Sampling Distributions II

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Parameter vs. Statistic

The mean GPA of all UCSB undergraduates is reported by the University to be 3.2 and the mean GPA of a sample of 500 UCSB undergraduates was found to be 2.96.

- (a) 3.2 is a parameter and 2.96 is a statistic
- (b) 3.2 is a statistic and 2.96 is a parameter
- (c) both are statistics
- (d) both are parameters

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Example: Average number of cars in US households

- Record the number of cars in each household in the US.
- Population: All US Households
 Population size: N = 312,615,000
 - Dataset: x_1, \ldots, x_N , where x_1 would be the number of cars in the 1st household in the population etc.
- Population Average:

$$\mu = \frac{\mathbf{x}_1 + \mathbf{x}_2 + \ldots + \mathbf{x}_N}{N}$$

Population Standard deviation:

$$\sigma = \sqrt{\frac{1}{N-1} \sum_{\text{for all } x} (x-\mu)^2}$$

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Sample

Sample Dataset: $\{X_1, \ldots, X_n\}$

Sample size: n

Goal

Based on the sample values, $\{X_1, \dots, X_n\}$, where n is the sample size, we try to reach some conclusions for the parameters of interest.

- The ideal sample:
 - representative
 - unbiased
- Choose the sample at random.

Simple random sample

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Simple Random Sample

$\{X_1, \dots, X_n\}$ is a Simple Random Sample if

- i. if a certain member of the population is chosen, this does not affect the chances of another member to be chosen.
- ii. every member of the population is equally likely to be chosen.

In other words...

- i. $\{X_1, \dots, X_n\}$ are independent
- ii. $\{X_1, \ldots, X_n\}$ are identically distributed (i.e. have the same pmf or pdf).

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Estimation of μ

- An estimator of a parameter is a statistic whose value in the sample is used to estimate this parameter.
 - An estimator for μ is the sample mean, \bar{x} .
 - An estimator for σ is the sample standard deviation, s.

Examples

- X_1, \ldots, X_n is a sample of n American households.
- An estimate for μ will be

$$\bar{X} = \frac{X_1 + \ldots + X_n}{n}$$

• An estimator for σ will be $s = \sqrt{\frac{1}{n-1} \sum_{\text{for all } x} (x - \bar{x})^2}$.

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Properties of \bar{x}

Question

Is \bar{x} a "good" estimator for μ ?

- This depends on <u>how</u> the sample X_1, \ldots, X_n is chosen.
- If X₁,..., X_n are iid, i.e. if we have a simple random sample (SRS), then the sample mean X̄ has some has some very nice properties.

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Remark

- \bar{X} is a random variable, thus
 - ▶ It has a distribution, which we will call "sampling distribution".

Example

- Suppose that many different researchers collect SRS of households (all of them with the same sample size n).
- Then, each researcher will compute a different sample mean \bar{x} , i.e. we will have the following table:

Researcher 1	<i>x</i> ₁
Researcher 2	\bar{x}_2

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Central Limit Theorem

Central Limit Theorem

 $X_1, X_2, ..., X_n$ are independent and identically distributed and n is large (i.e. n>30), then

$$\mathbb{E}(ar{x}) = \mu$$
 $\mathrm{s.}d.(ar{x}) = rac{\sigma}{\sqrt{n}}$

$$\bar{\mathbf{x}} \sim \mathcal{N}\left(\mu, \frac{\sigma}{\sqrt{\mathbf{n}}}\right)$$

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When do we use the CLT?

- How large n should be?
- The Central Limit Theorem does NOT say that every individual random variable is approximately normal.
- The Central Limit Theory applies independently of the distribution of X. It suffices to know its expectation and its standard deviation.

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Example: Exam scores

In a previous year of PSTAT 5A, the student scores on exams had mean 74 and standard deviation 14. The instructor gave a final exam in a class of 64 students.

 Approximate the probability that the average test scores in the class exceeds 80.

Remark!

In the problem, it is not mentioned that the exam scores follow a Normal distribution!

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Example: Exam scores

- X_i = test score of the *i*th student in the class of 64 students, i=1,...,64.
- Average test score: $\bar{x} = \frac{X_1 + ... + X_{64}}{64}$
- CLT assumptions? (iid, n>30)

$$P(\bar{x} > 80) = P\left(Z > \frac{80 - 74}{14/\sqrt{64}}\right) = P(Z > 3.429)$$
$$= 1 - P(Z \le 3.429) = 1 - 0.9997 = 0.0003$$

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Turn on your clickers!

We have a sample of 100 uniform random variables

 $X_1, X_2, \dots, X_{100} \sim \textit{U}[0, 10]$ with mean 5 and standard deviation 8.3.

The distribution of $\bar{x} = \frac{X_1 + X_2 + ... + X_{100}}{100}$ can be approximated by a

- (a) Normal(5, 8.3/10)
- (b) Uniform[0, 10]
- (c) Uniform[5, 8.3]
- (d) Normal(5, 8.3/100)

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Turn on your clickers!

We have a sample of 100 uniform random variables

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The distribution of $\bar{x} = \frac{X_1 + X_2 + ... + X_{100}}{100}$ can be approximated by a

- (a) Normal(5, 8.3/10)
- (b) Uniform[0, 10]
- (c) Uniform[5, 8.3]
- (d) Normal(500, 8.3/100)

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Population & Sample Proportion

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Survey

- Simple Random Sample
- 55% are in favor to a statement
- Repeat survey
 - ▶ 56% are in favor
 - 52% are in favor
 - **>**
- Sample is random

Example: Elections

- 55% support the Democrats
- 56% support the Democrats
- 52% support the Democrats

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Sampling Proportion

- $X \sim Binomial(n, p)$
 - p is a population parameter (typically unknown)
- If we have X but not p,
 - Sample proportion

$$\hat{p} = \frac{X}{n}$$

- ▶ Sample proportion ≈ Population proportion
- p̂ is an estimator for p.

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Sampling Proportion

Example: Elections

- X = number of respondents who support the Democrats
- n = total number of respondents in the survey
- $\hat{p} = \frac{X}{n}$ = proportion of respondents who support the Democrats
- We estimate the proportion of respondents in the population who support the democrats based on the sample proportion.

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Sampling Distribution of \hat{p}

Expectation

$$\mathbb{E}(\hat{p}) = \frac{n\,p}{n} = p$$

Standard deviation

$$s_{\hat{p}} = \sqrt{\frac{p(1-p)}{n}}$$

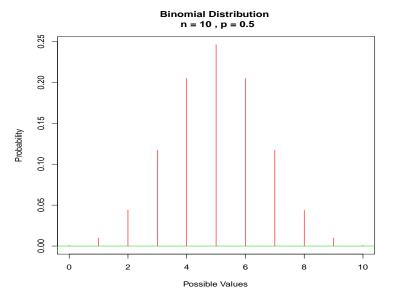
Standard error

s.e.
$$(\hat{p}) = \sqrt{\frac{\hat{p}(1-\hat{p})}{n}}$$

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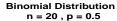
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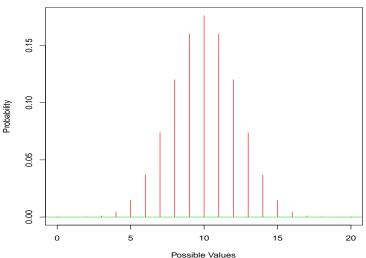
Let *n* get big: n = 10, p = 0.5



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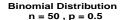
Let *n* get big: n = 20, p = 0.5

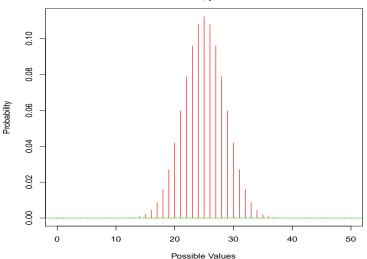




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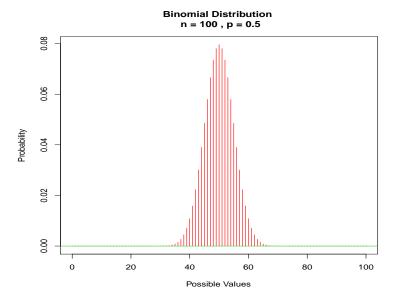
Let *n* get big: n = 50, p = 0.5





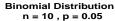
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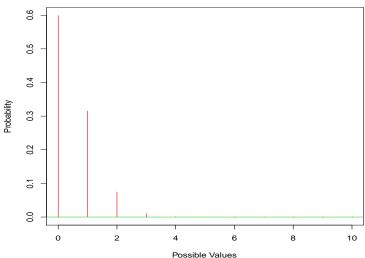
Let *n* get big: n = 100, p = 0.5



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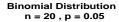
Let *n* get big: n = 10, p = 0.05

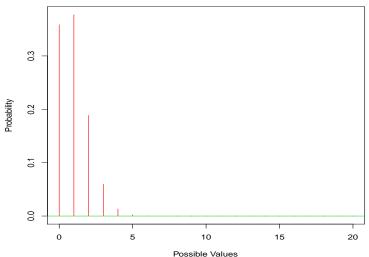




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Let *n* get big: n = 20, p = 0.05

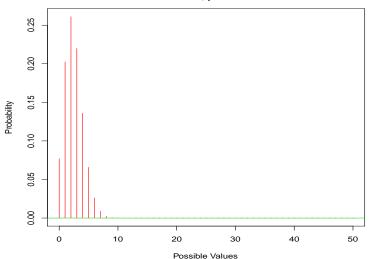




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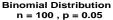
Let *n* get big: n = 50, p = 0.05

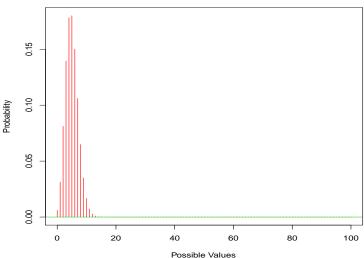
Binomial Distribution n = 50, p = 0.05



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Let *n* get big: n = 100, p = 0.05





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Law of Large Numbers

- As $n \to \infty$
- ullet Sample Proportion o Population proportion
- Standard Deviation → 0

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Central Limit Theorem

If n is large enough, the sample proportion \hat{p} behaves approximately as a normal random variable with

- Mean: μ = p
- Standard deviation: $\sigma = \sqrt{\frac{p(1-p)}{n}}$.

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Standardization for Binomial

In other words, if *n* is *large enough*

$$Z pprox rac{\hat{
ho} -
ho}{\sqrt{rac{
ho \; (1-
ho)}{n}}} pprox rac{\hat{
ho} -
ho}{\sqrt{rac{\hat{
ho} \; (1-\hat{
ho})}{n}}}$$
 $Z \sim \mathcal{N}(0,1)$

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When can we use this approximation?

Rule

- A population with a fixed proportion
- Random Sample

Independent Equally likely (equal chance)

Sample size is large

$$np > 10$$

 $n(1-p) > 10$

i.e. this is a binomial experiment with normal approximation.

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Example: Tossing a coin n times

Suppose we flip a coin 50 times with $\mathbb{P}(Heads) = 0.25$.

- (a) What is the distribution of the sample proportion?
- (b) What is the probability to have more than 50% Heads in the 50 tosses?

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Example: Tossing a coin n times

X =# of Heads in the 50 tosses

- We have a fixed number of tosses (n = 50).
- Each toss is independent of the others.
- There are two possible outcomes (Heads, Tails)

This is a Binomial experiment!

Check also

$$n p = 50 \cdot 0.25 = 12.5 > 10$$

 $n (1 - p) = 50 \cdot 0.75 = 37.5 > 10$

(This guarantees that the sample size is large enough.)

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Example: Tossing a coin n times

(a)
$$\mu = p = 0.25$$

 $\sigma = \sqrt{\frac{p(1-p)}{n}} = \sqrt{\frac{0.25(1-0.25)}{50}} = 0.06$
 $\hat{p} \sim \mathcal{N}(0.25, 0.06)$

(b)

$$\mathbb{P}(\hat{p} \ge 0.5) = \mathbb{P}\left(Z \ge \frac{0.5 - 0.25}{0.06}\right)$$
$$= \mathbb{P}(Z \ge 4.17) \approx 0$$

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