# ST5225: Statistical Analysis of Networks Lecture 11: Latent Space Model

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## Review: ERGM



- Model
- Edge prob., MLE
- $\blacksquare$  Example: p1 model

Note: the content in Lecture 11 is not covered in the final exam.

## **ERGM**



Recall our examples for ERGM:

- $\blacksquare$  Random graph model
- Stochastic block model
- many other models...

Question. Are there any structure that ERGM cannot catch?

Answer. Yes!

## Latent Space Model

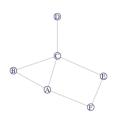


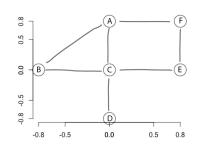
Question: What if the structure in a network is being driven by some 'simpler? form of relationship?

- For stochastic block model, we assume the edges are dependent on the community labels
- Now we generalize the community labels to a broader notion: each node can be represented by a point in the some space, and the edges are dependent on these points
- Given the underlying points, the edges are independent with each other.
- Without the info. of these points, the edges "seems" to be dependent

## Latent Space Model







- Left: a network. Right: Place the network in a 2-dim space
- In this 2-dim space, dist(i,j) < 1 if  $(i,j) \in E$ , and dist(i,j) > 1 if  $(i,j) \neq E$  with one exception.
- The positions of the nodes are the underlying factors that the network is formed
- The positions of each node is unknown

## Latent Space Model



#### These points are unobserved.

- Recall that for the stochastic block model, it is quite usual that the community labels are unknown, and it is hard to estimate.
- For general models, the underlying points are also unobserved (latent)
- So we call it as a latent space model
- How do we model the network with these latent nodes?
- What is the space? Is it 1-dim? 2-dim? or even higher dimensions?
- How to figure out the representation of each node?

## Model assumptions



#### How do we model the network if we have the latent points?

- We are trying to model the edges  $A_{ij}$ , which is either 0 or 1
- The probability of  $A_{ij}$  is related to the latent representation for nodes i and j, denoted by  $Z_i = (z_1^{(i)}, z_2^{(i)}, \dots, z_d^{(i)})$  and  $Z_j = (z_1^{(j)}, z_2^{(j)}, \dots, z_d^{(j)})$ .
- When  $Z_{ij}$ 's are not given, the probability for  $A_{ij}$  is related to other edges
- When  $Z_i$ 's are given, the probability is related to the distance between  $Z_i$  and  $Z_j$

$$\log \left[ \frac{P(A_{ij} = 1 | Z_i = z_i, Z_j = z_j)}{P(A_{ij} = 0 | Z_i = z_i, Z_j = z_j)} \right] = \alpha - ||z_i - z_j||,$$

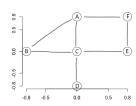
where  $\alpha$  is a parameter to estimate

## Remark



- Note that  $\log \left[ \frac{P(A_{ij}=1|Z_i=z_i,Z_j=z_j)}{P(A_{ij}=0|Z_i=z_i,Z_j=z_j)} \right] z = \alpha \|z_i z_j\|$ . When  $\|z_i z_j\|$  increases, logit $P(A_{ij}=1|Z_i,Z_j)$  decreases, and the probability for connection decreases.
- This model satisfies that the probability for two nodes to be connected is smaller when the latent distance between them is larger.
- Recall our example





■ For this example, when the Euclidean distance between  $z_i$  and  $z_j$  is smaller than 1, the probability of connection is 1. When  $dist(z_i, z_j) > 1$ , the probability of connection is small (but not 0).

### Identification



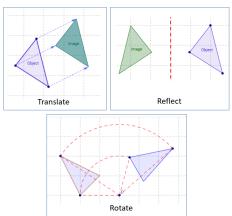
#### Can Z be identified (unique)?

- Why don't we take the function as  $\alpha \beta ||z_i z_j||$ ?
  - If we multiply all the  $z_i$ 's by the same scalar r, and multiply  $\beta$  by 1/r, the result does not change
  - Therefore, to make sure that  $z_i$ 's are unique, we fix  $\beta = 1$ .
- The  $Z_i$ 's are still not identified:
  - Nothing changes if rotate all the  $Z_i$ 's the same way
  - $\blacksquare$  Or if translate all the  $Z_i$ 's along the same vector
  - Or if reflect all the  $Z_i$ 's about the same plane
  - Or combine rotations, translations and reflections

## Isometry



#### **Transformations**



- Isometry = translations which leave all distances the same
  - For Euclidean space, isometry group built from rotations, translations and reflections
- The  $Z_i$ 's are "identified up to isometry"

## Likelihood function



■ Given  $Z_i$ 's, we have that

$$P(A_{ij} = 1 | Z_i = z_i, Z_j = z_j) = \frac{\exp\{\alpha - ||z_i - z_j||\}}{1 + \exp\{\alpha - ||z_i - z_j||\}}$$

So, the joint density function is the likelihood function

$$L(z_{1}, \dots, z_{|V|}, \alpha) = \prod_{i,j} \left[ \frac{\exp\{\alpha - ||z_{i} - z_{j}||\}}{1 + \exp\{\alpha - ||z_{i} - z_{j}||\}} \right]^{A_{ij}} \left[ \frac{1}{1 + \exp\{\alpha - ||z_{i} - z_{j}||\}} \right]^{1 - A_{ij}}$$

$$= \prod_{i,j} \left[ \frac{\exp\{A_{ij}[\alpha - ||z_{i} - z_{j}||]\}}{1 + \exp\{\alpha - ||z_{i} - z_{j}||\}} \right]$$

$$= \frac{\exp\{\sum_{i,j} A_{ij}[\alpha - ||z_{i} - z_{j}||]\}}{\prod_{ij} [1 + \exp\{\alpha - ||z_{i} - z_{j}||\}]}$$

■ Note that it does not belong to ERGM. If we let  $\theta_{ij} = ||z_i - z_j||$ , then the statistics are the adjacency matrix A (all the details are required), and  $\theta_{ij}$ 's are not independent. They have restrictions.

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## MLE



- Assume that the latent variables are "observed", then the likelihood function is about  $\alpha$  only.
- The log-likelihood function is

$$l(\alpha) = \sum_{i,j} A_{ij} [\alpha - ||z_i - z_j||] - \sum_{i,j} \log[1 + \exp{\{\alpha - ||z_i - z_j||\}}]$$

■ Take the derivative w.r.t. 0, and let it equals to 0,

$$\frac{dl(\alpha)}{d\alpha} = \sum_{i,j} A_{ij} - \sum_{i,j} \frac{\exp\{\alpha - \|z_i - z_j\|\}}{1 + \exp\{\alpha - \|z_i - z_j\|\}} = 0.$$

It can be found there is a unique solution for this equation, and that is the MLE.

- However, no explicit formula
- What's more,  $z_i$ 's are unknown....

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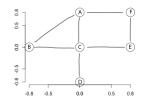
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# Example







Assume the network on the left follows latent variable model, where the link function is

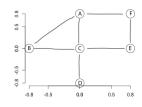
$$P(A_{ij} = 1 | Z_i = z_i, Z_j = z_j) = \frac{\exp\{\alpha - \|z_i - z_j\|\}}{1 + \exp\{\alpha - \|z_i - z_j\|\}}.$$

Assume it is known that the latent points for the nodes are  $z_A = (0, 0.8), z_B = (-0.8, 0), z_C = (0, 0), z_D = (0, -0.8),$   $z_E = (0.8, 0), z_F = (0.8, 0.8)$ . What is the MLE for  $\alpha$ ?

# Example, II







There are 6 nodes, so 15 pairs of nodes in this network. According to the formula for likelihood function, we have

$$L(\alpha) = \frac{\exp\{\sum_{i,j} A_{ij}(\alpha - ||z_i - z_j||)\}}{\prod_{i,j} [1 + \exp\{\alpha - ||z_i - z_j||\}]},$$

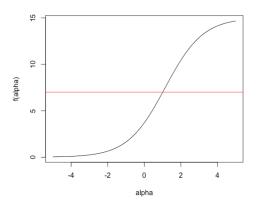
and we should solve the equation

$$\sum_{i,j} \frac{\exp\{\alpha - \|z_i - z_j\|\}}{1 + \exp\{\alpha - \|z_i - z_j\|\}} = \sum_{i,j} A_{ij} = 7.$$

# Example, III



Define  $f(\alpha) = \sum_{i,j} \frac{\exp\{\alpha - \|z_i - z_j\|\}}{1 + \exp\{\alpha - \|z_i - z_j\|\}}$ , we want to find the intersection of it and the horizontal line y = 7.



- The numerical solution is  $\hat{\alpha} = 1.01$ .
- The prob. given is 0.0002050468

#### Estimate



- For random graph model, SBM and ERGM, we calculate MLE
- For this latent space model, MLE is hard to calculate
- We further assume the latent points  $Z_i$ 's also follow some distribution, and include that in the likelihood. In this way,  $z_i$ 's are random variables, not the parameters to estimate
- With data, we are interested in the corrected distribution (posterior) for  $Z_i$ 's, and also the estimate of  $Z_i$ 's and  $\alpha$

## Bayesian Approach



■ Now we assume a prior distribution for the latent points  $Z_i \in \mathcal{R}^d$ 

$$Z_i \overset{i.i.d}{\sim} N(0, I_d), \qquad i = 1, 2, \cdots, |V|,$$

where  $I_d$  is the  $d \times d$  identity matrix, and  $N(0, I_d)$  denotes the multivariate normal distribution.

■ Multivariate Normal distribution: We say  $\mathbf{X} \sim N(\mu, \Sigma)$ , where  $\mathbf{X} = (X_1, X_2, \dots, X_d)'$ ,  $\mu = (\mu_1, \mu_2, \dots, \mu_d)$  is the mean vector, and  $\Sigma$  is a  $d \times d$  covariance matrix.

With these parameters, the joint density is

$$f(x_1, x_2, \dots, x_d) = (2\pi)^{-d/2} |\Sigma|^{-1/2} \exp\{-\frac{1}{2}(\mathbf{x} - \mu)' \Sigma^{-1}(\mathbf{x} - \mu)\},$$

where  $\mathbf{x} = (x_1, x_2, \cdots, x_d)$ .

■ For this example,  $f(x_1, x_2, \dots, x_d) = (2\pi)^{-d/2} \exp\{-\frac{1}{2}(\mathbf{x}'\mathbf{x})\}.$ 

#### Joint distribution



- To figure out the joint dist., usually we assume the latent space is known, say, d = 2.
- Therefore, we have the joint dist.,

$$P(A) = \left[\frac{\exp\{\sum_{i,j} A_{ij} [\alpha - \|\mathbf{z_i} - \mathbf{z_j}\|]\}}{\prod_{i,j} [1 + \exp\{\alpha - \|\mathbf{z_i} - \mathbf{z_j}\|\}]}\right] \prod_{i} [(2\pi)^{-d/2} \exp\{-\frac{1}{2} (\mathbf{z_i}' \mathbf{z_i})\}]$$

- The only unknown parameter is  $\alpha$ , and obviously it does not belong to ERGM.
- To solve the problem, we may
  - For given  $\alpha$ , find the posterior and the estimate for  $Z_i$ 's
  - With the estimation of  $Z_i$ 's, find the MLE of  $\alpha$
  - Repeat until convergence



#### Generalization on the link function:

■ Recall that we link the probability with the function

$$logit P(A_{ij} = 1 | z_i, z_j, \alpha) = \alpha - ||z_i - z_j||.$$

This is the most basic one. According to the data, we may consider some other choices.

■ For some data sets, such as our toy example, the nodes can be embedded in a way that the relationship between  $dist(Z_i, Z_j)$  and 1 largely impacts the probability of connection. For this case, a proper function is

$$logit P(A_{ij} = 1 | z_i, z_j, \alpha) = \alpha (1 - ||z_i - z_j||).$$

■ Projection methods. For some data sets, it can be found that node i is connected to many nodes, and node j is connected to a small subset of the neighbors of node i. In this case, we model that both i and j are "similar" but i is more "socially active".

$$logit P(A_{ij} = 1 | z_i, z_j, \alpha) = \alpha + \frac{z_i' z_j}{\|z_i\|}.$$



#### Additional covariate information:

- Some networks contain the covariate information matrix X. The covariate information may also impact the probability of connection.
- In this case, we model it as

$$logit P(A_{ij} = 1 | z_i, z_j, \alpha) = \alpha + \beta' x_{i,j} - ||z_i - z_j||,$$

where  $x_{i,j}$  is the covariate information.

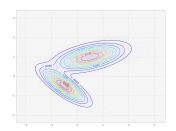
Of course, the covariate information can be added to other linkage functions:

$$logit P(A_{ij} = 1 | z_i, z_j, \alpha) = \alpha + \beta x_{i,j} + \frac{z_i' z_j}{\|z_j\|}.$$



#### Prior distribution:

- In our model, we assume the prior dist. for  $Z_i$ 's are multivariate normal dist.
- Combine it with the community recovery problem, then we assume the latent points also have community structure



■ The latents points follow a mixture normal dist., where

$$Z_i \sim \sum_{k=1}^K \lambda_k N(\mu_k, \sigma_k^2 I_d), \qquad \sum_{k=1}^K \lambda_k = 1.$$



#### Generalization on the distance measurement:

- In the model, we consider the Euclidean distance between two latent points.
- With the projection method, we consider the projection of one latent vector on the other latent vector
- There are more possibilities:
  - Geodesic distance: for points on some surfaces, other than the Euclidean distance, there is also the geodesic distance, which measures the distance on the surface, e.g., the airlines



■ The latent space is a manifold

## Summary



- Latent space model assume latent points for every node. The prob. of connection depends on the latent points only.
- The model based on the latent points varies. We mainly discuss the distance model
- MLE is hard to calculate; latent points are hard to get
- Bayesian approach
- Generalizations