

AMAT 362 - PROBABILITY for STATISTICS

LECTURE 23

: **CONDITIONAL DISTRIBUTIONS**
for DISCRETE RV's
+ Intro to BAYESIAN INFERENCE

Law of Iterated Expectations

$$E[E[X|Y]] = E[X]$$

Q: Suppose I roll

a 6-sided fair die

$$\Rightarrow X \sim \text{Unif}\{1, \dots, 6\}$$

Now, based on outcome of the roll
flip a fair coin that # of times
denote the # of heads by Y .

$$P(Y=y \text{ heads} | X=x \text{ on the die roll}) = \binom{x}{y} \underbrace{\left(\frac{1}{2}\right)^y \left(\frac{1}{2}\right)^{x-y}}_{\text{CONDITIONAL DISTRIBUTION}} = \left(\frac{1}{2}\right)^x = 2^{-x}$$

AKA

$$P(Y=y | X=x) = \binom{x}{y} 2^{-x}$$

Q: What is the (unconditional) PMF for Y ?
FOR Y GIVEN X

$$P(Y=y) = \sum_x P(Y=y | X=x) P(X=x)$$

LAW OF TOTAL PROBABILITY / RULE OF AVERAGED CONDITIONAL

Know that $X \sim \text{Unif}\{1, \dots, 6\}$

$$P(Y=y) = \sum_{x=1}^6 P(Y=y | X=x) P(X=x)$$

$$= \binom{x}{y} 2^{-x} = \frac{1}{6}$$

$$= \frac{1}{6} \sum_{x=1}^6 \binom{x}{y} \frac{1}{2^x}$$

$$P(Y=0) = \frac{1}{6} \cdot \frac{1}{2} + \frac{1}{6} \left(\frac{1}{2}\right)^2 + \frac{1}{6} \left(\frac{1}{2}\right)^3 + \dots + \frac{1}{6} \cdot \left(\frac{1}{2}\right)^6$$

$\uparrow \quad \uparrow$
 $P(X=1) \quad P(\text{no heads} | X=1)$

$$P(Y=0) = \frac{1}{6} \cdot \frac{63}{64}$$

IN GENERAL ... From Pitman 6.1

TABLE 1. Probability $P(Y = y)$ of getting y heads.

y	0	1	2	3	4	5	6
$P(Y = y)$	$\frac{63}{384}$	$\frac{120}{384}$	$\frac{99}{384}$	$\frac{64}{384}$	$\frac{29}{384}$	$\frac{8}{384}$	$\frac{1}{384}$

This essentially Marginalization
of the joint PMF $P(X=x, Y=y)$

Joint PMF for $P(\text{dia roll} = x, \# \text{heads} = y)$

TABLE 2. Joint distribution table for (X, Y) .

		Possible values x for X						Marginal distn. of Y
		1	2	3	4	5	6	
Possible values y for Y	0	$\frac{1}{6} \frac{1}{2}$	$\frac{1}{6} \frac{1}{4}$	$\frac{1}{6} \frac{1}{8}$	$\frac{1}{6} \frac{1}{16}$	$\frac{1}{6} \frac{1}{32}$	$\frac{1}{6} \frac{1}{64}$	$\frac{63}{384}$
	1	$\frac{1}{6} \frac{1}{2}$	$\frac{1}{6} \frac{1}{4}$	$\frac{1}{6} \frac{3}{8}$	$\frac{1}{6} \frac{4}{16}$	$\frac{1}{6} \frac{5}{32}$	$\frac{1}{6} \frac{6}{64}$	$\frac{120}{384}$
	2	0	$\frac{1}{6} \frac{1}{4}$	$\frac{1}{6} \frac{3}{8}$	$\frac{1}{6} \frac{6}{16}$	$\frac{1}{6} \frac{10}{32}$	$\frac{1}{6} \frac{15}{64}$	$\frac{99}{384}$
	3	0	0	$\frac{1}{6} \frac{1}{8}$	$\frac{1}{6} \frac{4}{16}$	$\frac{1}{6} \frac{10}{32}$	$\frac{1}{6} \frac{20}{64}$	$\frac{64}{384}$
	4	0	0	0	$\frac{1}{6} \frac{1}{16}$	$\frac{1}{6} \frac{1}{32}$	$\frac{1}{6} \frac{6}{64}$	$\frac{29}{384}$
	5	0	0	0	0	$\frac{1}{6} \frac{1}{32}$	$\frac{1}{6} \frac{6}{64}$	$\frac{8}{384}$
	6	0	0	0	0	0	$\frac{1}{6} \frac{1}{64}$	$\frac{1}{384}$
	Marginal distn. of X	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$	$\frac{1}{6}$	1

$\rightarrow P(Y=0)$

RULE OF AVERAGED CONDITIONALS

$$P(Y=y) = \sum_x P(Y=y | X=x) P(X=x)$$

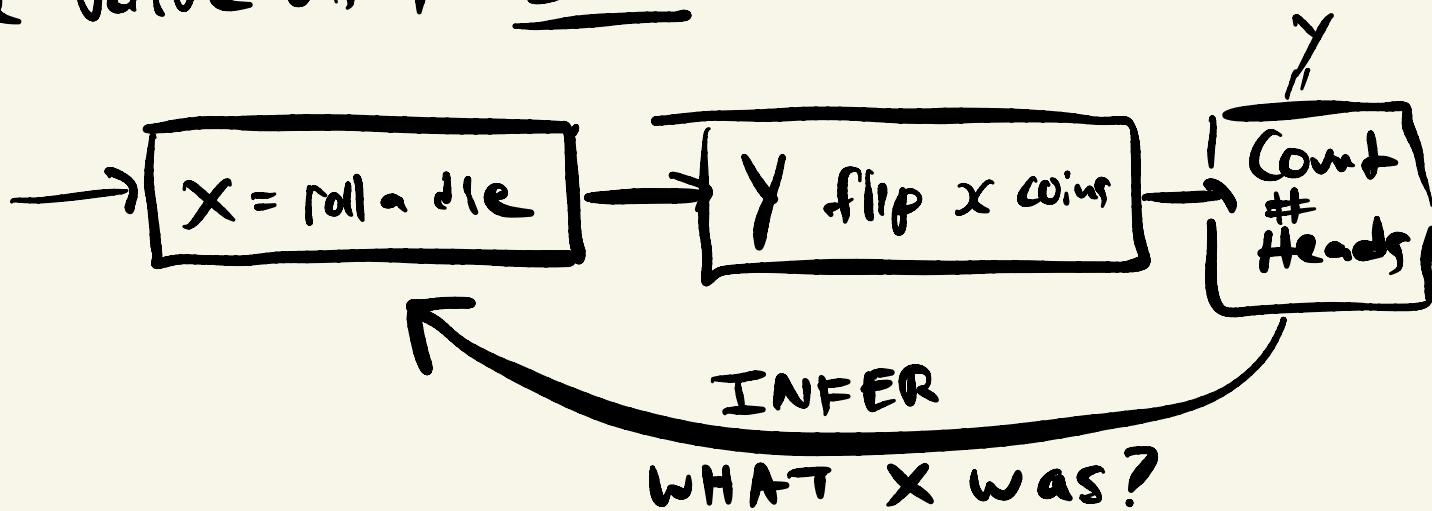
$$= \sum_x \frac{P(Y=y, X=x)}{P(X=x)} \cdot P(X=x)$$

RULE
for Marginalization

$$= \sum_x P(Y=y, X=x)$$

EXAMPLE OF BAYES' RULE

Suppose we've told the result of this two step experiment, i.e. we've told the value of Y BUT we don't know X



Q: Find $P(X=x | Y=y)$

value of roll # heads observed

A: BAYES RULE

$$P(X=x | Y=y) = \frac{P(Y=y | X=x) P(X=x)}{P(Y=y)}$$

Say... $y=2$... $P(X=1 | Y=2) = 0$ B/c had 2 heads $\Rightarrow x \geq 2$

$$P(X=2 | Y=2) = \binom{2}{2} \left(\frac{1}{2}\right)^2 \cdot \frac{1}{6}$$

$$\sum_{x=2}^6 \binom{x}{2} \left(\frac{1}{2}\right)^2 \left(\frac{1}{2}\right)^{x-2} \cdot \frac{1}{6} = P(Y=2)$$

$$= \frac{99}{384}$$

$$P(X=2 | Y=2) = \frac{16}{99}$$

Repeating this calculation for $x = 3, 4, 5, 6$
 conditionally $Y = 2$

TABLE 3. Conditional distribution of X given $Y = 2$.

x	1	2	3	4	5	6
$P(X = x Y = 2)$	0	$\frac{16}{99}$	$\frac{24}{99}$	$\frac{24}{99}$	$\frac{20}{99}$	$\frac{15}{99}$

STATISTICAL ASIDE

The principle of Maximum Likelihood

(Maximum Likelihood Estimator = MLE)

Says that you should assume the value of x that makes $Y=2$ maximally likely.

i.e. $x=3$ or 4 is the most likely explanation for the data.

OCCAM'S RAZOR



BUT OTHER STATISTICS OPTIMIZE DIFFERENT REWARD SCHEMES, i.e. MEDIAN { MEAN

For this example instead of guessing

$x=3$ or 4 given $Y=2$

$$E(X | Y=2) = \sum x P(X=x | Y=2)$$

could be another choice.

EXPECTATION OF CONDITIONAL

(LAW OF ITERATED EXPECTATIONS) RVS

Q: Suppose we want to know $E(Y)$, but we know Y is conditioned on X , so ..

$$E(Y|X=x) = \sum_y y P(Y=y|X=x)$$

Binomial
R.V. w/
 $n=x$
 $p=1/2$

$$= np$$

$$= x \cdot \frac{1}{2}$$

$$E(Y) = \sum E(Y|X=x) \cdot P(X=x)$$

This is just
a function of
RV X ? so. $E[g(x)]$

$$= \sum g(x) P(X=x)$$

$$= \sum_{x=1}^6 \left(\frac{x}{2}\right) \frac{1}{2}$$

$$= \frac{1}{2} \sum_{x=1}^6 x \cdot P(X=x)$$

$$E(X) = \frac{7}{2} = 3.5$$

$$E(Y) = \frac{E(X)}{2} = \frac{7}{4} = 1.75$$

LAW OF ITERATED EXPECTATION

$$E[E[\underline{Y|X}]] = E[Y]$$

that is dependent
on value x of X
i.e. $g(x)$

Pf

$$\begin{aligned} E[E[Y|X]] &= \sum_x E[Y|X=x] P(X=x) \\ &= \sum_x \sum_y y P(Y=y|X=x) P(X=x) \\ &= \sum_x \sum_y y \frac{P(X=x, Y=y)}{P(X=x)} P(X=x) \\ &\stackrel{\text{Reorder}}{\hookrightarrow} = \sum_y \sum_x y P(X=x, Y=y) \\ &= \sum_y y \underbrace{P(Y=y)}_{\text{MARGINALIZES}} \\ &= E[Y] \quad \blacksquare \end{aligned}$$

Turn this result on its head...

Recall this "simpler" problem, but hard to calculate

$$E[X|Y=2] = \sum x P(X=x | Y=2)$$

= ...

OK, so what is ...

$$E[E[X|Y]] = E[X] = \frac{7}{2}$$

function
of y

WHAT IS GOING ON HERE?

$$P(Y=n_H | X=N, P=\frac{1}{2}) = \binom{N}{n_H} P^{n_H} q^{N-n_H}$$

Now fix $n_H=2$, $P=\frac{1}{2}$, let X vary...

$$P(X=x | n_H=2, P=\frac{1}{2}) = \left[\binom{x}{2} P^x q^{x-2} \right] \cdot \frac{1}{6}$$

Computing
total probability
of $n_H=2$

$$\sum_{x=2}^6 \left(\binom{x}{2} P^x q^{x-2} \cdot \frac{1}{6} \right)$$

FUNDAMENTAL PARADIGM for BAYESIAN STATS

$$\text{POSTERIOR PROBABILITY} = \frac{\text{Likelihood} \times \text{Prior Probability}}{P(\text{Evidence})}$$