CSE 515T (Spring 2015) Assignment 2

Due Monday, 16 February 2015

1. (Curse of dimensionality.) Consider a d-dimensional, zero-mean, spherical multivariate Gaussian distribution:

$$p(\mathbf{x}) = \mathcal{N}(\mathbf{x}; \mathbf{0}, \mathbf{I}_d).$$

Equivalently, each entry of x is drawn iid from a univariate standard normal distribution.

In familiar small dimensions ($d \leq 3$), "most" of the vectors drawn from a multivariate Gaussian distribution will lie near the mean. For example, the famous 68–95–99.7 rule for d=1 indicates that large deviations from the mean are unusual. Here we will consider the behavior in larger dimensions.

- Draw 10 000 samples from $p(\mathbf{x})$ for each dimension in $d \in \{1, 5, 10, 50, 100\}$, and compute the length of each vector drawn: $y_d = \sqrt{\mathbf{x}^\top \mathbf{x}} = (\sum_i^d x_i^2)^{1/2}$. Estimate the distribution of each y_d using either a histogram or a kernel density estimate (in MATLAB, hist and ksdensity, respectively). Plot your estimates. (Please do not hand in your raw samples!) Summarize the behavior of this distribution as d increases.
- The true distribution of y_d^2 is a chi-square distribution with d degrees of freedom (the distribution of y_d itself is the less-commonly seen chi distribution). Use this fact to compute the probability that $y_d < 5$ for each of the dimensions in the last part.
- For $d=1\,000$, compute the 5th and 95th percentiles of y_d . Is the mean $\mathbf{x}=\mathbf{0}$ a representative summary of the distribution in high dimensions? This behavior has been called "the curse of dimensionality."
- 2. (Bayesian linear regression.) Consider the following data:

$$\mathbf{x} = [-2.26, -1.31, -0.43, 0.32, 0.34, 0.54, 0.86, 1.83, 2.77, 3.58]^{\top};$$

$$\mathbf{y} = [1.03, 0.70, -0.68, -1.36, -1.74, -1.01, 0.24, 1.55, 1.68, 1.53]^{\top}.$$

Fix the noise variance at $\sigma^2 = 0.5^2$.

• Perform Bayesian linear regression for these data using the polynomial basis functions $\phi_k(x) = [1, x, x^2, \dots x^k]^\top$ for $k \in \{1, 2, 3\}$, in each case using the parameter prior $p(\mathbf{w}) = \mathcal{N}(\mathbf{w}; \mathbf{0}, \mathbf{I})$. Evaluate and plot the posterior means $\mathbb{E}[\mathbf{y}_* \mid \mathbf{X}_*, \mathcal{D}, \sigma^2]$ on the interval $x_* \in [-4, 4]$ for each model. Also plot the posterior mean plus-or-minus two times the posterior standard deviation:

$$\mathbb{E}[\mathbf{y}_* \mid \mathbf{X}_*, \mathcal{D}, \sigma^2] \pm 2\sqrt{\mathrm{var}[\mathbf{y}_* \mid \mathbf{X}_*, \mathcal{D}, \sigma^2]}.$$

This is a pointwise 95% credible interval for the regression function. Where is the pointwise uncertainty the largest?

- Compute the marginal likelihood of the data for each of the basis expansions above: $p(\mathbf{y} \mid \mathbf{X}, k, \sigma^2)$. Which model explains the data the best?
- 3. (Optimal design for Bayesian linear regression.) Consider the data from the last problem, and suppose we have selected the quadratic model corresponding to k=2 (do not assume that this is the answer to the last part of the last question). Imagine we are allowed to evaluate

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the function at a point x' of our choosing, giving a new dataset $\mathcal{D}' = \mathcal{D} \cup \left\{ (x', y') \right\}$ and a new posterior for the parameters $p(\mathbf{w} \mid \mathcal{D}', \sigma^2) = \mathcal{N}(\mathbf{w}; \boldsymbol{\mu}_{\mathbf{w} \mid \mathcal{D}'}, \boldsymbol{\Sigma}_{\mathbf{w} \mid \mathcal{D}'})$. We hope to select the location x' to best improve our current model, under some quality measure.

Assume that we ultimately wish to predict the function at a grid of points

$$\mathbf{x}_* = [-4, -3.5, -3, \dots, 3.5, 4]^{\top}.$$

We select the squared loss for a set of predictions $\hat{\mathbf{y}}_*$ at these points:

$$\ell(\mathbf{y}_*, \hat{\mathbf{y}}_*) = \sum_{i} ((y_*)_i - (\hat{y}_*)_i)^2;$$

therefore, we will predict using the new posterior mean $\hat{\mathbf{y}}_* = \mathbf{X}_* \boldsymbol{\mu}_{\mathbf{w}|\mathcal{D}'}$.

- Given a potential observation location x', derive a closed-form expression for the expected loss $\mathbb{E}[\ell(\mathbf{y}_*, \hat{\mathbf{y}}_*) \mid x', \mathcal{D}]$. Note: this does not require integration over y'! (What is the expected squared deviation from the mean?)
- Plot the expected loss over the interval $x' \in [-4, 4]$. Where is the optimal location to sample the function?

Note: this approach of actively selecting where to sample a function to maximize some utility function is known as *active learning* in machine learning and *optimal experimental design* in statistics. Bayesian decision theory provides a convenient and consistent framework for performing active learning with a variety of objectives.

4. (Woodbury matrix identity.) The *Woodbury matrix identity* is a very useful result. Let **A** be an $(n \times n)$ matrix, let **U** and **V** be $(n \times k)$ matrices, and let **C** be a $(k \times k)$ matrix. Then:

$$(\mathbf{A} + \mathbf{U}\mathbf{C}\mathbf{V}^{\top})^{-1} = \mathbf{A}^{-1} - \mathbf{A}^{-1}\mathbf{U}(\mathbf{C}^{-1} + \mathbf{V}^{\top}\mathbf{A}^{-1}\mathbf{U})^{-1}\mathbf{V}^{\top}\mathbf{A}^{-1}.$$

This result is useful when you already have the inverse of a matrix **A** and want to know the inverse after a rank-k adjustment. When $k \ll n$, the Woodbry matrix identity can be considerably faster than direct inversion!

- Prove this result.
- Use this result to rewrite the posterior covariance of the weight vector **w** in Bayesian linear regression (as written in the notes to lecture 5) in a simpler form.
- $_{\rm 5.}$ (Laplace approximation.) Find a Laplace approximation to the Gamma distribution:

$$p(\theta \mid \alpha, \beta) = \frac{1}{Z} \theta^{\alpha - 1} \exp(-\beta \theta).$$

Plot the approximation against the true density for $(\alpha, \beta) = (2, 1/2)$.

The true value of the normalizing constant is

$$Z = \frac{\Gamma(\alpha)}{\beta^{\alpha}}.$$

If we fix $\beta = 1$, then $Z = \Gamma(\alpha)$, so we may use the Laplace approximation to estimate the Gamma function. Analyze the quality of this approximation as a function of α .