

**R V COLLEGE OF ENGINEERING**

(An autonomous institution affiliated to VTU, Belgaum)

Department of Mathematics**MAT231CT : Linear Algebra and Probability Theory****Unit 3: Random Variables****Topic Learning Objectives:**

- To apply the knowledge of the statistical analysis and theory of probability in the study of uncertainties.
- Define the degree of dependence between two random variables and measuring the same .

Prerequisites:

If an experiment is repeated under essentially homogeneous and similar conditions one generally comes across two types of situations:

- (i) The result or what is usually known as the 'outcome' is unique or certain.
- (ii) The result is not unique but may be one of the several possible outcomes.

The phenomena covered by (i) are known as 'deterministic' or 'predictable' phenomena. By a deterministic phenomenon the result can be predicted with certainty.

For example:

- (a) The velocity 'v' of a particle after time 't' is given by $v = u + at$ where 'u' is the initial velocity and 'a' is the acceleration. This equation uniquely determines 'v' if the right-hand quantities are known.
- (b) Ohm's Law , viz., $C = E/R$ where C is the flow of current, E the potential difference between the two ends of the conductor and R the resistance, uniquely determines the value C as soon as E and R are given.

A deterministic model is defined as a model which stipulates that the conditions under which an experiment is performed to determine the outcome of the experiment. For a number of situations the deterministic model suffices. However, there are phenomena (as covered by (ii) above) which do not lend themselves to deterministic approach and are known as 'unpredictable' or 'probabilistic' phenomena. For example:

- (a) In tossing of a coin one is not sure if a head or tail will be obtained.
- (b) If a light tube has lasted for t hours, nothing can be said about its further life. It may fail to function any moment.

In such cases chance or probability comes into picture which is taken to be a quantitative measure of uncertainty.

Some basic definitions:

Trial and Event: Consider an experiment which, though repeated under essentially identical conditions, does not give unique results but may result in any one of the several possible outcomes. The experiment is known as a trial and the outcomes are known as events or casts.

For example:

- (i) Throwing of a die is a trial and getting 1 (or 2 or 3, ... or 6) is an event.
- (ii) Tossing of a coin is a trial and getting head (H) or tail (T) is an event.
- (iii) Drawing two cards from a pack of well-shuffled cards is a trial and getting a king and a queen are events.

Exhaustive Events: The total number of possible outcomes in any trial is known as exhaustive events or exhaustive cases. For example:

- (i) In tossing of a coin there are two exhaustive cases, viz., head and tail.
- (ii) In throwing of a die, there are six, exhaustive cases since anyone of the 6 faces 1, 2, ..., 6 may come uppermost.
- (iii) In drawing two cards from a pack of cards the exhaustive number of cases is 52_{C_2} , since 2 cards can be drawn out of 52 cards in 52_{C_2} ways.

Favourable Events or Cases: The number of cases favourable to an event in a trial is the number of outcomes which entail the happening of the event. For example:

- (i) In drawing a card from a pack of cards the number of cases favourable to drawing of an ace is 4, for drawing a spade 13 and for drawing a red card is 26.
- (ii) In throwing of two dice, the number of cases favourable to getting the sum 5 is : (1,4) (4,1) (2,3) (3,2), i.e., 4.

Mutually exclusive events: Events are said to be mutually exclusive or incompatible if the happening of any one of them precludes the happening of all the others, i.e., if no two or more of them can happen simultaneously in the same trial. For example:

- (i) In throwing a die all the 6 faces numbered 1 to 6 are mutually exclusive since if any one of these faces comes, the possibility of others, in the same trial, is ruled out.
- (ii) Similarly in tossing a coin the events head and tail are mutually exclusive.

Equally likely events: Outcomes of a trial are set to be equally likely if taking into consideration all the relevant evidences, there is no reason to expect one in preference to the others. For example:

- (i) In tossing an unbiased or uniform coin, head or tail are likely events.
- (ii) In throwing an unbiased die, all the six faces are equally likely to come.



Independent events: Several events are said to be independent if the happening (or non-happening) of an event is not affected by the supplementary knowledge concerning the occurrence of any number of the remaining events. For example:

- (i) In tossing an unbiased coin the event of getting a head in the first toss is independent of getting a head in the second, third and subsequent throws.
- (ii) If one draws a card from a pack of well-shuffled cards and replace it before drawing the second card, the result of the second draw is independent of the first draw. But, however, if the first card drawn is not replaced then the second draw is dependent on the first draw.

There are three systematic approaches to the study of probability as mentioned below.

Mathematical or Classical or ‘a priori’ Probability: If a trial results in n exhaustive, mutually exclusive and equally likely cases and m of them are favourable to the happening of an event E then the probability ' p ' of happening of E is given by

$$p = P(E) = \frac{\text{Favourable number of cases}}{\text{Exhaustive number of cases}} = \frac{m}{n}$$

Since the number of cases favourable to the 'non-happening' of the event E are $(n - m)$, the probability ' q ' that E will not happen is given by

$$q = \frac{n - m}{n} = 1 - \frac{m}{n} = 1 - p \text{ gives } p + q = 1$$

Obviously p as well as q are non-negative and cannot exceed unity, i.e,

$$0 \leq p \leq 1, 0 \leq q \leq 1.$$

Statistical or Empirical Probability: If a trial is repeated a number of times under essentially homogeneous and identical conditions, then the limiting value of the ratio of the number of times the event happens to the number of trials, as the number of trials become indefinitely large, is called the probability of happening of the event. (It is assumed that the limit is finite and unique).

Symbolically, if in n trials an event E happens m times, then the probability ' p ' of the happening of E is given by $p = P(E) = \lim_{n \rightarrow \infty} \frac{m}{n}$.

Axiomatic Probability: Let A be any event in the sample space S , then $P(A)$ is called the probability of event A , if the following axioms are satisfied.

Axiom 1: $P(A) \geq 0$

Axiom 2: $P(S) = 1$, S being the sure event

Axiom 3: For two mutually exclusive events A & B , $P(A \cup B) = P(A) + P(B)$

Some important results:

1. The probability of an event always lies between 0 and i.e., $0 \leq P(A) \leq 1$.
2. If A and A' are complementary events to each other defined on a random experiment then $P(A) + P(A') = 1$.
3. Addition Theorem: If A and B are any two events with respective probabilities P(A) and P(B), then the probability of occurrence of at least one of the events is given by $P(A \cup B) = P(A) + P(B) - P(A \cap B)$.
4. The probability of null event is zero i.e., $P(\emptyset) = 0$.
5. For any two events A and B of a sample space S
 - (i) $P(A - B) = P(A) - P(A \cap B)$
 - (ii) $P(B - A) = P(B) - P(A \cap B)$
 - (iii) $P(\bar{A} \cap B) = P(B) - P(A \cap B) = P(A \cup B) - P(A)$
 - (iv) $P[(A - B) \cup (B - A)] = P(A) + P(B) - 2P(A \cap B)$
6. Addition Theorem for three events: If A, B and C are any three events with respective probabilities P(A), P(B) and P(C), then the probability of occurrence of at least one of the events is given by $P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(A \cap B) - P(B \cap C) - P(A \cap C) + P(A \cap B \cap C)$.

Random variable

Random variable is a real number X connected with the outcome of a random experiment E. For example, if E consists of three tosses of a coin, one can consider the random variable which is the number of heads (0, 1, 2 or 3).

Outcome	HHH	HHT	HTH	THH	TTH	THT	HTT	TTT
Value of X	3	2	2	2	1	1	1	0

Let S denote the sample space of a random experiment. A random variable means it is a rule which assigns a numerical value to each and every outcome of the experiment. Thus, random variable is a function $X(\omega)$ with domain S and range $(-\infty, \infty)$ such that for every real number a , the event $[X(\omega) \leq a] \in B$ the field of subsets in S. It is denoted as $f: S \rightarrow R$.

Note that all the outcomes of the experiment are associated with a unique number. Therefore, f is an example of a random variable. Usually, a random variable is denoted by letters such as X, Y, Z etc. The image set of the random variable may be written as $f(S) = \{0, 1, 2, 3\}$.

There are two types of random variables. They are;

1. Discrete Random Variable (DRV)
2. Continuous Random Variable (CRV).



Discrete Random Variable: A discrete random variable is one which takes only a countable number of distinct values such as 0, 1, 2, 3, Discrete random variables are usually (but not necessarily) counts. If a random variable takes at most a countable number of values, it is called a **discrete random variable**. In other words, a real valued function defined on a discrete sample space is called a discrete random variable.

Examples of Discrete Random Variable:

- (i) In the experiment of throwing a die, define X as the number that is obtained. Then X takes any of the values 1 – 6. Thus, $X(S) = \{1, 2, 3, \dots, 6\}$ which is a finite set and hence X is a DRV.
- (ii) If X be the random variable denoting the number of marks scored by a student in a subject of an examination, then $X(S) = \{0, 1, 2, 3, \dots, 100\}$. Then, X is a DRV.
- (iii) The number of children in a family is a DRV.
- (iv) The number of defective light bulbs in a box of ten is a DRV.

Probability Mass Function: Suppose X is a one-dimensional discrete random variable taking at most a countably infinite number of values x_1, x_2, \dots . With each possible outcome x_i , one can associate a number $p_i = P(X = x_i) = p(x_i)$, called the probability of x_i .

The numbers $p(x_i); i = 1, 2, \dots$ must satisfy the following conditions:

- (i) $p(x_i) \geq 0 \forall i$,
- (ii) $\sum_{i=1}^{\infty} p(x_i) = 1$.

This function p is called the **probability mass function** of the random variable X and the set $\{x_i, p(x_i)\}$ is called the probability distribution of the random variable X .

Remarks:

1. The set of values which X takes is called the spectrum of the random variable.
2. For discrete random variable, knowledge of the probability mass function enables us to compute probabilities of arbitrary events. In fact, if E is a set of real numbers, $(X \in E) = \sum_{x \in E \cap S} p(x)$, where S is the sample space.

Discrete Distribution Function: For the random variable $X = \{x_1, x_2, x_3, \dots\}$. The cumulative distribution function is given by $F(x) = P(X \leq x) = \sum_{x_i \leq x} p(x_i)$.

Mean/Expected Value, Variance and Standard Deviation of DRV:

The **mean or expected value** of a DRV X is defined as

$$E(X) = \mu = \sum x_i P(X = x_i) = \sum p(x_i)x_i .$$

The **variance** of a DRV X is defined as

$$Var(X) = \sigma^2 = \sum P(X = x_i) (x_i - \mu)^2 = \sum p_i (x_i - \mu)^2 = \sum p_i (x_i^2 - \mu^2) .$$

The **standard deviation** of DRV X is defined as

$$SD(X) = \sigma = \sqrt{\sigma^2} = \sqrt{Var(X)} .$$



Continuous Random Variable: A continuous random variable is not defined at specific values. Instead, it is defined over an interval of values, and is represented by the area under a curve. Thus, a random variable X is said to be continuous if it can take all possible values between certain limits. In other words, a random variable is said to be continuous when its different values cannot be put in 1-1 correspondence with a set of positive integers. Here, the probability of observing any single value is equal to zero, since the number of values which may be assumed by the random variable is infinite.

A continuous random variable is a random variable that (at least conceptually) can be measured to any desired degree of accuracy.

Examples of Continuous Random Variable:

- (i) Rainfall in a particular area can be treated as CRV.
- (ii) Age, height and weight related problems can be included under CRV.
- (iii) The amount of sugar in an orange is a CRV.
- (iv) The time required to run a mile is a CRV.

Important Remark: In case of DRV, the probability at a point i.e., $P(x = c)$ is not zero for some fixed c . However, in case of CRV the probability at a point is always zero.

i.e., $P(x = c) = 0$ for all possible values of c .

Probability Density Function: The probability density function (p.d.f) of a random variable X usually denoted by $f_x(x)$ or simply by $f(x)$ has the following obvious properties:

- i) $f(x) \geq 0, -\infty < x < \infty$
- ii) $\int_{-\infty}^{\infty} f(x)dx = 1$
- iii) The probability $P(E)$ given by $P(E) = \int f(x)dx$ is well defined for any event E .

If $f(x)$ is the p.d.f of x , then the probability that x belongs to A , where A is some interval (a, b) is given by the integral of $f(x)$ over that interval.

$$\text{i.e., } P(X \in A) = \int_a^b f(x)dx$$

Cumulative Density Function: Cumulative density function of a continuous random variable is defined as $F(x) = \int_{-\infty}^x f(t)dt$ for $-\infty < x < \infty$.

Mean/Expectation, Variance and Standard deviation of CRV:

The mean or expected value of a CRV X is defined as $\mu = E(X) = \int_{-\infty}^{\infty} x f(x)dx$

The variance of a CRV X is defined as $Var(X) = \sigma^2 = \int_{-\infty}^{\infty} x^2 f(x)dx - \mu^2$

The standard deviation of a CRV X is given by $= \sqrt{Var(X)}$.

Examples:

1. The probability density function of a discrete random variable X is given below:

x	0	1	2	3	4	5	6
$P(X=x) = f(x)$	K	$3k$	$5k$	$7k$	$9k$	$11k$	$13k$

Find (i) k ; (ii) $F(4)$; (iii) $P(X \geq 5)$; (iv) $P(2 \leq X < 5)$; (v) $E(X)$ and (vi) $\text{Var}(X)$.

Solution: To find the value of k , consider the sum of all the probabilities which equals to $49k$. Equating this to 1, $k = 1/49$. Therefore, distribution of X may now be written as

x	0	1	2	3	4	5	6
$P(X=x) = f(x)$	$1/49$	$3/49$	$5/49$	$7/49$	$9/49$	$11/49$	$13/49$

$$F(4) = P[X \leq 4] = P[X = 0] + P[X = 1] + P[X = 2] + P[X = 3] + P[X = 4] = \frac{25}{49}.$$

$$P[X \geq 5] = P[X = 5] + P[X = 6] = \frac{24}{49}.$$

$$P[2 \leq X < 5] = P[X = 2] + P[X = 3] + P[X = 4] = \frac{21}{49}.$$

Next to find $E(X)$, consider

$$E(X) = \sum_i x_i * f(x_i) = \frac{203}{49}.$$

To obtain Variance, it is necessary to compute

$$E(X^2) = \sum_i x_i^2 * f(x_i) = \frac{973}{49}.$$

Thus, Variance of X is obtained by using the relation,

$$\text{Var}(X) = E(X^2) - [E(X)]^2 = \frac{973}{49} - \left(\frac{203}{49}\right)^2.$$

2. A random variable, X , has the following distribution function.

X	-2	-1	0	1	2	3
$f(x_i)$	0.1	k	0.2	$2k$	0.3	k

Find (i) k ; (ii) $F(2)$; (iii) $P(-2 < X < 2)$; (iv) $P(-1 < X \leq 2)$; (v) $E(X)$ and (vi) Variance.

Solution: Consider the result, namely, sum of all the probabilities equals 1,

$0.1 + k + 0.2 + 2k + 0.3 + k = 1$ yields $k = 0.1$. In view of this, distribution function of X may be formulated as

X	-2	-1	0	1	2	3
$f(x_i)$	0.1	0.1	0.2	0.2	0.3	0.1

Note that $F(2) = P[X \leq 2]$

$$= P[X = -2] + P[X = -1] + P[X = 0] + P[X = 1] + P[X = 2] \\ = 0.9. \text{ The same also be obtained using the result,}$$

$$F(2) = P[X \leq 2] = 1 - P[X < 1] = 1 - \{P[X = -2] + P[X = -1] + P[X = 0]\} = 0.6.$$

$$\text{Next, } P(-2 < X < 2) = P[X = -1] + P[X = 0] + P[X = 1] = 0.5.$$

$$\text{Clearly, } P(-1 < X \leq 2) = 0.7.$$

$$\text{Now, consider } (X) = \sum_i x_i * f(x_i) = 0.8.$$

$$\text{Then } E(X^2) = \sum_i x_i^2 * f(x_i) = 2.8. \text{ } Var(X) = E(X^2) - \{E(X)\}^2 = 2.8 - 0.64 = 2.16.$$

3. A shipment of 20 similar laptop computers to a retail outlet contains 3 that are defective. If a school makes a random purchase of 2 of these computers, find the probability distribution for the number of defectives.

Solution: Let X be a random variable whose values x are the possible numbers of defective computers purchased by the school. Then x can only take the numbers 0, 1, and 2. Now

$$f(0) = P(X = 0) = \frac{\binom{3}{0}\binom{17}{2}}{\binom{20}{2}} = \frac{68}{95}, \quad f(1) = P(X = 1) = \frac{\binom{3}{1}\binom{17}{1}}{\binom{20}{2}} = \frac{51}{190}$$

$$f(2) = P(X = 2) = \frac{\binom{3}{2}\binom{17}{0}}{\binom{20}{2}} = \frac{3}{190}.$$

Thus, the probability distribution of X is

x	0	1	2
$f(x)$	68/95	51/190	3/190

4. If a car agency sells 50% of its inventory of a certain foreign car equipped with side airbags, find a formula for the probability distribution of the number of cars with side airbags among the next 4 cars sold by the agency.

Solution: Since the probability of selling an automobile with side airbags is 0.5, the $2^4 = 16$ points in the sample space are equally likely to occur. Therefore, the denominator for all probabilities, and also for our function, is 16. To obtain the number of ways of selling 3 cars with side airbags, it is required to consider the number of ways of partitioning 4 outcomes into two cells, with 3 cars with side airbags assigned to one cell and the model without side airbags assigned to the other. This can be done in $\binom{4}{3} = 4$ ways. In general, the event of

selling x models with side airbags and $4 - x$ models without side airbags can occur in $\binom{4}{x}$ ways, where x can be 0, 1, 2, 3, or 4. Thus, the probability distribution $f(x) = P(X = x)$ is

$$f(x) = \binom{1}{16} \binom{4}{x} \text{ for } x = 0, 1, 2, 3, 4.$$

5. The diameter of an electric cable, say X , is assumed to be a continuous random variable

with p.d.f $f(x) = \begin{cases} 6x(1-x) & 0 \leq x \leq 1 \\ 0 & \text{otherwise} \end{cases}$

(i) Check that above is p.d.f.

(ii) Find $P\left(\frac{2}{3} < x < 1\right)$

(iii) Determine a number b such that $P(X < b) = P(X > b)$.

Solution: (i) $f(x) \geq 0$ in the given interval.

$$\begin{aligned} \int_{-\infty}^{\infty} f(x)dx &= \int_{-\infty}^0 f(x)dx + \int_0^1 f(x)dx + \int_1^{\infty} f(x)dx \\ &= 0 + \int_0^1 6x(1-x) dx + 0 \\ &= \left\{ \frac{6x^2}{2} - \frac{6x^3}{3} \right\} \text{ by putting limits } x = 0 \text{ to } 1 \\ &= 1 \end{aligned}$$

(ii) $P\left(\frac{2}{3} < x < 1\right) = \int_{2/3}^1 f(x)dx = \int_{2/3}^1 (6x - 6x^2)dx = \frac{7}{27}$.

(iii) $P(X < b) = P(X > b)$

$$\begin{aligned} \int_0^b f(x)dx &= \int_b^1 f(x)dx \\ 6 \int_0^b x(1-x)dx &= 6 \int_b^1 x(1-x)dx \\ \left(\frac{b^2}{2} - \frac{b^3}{3} \right) &= \left[\left(\frac{1}{2} - \frac{1}{3} \right) - \left(\frac{b^2}{2} - \frac{b^3}{3} \right) \right] \\ 3b^2 - 2b^3 &= [1 - 3b^2 + 2b^3] \\ 4b^3 - 6b^2 + 1 &= 0 \\ (2b - 1)(2b^2 - 2b - 1) &= 0 \end{aligned}$$

From this $b = \frac{1}{2}$ is the only real value lying between 0 and 1 and satisfying the given condition.

6. Suppose that the error in the reaction temperature, in $^{\circ}\text{C}$, for a controlled laboratory experiment is a continuous random variable X having the probability density function

$$f(x) = \begin{cases} \frac{x^2}{3}, & -1 < x < 2 \\ 0, & \text{elsewhere} \end{cases}$$

(i) Verify that $f(x)$ is a probability density function.

(ii) Find $P(0 < X \leq 1)$.

Solution: a) $\int_{-\infty}^{\infty} f(x)dx = \int_{-1}^2 \frac{x^2}{3} dx = 1$. Hence the given function is a p.d.f.

$$\text{b) } P(0 < X \leq 1) = \int_0^1 \frac{x^2}{3} dx = \frac{1}{9}.$$

7. The length of time (in minutes) that a certain lady speaks on telephone is found to be a

random variable with probability function $f(x) = \begin{cases} Ae^{-\frac{x}{5}} & \text{for } x \geq 0 \\ 0 & \text{otherwise} \end{cases}$

(i) Find A

(ii) Find the probability that she will speak on the phone

- (a) more than 10 min (b) less than 5 min (c) between 5 & 10 min.

Solution: (i) Given $f(x)$ is p.d.f. i.e., $\int_{-\infty}^{\infty} f(x)dx = 1$

$$\int_{-\infty}^0 f(x)dx + \int_0^{\infty} f(x)dx = 1$$

$$\xrightarrow{\text{yields}} 0 + \int_0^{\infty} Ae^{-\frac{x}{5}} dx = 1$$

$$\xrightarrow{\text{yields}} A = \frac{1}{5}$$

$$(ii) (a) P(x > 10) = \int_{10}^{\infty} f(x)dx = \int_{10}^{\infty} \frac{1}{5} e^{-\frac{x}{5}} dx = e^{-2} = 0.1353$$

$$(b) P(x < 5) = \int_{-\infty}^5 f(x)dx = \int_0^5 \frac{1}{5} e^{-\frac{x}{5}} dx = -e^{-1} + 1 = 0.6322$$

$$(c) P(5 < x < 10) = \int_5^{10} f(x)dx = \int_5^{10} \frac{1}{5} e^{-\frac{x}{5}} dx = -e^{-2} + e^{-1} = 0.2325 .$$

8. Suppose X is a continuous random variable with the following probability density function $f(x) = 3x^2$ for $0 < x < 1$. Find the mean and variance of X.

Solution: Mean = $\mu = \int_{-\infty}^{\infty} xf(x)dx$

$$= \int_{-\infty}^0 xf(x)dx + \int_0^1 xf(x)dx + \int_1^{\infty} xf(x)dx$$

$$= 0 + \int_0^1 x * 3x^2 dx + 0 = \int_0^1 3x^3 dx = \frac{3}{4} .$$

$$\text{Variance} = \sigma^2 = \int_{-\infty}^{\infty} x^2 f(x)dx - \mu^2$$

$$= \int_0^1 x^2 f(x)dx - \mu^2$$

$$= \int_0^1 x^2 * 3x^2 dx - \left(\frac{3}{4}\right)^2$$

$$= \int_0^1 3x^4 dx - \left(\frac{3}{4}\right)^2 = \frac{3}{80} .$$

Exercise:

- Two cards are drawn randomly, simultaneously from a well shuffled deck of 52 cards. Find the variance for the number of aces.

2. If X is a discrete random variable taking values $1, 2, 3, \dots$ with $P(x) = \frac{1}{2} \left(\frac{2}{3}\right)^x$. Find $P(X$ being an odd number) by first establishing that $P(x)$ is a probability function.
3. The probability mass function of a random variable X is zero except the points $x = 0, 1, 2$. At these points it has the values $p(0) = 3c^3, p(1) = 4c - 10c^2$ and $p(2) = 5c - 1$ for some $c > 0$.
 - a) Determine the value of c .
 - b) Compute the probabilities $P(X < 2)$ and $P(1 < X \leq 2)$.
 - c) Find the largest x such that $F(x) < \frac{1}{2}$.
 - d) Find the smallest x such that $F(x) \geq \frac{1}{3}$.
4. If X is a random variable with $P(X = x) = \frac{1}{2^x}$, where $x = 1, 2, 3, \dots \infty$.
Find i) $P(X)$ (ii) $P(X = \text{even})$ (iii) $P(X = \text{divisible by } 3)$.
5. A continuous random variable has the density function

$$f(x) = \begin{cases} kx^2 & -3 < x < 3 \\ 0 & \text{otherwise} \end{cases}$$
 Find k and hence find $P(x < 3), P(x > 1)$.
6. Let X be a continuous random variable with p.d.f.

$$f(x) = \begin{cases} ax, & 0 \leq x \leq 1 \\ a, & 1 \leq x \leq 2 \\ -ax + 3a, & 2 \leq x \leq 3 \\ 0, & \text{otherwise} \end{cases}$$
 (i) Determine the constant. (ii) Compute $P(X \leq 1.5)$.

7. Find the mean and variance of the probability density function $f(x) = \frac{1}{2}e^{-|x|}$
8. A continuous distribution of a variable X in the range $(-3, 3)$ is defined by

$$f(x) = \begin{cases} \frac{1}{16}(3+x)^2, & -3 \leq x \leq -1 \\ \frac{1}{16}(6-2x^2), & -1 \leq x \leq 1 \\ \frac{1}{16}(3-x)^2, & 1 \leq x \leq 3 \end{cases}$$

- (i) Verify that the area under the curve is unity.
- (ii) Find the mean and variance of the above distribution.

Answers: 1) 0.1392 2) 3/5 3) 1/3, 1/3, 2/3, 1, 1 4) 1, 1/3, 1/7 5) 1/18, 1, 13/27
6) 1/2, 1/2 7) Mean = 0 and Variance = 2 8) Unit area and 0, 1.

JOINT PROBABILITY

Two or more random variables:

So far, only single random variables were considered. If one chooses a person at random and measures his or her height and weight, each measurement is a random variable – but taller people also tend to be heavier than shorter people, so the outcomes will be related. In order to deal with such probabilities, joint probability distribution of two random variables are studied in detail.

Joint Probability distribution for discrete random variables

Joint Probability Mass Function:

Let X and Y be random variables on the same sample space S with respective range spaces $R_X = \{x_1, x_2, \dots, x_n\}$ and $R_Y = \{y_1, y_2, \dots, y_m\}$. The joint distribution or joint probability function of X and Y is the function h on the product space $R_X \times R_Y$ defined by

$$h(x_i, y_j) \equiv P(X = x_i, Y = y_j) \equiv P(\{s \in S : X(s) = x_i, Y(s) = y_j\})$$

The function h has the following properties:

- (i) $h(x_i, y_j) \geq 0$
- (ii) $\sum_i \sum_j h(x_i, y_j) = 1$

Thus, h defines a probability space on the product space $R_X \times R_Y$.

$X \backslash Y$	y_1	y_1	...	y_j	...	y_m	$\sum_i y_i$
x_1	$h(x_1, y_1)$	$h(x_1, y_2)$...	$h(x_1, y_j)$...	$h(x_1, y_m)$	$f(x_1)$
x_2	$h(x_2, y_1)$	$h(x_2, y_2)$...	$h(x_2, y_j)$...	$h(x_2, y_m)$	$f(x_2)$
\vdots	\vdots	\vdots	\vdots	\vdots	\vdots	\vdots	\vdots
x_i	$h(x_i, y_1)$	$h(x_i, y_2)$...	$h(x_i, y_j)$...	$h(x_i, y_m)$	$f(x_i)$
\vdots	\vdots	\vdots	\vdots	\vdots	\vdots	\vdots	\vdots
x_m	$h(x_n, y_1)$	$h(x_n, y_2)$...	$h(x_n, y_j)$...	$h(x_n, y_m)$	$f(x_n)$
$\sum_i x_i$	$g(y_1)$	$g(y_2)$...	$g(y_j)$...	$g(y_m)$	

The functions f and g on the right side and the bottom side, respectively, of the joint distribution table are defined by

$$f(x_i) = \sum_j h(x_i, y_j) \text{ and } g(y_j) = \sum_i h(x_i, y_j).$$

That is, $f(x_i)$ is the sum of the entries in the i^{th} row and $g(y_j)$ is the sum of the entries in the j^{th} column. They are called the marginal distributions of X and Y , respectively.

Expectation: Consider a function $\varphi(X, Y)$ of X and Y . Then the function

$$E\{\varphi(X, Y)\} = \sum_i \sum_j h(x_i, y_j) \varphi(x_i, y_j)$$

is called the mathematical expectation of $\varphi(X, Y)$ in the joint distribution of X and Y .

Co-variance and Correlation: Let X and Y be random variables with the joint distribution $h(x, y)$, and respective means μ_X and μ_Y . The covariance of X and Y , is denoted by $cov(X, Y)$ and is defined as

$$\begin{aligned} cov(X, Y) &= \sum_{i,j} (x_i - \mu_X)(y_j - \mu_Y) h(x_i, y_j) \\ cov(X, Y) &= \sum_{i,j} x_i y_j h(x_i, y_j) - \mu_X \mu_Y \end{aligned}$$

The correlation of X and Y is defined by

$$\rho(X, Y) = \frac{cov(X, Y)}{\sigma_X \sigma_Y}$$

The correlation ρ is dimensionless and has the following properties:

- (i) $\rho(X, Y) = \rho(Y, X)$,
- (ii) $-1 \leq \rho \leq 1$,
- (iii) $\rho(X, X) = 1$, $\rho(X, -X) = -1$,
- (iv) $\rho(aX + b, cY + d) = \rho(X, Y)$ if $a, c \neq 0$.

Conditional probability distribution:

We know that the value x of the random variable X represents an event that is a subset of the sample space. If we use the definition of conditional probability,

$$P(B|A) = \frac{P(A \cap B)}{P(A)}, \text{ provided } P(A) > 0,$$

where A and B are now the events defined by $X = x$ and $Y = y$, respectively, then

$$P(Y = y|X = x) = \frac{P(X=x, Y=y)}{P(X=x)} = \frac{h(x, y)}{f(x)}, \text{ provided } f(x) > 0,$$

Where X and Y are discrete random variables.

Problem 1. A coin is tossed three times. Let X be equal to 0 or 1 according as a head or a tail occurs on the first toss. Let Y be equal to the total number of heads which occurs. Determine (i) the marginal distributions of X and Y , and (ii) the joint distribution of X and Y , (iii) expected values of $X, Y, X + Y$ and XY , (iv) σ_X and σ_Y , (v) $Cov(X, Y)$ and $\rho(X, Y)$.

Solution: Here the sample space is given by

$$S = \{HHH, HHT, HTH, HTT, THH, THT, TTH, TTT\}$$

(i) The distribution of the random variable X is given by the following table

X (First toss Head or Tail)	0 (First toss Head)	1 (First toss Tail)
P(X) (Probability of random variable X)	$\frac{4}{8}$	$\frac{4}{8}$

which is the marginal distribution of the random variable X .

The distribution of the random variable Y is given by the following table

Y (Total number of Heads)	0 (zero Heads)	1 (one Head)	2 (two Head)	3 (three Head)
P(Y) (Probability of random variable Y)	$\frac{1}{8}$	$\frac{3}{8}$	$\frac{3}{8}$	$\frac{1}{8}$

which is the marginal distribution of the random variable Y .

(ii) The joint distribution of the random variables X and Y is given by the following table

X \ Y	0 (zero Heads)	1 (one Head)	2 (two Head)	3 (three Head)
0 (First toss Head)	0	$\frac{1}{8}$	$\frac{2}{8}$	$\frac{1}{8}$
1 (First toss Tail)	$\frac{1}{8}$	$\frac{2}{8}$	$\frac{1}{8}$	0

$$(iii) E[X] = \mu_X = \sum x_i P(x_i) = x_1 P(x_1) + x_2 P(x_2) = 0 \times \frac{4}{8} + 1 \times \frac{4}{8} = \frac{4}{8}$$

$$E[Y] = \mu_Y = \sum y_j P(y_j) = y_1 P(y_1) + y_2 P(y_2) + y_3 P(y_3)$$

$$= 0 \times \frac{1}{8} + 1 \times \frac{3}{8} + 2 \times \frac{3}{8} + 3 \times \frac{1}{8} = \frac{12}{8}$$

$$\begin{aligned} E[X+Y] &= \sum \sum P_{ij}(x_i + y_j) \\ &= P_{11}(x_1 + y_1) + P_{12}(x_1 + y_2) + P_{13}(x_1 + y_3) + P_{14}(x_1 + y_4) + P_{21}(x_2 + y_1) \\ &\quad + P_{22}(x_2 + y_2) + P_{23}(x_2 + y_3) + P_{24}(x_2 + y_4) \\ &= 0(0+0) + \frac{1}{8}(0+1) + \frac{2}{8}(0+2) + \frac{2}{8}(1+1) + \frac{1}{8}(1+2) + 0(1+3) = \frac{16}{8} \\ &= 2. \end{aligned}$$

$$\begin{aligned} E[XY] &= \sum \sum P_{ij}(x_i y_j) \\ &= P_{11}(x_1 y_1) + P_{12}(x_1 y_2) + P_{13}(x_1 y_3) + P_{14}(x_1 y_4) + P_{21}(x_2 y_1) + P_{22}(x_2 y_2) \\ &\quad + P_{23}(x_2 y_3) + P_{24}(x_2 y_4) \\ &= 0(0 \times 0) + \frac{1}{8}(0 \times 1) + \frac{2}{8}(0 \times 2) + \frac{2}{8}(1 \times 1) + \frac{1}{8}(1 \times 2) + 0(1 \times 3) = 2. \end{aligned}$$

$$(iv) \sigma_X^2 = E[X^2] - \mu_X^2 = \sum x_i^2 P(x_i) - [E(X)]^2 = x_1^2 P(x_1) + x_2^2 P(x_2) = 0^2 \times \frac{4}{8} + 1^2 \times \frac{4}{8} - \left(\frac{4}{8}\right)^2 = \frac{1}{4}$$

$$\begin{aligned} \sigma_Y^2 &= E[Y^2] - \mu_Y^2 \\ &= \sum y_i^2 P(y_i) - [E(Y)]^2 = y_1^2 P(y_1) + y_2^2 P(y_2) + y_3^2 P(y_3) + y_4^2 P(y_4) \\ &= 0^2 \times \frac{1}{8} + 1^2 \times \frac{3}{8} + 2^2 \times \frac{3}{8} + 3^2 \times \frac{1}{8} - \left(\frac{2}{8}\right)^2 = \frac{3}{4} \end{aligned}$$

$$(v) \text{Cov}(X, Y) = E(XY) - \mu_X \mu_Y = \frac{1}{2} - \frac{1}{2} \times \frac{3}{2} = -\frac{1}{4}$$

$$\rho(X, Y) = \frac{\text{Cov}(X, Y)}{\sigma_X \sigma_Y} = \frac{-1/4}{(1/2)(\sqrt{3}/2)} = -\frac{1}{\sqrt{3}}.$$

Problem 2: The joint distribution of two random variables X and Y is given by the following table:

		Y	2	3	4
		X	1	2	3
X	1	0.06	0.15	0.09	
	2	0.14	0.35	0.21	

Determine the individual distributions of X and Y. Also, verify that X and Y are stochastically independent.

Solution: X takes values 1, 2 and Y takes the values 2, 3, 4. Also, $h_{11} = 0.06$, $h_{12} = 0.15$

$$h_{13} = 0.09, h_{21} = 0.14, h_{22} = 0.35, h_{23} = 0.21$$

$$\text{Therefore, } f_1 = h_{11} + h_{12} + h_{13} = 0.3, f_2 = h_{21} + h_{22} + h_{23} = 0.7,$$

$$g_1 = h_{11} + h_{21} = 0.2, g_2 = h_{12} + h_{22} = 0.5, g_3 = h_{13} + h_{23} = 0.3.$$

The distribution of X is given by

x_i	1	2
f_i	0.3	0.7

The distribution of Y is given by

y_j	2	3	4
g_j	0.2	0.5	0.3

$$f_1g_1 = 0.06 = h_{11}, \quad f_1g_2 = 0.15 = h_{12}, \quad f_1g_3 = 0.09 = h_{13},$$

$$f_2g_1 = 0.14 = h_{21}, \quad f_2g_2 = 0.35 = h_{22}, \quad f_2g_3 = 0.21 = h_{23},$$

Thus, $f_i g_j = h_{ij}$ for all values of i and j so, X and Y are stochastically independent.

Problem 3. The joint distribution of two random variables X and Y is given by the following table:

		Y	
		0	1
X	0	0.1	0.2
	1	0.4	0.2
2	0.1	0	

- (a) Find $P(X + Y > 1)$
- (b) Determine the individual (marginal) probability distributions of X and Y and verify that X and Y are not independent.
- (c) Find $P(X = 2|Y = 0)$.
- (d) Find the conditional distribution of X given $Y = 1$.

Solution: Note that X takes the values 0, 1, 2 and Y takes the values 0, 1

$$h_{11} = 0.1, h_{12} = 0.2, h_{21} = 0.4, h_{22} = 0.2, h_{31} = 0.1, h_{32} = 0,$$

(a) The event $X + Y > 1$ occurs only when the pair (X, Y) takes the values (1,1), (2,0) and (2,1). The probability that this event occurs is therefore

$$P(X + Y > 1) = h_{22} + h_{31} + h_{32} = 0.2 + 0.1 + 0 = 0.3.$$

$$(b) f_1 = h_{11} + h_{12} = 0.1 + 0.2 = 0.3.$$

$$f_2 = h_{21} + h_{22} = 0.4 + 0.2 = 0.6.$$

$$f_3 = h_{31} + h_{32} = 0.1 + 0 = 0.1.$$

$$g_1 = h_{11} + h_{21} + h_{31} = 0.6$$

$$g_2 = h_{12} + h_{22} + h_{32} = 0.4.$$

The distribution of X is given by

x_i	0	1	2
f_i	0.6	0.6	0.1

The distribution of Y is given by

y_j	0	1
g_j	0.6	0.4

It is verified that $f_1 g_1 = 0.18 \neq h_{11}$.

Therefore, X and Y are not stochastically independent.

$$(c) P(X = 2|Y = 0) = \frac{h(2,0)}{g(0)} = \frac{h_{31}}{g_1} = \frac{0.1}{0.6} = \frac{1}{6}$$

(d) Conditional distribution of X given $Y = 1$ is

$$P(X = x|Y = 1) = \frac{h(x, 1)}{g(1)} = \frac{h_{i2}}{g_2}$$

$$P(X = 0|Y = 1) = \frac{h(0,1)}{g(1)} = \frac{h_{12}}{g_2} = \frac{0.2}{0.4} = 0.5$$

$$P(X = 1|Y = 1) = \frac{h(1,1)}{g(1)} = \frac{h_{22}}{g_2} = \frac{0.2}{0.4} = 0.5$$

$$P(X = 2|Y = 1) = \frac{h(2,1)}{g(1)} = \frac{h_{32}}{g_2} = \frac{0}{0.4} = 0$$

Problem 4. The joint distribution of two random variables X and Y is given by $p_{ij} = k(i+j)$, $i = 1, 2, 3, 4; j = 1, 2, 3$. Find (i) k and (ii) the marginal distributions of X and Y .

Show that X and Y are not independent.

Solution: For the given p_{ij} ,

$$\sum_i \sum_j h_{ij} = \sum_{i=1}^4 \sum_{j=1}^3 h = k \sum_{i=1}^4 \sum_{j=1}^3 (i+j)$$

$$\begin{aligned}
 &= k \sum_{i=1}^4 \{(i+1) + (i+2) + (i+3)\} = k \sum_{i=1}^4 (3i+6) \\
 &= k \{(3+6) + (3 \times 2+6) + (3 \times 3+6) + (3 \times 4+6)\} = 54k
 \end{aligned}$$

Since

$$\sum_i \sum_j h_{ij} = 1, i.e., 54k = 1 \text{ or } k = 1/54$$

$$f_i = \sum_j h_{ij} = \sum_{j=1}^3 h_{ij} = k \sum_{j=1}^3 (i+j) = \frac{i+2}{18}$$

$$g_j = \sum_i h_{ij} = \sum_{i=1}^4 h_{ij} = k \sum_{i=1}^4 (i+j) = \frac{2j+5}{27}$$

Therefore, the marginal distributions of X and Y are

$$\{f_i\} = \left\{ \frac{i+2}{18} \right\}, i=1,2,3,4 \text{ and } \{g_j\} = \left\{ \frac{2j+5}{27} \right\}, j=1,2,3.$$

Finally note that $f_i \ g_j \neq h_{11}$, so X and Y are not independent.

Problem 5. The joint probability distribution of two random variables X and Y is given by the following table.

X \ Y	1	3	9
2	1/8	1/24	1/12
4	1/4	1/4	0
6	1/8	1/24	1/12

Find the marginal distribution of X and Y, and evaluate $\text{cov}(X, Y)$. Find $P(X = 4|Y = 3)$ and

$$P(Y = 3|X = 4)$$

Solution: From the table, note that

$$f_1 = \frac{1}{8} + \frac{1}{24} + \frac{1}{12} = \frac{1}{4}$$

$$f_2 = \frac{1}{4} + \frac{1}{4} + 0 = \frac{1}{2}$$

$$f_3 = \frac{1}{8} + \frac{1}{24} + \frac{1}{12} = \frac{1}{4}$$

$$g_1 = \frac{1}{8} + \frac{1}{4} + \frac{1}{8} = \frac{1}{2}$$

$$g_2 = \frac{1}{24} + \frac{1}{4} + \frac{1}{24} = \frac{1}{3}$$

$$g_3 = \frac{1}{12} + 0 + \frac{1}{12} = \frac{1}{6}$$

The marginal distribution of X is given by the table:

x_i	2	4	6
f_i	1/4	1/2	1/4

And the marginal distribution of Y is given by the table:

y_j	1	3	9
g_j	1/2	1/3	1/6

Therefore, the means of these distributions are respectively,

$$\mu_X = \sum x_i P(x_i) = \left(2 \times \frac{1}{4}\right) + \left(4 \times \frac{1}{2}\right) + \left(6 \times \frac{1}{4}\right) = 4$$

$$\mu_Y = \sum y_j P(y_j) = \left(1 \times \frac{1}{2}\right) + \left(3 \times \frac{1}{3}\right) + \left(9 \times \frac{1}{6}\right) = 3$$

$$\begin{aligned}
 E[XY] &= \sum_i \sum_j h_{ij} x_i y_j \\
 &= (h_{11}x_1y_1 + h_{12}x_1y_2 + h_{13}x_1y_3) + (h_{21}x_2y_1 + h_{22}x_2y_2 + h_{23}x_2y_3) \\
 &\quad + (h_{31}x_3y_1 + h_{32}x_3y_2 + h_{33}x_3y_3) \\
 &= \left(2 \times \frac{1}{8}\right) + \left(6 \times \frac{1}{24}\right) + \left(18 \times \frac{1}{12}\right) + \left(4 \times \frac{1}{4}\right) + \left(12 \times \frac{1}{4}\right) + 36 \times 0 + \left(6 \times \frac{1}{8}\right) + \\
 &\quad \left(18 \times \frac{1}{24}\right) + \left(54 \times \frac{1}{12}\right) \\
 &= 2 + 4 + 6 = 12
 \end{aligned}$$

$$Cov(X, Y) = E[XY] - \mu_X \mu_Y = 12 - 12 = 0$$

$$\rho(X, Y) = 0.$$

$$P(X = 4|Y = 3) = \frac{h(4,3)}{g(3)} = \frac{h_{22}}{g_2} = \frac{\frac{1}{4}}{\frac{1}{3}} = \frac{3}{4}$$

$$P(Y = 3|X = 4) = \frac{h(4,3)}{f(4)} = \frac{h_{22}}{f_2} = \frac{\frac{1}{4}}{\frac{1}{2}} = \frac{2}{4} = 0.5$$

Problems to practice:

- 1) The joint probability distribution of two random variables X and Y is given by the following table.

	Y	-2	-1	4	5
X					
1		0.1	0.2	0	0.3
2		0.2	0.1	0.1	0

- (a) Find the marginal distribution of X and Y, and evaluate $\text{cov}(X, Y)$.
- (b) Also determine whether μ_X and μ_Y .
- (c) Find $P(Y = -1|X = 1)$ and $P(X = 2|Y = 4)$
- 2) Two textbooks are selected at random from a shelf containing three statistics texts, two mathematics texts and three engineering texts. Denoting the number of books selected in each subject by S, M and E respectively, find (a) the joint distribution of S and M, (b) the marginal distributions of S, M and E, and (c) Find the correlation of the random variables S and M.
- 3) Consider an experiment that consists of 2 throws of a fair die. Let X be the number of 4s and Y be the number of 5s obtained in the two throws. Find the joint probability distribution of X and Y. Also evaluate $P(2X + Y < 3)$.

Joint Probability distribution for continuous random variables:

Let x and y be two continuous random variables. Suppose there exists a real valued function $h(x, y)$ of x and y such that the following conditions hold:

- (i) $h(x, y) \geq 0$ for all x, y
- (ii) $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x, y) dx dy$ exists and is equal to 1.

Then, $h(x, y)$ is called joint probability density function.

If $[a, b]$ and $[c, d]$ are any two intervals, then the probability that $x \in [a, b]$ and $y \in [c, d]$, is denoted by $P(a \leq x \leq b, c \leq y \leq d)$ is defined by the formula

$$P(a \leq x \leq b, c \leq y \leq d) = \int_a^b \int_c^d h(x, y) dy dx.$$

For any specified real numbers u, v, the function

$$F(u, v) = \int_{-\infty}^u \int_{-\infty}^v h(x, y) dy dx$$

is called the joint or the compound cumulative distribution function.

Where $F(u, v) = P(-\infty < x \leq u, -\infty < y \leq v)$ and $\frac{\partial^2 F}{\partial u \partial v} = p(u, v)$.

Further, the function $h_1(x) = \int_{-\infty}^{\infty} h(x, y) dy$ is called marginal density function of x , and the function $h_2(y) = \int_{-\infty}^{\infty} h(x, y) dx$ is called marginal density function of y . $h_1(x)$ is the density function of x and $h_2(y)$ is the density function of y .

The variables x and y are said to stochastically independent if $h_1(x)h_2(y) = h(x, y)$.

If $\varphi(x, y)$ is a function of x and y , then the expectation of $\varphi(x, y)$ is defined by

$$E\{\varphi(x, y)\} = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} \varphi(x, y) h(x, y) dx dy.$$

The covariance between x and y is defined as

$$Cov(x, y) = E\{xy\} - E\{x\} E\{y\}.$$

Conditional Probability:

The idea of conditional probability function of discrete random variables of extended to the case of continuous random variables.

If X and Y are continuous random variables, then the conditional probability distribution Y given X is $f(y|x) = \frac{h(x,y)}{h_1(x)}$

where $h(x, y)$ is the joint density function of X and Y , $h_1(x)$ is the marginal density function of X .

$$P(c < Y < d | a < x < b) = \frac{P(a < X < b, c < Y < d)}{P(a < x < b)}$$

Problem 6: Find the constant ' k ' so that

$$h(x, y) = \begin{cases} k(x+1)e^{-y}, & 0 < x < 1, y > 0 \\ 0, & \text{elsewhere} \end{cases}$$

is a joint probability density function. Are x and y independent?

Solution: Observe that $h(x, y) \geq 0$ for x, y , if $k \geq 0$

$$\begin{aligned} \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x, y) dx dy &= \int_{y=0}^{\infty} \int_{x=0}^1 h(x, y) dx dy \\ &= k \left\{ \int_0^1 (x+1) dx \right\} \left\{ \int_0^{\infty} e^{-y} dy \right\} \\ &= k \left\{ \frac{2^2 - 1^2}{2} \right\} \{0 + 1\} = \frac{3}{2} k. \end{aligned}$$

Hence $\int_{-\infty}^{\infty} \int_{-\infty}^{\infty} h(x, y) dx dy = 1$ if $k = \frac{2}{3}$.

Therefore, $h(x, y)$ is a joint probability density function if $k = \frac{2}{3}$.

With $k = \frac{2}{3}$, the marginal density functions are

$$\begin{aligned} h_1(x) &= \int_{-\infty}^{\infty} h(x, y) dy, \quad 0 < x < 1 \\ &= \frac{2}{3}(x+1) \int_0^{\infty} e^{-y} dy \\ &= \frac{2}{3}(x+1)(0+1). \\ &= \frac{2}{3}(x+1), \quad 0 < x < 1 \end{aligned}$$

Next,

$$\begin{aligned} h_2(y) &= \int_{-\infty}^{\infty} h(x, y) dx, \quad y > 0 \\ &= \frac{2}{3} e^{-y} \int_0^1 (x+1) dx = \frac{2}{3} e^{-y} \left\{ \frac{2^2}{2} - \frac{1}{2} \right\} \\ &= e^{-y}, \quad y > 0. \end{aligned}$$

Therefore, $h_1(x)h_2(y) = h(x, y)$ and hence x and y are stochastically independent.

Problem 7: The life time x and brightness y of a light bulb are modeled as continuous random variables with joint density function

$$h(x, y) = \alpha\beta e^{-(\alpha x + \beta y)}, \quad 0 < x < \infty, 0 < y < \infty.$$

Where α and β are appropriate constants. Find (i) the marginal density functions of x and y , and (ii) the compound cumulative distributive function.

Solution: For the given distribution, the marginal density function of x is

$$\begin{aligned} h_1(x) &= \int_{-\infty}^{\infty} h(x, y) dy = h_1(x) = \int_0^{\infty} \alpha\beta e^{-(\alpha x + \beta y)} dy \\ &= \alpha\beta e^{-\alpha x} \int_0^{\infty} e^{-\beta y} dy = \alpha e^{-\alpha x}, \quad 0 < x < \infty \end{aligned}$$

the marginal density function of y is

$$h_2(y) = \int_{-\infty}^{\infty} h(x, y) dx = \beta e^{-\beta y}, \quad 0 < y < \infty.$$

Further, the compound cumulative distribution function is

$$F(u, v) = \int_{-\infty}^u \int_{-\infty}^v h(x, y) dy dx = \int_0^u \int_0^v \alpha\beta e^{-(\alpha x + \beta y)} dy dx$$

$$\begin{aligned}
 &= \alpha\beta \left\{ \int_0^u e^{-\alpha x} dx \right\} \left\{ \int_0^v e^{-\beta y} dy \right\} \\
 &= \alpha\beta \left\{ \frac{1}{\alpha} (1 - e^{-\alpha u}) \right\} \left\{ \frac{1}{\beta} (1 - e^{-\beta v}) \right\} \\
 &= (1 - e^{-\alpha u})(1 - e^{-\beta v}), \quad 0 < u < \infty, \quad 0 < v < \infty.
 \end{aligned}$$

Problem 8: The joint probability density function of two random variables x and y is given by $h(x, y) = \begin{cases} 2, & 0 < x < y < 1 \\ 0, & \text{elsewhere} \end{cases}$

(i) Find the covariance between x and y .

Solution: The marginal density function of x is

$$h_1(x) = \int_{-\infty}^{\infty} h(x, y) dy = \begin{cases} \int_x^1 2 dy = 2(1-x), & 0 < x < 1 \\ 0 & \text{elsewhere} \end{cases}$$

The marginal density function of y is

$$\begin{aligned}
 h_2(y) &= \int_{-\infty}^{\infty} h(x, y) dx = \begin{cases} \int_0^y 2 dx = 2y, & 0 < y < 1 \\ 0 & \text{elsewhere} \end{cases} \\
 E[x] &= \int_{-\infty}^{\infty} x h_1(x) dx = \int_0^1 x \{2(1-x)\} dx \\
 &= 2 \int_0^1 (x - x^2) dx = 2 \left(\frac{1}{2} - \frac{1}{3} \right) = \frac{1}{3}, \\
 E[y] &= \int_{-\infty}^{\infty} y h_2(y) dy = \int_0^1 y(2y) dy = \frac{2}{3}, \\
 E[xy] &= \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} xy h(x, y) dx dy = \int_0^1 2y \left\{ \int_0^y x dx \right\} dy = \int_0^1 y^3 dy = \frac{1}{4}.
 \end{aligned}$$

Therefore,

$$Cov(x, y) = E[xy] - E[x]E[y] = \frac{1}{4} - \frac{1}{3} \cdot \frac{2}{3} = \frac{1}{36}.$$

Problem 9: Verify that $f(x, y) = \begin{cases} e^{-(x+y)}, & x \geq 0, y \geq 0 \\ 0, & \text{elsewhere} \end{cases}$ is a density function of a joint probability distribution. Then evaluate the following:

- (i) $P\left(\frac{1}{2} < x < 2, 0 < y < 4\right)$
- (ii) $P(x < 1)$
- (iii) $P(x > y)$
- (iv) $P(x + y \leq 1)$,
- (v) $P(0 < x < 1 | y = 2)$.

Solution: Given $f(x, y) \geq 0$

$$f(x, y) = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} f(x, y) dx dy = \int_{-\infty}^{\infty} \int_{-\infty}^{\infty} e^{-(x+y)} dx dy = \int_{-\infty}^{\infty} e^{-x} dx \int_0^{\infty} e^{-y} dy$$

$$= (0 + 1)(0 + 1) = 1.$$

Therefore, $f(x, y)$ is a density function.

$$\begin{aligned} \text{(i)} \quad P\left(\frac{1}{2} < x < 2, 0 < y < 4\right) &= \int_{1/2}^2 \int_0^4 f(x, y) dy dx = \int_{1/2}^2 \int_0^4 e^{-(x+y)} dy dx \\ &= \int_{1/2}^2 e^{-x} dx \int_0^4 e^{-y} dy = (e^{-1/2} - e^{-2})(1 - e^{-4}). \end{aligned}$$

(ii) The marginal density function of x is

$$h_1(x) = \int_{-\infty}^{\infty} f(x, y) dy = \int_0^{\infty} e^{-(x+y)} dy = e^{-x} \int_0^{\infty} e^{-y} dy = e^{-x}$$

$$\text{Therefore, } P(x < 1) = \int_0^1 h_1(x) dx = \int_0^1 e^{-x} dx = 1 - \frac{1}{e}.$$

$$\begin{aligned} \text{(iii)} \quad P(x \leq y) &= \int_0^{\infty} \left\{ \int_0^y f(x, y) dx \right\} dy = \int_0^{\infty} \left\{ \int_0^y e^{-(x+y)} dx \right\} dy \\ &= \int_0^{\infty} e^{-y} \left(\int_0^y e^{-x} dx \right) dy = \int_0^{\infty} e^{-y} (1 - e^{-y}) dy \\ &= \int_0^{\infty} (e^{-y} - e^{-2y}) dy = 1 - \frac{1}{2} = \frac{1}{2} \end{aligned}$$

$$\text{Therefore, } P(x > y) = 1 - P(x \leq y) = 1 - \frac{1}{2} = \frac{1}{2}.$$

$$\begin{aligned} \text{(iv)} \quad P(x + y \leq 1) &= \iint_A f(x, y) dA \\ &= \int_{x=0}^1 \int_{y=0}^{1-x} f(x, y) dy dx = \int_0^1 \left\{ \int_0^{1-x} e^{-(x+y)} dy \right\} dx \\ &= \int_0^1 e^{-x} \left\{ \int_0^{1-x} e^{-y} dy \right\} dx = \int_0^1 e^{-x} \{1 - e^{-(1-x)}\} dx \\ &= \int_0^1 (e^{-x} - e^{-1}) dx = 1 - \frac{2}{e}. \end{aligned}$$

$$\text{(v)} \quad P(0 < x < 1 | y = 2) = \frac{P(0 < x < 1 | y = 2)}{P(y = 2)}$$

[putting $y = 2$]

$$P(0 < x < 1 | y = 2) = \frac{\int_0^1 e^{-(x+2)} dx}{\int_0^{\infty} e^{-(x+2)} dx} = 1 - \frac{1}{e} = 0.63$$

Problems to practice:

- 1) If the joint probability function for $f(x, y)$ is

$$f(x,y) = \begin{cases} c(x^2 + y^2), & 0 \leq x \leq 1, 0 \leq y \leq 1, c \geq 0 \\ 0, & \text{elsewhere} \end{cases}$$

is a density function of a joint probability distribution. Then evaluate the following:

(i) the value of the constant c . (ii) the marginal density functions of x and y .

$$(iii) P\left(x < \frac{1}{2}, y > \frac{1}{2}\right) \quad (iv) P\left(\frac{1}{4} < x < \frac{3}{4}\right) \quad (v) P\left(y < \frac{1}{2}\right).$$

2) For the distribution given by the density function

$$f(x,y) = \begin{cases} \frac{1}{96}xy, & 0 < x < 4, 1 < y < 5 \\ 0, & \text{elsewhere} \end{cases}$$

evaluate (i) $P(1 < x < 2, 2 < y < 3)$, (ii) $P(x > 3, y \leq 2)$ (iii) $P(y \leq x)$, (iv) $(x + y \leq 3)$

3) For the distribution defined by the density function

$$f(x,y) = \begin{cases} 3xy(x+y), & 0 \leq x \leq 1, 0 \leq y \leq 1 \\ 0, & \text{elsewhere} \end{cases}$$

find the covariance between x and y .

4) For the distribution defined by the density function

$$f(x,y) = \begin{cases} \frac{1}{8}(6-x-y), & 0 < x < 2, 0 < y < 4 \\ 0, & \text{elsewhere} \end{cases}$$

Evaluate (i) $P(x < 1, y < 3)$, (ii) $P(x + y < 3)$, (iii) the covariance between x and y and (iv) $P(x < 1 | y < 3)$

Video Links:

<https://www.youtube.com/watch?v=82Ad1orN-NA>

<https://www.youtube.com/watch?v=eYthpvmqcf0>

<https://www.youtube.com/watch?v=L0zWnBrjhng>

<https://www.youtube.com/watch?v=Om68Hkd7pfw>



<https://www.youtube.com/watch?v=RYIb1u3C13I>

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