3 Empirical properties of financial data

3.1 Stylized facts of financial return series

3.2 Multivariate stylized facts

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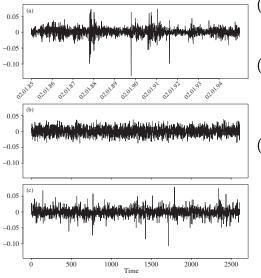
3.1 Stylized facts of financial return series

- The stylized facts are a collection of empirical observations and related inferences, which apply to many time series of risk-factor changes (e.g. log-returns on equities, indices, exchange rates, commodity prices).
- The best-known stylized facts apply to daily log-returns (also to intradaily, weekly, monthly). Tick-by-tick (high-frequency) data have their own stylized facts (not discussed here) and annual return (low-frequency) data are more difficult to investigate (data sparseness; non-stationarity).
- Consider discrete-time risk-factor changes $X_t = Z_t Z_{t-1}$ for a log-price or rate $Z_t = \log S_t$. In this case

$$X_t = \log(S_t/S_{t-1}) \approx S_t/S_{t-1} - 1 = (S_t - S_{t-1})/S_{t-1};$$

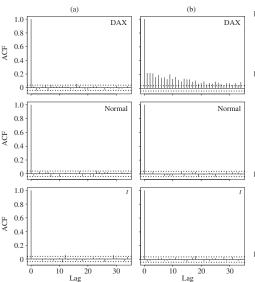
the former is often called a (log-)return, the latter a relative return.

3.1.1 Volatility Clustering



- (a) Log-returns for the DAX index from 1985-01-02 to 1994-12-30 (n=2608).
- (b) Simulated iid data from a fitted normal with $\hat{\mu} = \bar{X}_n$, $\hat{\sigma}^2 = \frac{1}{n} \sum_{i=1}^n (X_i \bar{X}_n)^2$ show too few extremes.
- (c) Simulated iid data from a fitted $t_{3.8}$. Better range of values but still no volatility clustering (= tendency for extreme returns to be followed by extreme returns).

Estimated autocorrelation function (ACF) $\rho(h) = \operatorname{corr}(X_0, X_h), h \in \mathbb{Z}$

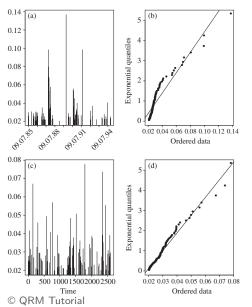


- Estimated (a) ACF of $(X_t)_{t\in\mathbb{Z}}$ (b) ACF of $(|X_t|)_{t\in\mathbb{Z}}$
- Non-zero ACF at lag 1 implies a tendency for a return to be followed by a return of equal sign; not the case here. \Rightarrow Predicted return ≈ 0

 - For iid process $(X_t)_{t\in\mathbb{Z}}$ $\rho_X(h) = \rho_{|X|}(h) = I_{\{h=0\}};$ not the case here. (Can confirm with Ljung-Box tests.)
- $(Z_t)_{t\in\mathbb{Z}}$ not a random walk $(S_t)_{t\in\mathbb{Z}}$ not GBM.

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Concerning clustering of extremes, consider the 100 largest losses of the. . .



- (a) ... DAX index (c) ... simulated fitted $t_{3.8}$
- (b), (d) Q-Q plots of waiting times between these large losses (should be Exp(λ) for iid data; see EVT chapter).
- The DAX data shows shorter and longer waiting times than the iid data
 - \Rightarrow clustering of extremes.

3.1.2 Non-normality and heavy tails

Formal statistical tests of normality

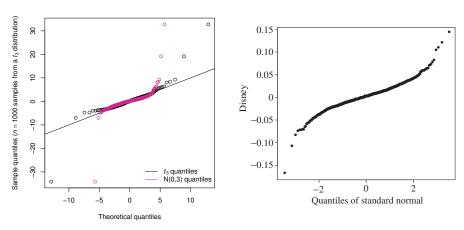
- For general univariate df *F*:
 - ► Kolmogorov–Smirnov (test statistic $T_n = \sup_x |\hat{F}_n(x) F(x)|$)
 - ▶ Cramér–von Mises $(T_n = n \int_{-\infty}^{\infty} (\hat{F}_n(x) F(x))^2 dF(x))$
 - ▶ Anderson–Darling $(T_n = n \int_{-\infty}^{\infty} \frac{(\bar{F}_n(x) F(x))^2}{F(x)(1 F(x))} dF(x)$; recommended by D'Agostino and Stephens (1986))
- For $F = N(\mu, \sigma^2)$:
 - Shapiro–Wilk (idea: quantify Q-Q plot in one number)
 - D'Agostino (based on skewness and kurtosis as Jarque–Bera)
 - ▶ Jarque–Bera test: Compares skewness $\beta = \frac{\mathbb{E}((X-\mu)^3)}{\sigma^3}$ and kurtosis $\kappa = \frac{\mathbb{E}((X-\mu)^4)}{\sigma^4}$ with sample versions. The test statistic is

$$T_n = \frac{n}{6} (\hat{\beta}^2 + \frac{1}{4} (\hat{\kappa} - 3)^2) \stackrel{H_0}{\sim} \underset{n \, \text{large}}{\sim} \chi_2^2.$$

Graphical tests

- Suppose we want to graphically test whether $X_1, \ldots, X_n \sim F$ for some df F based on realizations of iid X_1, \ldots, X_n .
- Let $X_{(1)} \leq \cdots \leq X_{(n)}$ denote the corresponding order statistics and note that $\hat{F}_n(x) = \frac{1}{n} \sum_{i=1}^n I_{\{X_i \leq x\}} = \frac{1}{n} \sum_{i=1}^n I_{\{X_{(i)} \leq x\}}$, $x \in \mathbb{R}$, i.e. the order statistics contain all relevant information about X_1, \ldots, X_n .
- Possible graphical tests (see also the appendix):
 - ▶ P-P plot: Plot $\{(p_i, F(X_{(i)})) : i = 1, \ldots, n\}$, where $p_i = \frac{i-1/2}{n} \approx \frac{i}{n} \approx \hat{F}_n(X_{(i)})$. If $F \approx \hat{F}_n$, the points lie roughly on a line with slope 1; this also applies to Q-Q plots.
 - ▶ Q-Q plot: Plot $\{(F^{\leftarrow}(p_i), X_{(i)}) : i = 1, ..., n\}$ (tail differences better visible).

Interpreting Q-Q plots (S-shape hints at heavier tails than $N(\mu, \sigma^2)$):



Daily returns typically have kurtosis $\kappa>3$ (leptokurtic; narrower center, heavier tails than $N(\mu,\sigma^2)$ for which $\kappa=3$). They have power-like tails rather than exponential.

3.1.3 Longer-interval return series

- By going from daily to weekly, monthly, quarterly and yearly data, these effects become less pronounced (returns look more iid, less heavy-tailed).
- \blacksquare The (non-overlapping) h-period log-return at $t\in\{h,2h,\ldots,\lfloor\frac{n}{h}\rfloor h\}$ is

$$X_t^{(h)} = \log\left(\frac{S_t}{S_{t-h}}\right) = \log\left(\frac{S_t}{S_{t-1}} \frac{S_{t-1}}{S_{t-2}} \dots \frac{S_{t-h+1}}{S_{t-h}}\right) = \sum_{k=0}^{h-1} X_{t-k}$$

A Central Limit Theorem (CLT) effect takes place (less heavy-tailed, less evidence of serial correlation).

- \blacksquare Problem: the larger h, the less data are available.
- Possible remedy: Consider overlapping returns

$$\left\{X_t^{(h)} : t \in \left\{h, h+k, \dots, h + \left\lfloor \frac{n-h}{k} \right\rfloor k\right\}\right\}, \quad 1 \le k < h.$$

⇒ More data but serially dependent now.

To summarize, we can infer the following stylized facts about univariate financial return series:

- (U1) Return series are not iid although they show little serial correlation;
- (U2) Series of absolute or squared returns show profound serial correlation;
- (U3) Conditional expected returns are close to zero;
- (U4) Volatility (conditional standard deviation) appears to vary over time;
- (U5) Extreme returns appear in clusters;
- (U6) Return series are leptokurtic or heavy-tailed (power-like tail).

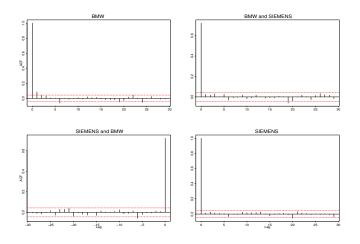
3.2 Multivariate stylized facts

Consider multivariate log-return data X_1, \ldots, X_n .

3.2.1 Correlation between series

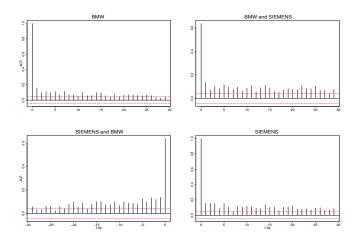
- By (U1), the returns of stock A at t and t+h show little correlation. The same applies to the returns of stock A at t and stock B at t+h, h>0. Stock A and stock B on day t may be correlated due to factors that affect the whole market (*contemporaneous dependence*).
- Correlations of returns at t vary over time (difficult to detect whether changes are continual or constant within regimes; fit different models for changing correlation, then make a formal comparison).
- Periods of high/low volatility are typically common to more than one stock \Rightarrow Returns of large magnitude in A at t may be followed by returns of large magnitude in A and B at t+h.

Estimated correlations between/within series:



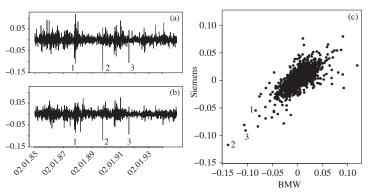
2000 values from period 1985-01-23 to 1994-09-22

Estimated correlations between/within series of absolute values:



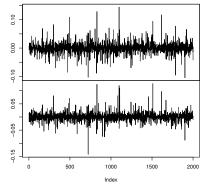
3.2.2 Tail dependence

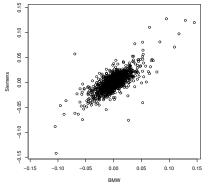
(BMW, Siemens) log-returns from 1985-01-23 to 1994-09-22 (n=2000)



In volatile/extreme periods, dependence is stronger (1: 1987-10-19 Black Monday (DJ drop by 22%); 2: 1989-10-16 Monday demonstrations in Leipzig (Wende); 3: 1991-08-19 coup against soviet president M. Gorbachev).

Simulated log-returns from a fitted bivariate t distribution (n=2000; $\rho=0.72,\ \nu=2.8$ both fitted to (BMW, Siemens))





- The multivariate t distribution can replicate joint large gains/losses but in a symmetric way.
- The multivariate normal distribution cannot replicate such behaviour, known as tail dependence; see Chapter 7.

To summarize, we can infer the following stylized facts about multivariate financial return series:

- (M1) Multivariate return series show little evidence of cross-correlation, except for contemporaneous returns (i.e. at the same t);
- (M2) Multivariate series of absolute returns show profound cross-correlation;
- (M3) Correlations between contemporaneous returns vary over time;
- (M4) Extreme returns in one series often coincide with extreme returns in several other series (e.g. tail dependence).