

# Close Elections, Campaign Contributions, and Financial Deregulation

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## Abstract

This paper builds upon Igan and Mishra (2014) Igan and Mishra (2014) on vote switches towards financial deregulation by US legislators. I measure the effect of close elections on US legislators on switching their votes towards financial deregulation in Congress bills. I aim to distinguish between vote switches towards financial deregulation because of voters' general interests (especially after the Global Financial Crisis of 2007) versus the financial industry's special interests in the form of campaign contributions and lobbying expenditures towards legislators in close elections.

**Keywords** Financial Deregulation, Close Elections

**JEL codes** D81, D91, E21

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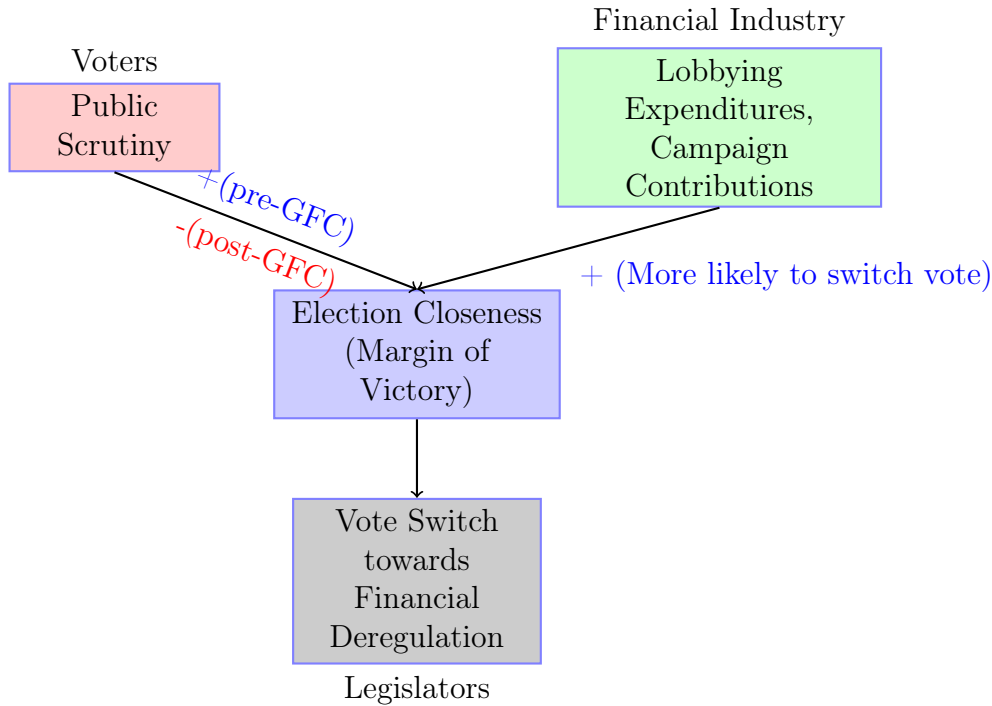
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The dashboard lets users see consequences of alternative parameters in an interactive framework.

The paper's results can be automatically reproduced using the Econ-ARK/HARK toolkit, which can be cited per our references (Carroll, Kaufman, Kazil, Palmer, and White (2018)); for reference to the toolkit itself see Acknowledging Econ-ARK. Thanks to the Consumer Financial Protection Bureau for funding the original creation of the Econ-ARK toolkit; and to the Sloan Foundation for funding Econ-ARK's extensive further development that brought it to the point where it could be used for this project. The toolkit can be cited with its digital object identifier, 10.5281/zenodo.1001067, as is done in the paper's own references as Carroll, Kaufman, Kazil, Palmer, and White (2018). Thanks to Will Du, James Feigenbaum, Joseph Kaboski, Miles Kimball, Qingyin Ma, Misuzu Otsuka, Damiano Sandri, John Stachurski, Adam Szeidl, Alexis Akira Toda, Metin Uyanik, Mateo Velásquez-Giraldo, Weifeng Wu, Jiaxiong Yao, and Xudong Zheng for comments on earlier versions of this paper, John Boyd for help in applying his weighted contraction mapping theorem, Ryoji Hiraguchi for extraordinary mathematical insight that improved the paper greatly, David Zervos for early guidance to the literature, and participants in a seminar at the Johns Hopkins University, a presentation at the 2009 meetings of the Society of Economic Dynamics for their insights, and at a presentation at the Australian National University.

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**Figure 1** Two Channels in which Legislators in Close Elections Switch their Votes Toward Financial Deregulation

## 1 Introduction

Sample Text for ProjectKWK: Lorem Ipsum et cetera. Try citing Igan and Mishra (2014) Igan and Mishra (2014).

The next theoretical contribution demonstrates that a corresponding model with an ‘artificial’ liquidity constraint (a model that prohibits borrowing by consumers who could certainly repay) is a limiting case of the model without constraints. The analytical appeal of the unconstrained model is that it is both mathematically convenient (e.g., the consumption function is twice continuously differentiable), and arbitrarily close (cf. Section 2.10) to less tractable models. The congenial environment makes the proof easier, and we define the analogous proposition as holding (in the limit) if it continues to hold as the horizon extends to infinity.

In proving the remaining theorems, the **next section** examines the key properties of the model. First, as **cash approaches infinity** the expected growth rate of consumption and the marginal propensity to consume (MPC) converge to their values in the perfect foresight case. Second, as **cash approaches zero** the expected growth rate of consumption approaches infinity, and the MPC approaches a simple analytical limit. Next, the central theorems articulate conditions under which different measures of ‘growth impatience’ imply useful conclusions about points of stability (‘target’ or ‘balanced growth’ points).

The final section elaborates the conditions under which, even with a fixed aggregate interest rate that differs from the time preference rate, a small open economy populated by buffer stock consumers has a balanced growth equilibrium in which growth rates of consumption, income, and wealth match the exogenous growth rate of permanent income (equivalent, here, to productivity growth). In the terms of Schmitt-Grohé and Uribe (2003), buffer stock saving is an appealing method of ‘closing’ a small open economy model, because it requires no ad-hoc assumptions. Not even liquidity constraints.<sup>1</sup>

## 2 The Problem

### 2.1 Setup

The infinite horizon solution is the (limiting) first-period solution to a sequence of finite-horizon problems as the horizon (the last period of life) becomes arbitrarily distant.

That is, for the value function, fixing a terminal date  $T$ , we are interested in the term  $\mathbf{v}_{T-n}$  in the sequence of value functions  $\{\mathbf{v}_T, \mathbf{v}_{T-1}, \dots, \mathbf{v}_{T-n}\}$ . We will say that the problem has a ‘nondegenerate’ infinite horizon solution if, corresponding to that  $\mathbf{v}$ , as  $n \uparrow \infty$  there is a limiting consumption function  $c(m) = \lim_{n \uparrow \infty} c_{T-n}$  which is neither  $c(m) = 0$  everywhere (for all  $m$ ) nor  $c(m) = \infty$  everywhere.

Concretely, a consumer born  $n$  periods before date  $T$  solves the problem

$$\mathbf{v}_{T-n} = \max \mathbb{E}_t \left[ \sum_{i=0}^n \beta^i u(\mathbf{c}_{t+i}) \right]$$

where the Constant Relative Risk Aversion (CRRA) utility function

$$u(\bullet) = \bullet^{1-\rho} / (1-\rho) \quad (1)$$

exhibits relative risk aversion  $\rho > 1$ .<sup>2</sup> The consumer’s initial condition is defined by market resources  $\mathbf{m}_t$  and permanent noncapital income  $\mathbf{p}_t$ , which both are positive,

$$\{\mathbf{p}_t, \mathbf{m}_t\} \in (0, \infty), \quad (2)$$

and the consumer cannot die in debt,

$$\mathbf{c}_T \leq \mathbf{m}_T. \quad (3)$$

In the usual treatment, a dynamic budget constraint (DBC) incorporates several elements that jointly determine next period’s  $\mathbf{m}$  (given this period’s choices); for the detailed analysis here, it will be useful to disarticulate and describe every step:

$$\begin{aligned} \mathbf{a}_t &= \mathbf{m}_t - \mathbf{c}_t \\ \mathbf{k}_{t+1} &= \mathbf{a}_t \\ \mathbf{b}_{t+1} &= \mathbf{k}_t R \end{aligned} \quad (4)$$

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<sup>1</sup>The paper’s insights are instantiated in the Econ-ARK toolkit, whose `buffer stock saving` module flags parametric choices under which a problem is degenerate or under which stable ratios of wealth to income may not exist.

<sup>2</sup>The main results also hold for logarithmic utility which is the limit as  $\rho \rightarrow 1$  but incorporating the logarithmic special case in the proofs is omitted because it would be cumbersome.

$$\begin{aligned}\mathbf{p}_{t+1} &= \mathbf{p}_t \underbrace{\Gamma \psi_{t+1}}_{\equiv \Gamma_{t+1}} \\ \mathbf{m}_{t+1} &= \mathbf{b}_{t+1} + \mathbf{p}_{t+1} \xi_{t+1},\end{aligned}$$

where  $\mathbf{a}_t$  indicates the consumer's assets at the end of period  $t$ , which translate one-for-one into capital  $\mathbf{k}_{t+1}$  at the beginning of the next period, which (before the consumption choice) grows by a fixed interest factor  $\mathbf{R} = (1+r)$ , so that  $\mathbf{b}_{t+1}$  is the consumer's financial ('bank') balances before next period's consumption choice;<sup>3</sup>  $\mathbf{m}_{t+1}$  ('market resources') is the sum of financial wealth  $\mathbf{b}_{t+1}$  and noncapital income  $\mathbf{p}_{t+1}\xi_{t+1}$  (permanent noncapital income  $\mathbf{p}_{t+1}$  multiplied by a mean-one iid transitory income shock factor  $\xi_{t+1}$ ; transitory shocks are assumed to satisfy  $\mathbb{E}_t[\xi_{t+n}] = 1 \ \forall n \geq 1$ ). Permanent noncapital income in  $t+1$  is equal to its previous value, multiplied by a growth factor  $\Gamma$ , modified by a mean-one iid shock  $\psi_{t+1}$ ,  $\mathbb{E}_t[\psi_{t+n}] = 1 \ \forall n \geq 1$  satisfying  $\psi \in [\underline{\psi}, \bar{\psi}]$  for  $0 < \underline{\psi} \leq 1 \leq \bar{\psi} < \infty$  (and  $\underline{\psi} = \bar{\psi} = 1$  is the degenerate case with no permanent shocks).

Following Zeldes (1989), in future periods  $t+n \ \forall n \geq 1$  there is a small probability  $\wp$  that income will be zero (a 'zero-income event'),

$$\xi_{t+n} = \begin{cases} 0 & \text{with probability } \wp > 0 \\ \theta_{t+n}/(1-\wp) & \text{with probability } (1-\wp) \end{cases} \quad (5)$$

where  $\theta_{t+n}$  is an iid mean-one random variable ( $\mathbb{E}_t[\theta_{t+n}] = 1 \ \forall n > 0$ ) whose distribution satisfies  $\theta \in [\underline{\theta}, \bar{\theta}]$  where  $0 < \underline{\theta} \leq 1 \leq \bar{\theta} < \infty$ .<sup>4</sup> Call the cumulative distribution functions  $\mathcal{F}_\psi$  and  $\mathcal{F}_\theta$  (where  $\mathcal{F}_\xi$  is derived trivially from (5) and  $\mathcal{F}_\theta$ ). For quick identification in tables and graphs, we will call this the 'Friedman/Muth' model because it is a specific implementation of the Friedman (1957) model as interpreted by Muth (1960).

The model looks more special than it is. In particular, a positive probability of zero-income events may seem objectionable (despite empirical support).<sup>5</sup> But a model with a nonzero minimum value of  $\xi$  (motivated, say, by the existence of unemployment insurance) can be redefined by capitalizing the PDV of minimum income into current market assets,<sup>6</sup> transforming that model back into this one. And no key results would change if the transitory shocks were persistent but mean-reverting, instead of IID. Also, the assumption of a positive point mass for the worst realization of the transitory shock is inessential, but simplifies the proofs and is a powerful aid to intuition.

This model differs from Bewley's (1977) classic formulation in several ways. The CRRA utility function does not satisfy Bewley's assumption that  $u(0)$  is well-defined, or that  $u'(0)$  is well defined and finite; indeed, neither the value function nor the marginal value function will be bounded. It differs from Schectman and Escudero (1977) in that they impose liquidity constraints and positive minimum income. It differs from both of these in that it permits permanent growth in income, and also permanent shocks to

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<sup>3</sup>Allowing a stochastic interest factor is straightforward but adds little insight for our purposes; however, see Benhabib, Bisin, and Zhu (2015), Ma and Toda (2020), and Ma, Stachurski, and Toda (2020) for the implications of capital income risk for the distribution of wealth and other interesting questions not considered here.

<sup>4</sup>Rabault (2002) and Li and Stachurski (2014) analyze cases where the shock processes have unbounded support.

<sup>5</sup>We calibrate this probability to 0.005 to match data from the Panel Study of Income Dynamics (Carroll (1992)).

<sup>6</sup>So long as unemployment benefits are proportional to  $\mathbf{p}_t$ ; see the discussion in Section 2.11.

income, which a large empirical literature finds to be of dominant importance in micro data.<sup>7</sup> It differs from Deaton (1991) because liquidity constraints are absent; there are separate transitory and permanent shocks (*a la* Muth (1960)); and the transitory shocks here can occasionally cause income to reach zero. It differs from models found in Stokey et. al. (1989) because neither liquidity constraints nor bounds on utility or marginal utility are imposed.<sup>8,9</sup> Li and Stachurski (2014) show how to allow unbounded returns by using policy function iteration, but also impose constraints.

The paper with perhaps the most in common is Ma, Stachurski, and Toda (2020), henceforth MST, who establish the existence and uniqueness of a solution to a general income fluctuation problem in a Markovian setting. The most important differences are that MST impose liquidity constraints, assume that  $u'(0) = 0$ , and that expected marginal utility of income is finite ( $\mathbb{E}[u'(Y)] < \infty$ ). These assumptions are not consistent with the combination of CRRA utility and income dynamics used here, whose joint properties are key to the results.<sup>10</sup>

## 2.2 The Problem Can Be Normalized By Permanent Income

We establish a bit more notation by reviewing the familiar result that in such problems (CRRA utility, permanent shocks) the number of states can be reduced from two ( $\mathbf{m}$  and  $\mathbf{p}$ ) to one ( $m = \mathbf{m}/\mathbf{p}$ ). Value in the last period is  $u(\mathbf{m}_T)$ ; using (in the last line in (6) below) the fact that for our CRRA utility function,  $u(xy) = x^{1-\rho}u(y)$ , and generically defining nonbold variables as the boldface counterpart normalized by  $\mathbf{p}_t$  (as with  $m = \mathbf{m}/\mathbf{p}$ ), consider the problem in the second-to-last period,

$$\begin{aligned} \mathbf{v}_{T-1}(\mathbf{m}_{T-1}, \mathbf{p}_{T-1}) &= \max_{\mathbf{c}_{T-1}} u(\mathbf{c}_{T-1}) + \beta \mathbb{E}_{T-1}[u(\mathbf{m}_T)] \\ &= \max_{c_{T-1}} u(\mathbf{p}_{T-1}c_{T-1}) + \beta \mathbb{E}_{T-1}[u(\mathbf{p}_T m_T)] \\ &= \mathbf{p}_{T-1}^{1-\rho} \left\{ \max_{c_{T-1}} u(c_{T-1}) + \beta \mathbb{E}_{T-1}[u(\Gamma_T m_T)] \right\}. \end{aligned} \quad (6)$$

Now, in a one-time deviation from the notational convention established in the last sentence, define nonbold ‘normalized value’ not as  $\mathbf{v}_t/\mathbf{p}_t$  but as  $v_t = \mathbf{v}_t/\mathbf{p}_t^{1-\rho}$ , because this allows us to exploit features of the related problem,

$$v_t(m_t) = \max_{\{c\}_t^T} u(c_t) + \beta \mathbb{E}_t[\Gamma_{t+1}^{1-\rho} v_{t+1}(m_{t+1})]$$

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<sup>7</sup>MaCurdy (1982); Abowd and Card (1989); Carroll and Samwick (1997); Jappelli and Pistaferri (2000). Much of the literature instead incorporates highly ‘persistent’ but not completely permanent shocks, but Daly, Hryshko, and Manovskii (2016) show that when measurement problems are handled correctly data yields serial correlation coefficients 0.98 – 1.00; and Hryshko and Manovskii (2020) suggests that survey data support the same conclusion.

<sup>8</sup>Similar restrictions are made in the well known papers by Scheinkman and Weiss (1986), Clarida (1987), and Chamberlain and Wilson (2000). See Toche (2005) for an elegant analysis of a related but simpler continuous-time model.

<sup>9</sup>Alvarez and Stokey (1998) relaxed the bounds on the return function, but they address only the deterministic case.

<sup>10</sup>The incorporation of permanent shocks rules out application of the tools of Matkowski and Nowak (2011), who followed and corrected an error in the fundamental work on the local contraction mapping method developed in Rincón-Zapatero and Rodríguez-Palmero (2003). Martins-da Rocha and Vailakis (2010) provide a correction to Rincón-Zapatero and Rodríguez-Palmero (2003), that works under easier conditions to verify, but only addresses the deterministic case.

$$\begin{aligned}
& \text{s.t.} \\
& a_t = m_t - c_t \\
& k_{t+1} = a_t / \Gamma_{t+1} \\
& b_{t+1} = k_{t+1} R = (R / \Gamma_{t+1}) a_t = \mathcal{R}_{t+1} a_t \\
& m_{t+1} = b_{t+1} + \xi_{t+1},
\end{aligned} \tag{7}$$

where  $\mathcal{R}_{t+1} \equiv (R / \Gamma_{t+1})$  is a ‘permanent-income-growth-normalized’ return factor, and the reformulated problem’s first order condition is<sup>11</sup>

$$c_t^{-\rho} = R\beta \mathbb{E}_t[\Gamma_{t+1}^{-\rho} c_{t+1}^{-\rho}]. \tag{8}$$

Since  $v_T(m_T) = u(m_T)$ , defining  $v_{T-1}(m_{T-1})$  from (7), we obtain

$$\mathbf{v}_{T-1}(\mathbf{m}_{T-1}, \mathbf{p}_{T-1}) = \mathbf{p}_{T-1}^{1-\rho} \underbrace{v_{T-1}(\mathbf{m}_{T-1} / \mathbf{p}_{T-1})}_{=m_{T-1}}.$$

This logic induces to earlier periods; if we solve the normalized one-state-variable problem (7), we will have solutions to the original problem for any  $t < T$  from:

$$\begin{aligned}
\mathbf{v}_t(\mathbf{m}_t, \mathbf{p}_t) &= \mathbf{p}_t^{1-\rho} v_t(m_t), \\
\mathbf{c}_t(\mathbf{m}_t, \mathbf{p}_t) &= \mathbf{p}_t c_t(m_t).
\end{aligned}$$

## 2.3 Definition of a Nondegenerate Solution

The problem has a nondegenerate solution if as the horizon  $n$  gets arbitrarily large the solution in the first period of life  $c_{T-n}(m)$  gets arbitrarily close to a limiting  $c(m)$ :

$$c(m) \equiv \lim_{n \rightarrow \infty} c_{T-n}(m) \tag{9}$$

that satisfies

$$0 < c(m) < \infty \tag{10}$$

for every  $0 < m < \infty$ .

## 2.4 Perfect Foresight Benchmarks

The familiar analytical solution to the perfect foresight model, obtained by setting  $\wp = 0$  and  $\underline{\theta} = \bar{\theta} = \underline{\psi} = \bar{\psi} = 1$ , allows us to define some remaining notation and terminology.

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<sup>11</sup>Leaving aside their assumptions about the marginal utility function and liquidity constraints, it is tempting to view this as a special case of the model of MST, with our  $\mathcal{R}_{t+1} = R / \Gamma_{t+1}$  (defined below equation (7)) corresponding to their stochastic rate of return on capital and the FVAF  $\beta \Gamma_{t+1}^{1-\rho}$  defined below (39) corresponding to their stochastic discount factor. A caveat is that, here,  $\mathcal{R}_{t+1}$  and the modified discount factor are intimately (through  $\Gamma_{t+1}$ ), which has profound effects. It would be interesting, and should not be too difficult, to examine the case with independent shocks to the rate of return and productivity growth.

### 2.4.1 Human Wealth

The dynamic budget constraint, strictly positive marginal utility, and the can't-die-in-debt condition (3) imply an exactly-holding intertemporal budget constraint (IBC):

$$\text{PDV}_t(\mathbf{c}) = \overbrace{\mathbf{m}_t - \mathbf{p}_t}^{\mathbf{b}_t} + \overbrace{\text{PDV}_t(\mathbf{p})}^{\mathbf{h}_t}, \quad (11)$$

where  $\mathbf{b}$  is nonhuman wealth, and with a constant  $\mathcal{R} \equiv R/\Gamma$  ‘human wealth’ is

$$\begin{aligned} \mathbf{h}_t &= \mathbf{p}_t + \mathcal{R}^{-1}\mathbf{p}_t + \mathcal{R}^{-2}\mathbf{p}_t + \cdots + \mathcal{R}^{t-T}\mathbf{p}_t \\ &= \underbrace{\left( \frac{1 - \mathcal{R}^{-(T-t+1)}}{1 - \mathcal{R}^{-1}} \right)}_{\equiv h_t} \mathbf{p}_t. \end{aligned} \quad (12)$$

For  $h \equiv \lim_{n \rightarrow \infty} h_{T-n}$  to be finite, need the Finite Human Wealth Condition (‘FHW’):

$$\underbrace{\Gamma/R}_{\equiv \mathcal{R}^{-1}} < 1. \quad (13)$$

Intuitively, finite human wealth requires a growth rate of (noncapital) income smaller than the interest rate at which that income is being discounted.

### 2.4.2 When Does the Perfect Foresight Unconstrained Solution Exist?

Without constraints, the consumption Euler equation always holds; with  $u'(\mathbf{c}) = \mathbf{c}^{-\rho}$ ,

$$\mathbf{c}_{t+1}/\mathbf{c}_t = (R\beta)^{1/\rho} \equiv \mathbf{P} \quad (14)$$

where the archaic letter ‘**thorn**’ represents what we will call the ‘Absolute Patience Factor’ or APF:

$$\mathbf{P} = (R\beta)^{1/\rho}. \quad (15)$$

$\mathbf{P}$  captures ‘patience’ because, if the ‘absolute impatience condition’ (AIC) holds,<sup>12</sup>

$$\mathbf{P} < 1, \quad (16)$$

the consumer’s level of spending will be too large to sustain indefinitely. We call such a consumer ‘absolutely impatient.’

A ‘Return Patience Factor’ (RPF) relates absolute patience to the return factor:

$$\mathbf{P}_R \equiv \mathbf{P}/R \quad (17)$$

and since consumption is growing by  $\mathbf{P}$  but discounted by  $R$ :

$$\text{PDV}_t(\mathbf{c}) = \left( \frac{1 - \mathbf{P}_R^{T-t+1}}{1 - \mathbf{P}_R} \right) \mathbf{c}_t \quad (18)$$

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<sup>12</sup>Impatience conditions have figured in intertemporal optimization problems since the beginning, e.g. in Ramsey (1928). These issues are so central that it would be hopeless to attempt to cite conditions in every other paper that correspond to conditions named and briefly exposted here. I make no claim to novelty for any condition aside from those implicated in my theorems, whose parallels *will* be articulated.

from which the IBC (11) implies

$$\mathbf{c}_t = \overbrace{\left( \frac{1 - \mathbf{p}_R}{1 - \mathbf{p}_R^{T-t+1}} \right)}^{\equiv \underline{\kappa}_t} (\mathbf{b}_t + \mathbf{h}_t) \quad (19)$$

which defines a normalized finite-horizon perfect foresight consumption function

$$\bar{\mathbf{c}}_{T-n}(m_{T-n}) = \overbrace{(m_{T-n} - 1 + h_{T-n})}^{\equiv b_{T-n}} \underline{\kappa}_{T-n} \quad (20)$$

where  $\underline{\kappa}_t$  is the marginal propensity to consume (MPC) — it answers the question ‘if the consumer had an extra unit of resources, how much more spending would occur.’ ( $\bar{\mathbf{c}}$ ’s overbar signifies that  $\bar{\mathbf{c}}$  will be an upper bound as we modify the problem to incorporate constraints and uncertainty; analogously,  $\underline{\kappa}$  is a lower bound for the MPC).

The denominator of (19) is the reason that, for  $\underline{\kappa}$  to be strictly positive as  $n = T - t$  goes to infinity, we must impose the Return Impatience Condition (RIC):

$$\mathbf{p}_R < 1, \quad (21)$$

so that

$$0 < \underline{\kappa} \equiv 1 - \mathbf{p}_R = \lim_{n \rightarrow \infty} \underline{\kappa}_{T-n}. \quad (22)$$

The RIC thus implies that the consumer cannot be so pathologically patient as to wish, in the limit as the horizon approaches infinity, to spend nothing today out of an increase in current wealth (the RIC rules out the degenerate limiting solution  $\bar{\mathbf{c}}(m) = 0$ ). We call a consumer who satisfies the RIC ‘return impatient.’

Given that the RIC holds, and (as before) defining limiting objects by the absence of a time subscript, the limiting upper bound consumption function will be

$$\bar{\mathbf{c}}(m) = (m + h - 1) \underline{\kappa}, \quad (23)$$

and so in order to rule out the degenerate limiting solution  $\bar{\mathbf{c}}(m) = \infty$  we need  $h$  to be finite; that is, we must impose the Finite Human Wealth Condition (13).

Because  $u(xy) = x^{1-\rho}u(y)$  the value the consumer would achieve by spending permanent income  $\mathbf{p}$  in every period is:

$$\begin{aligned} \mathbf{v}_t^{\text{autarky}} &= u(\mathbf{p}_t) + \beta u(\mathbf{p}_t \Gamma) + \beta^2 u(\mathbf{p}_t \Gamma^2) + \dots \\ &= u(\mathbf{p}_t) \left( 1 + \beta \Gamma^{1-\rho} + (\beta \Gamma^{1-\rho})^2 + \dots \right) \\ &= u(\mathbf{p}_t) \left( \frac{1 - (\beta \Gamma^{1-\rho})^{T-t+1}}{1 - \beta \Gamma^{1-\rho}} \right) \end{aligned} \quad (24)$$

which (for  $\Gamma > 0$ ) asymptotes to a finite number as  $n = T - t$  approaches  $+\infty$  if any of these equivalent conditions holds:

$$\begin{aligned} \overbrace{\beta \Gamma^{1-\rho}}^{\equiv \exists} &< 1 \\ \beta R \Gamma^{-\rho} &< R/\Gamma \end{aligned} \quad (25)$$



$$\mathbf{P}_R < (\Gamma/R)^{1-1/\rho},$$

where we call  $\mathbf{\beth}^{13}$  the ‘Perfect Foresight Value Of Autarky Factor’ (PF-VAF), and the variants of (25) constitute alternative versions of the Perfect Foresight Finite Value of Autarky Condition, PF-FVAC; they guarantee that a consumer who always spends all permanent income ‘has finite autarky value’ (in the perfect foresight case).<sup>14</sup>

If the FHWC is satisfied, the PF-FVAC implies that the RIC is satisfied.<sup>15</sup> Likewise, if the FHWC and the GIC are both satisfied, PF-FVAC follows:

$$\begin{aligned} \mathbf{P} &< \Gamma < R \\ \mathbf{P}_R &< \Gamma/R < (\Gamma/R)^{1-1/\rho} < 1 \end{aligned} \tag{27}$$

(the last line holds because  $\text{FHWC} \Rightarrow 0 \leq (\Gamma/R) < 1$  and  $\rho > 1 \Rightarrow 0 < 1 - 1/\rho < 1$ ).

The first panel of Table 4 summarizes: The PF-Unconstrained model has a non-degenerate limiting solution if we impose the RIC and FHWC (these conditions are necessary as well as sufficient). Together the PF-FVAC and the FHWC imply the RIC, so PF-FVAC and FHWC are jointly sufficient. If we impose the GIC and the FHWC, both the PF-FVAC and the RIC follow, so GIC+FHWC are also sufficient. But there are circumstances under which the RIC and FHWC can hold while the PF-FVAC fails (which we write ~~PF-FVAC~~). For example, if  $\Gamma = 0$ , the problem is a standard ‘cake-eating’ problem with a nondegenerate solution under the RIC.

Perhaps more useful than prose or a table, the relations of the conditions for the unconstrained perfect foresight case are presented diagrammatically in Figure 2. Each node represents a quantity considered in the foregoing analysis. The arrow associated with each inequality reflects the imposition of that condition. For example, one way we wrote the PF-FVAC in equation (25) is  $\mathbf{P} < R^{1/\rho}\Gamma^{1-1/\rho}$ , so imposition of the PF-FVAC is captured by the diagonal arrow connecting  $\mathbf{P}$  and  $R^{1/\rho}\Gamma^{1-1/\rho}$ . Traversing the boundary of the diagram clockwise starting at  $\mathbf{P}$  involves imposing first the GIC then the FHWC, and the consequent arrival at the bottom right node tells us that these two conditions jointly imply that the PF-FVAC holds. Reversal of a condition reverses the arrow’s direction; so, for example, the bottom-most arrow going from  $R$  to  $R^{1/\rho}\Gamma^{1-1/\rho}$  imposes ~~FHWC~~; but we can cancel the cancellation and reverse the arrow. This would allow us to traverse the diagram in a clockwise direction from  $\mathbf{P}$  to  $R$ , revealing that imposition of GIC and FHWC (and, redundantly, FHWC again) let us conclude that the RIC holds because the starting point is  $\mathbf{P}$  and the endpoint is  $R$ . (Consult Appendix K for a detailed exposition of diagrams of this type).

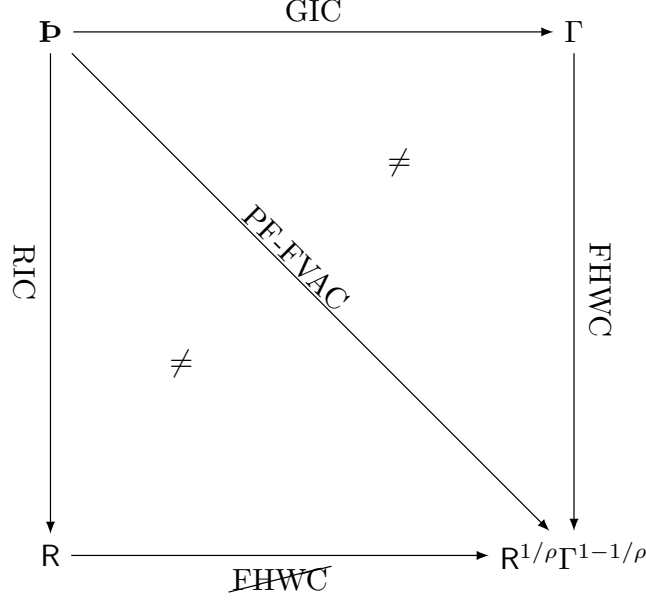
<sup>13</sup>This is another kind of discount factor, so we use the Hebrew ‘bet’ which is a cognate of the Greek ‘beta’.

<sup>14</sup>This is related to the key impatience condition in Alvarez and Stokey (1998).

<sup>15</sup>Divide both sides of the second inequality in (25) by  $R$ :

$$\mathbf{P}/R < (\Gamma/R)^{1-1/\rho} \tag{26}$$

and  $\text{FHWC} \Rightarrow$  the RHS is  $< 1$  because  $(\Gamma/R) < 1$  (and the RHS is raised to a positive power (because  $\rho > 1$ )).



**Figure 2** PF Unconstrained Model: Relation of GIC, FHWC, RIC, and PF-FVAC

An arrowhead points to the larger of the two quantities being compared. For example, the diagonal arrow indicates that  $\mathbf{P} < R^{1/\rho}\Gamma^{1-1/\rho}$ , which is one way of writing the PF-FVAC, equation (25)

#### 2.4.3 PF Constrained Solution Exists Under RIC or Under $\{\text{RIC}, \text{GIC}\}$

We next sketch the perfect foresight constrained solution because it is a useful benchmark (and limit) for the unconstrained problem with uncertainty which is our ultimate interest.

If a liquidity constraint requiring  $b \geq 0$  is ever to be relevant, it must be relevant at the lowest possible level of market resources,  $m_t = 1$ , defined by the lower bound for entering the period,  $b_t = 0$  (if it were relevant at any higher point, it would certainly be relevant at this point). The constraint is ‘relevant’ if it prevents the choice that would otherwise be optimal; at  $m_t = 1$  the constraint is relevant if the marginal utility from spending all of today’s resources  $c_t = m_t = 1$ , exceeds the marginal utility from doing the same thing next period,  $c_{t+1} = 1$ ; that is, if such choices would violate the Euler equation (8):

$$1^{-\rho} > R\beta\Gamma^{-\rho}1^{-\rho}. \quad (28)$$

By analogy to the RPF, we therefore define a ‘growth patience factor’ (GPF) as

$$\mathbf{P}_\Gamma = \mathbf{P}/\Gamma, \quad (29)$$

and define a ‘growth impatience condition’ (GIC)

$$\mathbf{P}_\Gamma < 1 \quad (30)$$

which is equivalent to (28) (exponentiate both sides by  $1/\rho$ ).

We now examine implications of possible configurations of the conditions.

*GIC and RIC.* If the GIC fails but the RIC (21) holds, Appendix A shows that, for some  $0 < m_\# < 1$ , an unconstrained consumer behaving according to (23) would choose

$c < m$  for all  $m > m_{\#}$ . In this case the solution to the constrained consumer's problem is simple: For any  $m \geq m_{\#}$  the constraint does not bind (and will never bind in the future); for such  $m$  the constrained consumption function is identical to the unconstrained one. If the consumer were somehow<sup>16</sup> to arrive at an  $m < m_{\#} < 1$  the constraint would bind and the consumer would consume  $c = m$ . Using the  $\dot{\cdot}$  accent for the version of a function  $\bullet$  in the presence of constraints (and recalling that  $\bar{c}(m)$  is the unconstrained perfect foresight solution):

$$\dot{c}(m) = \begin{cases} m & \text{if } m < m_{\#} \\ \bar{c}(m) & \text{if } m \geq m_{\#}. \end{cases} \quad (31)$$

*GIC and RIC.* More useful is the case where the return impatience and GIC conditions both hold. In this case Appendix A shows that the limiting constrained consumption function is piecewise linear, with  $\dot{c}(m) = m$  up to a first 'kink point' at  $m_{\#}^1 > 1$ , and with discrete declines in the MPC at a set of kink points  $\{m_{\#}^1, m_{\#}^2, \dots\}$ . As  $m \uparrow \infty$  the constrained consumption function  $\dot{c}(m)$  becomes arbitrarily close to the unconstrained  $\bar{c}(m)$ , and the marginal propensity to consume function  $\dot{\kappa}(m) \equiv \dot{c}'(m)$  limits to  $\underline{\kappa}$ .<sup>17</sup> Similarly, the value function  $\dot{v}(m)$  is nondegenerate and limits into the value function of the unconstrained consumer.

This logic holds even when the finite human wealth condition fails ( $\overline{\text{EHWC}}$ ), because the constraint prevents the (limiting) consumer<sup>18</sup> from borrowing against unbounded human wealth to finance unbounded current consumption. Under these circumstances, the consumer who starts with any  $b_t > 1$  will, over time, run those resources down so that by some finite number of periods  $\tau$  in the future the consumer will reach  $b_{t+\tau} = 0$ , and thereafter will set  $\mathbf{c} = \mathbf{p}$  for eternity (which the PF-FVAC says yields finite value). Using the same steps as for equation (24), value of the interim program is also finite:

$$\mathbf{v}_{t+\tau} = \Gamma^{\tau(1-\rho)} \mathbf{u}(\mathbf{p}_t) \left( \frac{1 - (\beta \Gamma^{1-\rho})^{T-(t+\tau)+1}}{1 - \beta \Gamma^{1-\rho}} \right).$$

So, if  $\overline{\text{EHWC}}$ , the limiting consumer's value for any finite  $m$  will be the sum of two finite numbers: The component due to the unconstrained consumption choice made over the finite horizon leading up to  $b_{t+\tau} = 0$ , and the finite component due to the value of consuming all  $\mathbf{p}_{t+\tau}$  thereafter.

*GIC and  $\overline{\text{RIC}}$ .* The most peculiar possibility occurs when the RIC fails. Under these circumstances the  $\overline{\text{FHWC}}$  must also fail (Appendix A), and the constrained consumption function is nondegenerate. (See appendix Figure 9 for a numerical example). While it is true that  $\lim_{m \uparrow \infty} \dot{\kappa}(m) = 0$ , nevertheless the limiting constrained consumption function  $\dot{c}(m)$  is strictly positive and strictly increasing in  $m$ . This result interestingly reconciles the conflicting intuitions from the unconstrained case, where  $\overline{\text{RIC}}$  would suggest a degenerate limit of  $\dot{c}(m) = 0$  while  $\overline{\text{EHWC}}$  would suggest a degenerate limit of  $\dot{c}(m) = \infty$ .

Tables 3 and 4 (and appendix table 5) codify.

<sup>16</sup>"Somehow" because  $m < 1$  could only be obtained by entering the period with  $b < 0$  which the constraint forbids.

<sup>17</sup>See Carroll, Holm, and Kimball (2019) for details.

<sup>18</sup>That is, one obeying  $c(m) = \lim_{n \uparrow \infty} c_{T-n}(m)$ .

We now examine the case with uncertainty but without constraints, which will turn out to be a close parallel to the model with constraints but without uncertainty.

## 2.5 Uncertainty-Modified Conditions

### 2.5.1 Impatience

When uncertainty is introduced, the expectation of beginning-of-period bank balances  $b_{t+1}$  can be rewritten as:

$$\mathbb{E}_t[b_{t+1}] = a_t \mathbb{E}_t[(R/\Gamma_{t+1})] = a_t (R/\Gamma) \mathbb{E}_t[\psi_{t+1}^{-1}] \quad (32)$$

where Jensen's inequality guarantees that the expectation of the inverse of the permanent shock is greater than one. It will be convenient to define

$$\underline{\psi} \equiv (\mathbb{E}[\psi^{-1}])^{-1} \quad (33)$$

which satisfies  $\underline{\psi} < 1$  (thanks again to Mr. Jensen), so we can define

$$\underline{\Gamma} \equiv \Gamma \underline{\psi} < \Gamma \quad (34)$$

which allows us to write uncertainty-adjusted versions of equations and conditions in a manner exactly parallel to those for the perfect foresight case; for example, we define a normalized Growth Patience Factor (GPF-Nrm):

$$\mathbf{P}_{\underline{\Gamma}} = \mathbf{P}/\underline{\Gamma} = \mathbb{E}[\mathbf{P}/(\Gamma\psi)] \quad (35)$$

and a normalized version of the Growth Impatience Condition, GIC-Nrm:

$$\mathbf{P}_{\underline{\Gamma}} < 1, \quad (36)$$

that is stronger than the perfect foresight version (30) because  $\underline{\Gamma} < \Gamma$ .

### 2.5.2 Autarky Value

Analogously to (24), value for a consumer who spent exactly their permanent income every period would reflect the product of the expectation of the (independent) future shocks to permanent income:

$$\begin{aligned} \mathbf{v}_t &= \mathbb{E}_t [\mathbf{u}(\mathbf{p}_t) + \beta \mathbf{u}(\mathbf{p}_t \Gamma_{t+1}) + \cdots + \beta^{T-t} \mathbf{u}(\mathbf{p}_t \Gamma_{t+1} \cdots \Gamma_T)] \\ &= \mathbf{u}(\mathbf{p}_t) \left( \frac{1 - (\beta \Gamma^{1-\rho} \mathbb{E}[\psi^{1-\rho}])^{T-t+1}}{1 - \beta \Gamma^{1-\rho} \mathbb{E}[\psi^{1-\rho}]} \right), \end{aligned}$$

suggesting the definition of a utility-compensated equivalent of the permanent shock,

$$\underline{\underline{\psi}} = (\mathbb{E}[\psi^{1-\rho}])^{1/(1-\rho)}, \quad (37)$$

which will satisfy  $\underline{\underline{\psi}} < 1$  for  $\rho > 1$  and nondegenerate  $\psi$ . Defining

$$\underline{\underline{\Gamma}} = \Gamma \underline{\underline{\psi}}, \quad (38)$$

**Table 1** Microeconomic Model Calibration

Calibrated Parameters			
Description	Parameter	Value	Source
Permanent Income Growth Factor	$\Gamma$	1.03	PSID: Carroll (1992)
Interest Factor	$R$	1.04	Conventional
Time Preference Factor	$\beta$	0.96	Conventional
Coefficient of Relative Risk Aversion	$\rho$	2	Conventional
Probability of Zero Income	$\wp$	0.005	PSID: Carroll (1992)
Std Dev of Log Permanent Shock	$\sigma_\psi$	0.1	PSID: Carroll (1992)
Std Dev of Log Transitory Shock	$\sigma_\theta$	0.1	PSID: Carroll (1992)

$\mathbf{v}_t$  will be positive and finite as  $T$  approaches  $\infty$  if

$$\begin{aligned}
 0 < \overbrace{\beta \Gamma^{1-\rho}}^{\equiv \underline{\underline{\Gamma}}} < 1 \\
 0 < \beta < \underline{\underline{\Gamma}}^{\rho-1}
 \end{aligned} \tag{39}$$

We call (39) the ‘finite value of autarky condition’ because it guarantees that value is finite for a consumer who always consumes their (now stochastic) permanent income (and we will call  $\underline{\underline{\Gamma}}$  the ‘Value of Autarky Factor’ (or ‘VAF’)).<sup>19</sup> For nondegenerate  $\psi$ , this condition is stronger (harder to satisfy in the sense of requiring lower  $\beta$ ) than the perfect foresight version (25) because  $\underline{\underline{\Gamma}} < \Gamma$ .<sup>20</sup>

## 2.6 The Baseline Numerical Solution

Figure 3, familiar from the literature, depicts the successive consumption rules that apply in the last period of life ( $c_T(m)$ ), the second-to-last period, and earlier periods under baseline parameter values listed in Table 2. (The 45 degree line is  $c_T(m) = m$  because in the last period of life it is optimal to spend all remaining resources.)

In the figure, the consumption rules appear to converge to a nondegenerate  $c(m)$ . Our next purpose is to show that this appearance is not deceptive.

<sup>19</sup>In a stationary environment — that is, with  $\underline{\underline{\Gamma}} = 1$  — this corresponds to an impatience condition imposed by Ma, Stachurski, and Toda (2020); but their remaining conditions do not correspond to those here, because their problem differs and their method of proof differs.

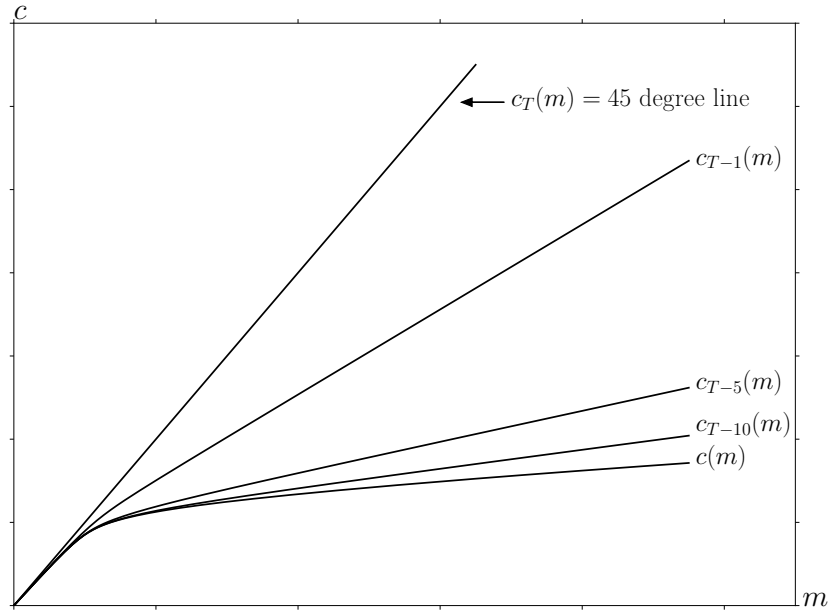
<sup>20</sup>To see this, rewrite (39) as

$$\begin{aligned}
 \beta R &< \underline{\underline{\Gamma}}^{\rho-1} \\
 (\beta R)^{1/\rho} &< R^{1/\rho} \underline{\underline{\Gamma}}^{1-1/\rho} \psi^{1-1/\rho} \\
 \mathbf{P}_\Gamma &< (R/\Gamma)^{1/\rho} \underline{\underline{\Gamma}}^{1-1/\rho}
 \end{aligned}$$

where the last equation is the same as the PF-FVAC condition except that the RHS is multiplied by  $\underline{\underline{\Gamma}}^{1-1/\rho}$  which is strictly less than 1.

**Table 2** Model Characteristics Calculated from Parameters

Description	Symbol and Formula	Approximate Calculated Value
Finite Human Wealth Factor	$\mathcal{R}^{-1} \equiv \Gamma/R$	0.990
PF Finite Value of Autarky Factor	$\sqsupset \equiv \beta\Gamma^{1-\rho}$	0.932
Growth Compensated Permanent Shock	$\underline{\psi} \equiv (\mathbb{E}[\psi^{-1}])^{-1}$	0.990
Uncertainty-Adjusted Growth	$\underline{\Gamma} \equiv \Gamma\underline{\psi}$	1.020
Utility Compensated Permanent Shock	$\underline{\underline{\psi}} \equiv (\mathbb{E}[\psi^{1-\rho}])^{1/(1-\rho)}$	0.990
Utility Compensated Growth	$\underline{\underline{\Gamma}} \equiv \Gamma\underline{\underline{\psi}}$	1.020
Absolute Patience Factor	$\mathfrak{P} \equiv (R\beta)^{1/\rho}$	0.999
Return Patience Factor	$\mathfrak{P}_R \equiv \mathfrak{P}/R$	0.961
Growth Patience Factor	$\mathfrak{P}_\Gamma \equiv \mathfrak{P}/\Gamma$	0.970
Normalized Growth Patience Factor	$\mathfrak{P}_{\underline{\Gamma}} \equiv \mathfrak{P}/\underline{\Gamma}$	0.980
Finite Value of Autarky Factor	$\sqsubseteq \equiv \beta\Gamma^{1-\rho}\underline{\underline{\psi}}^{1-\rho}$	0.941
Weak Impatience Factor	$\wp^{1/\rho}\mathfrak{P} \equiv (\wp\beta R)^{1/\rho}$	0.071



**Figure 3** Convergence of the Consumption Rules

## 2.7 Concave Consumption Function Characteristics

A precondition for the main proof is that the maximization problem defines a sequence of continuously differentiable strictly increasing strictly concave<sup>21</sup> functions  $\{c_T, c_{T-1}, \dots\}$ . The straightforward but tedious proof is relegated to Appendix B. For present purposes, the most important point is that the income process induces what Aiyagari (1994) dubbed a ‘natural borrowing constraint’:  $c_t(m) < m$  for all periods  $t < T$  because a consumer who spent all available resources would arrive in period  $t + 1$  with balances  $b_{t+1}$  of zero, and then might earn zero income over the remaining horizon, risking the possibility of a requirement to spend zero, yielding negative infinite utility. To avoid this disaster, the consumer never spends everything. Zeldes (1989) seems to have been the first to argue, based on his numerical results, that the natural borrowing constraint was a quantitatively plausible alternative to ‘artificial’ or ‘ad hoc’ borrowing constraints.<sup>22</sup>

Strict concavity and continuous differentiability of the consumption function are key elements in many of the arguments below, but are not characteristics of models with ‘artificial’ borrowing constraints – and we will see below that the analytical convenience of these features is considerable.

## 2.8 Bounds for the Consumption Functions

The consumption functions depicted in Figure 3 appear to have limiting slopes as  $m \downarrow 0$  and as  $m \uparrow \infty$ . This section confirms that impression and derives those slopes, which will be needed in the contraction mapping proof.<sup>23</sup>

Assume (as justified above) that a continuously differentiable concave consumption function exists in period  $t + 1$ , with an origin at  $c_{t+1}(0) = 0$ , a minimal MPC  $\underline{\kappa}_{t+1} > 0$ , and maximal MPC  $\bar{\kappa}_{t+1} \leq 1$ . (If  $t + 1 = T$  these will be  $\underline{\kappa}_T = \bar{\kappa}_T = 1$ ; for earlier periods they will exist by recursion.)

Under our imposed assumption that human wealth is finite, the MPC bound as wealth approaches infinity is easy to understand: As the *proportion* of consumption that will be financed out of human wealth approaches zero, the proportional difference between the solution to the model with uncertainty and the perfect foresight model shrinks to zero. In the course of proving this, Appendix G provides a useful recursive expression (used below) for the (inverse of the) limiting MPC:

$$\underline{\kappa}_t^{-1} = 1 + \mathbf{P}_R \underline{\kappa}_{t+1}^{-1}. \quad (40)$$

<sup>21</sup>With one obvious exception:  $c_T(m)$  is linear (and so only weakly concave).

<sup>22</sup>The same (numerical) point applies for infinite horizon models (calibrated to actual empirical data on household income dynamics); cf. Carroll (1992).

<sup>23</sup>Benhabib, Bisin, and Zhu (2015) show that the consumption function becomes linear as wealth approaches infinity in a model with capital income risk and liquidity constraints; Ma and Toda (2020) show that these results generalize to the limits derived here if capital income is added to the model.

### 2.8.1 Weak RIC Conditions

Appendix equation (93) presents a parallel expression for the limiting maximal MPC as  $m_t \downarrow 0$ :

$$\bar{\kappa}_t^{-1} = 1 + \wp^{1/\rho} \mathbf{P}_R \bar{\kappa}_{t+1}^{-1} \quad (41)$$

where  $\{\bar{\kappa}_{T-n}^{-1}\}_{n=0}^{\infty}$  is a decreasing convergent sequence if the ‘weak return patience factor’  $\wp^{1/\rho} \mathbf{P}_R$  satisfies:

$$0 \leq \wp^{1/\rho} \mathbf{P}_R < 1, \quad (42)$$

a condition we dub the ‘Weak Return Impatience Condition’ (WRIC) because with  $\wp < 1$  it will hold more easily (for a larger set of parameter values) than the RIC ( $\mathbf{P}_R < 1$ ). The essence of the argument is that as wealth approaches zero, the overriding consideration that limits consumption is the (recursive) fear of the zero-income events. (That is why the probability of the zero income event  $\wp$  appears in the expression.)

We are now in position to observe that the optimal consumption function must satisfy

$$\underline{\kappa}_t m_t \leq c_t(m_t) \leq \bar{\kappa}_t m_t \quad (43)$$

because consumption starts at zero and is continuously differentiable, is strictly concave,<sup>24</sup> and always exhibits a slope between  $\underline{\kappa}_t$  and  $\bar{\kappa}_t$  (the formal proof is in Appendix D).

## 2.9 Conditions Under Which the Problem Defines a Contraction Mapping

As mentioned above, standard theorems in the contraction mapping literature following Stokey et. al. (1989) require utility or marginal utility to be bounded over the space of possible values of  $m$ , which does not hold here because the possibility (however unlikely) of an unbroken string of zero-income events through the end of the horizon means that utility (and marginal utility) are unbounded as  $m \downarrow 0$ . Although a recent literature examines the existence and uniqueness of solutions to Bellman equations in the presence of ‘unbounded returns’ (see, e.g., Matkowski and Nowak (2011)), the techniques in that literature cannot be used to solve the problem here because the required conditions are violated by a problem that incorporates permanent shocks.<sup>25</sup>

Fortunately, Boyd (1990) provided a weighted contraction mapping theorem that Alvarez and Stokey (1998) showed could be used to address the homogeneous case (of which CRRA is an example) in a deterministic framework; later, Durán (2003) showed how to extend the Boyd (1990) approach to the stochastic case.

**Definition 1.** Consider any function  $\bullet \in \mathcal{C}(\mathcal{A}, \mathcal{B})$  where  $\mathcal{C}(\mathcal{A}, \mathcal{B})$  is the space of continuous functions from  $\mathcal{A}$  to  $\mathcal{B}$ . Suppose  $F \in \mathcal{C}(\mathcal{A}, \mathcal{B})$  with  $\mathcal{B} \subseteq \mathbb{R}$  and  $F > 0$ . Then  $\bullet$  is  $F$ -bounded if the  $F$ -norm of  $\bullet$ ,

$$\|\bullet\|_F = \sup_m \left[ \frac{|\bullet(m)|}{F(m)} \right], \quad (44)$$

<sup>24</sup>Carroll and Kimball (1996)

<sup>25</sup>See Yao (2012) for a detailed discussion of the reasons the existing literature up through Matkowski and Nowak (2011) cannot handle the problem described here.



is finite.

For  $\mathcal{C}_F(\mathcal{A}, \mathcal{B})$  defined as the set of functions in  $\mathcal{C}(\mathcal{A}, \mathcal{B})$  that are  $F$ -bounded;  $w, x, y$ , and  $z$  as examples of  $F$ -bounded functions; and using  $\mathbf{0}(m) = 0$  to indicate the function that returns zero for any argument, Boyd (1990) proves the following.

**Boyd's Weighted Contraction Mapping Theorem.** *Let  $\mathsf{T} : \mathcal{C}_F(\mathcal{A}, \mathcal{B}) \rightarrow \mathcal{C}(\mathcal{A}, \mathcal{B})$  such that<sup>26,27</sup>*

- (1)  $\mathsf{T}$  is non-decreasing, i.e.  $x \leq y \Rightarrow \{\mathsf{T}x\} \leq \{\mathsf{T}y\}$
- (2)  $\{\mathsf{T}\mathbf{0}\} \in \mathcal{C}_F(\mathcal{A}, \mathcal{B})$
- (3) There exists some real  $0 < \alpha < 1$  such that
$$\{\mathsf{T}(w + \zeta F)\} \leq \{\mathsf{T}w\} + \zeta \alpha F \quad \text{holds for all real } \zeta > 0 \quad .$$

Then  $\mathsf{T}$  defines a contraction with a unique fixed point.

For our problem, take  $\mathcal{A}$  as  $\mathbb{R}_{>0}$  and  $\mathcal{B}$  as  $\mathbb{R}$ , and define

$$\{\mathsf{E}z\}(a_t) = \mathbb{E}_t \left[ \Gamma_{t+1}^{1-\rho} z(a_t \mathcal{R}_{t+1} + \xi_{t+1}) \right] .$$

Using this, we introduce the mapping  $\mathcal{T} : \mathcal{C}_F(\mathcal{A}, \mathcal{B}) \rightarrow \mathcal{C}(\mathcal{A}, \mathcal{B})$ .<sup>28</sup>

We can show that our operator  $\mathcal{T}$  satisfies the conditions that Boyd requires of his operator  $\mathsf{T}$ , if we impose two restrictions on parameter values. The first is the WRIC necessary for convergence of the maximal MPC, equation (42) above. More serious is the Finite Value of Autarky condition, equation (39). (We discuss the interpretation of these restrictions in detail in Section 2.11 below.) Imposing these restrictions, we are now in position to state the central theorem of the paper.

**Theorem 1.**  *$\mathcal{T}$  is a contraction mapping if the restrictions on parameter values (42) and (39) are true (that is, if the weak return impatience condition and the finite value of autarky condition hold).*

Intuitively, Boyd's theorem shows that if you can find a  $F$  that is everywhere finite but goes to infinity 'as fast or faster' than the function you are normalizing with  $F$ , the normalized problem defines a contraction mapping. The intuition for the FVAC condition is that, with an infinite horizon, with any initial amount of bank balances  $b_0$ , in the limit your value can always be made greater than you would get by consuming exactly the sustainable amount (say, by consuming  $(r/R)b_0 - \epsilon$  for some small  $\epsilon > 0$ ).

The cumbersome details of the proof are relegated to Appendix D. Given that the value function converges, Appendix E.2 shows that the consumption functions converge.<sup>29</sup>

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<sup>26</sup>We will usually denote the function that results from the mapping as, e.g.,  $\{\mathsf{T}w\}$ .

<sup>27</sup>To non-theorists, this notation may be slightly confusing; the inequality relations in (1) and (3) are taken to mean 'for any specific element  $\bullet$  in the domain of the functions in question' so that, e.g.,  $x \leq y$  is short for  $x(\bullet) \leq y(\bullet) \forall \bullet \in \mathcal{A}$ . In this notation,  $\zeta \alpha F$  in (3) is a *function* which can be applied to any argument  $\bullet$  (because  $F$  is a function).

<sup>28</sup>Note that the existence of the maximum is assured by the continuity of  $\{\mathsf{E}z\}(a_t)$  (it is continuous because it is the sum of continuous  $F$ -bounded functions  $z$ ) and the compactness of  $[\underline{\kappa}m_t, \bar{\kappa}m_t]$ .

<sup>29</sup>MST's proof is for convergence of the consumption policy function directly, rather than of the value function, which is why their conditions are on  $u'$ , which governs behavior.

## 2.10 The Liquidity Constrained Solution as a Limit

This section explains why a related problem commonly considered in the literature (e.g., by Deaton (1991)), with a liquidity constraint and a positive minimum value of income, is the limit of the problem considered here as the probability  $\wp$  of the zero-income event approaches zero.

The ‘related’ problem makes two changes to the problem defined above:

1. An ‘artificial’ liquidity constraint is imposed:  $a_t \geq 0$
2. The probability of zero-income events is zero:  $\wp = 0$

The essence of the argument is simple. Imposing the artificial constraint without changing  $\wp$  would not change behavior at all: The possibility of earning zero income over the remaining horizon already prevents the consumer from ending the period with zero assets. So, for precautionary reasons, the consumer will save something.

But the *extent* to which the consumer feels the need to make this precautionary provision depends on the *probability* that it will turn out to matter. As  $\wp \downarrow 0$ , that probability becomes arbitrarily small, so the *amount* of precautionary saving induced by the zero-income events approaches zero as  $\wp \downarrow 0$ . But “zero” is the amount of precautionary saving that would be induced by a zero-probability event for the impatient liquidity constrained consumer.

Another way to understand this is just to think of the liquidity constraint reflecting a component of the utility function that is zero whenever the consumer ends the period with (strictly) positive assets, but negative infinity if the consumer ends the period with (weakly) negative assets.

See Appendix H for the formal proof justifying the foregoing intuitive discussion.<sup>30</sup>

The conditions required for convergence and nondegeneracy are thus strikingly similar between the liquidity constrained perfect foresight model and the model with uncertainty but no explicit constraints: The liquidity constrained perfect foresight model is just the limiting case of the model with uncertainty as the degree of all three kinds of uncertainty (zero-income events, other transitory shocks, and permanent shocks) approaches zero.

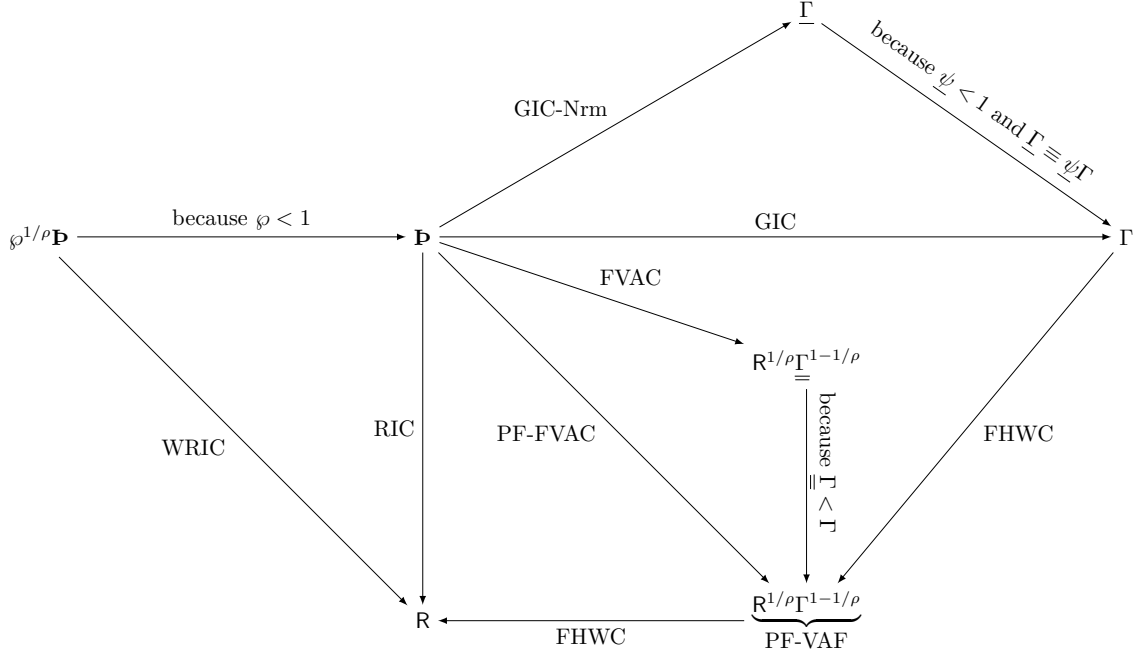
## 2.11 Discussion of Parametric Restrictions

The full relationship among conditions is represented in Figure 4. Though the diagram looks complex, it is merely a modified version of the earlier diagram (Figure 2) with further (mostly intermediate) inequalities inserted. (Arrows with a “because” now label relations that always hold under the model’s assumptions.)<sup>31</sup>

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<sup>30</sup>It seems likely that a similar argument would apply even in the context of a model like that of MST, perhaps with some weak restrictions on returns.

<sup>31</sup>Again, readers unfamiliar with such diagrams should see Appendix K for a more detailed exposition.



**Figure 4** Relation of All Inequality Conditions

See Table 2 for Numerical Values of Nodes Under Baseline Parameters

### 2.11.1 The WRIC

The ‘weakness’ of the additional condition sufficient for contraction beyond the FVAC, the WRIC, can be seen by asking ‘under what circumstances would the FVAC hold but the WRIC fail?’ Algebraically, the requirement is

$$\beta \Gamma^{1-\rho} \underline{\underline{\psi}}^{1-\rho} < 1 < (\varphi \beta)^{1/\rho} / \mathbf{R}^{1-1/\rho}. \quad (45)$$

If we require  $\mathbf{R} \geq 1$ , the WRIC is redundant because now  $\beta < 1 < \mathbf{R}^{\rho-1}$ , so that (with  $\rho > 1$  and  $\beta < 1$ ) the RIC (and WRIC) must hold. But neither theory nor evidence demand that  $\mathbf{R} \geq 1$ . We can therefore approach the question of the WRIC’s relevance by asking just how low  $\mathbf{R}$  must be for the condition to be relevant. Suppose for illustration that  $\rho = 2$ ,  $\underline{\underline{\psi}}^{1-\rho} = 1.01$ ,  $\Gamma^{1-\rho} = 1.01^{-1}$  and  $\varphi = 0.10$ . In that case (45) reduces to

$$\beta < 1 < (0.1\beta/\mathbf{R})^{1/2}$$

but since  $\beta < 1$  by assumption, the binding requirement is that

$$\mathbf{R} < \beta/10$$

so that for example if  $\beta = 0.96$  we would need  $\mathbf{R} < 0.096$  (that is, a perpetual riskfree rate of return of worse than -90 percent a year) in order for the WRIC to bind.

Perhaps the best way of thinking about this is to note that the space of parameter values for which the WRIC is relevant shrinks out of existence as  $\varphi \rightarrow 0$ , which Section 2.10 showed was the precise limiting condition under which behavior becomes

arbitrarily close to the liquidity constrained solution (in the absence of other risks). On the other hand, when  $\wp = 1$ , the consumer has no noncapital income (so that the FHC holds) and with  $\wp = 1$  the WRIC is identical to the RIC; but the RIC is the only condition required for a solution to exist for a perfect foresight consumer with no noncapital income. Thus the WRIC forms a sort of ‘bridge’ between the liquidity constrained and the unconstrained problems as  $\wp$  moves from 0 to 1.

### 2.11.2 When the RIC Fails

In the perfect foresight problem (Section 2.4.2), the RIC was necessary for existence of a nondegenerate solution. It is surprising, therefore, that in the presence of uncertainty, the much weaker WRIC is sufficient for nondegeneracy (assuming that the FVAC holds). We can directly derive the features the problem must exhibit (given the FVAC) under  $\text{RIC}$  (that is,  $R < (R\beta)^{1/\rho}$ ):

$$\begin{aligned}
R &< \overbrace{(R\beta)^{1/\rho} < (R(\Gamma\underline{\psi})^{\rho-1})^{1/\rho}}^{\text{implied by FVAC}} \\
R &< (R/\Gamma)^{1/\rho} \Gamma \underline{\psi}^{1-1/\rho} \\
R/\Gamma &< (R/\Gamma)^{1/\rho} \underline{\psi}^{1-1/\rho} \\
R/\Gamma &< \underline{\underline{\psi}}
\end{aligned} \tag{46}$$

but since  $\underline{\psi} < 1$  (cf. the argument below (37)), this requires  $R/\Gamma < 1$ ; so, given the FVAC, the RIC can fail only if human wealth is unbounded. As an illustration of the usefulness of our diagrams, note that this algebraically complicated conclusion could be easily reached diagrammatically in figure 4 by starting at the  $R$  node and imposing  $\text{RIC}$ , reversing the RIC arrow and then traversing the diagram along any clockwise path to the PF-VAF node at which point we realize that we *cannot* impose the FHC because that would let us conclude  $R > R$ .

As in the perfect foresight constrained problem, unbounded limiting human wealth ( $\text{FHC}$ ) here does not lead to a degenerate limiting consumption function (finite human wealth is not a condition required for the convergence theorem). But, from equation (40) and the discussion surrounding it, an implication of  $\text{RIC}$  is that  $\lim_{m \uparrow \infty} c'(m) = 0$ . Thus, interestingly, in the special  $\{\text{RIC}, \text{FHC}\}$  case (unavailable in the perfect foresight model) the presence of uncertainty both permits unlimited human wealth (in the  $n \uparrow \infty$  limit) and at the same time prevents unlimited human wealth from resulting in (limiting) infinite consumption at any finite  $m$ . Intuitively, in the presence of uncertainty, pathological patience (which in the perfect foresight model results in a limiting consumption function of  $c(m) = 0$  for finite  $m$ ) plus unbounded human wealth (which the perfect foresight model prohibits because it leads to a limiting consumption function  $c(m) = \infty$  for any finite  $m$ ) combine to yield a unique finite limiting (as  $n \uparrow \infty$ ) level of consumption and MPC for any finite value of  $m$ . Note the close parallel to the conclusion in the perfect foresight liquidity constrained model in the  $\{\text{GIC}, \text{RIC}\}$  case. There, too, the

tension between infinite human wealth and pathological patience was resolved with a nondegenerate consumption function whose limiting MPC was zero.<sup>32</sup>

### 2.11.3 When the RIC Holds

**FHWC.** If the RIC and FHWC both hold, a perfect foresight solution exists (see 2.4.2 above). As  $m \uparrow \infty$  the limiting consumption function and value function become arbitrarily close to those in the perfect foresight model, because human wealth pays for a vanishingly small portion of spending. This will be the main case analyzed in detail below.

**FHWC.** The more exotic case is where FHWC does not hold; in the perfect foresight model,  $\{\text{RIC}, \text{FHWC}\}$  is the degenerate case with limiting  $\bar{c}(m) = \infty$ . Here, since the FVAC implies that the PF-FVAC holds (traverse Figure 4 clockwise from  $\mathbf{P}$  by imposing FVAC and continue to the PF-VAF node): Reversing the arrow connecting the R and PF-VAF nodes implies that under  $\text{FHWC}$ :

$$\begin{array}{c} \text{PF-FVAC} \\ \hline \mathbf{P} < (R/\Gamma)^{1/\rho} \Gamma \\ \mathbf{P} < \Gamma \end{array}$$

where the transition from the first to the second lines is justified because  $\text{FHWC} \Rightarrow (R/\Gamma)^{1/\rho} < 1$ . So,  $\{\text{RIC}, \text{FHWC}\}$  implies the GIC holds. However, we are not entitled to conclude that the GIC-Nrm holds:  $\mathbf{P} < \Gamma$  does not imply  $\mathbf{P} < \psi \Gamma$  where  $\psi < 1$ .

We have now established the principal points of comparison between the perfect foresight solutions and the solutions under uncertainty; these are codified in the remaining parts of Tables 3 and 4.

## 3 Analysis of the Converged Consumption Function

Figures 5-7 capture the main properties of the converged consumption rule when the RIC, GIC-Nrm, and FHWC all hold.<sup>33</sup>

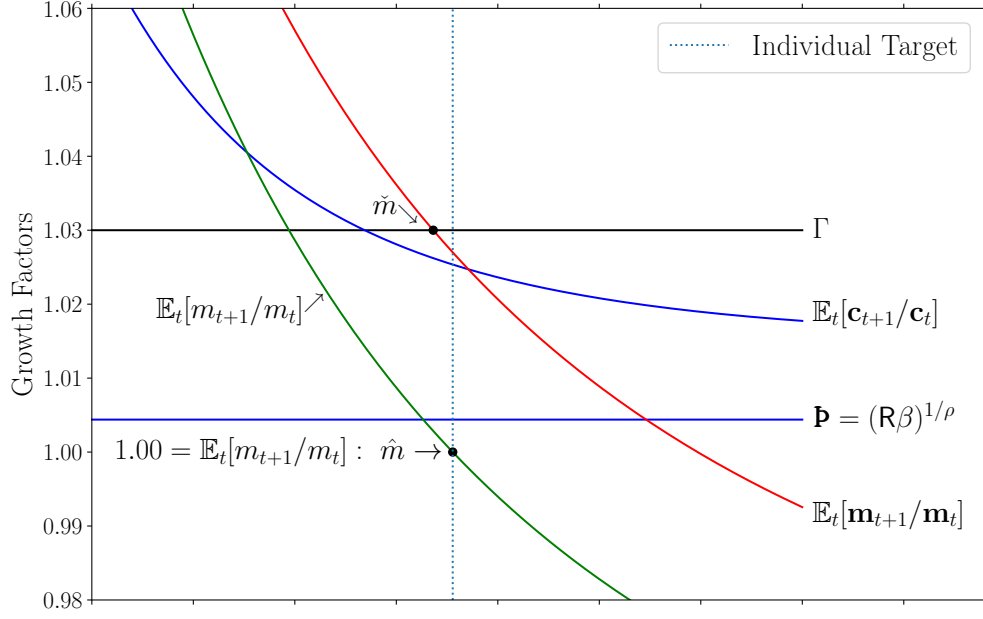
Figure 5 shows the expected growth factors for consumption, the level of market resources, and the market resources ratio,  $\mathbb{E}_t[\mathbf{c}_{t+1}/\mathbf{c}_t]$  and  $\mathbb{E}_t[\mathbf{m}_{t+1}/\mathbf{m}_t]$ , and  $\mathbb{E}_t[m_{t+1}/m_t]$ , for a consumer behaving according to the converged consumption rule, while Figures 6—7 illustrate theoretical bounds for the consumption function and the MPC.

Three points are worth highlighting.

First, as  $m_t \uparrow \infty$  the expected consumption growth factor goes to  $\mathbf{P}$ , indicated by the lower bound in Figure 5, and the marginal propensity to consume approaches  $\underline{\kappa} = (1 - \mathbf{P}_R)$  (see Figure 6) — the same as the perfect foresight MPC. Second, as  $m_t$  approaches zero the consumption growth factor approaches  $\infty$  (Figure 5) and the MPC approaches  $\bar{\kappa} = (1 - \wp^{1/\rho} \mathbf{P}_R)$  (Figure 6). Third, there is a value of the market

<sup>32</sup>Ma and Toda (2020) derive conditions under which the limiting MPC is zero in an even more general case where there is also capital income risk.

<sup>33</sup>These figures reflect the converged rule corresponding to the parameter values indicated in Table 2.



**Figure 5** ‘Stable’  $m$  Values and Expected Growth Factors

resources ratio  $m_t = \tilde{m}$  at which the expected growth rate of the level of market resources  $\mathbf{m}$  matches the expected growth rate of permanent income  $\Gamma$ , and a different (larger) target ratio  $\hat{m}$  where  $\mathbb{E}[m_{t+1}/m_t] = 1$  and the expected growth rate of consumption is lower than  $\Gamma$ . Thus, at the individual level, this model does not have a single  $m$  at which  $\mathbf{p}$ ,  $\mathbf{m}$ , and  $\mathbf{c}$  all grow at the same rate.<sup>34</sup>

### 3.1 Limits as $m$ approaches Infinity

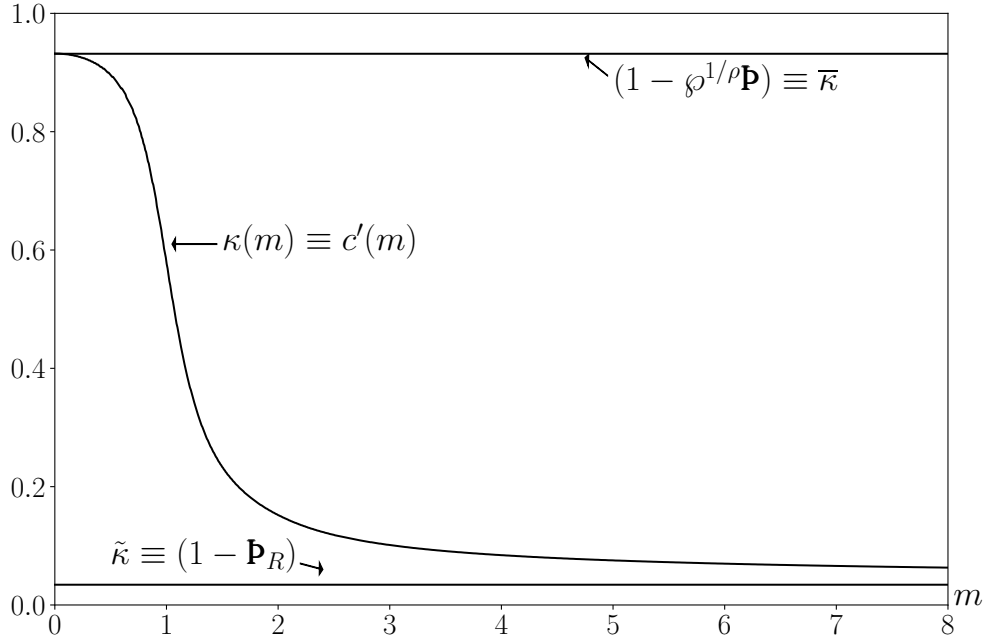
Define

$$\underline{c}(m) = \underline{\kappa}m$$

which is the solution to an infinite-horizon problem with no noncapital income ( $\xi_{t+n} = 0 \ \forall \ n \geq 1$ ); clearly  $\underline{c}(m) < c(m)$ , since allowing the possibility of future noncapital income cannot reduce current consumption. Our imposition of the RIC guarantees that  $\underline{\kappa} > 0$ , so this solution satisfies our definition of nondegeneracy, and because this solution is always available it defines a lower bound on both the consumption and value functions.

Assuming the FHWC holds, the infinite horizon perfect foresight solution (23) constitutes an upper bound on consumption in the presence of uncertainty, since the intro-

<sup>34</sup>A final proposition suggested by Figure 5 is that the expected consumption growth factor is declining in the level of the cash-on-hand ratio  $m_t$ . This turns out to be true in the absence of permanent shocks, but in extreme cases it can be false if permanent shocks are present; see Appendix L.



**Figure 6** Limiting MPC's

duction of uncertainty strictly decreases the level of consumption at any  $m$  (Carroll and Kimball (1996)). Thus, we can write

$$\begin{aligned} \underline{c}(m) &< c(m) < \bar{c}(m) \\ 1 &< c(m)/\underline{c}(m) < \bar{c}(m)/\underline{c}(m). \end{aligned} \tag{47}$$

But

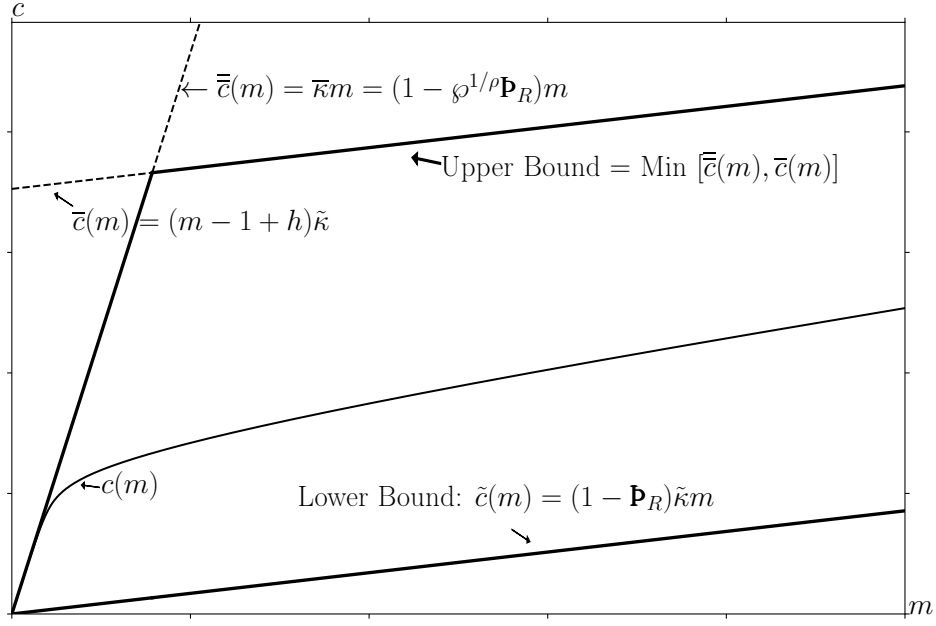
$$\begin{aligned} \lim_{m \uparrow \infty} \bar{c}(m)/\underline{c}(m) &= \lim_{m \uparrow \infty} (m - 1 + h)/m \\ &= 1, \end{aligned}$$

so as  $m \uparrow \infty$ ,  $c(m)/\underline{c}(m) \rightarrow 1$ , and the continuous differentiability and strict concavity of  $c(m)$  therefore implies

$$\lim_{m \uparrow \infty} c'(m) = \underline{c}'(m) = \bar{c}'(m) = \underline{\kappa}$$

because any other fixed limit would eventually lead to a level of consumption either exceeding  $\bar{c}(m)$  or lower than  $\underline{c}(m)$ .

Figure 6 confirms these limits visually. The top plot shows the converged consumption function along with its upper and lower bounds, while the lower plot shows the marginal propensity to consume.



**Figure 7** Upper and Lower Bounds on The Consumption Function

Next we establish the limit of the expected consumption growth factor as  $m_t \uparrow \infty$ :

$$\lim_{m_t \uparrow \infty} \mathbb{E}_t[\mathbf{c}_{t+1}/\mathbf{c}_t] = \lim_{m_t \uparrow \infty} \mathbb{E}_t[\Gamma_{t+1}c_{t+1}/c_t].$$

But

$$\mathbb{E}_t[\Gamma_{t+1}\underline{c}_{t+1}/\bar{c}_t] \leq \mathbb{E}_t[\Gamma_{t+1}c_{t+1}/c_t] \leq \mathbb{E}_t[\Gamma_{t+1}\bar{c}_{t+1}/\underline{c}_t]$$

and

$$\lim_{m_t \uparrow \infty} \Gamma_{t+1}\underline{c}(m_{t+1})/\bar{c}(m_t) = \lim_{m_t \uparrow \infty} \Gamma_{t+1}\bar{c}(m_{t+1})/\underline{c}(m_t) = \lim_{m_t \uparrow \infty} \Gamma_{t+1}m_{t+1}/m_t,$$

while (for convenience defining  $a(m_t) = m_t - c(m_t)$ ),

$$\begin{aligned} \lim_{m_t \uparrow \infty} \Gamma_{t+1}m_{t+1}/m_t &= \lim_{m_t \uparrow \infty} \left( \frac{\text{Ra}(m_t) + \Gamma_{t+1}\xi_{t+1}}{m_t} \right) \\ &= (\text{R}\beta)^{1/\rho} = \mathbf{P} \end{aligned} \quad (48)$$

because  $\lim_{m_t \uparrow \infty} a'(m) = \mathbf{P}_R$ <sup>35</sup> and  $\Gamma_{t+1}\xi_{t+1}/m_t \leq (\Gamma\bar{\psi}\bar{\theta}/(1 - \wp))/m_t$  which goes to zero as  $m_t$  goes to infinity.

Hence we have

$$\mathbf{P} \leq \lim_{m_t \uparrow \infty} \mathbb{E}_t[\mathbf{c}_{t+1}/\mathbf{c}_t] \leq \mathbf{P}$$

<sup>35</sup>This is because  $\lim_{m_t \uparrow \infty} a(m_t)/m_t = 1 - \lim_{m_t \uparrow \infty} c(m_t)/m_t = 1 - \lim_{m_t \uparrow \infty} c'(m_t) = \mathbf{P}_R$ .



so as cash goes to infinity, consumption growth approaches its value  $\mathbf{P}$  in the perfect foresight model.

### 3.2 Limits as $m$ Approaches Zero

Equation (41) shows that the limiting value of  $\bar{\kappa}$  is

$$\bar{\kappa} = 1 - R^{-1}(\wp R \beta)^{1/\rho}.$$

Defining  $e(m) = c(m)/m$  as before we have

$$\lim_{m \downarrow 0} e(m) = (1 - \wp^{1/\rho} \mathbf{P}_R) = \bar{\kappa}.$$

Now using the continuous differentiability of the consumption function along with L'Hôpital's rule, we have

$$\lim_{m \downarrow 0} c'(m) = \lim_{m \downarrow 0} e(m) = \bar{\kappa}.$$

Figure 6 confirms that the numerical solution obtains this limit for the MPC as  $m$  approaches zero.

For consumption growth, as  $m \downarrow 0$  we have

$$\begin{aligned} \lim_{m_t \downarrow 0} \mathbb{E}_t \left[ \left( \frac{c(m_{t+1})}{c(m_t)} \right) \Gamma_{t+1} \right] &> \lim_{m_t \downarrow 0} \mathbb{E}_t \left[ \left( \frac{\underline{c}(\mathcal{R}_{t+1} a(m_t) + \xi_{t+1})}{\bar{\kappa} m_t} \right) \Gamma_{t+1} \right] \\ &= \wp \lim_{m_t \downarrow 0} \mathbb{E}_t \left[ \left( \frac{\underline{c}(\mathcal{R}_{t+1} a(m_t))}{\bar{\kappa} m_t} \right) \Gamma_{t+1} \right] \\ &\quad + (1 - \wp) \lim_{m_t \downarrow 0} \mathbb{E}_t \left[ \left( \frac{\underline{c}(\mathcal{R}_{t+1} a(m_t) + \theta_{t+1}/(1 - \wp))}{\bar{\kappa} m_t} \right) \Gamma_{t+1} \right] \\ &> (1 - \wp) \lim_{m_t \downarrow 0} \mathbb{E}_t \left[ \left( \frac{\underline{c}(\theta_{t+1}/(1 - \wp))}{\bar{\kappa} m_t} \right) \Gamma_{t+1} \right] \\ &= \infty \end{aligned}$$

where the second-to-last line follows because  $\lim_{m_t \downarrow 0} \mathbb{E}_t \left[ \left( \frac{\underline{c}(\mathcal{R}_{t+1} a(m_t))}{\bar{\kappa} m_t} \right) \Gamma_{t+1} \right]$  is positive, and the last line follows because the minimum possible realization of  $\theta_{t+1}$  is  $\underline{\theta} > 0$  so the minimum possible value of expected next-period consumption is positive.<sup>36</sup>

### 3.3 Unique ‘Stable’ Points

Two theorems, whose substance is described here and whose details are in an appendix, articulate alternative (but closely related) stability criteria for the model.

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<sup>36</sup>None of the arguments in either of the two prior sections depended on the assumption that the consumption functions had converged. With more cumbersome notation, each derivation could have been replaced by the corresponding finite-horizon versions. This strongly suggests that it should be possible to extend the circumstances under which the problem can be shown to define a contraction mapping to the union of the parameter values under which {RIC,FHWC} hold and {FVAC,WRIC} hold. That extension is not necessary for our purposes here, so we leave it for future work.

### 3.3.1 ‘Individual Target Wealth’

One definition of a ‘stable’ point is what we will call a ‘target’ value  $\hat{m}$  such that if  $m_t = \hat{m}$ , then  $\mathbb{E}_t[m_{t+1}] = m_t$ . Existence of such a target turns out to require the GIC-Nrm condition.

**Theorem 2.** *For the nondegenerate solution to the problem defined in Section 2.1 when FVAC, WRIC, and GIC-Nrm all hold, there exists a unique cash-on-hand-to-permanent-income ratio  $\hat{m} > 0$  such that*

$$\mathbb{E}_t[m_{t+1}/m_t] = 1 \text{ if } m_t = \hat{m}. \quad (49)$$

Moreover,  $\hat{m}$  is a point of ‘wealth stability’ in the sense that

$$\begin{aligned} \forall m_t \in (0, \hat{m}), \quad \mathbb{E}_t[m_{t+1}] &> m_t \\ \forall m_t \in (\hat{m}, \infty), \quad \mathbb{E}_t[m_{t+1}] &< m_t. \end{aligned} \quad (50)$$

Since  $m_{t+1} = (m_t - c(m_t))\mathcal{R}_{t+1} + \xi_{t+1}$ , the implicit equation for  $\hat{m}$  is

$$\begin{aligned} \mathbb{E}_t[(\hat{m} - c(\hat{m}))\mathcal{R}_{t+1} + \xi_{t+1}] &= \hat{m} \\ (\hat{m} - c(\hat{m})) \underbrace{\mathcal{R} \mathbb{E}_t[\psi^{-1}]}_{\equiv \bar{\mathcal{R}}} + 1 &= \hat{m} \end{aligned} \quad (51)$$

### 3.3.2 The Unexpectedly Expected Individual Balanced Growth ‘pseudo steady state’

A traditional question in macroeconomic models is whether there is a ‘balanced growth’ equilibrium in which aggregate variables (income, consumption, market resources) all grow forever at the same rate. For our model, Figure 5 showed that there is no single  $m$  for which  $\mathbb{E}_t[\mathbf{m}_{t+1}]/\mathbf{m}_t = \mathbb{E}_t[\mathbf{c}_{t+1}]/\mathbf{c}_t = \Gamma$  for an individual consumer. Nevertheless, the next section will show that economies populated by collections of such consumers can exhibit balanced growth in the aggregate, and in the cross-section of households.

As an input to that analysis, we show here that if the GIC holds, the problem will exhibit what we call a ‘pseudo-steady-state’ point, by which we mean that there is some  $\check{m}$  such that, for all  $m_t > \check{m}$ ,  $\mathbb{E}_t[\mathbf{m}_{t+1}/\mathbf{m}_t] < \Gamma$ , and conversely if  $m_t < \check{m}$  then  $\mathbb{E}_t[\mathbf{m}_{t+1}/\mathbf{m}_t] > \Gamma$ .

The critical  $m$  will be the value  $\check{m}$  at which  $\mathbf{m}$  growth matches  $\Gamma$ :

$$\begin{aligned} \mathbb{E}_t[\mathbf{m}_{t+1}]/\mathbf{m}_t &= \mathbb{E}_t[\mathbf{p}_{t+1}]/\mathbf{p}_t \\ \mathbb{E}_t[m_{t+1}\Gamma\psi_{t+1}\mathbf{p}_t]/(m_t\mathbf{p}_t) &= \mathbb{E}_t[\mathbf{p}_t\Gamma\psi_{t+1}]/\mathbf{p}_t \\ \mathbb{E}_t \left[ \psi_{t+1} \underbrace{((m_t - c(m_t))\mathcal{R}/(\Gamma\psi_{t+1}))}_{m_{t+1}} + \xi_{t+1} \right] / m_t &= 1 \\ \mathbb{E}_t \left[ (\check{m} - c(\check{m})) \underbrace{\mathcal{R}}_{\bar{\mathcal{R}}} / \Gamma + \psi_{t+1}\xi_{t+1} \right] &= \check{m} \\ (\check{m} - c(\check{m}))\bar{\mathcal{R}} + 1 &= \check{m}. \end{aligned} \quad (52)$$

The only difference between (52) and (51) is the substitution of  $\mathcal{R}$  for  $\bar{\mathcal{R}}$ .

Our choice to call to this  $\check{m}$  as individual problem's 'pseudo-steady-state' is motivated by what happens in the case where all draws of all future shocks just happen to take on their expected value of 1.0. (They unexpectedly always take on their expected values). In that infinitely improbable case, the economy *would* exhibit balanced growth.<sup>37</sup>

Theorem 3 formally states the relevant proposition.

**Theorem 3.** *For the nondegenerate solution to the problem defined in Section 2.1 when FVAC, WRIC, and GIC all hold, there exists a unique pseudo-steady-state cash-on-hand-to-income ratio  $\check{m} > 0$  such that*

$$\mathbb{E}_t[\psi_{t+1}m_{t+1}/m_t] = 1 \text{ if } m_t = \check{m}. \quad (53)$$

Moreover,  $\check{m}$  is a point of stability in the sense that

$$\begin{aligned} \forall m_t \in (0, \check{m}), \quad \mathbb{E}_t[\mathbf{m}_{t+1}]/\mathbf{m}_t &> \Gamma \\ \forall m_t \in (\check{m}, \infty), \quad \mathbb{E}_t[\mathbf{m}_{t+1}]/\mathbf{m}_t &< \Gamma. \end{aligned} \quad (54)$$

The proofs of the two theorems are almost completely parallel; to save space, they are relegated to Appendix M. In sum, they involve three steps:

1. Existence and continuity of  $\mathbb{E}_t[m_{t+1}/m_t]$  or  $\mathbb{E}_t[m_{t+1}\psi_{t+1}/m_t]$ 
  - This follows from existence and continuity of the constituents
2. Existence of the equilibrium point
  - This follows from the upper and lower bound limiting MPC's, existence and continuity, and the Intermediate Value Theorem
3. Monotonicity of  $\mathbb{E}_t[m_{t+1} - m_t]$  or  $\mathbb{E}_t[m_{t+1}\psi_{t+1} - m_t]$ 
  - This follows from concavity of the consumption function

### 3.3.3 Example Where There Is An Expected-Balanced-Growth $\check{m}$ State But No Target

Because the equations defining target and pseudo-steady-state  $m$ , (51) and (52), differ only by substitution of  $\mathcal{R}$  for  $\bar{\mathcal{R}} = \mathcal{R} \mathbb{E}[\psi^{-1}]$ , if there are no permanent shocks ( $\psi \equiv 1$ ), the conditions are identical. For many parameterizations (e.g., under the baseline parameter values used for constructing figure 5),  $\hat{m}$  and  $\check{m}$  will not differ much.

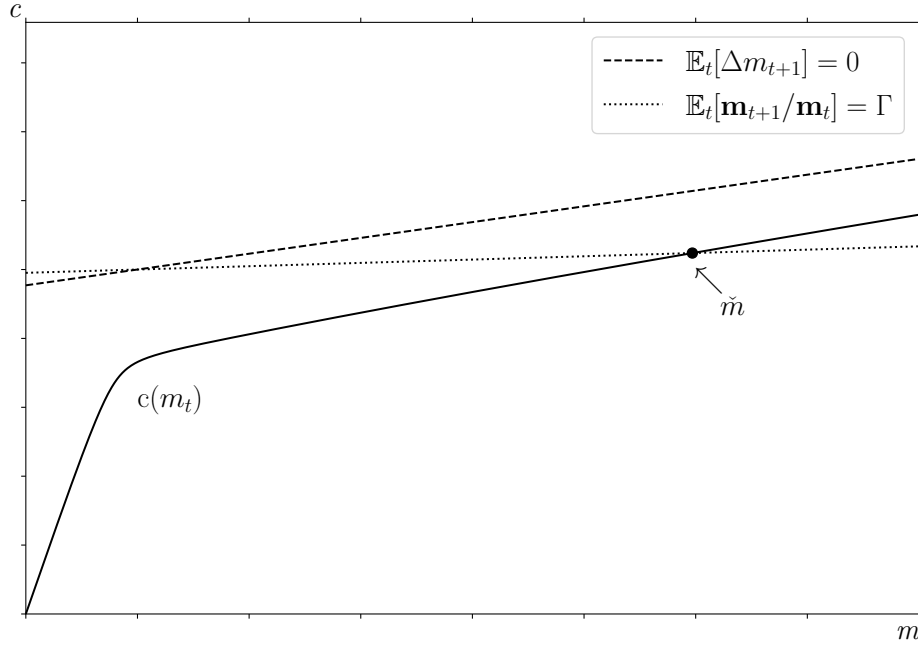
An illuminating exception is exhibited in Figure 8, which modifies the baseline parameter values by quadrupling the variance of the permanent shocks, enough to cause failure of the GIC-Nrm; now there is no target wealth level  $\hat{m}$  (precautionary motives keep  $c(m)$  everywhere below the level that would keep expected  $m$  constant).

The pseudo-steady-state still exists because it turns off realizations of the permanent shock. But the next section will show that an aggregate balanced growth equilibrium

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$$\mathbb{E}_t[m_{t+1}/m_t | \psi_{t+1} = \xi_{t+1} = 1] = \Gamma (\check{m} - c(\check{m})\mathcal{R} + 1) / m = \Gamma$$



**Figure 8**  $\{\text{FVAC}, \text{GIC}, \text{GIC-Nrm}\}$ : No  $\hat{m}$  Exists But  $\check{m}$  Does

will exist even when realizations of the permanent shock are not turned off, and instead implemented exactly as specified in the model: The required condition for aggregate balanced growth is the regular, not the normalized, GIC.

Before we get to the formal arguments below, the key insight can be understood here by considering the evolution of an economy that starts, at date  $t$ , with the entire population at  $m_t = \check{m}$ , but then evolves according to the model's assumed dynamics between  $t$  and  $t + 1$ . Equation (52) will still hold, so for this first period, at least, the economy will exhibit balanced growth: the growth factor for aggregate  $\mathbf{m}$  will match the growth factor for permanent income  $\Gamma$ . It is true that there will be people for whom  $b_{t+1} = a_t R / (\Gamma \psi_{t+1})$  is boosted by a small draw of  $\psi_{t+1}$ . But their contribution to the aggregate variable is given by  $\mathbf{b}_{t+1} = b_{t+1} \psi_{t+1}$ , so their  $\mathbf{b}_{t+1}$  is reweighted by an amount that exactly unwinds that boosting. This is how it is possible for each individual consumer to have a target wealth ratio of infinity, and yet for the economy as a whole to have a balanced growth equilibrium with a finite aggregate balanced growth steady state (which, for symmetry with the pseudo-steady-state, we will call  $\check{M}$ ). This is not numerically the same as the individual **pseudo-steady-state** ratio  $\check{m}$  because the nonlinearities involved in simulation and aggregation will have consequences.<sup>38</sup>

<sup>38</sup>Still, the pseudo-steady-state can be calculated immediately from the policy function without any simulation, and therefore would likely serve as an excellent and low-cost starting point for the numerical simulation process.

## 4 Invariant Aggregate, Idiosyncratic, and Covariance Relationships

Assume a continuum of *ex ante* identical households on the unit interval, with constant total mass normalized to one and indexed by  $i \in [0, 1]$ , all behaving according to the model specified above.

Szeidl (2013) proved that such a population will be characterized by invariant distributions of  $m$ ,  $c$ , and  $a$  under the condition<sup>39</sup>

$$\mathbf{P}_\Gamma < e^{\mathbb{E}[\log \psi]} \quad (55)$$

which is stronger than our GIC condition  $\mathbf{P}_\Gamma < 1$  (under the imposed assumption  $\mathbb{E}[\log \psi] < 0$ ).<sup>40</sup>

Harmenberg (2021) substitutes a clever change of probability measure into Szeidl (2013)’s proof to show that under the GIC, invariant *permanent-income-weighted* distributions exist.<sup>41</sup> This allows him to prove a conjecture from an earlier draft of this paper (Carroll (2020, Submitted)) that under the GIC, aggregate consumption grows at the same rate  $\Gamma$  as aggregate noncapital income in the long run (with the corollary that aggregate assets and market resources grow at the same rate).

Harmenberg (2021) shows<sup>42</sup> that his reformulation of the problem can reduce costs of calculation enormously.

The remainder of this section briefly draws out some implications of these points.

### 4.1 Individual Balanced Growth of Income, Consumption, and Wealth

Say that  $\mathbb{M}[\bullet]$  yields the mean of its argument in the population (as distinct from the expectations operator  $\mathbb{E}[\bullet]$  which represents beliefs about the future). Using boldface capitals for aggregates, the growth factor for aggregate noncapital income is:

$$\begin{aligned} \mathbf{Y}_{t+1}/\mathbf{Y}_t &= \mathbb{M}[\xi_{t+1}\Gamma\psi_{t+1}\mathbf{p}_t] / \mathbb{M}[\mathbf{p}_t\xi_t] \\ &= \Gamma \end{aligned}$$

---

<sup>39</sup>Szeidl (2013)’s equation (9), in our notation, is:

$$\begin{aligned} \mathbb{E} \log \mathbf{R}(1 - \kappa) &< \mathbb{E} \log \Gamma\psi \\ \mathbb{E} \log \mathbf{R}\mathbf{P}_\Gamma &< \mathbb{E} \log \Gamma\psi \\ \log \mathbf{P}_\Gamma &< \mathbb{E} \log \psi \end{aligned}$$

which, exponentiated, yields (55).

<sup>40</sup>Under our default (though not required) assumption that  $\log \psi \sim \mathcal{N}(-\sigma_\psi^2/2, \sigma_\psi^2)$ ; the GIC-Nrm in this case, is  $\mathbf{P}_\Gamma < \exp(-\sigma^2)$ , so if the GIC-Nrm holds then Szeidl’s condition will hold.

<sup>41</sup>Harmenberg in a nutshell: If  $\psi$  is described by density function  $f_\psi(\psi)$ , define the *permanent-income-neutral measure* by  $\tilde{f}_\psi(\psi) \equiv \psi f_\psi(\psi)$ . Following derivations parallel to those in the previous footnote, the condition for the existence of invariant permanent-income-weighted distributions is

$$\log \mathbf{P}_\Gamma < \tilde{\mathbb{E}} \log \psi$$

where the  $\tilde{\cdot}$  over the expectations operator indicates that the expectation is being taken with respect to  $f_\psi(\psi)$ . But by assumption  $\tilde{\mathbb{E}} \log \psi = 0$  so exponentiating both sides of this equation turns it into  $\mathbf{P}_\Gamma < 1$  which is the GIC.

<sup>42</sup>The test of the Harmenberg method currently implemented in the Econ-ARK toolkit shows that, for the same computational resources, Harmenberg’s method reduces the standard error of the simulated aggregate consumption series by a factor of more than 8; see the last part of `reproduce/computed.sh`.

because of the independence assumptions we have made about  $\xi$  and  $\psi$ .

Consider an economy that satisfies the Szeidl impatience condition (55) and has existed for long enough by date  $t$  that by some measure we can consider it as Szeidl-converged. In such an economy a microeconomist with a population-representative panel dataset could calculate the growth rate of consumption for each individual household, and take the average to obtain:

$$\begin{aligned}\mathbb{M}[\Delta \log \mathbf{c}_{t+1}] &= \mathbb{M}[\log c_{t+1} \mathbf{p}_{t+1} - \log c_t \mathbf{p}_t] \\ &= \mathbb{M}[\log \mathbf{p}_{t+1} - \log \mathbf{p}_t] + \mathbb{M}[\log c_{t+1} - \log c_t].\end{aligned}\quad (56)$$

Because this economy is Szeidl-converged, distributions of  $c_t$  and  $c_{t+1}$  will be identical, so that the second term in (56) disappears, reducing the result to the proposition that cross-sectional growth rates of consumption and permanent income are the same:

$$\mathbb{M}[\Delta \log \mathbf{c}_{t+1}] = \mathbb{M}[\Delta \log \mathbf{p}_{t+1}] = \log \Gamma. \quad (57)$$

In a Harmenberg-invariant economy (and therefore also any Szeidl-invariant economy), a similar proposition that is testable using cross-section data is a direct implication of the fact that in such an economy a constant proportion of aggregate permanent income is accounted for by the successive sets of consumers with any particular  $m$ . This is essentially Harmenberg's definition of the density of the permanent-income-invariant distribution of a variable (say,  $m$ ). Call  $\mathbf{c}_t(m)$  the total amount of consumption at date  $t$  by persons with market resources  $m$ , and note that in the invariant economy it is given by the converged consumption function  $c(m)$  multiplied by the amount of permanent income accruing to such, people  $p_t(m)P_t$ . But Harmenberg's proof shows that  $p(m)$  is invariant,<sup>43</sup> and we know from above that  $P_{t+1} = \Gamma P_t$ . So

$$\begin{aligned}\log \mathbf{c}_{t+1}(m) - \log \mathbf{c}_t(m) &= \log c(m)p(m)P_{t+1} - \log c(m)p(m)P_t \\ &= \log \Gamma.\end{aligned}\quad (58)$$

Thus we have microeconomic propositions that can be tested in either cross-section or panel microdata to judge the microfoundations that should be expected to hold for a macroeconomic analysis that requires balanced growth for its conclusions.

## 4.2 Balanced Growth and Covariances

Harmenberg's brief discussion of covariance relationships demonstrates that they do not shrink to zero, with the implication that they are another potential target of measurement for construction of microfoundations. Here we draw out some of those implications.

Consider again a Harmenberg-converged economy at date  $t = 0$ , and define  $\bar{c}_{t+n} \equiv \mathbb{M}[c_{t+n}]$ . Total consumption at  $t + 1$  and  $t + 2$  are

$$\begin{aligned}\mathbf{C}_{t+1} &= \mathbb{M}[c_{t+1} \mathbf{p}_{t+1}] = \bar{c}_{t+1} \Gamma^1 + \text{cov}_{t+1}(c_{t+1}, \mathbf{p}_{t+1}) \\ \mathbf{C}_{t+2} &= \mathbb{M}[c_{t+2} \mathbf{p}_{t+2}] = \bar{a}_{t+2} \Gamma^2 + \text{cov}_{t+2}(c_{t+2}, \mathbf{p}_{t+2})\end{aligned}\quad (59)$$

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<sup>43</sup>In his notation, this is  $\tilde{\psi}^m$ .

but Harmenberg's proof that  $\mathbf{C}_{t+2} - \Gamma \mathbf{C}_{t+1} = 0$  allows us to multiply the first of these by  $\Gamma$  and subtract from the second, yielding (dropping the arguments to the cov terms for clarity):

$$(\bar{c}_{t+2} - \bar{c}_{t+1}) \Gamma^2 = \text{cov}_{t+2} - \Gamma \text{cov}_{t+1}.$$

The less interesting possibility here is when the economy is Szeidl-invariant, in which case  $\bar{c}_{t+2} = \bar{c}_{t+1}$ , implying that the economy exhibits balanced growth in the covariance:

$$\text{cov}_{t+2} = \Gamma \text{cov}_{t+1}. \quad (60)$$

The slightly more interesting case is when the economy is Harmenberg- but not Szeidl-invariant. In that case, if the cov and the  $\mathbb{M}[c]$  terms have constant growth factors  $\Omega_{\text{cov}}$  and  $\Omega_c$ , an equation corresponding to (59) will hold in  $t + n$ :

$$\begin{aligned} (\overbrace{\Omega_c^n \bar{c}_t}^{\bar{c}_{t+n}} - \Omega_c^{n-1} \bar{c}_t) \Gamma^n &= (\Omega_{\text{cov}}^n - \Gamma \Omega_{\text{cov}}^{n-1}) \text{cov}_t \\ (\Omega_c \Gamma)^{n-1} (\Omega_c - 1) \bar{c}_t \Gamma &= \Omega_{\text{cov}}^{n-1} (\Omega_{\text{cov}} - \Gamma) \text{cov}_t \end{aligned} \quad (61)$$

so for the LHS and RHS to grow at the same rates we need

$$\Omega_{\text{cov}} = \Omega_c \Gamma. \quad (62)$$

This is intuitive. In the Szeidl-invariant economy, this just reproduces our result above that the covariance exhibits balanced growth because  $\Omega_c = 1$ . The revised result just says that in the Harmenberg case where the mean value of  $c$  can grow, the covariance must rise in proportion to any ongoing expansion of  $\bar{c}$  as well.

At first blush, these are reassuring conclusions; one of the most persuasive arguments for the agenda of building microfoundations of macroeconomics is that microeconomic 'big data' allow us to measure cross-sectional covariances with great precision, and we can use microeconomic natural experiments to disentangle questions that are hopelessly entangled in aggregate time-series data. Knowing that such covariances ought to be a stable feature of a stably growing economy is therefore encouraging.

But this discussion also highlights an uncomfortable point: In the model as specified, permanent income does not have a limiting distribution; it becomes ever more dispersed as the economy with infinite-horizon consumers continues to exist indefinitely.

A few microeconomic data sources attempt direct measurement of 'permanent income'; Carroll, Slacalek, Tokuoka, and White (2017), for example, show that their assumptions about the magnitude of permanent shocks (and mortality; see below) yield a distribution of permanent income that roughly matches answers in the *Survey of Consumer Finances* to a survey question that aims to elicit a direct measure of permanent income. Those results are used to calibrate a model that matches empirical facts about the distribution of permanent income and wealth also does a reasonable job of matching empirical facts about the marginal propensity to consume. The quantitative credibility of the paper's argument depends on the model's match to the distribution of permanent income inequality, which would not be possible in a model without a nondegenerate steady-state distribution of permanent income.

For the large and growing contingent of macroeconomists who want to build microfoun-

dations by comparing the microeconomic implications of their models to microeconomic data (in levels – not in ratios to permanent income, which is rarely measured in microeconomic datasets other than the SCF), it would be something of a challenge to determine how to construct either Monte Carlo or Markov-matrix-based simulated microeconomic data from a model of the kind sketched above that can be straightforwardly compared to data readily available from the ever-richer microeconomic sources (including, now, national registries from some countries).

### 4.3 Mortality and Redistribution

Death can solve the problem that infinite-horizon models do not have a limiting distribution of permanent income.

#### 4.3.1 *Blanchard Lives*

Most heterogeneous-agent models incorporate a constant positive probability of death, following Blanchard (1985). In the Blanchardian model, if the probability of death exceeds a threshold that depends on the size of the permanent shocks, Carroll, Slacalek, Tokunaka, and White (2017) show that the limiting distribution of permanent income has a finite variance. Blanchard (1985) assumes the existence of a universal annuitization scheme in which estates of dying consumers are redistributed to survivors in proportion to survivors' wealth, giving the recipients a higher effective rate of return. This treatment has several analytical advantages, most notably that the effect of mortality on the time preference factor is the exact inverse of its effect on the (effective) interest factor: If the probability of remaining alive (not dead) is  $\aleph$ , then the assumption that no utility accrues after death makes the effective discount factor  $\hat{\beta} = \beta\aleph$ , while the enhancement to the rate of return from the annuity scheme yields an effective interest rate of  $\hat{R}/\aleph$  (recall that because of white-noise mortality, the average wealth of the two groups is identical). Combining these, the effective patience factor in the new economy  $\hat{\mathbf{P}}$  is unchanged from its value in the infinite horizon model:

$$\hat{\mathbf{P}} \equiv (\beta\aleph R/\aleph)^{1/\rho} = (R\beta)^{1/\rho} \equiv \mathbf{P}. \quad (63)$$

The only adjustments this requires to the analysis above are therefore to the few elements that involve a role for  $R$  distinct from its contribution to  $\mathbf{P}$  (principally, the RIC).

#### 4.3.2 *Modigliani Lives*

Blanchard (1985)'s innovation was useful not only for the insight it provided but also because the principal alternative, the Life Cycle model of Modigliani (1966), was computationally challenging given the then-available technologies. Aside from its (considerable) conceptual value, there is no need for Blanchard's analytical solution today, when serious modeling incorporates uncertainty, constraints, and other features that rule out analytical solutions anyway.



The simplest alternative to Blanchard is to follow Modigliani in assuming that any wealth remaining at death occurs accidentally (not implausible, given the robust finding that for the great majority of households, bequests amount to less than 2 percent of lifetime earnings, Hendricks (2001, 2016)).

Even if bequests are accidental, a macroeconomic model must make some assumption about how they are disposed of: As windfalls to heirs, estate tax proceeds, etc. We again consider the simplest choice, because it again represents something of a polar alternative to Blanchard: Without a bequest motive, there are no behavioral effects of a 100 percent estate tax; we assume such a tax is imposed and that the revenues are effectively thrown in the ocean; the estate-related wealth effectively vanishes from the economy.

The chief appeal of this approach is the simplicity of the change it makes in the condition required for the economy to exhibit a balanced growth equilibrium. If  $\aleph$  is the probability of remaining alive, the condition changes from the plain GIC to a looser mortality-adjusted GIC:

$$\aleph \mathbf{D}_T < 1. \tag{64}$$

With no income growth, the condition required to prohibit unbounded growth in aggregate wealth would be the condition that prevents the per-capita wealth to income ratio of surviving consumers from growing faster than the rate at which mortality diminishes their collective population. With income growth, the aggregate wealth-to-income ratio will head to infinity only if a cohort of consumers is patient enough to make the desired rate of growth of wealth fast enough to counteract combined erosive forces of mortality and productivity.

## 5 Conclusions

Numerical solutions to optimal consumption problems, in both life cycle and infinite horizon contexts, have become standard tools since the first reasonably realistic models were constructed in the late 1980s. One contribution of this paper is to show that finite horizon (‘life cycle’) versions of the simplest such models, with assumptions about income shocks (transitory and permanent) dating back to Friedman (1957) and standard specifications of preferences — and without plausible (but computationally and mathematically inconvenient) complications like liquidity constraints — have attractive properties (like continuous differentiability of the consumption function, and analytical limiting MPC’s as resources approach their minimum and maximum possible values).

The main focus of the paper, though, is on the limiting solution of the finite horizon model as the horizon extends to infinity. The paper shows that the simple model has additional attractive properties: A ‘**Finite Value of Autarky**’ condition guarantees convergence of the consumption function, under the mild additional requirement of a ‘Weak Return Impatience Condition’ that will never bind for plausible parameterizations, but provides intuition for the bridge between this model and models with explicit liquidity constraints. The paper also provides a roadmap for the model’s relationships to the perfect foresight model without and with constraints. The constrained perfect foresight

model provides an upper bound to the consumption function (and value function) for the model with uncertainty, which explains why the conditions for the model to have a nondegenerate solution closely parallel those required for the perfect foresight constrained model to have a nondegenerate solution.

The main use of infinite horizon versions of such models is in heterogeneous agent macroeconomics. The paper articulates intuitive ‘**Growth Impatience Conditions**’ under which populations of such agents, with Blanchardian (tighter) or Modiglianian (looser) mortality will exhibit balanced growth. Finally, the paper provides the analytical basis for many results about buffer-stock saving models that are so well understood that even without analytical foundations researchers uncontroversially use them as explanations of real-world phenomena like the cross-sectional pattern of consumption dynamics in the Great Recession.

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**Table 3** Definitions and Comparisons of Conditions

Perfect Foresight Versions	Uncertainty Versions
Finite Human Wealth Condition (FWHC)	
$\Gamma/R < 1$  The growth factor for permanent income $\Gamma$ must be smaller than the discounting factor $R$ for human wealth to be finite.	$\Gamma/R < 1$  The model's risks are mean-preserving spreads, so the PDV of future income is unchanged by their introduction.
Absolute Impatience Condition (AIC)	
$\mathbf{P} < 1$  The unconstrained consumer is sufficiently impatient that the level of consumption will be declining over time:  $\mathbf{c}_{t+1} < \mathbf{c}_t$	$\mathbf{P} < 1$  <i>If wealth is large enough, the expectation of consumption next period will be smaller than this period's consumption:</i>  $\lim_{m_t \rightarrow \infty} \mathbb{E}_t[\mathbf{c}_{t+1}] < \mathbf{c}_t$
Return Impatience Conditions	
Return Impatience Condition (RIC)	Weak RIC (WRIC)
$\mathbf{P}/R < 1$  The growth factor for consumption $\mathbf{P}$ must be smaller than the discounting factor $R$ , so that the PDV of current and future consumption will be finite:  $c'(m) = 1 - \mathbf{P}/R < 1$	$\wp^{1/\rho} \mathbf{P}/R < 1$  If the probability of the zero-income event is $\wp = 1$ then income is always zero and the condition becomes identical to the RIC. Otherwise, weaker.  $c'(m) < 1 - \wp^{1/\rho} \mathbf{P}/R < 1$
Growth Impatience Conditions	
GIC	GIC-Nrm
$\mathbf{P}/\Gamma < 1$  For an unconstrained PF consumer, the ratio of $\mathbf{c}$ to $\mathbf{p}$ will fall over time. For constrained, guarantees the constraint eventually binds. Guarantees $\lim_{m_t \uparrow \infty} \mathbb{E}_t[\psi_{t+1} m_{t+1}/m_t] = \mathbf{P}\Gamma$	$\mathbf{P} \mathbb{E}[\psi^{-1}]/\Gamma < 1$  By Jensen's inequality stronger than GIC. Ensures consumers will not expect to accumulate $m$ unboundedly.  $\lim_{m_t \rightarrow \infty} \mathbb{E}_t[m_{t+1}/m_t] = \mathbf{P}\Gamma$
Finite Value of Autarky Conditions	
PF-FVAC	FVAC
$\beta \Gamma^{1-\rho} < 1$ equivalently $\mathbf{P} < R^{1/\rho} \Gamma^{1-1/\rho}$  The discounted utility of constrained consumers who spend their permanent income each period should be finite.	$\beta \Gamma^{1-\rho} \mathbb{E}[\psi^{1-\rho}] < 1$  By Jensen's inequality, stronger than the PF-FVAC because for $\rho > 1$ and nondegenerate $\psi$ , $\mathbb{E}[\psi^{1-\rho}] > 1$ .

**Table 4** Conditions for Nondegenerate<sup>‡</sup> Solution

Consumption Model(s)	Conditions	Comments
$\bar{c}(m)$ : PF Unconstrained $\underline{c}(m) = \underline{\kappa}m$  Section 2.4.2: Section 2.4.2: Eq (26): Eq (27):	RIC, FHWCo	RIC $\Rightarrow  v(m)  < \infty$ ; FHWCo $\Rightarrow 0 <  v(m) $ PF model with no human wealth ( $h = 0$ )  RIC prevents $\bar{c}(m) = \underline{c}(m) = 0$ FHWCo prevents $\bar{c}(m) = \infty$ PF-FVAC+FHWCo $\Rightarrow$ RIC GIC+FHWCo $\Rightarrow$ PF-FVAC
$\dot{c}(m)$ : PF Constrained Section 2.4.3:  Appendix A:  Appendix A:	<del>GIC</del> , RIC  GIC, RIC  GIC, <del>RIC</del>	FHWCo holds ( $\Gamma < \mathbf{D} < \mathbf{R} \Rightarrow \Gamma < \mathbf{R}$ ) $\dot{c}(m) = \bar{c}(m)$ for $m > m_{\#} < 1$ ( <del>RIC</del> would yield $m_{\#} = 0$ so $\dot{c}(m) = 0$ ) $\lim_{m \rightarrow \infty} \dot{c}(m) = \bar{c}(m)$ , $\lim_{m \rightarrow \infty} \dot{\kappa}(m) = \underline{\kappa}$ kinks where horizon to $b = 0$ changes* $\lim_{m \rightarrow \infty} \dot{\kappa}(m) = 0$ kinks where horizon to $b = 0$ changes*
$c(m)$ : Friedman/Muth  Section 2.9: Section 2.11.1: Figure 4: Section 2.11.3: Section 2.11.2: Section 3.3: Section 3.3.2: Section 3.3.1:	Section 3.1, Section 3.2 FVAC, WRIC	$\underline{c}(m) < c(m) < \bar{c}(m)$ $\underline{v}(m) < v(m) < \bar{v}(m)$ Sufficient for Contraction WRIC is weaker than RIC FVAC is stronger than PF-FVAC <del>FHWCo</del> +RIC $\Rightarrow$ GIC, $\lim_{m \rightarrow \infty} \kappa(m) = \underline{\kappa}$ <del>RIC</del> $\Rightarrow$ <del>FHWCo</del> , $\lim_{m \rightarrow \infty} \kappa(m) = 0$ “Buffer Stock Saving” Conditions GIC $\Rightarrow \exists 0 < \check{m} < \infty$ GIC-Nrm $\Rightarrow \exists 0 < \hat{m} < \infty$

<sup>‡</sup>For feasible  $m$  satisfying  $0 < m < \infty$ , a nondegenerate limiting consumption function defines a unique optimal value of  $c$  satisfying  $0 < c(m) < \infty$ ; a nondegenerate limiting value function defines a corresponding unique value of  $-\infty < v(m) < 0$ . °RIC, FHWCo are necessary as well as sufficient for the perfect foresight case. \*That is, the first kink point in  $c(m)$  is  $m_{\#}$  s.t. for  $m < m_{\#}$  the constraint will bind now, while for  $m > m_{\#}$  the constraint will bind one period in the future. The second kink point corresponds to the  $m$  where the constraint will bind two periods in the future, etc.\*\*In the Friedman/Muth model, the RIC+FHWCo are sufficient, but *not* necessary for nondegeneracy