LECTURE 8: PRECISION OF OLS AND HYPOTHESIS TESTING

ECON 480 - ECONOMETRICS - FALL 2018

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September 26, 2018



The Precision of OLS

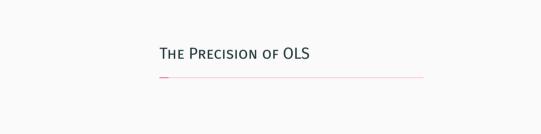
Hypothesis Testing About Regression

Digression: p-Values and the Philosophy of Science

Back to Our Hypothesis Test: The Test-Statistic

Reporting Regression Outputs with ${\tt stargazer}$

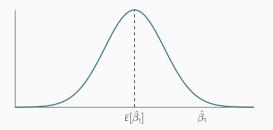




Recall: The Sampling Distribution of $\hat{eta}_{\!\scriptscriptstyle 1}$

$$\hat{eta}_1 \sim N(E[\hat{eta}_1], \sigma_{\hat{eta}_1})$$

· We want to know:

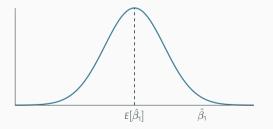




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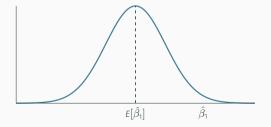




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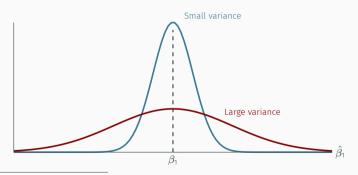
- · We want to know:
 - \cdot $E[\hat{eta}_1]$; what is the center of the distribution?
 - \cdot $\sigma_{\hat{eta}_{\!\scriptscriptstyle 1}}$; how precise is our estimate?





Precision: Variance or Standard Error

• How precise is our estimate $\hat{\beta}_1$?

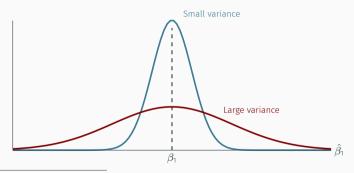


¹The "standard **error**" is the analogue of standard *deviation* for a sample statistic's sampling distribution. Recall the sampling distribution is the distribution of a statistic, like \bar{X} or $\hat{\beta}_1$ over many potential samples.



PRECISION: VARIANCE OR STANDARD ERROR

- How precise is our estimate $\hat{\beta}_1$?
- · We can talk of the variance $(\sigma_{\hat{\beta}_1}^2)$ or the standard error $(\sigma_{\hat{\beta}_1})$ of $\hat{\beta}_1^{1}$



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Variance of \hat{eta}_1

. The variance of $\hat{\beta}_1$ is

$$var(\hat{\beta}_1) = \frac{(SER)^2}{n \times var(X)}$$

where SER is the standard error of the regression (again):

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• The standard error of $\hat{\beta}_1$ is simply the square root of the variance

$$se(\hat{eta}_1) = \sqrt{var(\hat{eta}_1)}$$



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 - 3. **Variation** in *X*

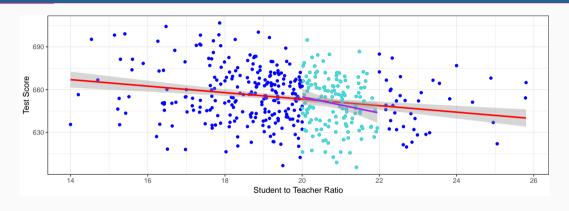


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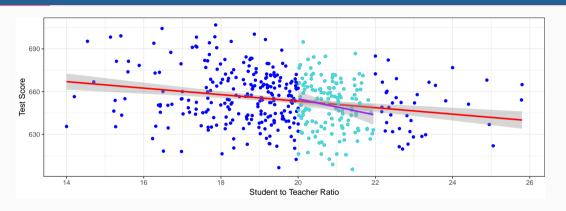
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- · Smaller var(X) (light dots only) \implies larger $var(\hat{\beta}_1)$: harder to determine precise slope!
- · Larger var(X) (all dots) \implies smaller $var(\hat{\beta}_1)$: easier to determine precise slope!



Hypothesis Testing About

REGRESSION

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- · All modern science is built upon statistical hypothesis testing, so understand it well!



Hypothesis Testing II

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 - i.e. if/when we've done our model right, the causal effect of X on Y



NULL AND ALTERNATIVE HYPOTHESES

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 population parameter
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- · Ask: "Does our data (sample) give us sufficient evidence to reject H_0 in favor of H_a ?"
 - Note: the test is always about H_0 ! See if we have sufficient evidence to reject the status quo



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 - · Note this means either $eta_{\mathrm{1}} < H_{\mathrm{0}}$ or $eta_{\mathrm{1}} > H_{\mathrm{0}}$



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 - Beyond the critical value is the "rejection region", sufficient evidence to reject ${\it H}_{\rm 0}$
- 4. A **conclusion** whether or not to reject H_0 in favor of H_a



Type I and Type II Errors

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- · We cannot distinguish between these two possibilities with any certainty



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 - · Believing we found nothing when there was truly a relationship to find



Type I and Type II Errors III

	H_0 is True	H_0 is False
Reject H ₀	Type I Error	Correct Outcome
	False Positive	True Positive
Don't Reject H ₀	Correct Outcome True Negative	Type II Error False Negative



Type I and Type II Errors IV

	Defendant is Innocent	Defendant is Guilty
Convict	Type Error	Correct Outcome
"I think he's guilty"	False Positive	True Positive
Don't Convict	Correct Outcome	Type II Error
"I think he's innocent"	True Negative	False Negative

 $\boldsymbol{\cdot}$ Depending on context, committing one type of error may be more serious than the other



TYPE I AND TYPE II ERRORS IV

	Defendant is	Defendant is
	Innocent	Guilty
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Don't Convict "I think he's innocent"	Correct Outcome True Negative	Type II Error False Negative

- · Depending on context, committing one type of error may be more serious than the other
- Common law *presumes* the defendant is innocent and a jury judges whether the evidence presented against the defendant would be plausible *if the defendant were in fact innocent*



TYPE I AND TYPE II ERRORS V

Example

For each of the following scenarios, identify the Type I error, Type II error, α and β , and which error is of greater consequence?



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- · H_0 : a rock climber's equipment is safe
- H_0 : a woman is not pregnant
- H_0 : a highway project will cost no more than \$10 million
- H_0 : an experimental cancer drug has a cure rate of at least 75%



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$$\alpha = P(\text{Reject } H_0 | H_0 \text{ is true})$$

$$\beta = P(Don't rejectH_0|H_0 is false)$$



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- We often specify in advance an α -level (0.10,0.05,0.01) with associated confidence level (90%, 95%, 99%)
- The probability of a **Type II error** is defined as β :





lpha and eta

	H_0 is True	H_0 is False
Reject H ₀	Type I Error	Correct Outcome
	α	$(1-\beta)$
Don't Reject H ₀	Correct Outcome	Type II Error
	$(1-\alpha)$	β



POWER AND *p*-VALUES

• The statistical power of the test is $1 - \beta$, the probability of correctly rejecting H_0 when H_0 is in fact false (e.g. not convicting an innocent person)

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- The *p*-value or significance probability is the probability that, given the null hypothesis is true, the test statistic from a random sample will be at least as extreme as the test statistic of our sample
- · If $p < \alpha$, the results are "statistically significant"



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- Compare p-value with pre-determined lpha (commonly, lpha= 0.05, 95% confidence level)
 - · If $p < \alpha$: statistically significant evidence sufficient to reject H_0 in favor of H_a



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 - If $p \ge \alpha$: insufficient evidence to reject H_0
 - Note this does **not** mean H_0 is true! We merely have failed to reject H_0



DIGRESSION: p-VALUES AND THE

PHILOSOPHY OF SCIENCE

"The null hypothesis is never proved or established, but is possibly disproved, in the course of experimentation. Every experiment may be said to exist only in order to give the facts a chance of disproving the null hypothesis."

(1931). The Design of Experiments



Sir Ronald A. Fisher (1890-1962)



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 - e.g. "All swans are white" is a hypothesis rejected upon discovery of a single black swan
- Note: economics is a very different kind of "science" with a different methodology!





Caution



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It is easy to misinterpret what statistical significance and p-values mean. THE FOLLOWING ARE FALSE:

• *p* is the probability that the alternative hypothesis is true (We can never *prove* an alternative hypothesis, only tentatively reject a null hypothesis)



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- p is the probability that the null hypothesis is false (We are not proving the null hypothesis false, only saying that it is very unlikely that under the null hypothesis, we obtain an event as rare as our sample)
- p is the probability that the observed effects were produced purely by random chance (p is computed under a specific model (assuming H_0 is true)
- p tells us how significant our finding is (p tells us nothing about the size or the real world significance of any effect deemed "statistically significant")

- Again, *p* is the probability that, assuming the null hypothesis is true, we obtain (by pure random chance) a test statistic at least as extreme as the one we estimated for our sample
 - · This will make more sense in context, when we discuss the nature of our test statistics



- Again, p is the probability that, assuming the null hypothesis is true, we obtain (by pure random chance) a test statistic at least as extreme as the one we estimated for our sample
 - · This will make more sense in context, when we discuss the nature of our test statistics
- Remember a low *p*-value means **either** that the null hypothesis is true and a highly improbable event has occurred or that the null hypothesis is false (we don't know which!)



STATISTICAL SIGNIFICANCE AND *p*-Values



SMBC 1623



STATISTICAL SIGNIFICANCE AND *p*-Values





SMBC 1623



STATISTICAL SIGNIFICANCE AND p-VALUES II







XKCD 882



Statistical Significance and p-Values III

WE FOUND NO LINK BETWEEN PURPLE JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN BROWN JELLY BEANS AND ACNE (P > 0.05)



WE FOUND NO LINK BETWEEN PINK JELLY BEANS AND ACNE (P>0.05).



WE FOUND NO LINK BETWEEN BLUE JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN TEAL JELLY BEANS AND ACNE (P>0.05).



WE FOUND NO LINK BETWEEN SALMON JELLY BEANS AND ACIE (P > 0.05).



WE FOUND NO LINK BETWEEN RED JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN TURQUOISE JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN MAGENTA JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN YELLOW JELLY BEANS AND ACNE (P > 0.05).



XKCD 882



Statistical Significance and p-Values IV

WE FOUND NO
LINK BETWEEN
GREY JELLY
BEANS AND POIL
(P > 0.05).

WE. FOUND NO LINK BETWEEN TAN JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN CYAN JELLY BEANS AND ACNE (P>0.05).



WE FOUND A LINK BETWEEN GREEN JELLY BEANS AND ACNE (P<0.05).



WE. FOUND NO LINK BETWEEN MAUVE JELLY BEANS AND ACNE (P>0.05).



WE. FOUND NO LINK BETWEEN BEIGE JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN LICAC JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN BLACK JELLY BEANS AND ACNE (P>0.05).



WE FOUND NO LINK BETWEEN PEACH JELLY BEANS AND ACNE (P > 0.05)



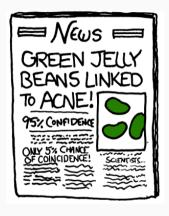
WE FOUND NO LINK BETWEEN ORANGE JELLY BEANS AND ACNE (P > 0.05).



XKCD 882



STATISTICAL SIGNIFICANCE AND p-Values V

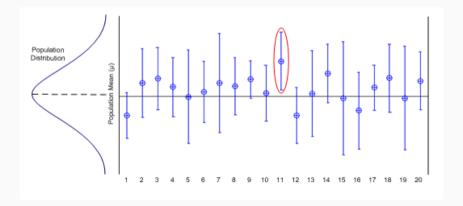


XKCD 882



STATISTICAL SIGNIFICANCE AND p-Values VI

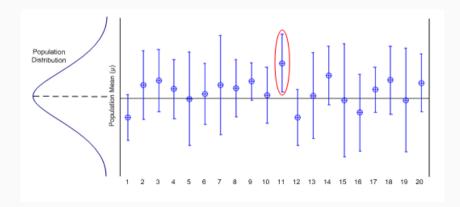
- Consider what "95% significance" or $\alpha =$ 0.05 means:





Statistical Significance and p-Values VI

- \cdot Consider what "95% significance" or $\alpha =$ 0.05 means:
 - If we repeat a procedure 20 times, we should expect 1/20 (5%) to produce a fluke result!





STATISTICAL SIGNIFICANCE AND p-Values VII

"The widespread use of "statistical significance" (generally interpreted as ($p \le 0.05$) as a license for making a claim of a scientific finding (or implied truth) leads to considerable distortion of the scientific process."





STATISTICAL SIGNIFICANCE AND p-Values VIII

How, and why, a journalist tricked news outlets into thinking chocolate makes you thin



Washington Post: How, and why, a journalist tricked news outlets into thinking chocolate makes you thin



BACK TO OUR HYPOTHESIS TEST: THE
TEST-STATISTIC

Distribution of H_0

• We next consider the population distribution assuming H_0 is true and calculate a test statistic, which takes the following form:

$$test\ statistic = \frac{sample\ statistic - hypothesized\ value}{standard\ error\ of\ the\ statistic}$$



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- \cdot We then compare our test statistic against a **critical value** to determine if we can reject H_0
- Essentially: test to see how likely a sample statistic at least as extreme as our discovery is if H_0 were true



Distribution of H_0 II

- We are testing our estimated $\hat{\beta}_{\rm 1}$ against a null hypothesis, e.g. $\beta_{\rm 1,0}=0$



DISTRIBUTION OF H_0 II

- · We are testing our estimated \hat{eta}_{1} against a null hypothesis, e.g. $eta_{1,0}=0$
- It would be nice if we could use normal distribution, our test statistic would just be *Z*-score:

$$Z = \frac{\hat{\beta}_1 - \beta_{1,0}}{SE(\hat{\beta}_1)}$$

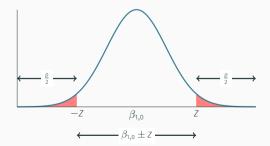


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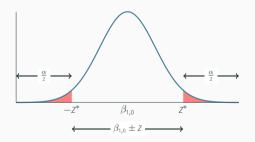
• p-value: area in the tail(s) of the distr. of $\hat{\beta}_1$ under H_0 beyond our Z score





DISTRIBUTION OF H_0 III

• The **critical value** Z^* is determined by our α level (e.g. 0.05)



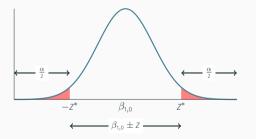
Critical values of Z^* with rejection regions in red



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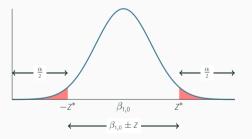
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- Any Z-score beyond ± 1.96 is in rejection region, sufficient evidence to reject H_0



Critical values of Z^* with rejection regions in red

HOOD

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- It would be nice if we *could* just use the normal distribution, and run a *Z*-test, as described above
- · Central Limit Theorem lets us if $n \geq 30$ and we know the population distribution μ, σ
- · We almost never know them...



STUDENT'S t-DISTRIBUTION

· Worked at Guinness testing beer quality





William Sealy Gosset (1876-1937)



STUDENT'S t-DISTRIBUTION

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- Using normal distributions with small sample sizes did not yield accurate estimates





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STUDENT'S t-DISTRIBUTION

- · Worked at Guinness testing beer quality
- Using normal distributions with small sample sizes did not yield accurate estimates
- Developed a new distribution, using the pseudonym "Student," to publish, the Student's t-distribution





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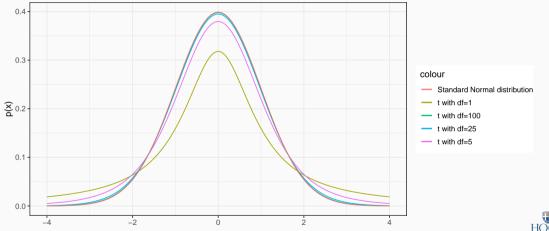


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- Exact shape of t depends on df: as \uparrow df, $t \rightarrow$ Normal distribution







CALCULATING t-SCORES: OLD-FASHIONED WAY

Two tail probability	0.20	0.10	0.05	
One tail probability	0.10	0.05	0.025	
Table T	df			
Values of t_{α}	1	3.078	6.314	12.706
ναιαεστια		1.886	2.920	4.303
	2	1.638	2.353	3.182
	4	1.533	2.333	2.776
		1.533	2.132	2.776
$\frac{\alpha}{2}$ $\frac{\alpha}{2}$	5	1.476	2.015	2.571
2	6	1.440	1.943	2.447
	7	1.415	1.895	2.365
$-t_{\alpha/2} = 0$ $t_{\alpha/2}$	8	1.397	1.860	2.306
Two tails	9	1.383	1.833	2.262
	10	1.372	1.812	2.228
	11	1.363	1.796	2.201
/ \	12	1.356	1.782	2.179
α	13	1.350	1.771	2.160
	14	1.345	1.761	2.145
$0 \qquad t_{\alpha}$	15	1.341	1.753	2.131
One tail	16	1.337	1.746	2.120
	17	1.333	1.740	2.110
	18	1.330	1.734	2.101
	19	1.328	1.729	2.093
	:	:	:	:
	:	:	:	:
	∞	1.282	1.645	1.960
Confidenc	80%	90%	95%	



CALCULATING t-SCORES: IN R

```
# use pt() command, needs t value and df
pt(2,df=5) #probability of t>2 with 5 df

## [1] 0.9490303
pt(2,df=40)# probability of t>2 with 40 df
```

[1] 0.9738388

pt(2, df=100) # probability of t>2 with 100 df

[1] 0.9758939



pnorm(2, mean=0, sd=1) # compare to normal distribution!

• So our test statistic is a *t*-score (instead of *Z*-score)

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 - Note: there will be a unique critical value for every value of n k 1!
 - \cdot R determines the critical t^* automatically with regression
- p-value= P(t < T)
- Reject ${\it H}_{\rm 0}$ if ${\it p}$ -value < lpha



Hypothesis Testing with *t*-distribution II

Depending on the desired alternative hypothesis:

Alternative	<i>p</i> -value	PDF
$H_a:eta_1>eta_{1,0}$	$P(T \ge t)$	μο τ
$H_a:eta_1$	$P(T \le t)$	t µo
$H_a:eta_1 eqeta_{1,0}$	$2P(T \ge t)$	μο τ



Example

We have an estimated regression line:

· Regression reporting format: Coefficients with their (standard errors) beneath them



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$$t = \frac{\hat{\beta}_1 - \beta_{1,0}}{SE(\hat{\beta}_1)} = \frac{-2.28 - 0}{0.48}$$



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$$t = \frac{\hat{\beta}_1 - \beta_{1,0}}{SE(\hat{\beta}_1)} = \frac{-2.28 - 0}{0.48} = -4.75$$

calculate p-value for t=-4.75, df=418



2*pt(-4.75,df=418) # x2 because we want both tails!

LOOKING AT R AGAIN

summary(school.regression)

```
##
## Call:
## lm(formula = testscr ~ str, data = CASchool)
##
## Residuals:
##
      Min
          10 Median 30 Max
## -47.727 -14.251 0.483 12.822 48.540
##
## Coefficients:
             Estimate Std. Error t value Pr(>|t|)
##
## (Intercept) 698.9330 9.4675 73.825 < 2e-16 ***
       -2.2798 0.4798 -4.751 2.78e-06 ***
## str
## ---
## Signif. codes: 0 '***' 0.001 '**' 0.05 '.' 0.1 ' ' 1
##
## Residual standard error: 18.58 on 418 degrees of freedom
## Multiple R-squared: 0.05124, Adjusted R-squared: 0.04897
## F-statistic: 22.58 on 1 and 418 DF. p-value: 2.783e-06
```

A Rule of Thumb

 \cdot If $|\hat{eta}_k| > 2 imes \mathit{SE}(\hat{eta}_k)$, the estimate is significant



A RULE OF THUMB

· If $|\hat{eta}_k| > 2 imes \mathit{SE}(\hat{eta}_k)$, the estimate is significant

· Since essentially $t=\frac{\hat{eta}_k}{\mathit{SE}(\hat{eta}_k)}$ and we roughly want $t\geq 2$ for 95% confidence level (lpha=0.05)



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- Random sampling variation implies that our estimated \hat{eta}_1 will vary slightly from sample to sample
- We can also calculate a confidence interval of $\hat{\beta}_1$ values that contain the true slope $\hat{\beta}_0$ with a specified probability (1 $-\alpha$)
- · In general, a confidence interval takes the form:

(point estimate — margin of error, point estimate + margin of error)



CONFIDENCE INTERVALS: INTUITION

• Recall the empirical 68-95-99.7 rule: approximately 95% of a normal distribution occurs within 2 standard deviations of the mean



CONFIDENCE INTERVALS: INTUITION

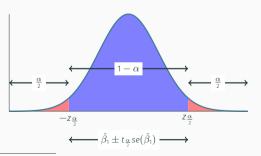
- Recall the empirical 68-95-99.7 rule: approximately 95% of a normal distribution occurs within 2 standard deviations of the mean
- Thus, in 95% of samples, the true slope (β_1) is likely to fall within about two standard errors of our estimated slope $(\hat{\beta}_1)$

$$(\hat{\beta}_1 - 2se(\hat{\beta}_1), \quad \hat{\beta}_1 - 2se(\hat{\beta}_1))$$

_ hota + 2σ . ____



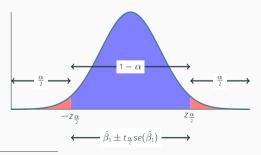
• We need to find the t-score³, $t_{\frac{\alpha}{2}}$ that puts an area equal to CL in the middle of the t-distribution $t \sim t_{n-k-1}$





³If we knew the population distribution, and if $n \ge 30$, we could just use Z-score and normal distribution

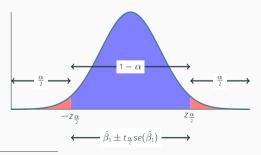
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- The specific $t_{\frac{\alpha}{2}}$ is (again) the critical value for a given confidence level
- $CL = 1 \alpha$, so $\frac{\alpha}{2}$ is the area split equally across the two tails of the distribution $\left(\frac{\alpha}{2}\right)$ in each tail



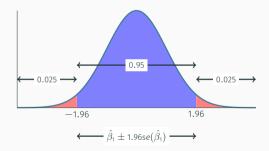


³If we knew the population distribution, and if $n \ge 30$, we could just use Z-score and normal distribution

CONFIDENCE INTERVALS: EXAMPLE

Example

For $\alpha = 0.05$ (or CL=95%) and large n, $t^* = Z^* = 1.96$





• The margin of error (MOE) is the critical value of t times the standard error of $\hat{\beta}_1$



 \cdot The margin of error (MOE) is the critical value of t times the standard error of \hat{eta}_1

$$MOE = t_{\frac{\alpha}{2}}se(\hat{\beta}_1)$$

• For example, for 95% confidence and large n:



 \cdot The margin of error (MOE) is the critical value of t times the standard error of \hat{eta}_1

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• For example, for 95% confidence and large n:

$$MOE = 1.96se(\hat{eta}_1)$$

· A confidence interval for eta_1 is then:



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• For example, for 95% confidence and large *n*:

$$\textit{MOE} = 1.96 se(\hat{\beta}_1)$$

• A confidence interval for β_1 is then:

$$\left(\hat{\beta}_1 - t_{\frac{\alpha}{2}} se(\hat{\beta}_1), \quad \hat{\beta}_1 + t_{\frac{\alpha}{2}} se(\hat{\beta}_1)\right)$$



· "We estimate with [1 $-\alpha$] confidence that the true population mean is between [...] and [...]."

COMMON CONFIDENCE INTERVALS

• For large *n*:

Confidence Level	Critical Value	Confidence Interval
90%	1.64	$\hat{eta}_1 \pm$ 1.64s $e(\hat{eta}_1)$
90%	1.96	$\hat{eta}_{ exttt{1}} \pm$ 1.96se $(\hat{eta}_{ exttt{1}})$
90%	2.58	$\hat{eta}_1 \pm 2.58$ se (\hat{eta}_1)



CONFIDENCE INTERVALS IN R

Generate confidence intervals with confint() command

```
confint(school.regression, level=0.90) # 90% confidence
##
                   5 % 95 %
## (Intercept) 683.325725 714.540180
## str -3.070804 -1.488812
confint(school.regression, level=0.95) # 95% confidence
      2.5 % 97.5 %
##
## (Intercept) 680.32313 717.542779
## str -3.22298 -1.336637
confint(school.regression, level=0.99) # 99% confidence
```



0.5 % 99.5 %

##

REPORTING REGRESSION OUTPUTS WITH

stargazer

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• Regressions are reported with a regression table (especially when we have many models!)



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- \cdot You can make this by hand, but it's much faster with a great package called ${\tt stargazer}$



REPORTING REGRESSION OUTPUTS WITH stargazer

- · Regressions are reported with a regression table (especially when we have many models!)
- · You can make this by hand, but it's much faster with a great package called stargazer

	Test Score
Class Size	-2.280***
	(0.480)
Constant	698.933***
	(9.467)
N	420
R^2	0.051
Residual Std. Error	18.581 (df = 418)
Notes:	***Significant at the 1 percent level.
	**Significant at the 5 percent level.



• Basic **stargazer** syntax is simple

```
library("stargazer") # load stargazer (you will need to install first!)
stargazer(reg.name,type="type")
```



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- Basic **stargazer** syntax is simple
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 - First include the name of the regression object to print (e.g. reg.name)

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library("stargazer") # load stargazer (you will need to install first!)
stargazer(reg.name,type="type")
```



- Basic **stargazer** syntax is simple
 - First you need to load with library(), and install if the first time!
 - First include the name of the regression object to print (e.g. reg.name)
 - Then specify the type of output ("html","latex" (for pdf), or "text")

```
library("stargazer") # load stargazer (you will need to install first!)
stargazer(reg.name,type="type")
```



Customizing stargazer

• stargazer allows for a lot of customization, e.g. the code for the above table:



stargazer Output

\hlino

• The raw output looks confusing, code that renders directly in html or pdf (via latex)

```
##
## \begingroup
## \scriptsize
## \begin{tabular}{@{\extracolsep{5pt}}lc}
## \\[-1.8ex]\hline
## \hline \\[-1.8ex]
## \\[-1.8ex] & Test Score \\
## \hline \\[-1.8ex]
   Class Size & $-$2.280$^{***}$ \\
##
    8 (0.480) \\
##
    8 \\
   Constant & 698.933$^{***}$ \\
##
    8 (9.467) \\
##
    8 \\
## \textit{N} & 420 \\
## R$^{2}$ & 0.051 \\
## Residual Std. Error & 18.581 (df = 418) \\
```

stargazer: Rendering Directly as Text

```
stargazer(school.regression, type="text")
```

```
##
                        Dependent variable:
##
##
##
                             testscr
                            -2.280***
## str
                             (0.480)
##
##
## Constant
                            698.933***
##
                              (9.467)
##
## Observations
                             420
## R2
                              0.051
## Adjusted R2
              0.049
## Residual Std. Error 18.581 (df = 418)
```