

σ -Subgaussian random variables

$$M_X(\lambda) = \mathbb{E}[e^{\lambda x}] \leq e^{\lambda^2 \sigma^2 / 2}$$

if X is σ -subgaussian

(i) $\mathbb{E}[X] = 0$

(ii) $\text{Var}[X] \leq \sigma^2$

(iii) cX is $|c|\sigma$ -subgaussian

(iv) if X_1 and X_2 are independent and

subgaussian with σ_1 and σ_2 , then

$X_1 + X_2$ is $\sqrt{\sigma_1^2 + \sigma_2^2}$ -subgaussian

Theorem: $\mathbb{P}(|X| \geq t) \leq 2 e^{-\left(\frac{t^2}{2\sigma^2}\right)} \quad \text{--- (1)}$

III

with probability greater than $1-\delta$, the

random variable takes values in the

(confidence) interval $(-\sqrt{2\sigma^2 \log(2/\delta)}, +\sqrt{2\sigma^2 \log(2/\delta)})$

(let $\delta = 2 e^{-\left(\frac{\epsilon^2}{2\sigma^2}\right)}$ in ①)

We need $IP(|\text{sample Mean} - \text{True Mean}| \geq \epsilon)$

$$\hat{\mu} = \frac{\sum_{i=1}^n x_i}{n}$$

Sample
Mean

$$, \quad IP(\hat{\mu} - \mu \geq \epsilon) ?$$

$$\mathbb{P}(\hat{\mu} - \mu \geq \varepsilon) = \mathbb{P}\left(\frac{S_n}{n} - \mu \geq \varepsilon\right)$$

$$= \mathbb{P}\left(\frac{S_n - n\mu}{n} \geq \varepsilon\right)$$

$$\frac{S_n - n\mu}{n} = \frac{(X_1 - \mu) + (X_2 - \mu) + \dots + (X_n - \mu)}{n}$$

$$= \underbrace{\frac{(X_1 - \mu)}{n}}_{\frac{\sigma}{n}} + \underbrace{\frac{(X_2 - \mu)}{n}}_{\frac{\sigma}{n}} + \dots + \underbrace{\frac{(X_n - \mu)}{n}}_{\frac{\sigma}{n}}$$

if $(X_1 - \mu)$ is σ -subgaussian, then

$\frac{X_1 - \mu}{n}$ is $\frac{\sigma}{n}$ -subgaussian

$$\frac{S_n - n\mu}{n} \text{ is } \sqrt{\left(\frac{\sigma}{n}\right)^2 + \dots + \left(\frac{\sigma}{n}\right)^2}$$

n

$$\underbrace{\sqrt{\frac{\sigma^2}{n^2}}}_{n \text{ times}}$$

$$\sqrt{n \frac{\sigma^2}{n^2}} = \frac{\sigma}{\sqrt{n}}$$

$\frac{S_n - n\mu}{n}$ is $\frac{\sigma}{\sqrt{n}}$ -sub gaussian

$$\begin{aligned} \mathbb{P}(\hat{\mu} - \mu \geq \varepsilon) &= \mathbb{P}\left(\frac{S_n - n\mu}{n} \geq \varepsilon\right) \\ &\leq e^{-\left(\frac{\varepsilon^2}{2(\sigma/\sqrt{n})^2}\right)} \\ &= e^{-\left(\frac{n\varepsilon^2}{2\sigma^2}\right)} \end{aligned}$$

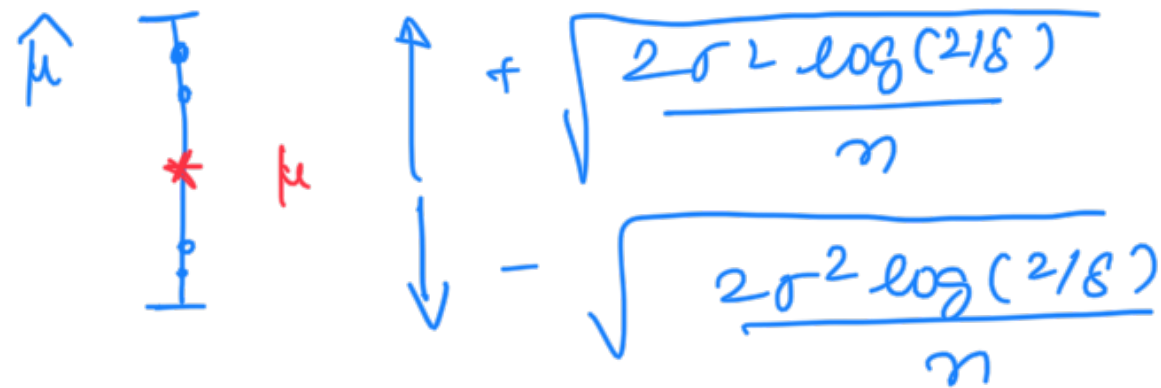
$$\mathbb{P}(|\hat{\mu} - \mu| \geq \varepsilon) \leq 2 e^{-\left(\frac{n\varepsilon^2}{2\sigma^2}\right)}$$

III

with probability greater than $1 - \delta$

$$\hat{\mu} \in \left[\mu - \sqrt{\frac{2\sigma^2 \log(2/\delta)}{n}}, \mu + \sqrt{\frac{2\sigma^2 \log(2/\delta)}{n}} \right]$$

Confidence interval



our next aim: $x \in [a, b]$ then x is

subgaussian

Lemma II: $\text{Var}[X] = \min_c \mathbb{E}[(X-c)^2]$

$$\begin{aligned}\mathbb{E}[(X-c)^2] &= \mathbb{E}[X^2] + \mathbb{E}[c^2] - 2\mathbb{E}[X \cdot c] \\ &= \mathbb{E}[X^2] + c^2 - 2c \mathbb{E}[X]\end{aligned}$$

differentiate w.r.t c

$$2c_* - 2\mathbb{E}[X] = 0 \quad \Rightarrow \quad c_* = \mathbb{E}[X]$$

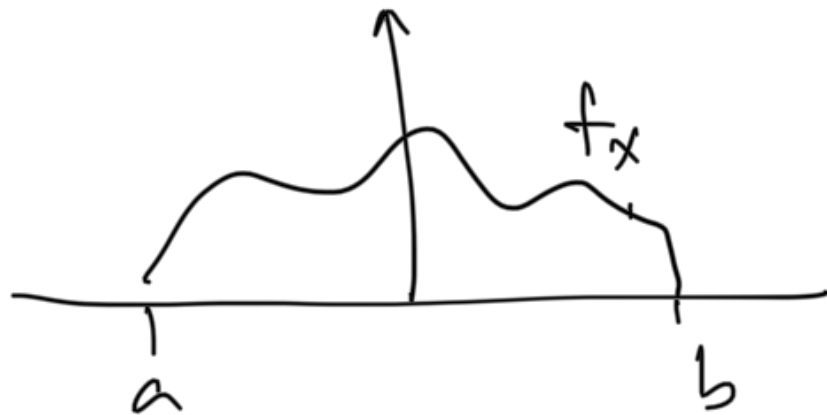
$$\text{Var}[X] = \mathbb{E}[(X - \mathbb{E}[X])^2]$$

Lemma I: $X \in [a, b]$ we have

$$\text{Var}[X] \leq \frac{(b-a)^2}{4}$$

Proof: $\text{Var}[X] = \min_c \mathbb{E}[(X-c)^2]$

(put $c = \frac{a+b}{2}$) $\leq \mathbb{E}\left[\left(X - \frac{b+a}{2}\right)^2\right]$



$$\begin{aligned} & \int_a^b \left(x - \frac{b+a}{2}\right)^2 f_X(x) dx \\ & \leq \int_a^b \max(e_1, e_2) f_X(x) dx \\ & = \max(e_1, e_2) \int_a^b f_X(x) dx \end{aligned}$$

extreme values of $\left(x - \frac{b+a}{2}\right)^2$ is

either $e_1 = \left(a - \frac{b+a}{2}\right)^2$ $e_2 = \left(b - \frac{b+a}{2}\right)^2$

$$\text{Var}[X] \leq \max \left\{ \left(a - \frac{b+a}{2}\right)^2, \left(b - \frac{b+a}{2}\right)^2 \right\}$$

$$= \frac{(b-a)^2}{4}$$

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Our next aim: $X \in [a, b]$ then X is

subgaussian (X is mean 0)

Proof: $M_X(\lambda) = \mathbb{E}[e^{\lambda X}]$

$$\psi(\lambda) = \log M_X(\lambda)$$

$$= \log \mathbb{E}[e^{\lambda x}]$$

$$\gamma(0) = \log \mathbb{E}[e^{0x}] = \log \mathbb{E}[1] = \log 1 = 0$$

$$\gamma'(\lambda) = \frac{\mathbb{E}[x e^{\lambda x}]}{\mathbb{E}[e^{\lambda x}]}, \quad \gamma'(0) = \frac{\mathbb{E}[x e^0]}{\mathbb{E}[e^0]} = \mathbb{E}[x] = 0$$

look at $\gamma'(\lambda) = \frac{1}{\mathbb{E}[e^{\lambda x}]} \int_{-\infty}^{\infty} x e^{\lambda x} f(x) dx$

forget the $\frac{1}{\mathbb{E}[e^{\lambda x}]}$ factor for a moment

$$\int_{-\infty}^{\infty} x e^{\lambda x} f(x) dx = \int_{-\infty}^{\infty} x \cdot g_{\lambda}(x) dx$$

$$= \mathbb{E}_f[x]$$

g_λ is a different density

$$g_\lambda(x) = M \cdot e^{\lambda x} f(x)$$

$$\begin{aligned} \int_{-\infty}^{\infty} g_\lambda(x) dx &= \int_{-\infty}^{\infty} M \cdot e^{\lambda x} f(x) dx \\ &= M \int_{-\infty}^{\infty} e^{\lambda x} f(x) dx \\ &= M \mathbb{E}_f[e^{\lambda x}] \end{aligned}$$

$$M \mathbb{E}_f[e^{\lambda x}] = 1, \quad M = \frac{1}{\mathbb{E}[e^{\lambda x}]}$$

$$\psi'(\lambda) = \int_{-\infty}^{\infty} x \cdot \underbrace{\frac{e^{-\lambda x} f(x)}{\mathbb{E}[e^{-\lambda x}]}}_{g_\lambda} dx$$

$$\psi'(\lambda) = \mathbb{E}_{g_\lambda}[x]$$

$$\psi'(\lambda) = \frac{\mathbb{E}[x e^{-\lambda x}]}{\mathbb{E}[e^{-\lambda x}]}$$

$$\psi''(\lambda) = \frac{\mathbb{E}[x^2 e^{-\lambda x}]}{\mathbb{E}[e^{-\lambda x}]} - \left(\frac{\mathbb{E}[x e^{-\lambda x}]}{\mathbb{E}[e^{-\lambda x}]} \right)^2$$

$$\psi''(\lambda) = \text{Var}_{g_\lambda}[x]$$

Involve Taylor theorem

$$\psi(\lambda) = \psi(0) + \psi'(0)\lambda + \psi''(\tilde{\lambda}) \frac{\lambda^2}{2}$$

$$\tilde{\lambda} \in [0, \lambda]$$

$$= 0 + 0 + \frac{\lambda^2}{2} \text{Var}_{g_{\tilde{\lambda}}} [x]$$

$$\leq \frac{\lambda^2}{2} \left(\frac{b-a}{4} \right)^2$$

$$= \frac{\lambda^2 (b-a)^2}{8}$$

$$M_x(\lambda) = e^{\psi(\lambda)}$$

$$\leq e^{\frac{\lambda^2 (b-a)^2}{8}}$$

$$\mathbb{E}[e^{\lambda x}] \leq e^{\frac{\lambda^2 (b-a)^2}{2 \times 4}}$$

$X \in [a, b]$ is subgaussian with

$$\sigma = \frac{(b-a)^2}{4}$$

Summary of concentration inequalities

Chebyshev

$$P(|\hat{\mu} - \mu| \geq \varepsilon) \leq \frac{\sigma^2}{n\varepsilon^2} \quad (\text{rate } \frac{1}{n})$$

Cramer
chernoff

$$P(|\hat{\mu} - \mu| \geq \varepsilon) \leq 2e^{-\left(\frac{n\varepsilon^2}{2\sigma^2}\right)} \quad (\text{rate is exponential})$$

Chebyshev

$$\hat{\mu} \in \left[\mu - \sqrt{\frac{\sigma^2}{n\delta}}, \mu + \sqrt{\frac{\sigma^2}{n\delta}} \right]$$

as $\delta \rightarrow 0$, $\rightarrow \infty$

Cramer
chernoff

$$\hat{\mu} \in \left[\mu - \sqrt{\frac{2\sigma^2 \log(2/\delta)}{n}}, \mu + \sqrt{\frac{2\sigma^2 \log(2/\delta)}{n}} \right]$$

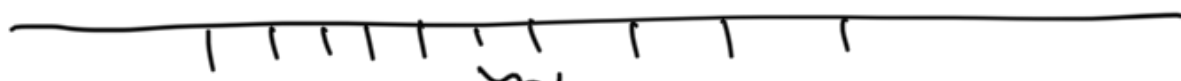
we are not doing any
better than a statistical rate, if I take
samples from a random variable the standard
deviation is of order $(\frac{1}{\sqrt{n}})$

$$\delta = 0.1 \quad \rightarrow \quad \delta = 0.01$$

$\log(2/\delta)$ becomes only 2 times
more

$$n = \frac{2 \sigma^2 \log(2/\delta)}{\epsilon^2}$$

(fixing ϵ and δ)



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