

RESEARCH REVIEW

Net effects of multiple stressors in freshwater ecosystems: a meta-analysis

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Abstract

The accelerating rate of global change has focused attention on the cumulative impacts of novel and extreme environmental changes (i.e. stressors), especially in marine ecosystems. As integrators of local catchment and regional processes, freshwater ecosystems are also ranked highly sensitive to the net effects of multiple stressors, yet there has not been a large-scale quantitative synthesis. We analysed data from 88 papers including 286 responses of freshwater ecosystems to paired stressors and discovered that overall, their cumulative mean effect size was less than the sum of their single effects (i.e. an antagonistic interaction). Net effects of dual stressors on diversity and functional performance response metrics were additive and antagonistic, respectively. Across individual studies, a simple vote-counting method revealed that the net effects of stressor pairs were frequently more antagonistic (41%) than synergistic (28%), additive (16%) or reversed (15%). Here, we define a reversal as occurring when the net impact of two stressors is in the opposite direction (negative or positive) from that of the sum of their single effects. While warming paired with eutrophication resulted in additive net effects, the overall mean net effect of warming combined with a second stressor was antagonistic. Most importantly, the mean net effects across all stressor pairs and response metrics were consistently antagonistic or additive, contrasting the greater prevalence of reported synergies in marine systems. Here, a possible explanation for more antagonistic responses by freshwater biota to stressors is that the inherent greater environmental variability of smaller aquatic ecosystems fosters greater potential for acclimation and co-adaptation to multiple stressors.

Keywords: antagonism, biodiversity, climate change, cumulative impacts, ecological surprises, functional resistance, reversals, synergy

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Introduction

The rise of ‘ecological surprises’ in the primary scientific literature highlights the growing uncertainty over the cumulative impacts of multiple novel and extreme environmental changes, or ‘stressors’ (e.g. Paine *et al.*, 1998; Christensen *et al.*, 2006; Lindenmayer *et al.*, 2010; Dehedin *et al.*, 2013; Harvey *et al.*, 2013). There is increasing evidence from marine environments that these stressors, such as rising temperatures, biological invasions and habitat destruction, act synergistically to exacerbate biodiversity loss and ecological degradation (Crain *et al.*, 2008; Harvey *et al.*, 2013; Przeslawski *et al.*, 2015). Interactions among stressors are at the core of these unexpected net ecological impacts (Sala *et al.*, 2000) as they can generate complex effects that lessen or amplify the direct single effect of each

stressor. The reported prevalence of nonadditive effects of stressors across many marine ecosystems (Crain *et al.*, 2008; Darling & Cote, 2008; Harvey *et al.*, 2013; Ban *et al.*, 2014) attests to an urgent need to fill knowledge gaps in freshwater ecosystems (Root *et al.*, 2003; Ormerod *et al.*, 2010; Staudt *et al.*, 2013; Hering *et al.*, 2015).

Empirical evidence of the net effects of multiple stressors on freshwaters remains very limited (but see Christensen *et al.*, 2006; Darling & Cote, 2008; Mantyka-Pringle *et al.*, 2014) despite their impacts being greatest on freshwater biodiversity (Jenkins, 2003; WWF, 2014). Freshwater ecosystems are particularly vulnerable to global change (Dudgeon *et al.*, 2006; Ormerod *et al.*, 2010) as they often occupy low points in landscapes, integrating the effects of local catchment and regional atmospheric processes (Williamson *et al.*, 2009). In comparison, recent meta-analyses of the marine literature show that the net impact of multiple stressors are frequently either greater than (i.e. a synergistic interaction; Crain *et al.*, 2008; Harvey *et al.*, 2013) or equal to (i.e. an

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additive effect; Ban *et al.*, 2014; Strain *et al.*, 2014) the sum of their single effects. Net effects of two or more stressors that were less than the potential additive outcome (i.e. an antagonistic interaction) were less common (Crain *et al.*, 2008; Harvey *et al.*, 2013). Such variation in the net effects of stressor combinations depends in part on how impact is measured, as different biological receptors will inherently vary in their responsiveness to environmental change (termed response diversity; Elmquist *et al.*, 2003). For example, compensatory species dynamics within a stressed community may result in measurable changes in biodiversity while muting changes in function (e.g. primary production; Vinebrooke *et al.*, 2003).

Theoretical models that predict the combined impact of stressor pairs on populations or communities are often based on an evaluation of the similarity of their independent impacts (Vinebrooke *et al.*, 2004). For instance, if stressors A and B are highly redundant and both extirpate or negatively influence the same set of species in a community, then their net impact on species richness or functional performance (e.g. productivity or abundance) should be less than the sum of their independent effects (an antagonistic interaction). In contrast, synergy between stressors A and B can occur if species are affected only upon exposure to both stressors, resulting in their combined impact being greater than the sum of their single effects (a synergistic interaction). If stressor A affects a different set of species than stressor B, then their net impact on the community can equal the sum of their direct effects (an additive effect). In some cases, the net effect of stressors A and B may actually be in the opposite direction (positive or negative) than predicted based on their independent effects (Piggott *et al.*, 2015). For instance, Christensen *et al.* (2006) found that warming reversed the positive effect of acidification on phytoplankton. We term such interactions as 'reversals', perhaps representing the greatest of all 'ecological surprises'.

Here, we synthesize findings from dual-stressor studies in freshwater ecosystems to address two main questions: (i) What is the cumulative mean interaction and frequency of interaction types across all studies? and (ii) How do interactions vary among response metrics and stressor pairs? We also focused on how higher temperatures associated with climate change interact with other key stressors to impact ecosystem properties. We used a meta-analytical approach to optimize our ability to both conduct a powerful quantitative test of the nature of interactions between stressors affecting freshwater ecosystems and identify testable hypotheses (Gurevitch *et al.*, 2000; Parmesan *et al.*, 2013; Hillebrand & Gurevitch, 2014).

Materials and methods

Data selection

We searched the primary scientific literature and identified papers in which the impacts of multiple stressors were compared, both in combination and alone, to a nonstressed control (see Data S1 for full search terms and methods). Reported stressors included acidification, higher temperatures, ultraviolet radiation (UVR), contamination (xenobiotics or salinity), eutrophication, habitat alteration (physical manipulation, sedimentation, altered flow regime or drought) and invasive species. We considered the following response currencies or metrics: (i) survival, (ii) growth/size, (iii) condition, (iv) reproductive, (v) behaviour, (vi) total biomass/abundance, (vii) diversity and (viii) leaf decomposition.

We used the term 'observation' to refer to individual responses used in our analyses, and the term 'paper' to refer to their source documents. In several cases, multiple observations were extracted from individual papers when either several experiments were conducted (i.e. using different sets of species, study locations or stressor combinations) or various organismal groups were measured (e.g. producers, invertebrates or vertebrates). If the response of a specific organismal group to dual stressors during a single experiment was assessed using multiple metrics (e.g. plant biomass and plant diversity), then we treated each as an independent observation for inclusion only in our 'full data set' ($n = 286$). The full data set was then used for our mixed effects response metric meta-analyses (detailed and pooled; Table 1). For the remainder of our comparisons, we excluded all diversity metrics ($n = 31$) and reduced our data set to include only the most inclusive response metrics per experiment for each organismal group. For experiments where multiple response metrics were reported, the most inclusive response metric was selected where community responses were preferred over population or organism-level responses, and metrics were selected in favour of biomass/abundance over survival, survival over growth/size, growth/size over condition, condition over reproductive and reproductive over behaviour. However, if the same experiment measured impact separately on multiple organism groups (e.g. producers and invertebrates), then each observation was retained. This 'most inclusive response metric data set' ($n = 230$) was used for the majority of our meta-analyses (i.e. those not specifically comparing response metrics; Table 1) to minimize data nonindependence. See Table S1 (Data S2) for a complete list of observations included in each data set. For each observation/stressor response, we extracted mean, standard deviation and sample size values for each treatment combination (stressor A; stressor B; stressor A and B; no stressor control). We also collected relevant categorical data (e.g. location and response metric used to measure impact) for each observation (Table S1).

Effect size calculations

Interaction effect sizes were calculated for each observation in our data set using Hedges d , an estimate of the standardized

Table 1 Data sets used for each categorical analysis (meta-analytic and vote counting) and the levels of each category (where $n \geq 8$). See Table S2 in Data S2 for full model terms

Data set	Categorical analyses
Full data set ($n = 286$)	Detailed response metric: Animal survival; animal growth/size; plant growth/size; animal condition; animal biomass/abundance; plant biomass/abundance; animal diversity; plant diversity; leaf decomposition
Full data set ($n = 286$)	Pooled response metric: Diversity; functional performance
Most inclusive response metric data set ($n = 230$)	Level of biological organization: Community; population; organism
Most inclusive response metric data set ($n = 230$)	Organism group: Vertebrate; invertebrate; producer
Most inclusive response metric data set ($n = 230$)	Stressor pair: Contamination \times Habitat Alteration; Contamination \times Invasion; Contamination \times Nutrification; Contamination \times Warming; Habitat Alteration \times Nutrification; Invasion \times Invasion; Invasion \times Nutrification; Nutrification \times UVR; Nutrification \times Warming; Warming \times UVR

mean difference not biased by small sample sizes (Gurevitch & Hedges, 2001). The interaction effect size for each observation was calculated by comparing the null predicted additive effect to the actual observed effect of both stressors. Each interaction effect size was therefore based on the absolute difference between the observed net impact of dual stressors against a hypothetical additive outcome based on the sum of their single independent effects (see Data S1 for equation details). We inverted the response direction (\mp) of interaction effect sizes for which the additive effects were negative (i.e. where both single effects were negative, or if in opposing directions, where the negative effect had the higher absolute value; Piggott *et al.*, 2015). This allowed us to compare interaction effect sizes regardless of their directionality (Piggott *et al.*, 2015). This means that an effect size (d) of zero represents an exact additive effect of the two stressors (i.e. their combined impact is equal to the sum of their single effects), while a positive d denotes a synergistic interaction (a combined impact greater than the sum of their single effects) and a negative d reflects either antagonism or a reversal interaction (a combined impact less than the sum of their single effects; Fig. 1). To distinguish between antagonistic and reversal interactions, we compared the direction (negative or positive, relative to the control) of the observed response to both stressors applied in combination with the direction of their predicted additive response, and assigned reversals where they were opposite. Interaction significance was assessed using 95% confidence intervals calculated around each effect size, such that any interactions with intervals crossing zero were deemed additive (Fig. 1).

Statistical analyses

Mean interaction effect sizes across studies were estimated from weighted meta-analyses. In each analysis, 'Observation ID' was treated as a random effect to account for the random component of effect size variation among observations and

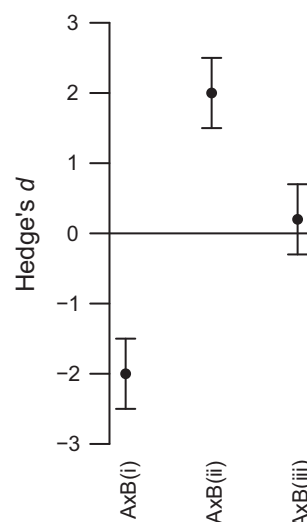


Fig. 1 The theoretical interactive effects of stressors A and B applied in combination, relative to their predicted additive response ($= 0$). Negative effect sizes (less than zero) represent antagonism or reversals (i) and positive effect sizes (greater than zero) represent synergistic interactions (ii), but only if their confidence intervals do not cross the x-axis. Interaction effect sizes with confidence intervals that overlap with zero were considered to be additive (iii).

calculate inverse unconditional variance effect size weights (Gurevitch & Hedges, 2001; see Data S1 for equations and model details). In addition to using random effects meta-analyses to assess the global mean interaction effect sizes across all observations included in our 'full' and 'most inclusive response metric' data sets, we conducted a series of mixed effects meta-analyses where selected categorical moderators were treated as fixed effects to assess mean interactions at each level of each category (where $n \geq 8$; see Table S2 in Data S2 for model terms).

Using our 'full data set', we conducted a detailed response metric analysis to evaluate the sensitivity of different response metrics to multiple stressors (Table 1). We followed this with a pooled response metric analysis, where response metrics were reassigned as either 'diversity' (plant or animal diversity) or 'functional performance' (all other response metrics considered), to assess the sensitivities of these broader response categories. We then used our reduced 'most inclusive response metric data set' to estimate mean effect sizes across receptor categories (response levels and organism groups) and stressor-pair combinations (Table 1). Percentile bootstrapped 95% confidence intervals were calculated around each mean interaction effect size to assess significance (Fig. 1). Similar to the assessment of interaction effect sizes for single observations, a positive mean effect reflects synergy, a negative mean effect reflects antagonism (reversals could not be distinguished with this method) and cases where the confidence intervals crossed zero were deemed additive.

In addition to the quantitative synthesis described above, we complemented each meta-analytic model with a vote-counting analysis to describe the frequencies of interaction types (including reversals) across individual observations. Randomization tests of independence (Monte Carlo approximation using 9999 permutations) were used to assess whether the frequencies of interaction types differed significantly among levels of each categorical moderator where $n \geq 8$ (Table 1).

Weighted meta-analyses were conducted in METAWIN version 2.1 (Rosenberg *et al.*, 2000), and the R computing program was used to perform independence tests and create figures (R Core Team, 2014). To assess the robustness of our results, we conducted several additional analyses to investigate potential publication bias and the sensitivity of our findings to variation in sample sizes and effect size outliers (Data S3). Although we found some evidence of asymmetry around our overall mean effect size estimate, we suspect this may be at least partially attributable to the considerable data heterogeneity observed. Nevertheless, the results of our sensitivity analyses indicate that our meta-analytic findings are robust to such variations.

Stressor interactions across response metrics

We found 88 articles representing 286 separate observations or biological responses to multiple stressors that met our selection criteria (Table S1). In addition, 11 articles fitting our criteria were not included because we were unable to extract the data or the study did not report margins of error (listed in Data S2). The majority of the research was carried out in North America (46 of 88 articles), followed by Europe (30) and New Zealand (7). All of the studies were conducted experimentally in laboratories (57), outdoor mesocosms (210) or *in situ* (19).

Individual observations in our full data set were most frequently antagonistic (40%; compared with 26% synergistic, 19% additive and 15% reversed), and the mean interaction effect size across all responses was

also significantly less than additive (i.e. antagonistic; Table S2). Multiple stressors exerted significant antagonistic effects on animal abundance/biomass, animal condition, animal growth/size, animal survival and plant diversity (Fig. 2a). Additive mean stressor effects were identified for the other four response metrics (decomposition, animal diversity, plant abundance/biomass and plant growth/size; Fig. 2a).

One possible explanation for widespread antagonistic interactions between freshwater stressors involves asymmetry of their single effect sizes. Here, the larger magnitude of the worst stressor completely overrides

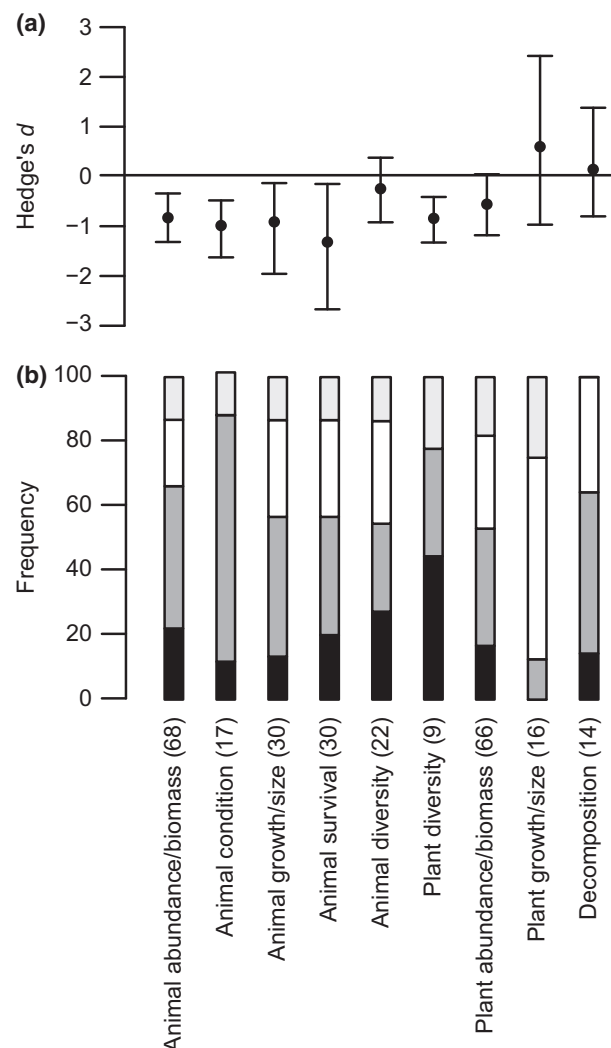


Fig. 2 The mean interaction effect sizes (Hedge's *d* and bootstrapped 95% confidence intervals; (a) and frequencies (%) of interaction types (b) for different response metric categories. Interaction types are additive (black), antagonistic (dark grey), synergistic (white) and reversals (light grey). The number of observations/studies included in each category is indicated in parentheses. Mean responses only presented where $n \geq 8$.

the effect of the weaker stressor, thereby negating its contribution to their net impact (Folt *et al.*, 1999; Sala *et al.*, 2000). The detected prevalence of antagonisms also suggests that exposure to one stressor often results in greater tolerance to the other (Vinebrooke *et al.*, 2004). Here, a potential mechanism involves hard selection for tolerant organisms that are co-adapted to both stressors, thereby reducing their combined impact. Alternatively, acclimation to each stressor may involve the same behavioural or physiological mechanism, which would result in exposure to one stressor inducing greater tolerance against the other.

Frequencies of interaction types varied significantly ($\chi^2 = 40.36$; $P = 0.019$; $df = 24$; $n = 272$) and nonadditive interactions were collectively more common than simple additive scenarios. Antagonisms occurred most often with animal condition (76.47%), synergies and reversals with plant growth/size (62.50% and 25.00%, respectively), and additive effects with plant diversity (44.44%; Fig. 2b). The highly variable nature of stressor interactions across these response metrics highlights the importance of currency selection when quantifying the net ecological impact of multiple stressors.

Stressors also exerted differing interactive and additive effects on functional performance and diversity responses, respectively. The mean interaction effect size for functional performance responses was antagonistic, while the mean effect of stressors on diversity was additive. Additive and reversal interactions occurred most frequently with diversity metrics (32.25% and 16.13%, respectively), while antagonistic and synergistic interactions occurred more frequently with functional performance metrics (41.57% and 27.06%, respectively); however, the frequencies of interaction types did not differ significantly ($\chi^2 = 4.87$, $P = 0.174$; $df = 3$; $n = 286$).

Compensatory species dynamics may explain the different mean interactive effects observed for stressor impacts on freshwater diversity and functional performance. The frequency of additive responses by diversity to dual stressors suggests that species eliminated by one stressor were often not the same that are eliminated by a second stressor. However, the prevalence of antagonism at the functional performance level suggests the remaining tolerant species may often compensate functionally for species loss, thereby reducing the net functional consequences of the stressors. Although the prevalence of functional species compensation has been debated in the literature (Houlahan *et al.*, 2007; Gonzalez & Loreau, 2009), several lines of evidence show it can help stabilize stressed freshwater communities (e.g. Klug *et al.*, 2000; Fischer *et al.*, 2001; Vinebrooke *et al.*, 2003; Downing *et al.*, 2008). Our findings support how functional resistance to stressors is not

simply a function of biodiversity, but more often indicative of species identity and associated traits (e.g. Smith & Knapp, 2003; Vaz-Pinto *et al.*, 2013). Thus, functional resistance should be related to the response diversity and functional redundancy within stressed communities (Elmqvist *et al.*, 2003; Nyström, 2006; Mori *et al.*, 2012). As a result, our findings point to freshwater biodiversity being more sensitive than functioning to the cumulative impacts of multiple stressors.

Stressor interactions across receptor categories

For analyses of receptor categories and stressor pairs (see following section), we considered only the most inclusive response metrics to avoid pseudoreplication. As a result, our data set was reduced to 230 observations for these analyses (Table 1; Table S1). The majority of the observations examined responses at the community level and the most frequently examined organisms were invertebrates (Fig. 3). The global mean interaction effect size was significantly antagonistic (Table S2) and of the 230 observations considered, 94 (40.87%) were antagonistic, 64 (27.83%) were synergistic, and 34 (14.78%) were reversals, while 38 (16.52%) were additive.

The cumulative mean interaction effect of stressors was significantly antagonistic at the community and organismal level but additive at the population level (Fig. 3a; Table S2). However, the frequencies of interaction types did not differ significantly among levels of biological organization ($\chi^2 = 11.39$; $P = 0.074$; $df = 6$; $n = 230$). While antagonistic interactions were most frequent at the organismal (65.22%) and community (40.88%) levels of biological organization, synergies and reversals occurred most frequently at the population level (37.14% and 17.14%, respectively) and additive interactions were most common at the community level (18.98%; Fig. 3b).

Dual stressors exerted significant antagonistic effects on invertebrates and vertebrates, while primary producers responded in an overall additive fashion (Fig. 3c; Table S2). However, frequencies of interaction types were similar across all organismal groups ($\chi^2 = 5.70$; $P = 0.457$; $df = 6$; $n = 224$). Antagonistic responses occurred most frequently for invertebrates (45.21%) and vertebrates (46.43%), synergies and reversals were most common with primary producers (34.74% and 16.84%, respectively), and additive interactions most often affected invertebrates (19.18%; Fig. 3d). These results were surprising because sensitivity to global change is often thought to increase with trophic position (e.g. Crain *et al.*, 2008; Petchey *et al.*, 1999), particularly with warming, as metabolic demands increase faster than ingestion rates with

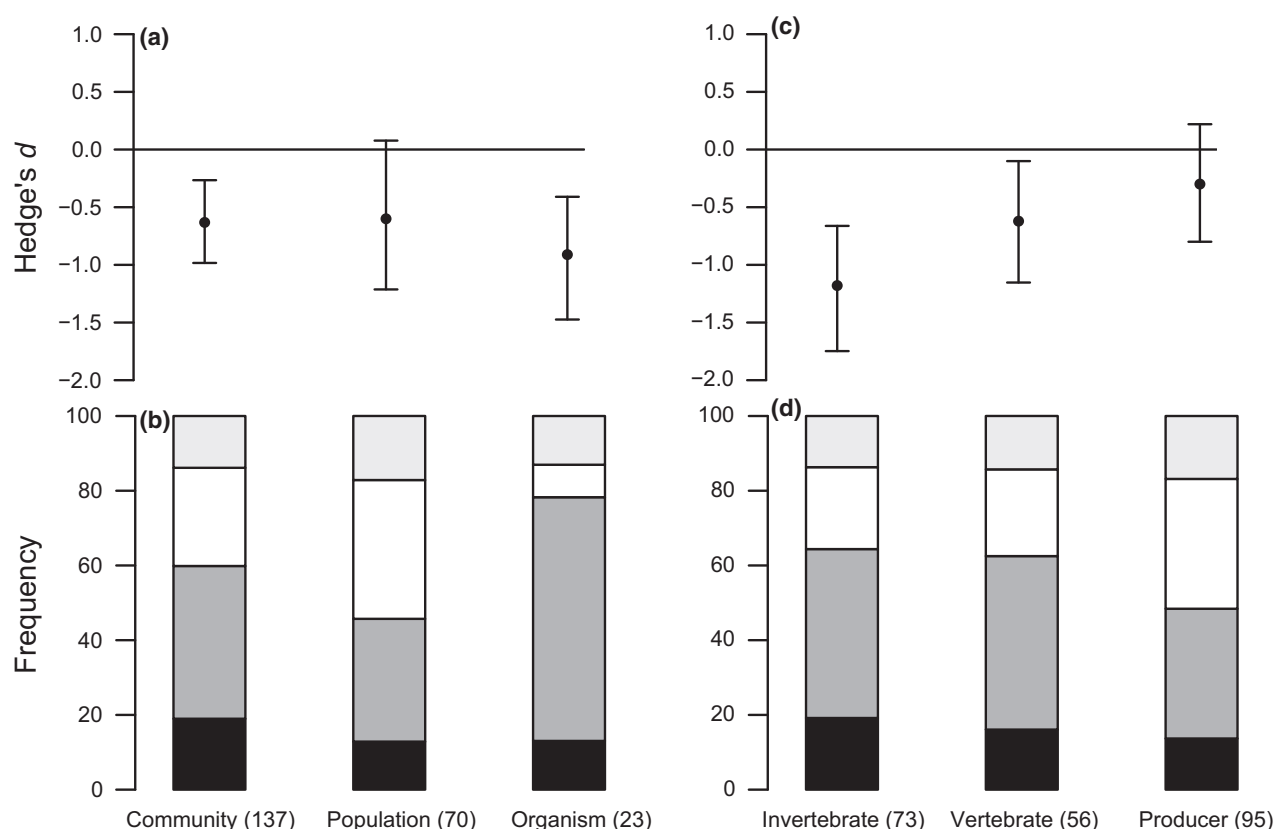


Fig. 3 The mean interaction effect sizes (Hedge's *d* and bootstrapped 95% confidence intervals; a, c) and frequencies (%) of interaction types (b, d) for different receptor categories, including level of biological organization (a, b) and organism group (c, d). Interaction types are additive (black), antagonistic (dark grey), synergistic (white) and reversals (light grey). The number of observations/studies included in each category is indicated in parentheses. Mean responses only presented where $n \geq 8$.

higher temperatures (Vucic-Pestic *et al.*, 2011). Here, the different responses of consumers and primary producers highlight the potential for multiple stressors to weaken trophic interactions, and promote algal blooms. Many of the synergistic responses by primary producers involved net positive effects by stressors such as eutrophication, UVR and warming. In fact, 36 of the 64 synergistic interactions in our analysis were positive, and of these, 21 showed an increase in producer performance. Globally, correlative evidence suggests that nutrients and climate interact synergistically to increase

the overall percentage of cyanobacteria in shallow lakes (Kosten *et al.*, 2012). Experimental evidence supports these observations, showing warming and nutrient enrichment can exert a synergistic positive effect on phytoplankton growth (e.g. Doyle *et al.*, 2005).

Stressor interactions across stressor pairs

Ten stressor pairs had sufficient observations ($n \geq 8$) for a comparison of their mean interaction effects (Table 2), which varied with their identity (Fig. 4a). Net

Table 2 The number of independent observations/studies meeting our criteria used in the stressor-pair analysis ($n = 230$)

	Acidification	Contamination	Habitat Alteration	Invasion	Nutritification	UVR	Warming
Acidification	0	3	2	0	0	3	5
Contamination		6	19	11	14	6	33
Habitat alteration			4	2	21	1	6
Invasion				13	10	0	7
Nutritification					0	10	41
UVR						0	13
Warming							0

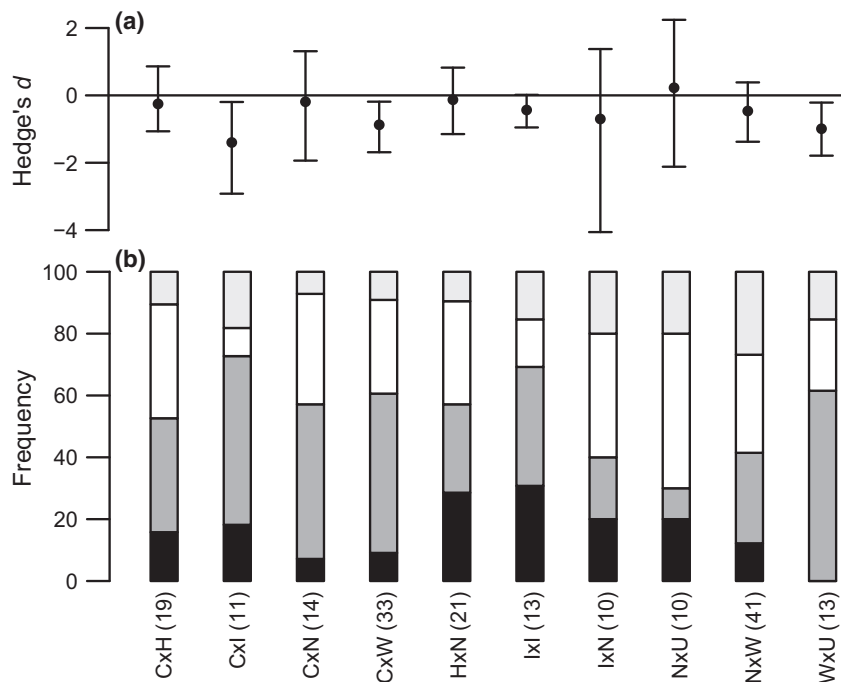


Fig. 4 The mean interaction effect sizes (Hedge's *d* and bootstrapped 95% confidence intervals; a) and frequencies (%) of interaction types (b) for different stressor-pair combinations. Interaction types are additive (black), antagonistic (dark grey), synergistic (white) and reversals (light grey). The number of observations/studies included in each category is indicated in parentheses. Mean responses only presented where $n \geq 8$. W = warming; C = contamination; H = habitat alteration; I = invasion; N = eutrophication; and U = ultraviolet light radiation.

effects were significantly antagonistic for contamination \times invasion, contamination \times warming and warming \times UVR; however, effects were additive for the remaining seven stressor pairs, including eutrophication paired with warming, habitat alteration, invasion and UVR (Fig. 4a). Although the frequencies of interaction types were not significantly different among stressor pairs ($\chi^2 = 28.25$; $P = 0.402$; $df = 27$; $n = 185$), antagonistic effects occurred most frequently when warming occurred with UVR (61.54%), synergistic interactions occurred most often with eutrophication and UVR (50.00%), reversal interactions were linked with warming and eutrophication (26.83%), and additive interactions were common with paired invasions (30.77%; Fig. 4b).

When higher temperature interacted with a second freshwater stressor, the mean interaction was antagonistic overall ($d = -0.68$; 95% CI = -1.1 to -0.3 ; $n = 105$). This finding is in contrast to studies of marine ecosystems where both Crain *et al.* (2008) and Harvey *et al.* (2013) found that warming most often interacted with a second stressor to produce a synergistic response. However, a recent re-analysis of the data presented by Crain *et al.* (2008) suggests that their original methods may have overrepresented synergies (Piggott *et al.*, 2015). Furthermore, Ban *et al.* (2014) found that the mean effect of multiple stressors in coral reefs

was additive overall, and it is important to note that different ecosystem types face different combinations of key stressors (Jenkins, 2003; Pratchett *et al.*, 2011). Lake (1990) suggested that benthic communities in freshwater and marine ecosystems may react differently to certain disturbances because of differences in the proportion of mobile versus sedentary biota. More general differences between freshwater and marine responses may be based on how specific stressors interact with inherent ecosystem properties. For example, Bancroft *et al.* (2007) predicted that UVR impacts should vary between marine and freshwater environments owing to differing optical qualities of the water; however, they were unable to detect significant differences from their meta-analysis. Additionally, the effects of some stressors (e.g. salinity and metal contaminants) may differ among freshwater and marine receptors based on physiological differences between biota (Hall & Anderson, 1995; Heugens *et al.*, 2001).

Higher environmental variability of smaller aquatic ecosystems may also foster greater species adaptation to stressors. Freshwaters generally experience much greater thermal variation than marine systems, so freshwater ectotherms might be better adapted to temperature changes than those from more thermally buffered marine ecosystems. For example, water fleas (*Daphnia*

spp.) that are often focal species in lakes and ponds have been shown to be highly responsive (Colbourne *et al.*, 2011) and capable of rapidly evolving in the face of environmental change (De Meester *et al.*, 2011). Aquatic organisms also tend to be most sensitive to multiple-stressor effects near their thermal tolerance limits (Heugens *et al.*, 2001), so more detrimental stressor interactions might be expected in marine ecosystems where species' ranges are often strongly aligned with their thermal limits (Pratchett *et al.*, 2011; Sunday *et al.*, 2012). Indeed, differences in how marine and freshwater ecosystems respond to similar stressors may depend on characteristics of the biological receptors and the environmental context, including the different communities, mechanisms, ecological networks and abiotic conditions present (Bancroft *et al.*, 2007; Tylanakis *et al.*, 2008; Segner *et al.*, 2014).

Three stressor-pair combinations had sufficient samples sizes ($n \geq 8$ for receptor categories within stressor pairs) for detailed analysis of interaction effects by level of biological organization or organismal type. The mean interaction effect size remained significantly additive for nitrification paired with warming or habitat alteration (Fig. 4a) regardless of level of biological organization or organism group. Contamination and warming had a significant antagonistic interaction overall (Fig. 4a) and at the organismal level ($d = -0.77$; 95% CI = -1.3 to -0.3 ; $n = 10$); however, the interaction was additive at the population ($d = -1.27$; 95% CI = -3.6 to 0.4 ; $n = 11$) and community ($d = -0.26$; 95% CI = -0.7 to 0.2 ; $n = 12$) levels. Similarly, the mean interaction between contamination and warming became additive when considering only studies which measured impacts on vertebrates ($d = -0.26$; 95% CI = -1.0 to 0.5 ; $n = 12$). These results suggest that the type of organism and level of biological organization are both important in determining and predicting the combined effects of specific stressor pairs.

Reversal interactions as extreme ecological surprises

Reversals (similar to 'mitigating synergisms' discussed by Piggott *et al.*, 2015) were found in 34 of 230 observations (14.78%) included in our stressor-pair analysis (i.e. the most inclusive response metric data set). Although they were the least common type of interaction detected in our data set of most inclusive end-points, reversal interactions warrant special consideration because they represent net effects that may differ markedly from those predicted by the typically assumed model of additivity (Piggott *et al.*, 2015). Reversal interactions often involve the weaker of two stressors inverting the effect of the strongest. For instance, application of excess nutrients surprisingly reversed

the toxic effect of atrazine on tadpoles as the additional resources likely permitted greater detoxification rates and stimulated growth, resulting in increased survival (Boone & Bridges-Britton, 2006).

Our findings showed that the stressor most commonly associated with reversal interactions was warming (19.05% of warming interactions; Fig. 4b). The greater likelihood of reversal interactions when a stressor is paired with higher temperatures might be related to the stimulatory effect of warming. As nearly all biological activity increases with warming (Brown *et al.*, 2004), temperature changes arguably have the greatest potential to mediate the effects of other more damaging stressors. For example, Thompson *et al.* (2008) found that warming reversed the negative effect of excess nitrogen supply on growth by alpine phytoplankton, possibly because higher temperatures stimulated enzymatic conversion of nitrate and ammonia. In contrast, Linton *et al.* (1997) showed that higher temperatures could reverse the stimulatory effects of sublethal ammonia enrichment on juvenile rainbow trout (*Oncorhynchus mykiss*) by increasing metabolic costs to where ammonia detoxification and growth rates were reduced. In these cases, warming directly altered the mechanisms by which the dominant stressors affected the biological receptors. However, like other nonadditive scenarios, reversals may also manifest from complex indirect interactions (e.g. Messner *et al.*, 2013). Given the complexity of ecological responses to temperature changes (Petchey *et al.*, 1999; O'Connor *et al.*, 2009; Dossena *et al.*, 2012; Stendera *et al.*, 2012) and their potential role in generating nonadditive interactions with other stressors (Crain *et al.*, 2008; Harvey *et al.*, 2013), we might then expect even more 'ecological surprises' in a warmer future.

Conclusions

We discovered a prevalence of antagonistic interactions between freshwater stressors across most receptor categories considered in our analysis (Table S2). Thus, there may exist a high potential for co-adaptation within freshwater ecosystems to minimize the net effects of multiple stressors. Alternatively, antagonism may be attributable to a high degree of asymmetry in the magnitude of independent effects between freshwater stressors (Folt *et al.*, 1999). In this case, ranking the worst stressor driving an antagonistic interaction would be essential to forecasting their cumulative impacts on a freshwater ecosystem (Sala *et al.*, 2000; Piggott *et al.*, 2015). However, our evidence of predominantly antagonistic responses by freshwater organisms should not lessen the need to reduce exposure to stressors as their net effects were still mostly negative. The urgency of

these findings is underscored by a recent global assessment that compared multiple-stressor-induced average population declines of 76% among freshwater species to 39% among terrestrial and marine species since 1970 (WWF, 2014).

Nonadditive interactions characterized 83% (192/230) of the cumulative impacts of multiple stressors in our most inclusive response metric data set (81% or 233/286 in our full data set). Mean interaction effect sizes varied significantly among stressor pairs and levels of receptor categories. Our analyses revealed different interactions for some stressor pairs (switching from antagonistic to additive, or vice versa) when only considering subsets of the data. This suggests that both stressor identity and characteristics of the ecological response (e.g. level of biological organization and organism type) are essential in predicting interactions between multiple stressors in freshwater ecosystems.

Our findings have implications for conservation management of freshwater ecosystems. For stressor pairs that generate additive or synergistic effects, management focusing on a single stressor should render a positive outcome (Brown *et al.*, 2013). However, in communities affected antagonistically by stressor pairs, both stressors may need to be removed or moderated to produce any substantial ecological recovery due to positive cotolerance (Brown *et al.*, 2013; Piggott *et al.*, 2015).

Our findings evoke several testable hypotheses for further investigation. Firstly, the observed trend of stressor synergies increasing the productivity of primary producers suggests that higher temperatures, UVR exposure and nutrient enrichment may jointly stimulate harmful algal blooms. Secondly, functional performance metrics appeared less sensitive overall than diversity metrics to dual stressors, highlighting the need for further investigation into the extent to which functional compensation occurs in stressed ecosystems. Thirdly, although we have demonstrated a clear predominance of antagonistic stressor interactions in freshwaters, further studies are needed to determine the specific underlying ecological mechanisms (e.g. asymmetry of stressor magnitudes, hard selection for co-adapted organisms, or similarity in behavioural or physiological acclimation). Finally, perhaps most interesting is our finding that multiple-stressor interactions differ between freshwaters and marine ecosystems and, although we have suggested several potential explanations, more research is needed to elucidate the specific physiological, genetic or environmental drivers behind these differences.

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Supporting Information

Additional Supporting Information may be found in the online version of this article:

Data S1. Detailed methods: Details of data search and selection criteria, effect size calculation and interpretation, weighted meta-analyses and vote-counting methodology.

Data S2. Meta-Analysis Tables and References: Interaction effect sizes and related information for each study used in our meta-analyses (Table S1), mean interaction effect sizes and details of meta-analytic models used in our analyses (Table S2) and a detailed list of data references.

Data S3. Robustness of meta-analytic results: Detailed analysis of potential publication bias and sensitivity of our results to variations in sample sizes and effect size outliers.