CENG 222Statistical Methods for Computer Engineering

Week 6

Chapter 5
Computer Simulations and Monte Carlo Methods

Outline

- Generation of random numbers from specific distributions
 - Discrete distributions
 - Continuous distributions
- Chebyshev's inequality (3.3.7)
- Solving problems by Monte Carlo methods
 - Estimating probabilities
 - Estimating means and standard deviations

Uniform Random Numbers

- Tables of random numbers
- Pseudo-random number generators
 - Long sequences of random-looking numbers
 - Seed: starting location in the sequence
 - May use system time as seed
- Many systems provide standard uniform random number generators
 - Uniform(0,1)
- Question: Can we generate random numbers from any distribution using Uniform(0,1) rvs?

Bernoulli

• Let *U* be Uniform(0,1)

•
$$X = \begin{cases} 1, & \text{if } U$$

• P (success) = P (U < p) = p

Binomial

- Sum of *n* independent Bernoulli variables.
- Example

```
n = 20; p = 0.68;
U = rand(n,1);
% generates an nx1 vector
% of uniform random numbers
X = sum(U < p);</pre>
```

Geometric

- Iterate and count the number of generated rvs until first success
- Example:

```
p = 0.16; X = 1;
while rand > p;
    X = X+1;
end;
X
```

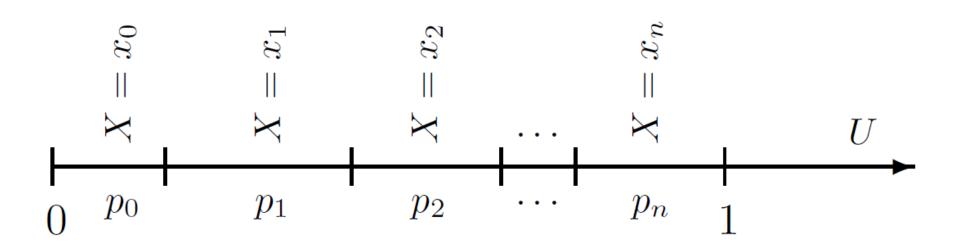
Negative Binomial

- Generate k independent Geometric(p) random numbers and sum them to get a NegativeBinomial(k,p) number.
- Example:

```
p = 0.16; X = 0; i = 0;
while i < k;
G = 1;
while rand > p;
G = G+1;
end;
X = X+G;
end;
```

• How efficient is generating a Binomial, a Geometric, or a Negative Binomial random number?

Arbitrary discrete distributions



Algorithm 5.1

- 1. Divide the interval [0,1] into subintervals A_i as follows:
 - $-A_i = [p_0 + p_1 + ... + p_{i-1}, p_0 + p_1 + ... + p_i)$
- 2. Generate *U*, a standard uniform number
- 3. If *U* belongs to A_i then $X = x_i$
- How efficient is this method?
 - If you want to generate many Xs, efficiency is important.
 - O(n), $O(\log n)$, O(1)?
 - Check out the *Alias Method*, if you need an O(1) method.

Poisson

- Using Algorithm 5.1 to generate Poisson numbers.
- Example:

Inverse transform method

- Theorem: $U = F_X(X)$ is Uniform(0,1)
- Proof:
 - Note that the standard uniform cdf is $F_U(u) = u$ (i.e., $F'_U(u) = f_U(u) = 1$). We will try to show this fact using the given definition of $U = F_X(X)$

$$-F_{U}(u) = P(U \le u)$$

$$= P(F_{X}(X) \le u)$$

$$= P(X \le F_{X}^{-1}(u))$$

$$= F_{X}(F_{X}^{-1}(u))$$

$$= u$$

Inverse transform method

- If $U = F_X(X)$ then $X = F_X^{-1}(U)$
- The method:
 - Generate a uniform random number
 - Plug it in F_X^{-1} to generate X (i.e. solve for X).
- Example 5.10 (Exponential):

$$-F_X(X) = 1 - e^{-\lambda X} = U$$

$$- \rightarrow X = -\frac{1}{\lambda} \ln(1 - U)$$

- Can also use $X = -\frac{1}{\lambda} \ln(U)$ since 1-*U* is also Uniform(0,1).

Inverse transform method

- Difficult to use if the inverse of the cdf is not easy to compute
- For example, for discrete distributions, $F_X^{-1}(U)$ does not exist. $U = F_X(X)$ has no roots, because X (hence $F_X(X)$) is finite and countable; whereas U is continuous.
- Therefore, for discrete rvs, we use the inverse method with a slight modification:
 - $-X = \min \{x \in S \text{ such that } F(x) > U \}$ where S is the set of possible values of X.

Example 5.12

• Using the inverse transform method for generating Geometric variables

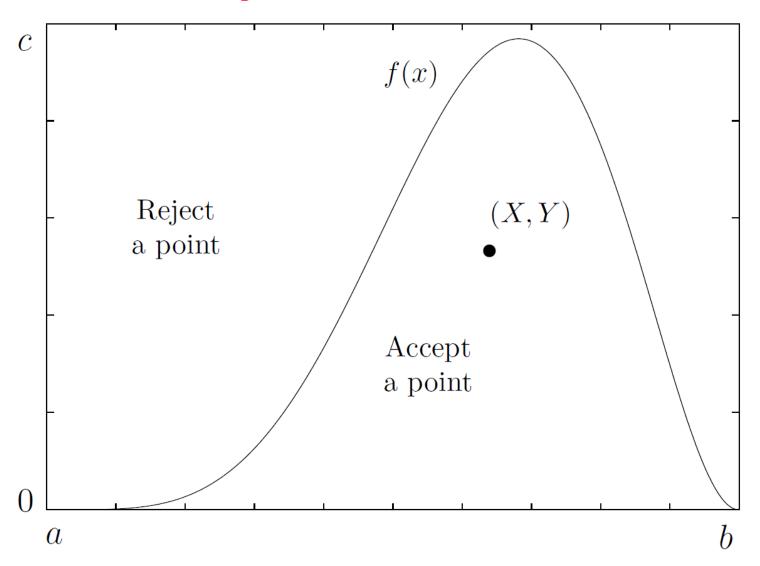
$$\bullet \ X = \left\lceil \frac{\ln(1-U)}{\ln(1-p)} \right\rceil$$

- The geometric variable is the ceiling of the exponential variable with $\lambda = -\ln(1-p)$
 - Exponential is the continuous analogue of geometric
 - Both have the memoryless property.

Rejection method

- When the cdf is difficult to solve for X and the pdf f_X is available, the rejection method can be used to generate random numbers from f_X .
- Idea:
 - Generate 2D uniform coordinates (X,Y) in the bounding box of f_X and if $Y \le f_X(X)$ output X.

Rejection method



Example

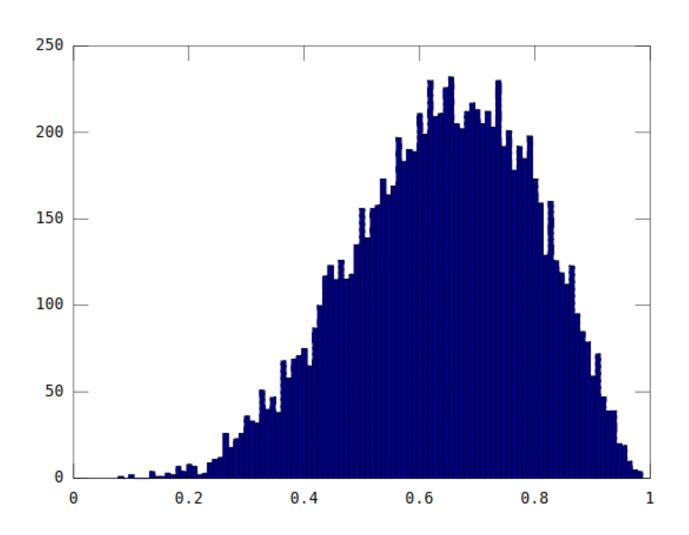
• The figure in the previous slide is the pdf of Beta(α =5.5, β =3.1)

$$-f_X = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} x^{\alpha-1} (1-x)^{\beta-1} \text{ for } 0 \le x \le 1$$

• Bounding box: m = 2.5, s = 0, t = 1.

```
a=5.5; b=3.1; s=0; t=1; m=2.5;
X = 0; Y = m;
F = gamma(a+b)/gamma(a)/gamma(b)*X^(a-1)*(1-X)^(b-1);
while (Y > F);
        U = rand; V = rand;
        X = s+(t-s)*U; Y = m*V;
        F = ... % same as above;
end; X
```

Example



Monte Carlo methods

- Generate many random variables from a distribution and estimate probabilities, means, standard deviations, etc. by simulating what happens in the long run.
- Question: How many numbers needed for acceptable results?
 - i.e., What will be the "size" of the Monte Carlo experiment?
 - Revisit Chebyshev's Inequality

Chebyshev's Inequality (3.3.7)

• For any distribution with expectation μ and variance σ^2 and for any positive ϵ

$$-P(|X-\mu|>\varepsilon)\leq \left(\frac{\sigma}{\varepsilon}\right)^2$$

In other words: any random variable X from the distribution is within ε distance of the μ with probability of at least $1 - (\sigma / \varepsilon)^2$

Estimating probabilities

- The probability $p=P(X \in A)$ can be estimated as \hat{p} by generating N random numbers and computing the proportion of random numbers that are in A.
- How accurate is the estimator?
 - What is $\mathbf{E}(\hat{p})$ and $\mathrm{Std}(\hat{p})$?
 - The number of X_i that are in A among the N generated random numbers is Binomial(N,p) with expectation Np and variance Np(1-p)
 - \rightarrow **E**(\hat{p}) = p (unbiased estimator)

Std(
$$\hat{p}$$
) = $\sqrt{\frac{p(1-p)}{N}}$ the error in \hat{p} decreases with $1/\sqrt{N}$

How large should N be?

- Given the error ε and the probability, α , to exceed this error limit
- If an intelligent guess p^* on the value of p is available:

$$-N \ge p^*(1-p^*)\left(\frac{z_{\alpha/2}}{\varepsilon}\right)^2$$

• If not:

$$-N \ge 0.25 \left(\frac{z_{\alpha/2}}{\varepsilon}\right)^2$$

Example 5.14

How large should N be?

- If the *N* returned by these equations are not large enough for Binomial approximation, we may use Chebyshev's inequality:
 - If an intelligent guess p^* on the value of p is available:

•
$$N \geq \frac{p^*(1-p^*)}{\alpha \varepsilon^2}$$

- If not:

•
$$N \ge \frac{1}{4\alpha\varepsilon^2}$$

Estimating means and standard deviations

$$\bullet \ \bar{X} = \frac{1}{N} (X_1 + \dots + X_N)$$

– also unbiased and its error decreases with $1/\sqrt{N}$

•
$$s^2 = \frac{1}{N-1} \sum_{i=1}^{N} (X_i - \bar{X})^2$$

-1/N-1 needed so that $\mathbf{E}(s^2) = \sigma^2$