Power and sample size calculations for interaction analysis

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Treatment-by-subgroup interactions

Subgrouping factor in clinical trials:

- treatment-by-centre interaction in multi-centre trials
- inconsistency in multi-regional trials
- subgroup heterogeneity in trials with pre-defined subgroups
- biomarker-by-treatment interactions
- **.** . . .

Cell means model in a randomized two-way layout

$$y_{ijk} = \mu_{ij} + \epsilon_{ijk}, \qquad \epsilon_{ijk} \sim N(0, \sigma^2)$$
 (1)
with $i = 1, ..., I$ (treatment factor)
 $j = 1, ..., J$ (sub-grouping factor)
 $k = 1, ..., N$, where $N = \sum_{i=1}^{I} \sum_{j=1}^{J} n_{ij}$

$$\mu = (\mu_{11}, \mu_{21}, \dots, \mu_{1J}, \mu_{2J}) = (\mu_{1}, \dots, \mu_{L})$$

Hypotheses to test for interactions

Product-type interaction contrasts:

$$H_0: \quad \boldsymbol{C}^{lnt}\boldsymbol{\mu} = \boldsymbol{\theta} \qquad H_A: \quad \boldsymbol{C}^{lnt}\boldsymbol{\mu} \neq \boldsymbol{\theta},$$
 (2)

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Ratios of treatment differences:

$$H_0^m: \quad \frac{\mathbf{h}_m \mu}{\mathbf{d}_m \mu} \ge \omega \qquad H_A^m: \quad \frac{\mathbf{h}_m \mu}{\mathbf{d}_m \mu} < \omega,$$
 (3)

where $\boldsymbol{C}_{Numerator} = (\boldsymbol{h}_1, \dots, \boldsymbol{h}_m)$ and $\boldsymbol{C}_{Denominator} = (\boldsymbol{d}_1, \dots, \boldsymbol{d}_m)$ denote the numerator and denominator interaction contrast matrix.

Test - Product-type interaction contrasts

The test-statistic to test the m local null hypothesis in Eq. 2 is given by

$$T_m = \frac{\sum_{l=1}^{L} c_l \hat{\mu}_l - \theta_m}{s \sqrt{\sum_{l=1}^{L} c_l^2 / n_l}} = \frac{\boldsymbol{c}_m \hat{\boldsymbol{\mu}} - \theta_m}{s \sqrt{\boldsymbol{c}_m' \boldsymbol{M} \boldsymbol{c}_m}},$$

where $\mathbf{M} = \operatorname{diag}\left(n_1^{-1},\ldots,n_L^{-1}\right)$ and $\mathbf{T} = (T_1,\ldots,T_M)$. Under H_0 \mathbf{T} follows a central multivariate t-distribution $Mt_{df=\nu,\mathbf{R}}$, with $\nu = \sum_{l=1}^L (n_l-1)$ and correlation matrix $\mathbf{R} = [\rho_{mm'}]$ with elements:

$$\rho_{mm'} = \frac{\sum_{l=1}^{L} c_{ml} c_{m'l} / n_{l}}{\sqrt{\left(\sum_{l=1}^{L} c_{ml}^{2} / n_{l}\right) \left(\sum_{l=1}^{L} c_{m'l}^{2} / n_{l}\right)}}.$$

Power - Product type interaction contrasts

The *m*th non-centrality parameter is given by [Bretz et al., 2001]

$$\delta_{m} = \frac{\boldsymbol{c}_{m}\boldsymbol{\mu} - \theta_{m}}{\sigma\sqrt{\boldsymbol{c}_{m}'\boldsymbol{M}\boldsymbol{c}_{m}}}.$$

Under the global null hypothesis $H_0: \mathbf{C}^{lnt}\mu = \theta$, the non-centrality parameter vector $\delta = \mathbf{0}$.

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Under the global null hypothesis H_0 : $\mathbf{C}^{lnt}\mu=\theta$, the non-centrality parameter vector $\delta=\mathbf{0}$.

The power of T_0 to detect H_A at significance level α is

$$P(T_0 \ge t_{1-\alpha,\nu,\mathbf{R}}|H_A) = 1 - \beta = 1 - P(|T_1| \le t_{1-\alpha,\nu,\mathbf{R}}, \dots, |T_M| \le t_{1-\alpha,\nu,\mathbf{R}}|H_A),$$

where $T_0 = \max(|T_1|, ..., |T_M|)$.

Software - Power for product-type interaction contrasts

R add-on package [R Core Team, 2013] MCPAN [Schaarschmidt et al., 2013] with its function powermcpt

Test - Ratios of treatment differences

The test-statistic to test the m local null hypothesis in Eq. 2 is given by

$$T_m^{\text{ratio}} = \frac{(\boldsymbol{h}_m - \omega \boldsymbol{d}_m) \, \hat{\boldsymbol{\mu}}}{s \sqrt{(\boldsymbol{h}_m - \omega \boldsymbol{d}_m) \, \boldsymbol{M} \, (\boldsymbol{h}_m - \omega \boldsymbol{d}_m)}}.$$

According to Dilba et al. [Dilba et al., 2006] under H_0 the vector of test statistics $\boldsymbol{T}^{Ratio} = (T_1^{Ratio}, \dots, T_M^{Ratio})$ follows a central multivariate t-distribution with ν degrees of freedom and a correlation matrix $\boldsymbol{R} = [\varsigma_{mm'}]$ with elements:

$$\varsigma_{mm'} = \frac{\left(\textbf{\textit{h}}_{m} - \omega \textbf{\textit{d}}_{m} \right) \textbf{\textit{M}} \left(\textbf{\textit{h}}_{m'} - \omega \textbf{\textit{d}}_{m'} \right)}{\sqrt{\left(\textbf{\textit{h}}_{m} - \omega \textbf{\textit{d}}_{m} \right) \textbf{\textit{M}} \left(\textbf{\textit{h}}_{m} - \omega \textbf{\textit{d}}_{m'} \right) } \sqrt{\left(\textbf{\textit{h}}_{m'} - \omega \textbf{\textit{d}}_{m'} \right) \textbf{\textit{M}} \left(\textbf{\textit{h}}_{m'} - \omega \textbf{\textit{d}}_{m'} \right)}}.$$

Power - Ratios of treatment differences

The non-centrality parameters are given by

$$\tau_{m} = \frac{(\boldsymbol{h}_{m} - \omega \boldsymbol{d}_{m}) \boldsymbol{\mu}}{\sigma \sqrt{(\boldsymbol{h}_{m} - \omega \boldsymbol{d}_{m}) \boldsymbol{M} (\boldsymbol{h}_{m} - \omega \boldsymbol{d}_{m})}}.$$

Under the global null hypothesis H_0 the non-centrality parameter vector $\tau = \mathbf{0}$ and T follows a central M variate t-distribution.

Power - Ratios of treatment differences

The non-centrality parameters are given by

$$\tau_{m} = \frac{\left(\boldsymbol{h}_{m} - \omega \boldsymbol{d}_{m}\right) \mu}{\sigma \sqrt{\left(\boldsymbol{h}_{m} - \omega \boldsymbol{d}_{m}\right) \boldsymbol{M} \left(\boldsymbol{h}_{m} - \omega \boldsymbol{d}_{m}\right)}}.$$

Under the global null hypothesis H_0 the non-centrality parameter vector $\tau = \mathbf{0}$ and \mathbf{T} follows a central M variate t-distribution. The power of T_0^{Ratio} to detect H_A at significance level α is

$$\begin{split} P\left(T_0^{\text{ratio}} > t_{1-\alpha,\nu,\textbf{\textit{R}}^{\text{ratio}}}|H_A\right) &= 1-\beta \\ &= 1-P\left(|T_1^{\text{ratio}}| \leq t_{1-\alpha,\nu,\textbf{\textit{R}}^{\text{ratio}}}, \dots, |T_M^{\text{ratio}}| \leq t_{1-\alpha,\nu,\textbf{\textit{R}}^{\text{ratio}}}|H_A\right) \end{split}$$

where $T_0^{\text{ratio}} = \max(|T_1^{\text{ratio}}|, \dots, |T_M^{\text{ratio}}|)$.

Software - Power ratios of treatment effects

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References I



Bretz, F., Genz, A., and Hothorn, L. (2001).

On the numerical availability of multiple comparison procedures. Biometrical Journal, 43(5):645–656.



Dilba, G., Bretz, F., Hothorn, L., and Guiard, V. (2006).

Power and sample size computations in simultaneous tests for non-inferiority based on relative margins. Statistics in Medicine, 25(7):1131–1147.



Kitsche, A. and Hothorn, L. A. (2014).

Testing for qualitative interaction using ratios of treatment differences. Statistics in Medicine, 13(9):1477–1489.



R Core Team (2013).

R: A Language and Environment for Statistical Computing.



Schaarschmidt, F., Gerhard, D., and Sill, M. (2013).

MCPAN: Multiple comparisons using normal approximation.

R package version 1.1-15.