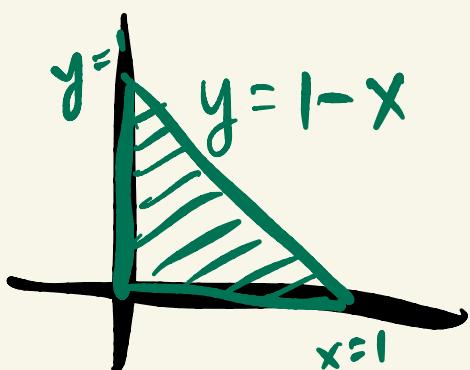


LECTURE 24 LIVE

Q1 $E[XY] = \dots$



$$E[g(x,y)] = \iint g(x,y) f(x,y) dx dy$$

$$= \int_{x=0}^{x=1} \int_{y=0}^{y=1-x} xy^2 dy dx$$

$$= 2 \int_0^1 x \int_{y=0}^{y=1-x} y dy dx = 2 \int_0^1 x \left[\frac{y^2}{2} \right]_{y=0}^{y=1-x} dx$$

$$= 2 \int_0^1 x \frac{(1-x)^2}{2} dx$$

$$= 2 \int_0^1 x (1-2x+x^2) dx$$

$$= \int_0^1 x - 2x^2 + x^3 dx = \left. \frac{x^2}{2} - \frac{2x^3}{3} + \frac{x^4}{4} \right|_0^1$$

$$\frac{3}{4} - \frac{2}{3} + \frac{1}{12}$$

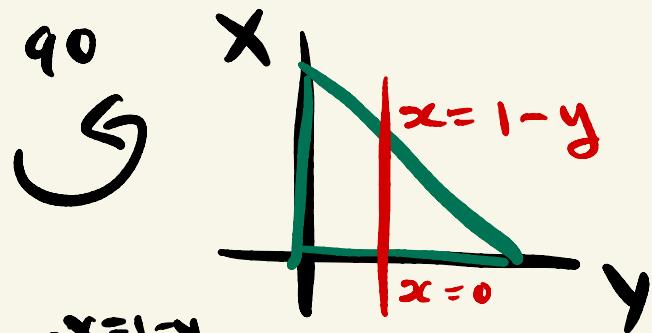
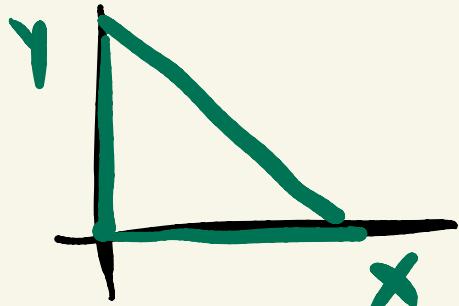
$$= \frac{1}{2} - \frac{2}{3} + \frac{1}{4}$$

$$\text{Q1(b)} \quad \text{Cov}(x, y) = E[xy] - E[x]E[y]$$

WHY IS THIS IMPORTANT?

$$\text{VAR}(x+y) = \text{VAR}(x) + 2\text{cov}(x,y) + \text{VAR}(y)$$

$$\begin{aligned}
 E[x] &= \int x f_X(x) dx = \int_0^1 x^2 dy dx \\
 &= 2 \int_0^1 x [y]_{y=0}^{y=1-x} dx = 2 \int_0^1 x(1-x) dx \\
 &= 2 \int_0^1 x - x^2 dx = 2 \left[\frac{x^2}{2} - \frac{x^3}{3} \right]_0^1 = 2 \left(\frac{1}{2} - \frac{1}{3} \right) \\
 &= \frac{2}{6} \checkmark
 \end{aligned}$$



$$E[y] = \int_{y=0}^{y=1} \int_{x=0}^{x=y} y^2 dx dy = \frac{2}{6} \text{ By Symmetry}$$

$$\text{Cov}(x_1, y) = \frac{1}{12} - \left(\frac{2}{6}\right)^2 = \frac{1}{3}$$

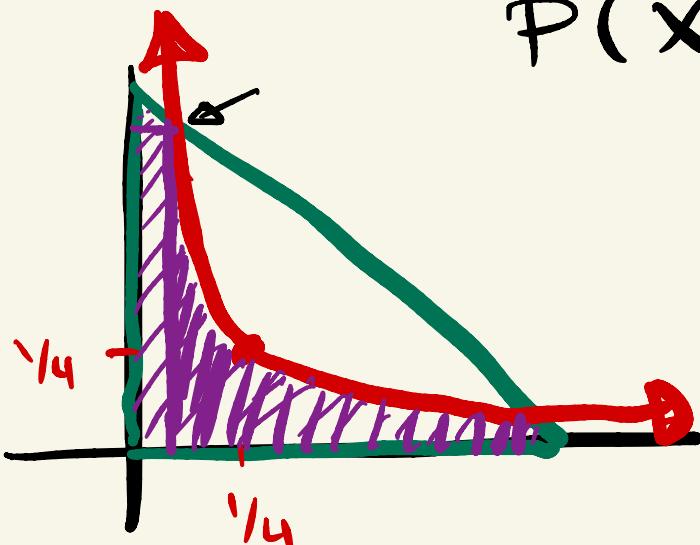
$$-\frac{1}{36}x^2 = \frac{1}{12} - \frac{1}{9} < 0$$

Q1(c)

Compute

$$P(\text{Area} = \frac{1}{16})$$

$$P(XY < \frac{1}{16})$$



$$y < \frac{1}{16x}$$

STEP 1 : DETERMINE
LIMITS
of Integration

$$\begin{aligned} xy &= \frac{1}{16} \\ y &= 1-x \end{aligned} \rightarrow x(1-x) = \frac{1}{16}$$

$$16x - 16x^2 = 1$$

$$16x^2 - 16x + 1 = 0$$

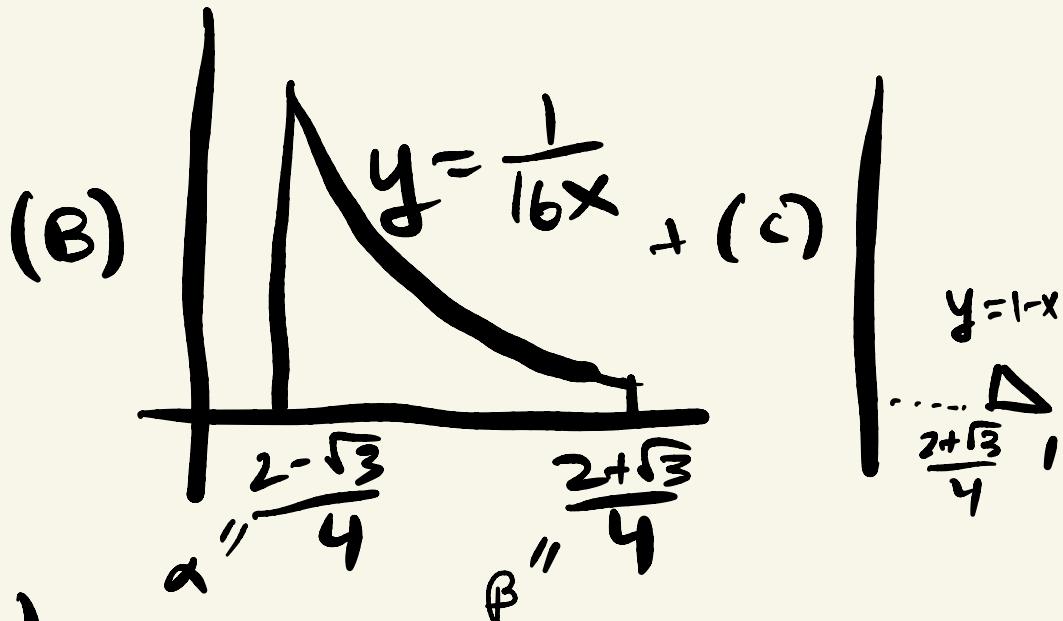
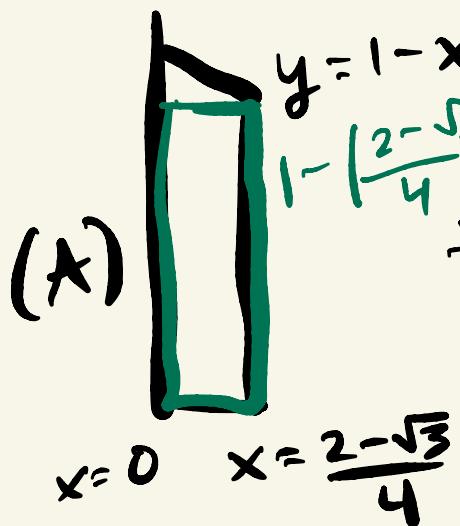
$$x^2 - x + \frac{1}{16}$$

$$ax^2 + bx + c = 0$$

$$x = \frac{1}{2} \pm \frac{\sqrt{1 - \frac{1}{4}}}{2} = \frac{\sqrt{3}}{2} = \frac{\sqrt{3}}{4}$$

$$x = \frac{2 \pm \sqrt{3}}{4}$$

3 Integrals to do



Integral (B)

$$\int_{\alpha}^{\beta} \frac{1}{16x} dx = \frac{1}{16} \ln x \Big|_{\alpha}^{\beta} = \frac{1}{16} [\ln \beta - \ln \alpha]$$

$$\frac{1}{16} \ln \left(\frac{2 + \sqrt{3}}{2 - \sqrt{3}} \right)$$

$$= \frac{1}{16} \ln \left(\frac{\beta}{\alpha} \right)$$

+ FINAL ANSWER

Integral (A) + Integral (C)

$$= \frac{2 - \sqrt{3}}{4}$$

$$= \boxed{1} \quad \text{base} = \frac{2 - \sqrt{3}}{4}$$

② Show up at Nikes to work @ 10am

Customers show up $\sim \text{Exp}(\lambda = \frac{1}{10})$

X = Service time $X \sim \text{Unif}(1, 6)$

Q: When do I expect to be done serving my first customer?

$$\rightarrow E[Y+X] = E[Y] + E[X]$$

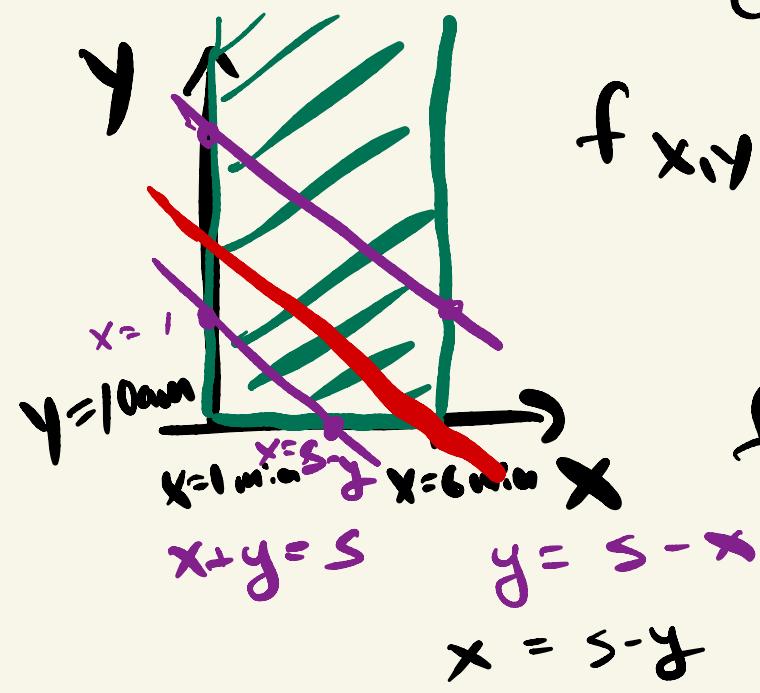
CAVEAT LECTOR

I drew this diagram wrong in class. So skip this problem

$$= 10:10\text{am} + \frac{6-1}{2}$$

$$= 10:12\text{am } 30\text{sec s}$$

OK Find Density of $Y+X = S$



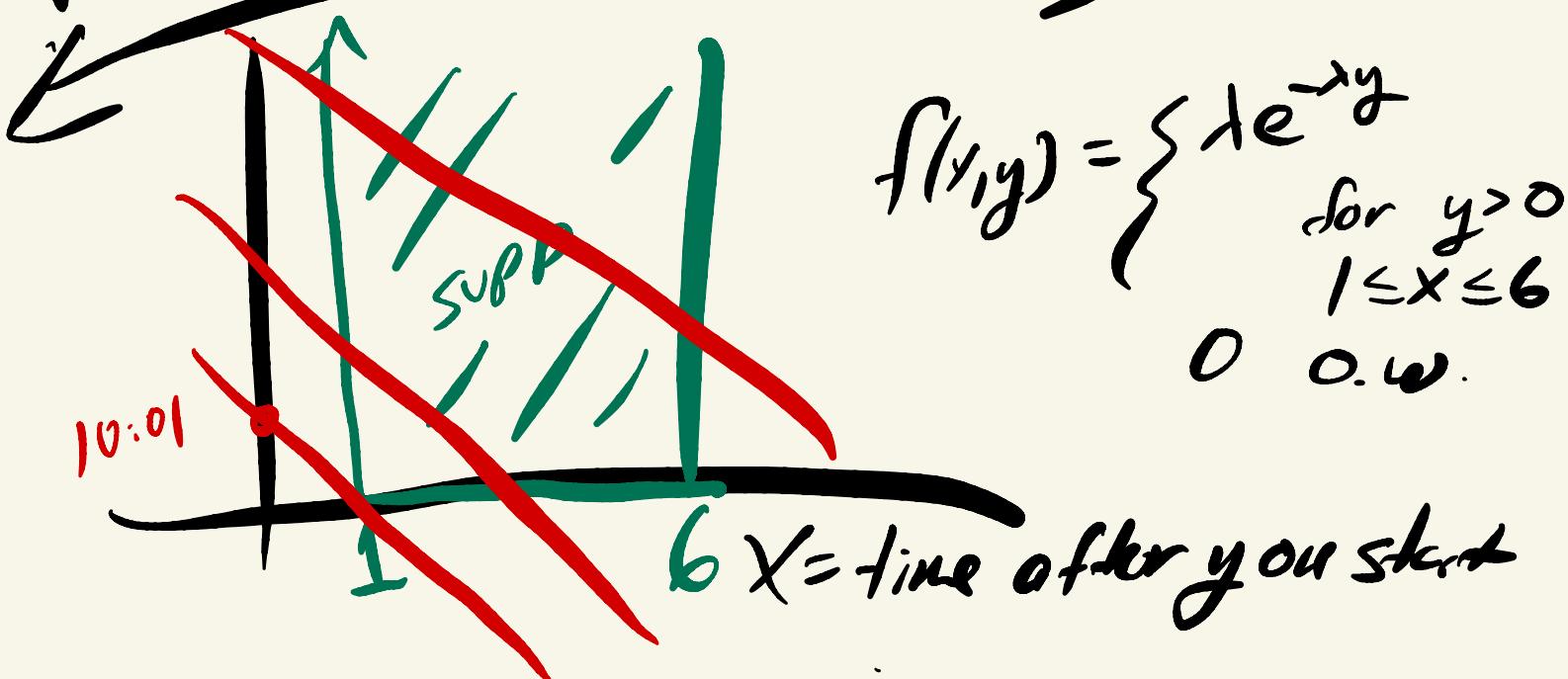
$$f_{X,Y} = \begin{cases} \frac{1}{10} e^{-y/10} & y > 10\text{am} \\ & 1 \leq x \leq 6 \text{ mins} \\ 0 & \text{o.w.} \end{cases}$$

$$f_S(s) = \int_{x=1}^{x=s-y} \frac{1}{10} e^{-y/10} dx$$

$$= \frac{1}{10} e^{-y/10} (s-y-1)$$

DO THIS LATER

Problem 2 - Right setup



"Don't worry about this problem
 We'll do something similar for
 the practice final on Mon. May 10."
 - Dr. Cus

③ $X \sim \text{Unif}(0,1)$

$$P(A | X=x) = x^2$$

Q: What is $P(A)$?

(Rule of Averaged Conditionals
(aka Law of Total Probability))

$$P(A) = \int P(A | X=x) f_X(x) dx$$

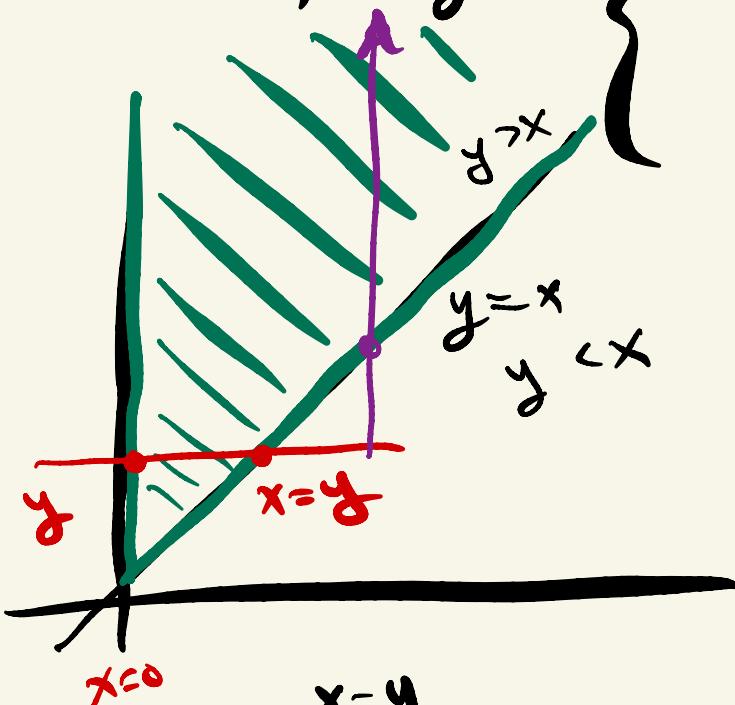
$\underbrace{f_X(x)}$
 $\mathbb{1}_{(0,1)}^{''}$ indicator function

$$= \int_0^1 P(A | X=x) dx$$

$$= \int_0^1 x^2 dx = \frac{x^3}{3} \Big|_0^1 = \boxed{\frac{1}{3}}$$

Q4

$$f(x,y) = \begin{cases} \lambda^3 x e^{-\lambda y} & 0 < x < y \\ 0 & \text{otherwise} \end{cases}$$



(a) PDF for Y

$$f_Y(y) = \int_{x=-\infty}^{x=\infty} f(x,y) dx$$

$$f_Y(y) = \int_{x=0}^{x=y} \lambda^3 x e^{-\lambda y} dx = \lambda^3 e^{-\lambda y} \left[\frac{x^2}{2} \right]_{x=0}^{x=y}$$

Gamma(3, λ)
Distribution

$$f_Y(y) = \frac{\lambda^3 y^2 e^{-\lambda y}}{2} \quad \text{for } 0 < y$$

So $E[Y] = \frac{3}{\lambda}$ ← IGNORE THE REST

JUST LOOK $E[\text{Gamma}]$ UP!

$$(b) E[Y] = \int y f_Y(y) dy = \int_{y=0}^{y=\infty} \frac{\lambda^3 y^3}{2} e^{-\lambda y} dy$$

ALTERNATIVELY

$$E[Y] = E[E[Y|X]]$$

So bottom of next page!
Integrate by Parts 3x! YUCK



$$E[Y|X=x] = \int y f_{Y|X=x}(y|x) dy$$

SKETCH THIS

$$f_{Y|X=x} = \frac{f(x,y)}{f(x)} = \frac{\lambda^3 x e^{-\lambda y}}{\lambda^2 x e^{-\lambda x}} = \lambda^{x-y}$$

$$f_X(x) = \int_{y=x}^{\infty} \lambda^3 x e^{-\lambda y} dy = \lambda^2 x \underbrace{\int_x^{\infty} e^{-\lambda y} dy}_{x}$$

$$f_X(x) = \lambda^2 x e^{-\lambda x} \quad 0 < x < \infty$$

SKETCH THIS

$$E[Y|X=x] = \int y \lambda e^{x-\lambda y} dy$$

ERR ... THIS
ISN'T MUCH
EASIER.

$$= e^{\lambda x} \int y \lambda e^{-\lambda y} dy$$

JUST USE \star
THIS
 $r=3$ so $E = \frac{3}{\lambda}$

Poisson Arrival Times (Gamma Distribution)

If T_r is the time of the r th arrival after time 0 in a Poisson process with rate λ , or if $T_r = W_1 + W_2 + \dots + W_r$ where the W_i are independent with exponential (λ) distribution, then T_r has the gamma (r, λ) distribution defined by either (1) or (2) for all $t \geq 0$:

GOT TO RECOGNIZE

(1) **Density:** $P(T_r \in dt) / dt = P(N_t = r-1) \lambda = e^{-\lambda t} \frac{(\lambda t)^{r-1}}{(r-1)!} \lambda$

where N_t , the number of arrivals by time t in the Poisson process with rate λ , has Poisson (λt) distribution. In words, the probability per unit time that the r th arrival comes around time t is the probability of exactly $r-1$ arrivals by time t multiplied by the arrival rate.

(2) **Right tail probability:** $P(T_r > t) = P(N_t \leq r-1) = \sum_{k=0}^{r-1} e^{-\lambda t} \frac{(\lambda t)^k}{k!}$

because $T_r > t$ if and only if there are at most $r-1$ arrivals in the interval $(0, t]$.

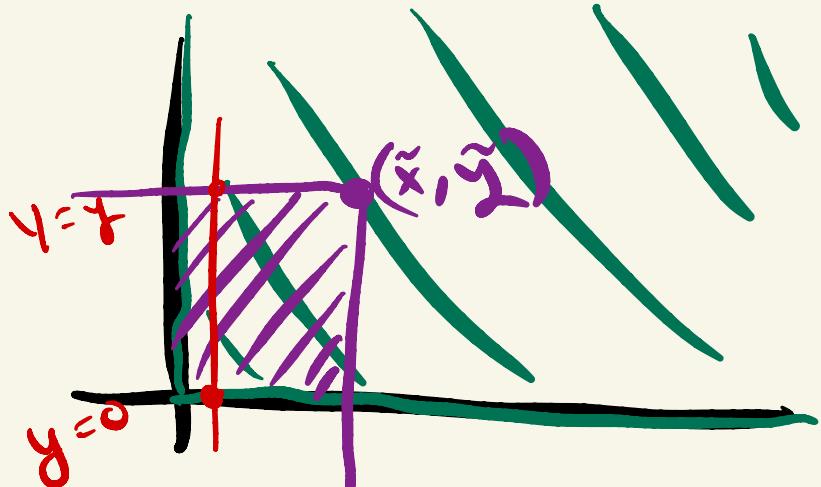
(3) **Mean and SD:** $E(T_r) = r/\lambda \quad SD(T_r) = \sqrt{r}/\lambda$

WS 22

$$f(x, y) = f_x(x) f_y(y)$$

Q6

$$f(x, y) = \frac{6e^{-2x-3y}}{\text{||}} \quad x, y \geq 0$$



$$= (2e^{-2x})(3e^{-3y})$$

$$\approx (1 - e^{-2\tilde{x}})(1 - e^{-3\tilde{y}})$$

GUESS...

$$\int_{x=0}^{x=\tilde{x}} \int_{y=0}^{y=\tilde{y}} (2e^{-2x})(3e^{-3y}) dy dx$$

$$= \int_{x=0}^{x=\tilde{x}} 2e^{-2x} \left[-e^{-3y} \right]_{y=0}^{y=\tilde{y}} dx = (1 - e^{-3\tilde{y}}) \int_{x=0}^{\tilde{x}} 2e^{-2x} dx \\ = (1 - e^{-3\tilde{y}}) \left(\frac{2}{2} e^{-2x} \right)_{x=0}^{\tilde{x}} = (1 - e^{-3\tilde{y}}) (1 - e^{-2\tilde{x}})$$

$$P(X \leq \tilde{x}, Y \leq \tilde{y}) = F_x(\tilde{x}) F_y(\tilde{y})$$

↑
BC Independent
 $X \nmid Y$