LECTURE 6: CORRELATION AND LINEAR REGRESSION BASICS

ECON 480 - ECONOMETRICS - FALL 2018

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September 17, 2018







 $\boldsymbol{\cdot}$ We looked at single variables for descriptive statistics



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- $\cdot \ \mathsf{Most} \ \mathsf{uses} \ \mathsf{of} \ \mathsf{statistics} \ \mathsf{in} \ \mathsf{economics} \ \mathsf{and} \ \mathsf{business} \ \mathsf{investigate} \ \mathsf{relationships} \ \mathsf{between} \ \mathsf{variables}$



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 - Immediate aim is to explore associations between variables, quantified with correlation and linear regression
 - · Later we want to develop more sophisticated tools to argue for causation



BIVARIATE DATA: SPREADSHEETS

econfreedom<-read.csv("~/Dropbox/Teaching/Hood College/ECON 480 - Econometrics/[
head(econfreedom)</pre>

##	ISO.Code	Country	${\tt Economic.Freedom.Summary.Index}$	GDP.Per.Capita
## 1	AGO	Angola	5.08	4153.146
## 2	ALB	Albania	7.40	4543.088
## 3	ARE	Unit. Arab Em.	7.98	39313.274
## 4	ARG	Argentina	4.81	10501.660
## 5	ARM	Armenia	7.71	3796.517
## 6	AUS	Australia	7.93	54688.446



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- · Rows are individual observations
- · Columns are variables on all individuals
- Let X be Economic Freedom and Y be GDP per capita



```
str(econfreedom)
```

\$ GDP.Per.Capita

##

'data.frame': 152 obs. of 4 variables:

```
## $ ISO.Code : Factor w/ 152 levels "AGO", "ALB", "ARE",...
## $ Country : Factor w/ 152 levels "Albania", "Algeria",
## $ Economic.Freedom.Summary.Index: num 5.08 7.4 7.98 4.81 7.71 7.93 7.56 6.
```

: num 4153 4543 39313 10502 3797 ...



summary(econfreedom)

##

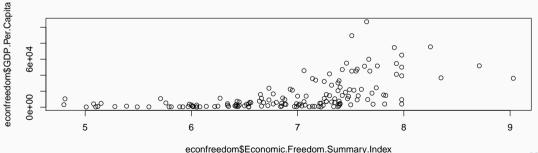
Median :

5719.3

```
ISO.Code
                                 Economic.Freedom.Summarv.Index
##
                      Country
                 Albania : 1
##
   AG0
             1
                                 Min.
                                        :4.800
##
   ALB
             1
                 Algeria : 1 1st Qu.:6.430
                 Angola : 1
                                 Median :7.050
##
   ARE
             1
             1
                 Argentina: 1
##
   ARG
                                 Mean
                                        :6.909
            1
##
   ARM
                 Armenia : 1
                                 3rd Ou.:7.428
##
   AUS
          : 1
                 Australia:
                                 Max.
                                        :9.030
    (Other):146
                 (Other) :146
##
##
   GDP.Per.Capita
##
   Min.
              206.7
##
   1st Qu.:
             1588.3
```

BIVARIATE DATA: SCATTERPLOTS

```
# syntax for plotting is similar to hist() and boxplot()
# just tell R "plot(df$x,df$y)"
plot(econfreedom$Economic.Freedom.Summary.Index, econfreedom$GDP.Per.Capita)
```

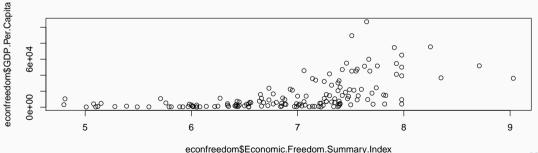




• The best way to visualize an association between two variables is with a scatterplot

BIVARIATE DATA: SCATTERPLOTS II

```
# syntax for plotting is similar to hist() and boxplot()
# just tell R "plot(df$x,df$y)"
plot(econfreedom$Economic.Freedom.Summary.Index, econfreedom$GDP.Per.Capita)
```

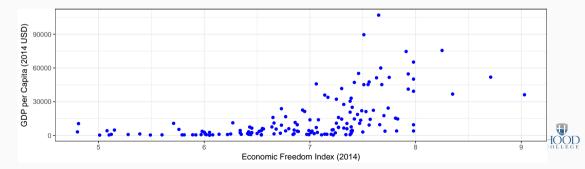




• Each point is a pair of variable values (X_i, Y_i) for observation i

BIVARIATE DATA: A BETTER-LOOKING SCATTERPLOT (WITH ggplot2)

```
library("ggplot2")
ggplot(econfreedom, aes(x=Economic.Freedom.Summary.Index,y=GDP.Per.Capita))+
   geom_point(color="blue")+theme_bw()+
   xlab("Economic Freedom Index (2014)")+ylab("GDP per Capita (2014 USD)")
```



 $\boldsymbol{\cdot}$ Look for $\boldsymbol{association}$ between independent and dependent variables



- $\boldsymbol{\cdot}$ Look for $\boldsymbol{association}$ between independent and dependent variables
 - 1. *Direction*: is the trend positive or negative?



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- Look for association between independent and dependent variables
 - 1. Direction: is the trend positive or negative?
 - 2. Form: is the trend linear, quadratic, something else, or no pattern?
 - 3. Strength: is the association strong or weak?
 - 4. Outliers: do any observations break the trends above?



$$s_{X,Y} = E[(X - \overline{X})(Y - \overline{Y})]$$



 $^{^{1}}$ Henceforth we limit to samples, for convenience. Population covariance is denoted $\sigma_{\rm X,Y}$

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 - to be above its mean also (X and Y covary positively)



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- Intuition: if X is above its mean, would we expect Y:
 - to be above its mean also (X and Y covary positively)
 - to be below its mean (X and Y covary negatively)
- Covariance is a common measure, but the units are meaningless, thus we rarely need to use it so don't worry about learning the formula



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CORRELATION

• More convenient to standardize covariance into a more intuitive concept: correlation (ρ or r), normalized to be between -1 and 1

$$r_{X,Y} = \frac{s_{X,Y}}{s_X s_Y} = \frac{cov(X,Y)}{sd(X)sd(Y)}$$



• More convenient to standardize covariance into a more intuitive concept: correlation (ρ or r), normalized to be between -1 and 1

$$r_{X,Y} = \frac{s_{X,Y}}{s_X s_Y} = \frac{cov(X,Y)}{sd(X)sd(Y)}$$

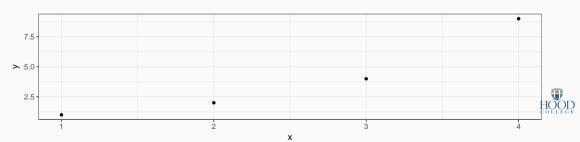
• Alternatively, sample correlation can be found by standardizing (finding the *Z*-score) X and Y and multiplying, for each (X, Y) pair, and then averaging (over n-1, due to sampling df, again):

$$r = \frac{1}{n-1} \sum_{i=1}^{n} \left(\frac{X_i - \bar{X}}{s_X} \right) \left(\frac{Y_i - \bar{Y}}{s_Y} \right)$$
$$= \frac{1}{n-1} \sum_{i=1}^{n} Z_X Z_Y$$



CORRELATION: EXAMPLE CALCULATION

Example



CORRELATION: EXAMPLE CALCULATION II

```
mean(corr.example$x) #find mean of x

## [1] 2.5

mean(corr.example$y) #find mean of y
```

sd(corr.example\$x) #find sd of x

Su(colling county in line su or x

sd(corr.example\$v) #find sd of v



[1] 1.290994

[1] 4

```
## x y z.product
## 1 1 1 0.9793959
## 2 2 2 0.2176435
## 3 3 4 0.0000000
## 4 4 9 1.6323265
```

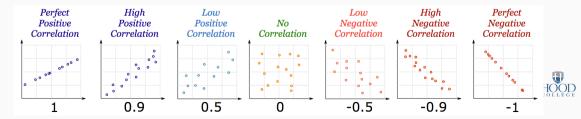


CORRELATION: EXAMPLE CALCULATION IV

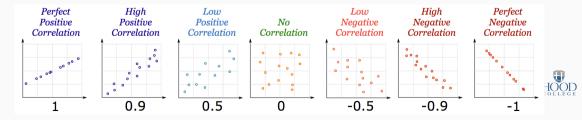
```
(sum(corr.example$z.product)/3) #average z products over n-1
## [1] 0.943122
cor(corr.example$x, corr.example$y) #compare our answer to cor() command
## [1] 0.9431191
cov(corr.example$x, corr.example$y) #just for kicks - covariance
```



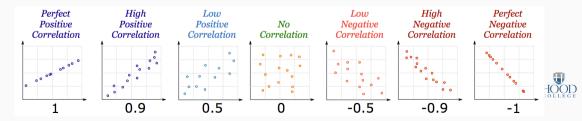
· Correlation is standardized to $-1 \le r \le 1$



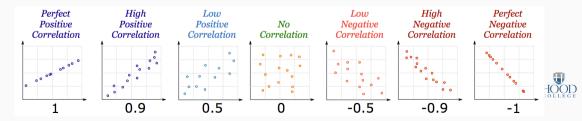
- Correlation is standardized to $-1 \le r \le 1$
 - \cdot Negative values \implies negative association



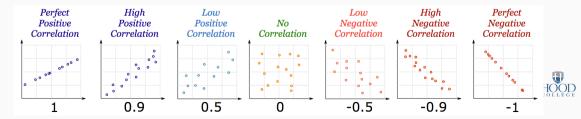
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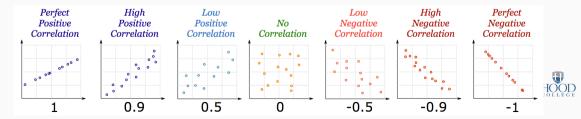
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 - \cdot As $|r| \to 1 \Longrightarrow$ the stronger the association



- Correlation is standardized to $-1 \le r \le 1$
 - \cdot Negative values \implies negative association
 - \cdot Positive values \implies positive association
 - · Correlation of $0 \implies$ no association
 - · As $|r| \rightarrow 1 \implies$ the stronger the association
 - · Correlation of $|r| = 1 \implies$ a perfect linear relationship



GUESS THE CORRELATION!



DEW GOME TWO PLOTERS SGORE BOORD OBOUT SETTINGS

HIGH SCORE (1)

Guess The Correlation Game



CORRELATION AND ENDOGENEITY

• Reminder: Correlation does not imply causation!



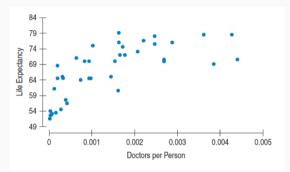
CORRELATION AND ENDOGENEITY

- · Reminder: Correlation does not imply causation!
- · See the **Handout** for more on Covariance and Correlation



CORRELATION AND ENDOGENEITY II

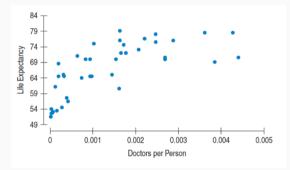
Example



• The correlation between Life Expectancy and Doctors Per Person is 0.705.



Example



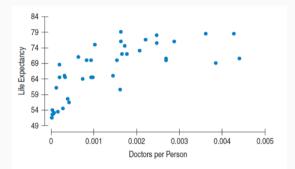






CORRELATION AND ENDOGENEITY II

Example

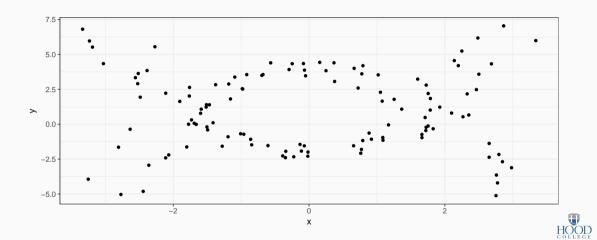




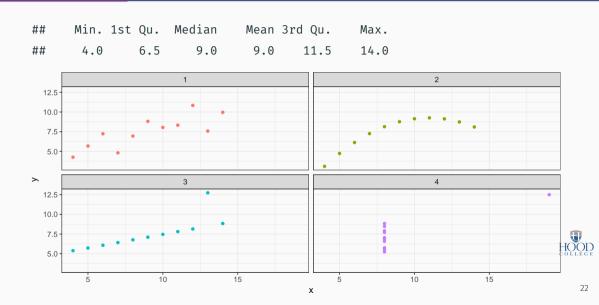




ALWAYS PLOT YOUR DATA!



ANSCOMBE'S QUARTET

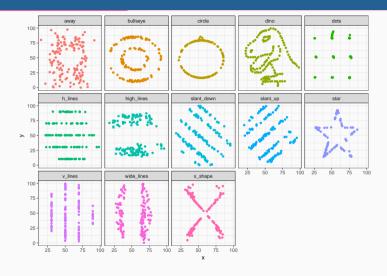


ANSCOMBE'S QUARTET: A MODERN RE-INTERPRATATION

```
##
     dataset
                           Х
   Length: 1846
                     Min. :15.56
                                     Min.
                                            : 0.01512
##
##
   Class :character
                     1st Qu.:41.07 1st Qu.:22.56107
   Mode :character
                     Median :52.59 Median :47.59445
##
##
                      Mean :54.27
                                     Mean :47.83510
                      3rd Qu.:67.28 3rd Qu.:71.81078
##
##
                            :98.29
                                     Max.
                                            :99.69468
                      Max.
```



ANSCOMBE'S QUARTET: A MODERN RE-INTERPRATATION II





See the Datasaurus

Model

LINEAR REGRESSION

• If an association appears linear, we can estimate the equation of a line that would "fit" the data



LINEAR REGRESSION

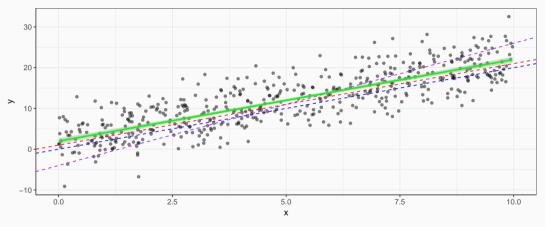
• If an association appears linear, we can estimate the equation of a line that would "fit" the data

$$Y = a + bX$$

- Recall a linear equation describing a line contains: - a: vertical intercept - b: slope - Note we will use different symbols for a and b, in line with standard econometric notation



LINEAR REGRESSION II



 $\boldsymbol{\cdot}$ How do we choose the equation that best fits the data? Process is called $\mbox{linear regression}$



 $\boldsymbol{\cdot}$ Linear regression lets us estimate the slope of the population regression line between X and Y



- \cdot Linear regression lets us estimate the slope of the population regression line between X and Y
- \cdot We can make **inferences** about the population slope coefficient



- · Linear regression lets us estimate the slope of the population regression line between X and Y
- · We can make **inferences** about the population slope coefficient
 - · eventually, a causal interpretation



- Linear regression lets us estimate the slope of the population regression line between X and Y
- · We can make **inferences** about the population slope coefficient
 - · eventually, a causal interpretation
 - slope $=\frac{\Delta Y}{\Delta X}$: for a 1-unit change in X, how many units will this cause Y to change?



 $\boldsymbol{\cdot}$ Statistically, we want to use the population regression model for:



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- 1. Estimation of the marginal effect of X on Y (slope of population regression line)



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- 2. Hypothesis Testing of the value of the marginal effect (slope)



- Statistically, we want to use the population regression model for:
- 1. Estimation of the marginal effect of X on Y (slope of population regression line)
- 2. Hypothesis Testing of the value of the marginal effect (slope)
- 3. Confidence Interval construction of a range for the true effect (slope)



AN EXTENDED EXAMPLE

Example

What is the relationship between class size and educational performance?

• Policy question: What is the effect of reducing class sizes by 1 student per class on test scores? 10 students?



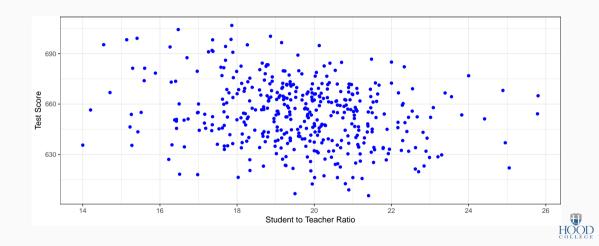


AN EXTENDED EXAMPLE: SCATTERPLOT

```
library("foreign") #for importing .dta files
CASchool<-read.dta("~/Dropbox/Teaching/Hood College/ECON 480 - Econometrics/Data
ca.scatter<-ggplot(CASchool, aes(str,testscr))+
   geom_point(color="blue",fill="blue")+
   xlab("Student to Teacher Ratio")+
   ylab("Test Score")+theme_bw()</pre>
```



AN EXTENDED EXAMPLE: SCATTERPLOT II



AN EXTENDED EXAMPLE: SLOPE

• If we *change* (Δ) the class size by an amount, what would we expect the *change* in test scores to be?

$$\beta_{\textit{ClassSize}} = \frac{\textit{change in test score}}{\textit{change in class size}} = \frac{\Delta \textit{test score}}{\Delta \textit{class size}}$$



AN EXTENDED EXAMPLE: SLOPE

• If we *change* (Δ) the class size by an amount, what would we expect the *change* in test scores to he?

$$\beta_{\textit{ClassSize}} = \frac{\textit{change in test score}}{\textit{change in class size}} = \frac{\Delta \textit{test score}}{\Delta \textit{class size}}$$

• If we knew $\beta_{ClassSize}$, we could say that changing class size by 1 student will change test scores by $\beta_{ClassSize}$



AN EXTENDED EXAMPLE: SLOPE II

· Rearranging:

$$\Delta$$
test score $=eta_{ extit{ClassSize}} imes \Delta$ class size



AN EXTENDED EXAMPLE: SLOPE II

· Rearranging:

$$\Delta$$
test score = $\beta_{ClassSize} \times \Delta$ class size

· Suppose $\beta_{ClassSize} = -0.6$. If we shrank class size by 2 students, our model predicts:



AN EXTENDED EXAMPLE: SLOPE II

· Rearranging:

$$\Delta$$
test score = $\beta_{\text{ClassSize}} imes \Delta$ class size

· Suppose $\beta_{ClassSize} = -0.6$. If we shrank class size by 2 students, our model predicts:

$$\Delta$$
test score = -0.6

$$\Delta$$
test score = \times - 2 = 1.2



test score =
$$\beta_0 + \beta_{ClassSize} \times$$
 class size

 $\boldsymbol{\cdot}$ The line relating class size and test scores has the above equation



test score
$$= eta_0 + eta_{ ext{ClassSize}} imes$$
 class size

- The line relating class size and test scores has the above equation
 - + $eta_{ ext{0}}$ is the vertical-intercept, test score where class size is 0



test score
$$= \beta_0 + \beta_{\text{ClassSize}} \times \text{class size}$$

- The line relating class size and test scores has the above equation
 - \cdot β_0 is the vertical-intercept, test score where class size is 0
 - \cdot $\beta_{ extit{ClassSize}}$ is the slope of the regression line



test score
$$= \beta_0 + \beta_{\text{ClassSize}} \times \text{class size}$$

- The line relating class size and test scores has the above equation
 - \cdot eta_0 is the vertical-intercept, test score where class size is 0
 - \cdot $\beta_{ extit{ClassSize}}$ is the slope of the regression line
- This relationship only holds **on average** for all districts in the population, individual districts are also affected by other factors



 \cdot To get an equation that holds for each district, we need to include other factors

test score
$$= \beta_0 + \frac{\beta_{\textit{ClassSize}}}{\beta_0} \times \text{class size} + \text{other factors}$$



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test score
$$= \beta_0 + \frac{\beta_{\textit{ClassSize}}}{\beta_0} \times \text{class size} + \text{other factors}$$

- · For now, we will ignore these until the next lesson
- · Thus, $\beta_0 + \beta_{\text{ClassSize}} imes$ class size gives the average effect of class sizes on scores



• To get an equation that holds for *each* district, we need to include other factors

test score =
$$\beta_0 + \beta_{ClassSize} \times class size + other factors$$

- · For now, we will ignore these until the next lesson
- \cdot Thus, $eta_0 + eta_{ extsf{ClassSize}} imes$ class size gives the **average effect** of class sizes on scores
- Later, we will want to estimate the marginal effect (causal effect) of each factor on an individual district's test score, holding all other factors constant



$$y = \beta_0 + \beta_1 x_1 + \beta_2 x_2 + \epsilon$$



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- · y is the dependent variable of interest
 - · AKA "response variable," "regressand," "Left-hand side (LHS) variable"



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- x_1 and x_2 are independent variables
 - AKA "explanatory variables," "regressors," "Right-hand side (RHS) variables," "covariates," "control variables"



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- $\cdot \beta_0, \beta_1$, and β_2 are unknown parameters to estimate



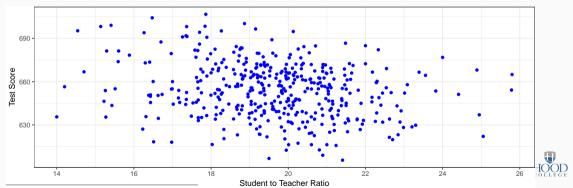
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- We have observed values of y, x_1 , and x_2 & "regress y on x_1 and x_2 "
- $\cdot \beta_0, \beta_1$, and β_2 are unknown parameters to estimate
- \cdot ϵ is the error term
 - It is **stochastic** (random)
 - · We can never measure the error term



THE POPULATION REGRESSION MODEL

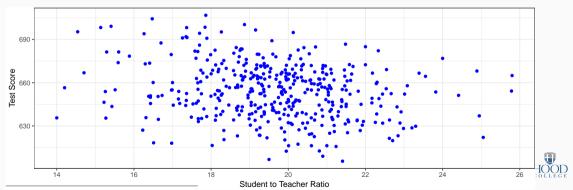
 \cdot How do we draw a line through the scatterplot? We do not know the true $eta_{ extit{ClassSize}}$



²Data is student-teacher-ratio and average test scores on Stanford 9 Achievement Test for 5th grade students for 420 K-6 and K-8 school districts in California in 1999, (Stock and Watson, 2015: p. 141)

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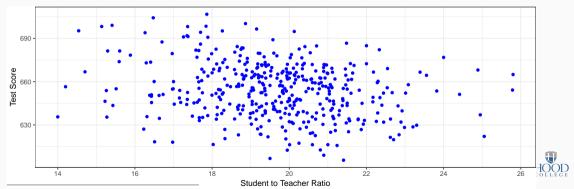
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- We do have data from a *sample* of class sizes and test scores²



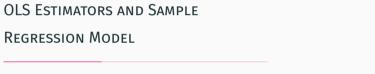
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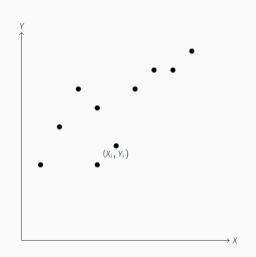
THE POPULATION REGRESSION MODEL

- \cdot How do we draw a line through the scatterplot? We do not know the true $eta_{ extit{ClassSize}}$
- We do have data from a sample of class sizes and test scores²
- So the real question is, how can we estimate β_0 and β_1 ?



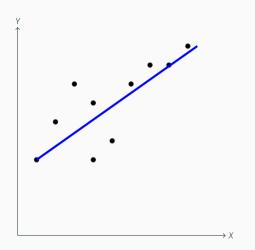
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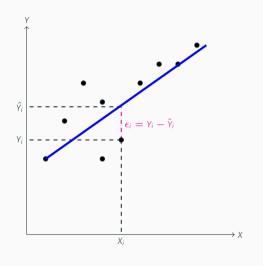
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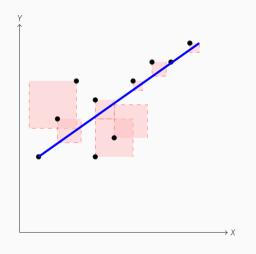




- Suppose we have a scatter plot of points (X_i, Y_i)
- · We can draw a "line of best fit" through our scatterplot
- The residual (ϵ_i) of each data point is the difference between actual and predicted value of Y given X

$$\epsilon_i = Y_i - \hat{Y}_i$$





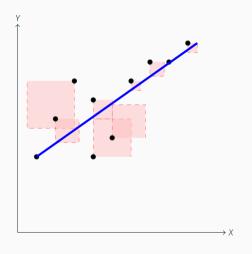
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· The line of best fit minimizes SSE

• I coded an example (using an application of R called shiny) to demonstrate how OLS tries to solve the problem by picking optimal line parameters



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$$\min_{\beta_0,\beta_1} \sum_{i=1}^n [Y_i - (\underbrace{\beta_0 + \beta_1 X_i}_{\hat{Y}_i})]^2$$

• OLS estimators minimize the average squared distance between the actual values (Y_i) and the predicted values (\hat{Y}_i) along the estimated regression line



• The OLS regression line or sample regression line is the linear function constructed using the OLS estimators:

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- The **predicted value** of Y given X, based on the regression, is $E(Y_i|X_i) = \hat{Y}_i$
- The **residual** or **prediction error** for the i^{th} observation is the difference between observed Y_i and its predicted value, $\hat{\epsilon_i} = Y_i \hat{Y}_i$



 $\cdot\,$ The solution to the SSE minimization problem yields: 3



³See **Handout** on Blackboard for proofs.

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- For $\hat{\beta}_0$:

$$\hat{\beta}_0 = \bar{Y} - \hat{\beta}_1 \bar{X}$$



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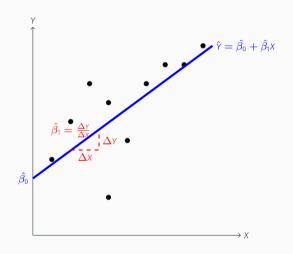
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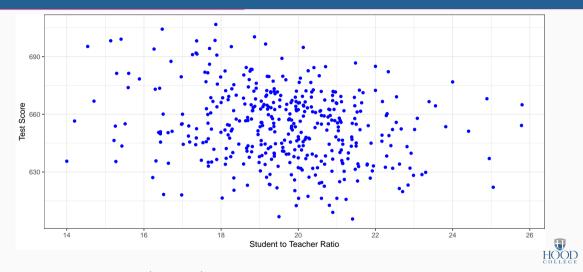


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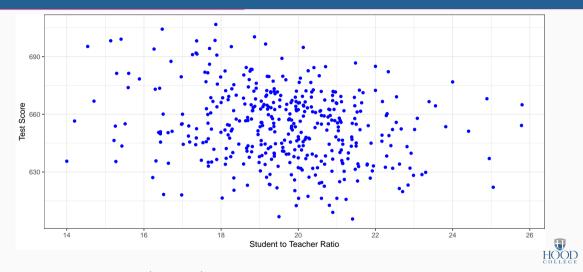


OLS Example: Class Size



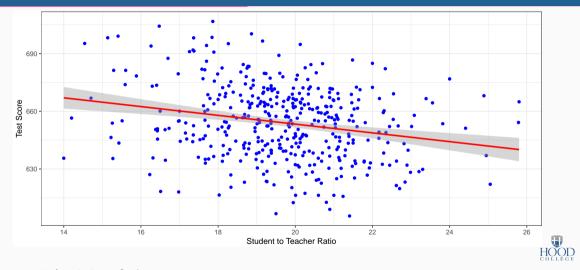
• There is some true (unknown) population relationship:

OLS Example: Class Size



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OLS Example: Class Size: OLS Estimation



 $\boldsymbol{\cdot}$ Using OLS, we find:

Test Score =
$$689.9 - 2.28 \times STR$$

$$\cdot$$
 Estimated slope: $\hat{eta}_1 = rac{\Delta ext{test score}}{\Delta ext{STR}} = -2.28$



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- Estimated slope: $\hat{\beta}_1 = \frac{\Delta \text{test score}}{\Delta \text{STR}} = -2.28$ Estimated intercept: $\hat{\beta}_0 = 689.9$



$$\widehat{\text{Test Score}} = 689.9 - 2.28 \times STR$$

- \cdot Estimated slope: $\hat{eta}_1 = rac{\Delta ext{test score}}{\Delta ext{STR}} = -2.28$
- \cdot Estimated intercept: $\hat{eta}_0 =$ 689.9
 - · Not always economically meaningful



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- Estimated slope: $\hat{eta}_1 = \frac{\Delta {\rm test \, score}}{\Delta {\rm STR}} = -2.28$
- \cdot Estimated intercept: $\hat{eta_0}=$ 689.9
 - · Not always economically meaningful
 - · Literally: "districts with 0 students have a predicted test score of 689.9"



OLS EXAMPLE: CLASS SIZE: PREDICTIONS

$$\widehat{\text{Test Score}} = 689.9 - 2.28 \times STR$$

• We can now make simple predictions with our model:



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 - $\boldsymbol{\cdot}$ For a district with 20 students per teacher, the predicted test score is:

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OLS EXAMPLE: CLASS SIZE: PREDICTIONS

Test Score =
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- $\boldsymbol{\cdot}$ We can now make simple predictions with our model:
 - For a district with 20 students per teacher, the predicted test score is:

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• Is this big or small? How **economically** meaningful is 644?



• Syntax for running a regression in **R** is simple:

```
# name an object e.g. "regression.name", "lm" stands for "linear model"
regression.name<-lm(y~x, data=data.frame.name)

# get simple (beta) coefficients by calling the object
regression.name

# get more detailed information with summary()
summary(regression.name)</pre>
```



```
school.regression<-lm(testscr~str, data=CASchool)</pre>
school.regression
##
## Call:
## lm(formula = testscr ~ str, data = CASchool)
##
## Coefficients:
## (Intercept)
                        str
        698.93
                      -2.28
##
```

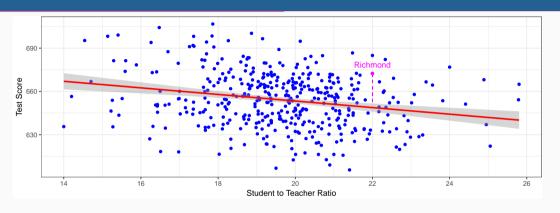


OLS EXAMPLE: CLASS SIZE: IN R II

summary(school.regression)

```
##
## Call:
## lm(formula = testscr ~ str. data = CASchool)
##
## Residuals:
##
      Min
          10 Median 30 Max
## -47.727 -14.251 0.483 12.822 48.540
##
## Coefficients:
             Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 698.9330 9.4675 73.825 < 2e-16 ***
       -2.2798 0.4798 -4.751 2.78e-06 ***
## str
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 18.58 on 418 degrees of freedom
## Multiple R-squared: 0.05124, Adjusted R-squared: 0.04897
## F-statistic: 22.58 on 1 and 418 DF. p-value: 2.783e-06
```

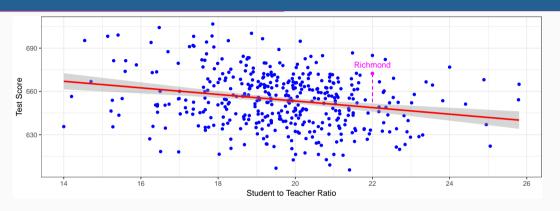
THE SAMPLE OLS REGRESSION MODEL: A DATA POINT



 $\boldsymbol{\cdot}$ One district in our sample is Richmond, CA with STR=22, Test Score=672



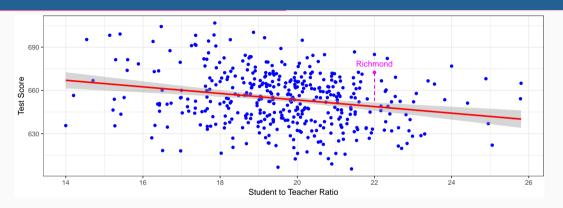
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THE SAMPLE OLS REGRESSION MODEL: A DATA POINT



- \cdot One district in our sample is Richmond, CA with STR=22, Test Score=672
- \cdot Predicted value: Test $\widehat{\text{Score}_{\textit{Richmond}}} = 698 2.28(22) \approx 647$
- Residual: $\widehat{\epsilon_{Richmond}} = 672 647 = 25$

