

Assignment Number: 2  
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From the discussion in class, we know about the following inequality, which I have used to prove some identities in this question.

If  $a, b \geq 0$  and  $a + b \geq c$ , then for all  $\eta \in [0, 1]$ ,

$$\eta \cdot a + (1 - \eta) \cdot b \geq c \cdot \min(\eta, 1 - \eta) \quad (1)$$

## PART 1

$$\begin{aligned} L^{0-1}(\eta) &= \min_{\hat{y}} \mathbb{E}_{Y \sim \eta} [l^{0-1}(\hat{y}, Y)] \\ &= \min_{\hat{y}} \mathbb{E}_{Y \sim \eta} [\mathbb{I}[\hat{y} \neq Y]] \\ &= \min_{\hat{y}} \mathbb{E}_{Y \sim \eta} [\mathbb{I}[\hat{y} = 1] \mathbb{I}[Y = -1] + \mathbb{I}[\hat{y} = -1] \mathbb{I}[Y = 1]] \\ &= \min_{\hat{y}} (1 - \eta) \mathbb{I}[\hat{y} = 1] + \eta \mathbb{I}[\hat{y} = -1] \\ &\geq \min(\eta, 1 - \eta) \end{aligned}$$

The last inequality is from the discussion in class. However, for an algorithm that predicts  $\hat{y}$  as follows,

$$\hat{y} = \begin{cases} +1 & \text{if } \eta > 0.5 \\ -1 & \text{else} \end{cases}$$

the equality is satisfied, that is  $L^{0-1}(\eta) = \min(\eta, 1 - \eta)$ .

Similarly for  $L^\sigma(\eta)$ ,

$$\begin{aligned} L^\sigma(\eta) &= \min_{\hat{y}} \mathbb{E}_{Y \sim \eta} [l^\sigma(\hat{y}, Y)] \\ &= \min_{\hat{y}} \mathbb{E}_{Y \sim \eta} \left[ \frac{1}{1 + \exp(Y\hat{y})} \right] \\ &= \min_{\hat{y}} (1 - \eta) \frac{\exp(\hat{y})}{1 + \exp(\hat{y})} + \eta \frac{1}{1 + \exp(\hat{y})} \end{aligned}$$

Again, from the discussion in class, we can say

$$L^\sigma(\eta) \geq \min(\eta, 1 - \eta)$$

We can find a prediction function such that this is an equality. The said prediction would be as follows

$$\hat{y} = \begin{cases} +\infty & \text{if } \eta > 0.5 \\ -\infty & \text{else} \end{cases}$$

Since for this predictor, the said inequality is indeed an equality, we can say that the minimum predictor gives  $L^\sigma(\eta) = \min(\eta, 1 - \eta)$ .

Hence, we have the result

$$L^{0-1}(\eta) = L^\sigma(\eta) = \min(\eta, 1 - \eta) \quad (2)$$

## PART 2

From the previous part, we have

$$\begin{aligned} L^{0-1}(\hat{y}, \eta) &= (1 - \eta) \mathbb{I}[\hat{y} = 1] + \eta \mathbb{I}[\hat{y} = -1] \\ L^{0-1}(\eta) &= \min(\eta, 1 - \eta) \end{aligned}$$

Therefore, there are four cases for  $L^{0-1}(\hat{y}, \eta) - L^{0-1}(\eta)$ , namely

$$L^{0-1}(\hat{y}, \eta) - L^{0-1}(\eta) = \begin{cases} 1 - 2\eta & \text{if } \eta \leq 0.5, \hat{y} = +1 \\ 0 & \text{if } \eta \leq 0.5, \hat{y} = -1 \\ 0 & \text{if } \eta > 0.5, \hat{y} = +1 \\ 2\eta - 1 & \text{if } \eta > 0.5, \hat{y} = -1 \end{cases}$$

From this, we can directly represent this in the form given in the question, *i.e.*

$$L^{0-1}(\hat{y}, \eta) - L^{0-1}(\eta) = |2\eta - 1| \cdot \mathbb{I}[\hat{y}(2\eta - 1) < 0]$$

## PART 3

From Part 1, we know

$$\begin{aligned} L^{0-1}(\text{sign}(\hat{y}), \eta) &= (1 - \eta) \mathbb{I}[\text{sign}(\hat{y}) = +1] + \eta \mathbb{I}[\text{sign}(\hat{y}) = -1] \\ &= (1 - \eta) \mathbb{I}[\hat{y} \geq 0] + \eta \mathbb{I}[\hat{y} < 0] \end{aligned}$$

Now, consider the term  $2L^\sigma(\hat{y}, \eta) - L^{0-1}(\text{sign}(\hat{y}), \eta)$ .

$$2L^\sigma(\hat{y}, \eta) - L^{0-1}(\text{sign}(\hat{y}), \eta) = (1 - \eta) \left[ \frac{2 \exp(\hat{y})}{1 + \exp(\hat{y})} - \mathbb{I}[\hat{y} \geq 0] \right] + \eta \left[ \frac{2}{1 + \exp(\hat{y})} - \mathbb{I}[\hat{y} < 0] \right]$$

Since for  $x \geq 0$ ,  $\frac{2 \exp(x)}{1 + \exp(x)} \geq 1$ , we have  $\frac{2 \exp(\hat{y})}{1 + \exp(\hat{y})} - \mathbb{I}[\hat{y} \geq 0] \geq 0$ . Similary, if  $x < 0$ ,  $\frac{2}{1 + \exp(x)} > 1$ . Therefore, we also have  $\frac{2}{1 + \exp(\hat{y})} - \mathbb{I}[\hat{y} < 0] \geq 0$ .

Hence, from Equation 1, we can say

$$\begin{aligned} 2L^\sigma(\hat{y}, \eta) - L^{0-1}(\text{sign}(\hat{y}), \eta) &\geq \left[ \frac{2 \exp(\hat{y}) + 2}{1 + \exp(\hat{y})} - \mathbb{I}[\hat{y} \geq 0] - \mathbb{I}[\hat{y} < 0] \right] \cdot \min(\eta, 1 - \eta) \\ &= \min(\eta, 1 - \eta) \\ &= 2L^\sigma(\eta) - L^{0-1}(\eta) \end{aligned}$$

The last equality comes from Equation 2. Therefore, we have the desired result, that is

$$L^{0-1}(\text{sign}(\hat{y}), \eta) - L^{0-1}(\eta) \leq 2(L^\sigma(\hat{y}, \eta) - L^\sigma(\eta)) \quad (3)$$

## PART 4

From part 3, equation 3, we have the following result

$$L^{0-1}(\text{sign}(\hat{y}), \eta) - L^{0-1}(\eta) \leq 2(L^\sigma(\hat{y}, \eta) - L^\sigma(\eta))$$

Suppose for some  $f : \mathcal{X} \rightarrow \mathbb{R}$ , we have  $\hat{y} = f(\mathbf{x})$ . Therefore, we can write

$$L^{0-1}(\text{sign}(f(\mathbf{x})), \eta) - L^{0-1}(\eta) \leq 2(L^\sigma(f(\mathbf{x}), \eta) - L^\sigma(\eta))$$

Since this inequality exists for all  $\mathbf{x} \in \mathcal{X}$ , we can take expectation on both sides over  $\mathbf{X} \sim \mathcal{D}$  without disturbing the inequality. Therefore

$$\mathbb{E}_{\mathbf{X} \sim \mathcal{D}} [L^{0-1}(\text{sign}(f(\mathbf{X})), \eta) - L^{0-1}(\eta)] \leq \mathbb{E}_{\mathbf{X} \sim \mathcal{D}} [2(L^\sigma(f(\mathbf{X}), \eta) - L^\sigma(\eta))]$$

Since both LHS and RHS are regret terms, we can replace them as follows,

$$\implies \mathcal{R}_{\mathcal{D}}^{0-1}[\text{sign} \circ f] \leq 2\mathcal{R}_{\mathcal{D}}^\sigma[f] \quad (4)$$

Hence, we have shown the inequality required.

## PART 5

We have

$$l_\alpha^\sigma = \frac{1}{1 + \exp(\alpha \cdot \hat{y}y)} \quad l^\sigma = \frac{1}{1 + \exp(\hat{y}y)}$$

Therefore, we can write  $l_\alpha^\sigma(\hat{y}, y)$  as

$$l_\alpha^\sigma(\hat{y}, y) = l^\sigma(\alpha \cdot \hat{y}, y)$$

Therefore, we have

$$L_\alpha^\sigma(\hat{y}, \eta) = L^\sigma(\alpha \cdot \hat{y}, \eta)$$

Now, borrowing the results from part 1, we can write

$$L_\alpha^\sigma(\hat{y}, \eta) = L^\sigma(\alpha \cdot \hat{y}, \eta) \geq \min\{\eta, 1 - \eta\}$$

**Note.** We can write this as the RHS is independent of the value of  $\mathbf{y}$  and therefore always holds

We can use the same predictor as we used in part 1 to show that there is in fact an optimal predictor which gives the lowest pointwise error ( $\min\{\eta, 1 - \eta\}$ ). The optimal predictor being

$$\hat{y} = \begin{cases} +\infty & \text{if } \eta > 0.5 \\ -\infty & \text{else} \end{cases}$$

Hence, we have

$$L_\alpha^\sigma(\eta) = \min\{\eta, 1 - \eta\} = L^\sigma(\eta)$$

From equation 4, we have

$$\begin{aligned} \mathbb{E}_{\mathbf{X} \sim \mathcal{D}} [L^{0-1}(\text{sign}(f(\mathbf{X})), \eta) - L^{0-1}(\eta)] &\leq 2 \mathbb{E}_{\mathbf{X} \sim \mathcal{D}} [L^\sigma(f(\mathbf{X}), \eta) - L^\sigma(\eta)] \\ \implies \mathbb{E}_{\mathbf{X} \sim \mathcal{D}} [L^{0-1}(\text{sign}(f(\mathbf{X})), \eta) - L^{0-1}(\eta)] &\leq 2 \mathbb{E}_{\mathbf{X} \sim \mathcal{D}} [L^\sigma(f(\mathbf{X}), \eta) - L_\alpha^\sigma(\eta)] \end{aligned}$$

Suppose we have some  $g : \mathcal{X} \rightarrow \infty$  such that  $f(\mathbf{x}) = \alpha \cdot g(\mathbf{x})$ . Clearly, such a  $g$  will exist for all  $f$ . Therefore, we can replace  $f$  by  $\alpha \cdot g$  as follows,

$$\mathbb{E}_{\mathbf{X} \sim \mathcal{D}} [L^{0-1}(\text{sign}(\alpha \cdot g(\mathbf{X})), \eta) - L^{0-1}(\eta)] \leq 2 \mathbb{E}_{\mathbf{X} \sim \mathcal{D}} [L^\sigma(\alpha \cdot g(\mathbf{X}), \eta) - L_\alpha^\sigma(\eta)]$$

Also, since  $\alpha > 0$ , we have  $\text{sign} \alpha \cdot g(\mathbf{x}) = \text{sign} g(\mathbf{x})$ , and  $L^\sigma(\alpha \cdot g(\mathbf{x}), \eta) = L_\alpha^\sigma(g(\mathbf{x}), \eta)$  for all  $\mathbf{x} \in \mathcal{X}$ . Hence, we can write this as

$$\begin{aligned} \mathbb{E}_{\mathbf{X} \sim \mathcal{D}} [L^{0-1}(\text{sign}(g(\mathbf{X})), \eta) - L^{0-1}(\eta)] &\leq 2 \mathbb{E}_{\mathbf{X} \sim \mathcal{D}} [L_\alpha^\sigma(g(\mathbf{X}), \eta) - L_\alpha^\sigma(\eta)] \\ \implies \mathcal{R}_{\mathcal{D}}^{0-1}[\text{sign} \circ g] &\leq 2\mathcal{R}_{\mathcal{D}}^{\sigma, \alpha}[g] \end{aligned}$$

Therefore, we have the exact same regret transfer bound for  $l_\alpha^\sigma$  as for  $l^\sigma$  with respect to  $l^{0-1}$ .