

Problem 1

Let X and Y be i.i.d. $\text{Expo}(\lambda)$, and $T = \log(X/Y)$. Find the CDF and PDF of T .

$$f_X(x) = \lambda e^{-\lambda x} \quad f_Y(y) = \lambda e^{-\lambda y}$$

$$P(T \leq t) = P\left(\ln\left(\frac{X}{Y}\right) \leq t\right) = P\left(\frac{X}{Y} \leq e^t\right) = P(X \leq e^t Y)$$

$$P(X \leq e^t Y) = \int_0^\infty \int_0^{e^t y} f_X(x) f_Y(y) dx dy$$

$$= \int_0^\infty f_Y(y) \int_0^{e^t y} \lambda e^{-\lambda x} dx dy$$

$$= \int_0^\infty f_Y(y) \left(-e^{-\lambda x} \Big|_0^{e^t y} \right) dy$$

$$= \int_0^\infty f_Y(y) \left(-e^{-\lambda e^t y} + 1 \right) dy$$

$$= \int_0^\infty \lambda e^{-\lambda y} \left(1 - e^{-\lambda y e^t} \right) dy$$

$$= \int_0^\infty \lambda e^{-\lambda y} dy - \int_0^\infty \lambda e^{-\lambda y e^t - \lambda y} dy$$

PULL CONSTANT OUT

$$= 1 + \lambda \int_0^\infty e^{y(-\lambda e^t - \lambda)} dy$$

$$u = y(-\lambda e^t - \lambda)$$

$\rightarrow = 1$ SINCE THIS
IS THE PDF OF
 $\text{Expo}(\lambda)$

$$\begin{aligned}
 & du = (-\lambda e^t - \lambda) dy \\
 & = 1 + \frac{\lambda}{-\lambda e^t - \lambda} \int_0^\infty e^v dv \\
 & = 1 + \frac{\lambda}{-\lambda e^t - \lambda} \left(e^v \Big|_0^\infty \right) \\
 & = 1 + \frac{\lambda}{-\lambda e^t - \lambda} \left(e^{y(-\lambda e^t - \lambda)} \Big|_0^\infty \right) \\
 & = 1 + \frac{\lambda}{-\lambda e^t - \lambda} (0 - 1) \\
 & = 1 + \frac{-\lambda}{\lambda e^t + \lambda} = 1 + \frac{-\lambda}{\lambda(e^t + 1)} = 1 - \frac{1}{e^t + 1} \\
 & = \frac{e^t + 1}{e^t + 1} - \frac{1}{e^t + 1} = \frac{e^t}{e^t + 1}
 \end{aligned}$$

CDF: $F_T(t) = \frac{e^t}{e^t + 1}$ For $t \in \mathbb{R}$
 0 otherwise

PDF: $f_T(t) = F_T(t) \frac{d}{dt} = \frac{e^t}{e^t + 1} \frac{d}{dt}$

$$\begin{aligned}
 & = \frac{e^t(e^t + 1) - e^t e^t}{(e^t + 1)^2} = \frac{e^{2t} + e^t - e^{2t}}{(e^t + 1)^2} \\
 & = \frac{e^t}{(e^t + 1)^2} \text{ For } t \in \mathbb{R} \\
 & \quad 0 \text{ otherwise}
 \end{aligned}$$

$g(x) = e^t \quad g'(x) = e^t$
 $h(x) = e^t + 1 \quad h'(x) = e^t$

Problem 2

Let X and Y be i.i.d. $\text{Expo}(\lambda)$, and transform them to $T = X + Y$, $W = X/Y$.

(a) Find the joint PDF of T and W . Are they independent?

(b) Find the marginal PDFs of T and W .

$$a) f_{T,W}(t,w) = f_{X,Y}(x,y) \left| \frac{\partial(x,y)}{\partial(t,w)} \right|$$

$$W = \frac{X}{Y} \quad T = X + Y$$

$$X = WY \quad Y = T - X$$

$$Y = \frac{X}{W} \quad Y = T - X$$

$$T = WY + Y = Y(W+1)$$

$$\Rightarrow Y = \frac{T}{W+1}$$

$$\Rightarrow X = W \left(\frac{T}{W+1} \right) = \frac{WT}{W+1}$$

$$\frac{\partial x}{\partial t} = \frac{WT}{W+1} \frac{d}{dt} = \frac{W}{W+1}$$

$$\frac{\partial x}{\partial w} = T \left(\frac{W}{W+1} \right) \frac{d}{dw} = T \left(\frac{W+1-1}{(W+1)^2} \right) = \frac{T}{(W+1)^2}$$

$$\frac{\partial y}{\partial t} = \frac{T}{W+1} \frac{d}{dt} = \frac{1}{W+1}$$

$$\frac{\partial y}{\partial w} = \frac{T}{W+1} \frac{d}{dw} = \frac{-T}{(W+1)^2}$$

$$\left| \frac{\partial(x,y)}{\partial(t,w)} \right| = \begin{vmatrix} \frac{\partial x}{\partial t} & \frac{\partial x}{\partial w} \\ \frac{\partial y}{\partial t} & \frac{\partial y}{\partial w} \end{vmatrix} = \begin{vmatrix} \frac{W}{W+1} & \frac{T}{(W+1)^2} \\ \frac{1}{W+1} & \frac{-T}{(W+1)^2} \end{vmatrix}$$

$$= \begin{vmatrix} \frac{W}{W+1} \cdot \frac{-T}{(W+1)^2} - \frac{T}{(W+1)^2} \cdot \frac{1}{W+1} \\ \frac{-WT}{(W+1)^3} - \frac{T}{(W+1)} \end{vmatrix}$$

$$= \begin{vmatrix} \frac{-WT-T}{(W+1)^3} \\ \frac{-T(W+1)}{(W+1)^3} \end{vmatrix} = \begin{vmatrix} \frac{-T}{(W+1)^2} \\ \frac{-T}{(W+1)^2} \end{vmatrix} = \frac{T}{(W+1)^2}$$

$$f_{T,W}(t,w) = f_x(x)f_y(y) |\vec{J}|$$

$$= \lambda e^{-\lambda x} \lambda e^{-\lambda y} |\vec{J}|$$

$$= \lambda^2 e^{-\lambda(x+y)} |\vec{J}| = \begin{cases} \lambda^2 e^{-\lambda t} \frac{t}{(w+1)^2} & \text{for } 0 < T < \infty; \\ & 0 < TW < \infty \\ 0 & \text{OTHERWISE} \end{cases}$$

INDEPENDENCE TEST BELOW

$$(b) f_T(t) = \int_0^\infty \lambda^2 e^{-\lambda t} \frac{t}{(w+1)^2} dw = \lambda^2 e^{-\lambda t} t \int_0^\infty \frac{1}{(w+1)^2} dw$$

$$= \lambda^2 e^{-\lambda t} t \left(\frac{-1}{w+1} \Big|_0^\infty \right)$$

$$= \lambda^2 e^{-\lambda t} t (0 - (-1))$$

$$= \lambda^2 e^{-\lambda t} t \quad \text{for } t \geq 0; \quad 0 \text{ OTHERWISE}$$

$$f_W(w) = \int_0^\infty \lambda^2 e^{-\lambda t} \frac{t}{(w+1)} dt = \lambda^2 \frac{1}{w+1} \int_0^\infty t e^{-\lambda t} dt$$

$$= \frac{\lambda^2}{w+1} \left(\frac{e^{-\lambda t} (\lambda t + 1)}{\lambda^2} \Big|_0^\infty \right)$$

$$= \frac{\lambda^2}{w+1} \cdot \left(\frac{1}{\lambda^2} \right) = \begin{cases} \frac{1}{w+1} & \text{for } w \geq 0 \\ 0 & \text{OTHERWISE} \end{cases}$$

INDEPENDENCE if $f_{T,W}(t,w) = f_T(t) f_W(w)$

$$\lambda^2 e^{-\lambda t} \frac{t}{(w+1)^2} = \lambda^2 e^{-\lambda t} t \cdot \frac{1}{w+1} \quad \therefore \text{INDEPENDENT}$$

Problem 3

Let $U \sim \text{Unif}(0, 1)$ and $X \sim \text{Expo}(\lambda)$, independently. Find the PDF of $U + X$.

$$\text{Let } T = U + X$$

$$U = T - X$$

$$\text{THEOREM 8.2.1} \Rightarrow f_T(t) = \int_{-\infty}^{\infty} f_U(t-x) f_X(x) dx$$

$$f_U(u) = \frac{1}{1-0} = \begin{cases} 1 & \text{For } u \in (0, 1) \\ 0 & \text{OTHERWISE} \end{cases}$$

$$f_X(x) = \begin{cases} \lambda e^{-\lambda x} & \text{for } x > 0 \\ 0 & \text{OTHERWISE} \end{cases}$$

$$\int_0^t 1 \cdot \lambda e^{-\lambda x} dx = \left| -e^{-\lambda x} \right|_0^t$$

$$= 1 - e^{-\lambda t} - (1 - 1) = 1 - e^{-\lambda t} \quad t > 0$$

$$\begin{cases} 0 & \text{OTHERWISE} \end{cases}$$

Problem 4

Let X and Y be i.i.d. $\text{Expo}(\lambda)$. Use a convolution integral to show that the PDF of $L = X - Y$ is

$$f_L(l) = \frac{\lambda}{2} e^{-\lambda|l|},$$

for all real l .

$$L = X - Y$$

$$f_L(l) = \frac{\lambda}{2} e^{-\lambda|l|} \quad l \in \mathbb{R}$$

As $X \perp Y$, we can apply THEOREM 8.2.1:

$$f_L(l) = \int_0^{\infty} f_Y(x-l) f_X(x) dx$$

$$= \int_0^{\infty} \lambda e^{-\lambda(x-l)} \lambda e^{-\lambda x} dx$$

$$= \int_0^{\infty} \lambda^2 e^{-\lambda x + \lambda l - \lambda x} dx$$

$$= \lambda e^{-\lambda l} \int_0^{\infty} \lambda e^{-2\lambda x} dx$$

$$= \lambda e^{-\lambda l} \int_0^{\infty} \lambda e^{-2\lambda x} dx$$

$$= \frac{\lambda}{2} e^{-\lambda l} \int_0^{\infty} 2\lambda e^{-2\lambda x} dx$$

$$= \frac{\lambda}{2} e^{-\lambda l}$$

For

$$l > 0$$

$= \text{Expo}(2\lambda)$, which
integrates to 1

$$= \begin{cases} \frac{\lambda}{2} e^{-\lambda |l|} & \text{for } l \in \mathbb{R} \\ 0 & \text{otherwise} \end{cases}$$

ABSOLUTE
VALUE TO ACCOUNT
FOR ALL l , NOT
JUST POSITIVE

Problem 5

Use a convolution integral to show that if $X \sim N(\mu_1, \sigma^2)$ and $Y \sim N(\mu_2, \sigma^2)$ are independent, then

$$T = X + Y \sim N(\mu_1 + \mu_2, 2\sigma^2).$$

You can use a standardization (location-scale) idea to reduce to the standard Normal case before setting up the integral. Hint: complete the square.

$$f_T(t) = \int_{-\infty}^{\infty} f_Y(t-x) f_X(x) dx$$

FORMULA TO FIND THE
PDF OF $T = X + Y$ WHERE

FIRST, STANDARDIZE X AND Y SUCH THAT

$$V = \frac{X - \mu_1}{\sigma}, \quad W = \frac{Y - \mu_2}{\sigma}, \quad \text{AND} \quad Z = V + W$$

$$\text{THUS, } f_Z(z) = \int_{-\infty}^{\infty} f_V(z-v) f_W(v) dv$$

THE PDFs OF V AND W ARE BOTH THAT OF

$$\text{A STANDARD NORMAL: } \frac{1}{\sqrt{2\pi}} e^{-\frac{z^2}{2}}$$

PLUGGING IN TO f_Z GIVES:

$$f_Z(z) = \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}} e^{-\frac{v^2}{2}} \frac{1}{\sqrt{2\pi}} e^{-\frac{(z-v)^2}{2}} dv$$

PULL THE CONSTANT
OUT AND
COMBINE EXPONENTS

$$= \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-\frac{1}{2}(v^2 + (z-v)^2)} dv$$

COMPLETE THE
SQUARES FOR THE
EXPONENT

$$-\frac{1}{2}v^2 - \frac{1}{2}(z-v)^2 = -\frac{1}{2}v^2 - \frac{1}{2}(z^2 - 2zv + v^2)$$

$$= -\frac{1}{2}v^2 - \frac{1}{2}z^2 + zv - \frac{1}{2}v^2$$

$$= -v^2 + zv - \frac{1}{2}z^2$$

$$= -\left(v^2 - zv + \frac{1}{2}z^2\right)$$

$$= -\left(v^2 - zv + \frac{1}{4}z^2 - \frac{1}{4}z^2 + \frac{1}{2}z^2\right)$$

$$= -\left[\left(v - \frac{1}{2}z\right)^2 + \frac{1}{4}z^2\right]$$

$$= -\left(v - \frac{1}{2}z\right)^2 - \frac{1}{4}z^2$$

Plug in results:

$$f_z(z) = \frac{1}{2\pi} \int_{-\infty}^{\infty} e^{-\left(v - \frac{1}{2}z\right)^2 - \frac{1}{4}z^2} dv$$

pull out $e^{-\frac{1}{4}z^2}$
FROM THE INTEGRAL

$$= \frac{e^{-\frac{1}{4}z^2}}{2\pi} \int_{-\infty}^{\infty} e^{-\left(v - \frac{1}{2}z\right)^2} dv$$

$$u = v - \frac{1}{2}z \quad du = dv$$

$$= \frac{e^{-\frac{1}{4}z^2}}{2\sqrt{\pi}} \int_{-\infty}^{\infty} \frac{1}{\sqrt{2\pi}\left(\frac{\sqrt{2}}{2}\right)} e^{-\frac{1}{2\left(\frac{\sqrt{2}}{2}\right)^2}u^2} du$$

THIS IS
EQUIVALENT TO

$$N\left(0, \frac{\sqrt{2}}{2}\right)$$

AND INTEGRATES
TO 1.

$$= \frac{1}{2\sqrt{\pi}} e^{-\frac{z^2}{4}} \sim N(0, 2); \quad Z = V + W \sim N(0, 2)$$

UNDO THE LOCATION-SCALE TRANSFORMATION:

$$T = X + Y = \sigma V + \mu_1 + \sigma W + \mu_2 = \sigma(V + W) + (\mu_1 + \mu_2)$$

$$\text{Hence, } T = X + Y \sim N(\mu_1 + \mu_2, 2\sigma^2)$$

Problem 6

Let W_1 and W_2 be two random variables with the joint distribution:

$$P(W_1 \leq w_1, W_2 \leq w_2) = \int_{-\infty}^{w_1} \int_{-\infty}^{w_2} \frac{1}{2\pi} \exp\left[-\frac{1}{2}(x^2 + y^2)\right] dx dy.$$

Consider two other random variables $Z_1 = |W_1|$ and $Z_2 = |W_2|$. In words, Z_1 is the absolute value of W_1 , Z_2 is the absolute value of W_2 .

(a) Show that Z_1 is independent of Z_2 .

(b) Show that Z_1 and Z_2 have the same distribution, and find that distribution.

$$\begin{aligned} \text{(a)} \quad P(W_1 \leq w_1, W_2 \leq w_2) &= \int_{-\infty}^{w_1} \int_{-\infty}^{w_2} \frac{1}{2\pi} e^{-\frac{1}{2}(x^2 + y^2)} dx dy \\ &= \int_{-\infty}^{w_1} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy \int_{-\infty}^{w_2} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \end{aligned}$$

$$= P(W_1 \leq w_1) P(W_2 \leq w_2)$$

GIVEN THIS, WE KNOW
 W_1 AND W_2 ARE INDEPENDENT
STANDARD NORMALS

IN THE CASE OF Z_1 AND Z_2 :

$$\begin{aligned} P(Z_1 \leq w_1, Z_2 \leq w_2) &= P(-w_1 \leq Z_1 \leq w_1, -w_2 \leq Z_2 \leq w_2) \\ &= \int_{-w_1}^{w_1} \int_{-w_2}^{w_2} \frac{1}{\sqrt{2\pi}} e^{(-\frac{1}{2}(x^2 + y^2))} dx dy \\ &= \int_{-w_1}^{w_1} \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}} dy \int_{-w_2}^{w_2} \frac{1}{\sqrt{2\pi}} e^{-\frac{x^2}{2}} dx \end{aligned}$$

$$= P(-w_1 \leq Z_1 \leq w_1) P(-w_2 \leq Z_2 \leq w_2)$$

$$= P(Z_1 \leq w_1) P(Z_2 \leq w_2)$$

\therefore INDEPENDENCE OF Z_1 AND Z_2

(b) Z_1 AND Z_2 ARE IDENTICAL EXCEPT FOR THEIR USE OF x OR y SO WE CAN JUST SOLVE THE PDF OF ONE TO GET THE OTHER:

$$P(Z_1 \leq w_1) = P(-w_1 \leq W_1 \leq w_1) = P(W_1 \leq w_1) - P(W_1 \leq -w_1)$$

$$f_{Z_1}(y) = f_{W_1}(y) + f_{W_1}(-y)$$

$$= 2 f_{W_1}(y)$$

DUE TO SYMMETRY OF N DIST

$$= 2 \left(\frac{1}{\sqrt{2\pi}} e^{-y^2/2} \right)$$

$$= \begin{cases} \sqrt{\frac{2}{\pi}} e^{-y^2/2} & \text{For } y \geq 0 \\ 0 & \text{OTHERWISE} \end{cases}$$

AND, SINCE THEY ARE IDENTICAL:

$$f_{Z_2}(x) = \begin{cases} \sqrt{\frac{2}{\pi}} e^{-x^2/2} & \text{For } x \geq 0 \\ 0 & \text{OTHERWISE} \end{cases}$$