

4 Statistical Methods

4.1 Choosing the Right Test

Different metrics require different statistical tests. The choice depends on:

1. **Metric type:** Binary, continuous, count, or ordinal
2. **Data distribution:** Normal, skewed, heavy-tailed
3. **Sample size:** Large (CLT applies) vs small
4. **Variance equality:** Equal vs unequal variances

Note

Decision Tree for Test Selection

Step 1: Identify metric type

- **Binary** (0/1, yes/no): Conversion, click, bounce → Proportion test (Section 4.2)
- **Continuous** (real numbers): Revenue, time on site, cart value → t-test or Mann-Whitney (Section 4.3)
- **Count** (non-negative integers): Page views, items purchased → Poisson test or Mann-Whitney (Section 4.4)
- **Categorical** (multiple categories): Product category, exit page → Chi-square test (Section 4.5)

Step 2: Check distributional assumptions

- If continuous and approximately normal: Use t-test (parametric)
- If continuous and heavily skewed: Use Mann-Whitney U (non-parametric)
- If count data: Consider distribution (Poisson vs negative binomial)

Step 3: Select specific test

- Two variants: Two-sample test
- > 2 variants: ANOVA or Kruskal-Wallis

Implementation: See `analyze_metric()` in `statistical_analysis.py` for automatic test selection.

4.2 Two-Sample Proportion Test

Use For: Binary metrics (conversion rate, click-through rate, bounce rate)

Key References:

- Wald, A. (1943). “Tests of Statistical Hypotheses Concerning Several Parameters.” *Transactions of the American Mathematical Society*, 54(3), 426–482.

4.2.1 The Test Statistic

For proportions \hat{p}_1 (treatment) and \hat{p}_2 (control):

$$Z = \frac{\hat{p}_1 - \hat{p}_2}{\sqrt{\hat{p}(1 - \hat{p}) \left(\frac{1}{n_1} + \frac{1}{n_2} \right)}} \quad (17)$$

where $\hat{p} = \frac{x_1 + x_2}{n_1 + n_2}$ is the pooled proportion.

Under $H_0 : p_1 = p_2$, we have $Z \sim N(0, 1)$ approximately (by CLT for large samples).

Decision Rule:

- Two-tailed test: Reject H_0 if $|Z| > Z_{\alpha/2}$ (e.g., $|Z| > 1.96$ for $\alpha = 0.05$)
- P-value: $p = 2 \times P(Z > |z_{\text{obs}}|)$

4.2.2 Confidence Interval

95% CI for difference in proportions:

$$(\hat{p}_1 - \hat{p}_2) \pm 1.96 \times \sqrt{\frac{\hat{p}_1(1 - \hat{p}_1)}{n_1} + \frac{\hat{p}_2(1 - \hat{p}_2)}{n_2}} \quad (18)$$

Note

Example: Conversion Rate Test

Data:

- Control: 500 conversions out of 10,000 users ($\hat{p}_2 = 0.05$)
- Treatment: 575 conversions out of 10,000 users ($\hat{p}_1 = 0.0575$)

Pooled proportion:

$$\hat{p} = \frac{500 + 575}{10000 + 10000} = \frac{1075}{20000} = 0.05375$$

Test statistic:

$$\begin{aligned} Z &= \frac{0.0575 - 0.05}{\sqrt{0.05375(0.94625) \left(\frac{1}{10000} + \frac{1}{10000} \right)}} \\ &= \frac{0.0075}{\sqrt{0.050896 \times 0.0002}} \\ &= \frac{0.0075}{0.003193} = 2.348 \end{aligned}$$

P-value: $p = 2 \times P(Z > 2.348) \approx 0.019$

Conclusion: $p < 0.05 \Rightarrow$ Reject H_0 , significant difference detected

Effect size: Relative lift = $(0.0575 - 0.05)/0.05 = 15\%$

Implementation: See `proportion_test()` in `statistical_analysis.py`

4.3 Two-Sample T-Test

Use For: Continuous metrics (revenue, time on site, cart value)

Key References:

- Student (W. S. Gosset). (1908). "The Probable Error of a Mean." *Biometrika*, 6(1), 1–25.
- Welch, B. L. (1947). "The Generalization of Student's Problem when Several Different Population Variances are Involved." *Biometrika*, 34(1/2), 28–35.

4.3.1 Welch's T-Test (Recommended)

Why Welch's t-test? It does NOT assume equal variances (robust to heteroscedasticity).

Test statistic:

$$t = \frac{\bar{X}_1 - \bar{X}_2}{\sqrt{\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}}} \quad (19)$$

Degrees of freedom (Welch-Satterthwaite approximation):

$$\nu = \frac{\left(\frac{s_1^2}{n_1} + \frac{s_2^2}{n_2}\right)^2}{\frac{(s_1^2/n_1)^2}{n_1-1} + \frac{(s_2^2/n_2)^2}{n_2-1}} \quad (20)$$

Decision Rule:

- Reject H_0 if $|t| > t_{\alpha/2, \nu}$
- P-value from t -distribution with ν degrees of freedom

Note

When Assumptions Fail: Robustness of T-Test

The t-test assumes:

1. Independence (satisfied by randomization)
2. Normality of sampling distribution
3. Equal variances (not required for Welch's version)

Good News: Thanks to the Central Limit Theorem (CLT), the t-test is robust to non-normality when:

- Sample sizes are large ($n > 30$ per group)
- Distributions aren't extremely skewed

E-Commerce Reality: Revenue data is often:

- Right-skewed (many small purchases, few large ones)
- Heavy-tailed (outliers exist)
- Zero-inflated (many users don't convert)

Solutions:

- For moderate skewness + large samples: Welch's t-test is still valid
- For severe skewness: Use Mann-Whitney U test (non-parametric, Section 4.4)
- For revenue: Consider winsorizing outliers or log-transformation

Implementation: See `continuous_metric_test()` in `statistical_analysis.py`

4.4 Mann-Whitney U Test (Non-Parametric)

Use For: Continuous or ordinal data when:

- Distributions are heavily skewed

- Outliers are present
- Sample sizes are small
- Normality assumption violated

Key References:

- Mann, H. B., & Whitney, D. R. (1947). “On a Test of Whether one of Two Random Variables is Stochastically Larger than the Other.” *Annals of Mathematical Statistics*, 18(1), 50–60.
- Wilcoxon, F. (1945). “Individual Comparisons by Ranking Methods.” *Biometrics Bulletin*, 1(6), 80–83.

4.4.1 How It Works

Instead of comparing means, Mann-Whitney compares *ranks*:

1. Pool all observations from both groups
2. Rank them from smallest to largest
3. Sum ranks for each group
4. Test if rank sums are significantly different

Null Hypothesis: The two distributions are identical

Alternative: One distribution is stochastically larger (higher values more likely)

Test Statistic:

$$U = n_1 n_2 + \frac{n_1(n_1 + 1)}{2} - R_1 \quad (21)$$

where R_1 is sum of ranks for group 1.

Note**Example: Revenue Test with Skewed Data****Scenario:** Revenue data with outliers**Control** (10 users): \$0, \$0, \$25, \$30, \$35, \$40, \$45, \$50, \$55, \$500

- Mean = \$78
- Median = \$37.50

Treatment (10 users): \$0, \$20, \$40, \$45, \$50, \$55, \$60, \$65, \$70, \$75

- Mean = \$48
- Median = \$52.50

Problem with t-test: Control mean (\$78) is inflated by \$500 outlier, would incorrectly suggest control is better!**Mann-Whitney Approach:**

- Ranks all 20 values
- Compares rank sums (robust to outliers)
- Correctly identifies treatment as better (higher median, more consistent performance)

Implementation: See `mann_whitney_test()` in `statistical_analysis.py`

4.5 ANOVA (Multiple Variant Tests)

Use For: Comparing > 2 variants on continuous metric**Key References:**

- Fisher, R. A. (1925). *Statistical Methods for Research Workers*. Oliver and Boyd.

4.5.1 One-Way ANOVA

Tests: $H_0: \mu_1 = \mu_2 = \dots = \mu_k$ (all group means equal)**Test Statistic** (F-ratio):

$$F = \frac{\text{Between-group variance}}{\text{Within-group variance}} = \frac{MS_{\text{between}}}{MS_{\text{within}}} \quad (22)$$

Decomposition of variance:

$$SS_{\text{total}} = SS_{\text{between}} + SS_{\text{within}} \quad (23)$$

$$SS_{\text{between}} = \sum_{i=1}^k n_i (\bar{X}_i - \bar{X})^2 \quad (24)$$

$$SS_{\text{within}} = \sum_{i=1}^k \sum_{j=1}^{n_i} (X_{ij} - \bar{X}_i)^2 \quad (25)$$

If $F > F_{\alpha, k-1, N-k}$, reject H_0 .

4.5.2 Post-Hoc Tests

If ANOVA rejects H_0 (at least one group differs), use post-hoc tests to identify which pairs differ:

Options:

- **Tukey HSD:** Controls family-wise error rate, best for all pairwise comparisons
- **Bonferroni:** Very conservative, simple
- **Dunnett:** Compares all treatments to single control

Note

Example: 3-Variant Product Slider Test

Variants:

- A: Social proof only
- B: Similar products
- C: Hybrid approach

Metric: Revenue per user

ANOVA Result: $F(2, 17997) = 12.5, p < 0.001$

Interpretation: At least one variant differs significantly

Post-Hoc (Tukey HSD):

- A vs B: $p = 0.023$ (significant)
- A vs C: $p = 0.001$ (significant)
- B vs C: $p = 0.412$ (not significant)

Business Conclusion: C (hybrid) performs significantly better than A, similar to B. Deploy C.

Implementation: See `anova_test()` in `statistical_analysis.py`

4.6 Chi-Square Test for Independence

Use For: Categorical outcomes (e.g., exit page, product category chosen)

Key References:

- Pearson, K. (1900). "On the Criterion that a Given System of Deviations." *Philosophical Magazine*, 50, 157–175.

4.6.1 Contingency Table Analysis

Tests independence of two categorical variables.

Example: Does variant affect which product category users browse?

	Electronics	Clothing	Home	Total
Control	150	200	150	500
Treatment	180	180	140	500
Total	330	380	290	1000

Expected Counts (under independence):

$$E_{ij} = \frac{(\text{Row}_i \text{ Total}) \times (\text{Column}_j \text{ Total})}{\text{Grand Total}} \quad (26)$$

Test Statistic:

$$\chi^2 = \sum_{i,j} \frac{(O_{ij} - E_{ij})^2}{E_{ij}} \quad (27)$$

Degrees of freedom: $(r - 1)(c - 1)$ where r = rows, c = columns.

Note

Implementation

See `chi_square_test()` in `statistical_analysis.py`

Cramér's V (effect size for chi-square):

$$V = \sqrt{\frac{\chi^2}{n \times \min(r - 1, c - 1)}} \quad (28)$$

Interpretation:

- $V < 0.1$: Weak association
- $0.1 \leq V < 0.3$: Moderate association
- $V \geq 0.3$: Strong association

4.7 Multiple Testing Correction

The Problem: When testing multiple metrics, false positive rate increases!

Example: Testing 10 independent metrics at $\alpha = 0.05$:

- Probability of NO false positives: $(1 - 0.05)^{10} = 0.599$
- Probability of ≥ 1 false positive: $1 - 0.599 = 0.401$ (40%!)

Key References:

- Bonferroni, C. (1936). *Teoria statistica delle classi e calcolo delle probabilità*.
- Holm, S. (1979). "A Simple Sequentially Rejective Multiple Test Procedure." *Scandinavian Journal of Statistics*, 6(2), 65–70.
- Benjamini, Y., & Hochberg, Y. (1995). "Controlling the False Discovery Rate." *Journal of the Royal Statistical Society B*, 57(1), 289–300.

4.7.1 Correction Methods

1. Bonferroni Correction (Most Conservative)

Adjust significance level: $\alpha_{\text{adjusted}} = \frac{\alpha}{m}$

For $m = 10$ tests and $\alpha = 0.05$: Use $\alpha = 0.005$ for each test.

Pros: Simple, controls family-wise error rate (FWER)

Cons: Very conservative (low power) when many tests

2. Holm-Bonferroni (Recommended for 5–10 tests)

Sequential procedure:

1. Order p-values: $p_{(1)} \leq p_{(2)} \leq \dots \leq p_{(m)}$
2. Test sequentially:
 - Compare $p_{(1)}$ to α/m
 - Compare $p_{(2)}$ to $\alpha/(m-1)$
 - Continue until first non-rejection

Pros: More powerful than Bonferroni, still controls FWER

3. Benjamini-Hochberg FDR (For > 10 tests)

Controls False Discovery Rate (proportion of false positives among rejections).

Procedure:

1. Order p-values: $p_{(1)} \leq \dots \leq p_{(m)}$
2. Find largest k where: $p_{(k)} \leq \frac{k}{m} \times \alpha$
3. Reject H_0 for all $i \leq k$

Pros: More powerful for many tests

Cons: Allows some false positives (by design)

i Note

Which Method to Use?

Our 5 E-Commerce Tests: We analyze 5–10 metrics per test

Recommendation: Use **Holm-Bonferroni**

- Appropriate for 5–10 tests
- Controls FWER (no false positives)
- More powerful than Bonferroni

Alternative: If testing 20+ metrics (e.g., full metric suite), use **Benjamini-Hochberg FDR**

Implementation: See `multiple_testing_correction()` in `validation.py`

Example Usage:

```
# Test 5 metrics, get p-values
pvalues = [0.023, 0.041, 0.087, 0.012, 0.156]

# Apply Holm correction
result = validator.multiple_testing_correction(
    pvalues,
    method='holm',
    alpha=0.05
)

# Result shows which tests remain significant
# after correction
```