RH: BEAULIEU ET AL.— Pop. Gen. Based Phylo.

- Population Genetics Based Phylogenetics Under Stabilizing Selection for an Optimal Amino Acid Sequence: A Nested Modeling Approach
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We present a novel phylogenetic approach rooted in the field of population genetics that more realistically models the evolution of protein-coding DNA under the assumption of 17 consistent, stabilizing selection for a gene specific, optimal amino acid sequence. In addition 18 to being consistent with the fundamental principles of population genetics, our new set of 19 models, which we collectively call SelAC (Selection on Amino acids and Codons), fit 20 phylogenetic data much better than popular models, suggesting strong potential for more 21 accurate inference of phylogenetic trees and branch lengths. SelAC also demonstrates that a large amount of biologically meaningful information is accessible when using a nested set 23 of mechanistic models. For example, for each position SelAC provides a probabilistic estimate of any given amino acid being optimal. SelAC also assumes the strength of selection is proportional to the expression level of a gene and, therefore, provides gene 26 specific estimates of protein synthesis rates. Finally, because SelAC is a nested approach 27 based on clearly stated biological assumptions, it can be expanded or simplified as needed.

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Phylogenetic analyses plays a critical role in most aspects of biology, particularly in
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   the fields of ecology, evolution, paleontology, medicine, and conservation. While the scale
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   and impact of phylogenetic studies has increased substantially over the past two decades,
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   the realism of the mathematical models on which these analyses are based has changed
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   relatively little by comparison. The most popular models of DNA substitution used by
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   molecular phylogentics are simple nucleotide models that are indifferent to the type of
   sequences to which they are applied. For example, when evaluating protein-coding
   sequences these models are inherently agnostic with regards to the different amino acid
   substitutions and their impact on gene function (e.g. F81, F84, HYK85, TN93, and GTR,
   see Yang (2014) for an overview) and, as a result, cannot describe the behavior of natural
   selection at the amino acid or protein level. To address this critical shortcoming, Goldman
   and Yang (1994) and Muse and Gaut (1994) independently introduced models which
   assumed that differences in the physico-chemical properties between amino acids, or
   physico-chemical distances for short, could affect substitution rates. These
   physico-chemical approaches as originally described have rarely been used for empirical
   data; instead these models have served as the basis for an array of simpler and, in turn,
   more popular models that, starting with (??), assume an equal fixation probability for all
   non-synonymous mutations. Thus, these simpler models initially employed a single term \omega
   to model the differences in fixation probability between nonsynonomous and synonomyous
   changes at all sites. To improve their realism, more complex forms have been developed
   that allow \omega to vary between sites or branches (as cited in ?) and include selection on
   different synonyms for the same amino acid (e.g. Yang and Nielsen 2008)
          Despite these extensions, the nature of selection these models describe are a far cry
   from how biologists intuitively envision the nature of natural selection on protein
   sequences. In Goldman and Yang (1994); ?); ? and later studies based on their work, \omega is
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   interpreted as indicating whether a sequence is under consistent 'purifying' (\omega < 1) or
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'diversifying' ($\omega > 1$) selection. However, the biological behavior the model describes is quite different¹. Because synonymous substitutions have a higher substitution rate than any possible non-synonymous substitutions when $\omega < 1$, the model behaves as if the 57 resident amino acid i at a given site is favored by natural selection. Even when ω is allowed 58 to vary between sites, the symmetrical nature of the model means that for any given site the strength of selection for the resident amino acid i over its 19 alternatives is equally strong. Paradoxically, the selection for i persists until a substitution for another amino acid, j, occurs. As soon as amino acid j fixes, but not before, selection now favors amino acid j equally over all other amino acids, including i. This is now the opposite scenario to when i was the resident. Similarly, when $\omega > 1$, synonymous substitutions have a lower substitution rate than any possible non-synonymous substitutions from the resident amino acid. Again due to the model's symmetry, the selection against the resident amino acid i is equally strong relative to alternative amino acids. The selection against the resident amino acid i persists until a substitution occurs at which point selection now favors i, as well as the 18 others, to the same degree it was previously disfavored. Thus, the simplest and most consistent interpretation of ω is that it represents the rate at which the selective environment itself changes, and this change in selection perfectly coincides with the 71 fixation of a new amino acid. This, in turn, implies that the rate of shifts in the optimal (or pessimal) amino acid is on the time scale as the rate of substitution. Contrary to popular belief, ω does not describe whether a site is evolving under a constant regime of stabilizing or diversifying selection, but instead how the selective environment changes over time under and even then under very limited conditions. Given this behavior, ω based models only reasonably approximate a subset of scenarios such as over-/under-dominance or positive/negative frequency dependent selection (Hughes and Nei 1988; Nowak 2006),

¹Although Goldman and Yang (1994) use a more complex, physico-chemical based distance function instead of ω , the criticisms below also apply to this work.

where selection is perfectly symmetrical.

Here we propose a new approach where selection is based explicitly on selection to 80 minimize the cost-benefit function η of a protein where protein function is determined soley 81 by the physico-chemical properties of the primary amino acid sequence. Our approach, 82 which we call SelAC (Selection on Amino acids and Codons), is developed in the same vein 83 as previous phylogenetic applications of the Wright-Fisher process (e.g. Muse and Gaut 1994; Halpern and Bruno 1998; Yang and Nielsen 2008; Rodrigue et al. 2005; Koshi and Goldstein 1997; Koshi et al. 1999; Dimmic et al. 2000; Thorne et al. 2012; Lartillot and Philippe 2004; Rodrigue and Lartillot 2014). Similar to Lartillot's work (Lartillot and Philippe 2004; Rodrigue and Lartillot 2014), we assume there is a finite set of rate matrices describing the substitution process and that each position within a protein is assigned to a particular rate matrix category. Unlike this previous work, we assume a priori there are 20 different families of rate matrices, one family for when a given amino acid is favored at a site. As a result, SelAC allows us to quantitatively evaluate the support for a particular amino acid being favored at a particular position within the protein encoded by a particular gene.

Biologically, we know protein-coding DNA sequences largely evolve through the introduction of new mutations that either become fixed or lost due to selection and/or drift. Selection on protein coding regions can take many forms, but can generally be related to the cost of producing the protein and the functional benefit (or potential harm) caused by the protein product. The gene specific cost of protein synthesis can be affected by the amino acids used, the direct and indirect costs of peptide assembly by the ribosome, the use of chaparones to aid in folding, and even the expected lifespan of the protein. Importantly, these costs can be computed to varying degrees of realism (e.g. Wagner 2005; ?). We have previously presented models of protein synthesis costs that, alternatively, take into account the cost of ribosome pausing (Shah and Gilchrist 2011) or premature

termination errors (Gilchrist and Wagner 2006; Gilchrist 2007; Gilchrist et al. 2009).

Protein benefit or 'function' can be affected by the amino acids at each sites and 106 their interactions. As a result, amino acid substitutions can affect the functionality at key 107 catalytic sites or, more broadly, the probability of particular protein fold and, in turn, the 108 expected functionality of the protein. Linking amino acid sequence to protein function is a 109 daunting task; thus for simplicity, we assume that for any given desired biological function 110 to be carried out by a protein, that (a) the biological importance of this protein function is 111 invariant across the tree, (b) single optimal amino acid sequence that carries out this 112 function best, and (c) the functionality of alternative amino acid sequences declines with 113 their physico-chemical distance from the optimum on a site by site basis. We readily 114 acknowledge that sequence space may have more than one local optimum, that 115 physico-chemical distance from the optimal primary amino acid sequence is likely a poor 116 model of protein function, and that the biological importance of a function can vary over time. Nevertheless, we believe our cost-benefit approach to be a substantial advance of the 118 more simplistic ω models, is complementary to the work of others in the field (e.g. Thorne 119 et al. 2012; Rodrigue and Lartillot 2014), and, in turn, lays the foundation for more 120 realistic work in the future. For instance, by assuming there is an optimal amino acid for 121 each site, SelAC naturally leads to a non-symmetrical and, thus, more cogent model of 122 protein sequence evolution. Because the strength of selection depends on an additive 123 function of amino acid physico-chemical properties, an amino acid more similar to the 124 optimum has a higher probability of replacing a more dissimilar amino acid than the 125 converse situation. Further, SelAC does not assume the system is always at the optimum 126 or pessimum point of the fitness landscape, as occurs when $\omega < 1$ or > 1, respectively. Importantly, the cost-benefit approach underlying SelAC allows us to link the 128 strength of selection on a protein sequence to its gene's expression level. Despite its well 129 recognized importance in determining the rate of protein evolution (e.g. Drummond et al.

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2005, 2006), phylogenetic models have ignored the fact that expression levels vary between genes. In order to link gene expression and the strength of stabilizing selection on protein sequences, we simply assume that the strength of selection on a gene is proportional to the average protein synthesis rate of the gene.

One possible mechanism consistent with this link between the strength of selection 135 and gene expression is the assumption of compensatory gene expression, where a reduction 136 in protein function is compensated for by increasing the abundance of the protein. For 137 example, a mutation which reduces the functionality of the protein to 90% of the optimal 138 protein, would require 1/0.9 = 1.11 of these suboptimal proteins be produced relative to 139 the optimal protein in order to maintain the same amount of that protein's functionality in 140 the cell. Because the energetic cost of an 11% increase in gene expression varies 141 proportionally with gene expression, our assumptions naturally link functionality and expression. Under what circumstances cells actually respond in this manner, remains to be 143 determined. The fact that our method allows us to explain 13-23\% of the variation in gene expression measured using RNA-Seq, suggests that this assumption is a reasonable starting 145 point. More importantly, by linking the strength of stabilizing selection for an optimal 146 amino acid sequence to gene expression, we can effectively weight the phylogenetic 147 information encoded in genes evolving at different rates. 148

We do note that because SelAC infers the optimal amino acid for each site, it is substantially more parameter rich than more commonly used models such as GTR+Γ,
GY94, and FMutSel. Despite this increase in number of model parameters, SelAC drastically outperforms these models with AICc values on the order of 10,000s to 100,000s.
We predict that SelAC's performance could be improved even further if we use a hierarchical approach where the optimal amino acid is not estimated on a per site basis, but rather as a vector of probability an amino acid is optimal at the gene level.

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Unlike ω based approaches, SelAC makes inferences about things other than branch

length and tree, and we can validate the assumptions indirectly by comparing our inferences to other empirical data, such as we do with protein synthesis data. More 158 generally, SelAC's assumptions lead to mechanistic and, thus, testable hypothesis about 159 the relationship between mutation, protein function, gene expression, and rates of 160 evolution. More importantly, alternative hypotheses could be used in place of ours and, in 161 turn, phylogenetic and other types of data could be used to evaluate the support of these 162 alternative models. Our hope is that by moving away from the more phenomenological 163 models we can better connect population genetics, molecular biology, and phylogenetics 164 allowing each area inform the others more effectively. 165

MATERIALS & METHODS

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We model the substitution process as a classic Wright-Fisher process which includes the forces of mutation, selection, and drift (Fisher 1930; Kimura 1962; Wright 1969; Iwasa 1988; Berg and Lässig 2003; Sella and Hirsh 2005; McCandlish and Stoltzfus 2014). For simplicity, we ignore linkage effects and, as a result of this and other assumptions, our method behaves in a site independent manner.

Because SelAC requires twenty families of 61×61 matrices, the number of parameters needed to implement SelAC would, without further assumptions, be extremely large (i.e. on the order of 74,420 parameters). To reduce the number of parameters needed, while still maintaining a high degree of biological realism, we construct our gene and amino acid specific substitution matrices using a submodel nested within our substitution model, similar to approaches in Gilchrist (2007); Shah and Gilchrist (2011); Gilchrist et al. (2015).

One advantage of a nested modeling framework is that it requires only a handful of genome-wide parameters such as nucleotide specific mutation rates (scaled by effective population size N_e), side chain physicochemical weighting parameters, and a shape

parameter describing the distribution of site sensitivities. In addition to these genome-wide parameters, SelAC requires a gene g specific expression parameter ψ_g which describes the 182 average rate at which the protein's functionality is produced by the organism or a gene's 183 'average functionality production rate' for short. (For notational simplicity, we will ignore 184 the gene specific indicator g, unless explicitly needed.) Currently, ψ is fixed across the 185 phylogeny, though relaxing this assumption is a goal of future work. The gene specific 186 parameter ψ is multiplied by additional model terms to make a composite term ψ' which 187 scales the strength and efficacy of selection for the optimal amino acid sequence relative to 188 drift (see Implementation below). In terms of the functionality of the protein encoded, we 189 assume that for any given gene there exists an optimal amino acid sequence \vec{a}^* and that, by 190 definition, is a complete, error free peptide consisting of \vec{a}^* and provides one unit of the 191 gene's functionality. We also assume that natural selection favors genotypes that are able 192 to synthesize their proteome more efficiently than their competitors and that each savings 193 of an high energy phosphate bond per unit time leads to a constant proportional gain in 194 fitness A_0 . SelAC also requires the specification (as part of parameter optimization) of an 195 optimal amino acid at each position or site within a coding sequence which, in turn, makes 196 it the largest category of parameters we estimate. Because we use a submodel to derive our 197 substitution matrices, SelAC requires the estimation of a fraction of the parameters 198 required when compared to approaches where the substitution rates are allowed to vary 199 independently (Halpern and Bruno 1998; Lartillot and Philippe 2004; Rodrigue and 200 Lartillot 2014). 201

As with other phylogenetic methods, we generate estimates of branch lengths and nucleotide specific mutation rates. In addition, because the math behind our model is mechanistically derived, our method can also be used to make quantitative inferences on the optimal amino acid sequence of a given protein as well as the realized average synthesis rate of each protein used in the analysis. The mechanistic basis of SelAC also means it can

be easily extended to include more biological realism and test more explicit hypotheses about sequence evolution.

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Mutation Rate Matrix μ

We begin with a 4x4 nucleotide mutation matrix that defines a model for mutation rates 210 between individual bases. For our purposes, we rely on the general unrestricted 211 model(Yang 1994, UNREST) because it makes no constraint on the instantaneous rate of 212 change between any pair of nucleotides. In our view, the flexibility and potential for strong 213 asymmetries in the transition among the different nucleotide states, and ultimately among 214 the different codon states, is more consistent with our model. We note, however, that more 215 constrained models, such as the Jukes-Cantor (JC), Hasegawa-Kishino-Yano (HKY), or the 216 general time-reversible model (GTR), can also be used. The 12 parameter UNREST model 217 defines the relative rates of change between a pair of nucleotides. Thus, we arbitrarily set 218 the G→T mutation rate to 1, resulting in 11 free mutation rate parameters in the 4x4 219 mutation nucleotide mutation matrix. The nucleotide mutation matrix is also scaled by a 220 diagonal matrix π whose entries, $\pi_{i,i} = \pi_i$, correspond to the equilibrium frequencies of 221 each base. These equilibrium nucleotide frequencies are determined by analytically solving 222 $\pi \times \mathbf{Q} = 0$. We use this **Q** to populate a 61×61 codon mutation matrix μ , whose entries $\mu_{i,j}$ describe the mutation rate from codon i to j under a "weak mutation" assumption, 224 such that evolution is mutation limited, codon substitutions only occur one nucleotide at a 225 time and, as a result, the rate of change between any pair of codons that differ by more 226 than one nucleotide is zero. 227 While the overall model does not assume equilibrium, we still need to scale our 228 mutation matrices μ by a scaling factor S. As traditionally done, we rescale our time units 229 such that at equilibrium, one unit of branch length represents one expected mutation per 230

site (which equals the substitution rate under neutrality, but would not with selection).

More explicitly, $S = \left(\sum_{i \in \text{codons}} \mu_i \pi_i\right)$ where the final mutation rate matrix is the original mutation rate matrix multiplied by 1/S.

Protein Synthesis Cost-Benefit Function η

SelAC links fitness to the product of the cost-benefit function of a gene η and the organism's average target synthesis rate of the functionality provided by gene ψ .

Compensatory changes that allow an organism to maintain functionality even with loss of one or both copies of a gene are widespread (reviewed in 2); here we assume that for finer scale problems than entire loss (for example, a 10% loss of functionality) the compensation is more production of the protein. This is because the average flux energy an organism spends to meet its target functionality provided by the gene is $\eta \times \psi$. In order to link genotype to our cost-benefit function $\eta = \mathbf{C}/\mathbf{B}$, we begin by defining our benefit function \mathbf{B} .

Benefit.— Our benefit function **B** measures the functionality of the amino acid sequence \vec{a}_i encoded by a set of codons \vec{c}_i , i.e. $a(\vec{c}_i) = \vec{a}_i$ relative to that of an optimal sequence \vec{a}^* . By definition, $\mathbf{B}(\vec{a}^*) = 1$ and $\mathbf{B}(\vec{a}_i|\vec{a}^*) < 1$ for all other sequences. We assume all amino acids within the sequence contribute to protein function and that this contribution declines as an inverse function of physicochemical distance between each amino acid and the optimal.

Formally, we assume that

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$$\mathbf{B}(\vec{a}_i|\vec{a}^*) = \left(\frac{1}{n}\sum_{p=1}^n \left(1 + G_p d(a_{i,p}, a_p^*)\right)\right)^{-1}$$
 (1)

where n is the length of the protein, $d(a_{i,p}, a_p^*)$ is a weighted physicochemical distance

²From Cruft: There is evidence of compensation for protein function. Metabolism with gene expression models (ME-models) link those factors to successfully make predictions about response to perturbations in a cell (?), (?). For example, an ME-model for *E. coli* successfully predicted gene expression levels in vivo (?).

between the amino acid encoded in gene i for position p and a_p^* is the optimal amino acid for that position of the protein. For simplicity, we define the distance between a stop codon 252 and a sense codon as effectively infinite and, as a result, nonsense mutations are effectively 253 lethal. The term G_p describes the sensitivity of the protein's function to physicochemical 254 deviation from the optimimum at site position p. There are many possible measures for 255 physiochemical distance; we use Grantham (1974) distances by default, though others may 256 be chosen. We assume that $G_p \sim \text{Gamma}(\alpha = \alpha_G, \beta = \alpha_G)$ in order to ensure $\mathbb{E}(G_p) = 1$. 257 Given the definition of the Gamma distribution, the variance in G_p is equal to 258 $\alpha/\beta^2 = 1/\alpha_G$. Further, at the limit of $\alpha_G \to \infty$, the model becomes equivalent to assuming 259 uniform site sensitivity where $G_p = 1$ for all positions p. Finally, we note that $\mathbf{B}(\vec{a}_i|\vec{a}^*)$ is 260 inversely proportional to the average physicochemical deviation of an amino acid sequence 261 \vec{a}_i from the optimal sequence \vec{a}^* weighted by each site's sensitivity to this deviation. $\mathbf{B}(\vec{a}_i|\vec{a}^*)$ can be generalized to include second and higher order terms of the distance 263 measure d. 264

Cost.— Protein synthesis involves both direct and indirect assembly costs. Direct costs consist of the high energy phosphate bonds $\sim P$ of ATP or GTP's used to assemble the ribosome on the mRNA, charge tRNA's for elongation, move the ribosome forward along the transcript, and terminate protein synthesis. As a result, direct protein assembly costs are the same for all proteins of the same length. Indirect costs of protein assembly are potentially numerous and could include the cost of amino acid synthesis as well the cost and efficiency with which the protein assembly infrastructure such as ribosomes, aminoacyl-tRNA synthetases, tRNAs, and mRNAs are used. When these indirect costs are combined with sequence specific benefits, the probability of a mutant allele fixing is no longer independent of the rest of the sequence (Gilchrist et al. 2015) and, as a result, model fitting becomes substantially more complex. Thus for simplicity, in this study we ignore

indirect costs of protein assembly that vary between genotypes and define,

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$$\mathbf{C}(\vec{c_i}) = \text{Energetic cost of protein synthesis.}$$
 (2)

$$= A_1 + A_2 n \tag{3}$$

where, A_1 and A_2 represent the direct cost, in high energy phosphate bonds, of ribosome initiation and peptide elongation, respectively, where $A_1 = A_2 = 4 \sim P$.

Defining Physicochemical Distances

Assuming that functionality declines with an amino acid a_i 's physicochemical distance from the optimum amino acid a^* at each site provides a biologically defensible way of mapping genotype to protein function that requires relatively few free parameters. In addition, SelAC naturally lends itself to model selection since we can compare the quality of SelAC fits using different mixtures of physicochemical properties. Following Grantham (1974), we focus on using composition c, polarity p, and molecular volume v of each amino acid's side chain residue to define our distance function, but the model and its implementation can flexibly handle a variety of properties. We use the Euclidian distance between residue properties where each property c, p, and v has its own weighting term, α_c , α_p , α_v , respectively, which we refer to as 'Grantham weights'. Because physicochemical distance is ultimately weighted by a gene's specific average protein synthesis rate ψ , another parameter we estimate, there is a problem with parameter identifiablity. Ultimately, the scale of gene expression is affected by how we measure physicochemical distances which, in turn, is determined by our choice of Grantham weights. As a result, by default we set $\alpha_v = 3.990 \times 10^{-4}$, the value originally estimated by Grantham, and recognize that our

estimates of α_c and α_p and ψ are scaled relative to this choice for α_v . More specifically,

$$d(a_i, a^*) = (\alpha_c [c(a_i) - c(a^*)]^2 + \alpha_p [p(a_i) - p(a^*)]^2 + \alpha_v [v(a_i) - v(a^*)]^2)^{1/2}.$$

Linking Protein Synthesis to Allele Substitution

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Next we link the protein synthesis cost-benefit function η of an allele with its fixation 269 probability. First, we assume that each protein encoded within a genome provides some 270 beneficial function and that the organism needs that functionality to be produced at a 271 target average rate ψ . By definition, the optimal amino acid sequence for a given gene, \vec{a}^* , 272 produces one unit of functionality. Second, we assume that protein expression is regulated 273 by the organism to ensure that functionality is produced at rate ψ . As a result, the average 274 protein synthesis rate of a gene, ϕ , by definition, satisfies the equality $\phi = \psi/\mathbf{B}(\vec{a})$. In 275 other words, the average production rate of a protein \vec{a} with relative functionality $\mathbf{B}(\vec{a}) < 1$ 276 must be $1/\mathbf{B}(\vec{a})$ times higher than the production rate needed if the optimal amino acid 277 sequence \vec{a}^* was encoded since, by definition, $\mathbf{B}(\vec{a}^*) = 1$. 278 For example, a cell with an allele \vec{a} where $\mathbf{B}(\vec{a}) = 0.9$ would have to produce the 279 protein at rate $\phi = 10/9 \times \psi$. In contrast, a cell with the optimal allele \vec{a}^* would have to 280 produce the protein at rate $\phi = \psi$. Similarly, a cell with an allele \vec{a} where $\mathbf{B}(\vec{a}) = 1/2$ will 281 have to produce the protein at $\phi = 2\psi$. Simply put, the fitness cost for a genotype 282

Third, we assume that every additional high energy phosphate bond, $\sim P$, spent per unit time to meet the organism's target function synthesis rate ψ leads to a slight and

encoding a suboptimal protein sequence stems from the need to produce suboptimal

proteins at a higher rate in order to compensate for their lower functionality.

proportional decrease in fitness W. This assumption, in turn, implies

$$W_i(\vec{c}) \propto \exp\left[-A_0 \,\eta(\vec{c}_i)\psi\right]. \tag{4}$$

where A_0 , again, describes the decline in fitness with every $\sim P$ wasted per unit time.

Because A_0 shares the same time units as ψ and ϕ and only occurs in SelAC in conjunction with ψ , we do not need to explicitly identify our time units.

Correspondingly, the ratio of fitness between two genotypes is,

$$W_i/W_j = \exp\left[-A_0 \,\eta(\vec{c}_i)\psi\right] / \exp\left[-A_0 \,\eta(\vec{c}_j)\psi\right] \tag{5}$$

$$= \exp\left[-A_0 \left(\eta(\vec{c}_i) - \eta(\vec{c}_i)\right)\psi\right] \tag{6}$$

(7)

Given our formulations of \mathbf{C} and \mathbf{B} , the fitness effects between sites are multiplicative and, therefore, the substitution of an amino acid at one site can be modeled independently of the amino acids at the other sites within the coding sequence. As a result, the fitness ratio for two genotypes differing at a single site p simplifies to

$$\frac{W_i}{W_j} = \exp\left[-\frac{A_0 \left(A_1 + A_2 n_g\right)}{n_g} \times \sum_{p \in \mathbb{P}} \left[d\left(a_{i,p}, a_p^*\right) - d\left(a_{j,p}, a_p^*\right)\right] G_p \psi\right]$$

where \mathbb{P} represents the codon positions in which $\vec{c_i}$ and $\vec{c_j}$ differ. Fourth, we make a weak mutation assumption, such that alleles can differ at only one position at any given time, i.e. $|\mathbb{P}| = 1$, and that the population is evolving according to a Fisher-Wright process. As a result, the probability a new mutant, j, introduced via mutation into a resident population

i with effective size N_e will go to fixation is,

$$u_{i,j} = \frac{1 - (W_i/W_j)^b}{1 - (W_i/W_j)^{2N_e}}$$

$$= \frac{1 - \exp\left\{-\frac{A_0}{n_g} (A_1 + A_2 n_g) \left[d(a_i, a^*) - d(a_j, a^*)\right] G_p \psi b\right\}}{1 - \exp\left\{-\frac{A_0}{n_g} (A_1 + A_2 n_g) \left[d(a_i, a^*) - d(a_j, a^*)\right] G_p \psi 2N_e\right\}}$$

where b = 1 for a diploid population and 2 for a haploid population (Kimura 1962; Wright 1969; Iwasa 1988; Berg and Lässig 2003; Sella and Hirsh 2005). Finally, assuming a constant mutation rate between alleles i and j, $\mu_{i,j}$, the substitution rate from allele i to j can be modeled as,

$$q_{i,j} = \frac{2}{b}\mu_{i,j}N_e u_{i,j}.$$

where, given the substitution model's weak mutation assumption, $N_e\mu \ll 1$. In the end, 288 each optimal amino acid has a separate 64 x 64 substitution rate matrix \mathbf{Q}_a , which 289 incorporates selection for the amino acid (and the fixation rate matrix this creates) as well 290 as the common mutation parameters across optimal amino acids. This results in the 291 creation of 20 Q matrices, one for each amino acid and each with 3,721 entries which are 292 based on a relatively small number of model parameters (one to 11 mutation rates, two free 293 Grantham weights, the cost of protein assembly, A_1 and A_2 , the gene specific target 294 functionality synthesis rate ψ , and optimal amino acid at each position p, a_p^*). These model 295 parameters can either be specified a priori or estimated from the data. 296 Given our assumption of independent evolution among sites, it follows that the 297 probability of the whole data set is the product of the probabilities of observing the data at 298 each individual site. Thus, the likelihood \mathcal{L} of amino acid a being optimal at a given site 299

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$$\mathcal{L}\left(\mathbf{Q}_{a}|\mathbf{D}_{p},\mathbf{T}\right) \propto \mathbf{P}\left(\mathbf{D}_{p}|\mathbf{Q}_{a},\mathbf{T}\right) \tag{8}$$

In this case, the data, \mathbf{D}_p , are the observed codon states at position p for the tips of the 301 phylogenetic tree with topology T. For our purposes we take T as given but it could be 302 estimated as well. The pruning algorithm of Felsenstein (1981) is used to calculate 303 $\mathcal{L}(\mathbf{Q}_a|\mathbf{D}_p,\mathbf{T})$. The log of the likelihood is maximized by estimating the genome scale 304 parameters which consist of 11 mutation parameters which are implicitly scaled by $2N_e/b$, 305 and two Grantham distance parameters, α_c and α_p , and the sensitivity distribution 306 parameter α_G . Because A_0 and ψ_g always co-occur and are scaled by N_e , for each gene g307 we estimate a composite term $\psi'_g = \psi_g A_0 b N_e$ and the optimal amino acid for each position 308 a_p^* of the protein. When estimating α_G , the likelihood then becomes the average likelihood which we calculate using the generalized Laguerre quadrature with k=4 points 310 (Felsenstein 2001). 311 Finally, we note that because we infer the ancestral state of the system, our 312 approach does not rely on any assumptions of model stationarity. Nevertheless, as our 313 branch lengths grow the probability of observing a particular amino acid a at a given site 314 approaches a stationary value proportional to $W(a)^{2N_e-b}$ (Sella and Hirsh 2005).

Implementation

All methods described above are implemented in the new R package, selac available 317 through GitHub (https://github.com/bomeara/selac) [it will be uploaded to CRAN 318 once peer review has completed. Our package requires as input a set of fasta files that each 319 contain an alignment of coding sequence for a set of taxa, and the phylogeny depicting the 320 hypothesized relationships among them. In addition to the SelAC models, we implemented 321 the GY94 codon model of Goldman and Yang (1994), the FMutSel0 mutation-selection 322

model of Yang and Nielsen (2008), and the standard general time-reversible nucleotide model that allows for Γ distributed rates across sites. These likelihood-based models represent a sample of the types of popular models often fit to codon data.

For the SelAC models, the starting guess for the optimal amino acid at a site comes 326 from 'majority' rule, where the initial optimum is the most frequently observed amino acid 327 at a given site (ties resolved randomly). Our optimization routine utilizes a four stage hill 328 climbing approach. More specifically, within each stage a block of parameters are 329 optimized while the remaining parameters are held constant. The first stage optimizes the 330 block of branch length parameters. The second stage optimizes the block of gene specific 331 composite parameters $\psi'_g = A_0 \psi_g N_e$. The third stage optimizes the model SelAC's 332 parameters shared across the genome α_c and α_p , and the sensitivity distribution parameter 333 α_G . The fourth stage estimates the optimal amino acid at each site a^* . This entire four stage cycle is repeated six more times. For optimization of a given set of parameters, we 335 rely on a bounded subplex routine (Rowan 1990) in the package NLopt (Johnson 2012) to 336 maximize the log-likelihood function. To help the optimization navigate through local 337 peaks, we perform a set of independent analyses with different sets of naive starting points 338 with respect to the gene specific composite ψ' parameters, α_c , and α_p . Confidence in the 330 parameter estimates can be generated by an 'adaptive search' procedure that we 340 implemented to provide an estimate of the parameter space that is some pre-defined 341 likelihood distance (e.g., 2 lnL units) from the maximum likelihood estimate (MLE), which 342 follows Beaulieu and OMeara (2016); Edwards (1984). 343

We note that our current implementation of SelAC is painfully slow, and is best suited for data sets with relatively few number of taxa (i.e. < 10). This limitation is largely due to the size and quantity of matrices we create and manipulate to calculate the log-likelihood of an individual site. Ongoing work will address the need for speed, with the eventual goal of implementing the model-SelAC in popular phylogenetic inference toolkits,

Simulations

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We evaluated the performance of our codon model by simulating datasets and estimating 351 the bias of the inferred model parameters from these data. Our 'known' parameters under 352 a given generating model were based on fitting SelAC to the 106 gene data set and 353 phylogeny of Rokas et al. (2003). The tree used in these analyses is outdated with respect 354 to the current hypothesis of relationships within Saccharomyces, but we rely on it simply as 355 a training set that is separate from our empirical analyses (see section Analysis of yeast genomes below). Bias in the model parameters were assessed under two generating models: 357 one where we assumed a model of SelAC assuming uniform sensitivity across sites 358 (i.e. $G_p = 1$ for all sites, i.e. $\alpha_G = \infty$), and one where we estimated the Gamma 359 distribution parameter α_G from the data. Under each of these two scenarios, we used 360 parameter estimates from the corresponding empirical analysis and simulated 50 five-gene 361 data sets. For the gene specific composite parameter ψ_g' the 'known' values used for the 362 simulation were five evenly spaced points along the rank order of the estimates across the 363 106 genes. The MLE estimate for a given replicate were taken as the fit with the highest 364 log-likelihood after running five independent analyses with different sets of naive starting 365 points with respect to the composite ψ'_q parameter, α_c , and α_p . All analyses were carried out in our selac R package.

Analysis of yeast genomes and tests of model adequacy

We focus our empirical analyses on the large yeast data set and phylogeny of Salichos and Rokas (2013). The yeast genome is an ideal system to examine our phylogenetic estimates of gene expression and its connection to real world measurements of these data within

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individual taxa. The complete data set of Salichos and Rokas (2013) contain 1070
   orthologs, where we selected 100 at random for our analyses. We also focus our analyses
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   only on Saccharomyces sensu stricto, including their sister taxon Candida qlabrata, and we
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    rely on the phylogeny depicted in Fig. 1 of Salichos and Rokas (2013) for our fixed tree.
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    We fit the two SelAC models described above (i.e., SelAC and SelAC+\Gamma), as well as two
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    codon models, GY94 and FMutSel0, and a standard GTR + \Gamma nucleotide model. The
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   FMutSel0 model, which assumes that the amino acid frequencies are determined by
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   functional requirements of the protein. In all cases, we assumed that the model was
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    partitioned by gene, but with branch lengths linked across genes.
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           For SelAC, we compared our estimates of \phi' = \psi'/\mathbf{B}, which represents the average
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   protein synthesis rate of a gene, to estimates of gene expression from empirical data.
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   Specifically, we obtained gene expression data for five of the six species used - four species
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    were measured during log-growth phase, whereas the other was measured at the beginning
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   of the stationary phase (S. kudriavzevii) from the Gene Expression Omnibus (GEO). Gene
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   expression in this context corresponds to mRNA abundances which were measured using
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    either microarrays (C. glabrata, S. castellii, and S. kudriavzevii) or RNA-Seq (S. paradoxus,
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    S. mikatae, and S. cerevisiae).
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           For further comparison, we also predicted the average protein synthesis rate for each
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   gene \phi by analyzing gene and genome-wide patterns of synonymous codon usage using
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   ROC-SEMPPR (Gilchrist et al. 2015) for each individual genome. While, like SelAC,
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    ROC-SEMPPR uses codon level information, it does not rely on any inter-specific
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   comparisons and, unlike SelAC, uses only the intra- and inter-genic frequencies of
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   synonymous codon usage as its data. Nevertheless, ROC-SEMPPR predictions of gene
   expression \phi correlates strongly (r = 0.53 - 0.74) with a wide range of laboratory
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   measurements of gene expression (Gilchrist et al. 2015).
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While one of our main objectives was to determine the improvement of fit that

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SelAC has with respect to other standard phylogenetic models, we also evaluated the adequacy of SelAC. Model fit, measured with assessments such as the Akaike Information 399 Criterion (AIC), can tell which model is least bad as an approximation for the data, but it 400 does not reveal whether a model is actually doing a good job of representing the biological 401 processes. An adequate model does the latter, one measure of which is that data generated 402 under the model resemble real data (Goldman 1993). For example, Beaulieu et al. (2013) 403 assessed whether parsimony scores and the size of monomorphic clades of empirical data 404 were within the distributions of simulated under a new model and the best standard 405 model; if the empirical summaries were outside the range for each, it would have suggested 406 that neither model was adequately modeling this part of the biology. 407

For a given gene we first remove a particular taxon from the data set and the 408 phylogeny. A marginal reconstruction of the likeliest sequence across all remaining nodes is conducted under the model, including the node where the pruned taxon attached to the 410 tree. The marginal probabilities of each site are used to sample and assemble the starting 411 coding sequence. This sequence is then evolved along the branch, periodically being 412 sampled and its current functionality assessed. We repeat this process 100 times and 413 compare the distribution of trajectories against the observed functionality calculated for 414 the gene. For comparison, we also conducted the same test, by simulating the sequence 415 under the standard GTR + Γ nucleotide model, which is often used on these data but does 416 not account for the fact that the sequence codes for a specific protein, and under FMutSel0, 417 which includes selection on codons but in a fundamentally different way as our model. 418

The appropriate estimator of bias for AIC

As part of the model set described above, we also included a reduced form of each of the two SelAC models, SelAC and SelAC+ Γ . Specifically, rather than optimizing the amino acid at any given site, we assume the the most frequently observed amino acid at each site

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is the optimal amino acid a^* . We refer to these 'majority rule' models as SelAC_M and $SelAC_M + \Gamma$ and the majority rule parameterization greatly accelerates model fitting. 424 Since these majority rule models assume that the optimal amino acids are known 425 prior to fitting of our model, it is tempting to reduce the number of parameters in the 426 model by the number of total sites being analyzed. Despite having become standard 427 behavior in the field of phylogenetics, this reduction is statistically inappropriate due to the 428 fact that identification of the majority rule amino acid is made by examining the same data 429 as we fit to our model. Because the difference in K when counting or not counting the 430 number of nucleotide sites drops out when comparing nucleotide models with AIC, this 431 statistical issue does not apply to nucleotide models. It does, however, matter for AICc, 432 where the number of parameters, K, and the sample size, n, combine in the penalty term. 433 This also matters in our case, where the number of estimated parameters for the majority 434 rule estimation differs based on whether one is looking at codons or single nucleotides. 435 In phylogenetics two variants of AICc are used. In comparative methods 436 (e.g. Butler and King 2004; O'Meara et al. 2006; Beaulieu et al. 2013) the number of data 437 points, n, is taken as the number of taxa. More taxa allow the fitting of more complex 438 models, given more data. However, in DNA evolution, which is effectively the same as a 430 discrete character model used in comparative methods, the n is taken as the number of 440 sites. Obviously, both cannot be correct. 441 The original derivation of AICc by Hurvich and Tsai (1989) assumed a regression 442 model, where the true model was in the set of examined models, as well as approximations in the derivation itself. The appropriatness of this approximation for phylogenetic data, where data points independence between taxa, is unclear. In any case, we argue that for

phylogenetic data, a good estimate of data set size is number of taxa multiplied by number

of sites. First of all, this is what is conventionally seen as the size of the dataset in the field.

Second, when considering how likelihood is calculated, the likelihood for a given site is the

sum of the probabilities of each observed state at each tip, and this is then multiplied across sites. It is arguable that the conventional approach in comparative methods is calculating 450 AICc in this way: number of taxa multiplied by number of sites equals the number of taxa, 451 if only one site is examined, as remains remarkably common in comparative methods. (One 452 notable exception to this appoach to calculating AICc is the program SURFACE 453 implemented by Ingram and Mahler (2013), which uses multiple characters and taxa. 454 While its default is to use AIC to compare models, if one chooses to use AICc, the number 455 of samples is taken as the product of number of sites and number of taxa.) 456 Recently, Jhwueng et al. (2014) performed an analysis that investigated what 457 variant of AIC and AICc worked best as an estimator, but the results were inconclusive. 458 Here, we have adopted and extended the simulation approach of Jhwueng et al. (2014) in 459 order to examine a large set of different penalty functions and how well they approximate the remaining portion of the Kullback-Liebler (KL) divergence between two models after 461 accounting for the deviance (i.e., $-2\mathcal{L}$) (see Appendix 1 for more details). 462

RESULTS

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By linking transition rates $q_{i,j}$ to gene expression ψ , our approach allows use of the same model for genes under varying degrees of stabilizing selection. Specifically, we assume the strength of stabilizing selection for the optimal sequence, \vec{a}^* , is proportional to the average protein synthesis rate ϕ , which we can estimate for each gene. In regards to model fit, our results clearly indicated that linking the strength of stabilizing selection for the optimal sequence to gene expression substantially improves our model fit. Further, including the shape parameter α_G for the random effects term $G \sim \text{Gamma}(\alpha_G, \beta_g)$ to allow for heterogeneity in this selection between sites within a gene improves the ΔAICc of SelAC+ Γ over the simpler SelAC models by over 22,000 AIC units. Using either ΔAICc or AIC_w as our measure of model support, the SelAC models fit extraordinarily better than
GTR + Γ , GY94, or FMutSel0 (Table 1). This is in spite of the need for estimating the
optimal amino acid at each position in each protein, which accounts for 49,881 additional
model parameters. Even when compared to the next most parameter rich codon model in
our model set, FMutSel0, SelAC+ Γ model shows nearly 180,000 AIC unit improvement
over FMutSel0.

With respect to estimates of ϕ within SelAC, they were strongly correlated with 470 both our empirical (i.e. mRNA abundances) and model based (i.e. ROC-SEMPPR) 480 measurements of gene expression (Figure 1 and Figures S1-S2, respectively). In other 481 words, using only codon sequences, our model can predict which genes have high or low 482 expression levels. The estimate of the α_G parameter, which describes the site-specific 483 variation in sensitivity of the protein's functionality, indicated a moderate level of variation in gene expression among sites. Our estimate of $\alpha_G = 1.40$, produced a distribution of 485 sensitivity terms G ranged from 0.344-7.16, but with nearly 90% of the weight for a given 486 site-likelihood being contributed by the 0.344 and 1.48 rate categories. In simulation, 487 however, of all the parameters in the model, only α_G showed a consistent bias, in that the 488 MLE were generally lower than their actual values (see Supporting Materials). Other 480 parameters in the model, such as the Grantham weights, provide an indication as to the 490 physicochemical distance between amino acids. Our estimates of these weights only 491 strongly deviate from Grantham's 1974 original estimates in regards to composition weight, 492 α_c , which is the ratio of noncarbon elements in the end groups to the number of side 493 chains. Our estimate of the composition weighting factor of α_c =0.484 is 1/4th the value estimate by Grantham which suggests that the substitution process is less sensitive to this physicochemical property when shared ancestry and variation in stabilizing selection are taken into account.

It is important to note that the nonsynonymous/synonymous mutation ratio, or ω ,

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which we estimated for each gene under the FMutSel0 model strongly correlated with our estimates of $\phi' = \psi'/\mathbf{B}$ where \mathbf{B} depends on the sequence of each taxa. In fact, ω showed 500 similar, though slightly reduced correlations, with the same empirical estimates of gene 501 expression described above (Figure 2) This would give the impression that the same 502 conclusions could have been gleaned using a much simpler model, both in terms of the 503 number of parameters and the assumptions made. However, as we discussed earlier, not 504 only is this model greatly restricted in terms of its biological feasibility, SelAC clearly 505 performs better in terms of its fit to the data and biological realism. 506

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For example, when we simulated the sequence for S. cervisieae, starting from the ancestral sequence under both $GTR + \Gamma$ and FMutSel0, the functionality of the simulated sequence moves away from the observed sequence, whereas SelAC remains near the functionality of the observed sequence (Figure 3b). In a way, this is somewhat unsurprising, given that both GTR + Γ and FMutSel0 are agnostic to the functionality of the gene, but it does highlight the improvement in biological realism in amino acid sequence evolution that SelAC provides. We do note that the adequacy of the SelAC model does vary among individual taxa, and does not always match the observed functionality. For instance, S. castellii is simulated with consistently higher functionality than observed (Figure 3c). 515 We suspect this is an indication that assuming a single set of optimal amino acid across all 516 taxa may be too simplistic, but we cannot also rule out other potential simplifying assumptions in our model, such as a single set of Grantham weights and α_G values or the 518 simple, inverse relationship between physicochemical distance d and benefit **B**. 519

Finally, we note that our simulation analysis suggested that the best measure of dataset size for AICc uses a scaled value of the product of number of sites and number of characters was the best at estimating KL distance. The model comparison approach described above included this assumption. For more details on the simulation approach, see Appendix 1.

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One goal of A central goal in evolutionary biology is to quantify the nature, strength, and,
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    ultimately, shifts in the forces of natural selection relative to genetic drift and mutation.
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    As data set size and complexity increase, so does the amount of potential information on
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    these forces and their dynamics. As a result, there is a need for more complex and realistic
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   models (Goldman et al. 1996; Thorne et al. 1996; Goldman et al. 1998; Halpern and Bruno
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    1998; Lartillot and Philippe 2004) to accomplish this. Although extremely popular due to
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    their elegance and computational efficiency, the utility of \omega based models in helping us
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   reach this goal is substantially more limited than commonly recognized. Because they use
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    only a single substitution matrix, they are only applicable for situations in which the
   substitution process and shifts in the selective environment are intrinsicly, such as with
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    positive or negative frequency dependent selection; these models do not describe stabilizing
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    or diversifying selection as commonly envisioned (Oxford Encyclopedia Entry goes here).
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           Starting with Halpern and Bruno (1998), a number of researchers have developed
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   methods for linking site-specific selection on protein sequence and phylogenetics (e.g. Koshi
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   et al. 1999; Dimmic et al. 2000; Koshi and Goldstein 2000; Robinson et al. 2003; Lartillot
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   and Philippe 2004; Thorne et al. 2012; Rodrigue and Lartillot 2014) . Our work here
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    Halpern and Bruno (1998) calculated a vector of 19 expected amino acid frequencies for
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   each amino acid site, making it the most general and most parameter rich of these
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   methods. This generality, however, comes at the cost of being purely descriptive; there is
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   no explicit biological mechanism is proposed to explain the site specific amino frequencies
   estimated. By grouping together amino sites with similar evolutionary behaviors, Lartillot
   and colleagues retained the descriptive nature of Halpern and Bruno (1998) work while
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    greatly reduced the number of model parameters needed
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    (Lartillot and Philippe 2004; Rodrigue and Lartillot 2014). SelAC follows in this tradition
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of using multiple substitution matrices, but includes some key advances.

First, by nesting a model of a sequence's cost-benefit function C/B within a 551 broader model, SelAC allows us to formulate and test more specific biological hypotheses. 552 For example, using a hierarchical approach we were able a hierarchical, mechanistic models 553 of stabilizing selection. More precisely, our nested approach allows us to relax the 554 assumption that physico-chemical deviations from the optimal sequence \vec{a}^* are equally 555 disruptive at all sites within a protein. We found strong support for the SelAC's hypothesis 556 that the strength of stabilizing selection against physico-chemical deviations from \vec{a}^* varies 557 between sites (\triangle AICc XXXXXX). Second, because our substitution matrices are built on a 558 formal description of a sequence's cost-benefit function C/B, we are able to efficiently 559 parameterize 20 different matrices using a relatively small number of genome-wide 560 parameters – e.g. our physico-chemical weighting and G distribution shape parameters, and one gene specific gene expression parameter ψ . While the model we use to link amino acid 562 sequence and protein function is simplistic C/B function on which SelAC currently rests is 563 very simple, nevertheless, it leads to a substantial dramatic increase in our ability to 564 explain the sequence data used to parameterize the model. Our model is built on we 565 analyzed. Importantly, because SelACuses a formal description of a sequence's C/B, which 566 make modifying our assumptions replacing our assumptions with more sophisticated ones 567 in the future is relatively straightforward. Conceptually, our work lies in between that of 568 Lartillot's and Thorne's, where the latter is utilizing even more detailed models of protein 569 structure as a means of linking amino acid substitutions and stabilizing selection. Third, 570 our use of nested models also allows us to make biologically meaningful and testable 571 predictions. By linking a gene's expression level to the strength of purifying selection it is experience, we are able to provide coarse estimates of gene expression. When we apply this 573 insight to the standard models, we found that ω is best explained as a proxy for gene 574 expression, rather than the nature of selection on a sequence.

One simplifying assumption we make is that the organism can and does compensate for any reduction in protein function by simply increasing the protein's production rate.

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While this production compensation assumption will clearly not hold in many 578 situations, it does allow us to connect protein function and energetic costs in a simple and 579 biologically plausible manner. Of course, researchers could employ and test other 580 assumptions within our framework, namely, by utilizing more detailed, gene specific 581 knowledge about the relationship between protein function and organism fitness. For 582 example, suppose a protein for a glucose transporter is far less efficient than usual. One 583 organismal response, the one envisioned here, is that the protein is thus produced far more 584 to compensate. This would leave the overall ability to transport glucose unchanged. An 585 alternative is that the cell is just less able to transport glucose across membranes. In 586 biology, it is likely a mixture of such effects exists. However, the production compensation mechanism is likely to have the same costs across proteins, making it a useful first 588 approximation to model, while the same expression but reduced functionality will have 589 gene specific effects more difficult to model generally (e.g., how does the cost of having 590 glucose transport slow by half compare to the cost of underproducing an anthocyanin for 591 flower color or fewer taste receptor proteins?). Moreover, there is evidence that cells do 592 compensate for lower protein function by increasing gene expression (?, MANY GOOD 593 CITATIONS) Nevertheless, by assuming that fitness declines with extraneous energy flux, 594 SelAC explicitly links the variation in the strength of stabilizing selection for the optimal 595 protein sequence among genes, to the variation among genes in their target expression levels ψ . 597

Furthermore, by linking expression and selection, SelAC provides a natural framework for combining information from protein coding genes with very different rates of evolution with the low expression genes providing information on shallow branches and the high expression genes providing information on deep branches. This is in contrast to a

more traditional approach of concatenating gene sequences together, which is equivalent to assuming the same average protein synthesis rate ψ for all of the genes, or more recent 603 approaches where different models are fitted to different genes. Our results indicate that 604 including a gene specific ψ value vastly improves SelAC fits (Table 1). Perhaps more 605 convincingly, we find that the target expression level ψ and realized average protein 606 synthesis rate ϕ are reasonably well correlated with laboratory measurements of gene 607 expression (r = 0.34 - 0.65; Figures 1, S1, and S2). The idea that quantitative information 608 on gene expression is embedded within intra-genomic patterns of synonymous codon usage 609 is well accepted; our work shows that this information can also be extracted from 610 comparative data at the amino acid level. 611

Of course, given the general nature of SelAC and the complexity of biological 612 systems, other biological forces besides selection for reducing energy flux likely contribute to intergenic variation in the magnitude of stabilizing selection. Similarly, other 614 physicochemical properties besides composition, volume, and charge likely contribute to 615 site specific patterns of amino acid substitution. Thus, a larger and more informative set of 616 Grantham weights might improve our model fit and reduce the noise in our estimates of ϕ . 617 Even if other physicochemical properties are considered, the idea of a consistent, genome 618 wide Grantham weighting of these terms seems highly unlikely. Since the importance of an 619 amino acid's physicochemical properties likely changes with where it lies in a folded 620 protein, one way to incorporate such effects is to test whether the data supports multiple 621 sets of Grantham weights for either subsets of genes or regions within genes, rather than a 622 single set. 623

Both of these points highlight the advantage of the detailed, mechanistic modeling approach underlying SelAC. Because there is a clear link between protein expression, synthesis cost, and functionality, SelAC can be extended by increasing the realism of the mapping between these terms and the coding sequences being analyzed. For example,

SelAC currently assumes the optimal amino acid for any site is fixed along all branches. This assumption can be relaxed by allowing the optimal amino acid to change during the 629 course of evolution along a branch. From a computational standpoint, the additive nature 630 of selection between sites is desirable because it allows us to analyze sites within a gene 631 largely independently of each other. From a biological standpoint, this additivity between 632 site ignores any non-linear interactions between sites, such as epistasis, or between alleles, 633 such as dominance. Thus, our work can be considered a first step to modeling to these 634 more complex scenarios. For example, our current implementation ignores any selection on 635 synonymous codon usage bias (CUB) (c.f. Yang and Nielsen 2008; Pouyet et al. 2016). 636 Including such selection is tricky because introducing the site-specific cost effects of CUB, 637 which is consistent with the hypothesis that codon usage affects the efficiency of protein 638 assembly or C, into a model where amino acids affect protein function or B, results in a cost-benefit ratio C/B with epistatic interactions between all sites. These epistatic effects can likely be ignored under certain conditions or reasonably approximated based on an 641 expectation of codon specific costs (e.g. Kubatko et al. 2016). Nevertheless, it is difficult 642 to see how one could identify such conditions without modeling the way in which codon 643 and amino acid usage affects C/B. 644

This work also points out the potential importance of further investigation into model choice in phylogenetics. For likelihood models, use of AICc has become standard. However, how one determines the appropriate number of parameters estimated in a model is more complicated than generally recognized. Common sense suggests that dataset size is increased by adding taxa and/or sites. In other words, a dataset of 1000 taxa and 100 sites must have more information on substitution models than a dataset of 4 taxa and 100 sites. Our simple analyses agree that the number of observations in a dataset (number of sites × number of taxa) should be taken as the sample size for AICc, but this conclusion likely only applies when there is sufficient independence between taxa. For instance, one could

imagine a phylogeny where one taxon is sister to a polytomy of 99 taxa that have zero
length terminal branches. Absent measurement error or other intraspecific variation, one
would have 100 species but only two unique trait values, and the only information about
the process of evolution comes from what happens on the path connecting the lone taxon
to the polytomy. Although this is a rather extreme example, it seems prudent for
researchers to use a simulation based approach similar to the one we take here to determine
the appropriate means for calculating the effective number of data points in their data.

There are still significant deficiencies in the approach outlined here. Most worrisome 661 are biological flaws in the model SelAC. For example, at its heart, the model SelAC assumes 662 that suboptimal proteins can be compensated for, at a cost, simply by producing more of 663 them. However, this is likely only true for proteins reasonably close to the optimal 664 sequence. Different enough proteins will fail to function entirely: the active site will not sufficiently match its substrates, a protein will not properly pass through a membrane, and 666 so forth. Yet, in our model, even random sequences still permit survival, just requiring 667 more protein production. Other oversimplifications include the assumption of no selection 668 on codon usage, no change of optimal amino acids through time, and no change of the 669 effect of physiochemical properties on fitness through time. However, because we take a 670 mechanistic approach, all of these assumptions can be relaxed through further extension of 671 our model. For example, it is possible to extend our model to include a hidden Markov 672 modelling (HMM) approach, which would allow for shifts in the optimal/pessimal amino 673 sequence along any given branch in the phylogeny. Extending our model in this way, should 674 eventually allow researchers to describe not only the dynamic shifts in natural selection 675 throughout the phylogeny, but evaluate how well a given dataset supports such a model. 676

There are also deficiencies in our implementation. Though reasonable to use for a given topology with a modest number of species, it is currently too slow for practical use for tree search. It thus serves as a proof of concept, or of utility for targeted questions

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where a more realistic model may be of use (placement of particular taxa, for example).
   Future work will encode SelAC models into a variety of mature, popular tree-search
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    programs. SelAC also represents a hard optimization problem: the nested models reduce
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    parameter complexity vastly, but there are still numerous parameters to optimize,
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   including the discrete parameter of the optimal amino acid at each site. A different
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    implementation, more parameter-rich, would optimize values of three (or more)
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    physiochemical properties per site. This would have the practical advantage of continuous
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    parameter optimization rather than discrete, and biologically would be more realistic (as it
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   is the properties that selection "sees", not the identity of the amino acid itself).
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           Overall In spite of these difficulties, SelAC
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   represents an important step in uniting phylogenetic and population genetic models. It While
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    Koshi et al. (1999); Dimmic et al. (2000); Koshi and Goldstein (2000); Robinson et al. (2003); Lartillot as
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   all models of constant, stabilizing selection, SelACcan be generalized further to include
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   diversifying selection. Specifically, by letting SelAC's Grantham weighting term G, which
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    we now assume is > 0, to take on negative values, SelACnow behaves as if there is a
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    pessimal, rather than optimal, amino acid for the given site. In this diversifying selection
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    scenario, amino acids with physico-chemical qualities more dissimilar to the pessimal amino
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    acid are increasingly favored, potentially resulting in multiple fitness peaks. This ability to
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    extend our model and, in turn, sharpen our thinking about the nature of natural selection
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   on amino acid sequences illustrates the value of moving from descriptive to more
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   mechanistic models in general and phylogenetics in particular. How frequently diversifying
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   selection of this nature occurs is an open, but addressable, question. Regardless of the
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    frequency at which diversifying selection occurs, it leads to the question, "How often does
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   the optimal/pessimal amino sequence change along any given branch?" Due to its
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   mechanistic nature, SelACcan also be extended to include changes in the optimal/pessimal
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   sequence over a phylogeny using a hidden markov modelling approach. Extending SelACin
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these ways, will allow researchers to model only shifts in natural selection and, in turn, quantify their frequency and magnitude.

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In summary, SelAC allows biologically relevant population genetic parameters to be 708 estimated from phylogenetic information, while also dramatically improving fit and 709 accuracy of phylogenetic models. By explicitly modeling the optimal/pessimal sequence of 710 a gene, SelACcan be extended to include shifts in the optimal/pessimal sequence over 711 evolutionary time. Exending our model in this way, will allow researchers to describe not 712 only the dynamic shifts in natural selection, but evaluate how well a given dataset supports 713 such a model. Moreover, it demonstrates that there remains substantially more information 714 in the coding sequences used for phylogenetic analysis than other methods can access. 715 Given the enormous amount of efforts expended to generate sequence datasets, it makes 716 sense for researchers to continue developing more realistic models of sequence evolution in 717 order to extract the biological information embedded in these datasets. The cost-benefit model we develop here is just one of many possible paths of mechanistic model 719 development. 720

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730

731 REFERENCES

- Beaulieu, J. M., B. C. O'Meara, and M. J. Donoghue. 2013. Identifying Hidden Rate
- Changes in the Evolution of a Binary Morphological Character: The Evolution of Plant
- Habit in Campanulid Angiosperms. Systematic Biology 62:725–737.
- Beaulieu, J. M. and B. C. OMeara. 2016. Detecting Hidden Diversification Shifts in Models
- of Trait-Dependent Speciation and Extinction. Systematic Biology 65:583–601.
- Berg, J. and M. Lässig. 2003. Stochastic Evolution and Transcription Factor Binding Sites.
- ⁷³⁸ Biophysics 48:S36–S44.
- Butler, M. A. and A. A. King. 2004. Phylogenetic comparative analysis: a modeling
- approach for adaptive evolution. American Naturalist 164:683–695.
- Dimmic, M. W., D. P. Mindell, and R. A. Goldstein. 2000. Modeling evolution at the
- protein level using an adjustable amino acid fitness model. Pacific Symposium on
- Biocomputing 5:18–29.
- Drummond, D. A., J. D. Bloom, C. Adami, C. O. Wilke, and F. H. Arnold. 2005. Why
- highly expressed proteins evolve slowly. Proceedings of the National Academy of Sciences
- of the United States of America 102:14338–14343.
- Drummond, D. A., A. Raval, and C. O. Wilke. 2006. A single determinant dominates the
- rate of yeast protein evolution. Molecular Biology and Evolution 23:327–337.
- Edwards, A. 1984. Likelihood. Cambridge science classics Cambridge University Press.
- Felsenstein, J. 1981. Evolutionary trees from DNA-sequences a maximum-likelihood
- approach. Journal of Molecular Evolution 17:368–376.

- Felsenstein, J. 2001. Taking Variation of Evolutionary Rates Between Sites into Account in
 Inferring Phylogenies. Journal of Molecular Evolution 53:447–455.
- Fisher, S., Ronald A. 1930. The Genetical Theory of Natural Selection. Oxford University
 Press, Oxford.
- Gilchrist, M., P. Shah, and R. Zaretzki. 2009. Measuring and detecting molecular
 adaptation in codon usage against nonsense errors during protein translation. Genetics
 183:1493–1505.
- Gilchrist, M. A. 2007. Combining Models of Protein Translation and Population Genetics
 to Predict Protein Production Rates from Codon Usage Patterns. Molecular Biology and
 Evolution 24:2362–2373.
- Gilchrist, M. A., W.-C. Chen, P. Shah, C. L. Landerer, and R. Zaretzki. 2015. Estimating
 Gene Expression and Codon-Specific Translational Efficiencies, Mutation Biases, and
 Selection Coefficients from Genomic Data Alone. Genome Biology and Evolution
 7:1559–1579.
- Gilchrist, M. A. and A. Wagner. 2006. A model of protein translation including codon bias, nonsense errors, and ribosome recycling. Journal of Theoretical Biology 239:417–434.
- Goldman, N. 1993. Statistical tests of models of DNA substitution. Journal of molecular evolution 36:182–198.
- Goldman, N., J. L. Thorne, and D. T. Jones. 1996. Using Evolutionary Trees in Protein
 Secondary Structure Prediction and Other Comparative Sequence Analyses. Journal of
 Molecular Biology 263:196 208.
- Goldman, N., J. L. Thorne, and D. T. Jones. 1998. Assessing the Impact of Secondary

 Structure and Solvent Accessibility on Protein Evolution. Genetics 149:445–458.

- Goldman, N. and Z. H. Yang. 1994. Codon-based model of nucleotide substitution for
- protein-coding DNA-sequences. Molecular Biology and Evolution 11:725–736.
- Grantham, R. 1974. Amino acid difference formula to help explain protein evolution.
- ⁷⁷⁸ Science 185:862–864.
- Halpern, A. L. and W. J. Bruno. 1998. Evolutionary distances for protein-coding sequences:
- Modeling site-specific residue frequencies. Molecular Biology And Evolution 15:910–917.
- Hughes, A. L. and M. Nei. 1988. Pattern of nucleotide substitution at major
- histocompatibility complex class-i loci reveals overdominant selection. Nature
- 783 335:167–170.
- Hurvich, C. M. and C.-L. Tsai. 1989. Regression and time series model selection in small samples. Biometrika 76:297–307.
- Hhna, S., M. J. Landis, T. A. Heath, B. Boussau, N. Lartillot, B. R. Moore, J. P.
- Huelsenbeck, and F. Ronquist. 2016. RevBayes: Bayesian Phylogenetic Inference Using
- Graphical Models and an Interactive Model-Specification Language. Systematic Biology
- 789 65:726.
- Ingram, T. and D. L. Mahler. 2013. SURFACE: detecting convergent evolution from data
- by fitting Ornstein-Uhlenbeck models with stepwise Akaike Information Criterion.
- Methods in ecology and evolution 4:416–425.
- ⁷⁹³ Iwasa, Y. 1988. Free fitness that always increases in evolution. Journal of Theoretical
- ⁷⁹⁴ Biology 135:265–281.
- Jhwueng, D.-C., H. Snehalata, B. C. O'Meara, and L. Liu. 2014. Investigating the
- performance of AIC in selecting phylogenetic models. Statistical applications in genetics
- and moleculr biology 13:459–475.

- Johnson, S. G. 2012. The NLopt nonlinear-optimization package. Version 2.4.2 Released
 20 May 2014.
- Kimura, M. 1962. on the probability of fixation of mutant genes in a population. Genetics
 47:713-719.
- Koshi, J. M. and R. A. Goldstein. 1997. Mutation matrices and physical-chemical properties: Correlations and implications. Proteins-Structure Function And Genetics 27:336–344.
- Koshi, J. M. and R. A. Goldstein. 2000. Analyzing site heterogeneity during protein evolution. Pages 191–202 *in* Biocomputing 2001. World Scientific.
- Koshi, J. M., D. P. Mindell, and R. A. Goldstein. 1999. Using physical-chemistry-based substitution models in phylogenetic analyses of HIV-1 subtypes. Molecular biology and evolution 16:173–179.
- Kubatko, L., P. Shah, R. Herbei, and M. A. Gilchrist. 2016. A codon model of nucleotide
 substitution with selection on synonymous codon usage. Molecular Phylogenetics and
 Evolution 94:290 297.
- Lartillot, N. and H. Philippe. 2004. A Bayesian mixture model for across-site
 heterogeneities in the amino-acid replacement process. Molecular Biology And Evolution
 21:1095–1109.
- Mayrose, I., N. Friedman, and T. Pupko. 2005. A Gamma mixture model better accounts for among site rate heterogeneity. Bioinformatics 21:ii151–ii158.
- McCandlish, D. M. and A. Stoltzfus. 2014. Modeling evolution using the probability of fixation: History and implications. The Quarterly Review of Biology 89:225–252.

- Muse, S. V. and B. S. Gaut. 1994. A likelihood approach for comparing synonymous and
- nonsynonymous nucleotide substitution rates, with application to the chloroplast
- genome. Molecular Biology and Evolution 11:715–724.
- Nowak, M. A. 2006. Evolutionary Dynamics: Exploring the Equations of Life. Belknap of
- Harvard University Press, Cambridge, MA.
- O'Meara, B. C., C. Ane, M. J. Sanderson, and W. P.C. 2006. Testing for different rates of
- continuous trait evolution using likelihood. Evolution 60:922–933.
- Pouyet, F., M. Bailly-Bechet, D. Mouchiroud, and L. Guguen. 2016. SENCA: A
- Multilayered Codon Model to Study the Origins and Dynamics of Codon Usage. Genome
- Biology and Evolution 8:2427–2441.
- Robinson, D. M., D. T. Jones, H. Kishino, N. Goldman, and J. L. Thorne. 2003. Protein
- evolution with dependence among codons due to tertiary structure. Molecular Biology
- 832 And Evolution 20:1692–1704.
- Rodrigue, N. and N. Lartillot. 2014. Site-heterogeneous mutation-selection models within
- the PhyloBayes-MPI package. Bioinformatics 30:1020–1021.
- Rodrigue, N., N. Lartillot, D. Bryant, and H. Philippe. 2005. Site interdependence
- attributed to tertiary structure in amino acid sequence evolution. Gene 347:207–217.
- Rokas, A., B. L. Williams, N. King, and S. B. Carroll. 2003. Genome-scale approaches to
- resolving incongruence in molecular phylogenies. Nature 425:798–804.
- Rowan, T. 1990. Functional Stability Analysis of Numerical Algorithms. Ph.D. thesis
- University of Texas, Austin.
- Salichos, L. and A. Rokas. 2013. Inferring ancient divergences requires genes with strong
- phylogenetic signals. Nature 497:327–331.

- Sella, G. and A. E. Hirsh. 2005. The application of statistical physics to evolutionary
- biology. Proceedings of the National Academy of Sciences of the United States of
- 845 America 102:9541–9546.
- Shah, P. and M. A. Gilchrist. 2011. Explaining complex codon usage patterns with
- selection for translational efficiency, mutation bias, and genetic drift. Proceedings of the
- National Academy of Sciences of the United States of America 108:10231–10236.
- 849 Stamatakis, A. 2006. RAxML-VI-HPC: maximum likelihood-based phylogenetic analyses
- with thousands of taxa and mixed models. Bioinformatics 22:2688–2690.
- Thorne, J. L., N. Goldman, and D. T. Jones. 1996. Combining protein evolution and
- secondary structure. Molecular Biology and Evolution 13:666–673.
- Thorne, J. L., N. Lartillot, N. Rodrigue, and S. C. Choi. 2012. Codon models as a vehicle
- for reconciling population genetics with inter-specific sequence data. Codon Evolution:
- Mechanisms And Models Pages 97–110 D2 10.1093/acprof:osobl/9780199601165.001.0001
- 856 ER.
- Wagner, A. 2005. Energy constraints on the evolution of gene expression. Molecular
- Biology and Evolution 22:1365–1374.
- Wright, S. 1969. Evolution and the genetics of populations. Vol. 2. The theory of gene
- frequencies. vol. 2. University of Chicago Press.
- Yang, Z. 2014. Molecular Evolution: A Statistical Approach. Oxford University Press, New
- York.
- Yang, Z. H. 1994. Maximum-likelihood phylogenetic estimation from DNA-sequences with
- variable rates over sites approximate methods. Journal Of Molecular Evolution
- 39:306–314.

- Yang, Z. H. 2007. PAML 4: Phylogenetic analysis by maximum likelihood. Molecular Biology And Evolution 24:1586–1591.
- Yang, Z. H. and R. Nielsen. 2008. Mutation-selection models of codon substitution and 868
- their use to estimate selective strengths on codon usage. Molecular Biology and
- Evolution 25:568–579.

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TABLE

		Parameters				Model
Model	logLik	Estimated	AIC	AICc	$\Delta { m AICc}$	Weight
$GTR+\Gamma$	-655,166.4	610	1,311,553	1,311,554	287,415	< 0.001
GY94	-612,670.4	111	1,225,563	$1,\!225,\!785$	XXXX	< 0.001
YN98	-594,713.9	111	1,118,650	1,189,872	XXXX	< 0.001
FMutSel	-597,140.7	178	1,194,637	1,194,994	XXXX	< 0.001
SelAC_M	-478,302.4	50,004	1,056,613	1,176,682	XXXXX	< 0.001
SelAC	-464,114.8	50,004	1,028,238	1,148,307	XXXXX	< 0.001
$SelAC_M + \Gamma$	-465,106.9	50,005	1,030,189	1,150,296	XXXXX	< 0.001
$\mathrm{SelAC+}\Gamma$	-453,620.8	50,005	1,007,252	1,127,324	XXXXX	> 0.999

Table 1: Comparison of model fits using AIC, AICc, and AIC_w. Note the subscripts M indicate model fits where the most common or 'majority rule' amino acid was fixed as the optimal amino acid a^* for each site. As discussed in text, despite the fact that a^* for each site was not fitted by our algorithm, its value was determined by examining the data and, as a result, represent an additional parameter estimated from the data and are accounted for in our table.

FIGURES

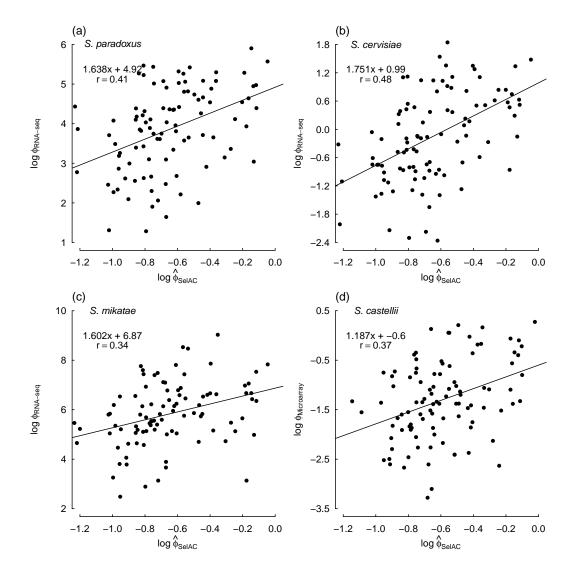


Figure 1: Comparisons between estimates of average protein translation rate $\hat{\phi}_{SelAC}$ obtained from SelAC+ Γ and direct measurements of expression for individual yeast taxa across the 100 selected genes from Salichos and Rokas (2013). Estimates of $\hat{\phi}_{SelAC}$ were generated by dividing the composite term ψ' by $\mathbf{B}(\vec{a}_i|\vec{a}^*)$. Gene expression was measured using either RNA-Seq (a)-(c) or microarray (d). The equations in the upper right hand corner of each panel represent the regression fit and correlation coefficient r.

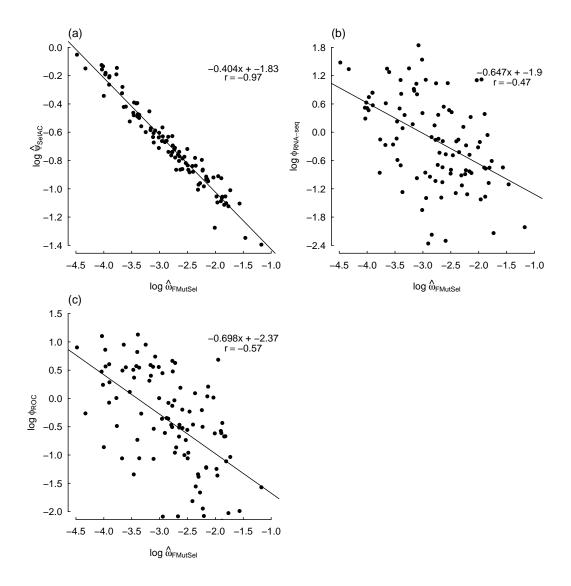


Figure 2: Comparisons between $\omega_{\rm FMutSel0}$, which is the nonsynonymous/synonymous mutation ratio in FMutSel0, SelAC+ Γ estimates of protein functionality production rates $\hat{\psi}_{\rm SelAC}$ (a), RNA-Seq based measurements of mRNA abundance $\phi_{\rm RNA-seq}$ (b), and ROC-SEMPPER's estimates of protein translation rates $\phi_{\rm ROC}$, which are based solely on *S. cerevisiae*'s patterns of codon usage bias (c), for *S. cerevisiae* across the 100 selected genes from Salichos and Rokas (2013). As in Figure 1, the equations in the upper left hand corner of each panel provide the regression fit and correlation coefficient.

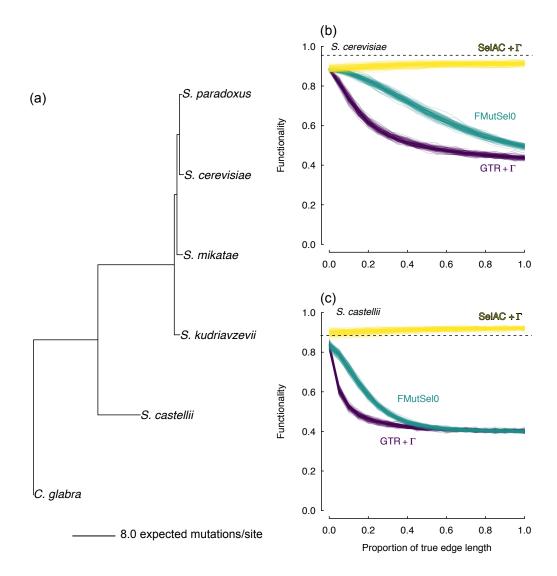


Figure 3: (a) Maximum likelihood estimates of branch lengths under SelAC+ Γ for 100 selected genes from Salichos and Rokas (2013). Tests of model adequacy for *S. cerevisiae* (b) and *S. castellii* (c) indicated that, when these taxa are removed from the tree, and their sequences are simulated, the parameters of SelAC+ Γ exhibit functionality $\mathbf{B}(\vec{a}_{\text{obs}}|\vec{a}^*)$ that is far closer to the observed (dashed black line) than data sets produced from parameters of either FMutSel0 or GTR + Γ .

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Supporting Materials for Population Genetics Based Phylogenetics Under Stabilizing

Selection for an Optimal Amino Acid Sequence: A Nested Modeling Approach by Beaulieu

et al. (In Review).

Comparisons of SelAC gene expression estimates with empirical measurements

In our model, the parameter ϕ measures the realized average protein synthesis rate 879 of a gene. We compared our estimates of ϕ to two separate measures of gene expression, 880 one empirical (Figure S1), and one model-based prediction that does not account for 881 shared ancestry, for individual yeast taxa across the same set of genes. Our estimates of ϕ 882 are positively correlated with both measures, which are also strongly correlated with each 883 other (Figure 1 - S2) On the whole, these comparisons indicate not only a high degree of 884 consistency among all three measures, but also, importantly, that estimates of ϕ obtained 885 from SelAC provide real biological insight into the expression level of a gene. 886

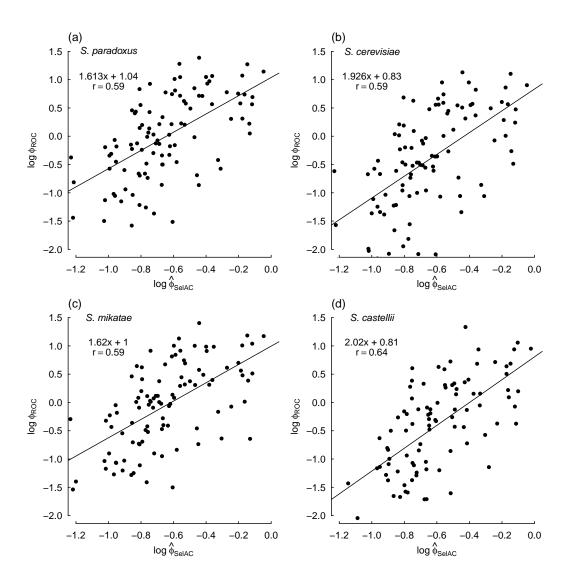


Figure S1: Comparisons between estimates of ϕ obtained from SelAC+ Γ and the predicted gene expression from the ROC SEMPER model (Gilchrist et al. (2015)) for individual yeast taxa across the 100 selected genes from Salichos and Rokas (2013). As with figures in the main text, estimates of ϕ were obtained by solving for ψ based on estimates of ψ' , and then dividing by $\mathbf{B}(\vec{a}_i|\vec{a}^*)$. The equations in the upper left hand corner of each panel represent the regression fit and correlation coefficient.

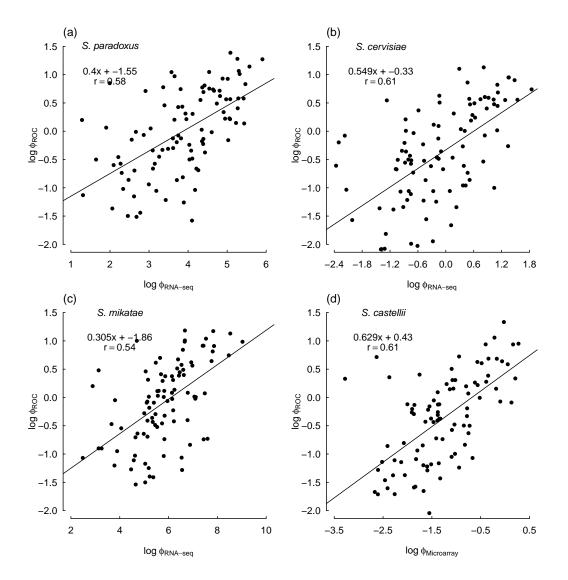


Figure S2: Comparisons of predicted gene expression from the ROC SEMPER model (Gilchrist et al. (2015)) and direct measurements of expression from RNA-Seq or microarray data for individual yeast taxa across the 100 selected genes from Salichos and Rokas (2013). The equations in the upper left hand corner of each panel represent the regression fit and correlation coefficient.

Simulations

Overall, the simulation results indicate that the SelAC model can reasonably recover the

known values of the generating model (Figure S3 - S6). This includes not only the

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parameters in the modelSelAC, but also the optimal amino acids for a given sequence as well as the estimates of the branch lengths. There are a few observations to note. First, the 891 ability to accurately recover the true optimal amino acid sequence will largely depend on 892 the magnitude of the realized average protein synthesis rate of the gene ϕ . This is, of 893 course, intuitive, given that ϕ sets the strength of stabilizing selection towards an optimal 894 amino acid at a site. However, the inclusion of α_G into the model SelAC, appears to 895 generally increase values of ϕ and generally improves the ability to recover the optimal 896 amino acids even for the gene with the lowest baseline ϕ . Second, we found a strong 897 downward bias in estimates of α_G , which actually translates to greater variation among the 898 rate categories. The choice of a gamma distribution to represent site-specific variation in 899 sensitivity was based on mathematical convenience and convention, rather than on 900 biological reality. Nevertheless, we suspect that this bias is in large part due to the difficulty in determining the baseline ψ for a given gene and the value of α_G that globally 902 satisfies the site-specific variation in sensitivity across all genes, as indicated by the slight 903 upward bias in estimates of ψ . A reviewer pointed out that it may also be difficulty for the 904 model-SelAC to account for changing amino-acid, which we agree may also play a role. It 905 has been suggested, in studies of the behavior of the gamma distribution in applications of 906 nucleotide substitution model, that increasing the number of rate categories can often 907 improve accuracy of the shape parameter (Mayrose et al. (2005)). Future work will address 908 this issue. 900



Figure S3: Summary of a 5-gene simulation for a SelAC model where we assume $\alpha_G = \infty$, and thus, no site-specific sensitivity in the generating model. The 'known' parameters were based on fitting the same SelAC to the 106 gene data set and phylogeny of Rokas et al. (2003), with gene choice being based on five evenly spaced points along the rank order of the gene specific composite parameter ψ'_g . The points and associated uncertainty in the estimates of the gene-specific average protein synthesis rate, or ψ (calculated from ψ')(a), nucleotide mutation rates under the UNREST model (b), proportion of correct optimal amino acids for a given gene (c), and estimates of the individual edge lengths are based the mean and 2.5% and 97.5% quantiles across all 50 simulated datasets (d). Gene index on the x-axis refers to the arbitrary number assigned to the simulated gene.



Figure S4: The distribution of estimates of the Grantham weights, α_c and α_p , in a SelAC model, where we assume $\alpha_G = \infty$, and thus no site-specific sensitivity in the generating model. The dashed line represents the value used in the generating model.



Figure S5: Same figure as in Figure S3, except the generating model includes site-specific sensitivity in the generating model (i.e., α_G).

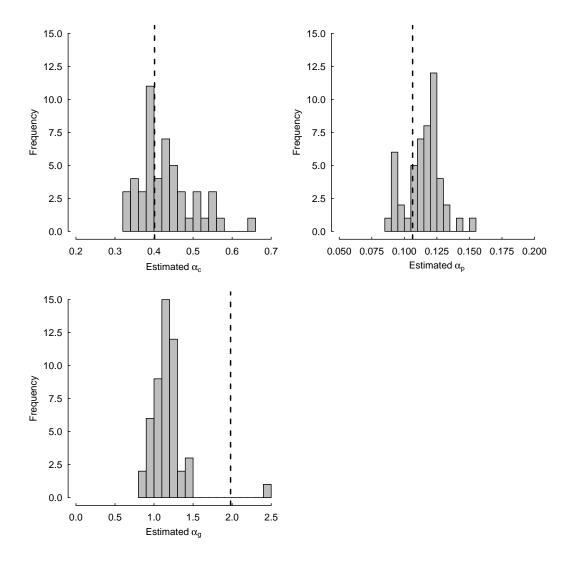


Figure S6: Same figure as in Figure S4, except the generating model includes site-specific sensitivity in the generating model (i.e., α_G). Unlike, Grantham weights, which showed no systematic bias, there is a downward bias in estimates of α_G .