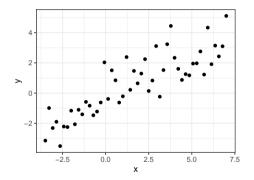
AutoML: Gaussian Processes

The Bayesian Linear Model

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Review: The Bayesian Linear Model I

Let $\mathcal{D}_{\text{train}} = \{(\mathbf{x}^{(1)}, y^{(1)}), ..., (\mathbf{x}^{(n)}, y^{(n)})\}$ be a training set of i.i.d. observations from some unknown distribution.



Let $\mathbf{y} = (y^{(1)},...,y^{(n)})^{\top}$ and $\mathbf{X} \in \mathbb{R}^{n \times p}$ be the design matrix where the i-th row contains vector $\mathbf{x}^{(i)}$.

Review: The Bayesian Linear Model II

The linear regression model is defined as

$$y^{(i)} = f(\mathbf{x}^{(i)}) + \epsilon^{(i)} = \boldsymbol{\theta}^{\top} \mathbf{x}^{(i)} + \epsilon^{(i)}, \text{ for all } i \in \{1, \dots, n\}.$$

The observed values $y^{(i)}$ differ from the function values $f(\mathbf{x}^{(i)})$ by some additive noise, which is assumed to be i.i.d. Gaussian

$$\epsilon^{(i)} \sim \mathcal{N}(0, \sigma^2).$$

Review: The Bayesian Linear Model III

• Let us assume we have **prior beliefs** about the parameter θ that are represented in a prior distribution $\theta \sim \mathcal{N}(\mathbf{0}, \tau^2 \mathbf{I}_p)$.

 Whenever data points are observed, we update the parameters' prior distribution according to Bayes' rule

$$\underbrace{p(\boldsymbol{\theta} \mid \mathbf{X}, \mathbf{y})}_{\text{posterior}} = \underbrace{\frac{p(\mathbf{y} \mid \mathbf{X}, \boldsymbol{\theta})}{p(\mathbf{y} \mid \mathbf{X}, \boldsymbol{\theta})}}_{\text{marginal}} \underbrace{\frac{p(\mathbf{y} \mid \mathbf{X})}{p(\mathbf{y} \mid \mathbf{X})}}_{\text{marginal}}$$

Review: The Bayesian Linear Model IV

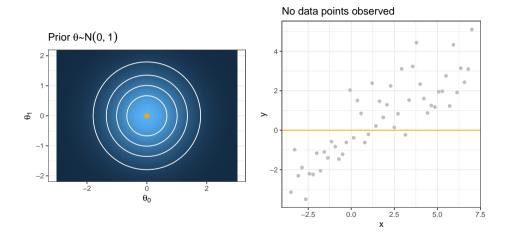
The posterior distribution of the parameter θ is again normal distributed (the Gaussian family is self-conjugate):

$$oldsymbol{ heta} \mid \mathbf{X}, \mathbf{y} \sim \mathcal{N}(\sigma^{-2} oldsymbol{A}^{-1} \mathbf{X}^{ op} \mathbf{y}, oldsymbol{A}^{-1})$$
, where $oldsymbol{A} := \sigma^{-2} \mathbf{X}^{ op} \mathbf{X} + rac{1}{ au^2} oldsymbol{I}_p$.

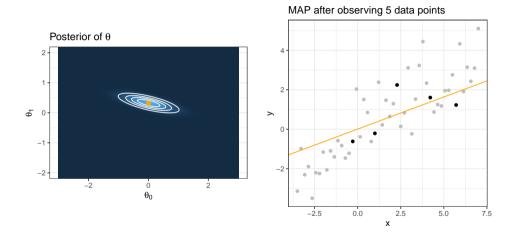
Note: If the posterior distributions $p(\theta \mid \mathbf{X}, \mathbf{y})$ are in the same probability distribution family as the prior $q(\theta)$, the prior and posterior are then called **conjugate distributions**, and the prior is called a **conjugate prior** for the likelihood function $p(\mathbf{y} \mid \mathbf{X}, \boldsymbol{\theta})$.

Note: The Gaussian family is **self-conjugate** with respect to a Gaussian likelihood function: choosing a Gaussian prior for a Gaussian likelihood ensures that the posterior is also Gaussian.

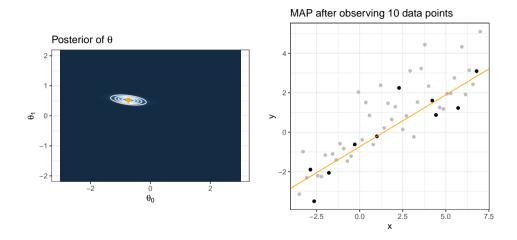
Review: The Bayesian Linear Model V



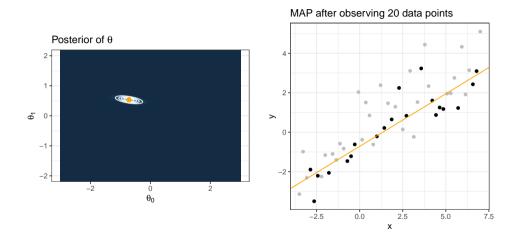
Review: The Bayesian Linear Model VI



Review: The Bayesian Linear Model VII



Review: The Bayesian Linear Model VIII



Review: The Bayesian Linear Model IX

Theorem:

• For a Gaussian prior on $\theta \sim \mathcal{N}(\mathbf{0}, \tau^2 \mathbf{I}_p)$ and a Gaussian likelihood $y \mid \mathbf{X}, \theta \sim \mathcal{N}(\mathbf{X}^\top \theta, \sigma^2 \mathbf{I}_n)$, the resulting posterior is Gaussian: $\mathcal{N}(\sigma^{-2} \mathbf{A}^{-1} \mathbf{X}^\top \mathbf{y}, \mathbf{A}^{-1})$, with $\mathbf{A} := \sigma^{-2} \mathbf{X}^\top \mathbf{X} + \frac{1}{\tau^2} \mathbf{I}_p$.

Proof:

Plugging in Bayes' rule and multiplying out yields

$$p(\boldsymbol{\theta} \mid \mathbf{X}, \mathbf{y}) \propto p(\mathbf{y} \mid \mathbf{X}, \boldsymbol{\theta}) q(\boldsymbol{\theta}) \propto \exp\left[-\frac{1}{2\sigma^{2}} (\mathbf{y} - \mathbf{X}\boldsymbol{\theta})^{\top} (\mathbf{y} - \mathbf{X}\boldsymbol{\theta}) - \frac{1}{2\tau^{2}} \boldsymbol{\theta}^{\top} \boldsymbol{\theta}\right]$$

$$= \exp\left[-\frac{1}{2} \left(\underbrace{\sigma^{-2} \mathbf{y}^{\top} \mathbf{y}}_{\text{doesn't depend on } \boldsymbol{\theta}} - 2\sigma^{-2} \mathbf{y}^{\top} \mathbf{X} \boldsymbol{\theta} + \sigma^{-2} \boldsymbol{\theta}^{\top} \mathbf{X}^{\top} \mathbf{X} \boldsymbol{\theta} + \tau^{-2} \boldsymbol{\theta}^{\top} \boldsymbol{\theta}\right)\right]$$

$$\propto \exp\left[-\frac{1}{2} \left(\sigma^{-2} \boldsymbol{\theta}^{\top} \mathbf{X}^{\top} \mathbf{X} \boldsymbol{\theta} + \tau^{-2} \boldsymbol{\theta}^{\top} \boldsymbol{\theta} - 2\sigma^{-2} \mathbf{y}^{\top} \mathbf{X} \boldsymbol{\theta}\right)\right]$$

$$= \exp\left[-\frac{1}{2} \boldsymbol{\theta}^{\top} \underbrace{\left(\sigma^{-2} \mathbf{X}^{\top} \mathbf{X} + \tau^{-2} \mathbf{I}_{p}\right)}_{\mathbf{x} = \mathbf{A}} \boldsymbol{\theta} + \sigma^{-2} \mathbf{y}^{\top} \mathbf{X} \boldsymbol{\theta}\right]$$

This expression resembles a normal density - except for the term in red!

Review: The Bayesian Linear Model X

Note: We need not worry about the normalizing constant since its mere role is to convert probability functions to density functions with a total probability of one.

We subtract a (not yet defined) constant c while compensating for this change by adding the respective terms ("adding 0"), emphasized in green:

$$p(\boldsymbol{\theta}|\mathbf{X}, \mathbf{y}) \propto \exp\left[-\frac{1}{2}(\boldsymbol{\theta} - c)^{\top}\mathbf{A}(\boldsymbol{\theta} - c) - c^{\top}\mathbf{A}\boldsymbol{\theta} + \underbrace{\frac{1}{2}c^{\top}\mathbf{A}c}_{\text{doesn't depend on }\boldsymbol{\theta}} + \sigma^{-2}\mathbf{y}^{\top}\mathbf{X}\boldsymbol{\theta}\right]$$
$$\propto \exp\left[-\frac{1}{2}(\boldsymbol{\theta} - c)^{\top}\mathbf{A}(\boldsymbol{\theta} - c) - c^{\top}\mathbf{A}\boldsymbol{\theta} + \sigma^{-2}\mathbf{y}^{\top}\mathbf{X}\boldsymbol{\theta}\right]$$

If we choose c such that $-c^{\top} \mathbf{A} \boldsymbol{\theta} + \sigma^{-2} \mathbf{y}^{\top} \mathbf{X} \boldsymbol{\theta} = 0$, the posterior is normal with mean c and covariance matrix \mathbf{A}^{-1} . Taking into account that \mathbf{A} is symmetric, this is if we choose

$$\sigma^{-2}\mathbf{y}^{\top}\mathbf{X} = c^{\top}\mathbf{A}$$

$$\Leftrightarrow \quad \sigma^{-2}\mathbf{y}^{\top}\mathbf{X}\mathbf{A}^{-1} = c^{\top}$$

$$\Leftrightarrow \quad c = \sigma^{-2}\mathbf{A}^{-1}\mathbf{X}^{\top}\mathbf{y}$$

as claimed.

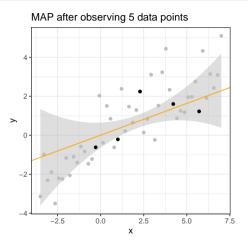
Review: The Bayesian Linear Model XI

- Based on the posterior destribution, $\boldsymbol{\theta} \mid \mathbf{X}, \mathbf{y} \sim \mathcal{N}(\sigma^{-2} \mathbf{A}^{-1} \mathbf{X}^{\top} \mathbf{y}, \mathbf{A}^{-1})$, we can derive the predictive distribution for a new observations \mathbf{x}_* .
- ullet The predictive distribution for the Bayesian linear model, i.e. the distribution of $oldsymbol{ heta}^ op \mathbf{x}_*$, is

$$y_* \mid \mathbf{X}, \mathbf{y}, \mathbf{x}_* \sim \mathcal{N}(\sigma^{-2}\mathbf{y}^\top \mathbf{X} \mathbf{A}^{-1} \mathbf{x}_*, \mathbf{x}_*^\top \mathbf{A}^{-1} \mathbf{x}_*).$$

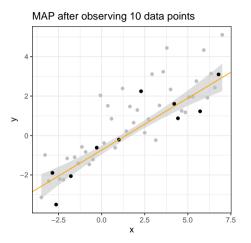
Note: This can be obtained by applying the rules for linear transformations of Gaussians.

Review: The Bayesian Linear Model XII



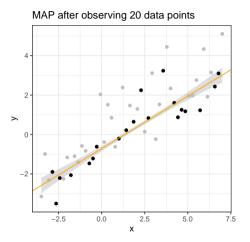
For every test input 5_* , we get a distribution over the prediction y_* . In particular, we get a posterior mean (orange) and a posterior variance (the grey region, which equals +/- two times the standard deviation).

Review: The Bayesian Linear Model XIII



For every test input 10_* , we get a distribution over the prediction y_* . In particular, we get a posterior mean (orange) and a posterior variance (the grey region, which equals +/- two times the standard deviation).

Review: The Bayesian Linear Model XIV



For every test input 20_* , we get a distribution over the prediction y_* . In particular, we get a posterior mean (orange) and a posterior variance (the grey region, which equals +/- two times the standard deviation).

Summary: The Bayesian Linear Model

- By switching to a Bayesian perspective, we have not only point estimation for the parameter θ but also whole **distributions**.
- From the posterior distribution of θ , we can derive a predictive distribution for $y_* = \theta^\top \mathbf{x}_*$.
- ullet We can perform online updates: the **posterior distribution** of ullet can be updated whenever new datapoints are observed.
- In the next step, we would like go beyond the linear funtions and develop a theory for functions with general shapes.