# Cost effective prediction of bodyfat

An example of project presentation slides

Aki Vehtari Aalto University

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Introduce yourself

Bodyfat percentage is related to many health outcomes

[Nice figures here]

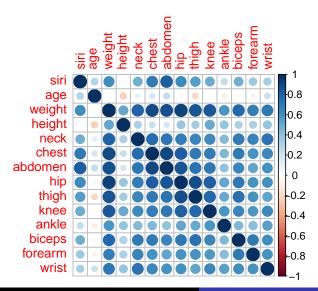
- Bodyfat percentage is related to many health outcomes
- Relatively accurate way to measure bodyfat is to weight a person in air and immersed in water
  - proportion of body fat can be derived from body density with Siri's (1956) formula
  - water immersion requires a big tub for the water and harness system for lowering a person to water

[Nice figures here]

- Bodyfat percentage is related to many health outcomes
- Relatively accurate way to measure bodyfat is to weight a person in air and immersed in water
  - proportion of body fat can be derived from body density with Siri's (1956) formula
  - water immersion requires a big tub for the water and harness system for lowering a person to water
- Can we estimate the bodyfat percentage with faster and a smaller equipment?
  - with just a scale and measure tape?
  - 252 subjects

[Nice figures here]

With just a scale and measure tape?



### Bodyfat predictive model

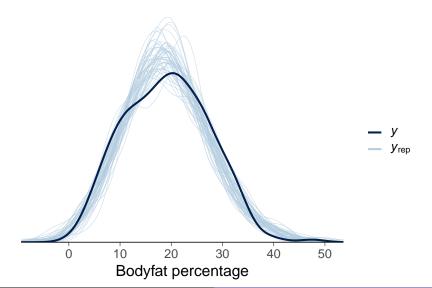
• Gaussian linear regression model with normal vs. regularized horseshoe prior ( $p_0 = 5$ ) on coefficients

#### Bodyfat predictive model

- Gaussian linear regression model with normal vs. regularized horseshoe prior ( $p_0 = 5$ ) on coefficients
- Model build with rstanarm and inference run with Stan
  - all convergence diagnostics were good

# Bodyfat model checking

Posterior predictive checking



#### Bodyfat model comparison

- Leave-one-out cross-validation comparison
  - no difference

```
\begin{array}{ccc} & & \text{elpd\_diff} & \text{se\_diff} \\ \text{RHS prior} & 0.0 & 0.0 \\ \text{Gaussian prior} & -1.1 & 2.2 \end{array}
```

#### Bodyfat model comparison

- Leave-one-out cross-validation comparison
  - no difference

```
elpd_diff se_diff
RHS prior 0.0 0.0
Gaussian prior -1.1 2.2
```

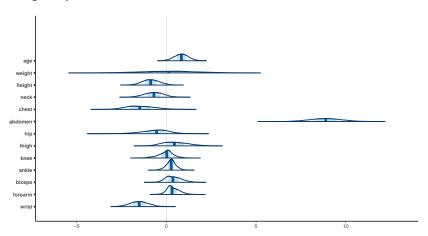
Computed from 4000 by 250 log-likelihood matrix

Monte Carlo SE of elpd\_loo is 0.1.

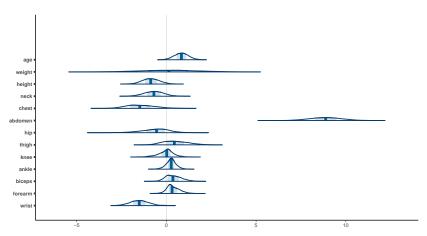
Pareto k diagnostic values:

```
Count Pct. Min. n_eff
(-Inf, 0.5] (good) 249 99.6% 1374
(0.5, 0.7] (ok) 1 0.4% 724
(0.7, 1] (bad) 0 0.0% <NA>
(1, Inf) (very bad) 0 0.0% <NA>
```

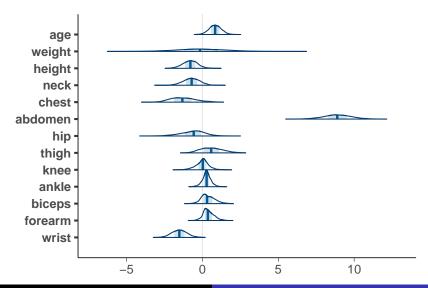
#### Marginal posteriors of coefficients



#### Check that the font in all figures is big enough!



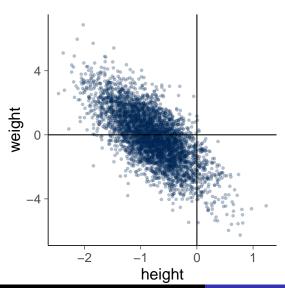
Marginal posteriors of coefficients (Much better!)



#### Figure font size

#### For example:

Bivariate marginal of weight and height



#### Bodyfat variable selection

- Do we need all the measurements?
- We find the model with a minimal set of variables which have similar predictive performance as the model with all variables

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- Do we need all the measurements?
- We find the model with a minimal set of variables which have similar predictive performance as the model with all variables
- We use projection predictive variable selection implemented in projpred package

### Projective predictive covariate selection

• The full model predictive distribution represents our best knowledge about future  $\tilde{y}$ 

$$p(\tilde{y}|D) = \int p(\tilde{y}|\theta)p(\theta|D)d\theta,$$

where  $\theta = (\beta, \sigma^2)$ ) and  $\beta$  is in general non-sparse (all  $\beta_i \neq 0$ )

- What is the best distribution  $q_{\perp}(\theta)$  given a constraint that only selected covariates have nonzero coefficient
- Optimization problem:

$$q_{\perp} = \arg\min_{q} \frac{1}{n} \sum_{i=1}^{n} \mathrm{KL} \bigg( p(\tilde{y}_{i} \mid D) \, \| \, \int p(\tilde{y}_{i} \mid \theta) q(\theta) d\theta \bigg)$$

 Optimal projection from the full posterior to a sparse posterior (with minimal predictive loss)

# For 10min presentation, too much information

ullet The full model predictive distribution represents our best knowledge about future  $\tilde{y}$ 

$$p(\tilde{y}|D) = \int p(\tilde{y}|\theta)p(\theta|D)d\theta,$$

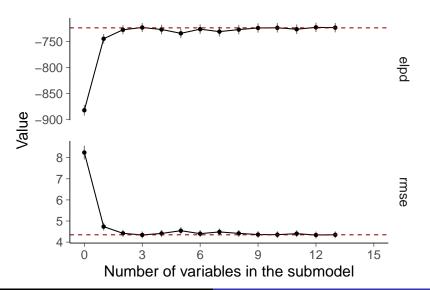
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- What is the best distribution  $q_{\perp}(\theta)$  given a constraint that only selected covariates have nonzero coefficient
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$$q_{\perp} = \arg\min_{q} \frac{1}{n} \sum_{i=1}^{n} KL \left( p(\tilde{y}_{i} \mid D) \parallel \int p(\tilde{y}_{i} \mid \theta) q(\theta) d\theta \right)$$

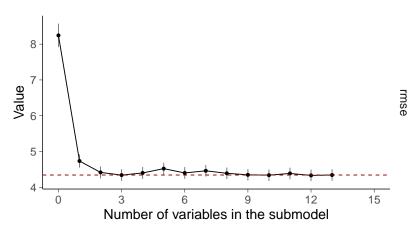
 Optimal projection from the full posterior to a sparse posterior (with minimal predictive loss)

The predictive performance of the full and submodels

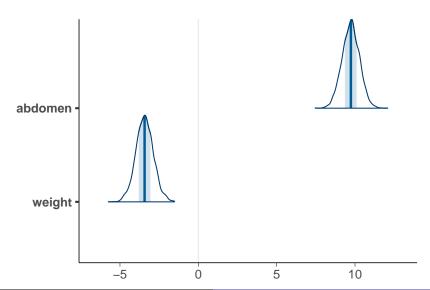


#### The predictive performance of the full and submodels

One of these plots is probably sufficient



Marginals of projected posterior



# Bodyfat - Conclusion

 Bodyfat percentage estimated using water immersion can be predicted using scale and tape measure

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- More results at avehtari.github.io/modelselection/bodyfat.html

# **THANKS!**

Don't ever end with a slide having just "THANKS"

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- "THANKS" slide has zero information content

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- Leave the conclusion slide or contact information slide

#### Conclusion

- Bodyfat percentage estimated using water immersion can be predicted using scale and tape measure
- The accuracy using mean of data is 16%-units (95% interval)
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#### Additional information

- You can have additional slides after the conclusion for supporting material to answer questions
  - for example, in this course, include Stan code and additional convergence and model checking results

#### Gaussian linear model with regularized horseshoe prior

```
// generated with brms 2.14.4
functions {
  vector horseshoe(vector z, vector lambda, real tau, real c2) {
    int K = rows(z);
    vector[K] lambda2 = square(lambda):
    vector[K] lambda tilde = sqrt(c2 * lambda2 ./ (c2 + tau^2 * lambda2));
    return z .* lambda tilde * tau;
data 4
 int < lower = 1 > N: // total number of observations
 vector[N] Y; // response variable
  int <lower=1> K; // number of population-level effects
  matrix[N, K] X; // population-level design matrix
  // data for the horseshoe prior
  real<lower=0> hs df; // local degrees of freedom
  real<lower=0> hs df global; // global degrees of freedom
  real < lower = 0 > hs df slab: // slab degrees of freedom
  real<lower=0> hs scale global; // global prior scale
  real<lower=0> hs scale slab; // slab prior scale
  int prior only: // should the likelihood be ignored?
```

#### Classification example: Pima Indians Diabetes

#### Predict diabetes based on

- Pregnancies
- Glucose
- Blood pressure
- Skin thickness
- Insulin
- BMI
- Diabetes Pedigree
- Age

768 observations

https://avehtari.github.io/modelselection/diabetes.html

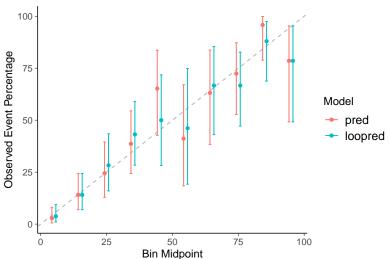
#### Classification example: Pima Indians Diabetes

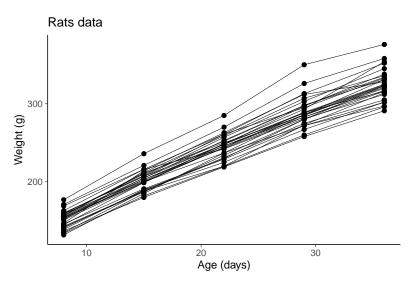
Leave-one-out cross-validation classification accuracy 78%

# Classification example: Pima Indians Diabetes

Leave-one-out cross-validation classification accuracy 78%

#### Calibration:





https://avehtari.github.io/modelselection/rats kcv.html

Simple linear model

```
fit_1 <- stan_glm(weight ~ age, data=dfrats)</pre>
```

Linear model with hierarchical intercept

```
fit_2 <- stan_glmer(weight ~ age + (1 | rat), data=dfrats)</pre>
```

Linear model with hierarchical intercept and slope

```
fit\_3 <- stan\_glmer(weight ~~age + (age | rat), data=dfrats)
```

#### Simple linear model

```
fit_1 <- stan_glm(weight ~ age, data=dfrats)</pre>
```

#### Linear model with hierarchical intercept

```
fit_2 \leftarrow stan_glmer(weight \sim age + (1 | rat), data=dfrats)
```

Linear model with hierarchical intercept and slope

```
fit\_3 <- stan\_glmer(weight ~~age + (age | rat), data=dfrats)
```

Instead of stan\_glm(er), use brm to get the Stan code, too.

#### Leave-one-out cross-validation

	elpd_diff	se_diff
hierarchical intercept and	slope 0.0	0.0
hierarchical intercept	-23.6	9.3
simple linear model	-109.6	13.3

#### Leave-one-out cross-validation

	elpd_diff	se_diff
hierarchical intercept and slope	0.0	0.0
hierarchical intercept	-23.6	9.3
simple linear model	-109.6	13.3

#### Leave-one-out cross-validation stacking model weights

			weight
hierarchical inte	rcept a	nd slope	0.93
hierarchical inte	rcept		0.07
simple linear mod	del		0.00