

Prob1(a) Since  $N(t)$  is a Poisson process with rate  $\lambda > 0$ ,  $E[N(t)] = \lambda t$ .

$$\text{Thus, } E[M(t)] = E[N(t) - \lambda t] = E[N(t)] - \lambda t = \lambda t - \lambda t = 0$$

For  $0 \leq s \leq t < \infty$ ,

$$\text{Cov}(M(t), M(s)) = E[(M(t) - E(M(t)))(M(s) - E(M(s)))]$$

$$= E[M(t) \cdot M(s)]$$

$$= E[(N(t) - \lambda t)(N(s) - \lambda s)]$$

$$= E[N(t)N(s)] - \lambda t E[N(s)] - \lambda s E[N(t)] + \lambda^2 st$$

since  $E[N(s)] = \lambda s$ ,  $E[N(t)] = \lambda t$

$$= E[N(t)N(s)] - \lambda^2 ts$$

$$\text{since } E[N(t)N(s)] = E[(N(t) - N(s))(N(s) - N(0))] + E[N^2(s)], N(0) = 0,$$

and  $(0, s)$ ,  $(s, t)$  are disjoint and therefore independent

$$E[N(t)N(s)] = E[N(t) - N(s)] \cdot E[N(s) - N(0)] + E[N^2(s)]$$

$$= \lambda^2(t-s)s + E[N^2(s)] = \lambda^2 t s - \lambda^2 s^2 + E[N^2(s)]$$

since  $E[X^2] = \text{Var}(X) + E^2(X)$ ,

$$\text{Cov}(M(t), M(s)) = \lambda^2 t s - \lambda^2 s^2 + E(N^2(s)) - \lambda^2 t s$$

$$= -\lambda^2 s^2 + E(N^2(s))$$

$$= \lambda^2 s^2 + \text{Var}(N(s)) + [E(N(s))]^2$$

$$= -\lambda^2 s^2 + s\lambda + (\lambda s)^2 = \lambda s, \text{ where } 0 \leq s \leq t < \infty$$

Thus, in general,  $\text{Cov}(N(s), N(t)) = \lambda \min\{s, t\}$ .

(b) Wiener Process.

For a Wiener Process, we have  $E[W(t)] = 0$ ,  $\text{Cov}_w(s, t) = \sigma^2 \min\{s, t\}$ .

if we let  $\sigma^2 = \lambda$ , then we have a Wiener Process which is equivalent to  $M(t)$

(c).  $M(t)$  is not a stationary process.

since the covariance function for  $M(t)$ ,  $\text{Cov}(M(s), M(t)) = \lambda \min\{s, t\}$ , depends on  $s$  or  $t$ , not on  $|s-t|$ .

(d) since  $M(1) = -\lambda$ ,  $N(1) = 0$ .

Thus, the MMSE estimate of  $T_1$  is that  $E(T_1 | N(1) = 0)$

$$\text{Since } \{N(1) = 0\} = \{T_1 > 1\}, E(T_1 | T_1 > 1) = 1 + E(T_1) \text{ by memorylessness}$$

$$= 1 + \frac{1}{\lambda} \quad T_1 \sim e^\lambda.$$

Thus,  $E(T_1 | M(1) = -\lambda) = 1 + \frac{1}{\lambda}$ .

$$(e) M(3) = 1 - 3\lambda \Rightarrow N(3) = 1$$

$$M(1) = -\lambda \Rightarrow N(1) = 0$$

$$E(N_6 | N_3, N_1) = E(N_6 | N_3 = 1) = E(N_3 | N_3 = 1) + E(N_6 - N_3 | N_3 = 1) = N_3 + E(N_6 - N_3).$$

since  $N_6 - N_3 \sim \text{Poisson}(\lambda(6-3) = 3\lambda)$ ,  $E[N_6 - N_3] = \lambda(3) = 3\lambda$ .

Thus,  $E(N_6 | N_3, N_1) = N_3 + E(N_6 - N_3) = 1 + 3\lambda$ .

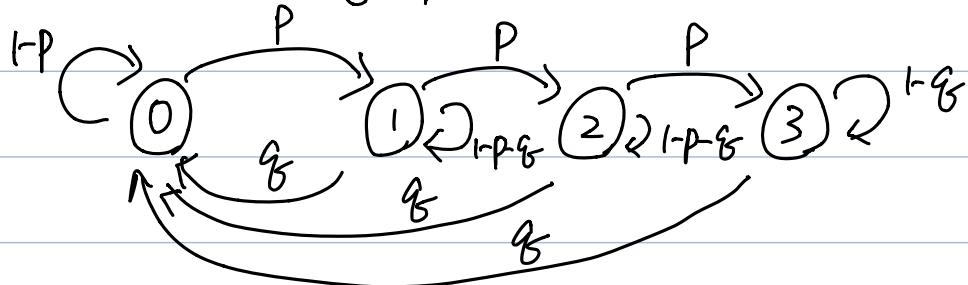
$$\therefore E(M_6 | M_3, M_1) = E(N_6 | N_3 - N_1) - 6\lambda = 1 + 3\lambda - 6\lambda = 1 - 3\lambda.$$

Prob 2. (a). when  $0 < x_i < K$ ,  $x_{i+1} = \begin{cases} x_i + 1 & \text{with probability } p \\ 0 & \text{with probability } q \\ x_i & \text{with probability } 1-p-q \end{cases}$

when  $x_i = 0$ ,  $x_{i+1} = \begin{cases} x_i + 1 & \text{with probability } p \\ 0 & \text{with probability } 1-p \end{cases}$

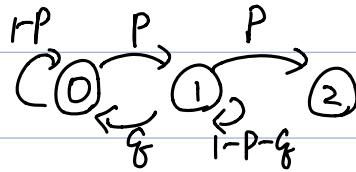
when  $x_i = K$ ,  $x_{i+1} = \begin{cases} K & \text{with probability } 1-q \\ 0 & \text{with probability } q \end{cases}$

(b). transition graph:



(c). Let  $T(i)$  be the expected time to first arrive at state 2, starting from  $i$ . Then, we will have

$$T(2) = 0$$



$$T(1) = 1 + pT(2) + (1-p-q)T(1) + qT(0)$$

$$T(0) = 1 + (1-p)T(0) + pT(1)$$

$$\text{Thus } T(0) = \frac{p+q}{p^2} + \frac{1}{p} \quad (k \geq 2)$$

$$(d) K \neq \infty, \quad p+q \leq 1, \quad p > 0, \quad q > 0.$$

$$P = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \\ \frac{1}{2} & \frac{1}{2} \end{bmatrix} \Rightarrow \pi = \begin{bmatrix} \frac{1}{2} & \frac{1}{2} \end{bmatrix}$$

$$(e) \pi \cdot P = \pi \quad \sum \pi_i = 1$$

$$P = \begin{bmatrix} 1-p & p & 0 & \dots & 0 \\ q & 1-p-q & p & \dots & 0 \\ \vdots & \ddots & \ddots & \ddots & \vdots \\ q & 0 & \dots & \dots & 1-q \end{bmatrix}$$

$$\text{Thus } T_{10} = \frac{q}{p+q}$$

$$T_{1i} = \frac{p^{i-1}q}{(p+q)} \quad i < k$$

$$\pi_k = \frac{p^k}{(p+q)^k}$$

(f). No.

the state of  $X_{n+1}$  depends on the state of  $X_n, X_{n-1}$  and  $X_{n-2}$ .

Prob 3. (a).  $\{X(t)\}$  is a zero-mean stationary process.

$$\text{Thus } K_X(s, t) = E[X(s)X(t)] = E[X(0)X(t-s)] = k_X(t-s).$$

$$\text{let } \tau = t-s, \text{ then } k_X(\tau) = E[X(t)X(t+\tau)]$$

From Cauchy-Schwarz inequality, if  $k_X(\tau) = k_X(0)$ , for some  $\tau$ ,  
then  $[E(X(t+\tau) - X(t)X(0))]^2 = [k_X(\tau+t) - k_X(t)]^2$

$$\begin{aligned} &\leq E[(X(t+\tau) - X(t))^2] E[(X(0))^2] \\ &= (E[(X(t+\tau))^2] + E[(X(t))^2] - 2E[(X(t+\tau))(X(t))]) \\ &= (2k_X(0) - 2k_X(\tau)) k_X(0) = 0 \end{aligned}$$

Thus  $k_X(t+\tau) = k_X(t)$ , for some  $\tau$

thus  $k_X(\tau)$  must be periodic with period  $\tau$ .

$$(b). E[(X(t+\tau) - X(t))^2]$$

$$\begin{aligned} &= E[(X(t+\tau))^2] + E[(X(t))^2] - 2E[X(t+\tau)X(t)] \\ &= 2k_X(0) - 2k_X(\tau) = 0. \end{aligned}$$

(c) Yes.

$$(d). E[X|Y] = E[X] + K_{XY} K_Y^{-1} (Y - E(Y))$$

$$\text{Thus } E[X(S+2\tau) | X(S)=x]$$

$$= E[X(S+2\tau)] + K_{XY} K_Y^{-1} (Y - E(Y)), \text{ where } Y = X(S) = x$$

$$E[X(S+2\tau)] = 0,$$

$$K_{XY} = K_X(S+2\tau, S) = k_X(2\tau) = k_X(0)$$

$$K_Y = K_X(S, S) = k_X(0)$$

$$\therefore E[X(S+2\tau) | X(S)=x] = 0 + k_X(0) \cdot K_Y^{-1}(x - 0) = \frac{k_X(0)}{K_Y} x = x$$

$$E[X(S+2\tau) - \hat{X}] = E[X(S+2\tau)] - E[\hat{X}]$$

$$= 0 - E[E[X(S+2\tau) | X(S)=x]] = 0.$$

$$\text{Var}[X(S+2\tau) - \hat{X}] = [K_X] - [K_{XY} \cdot K_Y^{-1} K_{XY}^T]$$

$$= k_X(0) - \frac{k_X^2(0)}{K_Y} = 0.$$

- $K_{x(10)}$
- if Covariance is non-singular  
 $\begin{bmatrix} K_x & K_{x,y} \\ K_{y,x} & K_y \end{bmatrix}$  their joint pdf doesn't exist.
- (i).  $K_x(t, t+2\tau) = K_x(2\tau) = K_x(0)$  not exist. (non-singular)
- (ii).  $K_x(t, t+\tau) = K_x(\tau) \neq K_x(0)$  exist.
- (iii).  $K_x(t+\tau, t+\tau) = K_x(\tau) = K_x(0)$  exist.  
 $K_x(t, t+\tau) = K_x(\tau), K_x(t+\tau, t+\tau) = K_x(\tau)$
- (iv).  $K_x(t, t+\tau) = K_x(0)$  not exist. (non-singular).

Prob4. (a). Yes.

$$\text{let } \delta = \frac{t+\alpha-t}{n} = \frac{\alpha}{n}$$

$$\begin{aligned} \text{since } X(t+\alpha) - X(t) &= [X(t+\alpha) - X(t+\alpha-\delta)] + [X(t+\alpha-\delta) - X(t+\alpha-2\delta)] \\ &\quad + \dots + [X(t+\alpha-(n-1)\delta) - X(t)] \\ &= \sum_{i=1}^n Y_i. \end{aligned}$$

when  $n \rightarrow \infty$ ,  $\delta \rightarrow 0$ , then  $Y_i = X[s - (i-1)\delta] - X[s - i\delta] \rightarrow 0$ .

Thus, according to CLT,  $X(t+\alpha) - X(t)$  follows gaussian distribution  
 $\{X(t)\}$  is a gaussian process.

(b). Yes.

$$\begin{aligned} \text{let } 0 < t_1 < t_2, X(t_2) - X(t_1) &= [W(t_2 + \alpha) - W(t_2)] - [W(t_1 + \alpha) - W(t_1)] \\ &= [W(t_2 + \alpha) - W(t_1 + \alpha)] - [W(t_2) - W(t_1)] \end{aligned}$$

Since Wiener process satisfies that the distribution of  $W(t_2) - W(t_1)$  is only related to  $|t_2 - t_1|$ ,  $X(t_2) - X(t_1)$  is only related to  $t_2 - t_1$ . Thus,  $X(t)$  is a stationary process.

$$(c). X(10) = W(10+\alpha) - W(10)$$

$$P(X(10) > 0) = P[W(10+\alpha) - W(10) > 0] = \frac{1}{2}$$

$$(d). X(n) \sim N(0, \alpha)$$

$$\therefore \frac{1}{N} \sum_{n=1}^N X_n \sim N(0, \frac{\alpha}{N^2}) = N(0, \frac{\alpha}{N})$$

Thus,  $N \rightarrow \infty$ ,  $\frac{\alpha}{N} \rightarrow 0$ .  $\therefore$  converge to 0 in probability.

(f).  $E[B(t)] = 0$

$$\begin{aligned} \text{if } s < t, \text{ Cov}[B(s), B(t)] &= \text{Cov}(W(s) - sW(1), W(t) - tW(1)) \\ &= \text{Cov}(W(s), W(t)) - t \text{Cov}(W(s), W(1)) \\ &\quad - s \text{Cov}(W(1), W(t)) + st \text{Var}(W(1)) \\ &= s - ts - st + st \\ &= s(1-t) \end{aligned}$$

$$\text{in which case, } \text{Var}(B(t)) = t(1-t)$$

$$\text{In general, } \text{Cov}[B(s), B(t)] = \min(s, t) - st.$$

(e) Yes.  $B(t) = W(t) - tW(1) \quad 0 \leq t \leq 1$  It is the linear combination of Gaussian.

(g). For  $0 \leq t \leq 1$

$$\begin{aligned} \text{Cov}(W(t) - tW(1), W(1)) &= E[(W(t) - tW(1))W(1)] \\ &= \min\{t, 1\} - t = 0 \end{aligned}$$

As their covariance is zero, and their joint distribution Gaussian,  $B(t) = W(t) - tW(1)$  and  $W(1)$  are independent.

[Solution 2] (a)

Since  $W(t)$  is a Wiener process,  $X(t) = W(t+\alpha) - W(t) \sim N(0, \alpha)$ .

for arbitrary chosen  $t_1, \dots, t_n$ , we have  $X(t_i) = W(t_i + \alpha) - W(t_i) \dots$

$X(t_n) = W(t_n + \alpha) - W(t_n)$ . Thus,  $(X(t_1), \dots, X(t_n))$  is the linear combination of  $(W(t_1), \dots, W(t_n), W(t_1 + \alpha), \dots, W(t_n + \alpha))$

Since Wiener process is a Gaussian Process,  $(W(t_1), \dots, W(t_n), W(t_1 + \alpha), \dots, W(t_n + \alpha))$

are  $2n$   $\vec{r}v$ . Thus,  $(X(t_1), \dots, X(t_n))$  is  $n \vec{w}$ ,

which means  $\{X(t_1), \dots, X(t_n)\}$  are joint Gaussian for arbitrary of

$t_1, \dots, t_n$  and  $n$ . Thus,  $\{X(t)\}$  is a Gaussian process.

[Solution 2] (b)

$$K_X(s, t) = E[W(s+\alpha) - W(s)] [W(t+\alpha) - W(t)].$$
$$= (s+\alpha)6^2 - \min(s+\alpha, t)6^2.$$

if  $s+\alpha \leq t$ ,  $K_X(s, t) = 0$ .

if  $s+\alpha > t$ ,  $K_X(s, t) = (s+\alpha - t)6^2 = s-t+\alpha$ .

Thus,  $K_X(s, t)$  is either 0 or depends on  $s-t$ .

$$E[X(t)] = E[W(t+\alpha)] - E[W(t)] = 0.$$

Thus,  $X(t)$  is a wide sense stationary for a Gaussian process, which means it is a stationary process.