

Mathematical Biostatistics Boot Camp 2: Lecture 8, Chi-Squared Tests

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September 30, 2013

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Chi-squared testing

- An alternative approach to testing equality of proportions uses the chi-squared statistic

$$\sum \frac{(\text{Observed} - \text{Expected})^2}{\text{Expected}}$$

- “Observed” are the observed counts
- “Expected” are the expected counts under the null hypothesis
- The sum is over all four cells
- This statistic follows a Chi-squared distribution with 1 df
- The Chi-squared statistic is exactly the square of the difference in proportions Score statistic

Example

| Trt | Side Effects | None | Total |
|-----|--------------|------|-------|
| X | 44 | 56 | 100 |
| Y | 77 | 43 | 120 |
| | 121 | 99 | 220 |

- p_1 and p_2 are the cure rates
- $H_0 : p_1 = p_2$

- The χ^2 statistic is $\sum \frac{(O-E)^2}{E}$
- $O_{11} = 44$, $E_{11} = \frac{121}{220} \times 100 = 55$
- $O_{21} = 77$, $E_{21} = \frac{121}{220} \times 120 = 66$
- $O_{12} = 56$, $E_{12} = \frac{99}{220} \times 100 = 45$
- $O_{22} = 43$, $E_{22} = \frac{99}{220} \times 120 = 54$

$$\chi^2 = \frac{(44 - 55)^2}{55} + \frac{(77 - 66)^2}{66} + \frac{(56 - 45)^2}{45} + \frac{(43 - 54)^2}{54}$$

Which turns out to be 8.96. Compare to a χ^2 with one degree of freedom (reject for large values).

```
pchisq(8.96, 1, lower.tail = FALSE)  
#result is 0.002
```

R code

```
dat <- matrix(c(44, 77, 56, 43), 2)
chisq.test(dat)
chisq.test(dat, correct = FALSE)
```

Notation reminder

| | | |
|--------------|--------------------|----------------|
| $n_{11} = x$ | $n_{12} = n_1 - x$ | $n_1 = n_{1+}$ |
| $n_{21} = y$ | $n_{22} = n_2 - y$ | $n_2 = n_{2+}$ |
| n_{+1} | n_{+2} | |

- Reject if the statistic is too large
- Alternative is two sided
- Do not divide α by 2
- A small χ^2 statistic implies little difference between the observed values and those expected under H_0
- The χ^2 statistic and approach generalizes to other kinds of tests and larger contingency tables
- Alternative computational form for the χ^2 statistic

$$\chi^2 = \frac{n(n_{11}n_{22} - n_{12}n_{21})^2}{n_{+1}n_{+2}n_{1+}n_{2+}}$$

- Notice that the statistic:

$$\chi^2 = \frac{n(n_{11}n_{22} - n_{12}n_{21})^2}{n_{+1}n_{+2}n_{1+}n_{2+}}$$

does not change if you transpose the rows and the columns of the table

- Surprisingly, the χ^2 statistic can be used
 - the rows are fixed (binomial)
 - the columns are fixed (binomial)
 - the total sample size is fixed (multinomial)
 - none are fixed (Poisson)
- For a given set of data, any of these assumptions results in the same value for the statistic

Testing independence

- Maternal age versus birthweight¹
- Cross-sectional sample, only the total sample size is fixed

| | Birthweight | | |
|------------|-------------|---------------|-------|
| Mat. Age | < 2500g | $\geq 2,500g$ | Total |
| < 20y | 20 | 80 | 100 |
| $\geq 20y$ | 30 | 270 | 300 |
| Total | 50 | 350 | 400 |

- H_0 : MA is independent of BW
- H_a : MA is not independent of BW

¹From Agresti Categorical Data Analysis second edition

Continued

- Under H_0 (est) $P(MA < 20) = \frac{100}{400} = .25$
- Under H_0 (est) $P(BW < 2500) = \frac{50}{400} = .125$
- Under H_0 (est)

$$P(MA < 20 \text{ and } BW < 2,500) = .25 \times .125$$

- Therefore
 - $E_{11} = \frac{100}{400} \times \frac{50}{400} \times 400 = 12.5$
 - $E_{12} = \frac{100}{400} \times \frac{350}{400} \times 400 = 87.5$
 - $E_{21} = \frac{300}{400} \times \frac{50}{400} \times 400 = 37.5$
 - $E_{22} = \frac{300}{400} \times \frac{350}{400} \times 400 = 262.5$
 - $\chi^2 = \frac{(20-12.5)^2}{12.5} + \frac{(80-87.5)^2}{87.5} + \frac{(30-37.5)^2}{37.5} + \frac{(270-262.5)^2}{262.5} = 6.86$
- Compare to critical value
`qchisq(.95, 1)=3.84`
- Or calculate P-value
`pchisq(6.86, 1, lower.tail = F)=.009`

Chi-squared testing cont'd

| Group | Alcohol use | | |
|------------|-------------|-----|-------|
| | High | Low | Total |
| Clergy | 32 | 268 | 300 |
| Educators | 51 | 199 | 250 |
| Executives | 67 | 233 | 300 |
| Retailers | 83 | 267 | 350 |
| Total | 233 | 967 | 1,200 |

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- Interest lies in testing whether or not the proportion of high alcohol use is the same in the four occupations
- $H_0 : p_1 = p_2 = p_3 = p_4 = p$
- $H_a : \text{at least two of the } p_j \text{ are unequal}$
- $O_{11} = 32, E_{11} = 300 \times \frac{233}{1200}$
- $O_{12} = 268, E_{12} = 300 \times \frac{967}{1200}$
- ...
- Chi-squared statistic $\sum \frac{(O-E)^2}{E} = 20.59$
- $df = (Rows - 1)(Columns - 1) = 3$
- Pvalue `pchisq(20.59, 3, lower.tail = FALSE)` ≈ 0

Word distributions

| Word | Book | | | Total |
|----------------|------|-----|-----|-------|
| | 1 | 2 | 3 | |
| <i>a</i> | 147 | 186 | 101 | 434 |
| <i>an</i> | 25 | 26 | 11 | 62 |
| <i>this</i> | 32 | 39 | 15 | 86 |
| <i>that</i> | 94 | 105 | 37 | 236 |
| <i>with</i> | 59 | 74 | 28 | 161 |
| <i>without</i> | 18 | 10 | 10 | 38 |
| Total | 375 | 440 | 202 | 1017 |

- H_0 : The probabilities of each word are the same for every book
- H_a : At least two are different
- $O_{11} = 147$ $E_{11} = 375 \times \frac{434}{1017}$
- $O_{12} = 186$ $E_{12} = 440 \times \frac{434}{1017}$
- ...
- $\sum \frac{(O-E)^2}{E} = 12.27$
- $df = (6 - 1)(3 - 1) = 10$

Independence cont'd

- H_0 : H and W ratings are independent
- H_a : not independent
- $P(H = N \text{ \& } W = A) = P(H = N)P(W = A)$
- $stat = \sum \frac{(O-E)^2}{E}$
- $O_{11} = 7 \quad E_{11} = 91 \times \frac{19}{91} \times \frac{12}{91} = 2.51$
- $E_{ij} = n_{i+}n_{+j}/n$
- $df = (Rows - 1)(Cols - 1)$

Independence cont'd

```
x<-matrix(c(7,7,2,3,  
            2,8,3,7,  
            1,5,4,9,  
            2,8,9,14),4)
```

```
chisq.test(x)
```

- $\sum \frac{(O-E)^2}{E} = 16.96$
- $df = (4 - 1)(4 - 1) = 9$
- $p\text{-value} = .049$
- Cell counts might be too small to use large sample approximation

- Equal distribution and independence test yield the same results
- Same test results if
 - row totals are fixed
 - column totals are fixed
 - total ss is fixed
 - none are fixed
- Note that this is common in statistics; mathematically equivalent results are applied in different settings, but result in different interpretations

- Chi-squared result requires large cell counts
- df is always $(Rows - 1)(Columns - 1)$
- Generalizations of Fishers exact test can be used or continuity corrections can be employed

- Permute either the W or H row
 - Recalculate the contingency table
 - Calculate the χ^2 statistic for each permutation
 - Percentage of times it is larger than the observed value is an exact P-value
- ```
chisq.test(x, simulate.p.value = TRUE)
```

## Chi-squared goodness of fit

### Results from R's RNG

|       | [0, .25) | [.25, .5) | [.5, .75) | [.75, 1) | Total |
|-------|----------|-----------|-----------|----------|-------|
| Count | 254      | 235       | 267       | 244      | 1000  |
| TP    | .25      | .25       | .25       | .25      | 1     |

- $H_0 : p_1 = .25, p_2 = .25, p_3 = .25, p_4 = .25$
- $H_a : \text{any } p_i \neq \text{it's hypothesized value}$

## Continued

- $O_1 = 254$   $E_1 = 1000 \times .25 = 250$
- $O_2 = 235$   $E_2 = 1000 \times .25 = 250$
- $O_3 = 267$   $E_3 = 1000 \times .25 = 250$
- $O_4 = 244$   $E_4 = 1000 \times .25 = 250$
- $\sum \frac{(O-E)^2}{E} = 2.264$
- $df = 3$
- $P - value = .52$

## Testing Mendel's hypothesis

|          | Phenotype |         | Total |
|----------|-----------|---------|-------|
|          | Yellow    | Green   |       |
| Observed | 6022      | 2001    | 8023  |
| TP       | .75       | .25     | 1     |
| Expected | 6017.25   | 2005.75 | 8023  |

- $H_0 : p_1 = .75, p_2 = .25$
- $\sum \frac{(O-E)^2}{E} = \frac{(6022-6017.25)^2}{6017.25} + \frac{(2001-2005.75)^2}{2005.75} = .015$



## Continued

- $df = 1$
- P-value = .90
- Fisher combined several of Mendel's tables
- $\sum \chi^2_{v_i} \sim \chi^2_{\sum v_i}$
- Statistic 42,  $df = 84$ , P-value = .99996
- Agreement with theoretical counts is perhaps too good?

## Notes on GOF

- Test of whether or not observed counts equal theoretical values
- Test statistic is  $\sum \frac{(O-E)^2}{E}$
- TS follows  $\chi^2$  distribution for large  $n$
- $df$  is the number of cells minus 1
- Undirected alternative is problematic
- Especially useful for testing RNGs
- Kolmogorov/Smirnov test is an alternative test that does not require discretization but often has low power