively. However, the link between ω and the key parameters found in standard population genetics models of Wright and Fisher or Moran, such as N_e , the distribution of s across genotype space, and mutation bias are far from clear. For example, ω is generally interpreted as indicating whether a sequence is under 'purifying' (ω < 1) or 'positive' (ω > 1) selection. However, the actually behavior of the model as is quite different. When ω < 1 the model behaves as if the resident amino acid i at a given site is favored by selection since synonymous substitutions have a higher substitution rate than any

possible non-synonymous substitutions. Paradoxically, this selection regime for the resident amino acid i persists until a substitution for another amino acid, j, occurs. As soon as amino acid j fixes, but not before, selection now favors amino acid j and all 18 others over amino acid i, the exact opposite of the case when i was the resident. Similarly, when $\omega > 1$, synonymous substitutions have a lower substitution rate than any possible non-synonymous substitutions the resident amino acid. Similar to the previous scenario, this selection against the resident amino acid i persists until a substitution occurs at which point selection now favors the former resident amino acid i as well as the 18 others. Thus, the simplest consistent interpretation of ω is that it describes the rate at which the selection regime itself changes and this change in selection perfectly coincides with the fixation of a new amino acid. As a result, ω based approaches likely only describe a subset of scenarios such as over/underdominance or frequency dependent selection (Hughes and Nei, 1988).

Fortunately, given the continual growth in computational power available to researchers, it is now possible to utilize a more general set of population genetics based models for the purpose of phylogenetic reconstruction (e.g. Halpern and Bruno, 1998; Robinson et al., 2003; Lartillot and Philippe, 2004; Rodrigue and Lartillot, 2014). This is desirable because population genetics is a mature, mathematically rigorous, and well established framework for describing biological evolution. One lesson from the field of population genetics is while there are only a few fundamental evolutionary forces at play, i.e. mutation, drift, selection, and linkage effects, describing the evolutionary behavior of a system in which there are non-linear interactions between different sites quickly extermely challenging. Fortunately, under the simplifying assumptions of additivity between and within sites, calculating stationary and substitution probabilities are relatively straightforward to calculate. As a result, fitting additive models of the evolutionary process to sequence data is computationally feasible. While critiques may argue such models are overly simplistic, in terms of fitting the data these simplistic models are still a fast improvement over standard approaches as we show below. Further, our 'simple' models are null hypotheses for more complex models that relax some of our more restrictive assumptions, such as substitution properties being independent of local structure or location. Another major advantage to our approach is that the parameters estimated are biologically meaningful. Here we show how a phylogenetic method based on fundamental results from the field of population genetics can be used to more accurately estimate branch lengths and nucleotide specific mutation rates. Because the math behind our model is mechanistically derived, our method can also be used to make quantitative inferences on the optimal amino acid sequence of a given protein as well as the average synthesis rate of each protein used in the analysis. The mechanistic basis of our model also means it can be easily extended to include more biological realism and test explicit hypotheses about sequence evolution.

We model the substitution process using the classic the Wright-Fisher (WF) model (Kimura, 1962; Wright, 1969; Iwasa, 1988; Berg and Lässig, 2003; Sella and Hirsh, 2005). For simplicity, we ignore linkage effects and, as a result of this and other assumptions, our model behaves in a site independent manner. Our approach is developed in the same vein as previous phylogenetic applications of the WF model (e.g. Muse and Gaut, 1994; Halpern and Bruno, 1998; Yang and Nielsen, 2008; Rodrigue et al., 2005; Koshi and Goldstein, 1997; Koshi et al., 1999; Dimmic et al., 2000; Thorne et al., 2012; Lartillot and Philippe, 2004; Rodrigue and Lartillot, 2014). Similar to Lartillot and Philippe (2004); Rodrigue and Lartillot (2014) we assume there is a finite set of rate matrices describing the substitution process and that each position within a protein must assigned to a particular rate matrix category. Unlike these other researchers, we assume a priori there are 20 different families of rate matrices, one family for when a given amino acid is favored at a site. As a result, our approach allows us to quantiatively evaluate the support for a particular amino acid being favored at a particular position within the protein encoded by a particular gene.

Because our approach requires twenty families of 61×61 matrices, the number of parameters needed to implement our model would, without further simplification, be extremely large. To reduce the number of parameters needed while still maintaining a high degree of biological realism, we construct our gene and amino acid specific substitution matrices using a submodel nested within our substitution model. As a result, our nested modeling framework requires only a handful of genome wide parameters such as nucleotide specific mutation rates (scaled by effective population size), side chain physiochemical weighting parameters, and a shape parameter describing the distribution of site sensitivities around a mean value of 1. In addition to these genome wide parameters, our model requires a gene specific expression parameter ψ , which describes the average synthesis rate of a protein and scales the strength of selection at the gene level. Our model also requires the designation of an optimal amino acid at each position or site within a coding sequence which, in turn, makes it the largest category of parameters we estimate. Because we use a submodel to derive our substitution matrices, our model requires the estimation of a fraction of the parameters required when compared to approaches where the substitution rates are allowed to vary independently (Halpern and Bruno, 1998; Lartillot and Philippe, 2004; Rodrigue and Lartillot, 2014).

The work we present here contributes to the field of phylogenetics and molecular evolution in a number of ways. Our model provides an complementary example to Thorne et al. (2012) studies of how models of molecular and evolutionary scales can be combined together in a nested manner. While the mapping between genotype and phenotype is more abstract than Thorne et al. (2012), our approach has the advantage of not requiring knowledge of a protein's native folding. Our use of model nesting also allows us to formulate and

test specific biological hypotheses. For example, we are able to compare a model formulation which assumes that physio-chemical deviations from the optimal sequence are equally disruptive at all sites within a protein to one which assumes they vary between sites. We've utilized the same nested, population genetic based approach in more traditional genomic analyses (e.g. Gilchrist, 2007; Shah and Gilchrist, 2011; Gilchrist et al., 2015). That work and our current work illustrates how more information can be extracted from sequence data when more biologically based models are used. By linking the strength of stabilizing selection to gene expression, we can weight the historical information encoded in different genes in a biologically plausible manner while simultaneously estimating their expression levels.

Results

Model Performance

- 1. Using Δ AIC as our measure, we see that even despite the need for estimating the optimal amino acid at each position in each protein, our model performs astronomically better than GTR, GY94, or YN08.
- 2. Including the random effects term G not only provides greater biological realism than assuming G = 1, it provides substantially better model fit and improves the ΔAIC score by over 16,000 units.

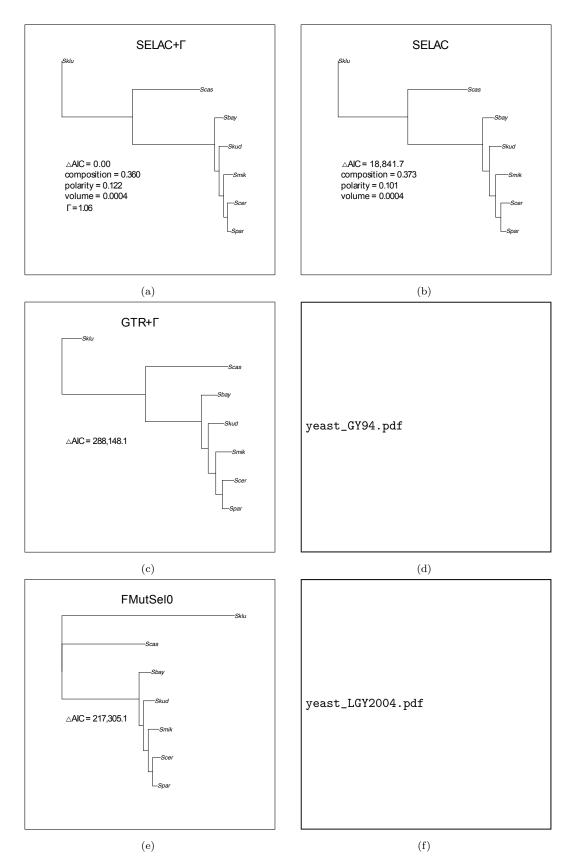


Figure 1: Maximum Likelihood Trees for (a) selac, (b) selac with uniform sensitivity G=1, (c) GTR, (d) GY94, and (e) YN08, (f) Lartillot and Philippe (2004).

		Parameters			Model
Model	$\log Lik$	Estimated	AIC	$\Delta { m AIC}$	Weight
GTR + Gamma	-557990.3	648	1,117,277	454,694	< 0.001
GY94	-509625.1	223	1,019,698	$357,\!115$	< 0.001
MutSel	-518090.6		1,120,945	$458,\!362$	< 0.001
SelAC: GTR	-382725.8		850,391	187,808	< 0.001
SelAC: UNREST	-381881.4		$678,\!816$	16,233	< 0.001
SelAC: UNREST + Gamma	-373765.6		$662,\!583$	0	0.999

Table 1: random effects, RE

- 3. Table of number of parameters, estimates for key parameters (or their summaries), and Δ AIC values.
- 4. Selac provides estimates of gene expression which are positively correlated with empirical estimates and explain 20-30% of the variation in the empirical measurements taken during log growth phase.¹

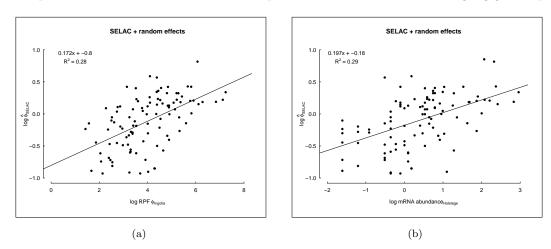


Figure 2: Comparison of log protein synthesis rate ψ estimated by selac and (a) estimates from ribosome profile footprint data of Ingolia et al. (2009) and Holstege et al. (1998).

5. By linking transition rates $q_{i,j}$ to gene expression ψ , our approach allows use the same model for genes under varying degrees of purifying selection. The traditional approach of concatenating gene sequences together is equivalent to assuming the same average protein synthesis rate ψ for all of the genes. By assuming the strength of stabilizing selection for the optimal sequence, \vec{a}_* , is proportional to ψ , our model allows us to estimate $\phi = \psi/\mathbf{B}$ for each gene. Our results clearly indicate that this information is available and accounting for it in our model substantially improves our model fit.

¹mikeg: We should replace the estimates of ψ with estimates of ϕ which is $\psi/\mathbf{B}(\vec{a}_{\text{obs}}|\vec{a}_*)$. I've decided it's better to replace ϕ with ψ rather than the other way around. Thus, $\phi = \psi/\mathbf{B}(\vec{a})$ is the target synthesis rate of error free protein \vec{a} and and $\phi = \psi$ when $\vec{a} = \vec{a}_*$ since, by definitioon, $\mathbf{B}(\vec{a}_*) = 1$.

Discussion

Key Points

- 1. Importance of phylomodels
- 2. Limitations of most common models
- 3. Pop gen alternative
 - (a) Link to previous workers
 - (b) Illustrate advantages of pop gen
 - i. Improved estimates of classical parmeters
 - ii. More realistic behavior over time
 - iii. Exstensible to other researchers approaches
 - A. Structure information
 - B. Selection on CUB
 - C. Approach can even be expanded to other types of sequence data
 - iv. Current results suggest such extensions will allow researchers access to even more information embedded in DNA coding sequences.
- 4. Lots of sequences available and in pipeline, let's get to it comrades!
- Work falls within the mutation-selection framework
- One important manner in which our approach moves beyond Halpern and Bruno (1998); Robinson et al. (2003); Lartillot and Philippe (2004); Thorne et al. (2012); Rodrigue and Lartillot (2014) is in that we parameterize the model and fit branch lengths simultaneously rather than in two separate steps.²
- Assumptions of additivity and no epistasis are unrealistic but can be viewed as a first order approximation to these more complex scenarios and a starting point for later relaxing these assumptions.
- Implicitly, our model assumes that all genes are essential because an organism that is homozygous for alleles with zero activity (i.e. no benefit) would have to spend an infinite amount of energy to achive a

 $^{^2}$ mikeg: Jeremy or Brian: I know this is true for Halpern and Bruno and am 95% sure it applies to Jeff Thorne's work. Can you confirm its true for the Lartillot references?

target functionality synthesis rate $\psi > 0$. Some ways this assumption could be relaxed include making fitness W a function of ψ such that $W(\psi = 0) > 0$ or incorporating functional overlap between proteins into our calculations.

- Our approach requires relatively few parameters to populate our transition matrices.
 - 1. Distance function $d(a_i, a_*)$: If n_d is the number of physiochemical properties examined, the number of parameters estimated is $n_d 1$
 - 2. Benefit function **B**: A single α_G , an optimal amino acid for each site, and an ancestral amino acid state for each site.
 - 3. Gene expression ψ : One ψ for each gene analyzed.
 - 4. Mutation bias: Depending on the choice of mutation model, from 1 to 8 global parameters
- Our approach can be expanded by allowing the optimal amino acid to change during the course of
 evolution. This would result in the need of additional parameters describing the rates at which the
 optimal amino acid switches at a site.
- By linking fitness to energy flux, selection is an increasing function of gene expression ψ . As a result, the strength of selection and rates of evolution vary between genes in a biologically plausable manner.
- Approach assumes rank order of amino acid function given a particular optimal AA does not change
 between sites. Since the importance of an amino acid's physiochemical properties likely changes with
 where it lies in a folde protein, one way to incorporate such effects is to test whether the data supports
 multiple sets of Grantham weights, rather than a single set.
- Because ψ is determined, in part, by our choice of a reference distance weighting $\alpha_v = 4 \times 10^{-4}$. A larger and more informative set of Grantham weights might reduce the noise in our estimates of ϕ .

Additional Points That Need to Be Mentioned

- Lartillot and Rodrigue's recent work
- Despite their widespread use, as defined by GoldmanAndYang1994, purifying and diversifying selection
 are very narrow categories of selection that mostly apply to cases of positive and negative frequency
 dependent selection at the level of a particular amino acid.

- Instead of focusing on detecting evidence for adaptation, our approach focuses on quantifying the strength of selection behind any adaptive (or non-adaptive) changes.
- We are estimating a physiochemical based fitness landscape using sequences at the tip of a tree.
- In this study we develop a model where the substitution rate of an allele is based on the substitution probability of an allele under selection, mutation bias, and genetic drift, per standard models of population genetics.
- In developing our model, we assume that for each protein coding gene there is a single amino acid sequence which executes its intended function better than any other sequence, i.e. is optimal.
- We assume the strength of selection for the optimal sequence increases with the target synthesis rate of the functionality the gene provides. That is genes with higher target expression levels are under stronger selection than genes with lower target expression levels.
- We also assume that the functionality of other amino acid sequences declines as the physiochemical properties of the sequence deviates from that of the optimal sequence.
- Because we assume that a protein's functionality is a declining function of the product of the physiochemical distances of each of the protein's amino acid from the optimal, we can treat the evolution at each amino acid position in a site independent manner. An approach which is almost universally used in other phylogenetic models.
- As a result, unlike most phylogenetic approaches, our model requires 20 different 64x64 rate matrices, one for when each amino acid is the optimal one.
- Even though our model requires a large number of matrices, because of our assumption that a protein's functionality is a declining function of physiochemical distance from the optimum, we are able to parameterize our 20 matrices using only a handful of parameters which we estimate from the data.
- Two additional key assumption of our model is that (a) the organism has an average target synthesis rate ψ for the functionality provided by each protein and (b) that protein synthesis is under some form of regulatory control such that the this average functionality production target is met. As a result, the relative rate of protein synthesis increases as the sequence's functionality declines due to deviation from the optimal sequence. This behavior, in turn, means that the energetic cost of protein synthesis for an allele deviating from the optimal sequence increases with the target synthesis rate ψ . For example, a

protein encoding allele which has a 10% reduction in functionality will have the same energetic burden and selective cost relative to its optimal sequence as a protein encoding allele of similar length which has a 20% reduction in functionality but whose target synthesis rate is 1/2 of the first protein.

• In its current formulation, our model is only applicable to protein coding sequences. However, it should be applicable to non-coding sequences so long as one has a mapping function between gene sequence and gene function.

Methods

We link genotype, phenotype, fitness, drift, and fixation, by extending the approach we have successfully used to quantify the evolutionary forces of fitness, drift, and fixation on to the evolution codon usage bias based on an organism's coding sequences (Gilchrist and Wagner, 2006; Gilchrist, 2007; Shah and Gilchrist, 2011; Gilchrist et al., 2015). More specifically, in order to link genotype, phenotype, and fitness, we assume that organisms have set of fixed, but a priori unspecified, metabolic requirements and the organism meets these requirements through the appropriate translation of its proteome. We assume that each protein has, on average, a target synthesis rate of ψ and, for now, that ψ is fixed over the tree. We also assume that natural selection favors genotypes that are able to synthesize their proteome efficiently than their competitors and that each savings of an high energy phosphate bond per unit time leads to a constant proportional gain in fitness q. In terms of the functionality of the protein encoded, we assume that for any given gene there exists an optimal amino acid sequence \vec{a}_* and that, by definition, a complete, error free peptide consisting of a_* provides one unit of the gene's functionality. Thus ψ for a given protein is determined by both the organism's metabolic requirements and the functionality of the protein encoded by \vec{a}_* . Our approach allows us to link amino acid sequence and gene expression directly to genotype fitness and, in turn, substitution rate in a general, yet simple and biologically plausible, manner. ³

The overall structure of our model involves a codon mutation model combined with a selection model based on the cost and benefits of translating a given genotype and the target gene expression rate of a gene. We fit the model using maximum likelihood and \dots ⁴

³mikeg: Moved from methods.

⁴mikeg: Jeremy, care to elaborate here?

Allele Substitution Model

Mutation Rate Matrix μ :

We begin by defining a time reversible model for mutation rates between individual bases. We use the existing drift driven substitution model JC Jukes and Cantor (1969), also referred to as a General Time Reversible (GTR) Tavare (1986); Yang (2014) which has 9 independent parameters, the most possible for a model that assumes mutations occur independently between nucleotides. This results in a 4x4 mutation matrix, where each entry describes the instantaneous rate of change between a pair of nucleotides for the most recent common ancestor at that site. This is converted to a 64 × 64 codon mutation matrix μ where entries $\mu_{i,j}$ describe the mutation rate from codon i to j under a 'weak mutation' assumption. That is, the rate of allele fixation is much greater than $N_e\mu$ and $N_e\mu \ll 1$, such that evolution is mutation limited, codon substitutions only occur one nucleotide at a time and, as a result, the rate of change between any pair of codons that differ by more than one nucleotide is zero. ⁵

Protein Synthesis Cost-Benefit Function η :

Our model links fitness to the product of the cost-benefit function of a gene g, η_g , and the organism's average target synthesis rate of the functionality provided by gene g, ψ_g . This is because the average flux energy an organism spends to met its target functionality provided by gene g is $\eta_g \times \psi_g$. In order to link genotype to our cost-benefit function η , we begin by defining our benefit function. This

Benefit: Our benefit function **B** measures the functionality of the amino acid sequence \vec{a}_i encoded by a set of codons \vec{c}_i , i.e. $a(\vec{c}_i) = \vec{a}_i$ relative to that of an optimal sequence \vec{a}_* . By definition, $\mathbf{B}(\vec{a}_*) = 1$ and $\mathbf{B}(\vec{a}_i|\vec{a}_*) < 1$ for all other sequences. We assume all amino acids within the sequence contribute to protein function and that this contribution declines as an inverse function of physiochemical distance between each amino acid and the optimal. Formally, we assume that

$$\mathbf{B}(\vec{a}_i|\vec{a}_*) = \left(\frac{1}{n_g} \sum_{p=1}^{n_g} (1 + G_p d(a_{i,p}, a_{*,p}))\right)^{-1}$$
(1)

where n_g is the length of the protein, $d(a_{i,p}, a_{*,p})$ is a weighted physiochemical distance between the amino acid encoded in gene i for position p and $a_{*,p}$ is the optimal amino acid for that position of the protein. For simplicity, we define the distance between a stop codon and a sense codon as infinite and, as a result,

⁵mikeg: Jeremy, please update for UNREST.

nonsense mutations are always lethal. The term G_p describes the sensitivity of the protein's function to deviation in Grantham's physiochemical space. We assume that $G_p \sim \text{Gamma} (\alpha = \alpha_G, \beta = 1/\alpha_G)$ in order to ensure $\mathbb{E}(G_p) = 1$. At the limit of $\alpha_G \to \infty$, the model collapses to a model with uniform sensitivity of $G_p = 1$ for all positions p. $\mathbf{B}(\vec{a}_i|\vec{a}_*)$ is inversely proportional to the average physiochemical deviation of an amino acid sequence \vec{a}_i from the optimal sequence \vec{a}_* weighted by each sites sensitivity to this deviation. $\mathbf{B}(\vec{a}_i|\vec{a}_*)$ can be generalized to include second and higher order terms of the distance measure d.

Cost: Protein synthesis involves both direct and indirect assembly costs. Direct costs consist of the high energy phosphate bonds in ATP or GTP's used to assemble the ribosome on the mRNA, charge tRNA's for elongation, move the ribosome forward along the transcript, and terminate protein synthesis. As a result, direct protein assembly costs are the same for all proteins of the same length. Indirect costs of protein assembly are potentially numerous and could include the cost of amino acid synthesis as well the cost and efficiency with which the protein assembly infrastructure such as ribosomes, aminoacyl-tRNA synthetases, tRNAs, and mRNAs are used. When these indirect costs are combined with sequence specific benefits, the probability of a mutant allele fixing is no longer independent of the rest of the sequence (Gilchrist et al., 2015) and, as a result, model fitting becomes substantially more complex. Thus for simplicity, in this study we ignore any indirect costs of protein assembly that vary between genotypes and define,

$$\mathbf{C}(\vec{c_i}) = \text{Energetic cost of protein synthesis.}$$
 (2)

$$=C_1+C_2n\tag{3}$$

where, C_1 and C_2 represent the direct and indirect costs in ATPs of ribosome initiation and peptide elongation, respectively, where $C_1 = C_2 = 4$ ATP. ⁶

Defining Physiochemical Distances: Assuming that functionality declines with an amino acid a_i 's physiochemical distance from the optimum amino acid a_* at each site provides a biologically defensible way of mapping genotype to protein function that requires relatively few free parameters. In addition, our approach naturally lends itself to model selection since we can compare the quality of our model fits using different mixtures of physiochemical properties. Following Grantham (1974), we focus on using composition c, polarity p, and molecular volume v of each amino acid's side chain residue to define our distance function, but emphasize that other properties could be used. We use the euclidian distance between residue properties

 $^{^6}$ mikeg: Jeremy, can we let C_1 vary as a factor of C_2 and then refit the model?. Answer: leave for later.

where each property c, p, and v has its own weighting term, α_c , α_p , α_v , respectively, which we refer to as 'Grantham weights'. Because physiochemical distance is ultimately weighted by a gene's specific average protein synthesis rate ψ , another parameter we estimate, there is a problem with parameter identifiablity. Ultimately, the scale of gene expression is affected by how we measure physiochemical distances which, in turn, is determined by our choice of Grantham weights. As a result, we set $\alpha_v = 4 \times 10^{-4}$, the value originally estimated by Grantham, and recognize that our our estimates of α_c and α_p and ψ are scaled relative to this choice for α_v . ⁷ More specifically,

$$d(a_i, a_*) = \sqrt{\alpha_c (c(a_i) - c(a_*))^2 + \alpha_p (p(a_i) - p(a_*))^2 + \alpha_v (v(a_i) - v(a_*))^2}.$$

Linking Cost of Protein Synthesis to Allele Substitution

Next we link the protein synthesis cost-benefit function η of an allele with its fixation probability. First, we assume that each protein encoded within a genome provides some beneficial function and that the organism needs that functionality to be produced at a target average rate ψ . By definition, the optimal amino acid sequence for a given gene, \vec{a}_* , produces one unit of functionality. Second, we assume that protein expression is regulated by the organism to ensure that functionality is produced at rate ψ . As a result, the realized average protein synthesis rate of a gene, ϕ , is equal to $\psi/\mathbf{B}(\vec{a})$ and the total energy flux allocated towards meeting the target functionality of a particular gene is $\eta(\vec{c})\psi$. As we shall show below, the fitness cost for a genotype encoding a suboptimal protein sequence stems from the need to produce $1/\phi$ proteins in order to produce the equivalent functionality of one protein consisting of the optimal amino acid sequence a_* .

Third, we assume that every additional ATP spent per unit time to meet the organism's target function synthesis rate ψ leads to a slight and proportional decrease in fitness W. This assumption, in turn, implies

$$W_i(\vec{c}) \propto \exp\left[-q\,\eta(\vec{c}_i)\psi\right].$$
 (4)

where q describes the decline in fitness with every ATP wasted per unit time used to measure ψ and ϕ .

⁷mikeg: Jeremy, is this correct? If not, please correct.

Correspondingly, the ratio of fitness between two genotypes is,

$$W_i/W_j = \exp\left[-q \,\eta(\vec{c}_i)\psi\right] / \exp\left[-q \,\eta(\vec{c}_j)\psi\right]$$
$$= \exp\left[-q \,(\eta(\vec{c}_i) - \eta(\vec{c}_j))\,\psi\right]$$

Given our formulations of \mathbf{C} and \mathbf{B} , the fitness effects between sites are multiplicative and, therefore, the substitution of an amino acid at one site can be modeled independently of the amino acids at the other sites within the coding sequence. As a result, the fitness ratio for two genotypes differing at a single site p simplifies to

$$W_i/W_j = \exp\left\{-q\left(C_1 + C_2 n\right) \frac{1}{n} \sum_{p \in \mathbb{P}} \left[d\left(a_{i,p}, a_{*,p}\right) - d\left(a_{j,p}, a_{*,p}\right)\right] \psi\right\}$$
 (5)

where \mathbb{P} represents the codon positions in which $\vec{c_i}$ and $\vec{c_j}$ differ. Fourth, we make a weak mutation assumption, such that alleles can differ at only one position at any given time, i.e. $|\mathbb{P}| = 1$, and that the population is evolving according to a Fisher-Wright process. As a result, the probability a new mutant j introduced via mutation into a resident population i with effective size N_e will go to fixation is,

$$u_{i,j} = \frac{1 - (W_i/W_j)^b}{1 - (W_i/W_j)^{N_e}}$$

$$= \frac{1 - \exp\left\{-\frac{q}{n}\left(C_1 + C_2n\right)\left[d\left(a_i, a_*\right) - d\left(a_j, a_*\right)\right]\psi b\right\}}{1 - \exp\left\{-\frac{q}{n}\left(C_1 + C_2n\right)\left[d\left(a_i, a_*\right) - d\left(a_j, a_*\right)\right]\psi 2N_e\right\}}$$

where b=1 for a diploid population and 2 for a haploid population (Kimura, 1962; Wright, 1969; Iwasa, 1988; Berg and Lässig, 2003; Sella and Hirsh, 2005). Finally, assuming a constant mutation rate between alleles i and j, $\mu_{i,j}$, the substitution rate from allele i to j can be modeled as,

$$q_{i,j} = \frac{2}{b}\mu_{i,j}N_e u_{i,j}.$$

where, given our weak mutation assumption, $\mu_{i,j} = 0$ when two codons differ by more than one nucleotide. In the end, each optimal amino acid has a separate 64 x 64 substitution rate matrix \mathbf{Q}_a , which incorporates selection for the amino acid (and the fixation rate matrix this creates) as well as the common mutation parameters across optimal amino acids. This results in the creation of 20 \mathbf{Q}_a matrices, one for each amino acid, with up to XXXXX unique rates, based on few parameters (one to six mutation rates, two free Grantham weights, the cost of protein production \mathbf{C} , the target functionality synthesis rate, and optimal amino acid at each site), which we infer from the data. Our model can be generalized to allow transitions between optimal amino acids as well as between codons, which would result in a $(21 \times 64) \times (21 \times 64) = 1344 \times 1344$ matrix.

While the overall model does not assume equilibrium, we still need to scale our substitution matrices \mathbf{Q} . Traditionally, it is rescaled such that at equilibrium, one unit of branch length represents one expected substitution per site. In our case, we want to do this scaling across all the matrices, since the branch lengths are used in common across the gene. One wrinkle is that this must be done taking optimal amino acid frequency into account. Because we are using 21 64 × 64 matrices rather than a single 1344 × 1344 matrix, our scaling done jointly across all the 21 matrices to allow branch lengths under the fixed optimal amino acid model to ensure that branch lengths are comparable. We calculate from the data a vector of 1344 empirical frequencies, π for each of the 64 codons observed when assuming each of 21 possible as the optimal amino acid (including stop codons). A scaling factor is then calculated as the average rate $-\sum_i \mu_i \pi_i = 1$, where i indexes a particular codon for a particular optimal amino acid. The final substitution-rate matrix is the original substitution-rate matrix multiplied by this scaling factor. This matrix can then be applied to all the sites to calculate the likelihood. Finally, in the usual manner the diagonal elements of \mathbf{Q} defining substitution for a given optimal amino acid are fixed so that the rows sum to zero, which allows for $P(t) = \exp{[\mathbf{Q}_a t]}$ to be calculated for the matrix π_j whose elements give the probabilities that codon j replaces codon i after time t.

Given our assumption of independent evolution among sites, the probability of the whole data set is the product of the probabilities of observing the data at each individual site. Thus, the log likelihood of an individual site is calculated as

$$L(Q_a) \propto P(D|Q_a, T)$$
 (6)

In this case, the data, \mathbf{D} , are the observed codon states at the tips of a phylogeny \mathbf{T} , whose topology are known. The pruning algorithm of Felsenstein (1981) is used to calculate $\mathbf{L}(\mathbf{Q}_a)$. The log likelihood is maximized by estimating the global parameters: $C q N_e$, 8 mutation parameters which are scaled by $2N_e/b$, and two Grantham distance parameters, α_c and α_p , and the sensitivity distribution parameter α_G . For each gene, we also estimate its target functionality synthesis rate ψ and the optimal amino acid for each position in the protein. When estimating α_G , the likelihood then becomes the average likelihood which we calculate using Laguerre quadrature with k=8 points (Yang, 1994; Felsenstein, 2001).

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