#### **Exponential Family and Generalized Linear Models**

Piyush Rai

Probabilistic Machine Learning (CS772A)

Aug 22, 2017

$$p(\mathbf{x}|\theta) = \frac{1}{Z(\theta)} h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x})]$$

$$p(\mathbf{x}|\theta) = \frac{1}{Z(\theta)}h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x})] = h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x}) - A(\theta)]$$

• Defines a class of distributions. An Exponential Family distribution is of the form

$$p(\mathbf{x}|\theta) = \frac{1}{Z(\theta)}h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x})] = h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x}) - A(\theta)]$$

•  $\mathbf{x} \in \mathcal{X}^m$  is the random variable being modeled (where  $\mathcal{X}$  denotes some space, e.g.,  $\mathbb{R}$  or  $\{0,1\}$ )

$$p(\mathbf{x}|\theta) = \frac{1}{Z(\theta)}h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x})] = h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x}) - A(\theta)]$$

- $m{\circ}$   $m{x} \in \mathcal{X}^m$  is the random variable being modeled (where  $\mathcal{X}$  denotes some space, e.g.,  $\mathbb{R}$  or  $\{0,1\}$ )
- $\theta \in \mathbb{R}^d$ : Natural parameters or canonical parameters defining the distribution

$$p(\mathbf{x}|\theta) = \frac{1}{Z(\theta)}h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x})] = h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x}) - A(\theta)]$$

- $m{\cdot}$   $m{x} \in \mathcal{X}^m$  is the random variable being modeled (where  $\mathcal{X}$  denotes some space, e.g.,  $\mathbb{R}$  or  $\{0,1\}$ )
- $oldsymbol{ heta} heta \in \mathbb{R}^d$ : Natural parameters or canonical parameters defining the distribution
- $\phi(\mathbf{x}) \in \mathbb{R}^d$ : Sufficient statistics (another random variable)

$$p(\mathbf{x}|\theta) = \frac{1}{Z(\theta)}h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x})] = h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x}) - A(\theta)]$$

- $m{v} \in \mathcal{X}^m$  is the random variable being modeled (where  $\mathcal{X}$  denotes some space, e.g.,  $\mathbb{R}$  or  $\{0,1\}$ )
- $oldsymbol{ heta} heta \in \mathbb{R}^d$ : Natural parameters or canonical parameters defining the distribution
- $\phi(\mathbf{x}) \in \mathbb{R}^d$ : Sufficient statistics (another random variable)
  - Why "sufficient":  $p(x|\theta)$  as a function of  $\theta$  depends on x only via  $\phi(x)$

$$p(\mathbf{x}|\theta) = \frac{1}{Z(\theta)}h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x})] = h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x}) - A(\theta)]$$

- $m{v} \in \mathcal{X}^m$  is the random variable being modeled (where  $\mathcal{X}$  denotes some space, e.g.,  $\mathbb{R}$  or  $\{0,1\}$ )
- $oldsymbol{ heta} \in \mathbb{R}^d$ : Natural parameters or canonical parameters defining the distribution
- $\phi(\mathbf{x}) \in \mathbb{R}^d$ : Sufficient statistics (another random variable)
  - Why "sufficient":  $p(x|\theta)$  as a function of  $\theta$  depends on x only via  $\phi(x)$
- $Z(\theta) = \int h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x})] d\mathbf{x}$ : Partition function

$$p(\mathbf{x}|\theta) = \frac{1}{Z(\theta)}h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x})] = h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x}) - A(\theta)]$$

- $m{v} \in \mathcal{X}^m$  is the random variable being modeled (where  $\mathcal{X}$  denotes some space, e.g.,  $\mathbb{R}$  or  $\{0,1\}$ )
- $oldsymbol{\theta} \in \mathbb{R}^d$ : Natural parameters or canonical parameters defining the distribution
- $\phi(\mathbf{x}) \in \mathbb{R}^d$ : Sufficient statistics (another random variable)
  - Why "sufficient":  $p(x|\theta)$  as a function of  $\theta$  depends on x only via  $\phi(x)$
- $Z(\theta) = \int h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x})] d\mathbf{x}$ : Partition function
- $A(\theta) = \log Z(\theta)$ : Log-partition function (also called the <u>cumulant function</u>)

$$\rho(\mathbf{x}|\theta) = \frac{1}{Z(\theta)}h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x})] = h(\mathbf{x})\exp[\theta^{\top}\phi(\mathbf{x}) - A(\theta)]$$

- $m{\cdot}$   $m{x} \in \mathcal{X}^m$  is the random variable being modeled (where  $\mathcal{X}$  denotes some space, e.g.,  $\mathbb{R}$  or  $\{0,1\}$ )
- $oldsymbol{ heta} \in \mathbb{R}^d$ : Natural parameters or canonical parameters defining the distribution
- $\phi(\mathbf{x}) \in \mathbb{R}^d$ : Sufficient statistics (another random variable)
  - Why "sufficient":  $p(x|\theta)$  as a function of  $\theta$  depends on x only via  $\phi(x)$
- $Z(\theta) = \int h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x})] d\mathbf{x}$ : Partition function
- $A(\theta) = \log Z(\theta)$ : Log-partition function (also called the <u>cumulant function</u>)
- h(x): A constant (doesn't depend on  $\theta$ )



• Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

• Recall the standard definition of the Binomial distribution

Binomial
$$(x|N,p) = \binom{N}{x} p^x (1-p)^{N-x}$$

Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Recall the standard definition of the Binomial distribution

Binomial
$$(x|N,p) = \binom{N}{x} p^x (1-p)^{N-x}$$

where N: number of trials

Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Recall the standard definition of the Binomial distribution

Binomial
$$(x|N,p) = \binom{N}{x} p^x (1-p)^{N-x}$$

where N: number of trials, x (a scalar): number of successes

• Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Recall the standard definition of the Binomial distribution

Binomial
$$(x|N,p) = \binom{N}{x} p^x (1-p)^{N-x}$$

where N: number of trials, x (a scalar): number of successes, p: probability is success in each trial

Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Recall the standard definition of the Binomial distribution

Binomial
$$(x|N,p) = \binom{N}{x} p^x (1-p)^{N-x}$$

where N: number of trials, x (a scalar): number of successes, p: probability is success in each trial

$$\binom{N}{x} \exp \left[ x \log \left( \frac{p}{1-p} \right) + N \log(1-p) \right]$$

Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

• Recall the standard definition of the Binomial distribution

Binomial
$$(x|N,p) = \binom{N}{x} p^x (1-p)^{N-x}$$

where N: number of trials, x (a scalar): number of successes, p: probability is success in each trial

• 
$$h(x) = \binom{N}{x}$$
  $\exp\left[x\log\left(\frac{p}{1-p}\right) + N\log(1-p)\right]$ 

Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Recall the standard definition of the Binomial distribution

Binomial
$$(x|N, p) = \binom{N}{x} p^x (1-p)^{N-x}$$

where N: number of trials, x (a scalar): number of successes, p: probability is success in each trial

$$\binom{N}{x} \exp \left[ x \log \left( \frac{p}{1-p} \right) + N \log(1-p) \right]$$

$$\bullet \ h(x) = \binom{N}{x}$$

• 
$$h(x) = \binom{N}{x}$$
  
•  $\theta = \log \left(\frac{p}{1-p}\right)$ 

• Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

• Recall the standard definition of the Binomial distribution

Binomial
$$(x|N, p) = \binom{N}{x} p^x (1-p)^{N-x}$$

where N: number of trials, x (a scalar): number of successes, p: probability is success in each trial

$$\binom{N}{x} \exp\left[x \log\left(\frac{p}{1-p}\right) + N \log(1-p)\right]$$

• 
$$h(x) = \binom{N}{x}$$
  
•  $\theta = \log\left(\frac{p}{1-p}\right)$ , and  $p = \frac{1}{1+\exp(-\theta)}$ 

Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Recall the standard definition of the Binomial distribution

Binomial
$$(x|N, p) = \binom{N}{x} p^x (1-p)^{N-x}$$

where N: number of trials, x (a scalar): number of successes, p: probability is success in each trial

- ullet  $heta=\log\left(rac{p}{1-p}
  ight)$ , and  $p=rac{1}{1+\exp(- heta)}$
- $\phi(x) = x$



Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Recall the standard definition of the Binomial distribution

Binomial
$$(x|N,p) = \binom{N}{x} p^x (1-p)^{N-x}$$

where N: number of trials, x (a scalar): number of successes, p: probability is success in each trial

$$\binom{N}{x} \exp\left[x\log\left(\frac{p}{1-p}\right) + N\log(1-p)\right]$$

- $h(x) = \binom{N}{x}$
- ullet  $\theta = \log\left(rac{p}{1-p}
  ight)$ , and  $p = rac{1}{1+\exp(- heta)}$
- $\phi(x) = x$
- $A(\theta) = -N \log(1-p)$



Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

• Recall the standard definition of the Binomial distribution

Binomial
$$(x|N, p) = \binom{N}{x} p^x (1-p)^{N-x}$$

where N: number of trials, x (a scalar): number of successes, p: probability is success in each trial

$$\binom{N}{x} \exp \left[ x \log \left( \frac{p}{1-p} \right) + N \log(1-p) \right]$$

- $\bullet \ h(x) = \binom{N}{x}$
- $\theta = \log\left(\frac{p}{1-p}\right)$ , and  $p = \frac{1}{1+\exp(-\theta)}$
- $\phi(x) = x$
- $A(\theta) = -N \log(1-p) = N \log(1 + \exp(\theta))$



• Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

• Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Recall the standard definition of a univariate Gaussian

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi\sigma^2}} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right] = \frac{1}{\sqrt{2\pi}} \exp\left[\frac{\mu}{\sigma^2}x - \frac{1}{2\sigma^2}x^2 - \frac{\mu^2}{2\sigma^2} - \log\sigma\right]$$
$$= \frac{1}{\sqrt{2\pi}} \exp\left[\left[-\frac{\mu}{\sigma^2}\right]^{\top} \begin{bmatrix} x \\ x^2 \end{bmatrix} - \left(\frac{\mu^2}{2\sigma^2} - \log\sigma\right)\right]$$

• Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Recall the standard definition of a univariate Gaussian

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right] = \frac{1}{\sqrt{2\pi}} \exp\left[\frac{\mu}{\sigma^2}x - \frac{1}{2\sigma^2}x^2 - \frac{\mu^2}{2\sigma^2} - \log\sigma\right]$$
$$= \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{\mu}{\sigma^2}\right]^{\top} \begin{bmatrix} x \\ x^2 \end{bmatrix} - \left(\frac{\mu^2}{2\sigma^2} - \log\sigma\right)$$

$$h(x) = \frac{1}{\sqrt{2\pi}}$$

• Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

• Recall the standard definition of a univariate Gaussian

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right] = \frac{1}{\sqrt{2\pi}} \exp\left[\frac{\mu}{\sigma^2}x - \frac{1}{2\sigma^2}x^2 - \frac{\mu^2}{2\sigma^2} - \log\sigma\right]$$
$$= \frac{1}{\sqrt{2\pi}} \exp\left[\left[-\frac{\mu}{\sigma^2}\right]^{\top} \begin{bmatrix} x \\ x^2 \end{bmatrix} - \left(\frac{\mu^2}{2\sigma^2} - \log\sigma\right)\right]$$

$$h(x) = \frac{1}{\sqrt{2\pi}}$$

$$\bullet \ \theta = \begin{bmatrix} \frac{\mu}{\sigma_1^2} \\ -\frac{1}{2\sigma^2} \end{bmatrix} = \begin{bmatrix} \theta_1 \\ \theta_2 \end{bmatrix}$$

• Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

• Recall the standard definition of a univariate Gaussian

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right] = \frac{1}{\sqrt{2\pi}} \exp\left[\frac{\mu}{\sigma^2}x - \frac{1}{2\sigma^2}x^2 - \frac{\mu^2}{2\sigma^2} - \log\sigma\right]$$
$$= \frac{1}{\sqrt{2\pi}} \exp\left[\left[-\frac{\mu}{\sigma^2}\right]^{\top} \begin{bmatrix} x \\ x^2 \end{bmatrix} - \left(\frac{\mu^2}{2\sigma^2} - \log\sigma\right)\right]$$

$$h(x) = \frac{1}{\sqrt{2\pi}}$$

$$\bullet \ \theta = \begin{bmatrix} \frac{\mu}{\sigma_1^2} \\ -\frac{1}{2\sigma^2} \end{bmatrix} = \begin{bmatrix} \theta_1 \\ \theta_2 \end{bmatrix}, \ \text{and} \ \begin{bmatrix} \mu \\ \sigma^2 \end{bmatrix} = \begin{bmatrix} -\frac{\theta_1}{2\theta_2} \\ -\frac{1}{2\theta_2} \end{bmatrix}$$

• Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

• Recall the standard definition of a univariate Gaussian

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right] = \frac{1}{\sqrt{2\pi}} \exp\left[\frac{\mu}{\sigma^2}x - \frac{1}{2\sigma^2}x^2 - \frac{\mu^2}{2\sigma^2} - \log\sigma\right]$$
$$= \frac{1}{\sqrt{2\pi}} \exp\left[\left[-\frac{\mu}{\sigma^2}\right]^{\top} \begin{bmatrix} x \\ x^2 \end{bmatrix} - \left(\frac{\mu^2}{2\sigma^2} - \log\sigma\right)\right]$$

• 
$$\theta = \begin{bmatrix} \frac{\mu}{\sigma^2} \\ -\frac{1}{2e^2} \end{bmatrix} = \begin{bmatrix} \theta_1 \\ \theta_2 \end{bmatrix}$$
, and  $\begin{bmatrix} \mu \\ \sigma^2 \end{bmatrix} = \begin{bmatrix} -\frac{\theta_1}{2\theta_2} \\ -\frac{1}{2\theta} \end{bmatrix}$ 

•  $h(x) = \frac{1}{\sqrt{2}}$ 



• Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Recall the standard definition of a univariate Gaussian

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right] = \frac{1}{\sqrt{2\pi}} \exp\left[\frac{\mu}{\sigma^2}x - \frac{1}{2\sigma^2}x^2 - \frac{\mu^2}{2\sigma^2} - \log\sigma\right]$$
$$= \frac{1}{\sqrt{2\pi}} \exp\left[-\frac{\mu}{\sigma^2}\right]^{\top} \begin{bmatrix} x \\ x^2 \end{bmatrix} - \left(\frac{\mu^2}{2\sigma^2} - \log\sigma\right)$$

$$\bullet \ \theta = \begin{bmatrix} \frac{\mu}{\sigma_1^2} \\ -\frac{1}{2\sigma^2} \end{bmatrix} = \begin{bmatrix} \theta_1 \\ \theta_2 \end{bmatrix}, \ \text{and} \ \begin{bmatrix} \mu \\ \sigma^2 \end{bmatrix} = \begin{bmatrix} -\frac{\theta_1}{2\theta_2} \\ -\frac{1}{2\theta_2} \end{bmatrix}$$

•  $h(x) = \frac{1}{\sqrt{2}}$ 

• 
$$A(\theta) = \frac{\mu^2}{2\sigma^2} + \log \sigma$$



• Let's try to write the Binomial distribution in the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp[\theta^{\top} \phi(\mathbf{x}) - A(\theta)]$$

Recall the standard definition of a univariate Gaussian

$$\mathcal{N}(x|\mu,\sigma^2) = \frac{1}{\sqrt{2\pi}\sigma^2} \exp\left[-\frac{(x-\mu)^2}{2\sigma^2}\right] = \frac{1}{\sqrt{2\pi}} \exp\left[\frac{\mu}{\sigma^2}x - \frac{1}{2\sigma^2}x^2 - \frac{\mu^2}{2\sigma^2} - \log\sigma\right]$$

$$= \frac{1}{\sqrt{2\pi}} \exp\left[\left[-\frac{\mu}{\sigma^2}\right]^{\top} \begin{bmatrix} x \\ x^2 \end{bmatrix} - \left(\frac{\mu^2}{2\sigma^2} - \log\sigma\right)\right]$$

$$\bullet h(x) = \frac{1}{\sqrt{2\pi}}$$

$$\bullet \ \theta = \begin{bmatrix} \frac{\mu}{\sigma_1^2} \\ -\frac{1}{2\sigma^2} \end{bmatrix} = \begin{bmatrix} \theta_1 \\ \theta_2 \end{bmatrix}, \ \text{and} \ \begin{bmatrix} \mu \\ \sigma^2 \end{bmatrix} = \begin{bmatrix} -\frac{\theta_1}{2\theta_2} \\ -\frac{1}{2\theta_2} \end{bmatrix}$$

• 
$$A(\theta) = \frac{\mu^2}{2\sigma^2} + \log \sigma = \frac{-\theta_1^2}{4\theta_2} - \frac{1}{2}\log(-2\theta_2) - \frac{1}{2}\log(2\pi)$$



#### A General Trick

A general trick to represent any distribution (assuming it is exp-family dist.) in exp-family form

• Write the given distribution as  $\exp(\log p())$  and simplify, e.g., for the Binomial

$$\exp\left(\log \operatorname{Binomial}(x|N,p)\right) = \exp\left(\log \binom{N}{x} p^{x} (1-p)^{N-x}\right)$$

$$= \exp\left(\log \binom{N}{x} + x \log p + (N-x) \log(1-p)\right)$$

$$= \binom{N}{x} \exp\left(x \log \frac{p}{1-p} - N \log(1-p)\right)$$

Now compare the resulting expression with the exponential family form

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp(\theta^{\top} \phi(\mathbf{x}) - A(\theta))$$

.. to identify the natural parameters, sufficient statistics, log-partition function, etc.



#### **Other Examples**

- Many other distribution belong to the exponential family
  - Bernoulli
  - Beta
  - Gamma
  - Multinoulli/Multinomial
  - Dirichlet
  - Multivariate Gaussian
  - .. and many more ( https://en.wikipedia.org/wiki/Exponential\_family )

#### **Other Examples**

- Many other distribution belong to the exponential family
  - Bernoulli
  - Beta
  - Gamma
  - Multinoulli/Multinomial
  - Dirichlet
  - Multivariate Gaussian
  - .. and many more ( https://en.wikipedia.org/wiki/Exponential\_family )
- Note: Not all distributions belong to the exponential family, e.g.,
  - Uniform distribution  $(x \sim \text{Unif}(a, b))$
  - Student-t distribution
  - Mixture distributions (e.g., mixture of Gaussians)

#### Other Examples

- Many other distribution belong to the exponential family
  - Bernoulli
  - Beta
  - Gamma
  - Multinoulli/Multinomial
  - Dirichlet
  - Multivariate Gaussian
  - .. and many more ( https://en.wikipedia.org/wiki/Exponential\_family )
- Note: Not all distributions belong to the exponential family, e.g.,
  - Uniform distribution  $(x \sim \text{Unif}(a, b))$
  - Student-t distribution
  - Mixture distributions (e.g., mixture of Gaussians)
- If the support of the distribution depends on its parameters, then it is not an exp. family dist.

#### **Log-Partition Function**

•  $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function

#### **Log-Partition Function**

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function
- Derivatives of  $A(\theta)$  can be used to generate the cumulants of the sufficient statistics  $\phi(x)$

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function
- Derivatives of  $A(\theta)$  can be used to generate the cumulants of the sufficient statistics  $\phi(x)$
- Exercise: Assume  $\theta$  to be a scalar (thus  $\phi(\mathbf{x})$  is also scalar).

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function
- Derivatives of  $A(\theta)$  can be used to generate the cumulants of the sufficient statistics  $\phi(x)$
- Exercise: Assume  $\theta$  to be a scalar (thus  $\phi(\mathbf{x})$  is also scalar). Show that the first and the second derivatives of  $A(\theta)$  are

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function
- Derivatives of  $A(\theta)$  can be used to generate the cumulants of the sufficient statistics  $\phi(x)$
- Exercise: Assume  $\theta$  to be a scalar (thus  $\phi(\mathbf{x})$  is also scalar). Show that the first and the second derivatives of  $A(\theta)$  are

$$\frac{dA}{d\theta} = \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})]$$

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function
- Derivatives of  $A(\theta)$  can be used to generate the cumulants of the sufficient statistics  $\phi(x)$
- Exercise: Assume  $\theta$  to be a scalar (thus  $\phi(\mathbf{x})$  is also scalar). Show that the first and the second derivatives of  $A(\theta)$  are

$$\frac{dA}{d\theta} = \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})] = \text{mean}[\phi(\mathbf{x})]$$

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function
- Derivatives of  $A(\theta)$  can be used to generate the cumulants of the sufficient statistics  $\phi(x)$
- Exercise: Assume  $\theta$  to be a scalar (thus  $\phi(\mathbf{x})$  is also scalar). Show that the first and the second derivatives of  $A(\theta)$  are

$$\begin{array}{rcl} \frac{dA}{d\theta} & = & \mathbb{E}_{\rho(\mathbf{x}|\theta)}[\phi(\mathbf{x})] = \mathsf{mean}[\phi(\mathbf{x})] \\ \\ \frac{d^2A}{d\theta^2} & = & \mathbb{E}_{\rho(\mathbf{x}|\theta)}[\phi^2(\mathbf{x})] - \left[\mathbb{E}_{\rho(\mathbf{x}|\theta)}[\phi(\mathbf{x})]\right]^2 \end{array}$$

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function
- Derivatives of  $A(\theta)$  can be used to generate the cumulants of the sufficient statistics  $\phi(x)$
- Exercise: Assume  $\theta$  to be a scalar (thus  $\phi(\mathbf{x})$  is also scalar). Show that the first and the second derivatives of  $A(\theta)$  are

$$\begin{array}{rcl} \frac{dA}{d\theta} & = & \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})] = \mathsf{mean}[\phi(\mathbf{x})] \\ \\ \frac{d^2A}{d\theta^2} & = & \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi^2(\mathbf{x})] - \left[\mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})]\right]^2 = \mathsf{var}[\phi(\mathbf{x})] \end{array}$$

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function
- Derivatives of  $A(\theta)$  can be used to generate the cumulants of the sufficient statistics  $\phi(x)$
- Exercise: Assume  $\theta$  to be a scalar (thus  $\phi(\mathbf{x})$  is also scalar). Show that the first and the second derivatives of  $A(\theta)$  are

$$\begin{array}{rcl} \frac{dA}{d\theta} & = & \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})] = \mathsf{mean}[\phi(\mathbf{x})] \\ \\ \frac{d^2A}{d\theta^2} & = & \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi^2(\mathbf{x})] - \left[\mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})]\right]^2 = \mathsf{var}[\phi(\mathbf{x})] \end{array}$$

• Note: The above result also holds when  $\theta$  and  $\phi(x)$  are vector-valued (the "var" will be "covar")

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function
- Derivatives of  $A(\theta)$  can be used to generate the cumulants of the sufficient statistics  $\phi(x)$
- Exercise: Assume  $\theta$  to be a scalar (thus  $\phi(\mathbf{x})$  is also scalar). Show that the first and the second derivatives of  $A(\theta)$  are

$$\begin{array}{lcl} \frac{dA}{d\theta} & = & \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})] = \mathsf{mean}[\phi(\mathbf{x})] \\ \frac{d^2A}{d\theta^2} & = & \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi^2(\mathbf{x})] - \left[\mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})]\right]^2 = \mathsf{var}[\phi(\mathbf{x})] \end{array}$$

- Note: The above result also holds when  $\theta$  and  $\phi(x)$  are vector-valued (the "var" will be "covar")
- Important:  $A(\theta)$  is a **convex function** of  $\theta$ .

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function
- Derivatives of  $A(\theta)$  can be used to generate the cumulants of the sufficient statistics  $\phi(x)$
- Exercise: Assume  $\theta$  to be a scalar (thus  $\phi(\mathbf{x})$  is also scalar). Show that the first and the second derivatives of  $A(\theta)$  are

$$\begin{array}{lcl} \frac{dA}{d\theta} & = & \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})] = \mathsf{mean}[\phi(\mathbf{x})] \\ \\ \frac{d^2A}{d\theta^2} & = & \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi^2(\mathbf{x})] - \left[\mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})]\right]^2 = \mathsf{var}[\phi(\mathbf{x})] \end{array}$$

- Note: The above result also holds when  $\theta$  and  $\phi(x)$  are vector-valued (the "var" will be "covar")
- Important:  $A(\theta)$  is a **convex function** of  $\theta$ . Why?

- $A(\theta) = \log Z(\theta) = \log \int h(x) \exp[\theta^{\top} \phi(x)] dx$  is the log-partition function
- $A(\theta)$  is also called the cumulant function
- Derivatives of  $A(\theta)$  can be used to generate the cumulants of the sufficient statistics  $\phi(x)$
- Exercise: Assume  $\theta$  to be a scalar (thus  $\phi(\mathbf{x})$  is also scalar). Show that the first and the second derivatives of  $A(\theta)$  are

$$\begin{array}{lcl} \frac{dA}{d\theta} & = & \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})] = \mathsf{mean}[\phi(\mathbf{x})] \\ \\ \frac{d^2A}{d\theta^2} & = & \mathbb{E}_{p(\mathbf{x}|\theta)}[\phi^2(\mathbf{x})] - \left[\mathbb{E}_{p(\mathbf{x}|\theta)}[\phi(\mathbf{x})]\right]^2 = \mathsf{var}[\phi(\mathbf{x})] \end{array}$$

- Note: The above result also holds when  $\theta$  and  $\phi(x)$  are vector-valued (the "var" will be "covar")
- Important:  $A(\theta)$  is a **convex function** of  $\theta$ . Why?
- Exercise: For Binomial, using its expression of  $A(\theta)$ , derive the first and second cumulants of  $\phi(x)$

ullet Suppose we have data  $\mathcal{D} = \{ oldsymbol{x}_1, \dots, oldsymbol{x}_N \}$  drawn i.i.d. from an exponential family distribution

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp \left[\theta^{\top} \phi(\mathbf{x}) - A(\theta)\right]$$

ullet Suppose we have data  $\mathcal{D} = \{ oldsymbol{x}_1, \dots, oldsymbol{x}_N \}$  drawn i.i.d. from an exponential family distribution

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp \left[\theta^{\top} \phi(\mathbf{x}) - A(\theta)\right]$$

$$p(\mathcal{D}|\theta) = \prod_{i=1}^{N} p(x_i|\theta)$$

• Suppose we have data  $\mathcal{D} = \{x_1, \dots, x_N\}$  drawn i.i.d. from an exponential family distribution

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp \left[\theta^{\top} \phi(\mathbf{x}) - A(\theta)\right]$$

$$p(\mathcal{D}|\theta) = \prod_{i=1}^{N} p(\mathbf{x}_i|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \sum_{i=1}^{N} \phi(\mathbf{x}_i) - NA(\theta)\right]$$

ullet Suppose we have data  $\mathcal{D} = \{ oldsymbol{x}_1, \dots, oldsymbol{x}_N \}$  drawn i.i.d. from an exponential family distribution

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp \left[\theta^{\top} \phi(\mathbf{x}) - A(\theta)\right]$$

$$p(\mathcal{D}|\theta) = \prod_{i=1}^{N} p(\mathbf{x}_i|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \sum_{i=1}^{N} \phi(\mathbf{x}_i) - NA(\theta)\right] = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) - NA(\theta)\right]$$

• Suppose we have data  $\mathcal{D} = \{x_1, \dots, x_N\}$  drawn i.i.d. from an exponential family distribution

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp \left[\theta^{\top} \phi(\mathbf{x}) - A(\theta)\right]$$

To do MLE, we need the overall likelihood. This is simply a product of the individual likelihoods

$$p(\mathcal{D}|\theta) = \prod_{i=1}^{N} p(\mathbf{x}_i|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \sum_{i=1}^{N} \phi(\mathbf{x}_i) - NA(\theta)\right] = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) - NA(\theta)\right]$$

• To estimate  $\theta$  (as we'll see shortly), we only need  $\phi(\mathcal{D}) = \sum_{i=1}^{N} \phi(\mathbf{x}_i)$  and N

• Suppose we have data  $\mathcal{D} = \{x_1, \dots, x_N\}$  drawn i.i.d. from an exponential family distribution

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp \left[\theta^{\top} \phi(\mathbf{x}) - A(\theta)\right]$$

$$p(\mathcal{D}|\theta) = \prod_{i=1}^{N} p(\mathbf{x}_i|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \sum_{i=1}^{N} \phi(\mathbf{x}_i) - NA(\theta)\right] = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) - NA(\theta)\right]$$

- ullet To estimate heta (as we'll see shortly), we only need  $\phi(\mathcal{D}) = \sum_{i=1}^N \phi(\mathbf{x}_i)$  and N
- Size of  $\phi(\mathcal{D}) = \sum_{i=1}^{N} \phi(\mathbf{x}_i)$  does not grow with N (same as the size of each  $\phi(\mathbf{x}_i)$ )

• Suppose we have data  $\mathcal{D} = \{x_1, \dots, x_N\}$  drawn i.i.d. from an exponential family distribution

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp \left[\theta^{\top} \phi(\mathbf{x}) - A(\theta)\right]$$

$$p(\mathcal{D}|\theta) = \prod_{i=1}^{N} p(\mathbf{x}_i|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \sum_{i=1}^{N} \phi(\mathbf{x}_i) - NA(\theta)\right] = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) - NA(\theta)\right]$$

- To estimate  $\theta$  (as we'll see shortly), we only need  $\phi(\mathcal{D}) = \sum_{i=1}^{N} \phi(\mathbf{x}_i)$  and N
- Size of  $\phi(\mathcal{D}) = \sum_{i=1}^{N} \phi(\mathbf{x}_i)$  does not grow with N (same as the size of each  $\phi(\mathbf{x}_i)$ )
- Only exponential family distributions have finite-sized sufficient statistics

• Suppose we have data  $\mathcal{D} = \{x_1, \dots, x_N\}$  drawn i.i.d. from an exponential family distribution

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp \left[\theta^{\top} \phi(\mathbf{x}) - A(\theta)\right]$$

$$p(\mathcal{D}|\theta) = \prod_{i=1}^{N} p(\mathbf{x}_i|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \sum_{i=1}^{N} \phi(\mathbf{x}_i) - NA(\theta)\right] = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) - NA(\theta)\right]$$

- ullet To estimate heta (as we'll see shortly), we only need  $\phi(\mathcal{D}) = \sum_{i=1}^N \phi(m{x}_i)$  and N
- Size of  $\phi(\mathcal{D}) = \sum_{i=1}^{N} \phi(\mathbf{x}_i)$  does not grow with N (same as the size of each  $\phi(\mathbf{x}_i)$ )
- Only exponential family distributions have finite-sized sufficient statistics
  - No need to store all the data; can simply store and recursively update the sufficient statistics with more and more data

• Suppose we have data  $\mathcal{D} = \{x_1, \dots, x_N\}$  drawn i.i.d. from an exponential family distribution

$$p(\mathbf{x}|\theta) = h(\mathbf{x}) \exp \left[\theta^{\top} \phi(\mathbf{x}) - A(\theta)\right]$$

$$p(\mathcal{D}|\theta) = \prod_{i=1}^{N} p(\mathbf{x}_i|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \sum_{i=1}^{N} \phi(\mathbf{x}_i) - NA(\theta)\right] = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) - NA(\theta)\right]$$

- To estimate  $\theta$  (as we'll see shortly), we only need  $\phi(\mathcal{D}) = \sum_{i=1}^{N} \phi(\mathbf{x}_i)$  and N
- Size of  $\phi(\mathcal{D}) = \sum_{i=1}^{N} \phi(\mathbf{x}_i)$  does not grow with N (same as the size of each  $\phi(\mathbf{x}_i)$ )
- Only exponential family distributions have finite-sized sufficient statistics
  - No need to store all the data; can simply store and recursively update the sufficient statistics with more and more data
  - Very useful when doing probabilistic/Bayesian inference with large-scale data sets. Also useful in online parameter estimation problems.



- The likelihood is of the form  $p(\mathcal{D}|\theta) = \left[\prod_{i=1}^N h(\mathbf{x}_i)\right] \exp\left[\theta^\top \phi(\mathcal{D}) NA(\theta)\right]$
- The log-likelihood is (ignoring constant w.r.t.  $\theta$ ):  $\log p(\mathcal{D}|\theta) = \theta^{\top} \phi(\mathcal{D}) NA(\theta)$

- The likelihood is of the form  $p(\mathcal{D}|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) NA(\theta)\right]$
- The log-likelihood is (ignoring constant w.r.t.  $\theta$ ):  $\log p(\mathcal{D}|\theta) = \theta^{\top} \phi(\mathcal{D}) NA(\theta)$
- ullet Note: This is concave in heta (since -A( heta) is concave). Maximization will yield a global maxima of heta

- The likelihood is of the form  $p(\mathcal{D}|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) NA(\theta)\right]$
- The log-likelihood is (ignoring constant w.r.t.  $\theta$ ):  $\log p(\mathcal{D}|\theta) = \theta^{\top} \phi(\mathcal{D}) NA(\theta)$
- Note: This is concave in  $\theta$  (since  $-A(\theta)$  is concave). Maximization will yield a global maxima of  $\theta$
- MLE for exp-fam distributions can <u>also</u> be seen as doing <u>moment-matching</u>. To see this, note that

$$abla_{ heta} \left[ heta^ op \phi(\mathcal{D}) - \mathit{NA}( heta) 
ight]$$

- The likelihood is of the form  $p(\mathcal{D}|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) NA(\theta)\right]$
- The log-likelihood is (ignoring constant w.r.t.  $\theta$ ):  $\log p(\mathcal{D}|\theta) = \theta^{\top} \phi(\mathcal{D}) NA(\theta)$
- Note: This is concave in  $\theta$  (since  $-A(\theta)$  is concave). Maximization will yield a global maxima of  $\theta$
- MLE for exp-fam distributions can <u>also</u> be seen as doing <u>moment-matching</u>. To see this, note that

$$\nabla_{\theta} \left[ \theta^{\top} \phi(\mathcal{D}) - NA(\theta) \right] = \phi(\mathcal{D}) - N\nabla_{\theta} [A(\theta)]$$

- The likelihood is of the form  $p(\mathcal{D}|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) NA(\theta)\right]$
- The log-likelihood is (ignoring constant w.r.t.  $\theta$ ):  $\log p(\mathcal{D}|\theta) = \theta^{\top} \phi(\mathcal{D}) NA(\theta)$
- ullet Note: This is concave in heta (since -A( heta) is concave). Maximization will yield a global maxima of heta
- MLE for exp-fam distributions can <u>also</u> be seen as doing <u>moment-matching</u>. To see this, note that

$$\nabla_{\theta} \left[ \theta^{\top} \phi(\mathcal{D}) - \mathsf{NA}(\theta) \right] \quad = \quad \phi(\mathcal{D}) - \mathsf{N} \nabla_{\theta} [\mathsf{A}(\theta)] \quad = \quad \phi(\mathcal{D}) - \mathsf{N} \mathbb{E}_{\rho(\mathbf{x} \mid \theta)} [\phi(\mathbf{x})]$$

- The likelihood is of the form  $p(\mathcal{D}|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) NA(\theta)\right]$
- The log-likelihood is (ignoring constant w.r.t.  $\theta$ ):  $\log p(\mathcal{D}|\theta) = \theta^{\top} \phi(\mathcal{D}) NA(\theta)$
- Note: This is concave in  $\theta$  (since  $-A(\theta)$  is concave). Maximization will yield a global maxima of  $\theta$
- MLE for exp-fam distributions can <u>also</u> be seen as doing moment-matching. To see this, note that

$$\nabla_{\theta} \left[ \theta^{\top} \phi(\mathcal{D}) - NA(\theta) \right] = \phi(\mathcal{D}) - N\nabla_{\theta} [A(\theta)] = \phi(\mathcal{D}) - N\mathbb{E}_{\rho(\mathbf{x}|\theta)} [\phi(\mathbf{x})] = \sum_{i=1}^{N} \phi(\mathbf{x}_{i}) - N\mathbb{E}_{\rho(\mathbf{x}|\theta)} [\phi(\mathbf{x})]$$

ullet Therefore, at the "optimal" (i.e., MLE)  $\hat{ heta}$ , where the derivative is 0, the following must hold

$$\mathbb{E}_{\rho(\boldsymbol{x}|\theta)}[\phi(\boldsymbol{x})] = \frac{1}{N} \sum_{i=1}^{N} \phi(\boldsymbol{x}_i)$$



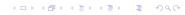
- The likelihood is of the form  $p(\mathcal{D}|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) NA(\theta)\right]$
- The log-likelihood is (ignoring constant w.r.t.  $\theta$ ):  $\log p(\mathcal{D}|\theta) = \theta^{\top} \phi(\mathcal{D}) NA(\theta)$
- Note: This is concave in  $\theta$  (since  $-A(\theta)$  is concave). Maximization will yield a global maxima of  $\theta$
- MLE for exp-fam distributions can <u>also</u> be seen as doing moment-matching. To see this, note that

$$\nabla_{\theta} \left[ \theta^{\top} \phi(\mathcal{D}) - NA(\theta) \right] = \phi(\mathcal{D}) - N\nabla_{\theta} [A(\theta)] = \phi(\mathcal{D}) - N\mathbb{E}_{\rho(\mathbf{x}|\theta)} [\phi(\mathbf{x})] = \sum_{i=1}^{N} \phi(\mathbf{x}_{i}) - N\mathbb{E}_{\rho(\mathbf{x}|\theta)} [\phi(\mathbf{x})]$$

ullet Therefore, at the "optimal" (i.e., MLE)  $\hat{ heta}$ , where the derivative is 0, the following must hold

$$\boxed{\mathbb{E}_{\rho(\boldsymbol{x}|\theta)}[\phi(\boldsymbol{x})] = \frac{1}{N} \sum_{i=1}^{N} \phi(\boldsymbol{x}_i)}$$

• This is basically <u>matching</u> the <u>expected</u> moments of the distribution with <u>empirical</u> moments ("empirical" here means what we compute using the observed data)



• Given N observations  $x_1, \ldots, x_N$  from a univariate Gaussian  $N(x|\mu, \sigma^2)$ , doing moment-matching

$$\mathbb{E}[\phi(x)] = \frac{1}{N} \sum_{i=1}^{N} \phi(x_i)$$

• Given N observations  $x_1, \ldots, x_N$  from a univariate Gaussian  $N(x|\mu, \sigma^2)$ , doing moment-matching

$$\mathbb{E}[\phi(x)] = \frac{1}{N} \sum_{i=1}^{N} \phi(x_i)$$

$$\mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix} = \begin{bmatrix} \frac{1}{N} \sum_{i=1}^{N} x_i \\ \frac{1}{N} \sum_{i=1}^{N} x_i^2 \end{bmatrix}$$

• Given N observations  $x_1, \ldots, x_N$  from a univariate Gaussian  $N(x|\mu, \sigma^2)$ , doing moment-matching

$$\mathbb{E}[\phi(x)] = \frac{1}{N} \sum_{i=1}^{N} \phi(x_i)$$

• The "true", i.e., expected moments:  $\mathbb{E}[\phi(x)] = \mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix}$ . Therefore

$$\mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix} = \begin{bmatrix} \frac{1}{N} \sum_{i=1}^{N} x_i \\ \frac{1}{N} \sum_{i=1}^{N} x_i^2 \end{bmatrix}$$

ullet For a univariate Gaussian, note that  $\mathbb{E}[x]=\mu$ 

• Given N observations  $x_1, \ldots, x_N$  from a univariate Gaussian  $N(x|\mu, \sigma^2)$ , doing moment-matching

$$\mathbb{E}[\phi(x)] = \frac{1}{N} \sum_{i=1}^{N} \phi(x_i)$$

ullet The "true", i.e., expected moments:  $\mathbb{E}[\phi(x)] = \mathbb{E} \left[ egin{matrix} x \\ x^2 \end{array} \right]$ . Therefore

$$\mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix} = \begin{bmatrix} \frac{1}{N} \sum_{i=1}^{N} x_i \\ \frac{1}{N} \sum_{i=1}^{N} x_i^2 \end{bmatrix}$$

ullet For a univariate Gaussian, note that  $\mathbb{E}[x]=\mu$  and  $\mathbb{E}[x^2]=\mathsf{var}[x]+\mathbb{E}[x]^2$ 

• Given N observations  $x_1, \ldots, x_N$  from a univariate Gaussian  $N(x|\mu, \sigma^2)$ , doing moment-matching

$$\mathbb{E}[\phi(x)] = \frac{1}{N} \sum_{i=1}^{N} \phi(x_i)$$

• The "true", i.e., expected moments:  $\mathbb{E}[\phi(x)] = \mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix}$ . Therefore

$$\mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix} = \begin{bmatrix} \frac{1}{N} \sum_{i=1}^{N} x_i \\ \frac{1}{N} \sum_{i=1}^{N} x_i^2 \end{bmatrix}$$

• For a univariate Gaussian, note that  $\mathbb{E}[x] = \mu$  and  $\mathbb{E}[x^2] = \text{var}[x] + \mathbb{E}[x]^2 = \sigma^2 + \mu^2$ 

• Given N observations  $x_1, \ldots, x_N$  from a univariate Gaussian  $N(x|\mu, \sigma^2)$ , doing moment-matching

$$\mathbb{E}[\phi(x)] = \frac{1}{N} \sum_{i=1}^{N} \phi(x_i)$$

$$\mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix} = \begin{bmatrix} \frac{1}{N} \sum_{i=1}^{N} x_i \\ \frac{1}{N} \sum_{i=1}^{N} x_i^2 \end{bmatrix}$$

- For a univariate Gaussian, note that  $\mathbb{E}[x] = \mu$  and  $\mathbb{E}[x^2] = \text{var}[x] + \mathbb{E}[x]^2 = \sigma^2 + \mu^2$
- Thus we have two equations and two unknowns

• Given N observations  $x_1, \ldots, x_N$  from a univariate Gaussian  $N(x|\mu, \sigma^2)$ , doing moment-matching

$$\mathbb{E}[\phi(x)] = \frac{1}{N} \sum_{i=1}^{N} \phi(x_i)$$

$$\mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix} = \begin{bmatrix} \frac{1}{N} \sum_{i=1}^{N} x_i \\ \frac{1}{N} \sum_{i=1}^{N} x_i^2 \end{bmatrix}$$

- For a univariate Gaussian, note that  $\mathbb{E}[x] = \mu$  and  $\mathbb{E}[x^2] = \text{var}[x] + \mathbb{E}[x]^2 = \sigma^2 + \mu^2$
- Thus we have two equations and two unknowns
- From the first equation, we immediately get  $\mu = \frac{1}{N} \sum_{i=1}^{N} x_i$



• Given N observations  $x_1, \ldots, x_N$  from a univariate Gaussian  $N(x|\mu, \sigma^2)$ , doing moment-matching

$$\mathbb{E}[\phi(x)] = \frac{1}{N} \sum_{i=1}^{N} \phi(x_i)$$

$$\mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix} = \begin{bmatrix} \frac{1}{N} \sum_{i=1}^{N} x_i \\ \frac{1}{N} \sum_{i=1}^{N} x_i^2 \end{bmatrix}$$

- For a univariate Gaussian, note that  $\mathbb{E}[x] = \mu$  and  $\mathbb{E}[x^2] = \text{var}[x] + \mathbb{E}[x]^2 = \sigma^2 + \mu^2$
- Thus we have two equations and two unknowns
- From the first equation, we immediately get  $\mu = \frac{1}{N} \sum_{i=1}^{N} x_i$
- From the second equation, we get  $\sigma^2 = \mathbb{E}[x^2] \mu^2$



• Given N observations  $x_1, \ldots, x_N$  from a univariate Gaussian  $N(x|\mu, \sigma^2)$ , doing moment-matching

$$\mathbb{E}[\phi(x)] = \frac{1}{N} \sum_{i=1}^{N} \phi(x_i)$$

$$\mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix} = \begin{bmatrix} \frac{1}{N} \sum_{i=1}^{N} x_i \\ \frac{1}{N} \sum_{i=1}^{N} x_i^2 \end{bmatrix}$$

- For a univariate Gaussian, note that  $\mathbb{E}[x] = \mu$  and  $\mathbb{E}[x^2] = \text{var}[x] + \mathbb{E}[x]^2 = \sigma^2 + \mu^2$
- Thus we have two equations and two unknowns
- From the first equation, we immediately get  $\mu = \frac{1}{N} \sum_{i=1}^{N} x_i$
- $\bullet$  From the second equation, we get  $\sigma^2=\mathbb{E}[\mathbf{x}^2]-\mu^2=\frac{1}{N}\sum_{i=1}^N x_i^2-\mu^2$



# Moment Matching: An Example

• Given N observations  $x_1, \ldots, x_N$  from a univariate Gaussian  $N(x|\mu, \sigma^2)$ , doing moment-matching

$$\mathbb{E}[\phi(x)] = \frac{1}{N} \sum_{i=1}^{N} \phi(x_i)$$

• The "true", i.e., expected moments:  $\mathbb{E}[\phi(x)] = \mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix}$ . Therefore

$$\mathbb{E}\begin{bmatrix} x \\ x^2 \end{bmatrix} = \begin{bmatrix} \frac{1}{N} \sum_{i=1}^{N} x_i \\ \frac{1}{N} \sum_{i=1}^{N} x_i^2 \end{bmatrix}$$

- For a univariate Gaussian, note that  $\mathbb{E}[x] = \mu$  and  $\mathbb{E}[x^2] = \text{var}[x] + \mathbb{E}[x]^2 = \sigma^2 + \mu^2$
- Thus we have two equations and two unknowns
- From the first equation, we immediately get  $\mu = \frac{1}{N} \sum_{i=1}^{N} x_i$
- From the second equation, we get  $\sigma^2 = \mathbb{E}[x^2] \mu^2 = \frac{1}{N} \sum_{i=1}^{N} x_i^2 \mu^2 = \frac{1}{N} \sum_{i=1}^{N} (x_i \mu)^2$



ullet We saw that the total likelihood given N i.i.d. observations  $\mathcal{D}\{m{x_1},\dots,m{x_N}\}$ 

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - NA(\theta)\right]$$
 where  $\phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(\mathbf{x}_i)$ 

ullet We saw that the total likelihood given N i.i.d. observations  $\mathcal{D}\{m{x}_1,\dots,m{x}_N\}$ 

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - \textit{NA}(\theta)
ight] \qquad ext{where} \quad \phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(m{x}_i)$$

ullet Let's choose the following prior (note: it looks similar in terms of heta within the exponent)

$$\left| p(\theta|\nu_0, \boldsymbol{\tau}_0) = h(\theta) \exp\left[\theta^{\top} \boldsymbol{\tau}_0 - \boldsymbol{\nu}_0 A(\theta) - A_c(\nu_0, \boldsymbol{\tau}_0)\right] \right|$$

• We saw that the total likelihood given N i.i.d. observations  $\mathcal{D}\{x_1,\ldots,x_N\}$ 

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - NA(\theta)\right]$$
 where  $\phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(\mathbf{x}_i)$ 

ullet Let's choose the following prior (note: it looks similar in terms of heta within the exponent)

$$oxed{p( heta|
u_0,oldsymbol{ au}_0) = h( heta) \exp\left[ heta^ op_{oldsymbol{ au}} A( heta) - oldsymbol{\lambda}_c(
u_0,oldsymbol{ au}_0)
ight]}$$

• Ignoring the prior's log-partition function  $A_c(\nu_0, \tau_0) = \log \int_{\theta} h(\theta) \exp \left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right] d\theta$ 

ullet We saw that the total likelihood given N i.i.d. observations  $\mathcal{D}\{m{x}_1,\dots,m{x}_N\}$ 

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{ op}\phi(\mathcal{D}) - \textit{NA}(\theta)
ight] \qquad ext{where} \quad \phi(\mathcal{D}) = \sum_{i=1}^N \phi(\pmb{x}_i)$$

ullet Let's choose the following prior (note: it looks similar in terms of heta within the exponent)

$$oxed{p( heta|
u_0,oldsymbol{ au}_0) = h( heta) \exp\left[ heta^ op_{oldsymbol{ au}_0} - oldsymbol{
u}_0 A( heta) - A_c(
u_0,oldsymbol{ au}_0)
ight]}$$

• Ignoring the prior's log-partition function  $A_c(\nu_0, \tau_0) = \log \int_{\theta} h(\theta) \exp \left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right] d\theta$ 

$$oxed{p( heta|
u_0, oldsymbol{ au}_0) \propto h( heta) \exp\left[ heta^ op oldsymbol{ au}_0 - oldsymbol{
u}_0 A( heta)
ight]}$$

ullet We saw that the total likelihood given N i.i.d. observations  $\mathcal{D}\{m{x}_1,\ldots,m{x}_N\}$ 

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{ op}\phi(\mathcal{D}) - \textit{NA}(\theta)
ight] \qquad ext{where} \quad \phi(\mathcal{D}) = \sum_{i=1}^N \phi(\pmb{x}_i)$$

ullet Let's choose the following prior (note: it looks similar in terms of heta within the exponent)

$$oxed{p( heta|
u_0,oldsymbol{ au}_0) = h( heta) \exp\left[ heta^ op_{oldsymbol{ au}_0} - oldsymbol{
u}_0 A( heta) - A_c(
u_0,oldsymbol{ au}_0)
ight]}$$

• Ignoring the prior's log-partition function  $A_c(\nu_0, \tau_0) = \log \int_{\theta} h(\theta) \exp \left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right] d\theta$ 

$$p( heta|
u_0, oldsymbol{ au}_0) \propto h( heta) \exp\left[ heta^ op oldsymbol{ au}_0 - oldsymbol{
u}_0 A( heta)
ight]$$

• Comparing the prior's form with the likelihood, we notice that

ullet We saw that the total likelihood given N i.i.d. observations  $\mathcal{D}\{m{x}_1,\dots,m{x}_N\}$ 

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{ op}\phi(\mathcal{D}) - \textit{NA}(\theta)
ight] \qquad ext{where} \quad \phi(\mathcal{D}) = \sum_{i=1}^N \phi(oldsymbol{x}_i)$$

ullet Let's choose the following prior (note: it looks similar in terms of heta within the exponent)

$$oxed{p( heta|
u_0,oldsymbol{ au}_0) = h( heta) \exp\left[ heta^ op_{oldsymbol{ au}_0} - oldsymbol{
u}_0 A( heta) - A_c(
u_0,oldsymbol{ au}_0)
ight]}$$

• Ignoring the prior's log-partition function  $A_c(\nu_0, \tau_0) = \log \int_{\theta} h(\theta) \exp \left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right] d\theta$ 

$$igg| p( heta|
u_0, oldsymbol{ au}_0) \propto h( heta) \exp\left[ heta^ op oldsymbol{ au}_0 - oldsymbol{
u}_0 A( heta)
ight]$$

- Comparing the prior's form with the likelihood, we notice that
  - $\bullet$   $\nu_0$  is like the number of "pseudo-observations" coming from the prior



• We saw that the total likelihood given N i.i.d. observations  $\mathcal{D}\{x_1,\ldots,x_N\}$ 

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{ op}\phi(\mathcal{D}) - \textit{NA}(\theta)
ight] \qquad ext{where} \quad \phi(\mathcal{D}) = \sum_{i=1}^N \phi(oldsymbol{x}_i)$$

ullet Let's choose the following prior (note: it looks similar in terms of heta within the exponent)

$$oxed{p( heta|
u_0,oldsymbol{ au}_0) = h( heta) \exp\left[ heta^ op_{oldsymbol{ au}_0} - oldsymbol{
u}_0 A( heta) - A_c(
u_0,oldsymbol{ au}_0)
ight]}$$

• Ignoring the prior's log-partition function  $A_c(\nu_0, \tau_0) = \log \int_{\theta} h(\theta) \exp \left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right] d\theta$ 

$$p( heta|
u_0, oldsymbol{ au}_0) \propto h( heta) \exp\left[ heta^ op oldsymbol{ au}_0 - oldsymbol{
u}_0 A( heta)
ight]$$

- Comparing the prior's form with the likelihood, we notice that
  - $\bullet$   $\nu_0$  is like the number of "pseudo-observations" coming from the prior
  - $\tau_0$  is the <u>total sufficient statistics</u> of these  $\nu_0$  pseudo-observations



As we saw, the likelihood is

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - NA(\theta)\right]$$
 where  $\phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(\mathbf{x}_i)$ 

• And the prior we chose is

$$p(\theta|
u_0, oldsymbol{ au}_0) \propto h(\theta) \exp\left[ heta^ op oldsymbol{ au}_0 - oldsymbol{
u}_0 A( heta)
ight]$$

As we saw, the likelihood is

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - \mathit{NA}(\theta)\right]$$
 where  $\phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(x_i)$ 

And the prior we chose is

$$p(\theta|\nu_0, \boldsymbol{\tau}_0) \propto h(\theta) \exp\left[\theta^{\top} \boldsymbol{\tau}_0 - \boldsymbol{\nu}_0 A(\theta)\right]$$

$$\boxed{\rho(\theta|\mathcal{D}) \propto h(\theta) \exp\left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta)\right]}$$

As we saw, the likelihood is

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - \mathit{NA}(\theta)\right]$$
 where  $\phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(x_i)$ 

And the prior we chose is

$$p(\theta|\nu_0, oldsymbol{ au}_0) \propto h(\theta) \exp\left[ heta^ op oldsymbol{ au}_0 - oldsymbol{
u}_0 A( heta) 
ight]$$

• For this form of the prior, the posterior  $p(\theta|\mathcal{D}) \propto p(\theta)p(\mathcal{D}|\theta)$  will be

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp \left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta)\right]$$

Note that the posterior has the same form as the prior

As we saw, the likelihood is

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - \mathit{NA}(\theta)\right]$$
 where  $\phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(x_i)$ 

• And the prior we chose is

$$p(\theta|\nu_0, oldsymbol{ au}_0) \propto h(\theta) \exp\left[ heta^ op oldsymbol{ au}_0 - oldsymbol{
u}_0 A( heta) 
ight]$$

• For this form of the prior, the posterior  $p(\theta|\mathcal{D}) \propto p(\theta)p(\mathcal{D}|\theta)$  will be

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp\left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta)\right]$$

Note that the posterior has the same form as the prior; such a prior is called a conjugate prior

As we saw, the likelihood is

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - \mathit{NA}(\theta)\right]$$
 where  $\phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(\mathbf{x}_i)$ 

• And the prior we chose is

$$p(\theta|\nu_0, oldsymbol{ au}_0) \propto h(\theta) \exp\left[ heta^ op oldsymbol{ au}_0 - oldsymbol{
u}_0 A( heta) 
ight]$$

• For this form of the prior, the posterior  $p(\theta|\mathcal{D}) \propto p(\theta)p(\mathcal{D}|\theta)$  will be

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp \left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta)\right]$$

 Note that the posterior has the same form as the prior; such a prior is called a conjugate prior (note: all exponential family distributions have a conjugate prior having a form shown as above)

As we saw, the likelihood is

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - NA(\theta)\right]$$
 where  $\phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(\mathbf{x}_i)$ 

• And the prior we chose is

$$p(\theta|\nu_0, \boldsymbol{\tau}_0) \propto h(\theta) \exp\left[\theta^{\top} \boldsymbol{\tau}_0 - \boldsymbol{\nu}_0 A(\theta)\right]$$

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp \left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta)\right]$$

- Note that the posterior has the same form as the prior; such a prior is called a conjugate prior (note: all exponential family distributions have a conjugate prior having a form shown as above)
- ullet Thus posterior hyperparams  $u_0{}', au_0{}'$  are obtained by simply adding "stuff" to prior's hyperparams

As we saw, the likelihood is

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - NA(\theta)\right]$$
 where  $\phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(\mathbf{x}_i)$ 

• And the prior we chose is

$$p(\theta|\nu_0, oldsymbol{ au}_0) \propto h(\theta) \exp\left[ heta^ op oldsymbol{ au}_0 - oldsymbol{
u}_0 A( heta) 
ight]$$

$$\boxed{p(\theta|\mathcal{D}) \propto h(\theta) \exp\left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\textcolor{red}{\nu_0} + N)A(\theta)\right]}$$

- Note that the posterior has the same form as the prior; such a prior is called a conjugate prior (note: all exponential family distributions have a conjugate prior having a form shown as above)
- Thus posterior hyperparams  $\nu_0{}', \tau_0{}'$  are obtained by simply adding "stuff" to prior's hyperparams  $\nu_0{}' \leftarrow \nu_0 + N$  (no. of pseudo-obs + no. of actual obs)

As we saw, the likelihood is

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - \mathit{NA}(\theta)\right]$$
 where  $\phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(\mathbf{x}_i)$ 

• And the prior we chose is

$$p(\theta|\nu_0, oldsymbol{ au}_0) \propto h(\theta) \exp\left[ heta^ op oldsymbol{ au}_0 - oldsymbol{
u}_0 A( heta) 
ight]$$

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp \left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta)\right]$$

- Note that the posterior has the same form as the prior; such a prior is called a **conjugate prior** (note: all exponential family distributions have a conjugate prior having a form shown as above)
- Thus posterior hyperparams  $\nu_0{'}, \tau_0{'}$  are obtained by simply adding "stuff" to prior's hyperparams  $\nu_0{'} \leftarrow \nu_0 + N$  (no. of pseudo-obs + no. of actual obs)  $\tau_0{'} \leftarrow \tau_0 + \phi(\mathcal{D})$  (total suff-stats from pseudo-obs + total suff-stats from actual obs)



As we saw, the likelihood is

$$p(\mathcal{D}|\theta) \propto \exp\left[\theta^{\top}\phi(\mathcal{D}) - \mathit{NA}(\theta)\right]$$
 where  $\phi(\mathcal{D}) = \sum_{i=1}^{N}\phi(x_i)$ 

• And the prior we chose is

$$p(\theta|\nu_0, \boldsymbol{ au}_0) \propto h(\theta) \exp\left[\theta^{\top} \boldsymbol{ au}_0 - \boldsymbol{
u}_0 A(\theta)\right]$$

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp \left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta)\right]$$

- Note that the posterior has the same form as the prior; such a prior is called a **conjugate prior** (note: all exponential family distributions have a conjugate prior having a form shown as above)
- Thus posterior hyperparams  $\nu_0{'}, \tau_0{'}$  are obtained by simply adding "stuff" to prior's hyperparams  $\nu_0{'} \leftarrow \nu_0 + N$  (no. of pseudo-obs + no. of actual obs)  $\tau_0{'} \leftarrow \tau_0 + \phi(\mathcal{D})$  (total suff-stats from pseudo-obs + total suff-stats from actual obs)
- Note: Prior's log-partition function  $A_c(\nu_0, \tau_0)$  updates to posterior's:  $A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D}))$

• Assuming the prior  $p(\theta|\nu_0, \tau_0) \propto h(\theta) \exp\left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right]$ , the posterior was

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp\left[\theta^{\top}( au_0 + \phi(\mathcal{D})) - (
u_0 + N)A(\theta)\right]$$

• Assuming the prior  $p(\theta|\nu_0, \tau_0) \propto h(\theta) \exp\left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right]$ , the posterior was

$$p( heta|\mathcal{D}) \propto h( heta) \exp\left[ heta^ op ( au_0 + \phi(\mathcal{D})) - (oldsymbol{
u}_0 + N)A( heta)
ight]$$

• Assuming  $\tau_0 = \nu_0 \bar{\tau}_0$ , we can also write the prior as  $p(\theta|\nu_0, \bar{\tau}_0) \propto \exp\left[\theta^\top \nu_0 \bar{\tau}_0 - \nu_0 A(\theta)\right]$ 

• Assuming the prior  $p(\theta|\nu_0, \tau_0) \propto h(\theta) \exp\left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right]$ , the posterior was

$$p( heta|\mathcal{D}) \propto h( heta) \exp\left[ heta^ op ( au_0 + \phi(\mathcal{D})) - (oldsymbol{
u}_0 + oldsymbol{N})A( heta)
ight]$$

- Assuming  $\tau_0 = \nu_0 \bar{\tau}_0$ , we can also write the prior as  $p(\theta|\nu_0, \bar{\tau}_0) \propto \exp\left[\theta^\top \nu_0 \bar{\tau}_0 \nu_0 A(\theta)\right]$
- Can think of  $\bar{\tau}_0 = \tau_0/\nu_0$  as the average sufficient statistics per pseudo-observation

• Assuming the prior  $p(\theta|\nu_0, \tau_0) \propto h(\theta) \exp\left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right]$ , the posterior was

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp\left[\theta^{\top}( au_0 + \phi(\mathcal{D})) - (
u_0 + N)A(\theta)\right]$$

- Assuming  $\tau_0 = \nu_0 \bar{\tau}_0$ , we can also write the prior as  $p(\theta|\nu_0, \bar{\tau}_0) \propto \exp\left[\theta^\top \nu_0 \bar{\tau}_0 \nu_0 A(\theta)\right]$
- ullet Can think of  $ar{ au}_0= au_0/
  u_0$  as the average sufficient statistics per pseudo-observation
- The posterior can be written as

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp \left[ \theta^{\top} (\mathbf{\nu_0} + N) \frac{\nu_0 \bar{\tau}_0 + \phi(\mathcal{D})}{\nu_0 + N} - (\mathbf{\nu_0} + N) A(\theta) \right]$$

• Assuming the prior  $p(\theta|\nu_0, \tau_0) \propto h(\theta) \exp\left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right]$ , the posterior was

$$p( heta|\mathcal{D}) \propto h( heta) \exp\left[ heta^ op ( au_0 + \phi(\mathcal{D})) - (oldsymbol{
u}_0 + oldsymbol{N})A( heta)
ight]$$

- Assuming  $\tau_0 = \nu_0 \bar{\tau}_0$ , we can also write the prior as  $p(\theta|\nu_0, \bar{\tau}_0) \propto \exp\left[\theta^\top \nu_0 \bar{\tau}_0 \nu_0 A(\theta)\right]$
- ullet Can think of  $ar{ au}_0= au_0/
  u_0$  as the average sufficient statistics per pseudo-observation
- The posterior can be written as

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp \left[ heta^{ op} ( oldsymbol{
u}_0 + N ) rac{
u_0 ar{ au}_0 + \phi(\mathcal{D})}{
u_0 + N} - (
u_0 + N) A( heta) 
ight]$$

ullet Denoting  $ar{\phi}=rac{\phi(D)}{N}$  as the average suff-stats per real observation, the posterior updates are

• Assuming the prior  $p(\theta|\nu_0, \tau_0) \propto h(\theta) \exp\left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right]$ , the posterior was

$$p( heta|\mathcal{D}) \propto h( heta) \exp\left[ heta^ op ( au_0 + \phi(\mathcal{D})) - (oldsymbol{
u}_0 + N)A( heta)
ight]$$

- Assuming  $\tau_0 = \nu_0 \bar{\tau}_0$ , we can also write the prior as  $p(\theta|\nu_0, \bar{\tau}_0) \propto \exp\left[\theta^\top \nu_0 \bar{\tau}_0 \nu_0 A(\theta)\right]$
- ullet Can think of  $ar{ au}_0= au_0/
  u_0$  as the average sufficient statistics per pseudo-observation
- The posterior can be written as

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp \left[ \theta^{\top} (\nu_0 + N) \frac{\nu_0 \bar{\tau}_0 + \phi(\mathcal{D})}{\nu_0 + N} - (\nu_0 + N) A(\theta) \right]$$

ullet Denoting  $ar{\phi}=rac{\phi(D)}{N}$  as the average suff-stats per real observation, the posterior updates are

$$\nu_0' \leftarrow \nu_0 + N$$

• Assuming the prior  $p(\theta|\nu_0, \tau_0) \propto h(\theta) \exp\left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right]$ , the posterior was

$$p( heta|\mathcal{D}) \propto h( heta) \exp\left[ heta^ op ( au_0 + \phi(\mathcal{D})) - (oldsymbol{
u}_0 + oldsymbol{N})A( heta)
ight]$$

- Assuming  $\tau_0 = \nu_0 \bar{\tau}_0$ , we can also write the prior as  $p(\theta|\nu_0, \bar{\tau}_0) \propto \exp\left[\theta^\top \nu_0 \bar{\tau}_0 \nu_0 A(\theta)\right]$
- ullet Can think of  $ar{ au}_0= au_0/
  u_0$  as the average sufficient statistics per pseudo-observation
- The posterior can be written as

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp \left[ \theta^{\top} (\nu_0 + N) \frac{\nu_0 \bar{\tau}_0 + \phi(\mathcal{D})}{\nu_0 + N} - (\nu_0 + N) A(\theta) \right]$$

ullet Denoting  $ar{\phi}=rac{\phi(D)}{N}$  as the average suff-stats per real observation, the posterior updates are

$$\overline{\tau}_0' \leftarrow \nu_0 + N$$

$$\overline{\tau}_0' \leftarrow \frac{\nu_0 \overline{\tau}_0 + N \overline{\phi}}{\nu_0 + N}$$

• Assuming the prior  $p(\theta|\nu_0, \tau_0) \propto h(\theta) \exp\left[\theta^{\top} \tau_0 - \nu_0 A(\theta)\right]$ , the posterior was

$$p( heta|\mathcal{D}) \propto h( heta) \exp\left[ heta^ op ( au_0 + \phi(\mathcal{D})) - (oldsymbol{
u}_0 + oldsymbol{N})A( heta)
ight]$$

- Assuming  $\tau_0 = \nu_0 \bar{\tau}_0$ , we can also write the prior as  $p(\theta|\nu_0, \bar{\tau}_0) \propto \exp\left[\theta^\top \nu_0 \bar{\tau}_0 \nu_0 A(\theta)\right]$
- ullet Can think of  $ar{ au}_0= au_0/
  u_0$  as the average sufficient statistics per pseudo-observation
- The posterior can be written as

$$p(\theta|\mathcal{D}) \propto h(\theta) \exp\left[\theta^{\top} (\nu_0 + N) \frac{\nu_0 \bar{\tau}_0 + \phi(\mathcal{D})}{\nu_0 + N} - (\nu_0 + N) A(\theta)\right]$$

ullet Denoting  $ar{\phi}=rac{\phi(D)}{N}$  as the average suff-stats per real observation, the posterior updates are

$$\begin{array}{ccc} {\nu_0}' & \leftarrow & \nu_0 + N \\ \bar{\tau}_0' & \leftarrow & \frac{\nu_0 \bar{\tau}_0 + N \bar{\phi}}{\nu_0 + N} \end{array}$$

• Note that the posterior hyperparam  $\bar{\tau}_0'$  is a convex combination of the average suff-stats  $\bar{\tau}_0$  of the  $\nu_0$  pseudo-observations and the average suff-stats  $\bar{\phi}$  of the N actual observations

ullet Assume some past (training) data  $\mathcal{D} = \{ oldsymbol{x}_1, \dots, oldsymbol{x}_N \}$  generated from an exp. family distribution

- ullet Assume some past (training) data  $\mathcal{D} = \{ oldsymbol{x}_1, \dots, oldsymbol{x}_N \}$  generated from an exp. family distribution
- ullet Assme some test data  $\mathcal{D}' = \{ ilde{m{x}}_1, \dots, ilde{m{x}}_{N'}\}$  from the same distribution  $(N' \geq 1)$

- ullet Assume some past (training) data  $\mathcal{D} = \{ oldsymbol{x}_1, \dots, oldsymbol{x}_N \}$  generated from an exp. family distribution
- Assme some test data  $\mathcal{D}' = \{\tilde{\mathbf{x}}_1, \dots, \tilde{\mathbf{x}}_{N'}\}$  from the same distribution  $(N' \geq 1)$
- ullet The posterior predictive distribution of  $\mathcal{D}'$  (probability distribution of new data given old data)

$$p(\mathcal{D}'|\mathcal{D}) = \int p(\mathcal{D}'| heta)p( heta|\mathcal{D})d heta$$

- Assume some past (training) data  $\mathcal{D} = \{x_1, \dots, x_N\}$  generated from an exp. family distribution
- Assme some test data  $\mathcal{D}' = \{\tilde{\pmb{x}}_1, \dots, \tilde{\pmb{x}}_{N'}\}$  from the same distribution  $(N' \geq 1)$
- ullet The posterior predictive distribution of  $\mathcal{D}'$  (probability distribution of new data given old data)

$$p(\mathcal{D}'|\mathcal{D}) = \int p(\mathcal{D}'| heta)p( heta|\mathcal{D})d heta$$

• We've already seen some specific examples of computing the posterior predictive dist., e.g.,

- ullet Assume some past (training) data  $\mathcal{D} = \{ oldsymbol{x}_1, \dots, oldsymbol{x}_N \}$  generated from an exp. family distribution
- Assme some test data  $\mathcal{D}' = \{\tilde{\pmb{x}}_1, \dots, \tilde{\pmb{x}}_{N'}\}$  from the same distribution  $(N' \geq 1)$
- ullet The posterior predictive distribution of  $\mathcal{D}'$  (probability distribution of new data given old data)

$$p(\mathcal{D}'|\mathcal{D}) = \int p(\mathcal{D}'| heta)p( heta|\mathcal{D})d heta$$

- We've already seen some specific examples of computing the posterior predictive dist., e.g.,
  - Beta-Bernoulli case: Posterior predictive distribution of next coin toss

- ullet Assume some past (training) data  $\mathcal{D} = \{ oldsymbol{x}_1, \dots, oldsymbol{x}_N \}$  generated from an exp. family distribution
- Assme some test data  $\mathcal{D}' = \{\tilde{\pmb{x}}_1, \dots, \tilde{\pmb{x}}_{N'}\}$  from the same distribution  $(N' \geq 1)$
- ullet The posterior predictive distribution of  $\mathcal{D}'$  (probability distribution of new data given old data)

$$p(\mathcal{D}'|\mathcal{D}) = \int p(\mathcal{D}'| heta)p( heta|\mathcal{D})d heta$$

- We've already seen some specific examples of computing the posterior predictive dist., e.g.,
  - Beta-Bernoulli case: Posterior predictive distribution of next coin toss
  - Bayesian linear regression: Posterior predictive distribution of the response  $y_*$  of test input  $x_*$

- ullet Assume some past (training) data  $\mathcal{D} = \{ oldsymbol{x}_1, \dots, oldsymbol{x}_N \}$  generated from an exp. family distribution
- Assme some test data  $\mathcal{D}' = \{\tilde{\pmb{x}}_1, \dots, \tilde{\pmb{x}}_{N'}\}$  from the same distribution  $(N' \geq 1)$
- ullet The posterior predictive distribution of  $\mathcal{D}'$  (probability distribution of new data given old data)

$$p(\mathcal{D}'|\mathcal{D}) = \int p(\mathcal{D}'| heta)p( heta|\mathcal{D})d heta$$

- We've already seen some specific examples of computing the posterior predictive dist., e.g.,
  - Beta-Bernoulli case: Posterior predictive distribution of next coin toss
  - ullet Bayesian linear regression: Posterior predictive distribution of the response  $y_*$  of test input  $x_*$
- Nice Property: If the likelihood is an exponential family distribution, prior is conjugate (and thus is the posterior), the posterior predictive always has a closed form expression (shown next)



• Recall the form of the likelihood  $p(\mathcal{D}|\theta)$  for exp. family dist.

$$p(\mathcal{D}|\theta) = \left[\prod_{i=1}^{N} h(x_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) - NA(\theta)\right]$$

• The conjugate prior was

$$p( heta|
u_0,oldsymbol{ au}_0) = h( heta) \exp\left[ heta^ op au_0 - 
u_0 A( heta) - A_c(
u_0,oldsymbol{ au}_0)
ight]$$

• Recall the form of the likelihood  $p(\mathcal{D}|\theta)$  for exp. family dist.

$$p(\mathcal{D}|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) - NA(\theta)\right]$$

• The conjugate prior was

$$p( heta|
u_0,oldsymbol{ au}_0) = h( heta) \exp\left[ heta^ op au_0 - 
u_0 A( heta) - A_c(
u_0,oldsymbol{ au}_0)
ight]$$

• For this choice of the conjugate prior, the posterior was shown to be

$$p(\theta|\mathcal{D}) = h(\theta) \exp \left[ \theta^{\top} (\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N) A(\theta) - A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D})) \right]$$

• Recall the form of the likelihood  $p(\mathcal{D}|\theta)$  for exp. family dist.

$$p(\mathcal{D}|\theta) = \left[\prod_{i=1}^{N} h(\mathbf{x}_i)\right] \exp\left[\theta^{\top} \phi(\mathcal{D}) - NA(\theta)\right]$$

The conjugate prior was

$$p( heta|
u_0,oldsymbol{ au}_0) = h( heta) \exp\left[ heta^ op au_0 - 
u_0 A( heta) - A_c(
u_0,oldsymbol{ au}_0)
ight]$$

• For this choice of the conjugate prior, the posterior was shown to be

$$p(\theta|\mathcal{D}) = h(\theta) \exp \left[ \theta^{\top} (\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N) A(\theta) - A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D})) \right]$$

ullet For the test data  $\mathcal{D}'$ , the likelihood will be

$$p(\mathcal{D}'|\theta) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[\theta^\top \phi(\mathcal{D}') - N'A(\theta)\right] \qquad \text{where} \quad \phi(\mathcal{D}') = \sum_{i=1}^{N'} \phi(\tilde{\mathbf{x}}_i)$$



• Therefore the posterior predictive distribution will be

$$p(\mathcal{D}'|\mathcal{D}) = \int p(\mathcal{D}'|\theta)p(\theta|\mathcal{D})d\theta$$

• Therefore the posterior predictive distribution will be

$$p(\mathcal{D}'|\mathcal{D}) = \int p(\mathcal{D}'|\theta)p(\theta|\mathcal{D})d\theta$$

$$= \int \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[\theta^{\top}\phi(\mathcal{D}') - N'A(\theta)\right]h(\theta) \exp\left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta) - \underbrace{A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D}))}_{\text{constant w.r.t. }\theta}\right]d\theta$$

• Therefore the posterior predictive distribution will be

$$p(\mathcal{D}'|\mathcal{D}) = \int p(\mathcal{D}'|\theta)p(\theta|\mathcal{D})d\theta$$

$$= \int \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[\theta^{\top}\phi(\mathcal{D}') - N'A(\theta)\right]h(\theta) \exp\left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta) - \underbrace{A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D}))}_{\text{constant w.r.t. }\theta}\right]d\theta$$

• The above gets simplified further into

• Therefore the posterior predictive distribution will be

$$\rho(\mathcal{D}'|\mathcal{D}) = \int \rho(\mathcal{D}'|\theta)\rho(\theta|\mathcal{D})d\theta$$

$$= \int \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[\theta^{\top}\phi(\mathcal{D}') - N'A(\theta)\right]h(\theta) \exp\left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta) - \underbrace{A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D}))}_{\text{constant w.r.t. }\theta}\right]d\theta$$

The above gets simplified further into

$$p(\mathcal{D}'|\mathcal{D}) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{\int h(\theta) \exp\left[\theta^{\top}(\tau_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) - (\nu_0 + N + N')A(\theta)\right] d\theta}{\exp\left[A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D}))\right]}$$

• Therefore the posterior predictive distribution will be

$$\rho(\mathcal{D}'|\mathcal{D}) = \int \rho(\mathcal{D}'|\theta)\rho(\theta|\mathcal{D})d\theta$$

$$= \int \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[\theta^{\top}\phi(\mathcal{D}') - N'A(\theta)\right]h(\theta) \exp\left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta) - \underbrace{A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D}))}_{\text{constant w.r.t. }\theta}\right]d\theta$$

The above gets simplified further into

$$p(\mathcal{D}'|\mathcal{D}) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{\int h(\theta) \exp\left[\theta^{\top}(\tau_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) - (\nu_0 + N + N')A(\theta)\right] d\theta}{\exp\left[A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D}))\right]}$$
$$= \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{Z_c(\nu_0 + N + N', \tau_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}'))}{\exp\left[A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D}))\right]}$$

Therefore the posterior predictive distribution will be

$$p(\mathcal{D}'|\mathcal{D}) = \int p(\mathcal{D}'|\theta)p(\theta|\mathcal{D})d\theta$$

$$= \int \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[\theta^{\top}\phi(\mathcal{D}') - N'A(\theta)\right]h(\theta) \exp\left[\theta^{\top}(\tau_0 + \phi(\mathcal{D})) - (\nu_0 + N)A(\theta) - \underbrace{A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D}))}_{\text{constant w.r.t. }\theta}\right]d\theta$$

• The above gets simplified further into

$$p(\mathcal{D}'|\mathcal{D}) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{\int h(\theta) \exp\left[\theta^{\top}(\tau_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) - (\nu_0 + N + N')A(\theta)\right] d\theta}{\exp\left[A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D}))\right]}$$
$$= \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{Z_c(\nu_0 + N + N', \tau_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}'))}{\exp\left[A_c(\nu_0 + N, \tau_0 + \phi(\mathcal{D}))\right]}$$

where  $Z_c(\nu_0 + N + N', \tau_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) = \int h(\theta) \exp\left[\theta^\top (\tau_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) - (\nu_0 + N + N')A(\theta)\right] d\theta$ 

$$p(\mathcal{D}'|\mathcal{D}) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{Z_c(\nu_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}'))}{Z_c(\nu_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))}$$

$$p(\mathcal{D}'|\mathcal{D}) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{Z_c(\nu_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}'))}{Z_c(\nu_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))}$$

$$= \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[A_c(\nu_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) - A_c(\nu_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))\right]$$

• Since  $A_c = \log Z_c$  or  $Z_c = \exp(A_c)$ , we can write the posterior predictive distribution as

$$\rho(\mathcal{D}'|\mathcal{D}) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{Z_c(\nu_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}'))}{Z_c(\nu_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))}$$

$$= \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[A_c(\nu_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) - A_c(\nu_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))\right]$$

• Therefore the posterior predictive is proportional to ..

$$p(\mathcal{D}'|\mathcal{D}) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{Z_c(\boldsymbol{\nu}_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}'))}{Z_c(\boldsymbol{\nu}_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))}$$

$$= \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[A_c(\boldsymbol{\nu}_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) - A_c(\boldsymbol{\nu}_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))\right]$$

- Therefore the posterior predictive is proportional to ..
  - .. the ratio of two partition functions of two "posterior distributions" (one with N+N' examples and the other with N examples)

$$\rho(\mathcal{D}'|\mathcal{D}) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{Z_c(\nu_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}'))}{Z_c(\nu_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))}$$

$$= \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[A_c(\nu_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) - A_c(\nu_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))\right]$$

- Therefore the posterior predictive is proportional to ..
  - .. the ratio of two partition functions of two "posterior distributions" (one with N + N' examples and the other with N examples)
  - .. or exponential of the difference of the corresponding log-partition functions

$$\rho(\mathcal{D}'|\mathcal{D}) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{Z_c(\nu_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}'))}{Z_c(\nu_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))}$$

$$= \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[A_c(\nu_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) - A_c(\nu_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))\right]$$

- Therefore the posterior predictive is proportional to ..
  - .. the ratio of two partition functions of two "posterior distributions" (one with N + N' examples and the other with N examples)
  - .. or exponential of the difference of the corresponding log-partition functions
- Note that the form of  $Z_c$  (and  $A_c$ ) will simply depend on the chosen conjugate prior

$$\rho(\mathcal{D}'|\mathcal{D}) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{Z_c(\boldsymbol{\nu}_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}'))}{Z_c(\boldsymbol{\nu}_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))}$$

$$= \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[A_c(\boldsymbol{\nu}_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) - A_c(\boldsymbol{\nu}_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))\right]$$

- Therefore the posterior predictive is proportional to ..
  - .. the ratio of two partition functions of two "posterior distributions" (one with N + N' examples and the other with N examples)
  - .. or exponential of the difference of the corresponding log-partition functions
- Note that the form of  $Z_c$  (and  $A_c$ ) will simply depend on the chosen conjugate prior
- Very useful result. Also holds for N=0



$$\rho(\mathcal{D}'|\mathcal{D}) = \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \frac{Z_c(\nu_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}'))}{Z_c(\nu_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))}$$

$$= \left[\prod_{i=1}^{N'} h(\tilde{\mathbf{x}}_i)\right] \exp\left[A_c(\nu_0 + N + N', \boldsymbol{\tau}_0 + \phi(\mathcal{D}) + \phi(\mathcal{D}')) - A_c(\nu_0 + N, \boldsymbol{\tau}_0 + \phi(\mathcal{D}))\right]$$

- Therefore the posterior predictive is proportional to ..
  - .. the ratio of two partition functions of two "posterior distributions" (one with N + N' examples and the other with N examples)
  - .. or exponential of the difference of the corresponding log-partition functions
- Note that the form of  $Z_c$  (and  $A_c$ ) will simply depend on the chosen conjugate prior
- Very useful result. Also holds for N=0
  - In the N=0 case,  $p(\mathcal{D}')=\int p(\mathcal{D}'|\theta)p(\theta)d\theta$  is simply the marginal likelihood of  $\mathcal{D}'$

# Exponential Family and GLM

 $\bullet$  (Probabilistic) Linear regression: when response y is real-valued

$$p(y|\mathbf{x}, \mathbf{w}) = \mathcal{N}(\mathbf{w}^{\top}\mathbf{x}, \beta^{-1})$$

 $\bullet$  (Probabilistic) Linear regression: when response y is real-valued

$$p(y|\mathbf{x},\mathbf{w}) = \mathcal{N}(\mathbf{w}^{\top}\mathbf{x}, \beta^{-1})$$

• Logistic regression: when response y is binary (0/1)

$$p(y|\mathbf{x}, \mathbf{w}) = \text{Bernoulli}(\sigma(\mathbf{w}^{\top}\mathbf{x})) = [\sigma(\mathbf{w}^{\top}\mathbf{x})]^{y}[1 - \sigma(\mathbf{w}^{\top}\mathbf{x})]^{1-y}$$

where 
$$\sigma(\mathbf{w}^{\top}\mathbf{x}) = \frac{1}{1 + \exp(-\mathbf{w}^{\top}\mathbf{x})} = \frac{\exp(\mathbf{w}^{\top}\mathbf{x})}{1 + \exp(\mathbf{w}^{\top}\mathbf{x})}$$

 $\bullet$  (Probabilistic) Linear regression: when response y is real-valued

$$p(y|\mathbf{x},\mathbf{w}) = \mathcal{N}(\mathbf{w}^{\top}\mathbf{x}, \beta^{-1})$$

• Logistic regression: when response y is binary (0/1)

$$p(y|\mathbf{x}, \mathbf{w}) = \text{Bernoulli}(\sigma(\mathbf{w}^{\top}\mathbf{x})) = [\sigma(\mathbf{w}^{\top}\mathbf{x})]^{y}[1 - \sigma(\mathbf{w}^{\top}\mathbf{x})]^{1-y}$$

where 
$$\sigma(\mathbf{w}^{\top}\mathbf{x}) = \frac{1}{1 + \exp(-\mathbf{w}^{\top}\mathbf{x})} = \frac{\exp(\mathbf{w}^{\top}\mathbf{x})}{1 + \exp(\mathbf{w}^{\top}\mathbf{x})}$$

• In both, the model depends on the inputs x as  $\mathbf{w}^{\top} \mathbf{x}$ 

 $\bullet$  (Probabilistic) Linear regression: when response y is real-valued

$$p(y|\mathbf{x},\mathbf{w}) = \mathcal{N}(\mathbf{w}^{\top}\mathbf{x}, \beta^{-1})$$

• Logistic regression: when response y is binary (0/1)

$$p(y|\mathbf{x}, \mathbf{w}) = \text{Bernoulli}(\sigma(\mathbf{w}^{\top}\mathbf{x})) = [\sigma(\mathbf{w}^{\top}\mathbf{x})]^{y}[1 - \sigma(\mathbf{w}^{\top}\mathbf{x})]^{1-y}$$

where 
$$\sigma(\mathbf{w}^{\top}\mathbf{x}) = \frac{1}{1 + \exp(-\mathbf{w}^{\top}\mathbf{x})} = \frac{\exp(\mathbf{w}^{\top}\mathbf{x})}{1 + \exp(\mathbf{w}^{\top}\mathbf{x})}$$

- In both, the model depends on the inputs x as  $w^{\top}x$
- Can we extend it to other type of outputs?

 $\bullet$  (Probabilistic) Linear regression: when response y is real-valued

$$p(y|\mathbf{x}, \mathbf{w}) = \mathcal{N}(\mathbf{w}^{\top}\mathbf{x}, \beta^{-1})$$

• Logistic regression: when response y is binary (0/1)

$$p(y|\mathbf{x}, \mathbf{w}) = \text{Bernoulli}(\sigma(\mathbf{w}^{\top}\mathbf{x})) = [\sigma(\mathbf{w}^{\top}\mathbf{x})]^{y}[1 - \sigma(\mathbf{w}^{\top}\mathbf{x})]^{1-y}$$

where 
$$\sigma(\mathbf{w}^{\top}\mathbf{x}) = \frac{1}{1 + \exp(-\mathbf{w}^{\top}\mathbf{x})} = \frac{\exp(\mathbf{w}^{\top}\mathbf{x})}{1 + \exp(\mathbf{w}^{\top}\mathbf{x})}$$

- In both, the model depends on the inputs x as  $w^{\top}x$
- Can we extend it to other type of outputs?
- Solution: Model the output using an exp-fam distribution (Gaussian and Bernoulli already are!)

$$p(y|\eta) = h(y) \exp(\eta y - A(\eta))$$
 (Generalized Linear Model (GLM))



• (Probabilistic) Linear regression: when response y is real-valued

$$p(y|\mathbf{x},\mathbf{w}) = \mathcal{N}(\mathbf{w}^{\top}\mathbf{x}, \beta^{-1})$$

• Logistic regression: when response y is binary (0/1)

$$p(y|\mathbf{x}, \mathbf{w}) = \text{Bernoulli}(\sigma(\mathbf{w}^{\top}\mathbf{x})) = [\sigma(\mathbf{w}^{\top}\mathbf{x})]^{y}[1 - \sigma(\mathbf{w}^{\top}\mathbf{x})]^{1-y}$$

where 
$$\sigma(\mathbf{w}^{\top}\mathbf{x}) = \frac{1}{1 + \exp(-\mathbf{w}^{\top}\mathbf{x})} = \frac{\exp(\mathbf{w}^{\top}\mathbf{x})}{1 + \exp(\mathbf{w}^{\top}\mathbf{x})}$$

- In both, the model depends on the inputs x as  $w^{\top}x$
- Can we extend it to other type of outputs?
- Solution: Model the output using an exp-fam distribution (Gaussian and Bernoulli already are!)

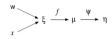
$$p(y|\eta) = h(y) \exp(\eta y - A(\eta))$$
 (Generalized Linear Model (GLM))

.. where  $\eta$  is a scalar-valued natural parameter (depends on  $\mathbf{x}$ ) and sufficient statistics  $\phi(\mathbf{y}) = \mathbf{y}$ 

• The GLM is of the form  $p(y|\eta) = h(y) \exp(\eta y - A(\eta))$  where  $\eta$  depends on  $\boldsymbol{x}$ 

- The GLM is of the form  $p(y|\eta) = h(y) \exp(\eta y A(\eta))$  where  $\eta$  depends on  $\boldsymbol{x}$
- The inputs  ${\pmb x}$  appear in the model only as a linear combination  ${\pmb \xi} = {\pmb w}^{ op} {\pmb x}$

- The GLM is of the form  $p(y|\eta) = h(y) \exp(\eta y A(\eta))$  where  $\eta$  depends on  $\boldsymbol{x}$
- The inputs x appear in the model only as a linear combination  $\xi = w^{\top}x$



- The GLM is of the form  $p(y|\eta) = h(y) \exp(\eta y A(\eta))$  where  $\eta$  depends on  $\boldsymbol{x}$
- The inputs  ${\pmb x}$  appear in the model only as a linear combination  ${\pmb \xi} = {\pmb w}^{ op} {\pmb x}$

$$\xi \xrightarrow{\psi} \eta$$

ullet Conditional mean  $\mu$  of a response y's distribution is modeled via a response function f

$$\mu = \mathbb{E}[y] = f(\xi) = f(\mathbf{w}^{\top} \mathbf{x})$$

- The GLM is of the form  $p(y|\eta) = h(y) \exp(\eta y A(\eta))$  where  $\eta$  depends on  $\boldsymbol{x}$
- ullet The inputs  $oldsymbol{x}$  appear in the model only as a linear combination  $oldsymbol{\xi} = oldsymbol{w}^{ op} oldsymbol{x}$

$$\xi \xrightarrow{f} \mu \xrightarrow{\psi} \eta$$

ullet Conditional mean  $\mu$  of a response y's distribution is modeled via a response function f

$$\mu = \mathbb{E}[y] = f(\xi) = f(\mathbf{w}^{\top} \mathbf{x})$$

- for (probabilistic) linear regression, f is identity, i.e.,  $\mu = f(\mathbf{w}^{\top}\mathbf{x}) = \mathbf{w}^{\top}\mathbf{x}$ ,

- The GLM is of the form  $p(y|\eta) = h(y) \exp(\eta y A(\eta))$  where  $\eta$  depends on  $\boldsymbol{x}$
- The inputs x appear in the model only as a linear combination  $\xi = w^{\top}x$

$$\xi \xrightarrow{f} \mu \xrightarrow{\psi} \eta$$

• Conditional mean  $\mu$  of a response y's distribution is modeled via a response function f

$$\mu = \mathbb{E}[y] = f(\xi) = f(\mathbf{w}^{\top} \mathbf{x})$$

- for (probabilistic) linear regression, f is identity, i.e.,  $\mu = f(\mathbf{w}^{\top}\mathbf{x}) = \mathbf{w}^{\top}\mathbf{x}$ ,
- for logistic regression, f is sigmoid, i.e.,  $\mu = f(\boldsymbol{w}^{\top}\boldsymbol{x}) = \exp(\boldsymbol{w}^{\top}\boldsymbol{x})/(1+\exp(\boldsymbol{w}^{\top}\boldsymbol{x}))$

- The GLM is of the form  $p(y|\eta) = h(y) \exp(\eta y A(\eta))$  where  $\eta$  depends on  $\boldsymbol{x}$
- The inputs x appear in the model only as a linear combination  $\xi = w^{\top}x$

$$\xi \xrightarrow{f} \mu \xrightarrow{\psi} \eta$$

ullet Conditional mean  $\mu$  of a response y's distribution is modeled via a response function f

$$\mu = \mathbb{E}[y] = f(\xi) = f(\mathbf{w}^{\top} \mathbf{x})$$

- for (probabilistic) linear regression, f is identity, i.e.,  $\mu = f(\mathbf{w}^{\top}\mathbf{x}) = \mathbf{w}^{\top}\mathbf{x}$ ,
- for logistic regression, f is sigmoid, i.e.,  $\mu = f(\boldsymbol{w}^{\top}\boldsymbol{x}) = \exp(\boldsymbol{w}^{\top}\boldsymbol{x})/(1+\exp(\boldsymbol{w}^{\top}\boldsymbol{x}))$
- The natural parameter  $\eta=\psi(\mu)$  where  $\psi$  is the link function

- The GLM is of the form  $p(y|\eta) = h(y) \exp(\eta y A(\eta))$  where  $\eta$  depends on  $\boldsymbol{x}$
- The inputs x appear in the model only as a linear combination  $\xi = \mathbf{w}^{\top} \mathbf{x}$

$$\xi \xrightarrow{f} \mu \xrightarrow{\psi} \eta$$

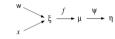
ullet Conditional mean  $\mu$  of a response y's distribution is modeled via a response function f

$$\mu = \mathbb{E}[y] = f(\xi) = f(\mathbf{w}^{\top} \mathbf{x})$$

- for (probabilistic) linear regression, f is identity, i.e.,  $\mu = f(\mathbf{w}^{\top}\mathbf{x}) = \mathbf{w}^{\top}\mathbf{x}$ ,
- for logistic regression, f is sigmoid, i.e.,  $\mu = f(\boldsymbol{w}^{\top}\boldsymbol{x}) = \exp(\boldsymbol{w}^{\top}\boldsymbol{x})/(1+\exp(\boldsymbol{w}^{\top}\boldsymbol{x}))$
- ullet The natural parameter  $\eta=\psi(\mu)$  where  $\psi$  is the link function
- Note: Some GLM can be represented as  $p(y|\eta,\phi) = h(y,\phi) \exp(\frac{\eta y A(\eta)}{\phi})$  where  $\phi$  is a dispersion parameter (Gaussian/gamma GLMs use this representation)

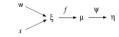


ullet A GLM has a canonical response function f if  $f=\psi^{-1}$ 



• For such a GLM,  $\eta = \psi(\mu) = \psi(f(\mathbf{w}^{\top}\mathbf{x})) = \mathbf{w}^{\top}\mathbf{x}$ 

• A GLM has a canonical response function f if  $f = \psi^{-1}$ 



- For such a GLM,  $\eta = \psi(\mu) = \psi(f(\boldsymbol{w}^{\top}\boldsymbol{x})) = \boldsymbol{w}^{\top}\boldsymbol{x}$
- $\bullet$  E.g., for logistic regression  $\eta = \log \frac{\mu}{1-\mu} = \mathbf{w}^{\top} \mathbf{x}_{\mathit{n}}$

• A GLM has a canonical response function f if  $f = \psi^{-1}$ 

$$\xi \xrightarrow{f} \mu \xrightarrow{\psi} \eta$$

- For such a GLM,  $\eta = \psi(\mu) = \psi(f(\boldsymbol{w}^{\top}\boldsymbol{x})) = \boldsymbol{w}^{\top}\boldsymbol{x}$
- E.g., for logistic regression  $\eta = \log \frac{\mu}{1-\mu} = {\bf w}^{\top} {\bf x}_n$
- Thus, for Canonical GLMs

$$p(y|\eta) = h(y) \exp(\eta y - A(\eta))$$
  
=  $h(y) \exp(y \mathbf{w}^{\top} \mathbf{x} - A(\eta))$ 

• A GLM has a canonical response function f if  $f = \psi^{-1}$ 

$$\xi \xrightarrow{f} \mu \xrightarrow{\psi} \eta$$

- For such a GLM,  $\eta = \psi(\mu) = \psi(f(\mathbf{w}^{\top}\mathbf{x})) = \mathbf{w}^{\top}\mathbf{x}$
- E.g., for logistic regression  $\eta = \log \frac{\mu}{1-\mu} = {\bf w}^{\top} {\bf x}_n$
- Thus, for Canonical GLMs

$$p(y|\eta) = h(y) \exp(\eta y - A(\eta))$$
  
=  $h(y) \exp(y \mathbf{w}^{\top} \mathbf{x} - A(\eta))$ 

• This form makes parameter estimation in canonical GLM easy (e.g., gradients easy to compute)

• A GLM has a canonical response function f if  $f = \psi^{-1}$ 

$$\xi \xrightarrow{f} \mu \xrightarrow{\psi} \eta$$

- For such a GLM,  $\eta = \psi(\mu) = \psi(f(\mathbf{w}^{\top}\mathbf{x})) = \mathbf{w}^{\top}\mathbf{x}$
- E.g., for logistic regression  $\eta = \log \frac{\mu}{1-\mu} = {\bf w}^{\top} {\bf x}_n$
- Thus, for Canonical GLMs

$$p(y|\eta) = h(y) \exp(\eta y - A(\eta))$$
  
=  $h(y) \exp(y \mathbf{w}^{\top} \mathbf{x} - A(\eta))$ 

- This form makes parameter estimation in canonical GLM easy (e.g., gradients easy to compute)
- We will focus on canonical GLMs only (these are the most common)



### MLE for GLM

Log likelihood

$$L(\eta) = \log p(Y|\eta) = \log \prod_{n=1}^{N} h(y_n) \exp(y_n \mathbf{w}^{\top} \mathbf{x}_n - A(\eta_n))$$

#### MLE for GLM

#### Log likelihood

$$L(\eta) = \log p(Y|\eta) = \log \prod_{n=1}^{N} h(y_n) \exp(y_n \mathbf{w}^{\top} \mathbf{x}_n - A(\eta_n)) = \sum_{n=1}^{N} \log h(y_n) + \mathbf{w}^{\top} \sum_{n=1}^{N} y_n \mathbf{x}_n - \sum_{n=1}^{N} A(\eta_n)$$

Log likelihood

$$L(\eta) = \log p(Y|\eta) = \log \prod_{n=1}^{N} h(y_n) \exp(y_n \mathbf{w}^{\top} \mathbf{x}_n - A(\eta_n)) = \sum_{n=1}^{N} \log h(y_n) + \mathbf{w}^{\top} \sum_{n=1}^{N} y_n \mathbf{x}_n - \sum_{n=1}^{N} A(\eta_n)$$

ullet Convexity of  $A(\eta)$  guarantees a global optima. Taking derivative w.r.t.  $oldsymbol{w}$ 

$$\sum_{n=1}^{N} \left( y_n \mathbf{x}_n - A'(\eta_n) \frac{d\eta_n}{d\mathbf{w}} \right)$$

Log likelihood

$$L(\eta) = \log p(Y|\eta) = \log \prod_{n=1}^{N} h(y_n) \exp(y_n \mathbf{w}^{\top} \mathbf{x}_n - A(\eta_n)) = \sum_{n=1}^{N} \log h(y_n) + \mathbf{w}^{\top} \sum_{n=1}^{N} y_n \mathbf{x}_n - \sum_{n=1}^{N} A(\eta_n)$$

• Convexity of  $A(\eta)$  guarantees a global optima. Taking derivative w.r.t.  $\boldsymbol{w}$ 

$$\sum_{n=1}^{N} \left( y_n \mathbf{x}_n - A'(\eta_n) \frac{d\eta_n}{d\mathbf{w}} \right) = \sum_{n=1}^{N} (y_n \mathbf{x}_n - \mu_n \mathbf{x}_n)$$

Log likelihood

$$L(\eta) = \log p(Y|\eta) = \log \prod_{n=1}^{N} h(y_n) \exp(y_n \mathbf{w}^{\top} \mathbf{x}_n - A(\eta_n)) = \sum_{n=1}^{N} \log h(y_n) + \mathbf{w}^{\top} \sum_{n=1}^{N} y_n \mathbf{x}_n - \sum_{n=1}^{N} A(\eta_n)$$

• Convexity of  $A(\eta)$  guarantees a global optima. Taking derivative w.r.t.  $\boldsymbol{w}$ 

$$\sum_{n=1}^{N} \left( y_n \mathbf{x}_n - A'(\eta_n) \frac{d\eta_n}{d\mathbf{w}} \right) = \sum_{n=1}^{N} (y_n \mathbf{x}_n - \mu_n \mathbf{x}_n) = \sum_{n=1}^{N} (y_n - \mu_n) \mathbf{x}_n$$

Log likelihood

$$L(\eta) = \log p(Y|\eta) = \log \prod_{n=1}^{N} h(y_n) \exp(y_n \mathbf{w}^{\top} \mathbf{x}_n - A(\eta_n)) = \sum_{n=1}^{N} \log h(y_n) + \mathbf{w}^{\top} \sum_{n=1}^{N} y_n \mathbf{x}_n - \sum_{n=1}^{N} A(\eta_n)$$

• Convexity of  $A(\eta)$  guarantees a global optima. Taking derivative w.r.t.  $\boldsymbol{w}$ 

$$\sum_{n=1}^{N} \left( y_n \mathbf{x}_n - A'(\eta_n) \frac{d\eta_n}{d\mathbf{w}} \right) = \sum_{n=1}^{N} (y_n \mathbf{x}_n - \mu_n \mathbf{x}_n) = \sum_{n=1}^{N} (y_n - \mu_n) \mathbf{x}_n$$

Log likelihood

$$L(\eta) = \log p(Y|\eta) = \log \prod_{n=1}^{N} h(y_n) \exp(y_n \mathbf{w}^{\top} \mathbf{x}_n - A(\eta_n)) = \sum_{n=1}^{N} \log h(y_n) + \mathbf{w}^{\top} \sum_{n=1}^{N} y_n \mathbf{x}_n - \sum_{n=1}^{N} A(\eta_n)$$

• Convexity of  $A(\eta)$  guarantees a global optima. Taking derivative w.r.t.  $\boldsymbol{w}$ 

$$\sum_{n=1}^{N} \left( y_n \mathbf{x}_n - A'(\eta_n) \frac{d\eta_n}{d\mathbf{w}} \right) = \sum_{n=1}^{N} (y_n \mathbf{x}_n - \mu_n \mathbf{x}_n) = \sum_{n=1}^{N} (y_n - \mu_n) \mathbf{x}_n$$

where  $\mu_n = f(\mathbf{w}^{\top} \mathbf{x}_n)$  and 'f'  $(= \psi^{-1})$  depends on type of response y, e.g.,

• Real-valued y (linear regression): f is identity, i.e.,  $\mu_n = \mathbf{w}^{\top} \mathbf{x}_n$ 

Log likelihood

$$L(\eta) = \log p(Y|\eta) = \log \prod_{n=1}^{N} h(y_n) \exp(y_n \mathbf{w}^{\top} \mathbf{x}_n - A(\eta_n)) = \sum_{n=1}^{N} \log h(y_n) + \mathbf{w}^{\top} \sum_{n=1}^{N} y_n \mathbf{x}_n - \sum_{n=1}^{N} A(\eta_n)$$

• Convexity of  $A(\eta)$  guarantees a global optima. Taking derivative w.r.t.  $\boldsymbol{w}$ 

$$\sum_{n=1}^{N} \left( y_n \mathbf{x}_n - A'(\eta_n) \frac{d\eta_n}{d\mathbf{w}} \right) = \sum_{n=1}^{N} (y_n \mathbf{x}_n - \mu_n \mathbf{x}_n) = \sum_{n=1}^{N} (y_n - \mu_n) \mathbf{x}_n$$

- Real-valued y (linear regression): f is identity, i.e.,  $\mu_n = \mathbf{w}^{\top} \mathbf{x}_n$
- Binary y (logistic regression): f is logistic function, i.e.,  $\mu_n = \frac{\exp(\mathbf{w}^{\top} \mathbf{x}_n)}{1 + \exp(\mathbf{w}^{\top} \mathbf{x}_n)}$

Log likelihood

$$L(\eta) = \log p(Y|\eta) = \log \prod_{n=1}^{N} h(y_n) \exp(y_n \mathbf{w}^{\top} \mathbf{x}_n - A(\eta_n)) = \sum_{n=1}^{N} \log h(y_n) + \mathbf{w}^{\top} \sum_{n=1}^{N} y_n \mathbf{x}_n - \sum_{n=1}^{N} A(\eta_n)$$

ullet Convexity of  $A(\eta)$  guarantees a global optima. Taking derivative w.r.t.  $oldsymbol{w}$ 

$$\sum_{n=1}^{N} \left( y_n \mathbf{x}_n - A'(\eta_n) \frac{d\eta_n}{d\mathbf{w}} \right) = \sum_{n=1}^{N} (y_n \mathbf{x}_n - \mu_n \mathbf{x}_n) = \sum_{n=1}^{N} (y_n - \mu_n) \mathbf{x}_n$$

- Real-valued y (linear regression): f is identity, i.e.,  $\mu_n = \mathbf{w}^{\top} \mathbf{x}_n$
- Binary y (logistic regression): f is logistic function, i.e.,  $\mu_n = \frac{\exp(\mathbf{w}^{\top} \mathbf{x}_n)}{1 + \exp(\mathbf{w}^{\top} \mathbf{x}_n)}$
- Count-valued y (Poisson regression):  $\mu_n = \exp(\mathbf{w}^\top \mathbf{x}_n)$

Log likelihood

$$L(\eta) = \log p(Y|\eta) = \log \prod_{n=1}^{N} h(y_n) \exp(y_n \mathbf{w}^{\top} \mathbf{x}_n - A(\eta_n)) = \sum_{n=1}^{N} \log h(y_n) + \mathbf{w}^{\top} \sum_{n=1}^{N} y_n \mathbf{x}_n - \sum_{n=1}^{N} A(\eta_n)$$

ullet Convexity of  $A(\eta)$  guarantees a global optima. Taking derivative w.r.t.  $oldsymbol{w}$ 

$$\sum_{n=1}^{N} \left( y_n \mathbf{x}_n - A'(\eta_n) \frac{d\eta_n}{d\mathbf{w}} \right) = \sum_{n=1}^{N} (y_n \mathbf{x}_n - \mu_n \mathbf{x}_n) = \sum_{n=1}^{N} (y_n - \mu_n) \mathbf{x}_n$$

- Real-valued y (linear regression): f is identity, i.e.,  $\mu_n = \mathbf{w}^{\top} \mathbf{x}_n$
- Binary y (logistic regression): f is logistic function, i.e.,  $\mu_n = \frac{\exp(\mathbf{w}^{\top} \mathbf{x}_n)}{1 + \exp(\mathbf{w}^{\top} \mathbf{x}_n)}$
- Count-valued y (Poisson regression):  $\mu_n = \exp(\mathbf{w}^\top \mathbf{x}_n)$
- Positive reals y (gamma regression):  $\mu_n = -(\mathbf{w}^{\top} \mathbf{x}_n)^{-1}$

Log likelihood

$$L(\eta) = \log p(Y|\eta) = \log \prod_{n=1}^{N} h(y_n) \exp(y_n \mathbf{w}^{\top} \mathbf{x}_n - A(\eta_n)) = \sum_{n=1}^{N} \log h(y_n) + \mathbf{w}^{\top} \sum_{n=1}^{N} y_n \mathbf{x}_n - \sum_{n=1}^{N} A(\eta_n)$$

• Convexity of  $A(\eta)$  guarantees a global optima. Taking derivative w.r.t.  $\boldsymbol{w}$ 

$$\sum_{n=1}^{N} \left( y_n \mathbf{x}_n - A'(\eta_n) \frac{d\eta_n}{d\mathbf{w}} \right) = \sum_{n=1}^{N} (y_n \mathbf{x}_n - \mu_n \mathbf{x}_n) = \sum_{n=1}^{N} (y_n - \mu_n) \mathbf{x}_n$$

- Real-valued y (linear regression): f is identity, i.e.,  $\mu_n = \mathbf{w}^{\top} \mathbf{x}_n$
- Binary y (logistic regression): f is logistic function, i.e.,  $\mu_n = \frac{\exp(\mathbf{w}^{\top} \mathbf{x}_n)}{1 + \exp(\mathbf{w}^{\top} \mathbf{x}_n)}$
- Count-valued y (Poisson regression):  $\mu_n = \exp(\mathbf{w}^\top \mathbf{x}_n)$
- Positive reals y (gamma regression):  $\mu_n = -(\mathbf{w}^{\top} \mathbf{x}_n)^{-1}$
- To estimate w via MLE (or MAP), either set the derivative to zero or use iterative methods (e.g., gradient descent, iteratively reweighted least squares, etc.)

• If likelihood is conjugate to the prior on  $\boldsymbol{w}$ , Bayesian inference can be done in closed form

- If likelihood is conjugate to the prior on  $\boldsymbol{w}$ , Bayesian inference can be done in closed form
  - ullet Example: Bayesian linear regression with Gaussian likelihood and Gaussian prior on  $oldsymbol{w}$

- If likelihood is conjugate to the prior on  $\boldsymbol{w}$ , Bayesian inference can be done in closed form
  - Example: Bayesian linear regression with Gaussian likelihood and Gaussian prior on w
- Otherwise, approximate Bayesian inference is needed (e.g., Laplace, MCMC, variational inf, etc.)

- If likelihood is conjugate to the prior on w, Bayesian inference can be done in closed form
  - Example: Bayesian linear regression with Gaussian likelihood and Gaussian prior on w
- Otherwise, approximate Bayesian inference is needed (e.g., Laplace, MCMC, variational inf, etc.)
  - ullet Example: Bayesian logistic regression with sigmoid-Bernoulli likelihood and Gaussian prior on ullet

- If likelihood is conjugate to the prior on w, Bayesian inference can be done in closed form
  - ullet Example: Bayesian linear regression with Gaussian likelihood and Gaussian prior on ullet
- Otherwise, approximate Bayesian inference is needed (e.g., Laplace, MCMC, variational inf, etc.)
  - ullet Example: Bayesian logistic regression with sigmoid-Bernoulli likelihood and Gaussian prior on ullet
- Interesting class project idea: Design simple inference algorithms for non-conjugate GLM

• Exp. family distributions are very useful for modeling diverse types of data/parameters

- Exp. family distributions are very useful for modeling diverse types of data/parameters
- Conjugate priors to exp. family distributions make parameter updates very simple

- Exp. family distributions are very useful for modeling diverse types of data/parameters
- Conjugate priors to exp. family distributions make parameter updates very simple
- Other quantities such as posterior predictive can be computed in closed form

- Exp. family distributions are very useful for modeling diverse types of data/parameters
- Conjugate priors to exp. family distributions make parameter updates very simple
- Other quantities such as posterior predictive can be computed in closed form
- Useful in designing generative classification models.

- Exp. family distributions are very useful for modeling diverse types of data/parameters
- Conjugate priors to exp. family distributions make parameter updates very simple
- Other quantities such as posterior predictive can be computed in closed form
- Useful in designing generative classification models. Choosing class-conditional from exponential family with conjugate priors helps in parameter estimation

- Exp. family distributions are very useful for modeling diverse types of data/parameters
- Conjugate priors to exp. family distributions make parameter updates very simple
- Other quantities such as posterior predictive can be computed in closed form
- Useful in designing generative classification models. Choosing class-conditional from exponential family with conjugate priors helps in parameter estimation
- Useful in designing generative models for unsupervised learning

- Exp. family distributions are very useful for modeling diverse types of data/parameters
- Conjugate priors to exp. family distributions make parameter updates very simple
- Other quantities such as posterior predictive can be computed in closed form
- Useful in designing generative classification models. Choosing class-conditional from exponential family with conjugate priors helps in parameter estimation
- Useful in designing generative models for unsupervised learning
- Also used in designing GLMs

- Exp. family distributions are very useful for modeling diverse types of data/parameters
- Conjugate priors to exp. family distributions make parameter updates very simple
- Other quantities such as posterior predictive can be computed in closed form
- Useful in designing generative classification models. Choosing class-conditional from exponential family with conjugate priors helps in parameter estimation
- Useful in designing generative models for unsupervised learning
- Also used in designing GLMs
- We will see several use cases when we discuss approximate inference algoritms (e.g., Gibbs sampling, and especially variational inference)