

## TOAs

$$t^{\text{obs}} = t^{\text{det}}(c_{\text{true}}) + n$$

contains everything except timing model

stochastic

## Residuals

$$s_t = t^{\text{obs}} - t^{\text{det}}(c_{\text{est}}) = t^{\text{det}}(c_{\text{true}}) - t^{\text{det}}(c_{\text{est}}) + n$$

$$\text{let } c_{\text{true}} = c_{\text{est}} + \varepsilon$$

$$\Rightarrow s_t = t^{\text{det}}(c_{\text{est}} + \varepsilon) - t^{\text{det}}(c_{\text{est}}) + n$$

$$= t^{\text{det}}(c_{\text{est}}) + \left. \frac{\partial t^{\text{det}}(c_{\text{est}} + \varepsilon)}{\partial \varepsilon} \right|_{\varepsilon=0} \varepsilon - t^{\text{det}}(c_{\text{est}}) + n + \mathcal{O}(\varepsilon^2)$$

$$\approx \left. \frac{\partial t^{\text{det}}(c_{\text{est}} + \varepsilon)}{\partial \varepsilon} \right|_{\varepsilon=0} \varepsilon + n$$

$$= M \varepsilon + n$$

design matrix

only depends on derivatives  
of timing model

\* in practice this procedure is iterated  
by using either a standard weighted least  
squares fit or a generalized least  
squares fit.

## Likelihood

$n$  follows Gaussian with unknown covariance matrix  $C(\phi)$

$$p(n|\phi) = \frac{1}{\sqrt{\det(2\pi C)}} \exp\left(-\frac{1}{2} n^T C^{-1} n\right)$$

where  $C(\phi) = \langle n n^T \rangle$

$\Rightarrow p(st|\phi, \varepsilon) = \frac{1}{\sqrt{\det(2\pi C)}} \exp\left(-\frac{1}{2} (st - M\varepsilon)^T C^{-1} (st - M\varepsilon)\right)$

← this is what is minimized in LS

$$\text{let } g = M^T C^{-1} st \quad + \quad \Gamma = M^T C^{-1} M$$

log-likelihood  $\rightarrow \log p(st|\phi, \varepsilon) \approx p(st|\phi, 0) = \varepsilon^T g - \frac{1}{2} \varepsilon^T \Gamma \varepsilon$   
 $= \log \Lambda$

maximize  $\log \Lambda \rightarrow 0 = \frac{\partial \log \Lambda}{\partial \varepsilon} = g - \Gamma \hat{\varepsilon}$

$$\text{ML } \hat{\varepsilon} = \Gamma^{-1} g = (M^T C^{-1} M)^{-1} M^T C^{-1} st$$

$$\sigma_{\hat{\varepsilon}} = \sqrt{\text{diag}(\Gamma^{-1})} = \sqrt{\text{diag}((M^T C^{-1} M)^{-1})}$$

so-called post-fit residuals

$$st^{\text{post}} = st - M \hat{\varepsilon} = \underbrace{(I - M(M^T C^{-1} M)^{-1} M^T C^{-1})}_{R} st = R st$$

# Rank-reduced noise model

residuals - 
$$s_t = M\varepsilon + F_a + U_j + n$$

$\downarrow$                        $\downarrow$                        $\downarrow$                        $\searrow$   
 timing                      real                      jitter                      white noise  
 model                      noise

$$F_a \rightarrow \sum_{n=1}^N \left[ b_n \sin(2\pi n/T) + c_n \cos(2\pi n/T) \right]$$

$$F = \begin{bmatrix} \sin(2\pi t_1/T) & \cos(2\pi t_1/T) & \dots & \sin(2\pi t_N/T) & \cos(\dots) \\ \vdots & \vdots & & & \\ \sin(2\pi t_{N_{\text{trn}}}/T) & \cos(2\pi t_{N_{\text{trn}}}/T) & \dots & \dots & \dots \end{bmatrix}$$

$$a = \begin{bmatrix} b_1 \\ c_1 \\ b_2 \\ c_2 \\ \vdots \\ \vdots \end{bmatrix}$$

resids  $\begin{bmatrix} \vdots & \vdots & \vdots \\ \vdots & \vdots & \vdots \\ \vdots & \vdots & \vdots \end{bmatrix}$  time

$[U] = N_{\text{trn}} \times N_{\text{epoch}}$

$$U = \begin{bmatrix} 1 & 0 & 0 \\ 1 & 0 & 0 \\ 1 & 0 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 1 & 0 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \\ 0 & 0 & 1 \end{bmatrix}$$

$\rightarrow$  maps TOAs to epochs

can describe jitter

White noise covariance

$N = \langle n n^T \rangle \rightarrow$  pdf of noise

$$p(n | \phi_{\text{white}}) = \frac{1}{\sqrt{\det(2\pi N)}} \exp\left(-\frac{1}{2} n^T N^{-1} n\right)$$

components of  $N \rightarrow N_{ij,k} = E_k^2 W_{ij} + Q_k^2 \delta_{ij}$

$ij$  - label TOA number,  $k$  labels backend  $\rightarrow$  diag( $\sigma_i^2$ )

full Likelihood function

$$p(st | \varepsilon, a, j, \phi_{\text{white}}) = \frac{1}{\sqrt{\det(2\pi N)}} \exp\left(-\frac{1}{2} (st - E_s - M_s - U_s)^T N^{-1} (st - E_s - M_s - U_s)\right)$$

Priors

$$p(j | J_k) = \frac{\exp\left(-\frac{1}{2} j^T J^{-1} j\right)}{\sqrt{\det(2\pi J)}} , \quad p(a | \Phi) = \frac{\exp\left(-\frac{1}{2} a^T \Phi^{-1} a\right)}{\sqrt{\det(2\pi \Phi)}}$$

$$p(\varepsilon | X) = \frac{\exp\left(-\frac{1}{2} \varepsilon^T X^{-1} \varepsilon\right)}{\sqrt{\det(2\pi X)}}$$

$$X \sim \begin{bmatrix} \infty & & \\ & \infty & \\ & & \infty \end{bmatrix}$$

$$J = J_k \delta_{ij} , \quad \Phi_{ij,m} = \begin{bmatrix} 10^{P_1} & & & & \\ & 10^{P_2} & & & \\ & & 10^{P_2} & & \\ & & & 10^{P_2} & \\ & & & & \dots & 10^{P_n} & 10^{P_n} \end{bmatrix}$$

posterior

$$p(\varepsilon, a, j, \phi_{\text{data}}, j_k, p_n) \propto p(st | \varepsilon, a, j, \phi_{\text{data}}) p(j | j_k) p(a | p_n)$$

$\varepsilon, j, a$  are nuisance parameters

$$\text{let } T = \begin{bmatrix} M & F & U \end{bmatrix} \quad b = \begin{bmatrix} \varepsilon \\ a \\ j \end{bmatrix}$$

prior is now

$$p(b | \phi) = \frac{\exp(-\frac{1}{2} b^T B^{-1} b)}{\sqrt{\det(2\pi B^{-1})}}$$

$$B = \begin{bmatrix} \sigma^2 & & \\ & \phi & \\ & & j \end{bmatrix}$$

likelihood  $p(st | b, \phi) = \frac{\exp(-\frac{1}{2} (st - Tb)^T N^{-1} (st - Tb))}{\sqrt{\det(2\pi N^{-1})}}$

posterior  $p(b, \phi | st) \propto \frac{\exp(-\frac{1}{2} b^T B^{-1} b)}{\sqrt{\det(2\pi B^{-1})}}$

just like before,  $d = T^T N^{-1} st$   
 $\Sigma = (T^T N^{-1} T + B^{-1})^{-1}$

$$\hat{b} = \Sigma^{-1} d$$

$$\sigma_{\hat{b}} = \sqrt{\text{diag}(\Sigma^{-1})}$$

can write likelihood as

$$p(st|b, \emptyset) = \frac{\exp\left(-\frac{1}{2} [st^T N^{-1} st - d^T \Sigma^{-1} d]\right)}{\sqrt{\det(2\pi N)} \sqrt{\det(2\pi B)}} \\ \times \exp\left(-\frac{1}{2} (b - \hat{b})^T \Sigma^{-1} (b - \hat{b})\right)$$

Marginalize

$$p(st|\emptyset) = \int db \, p(st|b, \emptyset) = \sqrt{\frac{(2\pi)^{\dim b}}{\det \Sigma}}$$

$\therefore$  Marginalized likelihood

$$p(st|\emptyset) = \frac{\exp\left(-\frac{1}{2} [st^T N^{-1} st - d^T \Sigma^{-1} d]\right)}{\sqrt{(2\pi)^{\dim N} \det(N) \det(B) \det(\Sigma)}}$$