

Definitions

- Sample Space: S is the set of all possible outcomes
 - eg. For rolling 2 dice: $S = \{(1, 1), (1, 2), \dots, (6, 5), (6, 6)\}$
- Sample Point: Any element/outcome in the sample space S
- Event: Any subset E of the sample space
- Sure Event: the sample space itself
- Null Event: empty set \emptyset

Counting

Choose k from n	Order Matters	Not Matter
With Replacement	n^k	$\binom{n+k-1}{k}$
Without Replacement	$\frac{n!}{(n-k)!}$	$\binom{n}{k}$

- In a circle: $(n-1)!$

Probability

Inclusion-Exclusion Principle

- $P(A \cup B) = P(A) + P(B) - P(A \cap B)$
- $P(A \cup B \cup C) = P(A) + P(B) + P(C) - [P(A \cap B) + P(A \cap C) + P(B \cap C)] + P(A \cap B \cap C)$

Independent Events

- $P(A \cap B) = P(A) \times P(B)$
- $P(A|B) = P(A)$
- $P(A) = P(A \cap B) + P(A \cap B^c)$

Mutually Exclusive Events

- $P(A \cap B) = 0$ (B cannot happen if A happens)
- $P(A|B) = 0$
- $P(A_1 \cup A_2 \cup \dots \cup A_n) = P(A_1) + P(A_2) + \dots + P(A_n)$

Two non-trivial ($P > 0$) events can only be independent, or mutually exclusive, or neither, but **never both** at the same time

Conditional Probability

- $P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{P(B|A)P(A)}{P(B)}$
- $P(A|B \cap C) = \frac{P(A \cap B \cap C)}{P(B \cap C)} = \frac{P(B|A \cap C)P(A|C)}{P(B|C)} = \frac{P(B \cap C|A)P(A)}{P(B \cap C)}$
- $P(A \cap B) = P(A)P(B|A) = P(B)(A|B)$

De Morgan's Law

- $(A \cup B)^c = A^c \cap B^c$
- $(A \cap B)^c = A^c \cup B^c$

Partition

- If B_1, B_2, \dots, B_n are mutually exclusive and exhaustive (they are disjoint and their union = S), then B_1, B_2, \dots, B_n is a partition of S

Law of Total Probability (Bayes' Formula 1)

Let B_1, \dots, B_n be a partition of S :

$$P(A) = \sum_{i=1}^n P(B_i \cap A) = \sum_{i=1}^n P(B_i)P(A|B_i)$$

With extra conditioning:

$$P(A|C) = \sum_{i=1}^n P(A|B_i \cap C)P(B_i|C) = \sum_{i=1}^n P(A \cap B_i|C)$$

Special case when B and B^c are the partitions:

- $P(A) = P(A|B)P(B) + P(A|B^c)P(B^c)$
- $P(A) = P(A \cap B) + P(A \cap B^c)$

Bayes' Theorem

Let B_1, \dots, B_n be a partition of S . $\forall k \in 1, \dots, n$,

$$P(B_k|A) = \frac{P(B_k)P(A|B_k)}{\sum_{i=1}^n P(B_i)P(A|B_i)}$$

Discrete Random Variables

Probability Mass Function, $f_X(x)$ (PMF)

- Probability that a discrete random variable = x
- Given by $f_X(x) = P(X = x)$
- When asked to find PMF: find $\forall x, P(X = x)$

Properties:

1. $0 \leq f_X(x) \leq 1$
2. $\sum_x f_X(x) = 1$
3. $P(X \in E) = \sum_{x \in E} f_X(x)$

Cumulative Distribution Function, $F_X(x)$ (CDF)

- Probability that a discrete random variable is $\leq x$
- $F_X(x) = P(X \leq x) = \sum_{t \leq x} P(X = t)$

Properties:

1. $F_X(x)$ is a non-decreasing function: $x_1 < x_2 \Rightarrow F_X(x_1) \leq F_X(x_2)$
2. $0 \leq F_X(x) \leq 1$

Continuous Random Variables

- $P(X = x) = 0$
- $P(a < X \leq b) = \int_a^b f_X(x) dx = F_X(b) - F_X(a)$

Probability Density Function, $f_X(x)$ (PDF)

The function f_X is the PDF of the continuous random variable $X \iff$

1. $\forall x, f_X(x) \geq 0$
2. $\int_{-\infty}^{\infty} f_X(x) dx = 1$

Cumulative Distribution Function, $F_X(x)$ (CDF)

CDF of a continuous random variable X with PDF $f_X(x)$ is given by

$$F_X(x) = \int_{-\infty}^x f_X(t) dt$$

Properties:

1. $F_X(x)$ is a non-decreasing function of x
2. $\lim_{x \rightarrow -\infty} F_X(x) = 0$ AND $\lim_{x \rightarrow \infty} F_X(x) = 1$

Mean & Variance

Mean, μ (1st Moment)

- Discrete: $E(X) = \sum_x xP(X = x)$ OR $X \in \mathbb{N} \Rightarrow \sum_{k=1}^{\infty} P(X \geq k)$
- Random: $E(X) = \int_{-\infty}^{\infty} x f_X(x) dx$
- LOTUS: $E[g(X)] = \sum_x g(x) f_X(x)$ OR $\int_{-\infty}^{\infty} g(x) f_X(x) dx$
 - eg. to find k^{th} moment: $E(X^k) = \sum_x (x)^k P(X = x)$

Properties:

1. $E(aX + bY + c) = aE(X) + bE(Y) + c$

Variance, σ^2

- Discrete: $V(X) = E[(X - \mu_X)^2] = \sum_x (x - \mu_X)^2 f_X(x)$
- Random: $V(X) = \int_{-\infty}^{\infty} (x - \mu_X)^2 f_X(x) dx$
- $SD(X) = \sqrt{V(X)}$

Properties:

1. $V(X) \geq 0$
2. $V(X) = E(X^2) - [E(X)]^2$
3. $V(X) = 0 \Rightarrow P(X = \mu_X) = 1$
4. $V(a + bX) = b^2 V(X)$

Chebyshev's Inequality

If a random variable X has mean, μ , and SD, σ , the probability of getting a value which deviates from μ by at least $k\sigma$ is at most $\frac{1}{k^2}$

- $P(|X - \mu| > k\sigma) \leq \frac{1}{k^2}$ OR $P(|X - \mu| \leq k\sigma) \geq 1 - \frac{1}{k^2}$
- Applying $k = 2$, we conclude that for any random variable X , there is at most $\frac{1}{4}$ chance that it is 2 SD or further away from its mean

Joint Distribution

Joint Probability Mass Function

1. $f_{(X,Y)}(x,y) \geq 0, \forall (x,y) \in R_{X,Y}$
2. $\sum_x \sum_y f_{X,Y}(x,y) = \sum_x \sum_y P(X = x, Y = y) = 1$
3. $P((X,Y) \in A) = \sum_{(x,y) \in A} f_{X,Y}(x,y)$

Joint Probability Density Function

1. $f_{X,Y}(x,y) \geq 0, \forall (x,y) \in R_{X,Y}$
2. $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) dy dx = 1$

Examples

Suppose it is given that

$$F(0) = 0.58, F(1) = 0.72, F(2) = 0.76, F(3) = 0.81, F(4) = 0.88, F(5) = 0.94.$$

Compute the follow probabilities.

$$(a) P(2 \leq X \leq 5)$$

$$(b) P(X = 3)$$

Solution:

$$(a)$$

$$P(2 \leq X \leq 5) = F(5) - F(2^-) = F(5) - F(1) = 0.22.$$

$$(b)$$

$$P(X = 3) = F(3) - F(3^-) = F(3) - F(2) = 0.05.$$

Suppose that the random variable X is continuous with the following probability mass function:

$$f(x) = \begin{cases} \frac{x}{225}, & 0 < x < 15 \\ \frac{30-x}{225}, & 15 \leq x \leq 30 \\ 0, & \text{otherwise} \end{cases}$$

Find $E(X)$ and $V(X)$.

Solution:

We first obtain

$$\begin{aligned} E(X) &= \int_0^{15} x \cdot \frac{x}{225} dx + \int_{15}^{30} x \cdot \frac{30-x}{225} dx \\ &= \frac{1}{225} \left(\left[\frac{x^3}{3} \right]_0^{15} + \left[15x^2 - \frac{x^3}{3} \right]_{15}^{30} \right) \\ &= 15. \end{aligned}$$

Now,

$$\begin{aligned} E(X^2) &= \int_0^{15} x^2 \cdot \frac{x}{225} dx + \int_{15}^{30} x^2 \cdot \frac{30-x}{225} dx \\ &= \frac{1}{225} \left(\left[\frac{x^4}{4} \right]_0^{15} + \left[10x^3 - \frac{x^4}{4} \right]_{15}^{30} \right) \\ &= \frac{525}{2}. \end{aligned}$$

Thus

$$V(X) = E(X^2) - [E(X)]^2 = \frac{525}{2} - 15^2 = 37.5.$$

Let X denote the amount of time for which a book on 2-hour reserve at the Science Library is checked out by a randomly selected student and suppose X has the probability density function

$$f(x)=\begin{cases}\frac{x}{2}, & 0 < x < 2 \\ 0, & \text{otherwise}\end{cases}.$$

- (a) Find $E(X)$, $V(X)$ and σ_X .
- (b) If the borrower is charged an amount $h(X) = X^2$ when checkout duration is X , compute the expected charge $E[h(X)]$.

Solution:

(a)

$$\begin{aligned} E(X) &= \int_0^2 x \cdot \frac{x}{2} \, dx = \frac{4}{3}. \\ E(X^2) &= \int_0^2 x^2 \cdot \frac{x}{2} \, dx = 2. \\ V(X) &= 2 - (4/3)^2 = 2/9. \\ \sigma_X &= \sqrt{2/9}. \end{aligned}$$

- (b) $E[h(X)] = E(X^2) = 2$.

The probability density function of a continuous random variable X , the total number of hours, in units of 100 hours, that a family runs a vacuum cleaner over a period of one year, is given as follows.

$$f(x)=\begin{cases}x, & 0 < x < 1 \\ 2-x, & 1 \leq x \leq 2 \\ 0, & \text{otherwise}\end{cases}.$$

Find the average number of hours per year that families run their vacuum cleaners.

Solution:

$$\begin{aligned} E(X) &= \int_{-\infty}^{\infty} xf(x) \, dx \\ &= \int_0^1 x^2 \, dx + \int_1^2 x(2-x) \, dx \\ &= \left[\frac{x^3}{3}\right]_0^1 + \left[x^2 - \frac{x^3}{3}\right]_1^2 = 1. \end{aligned}$$

Families run their vacuum cleaners 100 hours per year on average.

The probability density function $f(x)$ of a random variable X is

$$f(x)=\begin{cases}2x, & 0 < x < 1 \\ 0, & \text{otherwise}\end{cases}.$$

Find the cumulative distribution function of X .

Solution:

$$\begin{aligned} F(x) &= \int_{-\infty}^x f(t) \, dt \\ &= \begin{cases} \int_{-\infty}^x 0 \, dt, & \text{for } x < 0 \\ \int_{-\infty}^0 0 \, dt + \int_0^x 2t \, dt, & \text{for } 0 < x < 1 \\ \int_{-\infty}^0 0 \, dt + \int_0^1 2t \, dt + \int_1^x 0 \, dt, & \text{for } x \geq 1 \end{cases} \\ &= \begin{cases} 0, & \text{for } x < 0 \\ x^2, & \text{for } 0 < x < 1. \\ 1, & \text{for } x \geq 1 \end{cases} \end{aligned}$$

The daily production of electric motors at a certain factory averaged 120 with a standard deviation of 10. Use the Chebyshev’s Inequality to find an interval that contains at least 90% of the daily production levels.

Solution:

Let X be the daily production of electric motors at that factory. Cheby-shev’s Inequality gives

$$P(\mu - k\sigma < X < \mu + k\sigma) \geq 1 - \frac{1}{k^2}.$$

Setting $1 - \frac{1}{k^2} = 0.90$ we obtain $k = \sqrt{10}$, so

$$P\left(\mu - \sqrt{10}\sigma < X < \mu + \sqrt{10}\sigma\right) \geq 0.90.$$

The interval we seek is

$$\left(120 - \sqrt{10} \times 10, 120 + \sqrt{10} \times 10\right) = \textcolor{red}{(88.4, 151.6)}.$$

A company has 2 production lines, A and B, which produces at most 5 and 3 machines respectively. Assume that the number of machines produced is a random variable.

Let (X,Y) represent the 2-dimensional random variable yielding the numbers of machines produced by Line A and Line B respectively on a given day.

The joint probability function, $f_{X,Y}(x,y)$ of (X,Y) is given as follow

		$f_{X,Y}(x,y)$						
<div><div><div><div></div><div>x</div></div><div><div>y</div><div></div></div></div></div>		0	1	2	3	4	5	Row Total
0		0	0.01	0.02	0.05	0.06	0.08	0.22
1		0.01	0.03	0.04	0.05	0.05	0.07	0.25
2		0.02	0.03	0.05	0.06	0.06	0.07	0.29
3		0.02	0.04	0.03	0.04	0.06	0.05	0.24
Column Total		0.05	0.11	0.14	0.20	0.23	0.27	1

What is the probability that more chips are produced by Line A than by Line B on a given day?

Solution:

Let $B = \{X > Y\}$. Then

$$\begin{aligned} P(B) &= P(X > Y) \\ &= P((X,Y) = (1,0)) + P((X,Y) = (2,0)) \\ &\quad + P((X,Y) = (2,1)) + \cdots + P((X,Y) = (5,3)) \\ &= f_{X,Y}(1,0) + f_{X,Y}(2,0) + f_{X,Y}(2,1) + \cdots + f_{X,Y}(5,3) \\ &= 0.01 + 0.02 + 0.04 + \cdots + 0.06 + 0.05 = 0.73 \end{aligned}$$

Suppose the two-dimensional continuous random variable (X,Y) has the joint probability density function

$$f_{X,Y}(x,y)=\begin{cases}x^2+\frac{xy}{3}, & \text{for } 0 \leq x \leq 1, 0 \leq y \leq 2 \\ 0, & \text{otherwise}\end{cases}.$$

Verify that $f_{X,Y}$ is a joint probability density function and compute $P(X+Y \geq 1)$

Solution:

It is clear that $f_{X,Y}(x,y) \geq 0$ for all $0 \leq x \leq 1, 0 \leq y \leq 2$.

$$\begin{aligned} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f_{X,Y}(x,y) \, dx \, dy &= \int_0^2 \int_0^1 \left(x^2 + \frac{xy}{3}\right) \, dx \, dy \\ &= \int_0^2 \left[\frac{x^3}{3} + \frac{x^2y}{6}\right]_0^1 \, dy \\ &= \int_0^2 \left(\frac{1}{3} + \frac{y}{6}\right) \, dy = \left[\frac{y}{3} + \frac{y^2}{12}\right]_0^2 = 1. \end{aligned}$$

Thus $f_{X,Y}$ is a joint probability density function.

$$\begin{aligned} P(X+Y \geq 1) &= \int_0^1 \int_{1-x}^2 \left(x^2 + \frac{xy}{3}\right) \, dy \, dx \\ &= \int_0^1 \left[x^2y + \frac{xy^2}{6}\right]_{1-x}^2 \, dx \\ &= \int_0^1 \frac{1}{6} (5x^3 + 8x^2 + 3x) \, dx \\ &= \frac{1}{6} [5x^4/4 + 8x^3/3 + 3x^2/2]_0^1 = \textcolor{red}{\frac{65}{72}}. \end{aligned} \tag{3.1}$$