Chapter 3 Linear Regression

STAT303-2

Simple Linear Regression - Recap

In the last SLR lecture, we derived the optimum parameters for:

$$\hat{Y} = \hat{f}(X) = \hat{\beta}_0 + \hat{\beta}_1 X$$

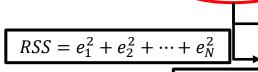
In terms of the data at hand:

$$X^1, X^2, ... X^N$$

$$Y^1, Y^2, ... Y^N$$

and used the formula with optimum parameters for prediction,

$$\widehat{Y} = \widehat{\beta}_0 + \widehat{\beta}_1 X$$



Find RMSE/MAE for test and training data

Training

Last week's

focus

One addition: Residual Standard Error (RSE)

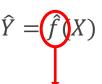
$$RSE = \sqrt{\frac{RSS}{n-2}}$$

- A common alternate stat. measure
- Used for assessing training data

Simple Linear Regression - Inference

Today, we will focus more on the estimated function itself

- Straightforward for SLR
- Only important parameter: \hat{eta}_1
 - The estimated slope
 - Determines the magnitude and direction of the linear association between X and Y
 - Change in Y = Change in X * $\hat{\beta}_1$



Look for insights:

- The parameters themselves
- How sure we are that the real function, f, is close to \hat{f} , and how close
- How much \hat{f} explains the variation in the data $\xrightarrow{\text{Coefficient of }}$

Jncertainty

Determination (R²)

Remember that the formula that we assumed to be linear

$$\hat{Y} = \hat{f}(X) = \hat{\beta}_0 + \hat{\beta}_1 X$$

is just an **estimation** of the underlying function f. \longrightarrow Still assumed to be $\beta_0 + \beta_1 X$

- f has no way to be known or analytically found
- f is only observed through the collected data

Question: How to actually assess how close the estimation is?

We discussed RMSE/MAE to assess the results using the data – Can we do anything a little more analytic about the function?

Find the confidence intervals (CIs)

Recall that the residual sum of squares error is:

$$RSS = (Y^{1} - \hat{\beta}_{0} - \hat{\beta}_{1}X^{1})^{2} + (Y^{2} - \hat{\beta}_{0} - \hat{\beta}_{1}X^{2})^{2} + \dots + (Y^{N} - \hat{\beta}_{0} - \hat{\beta}_{1}X^{N})^{2}, \text{ or}$$

$$RSS = e_{1}^{2} + e_{2}^{2} + \dots + e_{N}^{2}$$

- The coefficients $\widehat{\beta}_0$ and $\widehat{\beta}_1$ are estimated by minimizing RSS
- However, the observations $(X^1, Y^1), (X^2, Y^2), ... (X^N, Y^N)$ correspond to a particular sample of observations
- For a different sample, the values of (X^1, Y^1) , (X^2, Y^2) , ... (X^N, Y^N) are likely to be different, which implies that e_1^2 , e_2^2 ,..., e_n^2 are likely to be different
- Thus, the uncertainty in the estimates $\widehat{\beta}_0$ and $\widehat{\beta}_1$ is due to the uncertainties in the residuals e_1^2 , e_2^2 ,..., e_n^2

Assuming the variance of the residuals to be a constant, i.e.,

$$Var(\epsilon) = \sigma^2$$
,

it can be shown that the uncertainties in the coefficients $\widehat{\beta}_0$ and $\widehat{\beta}_1$ can be estimated as:

$$SE(\hat{\beta}_0)^2 = \sigma^2 \left[\frac{1}{N} + \frac{\bar{x}^2}{\sum_{i=1}^{N} (x_i - \bar{x})^2} \right] \qquad SE(\hat{\beta}_1)^2 = \frac{\sigma^2}{\sum_{i=1}^{N} (x_i - \bar{x})^2}$$

- Note that σ^2 is unknown.
- However, it can be estimated from the data as:

$$\sigma^2 = \sqrt{\frac{RSS}{N-2}}$$

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where
$$\sigma^2 = Var(\epsilon) = \sqrt{\frac{RSS}{n-2}}$$

- The standard errors can be used to compute the confidence intervals. (Cls)
- 95% CI → A range of values that contains the true unknown value of the parameter with 95% probability.
- 95% CI of a variable = variable $\pm 2 * SE$ (variable)
- For SLR:

$$\hat{\beta}_1 \pm 2 * SE(\hat{\beta}_1)$$

$$\hat{\beta}_0 \pm 2 * SE(\hat{\beta}_0)$$

We are 95% sure that the real intercept (β_0) and slope (β_1) values of the underlying function, are within the ranges of:

$$[\hat{\beta}_0 - 2 * SE(\hat{\beta}_0), \hat{\beta}_0 + 2 * SE(\hat{\beta}_0)]$$

 $[\hat{\beta}_1 - 2 * SE(\hat{\beta}_1), \hat{\beta}_1 + 2 * SE(\hat{\beta}_1)]$

95% CIs for SLR

$$SE(\hat{\beta}_{0})^{2} = \sigma^{2} \left[\frac{1}{N} + \frac{\bar{X}^{2}}{\sum_{i=1}^{N} (X^{i} - \bar{X})^{2}} \right] \qquad [\hat{\beta}_{0} - 2 * SE(\hat{\beta}_{0}), \hat{\beta}_{0} + 2 * SE(\hat{\beta}_{0})]$$

$$SE(\hat{\beta}_{1})^{2} = \frac{\sigma^{2}}{\sum_{i=1}^{N} (X^{i} - \bar{X})^{2}} \qquad [\hat{\beta}_{1} - 2 * SE(\hat{\beta}_{1}), \hat{\beta}_{1} + 2 * SE(\hat{\beta}_{1})]$$

$$\text{where } \sigma^{2} = Var(\epsilon) = \sqrt{\frac{RSS}{n-2}}$$

- The ols function from statsmodels calculates the CIs
- We still need to interpret them
 - Higher RSS higher standard errors, wider CI, less certainty
 - Lower variation in X higher standard errors, wider CI, less certainty
 - Higher variation in X lower standard errors, tighter CI, more certainty

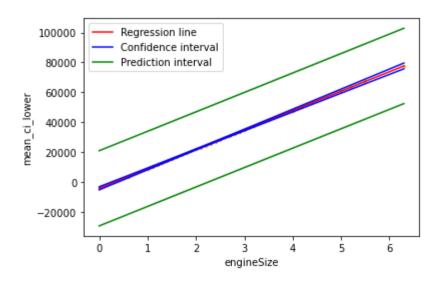
Very similar to the confidence intervals, there is also **prediction intervals**.

- A more conservative approach to uncertainty, that also takes into account the irreducible error in statistical modeling.
- Recall from Chapter 2 that:

$$E(Y - \hat{Y})^2 = [f(X) - \hat{f}(X)]^2 + Var(\in)$$
Reducible Irreducible error error

 For this reason, the prediction interval of a Linear Regression model is always wider than its confidence interval.

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- In this course, you only need to know the prediction interval conceptually as a more conservative approach that takes $Var(\in)$ into account.
- A model created by ols function in Python will easily return both prediction and confidence intervals

Another important application for Standard Errors: Hypothesis Testing

- A purely statistical concept covered in more detail in Statistics courses
- In Hypothesis testing, the starting point is a Null Hypothesis (H₀) and an Alternate Hypothesis (H_A)

$$\mathbf{H_0}$$
: $eta_1 = 0$ \longrightarrow X and Y are unrelated

$$\mathbf{H_A}: \beta_1 \neq 0 \longrightarrow \mathsf{X}$$
 and Y are related

The whole point of Hypothesis testing:

- Using the data to calculate a probability
- That probability of the calculated nonzero \hat{eta}_1 being due to random chance
- This probability is also called the **p-value**.

To find the probability, we first need to calculate a value called the t-statistic.

$$t = \frac{\hat{\beta}_1}{SE(\hat{\beta}_1)}$$

- t is a random variable that belongs to a distribution, called tdistribution.
- You can calculate the t-statistic plugging in $\hat{\beta}_1$ and $SE(\hat{\beta}_1)$
- You can then check (using Python) where the t-statistic falls in the distribution and find the **probability** of that value t-statistic or a higher/lower value occurring. (Again, using Python)

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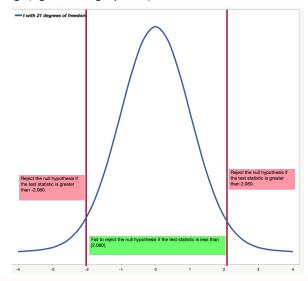
- This probability is also called the **p-value**.
 - If p-value is below 0.05 (5%) the Null Hypothesis (H_0) is rejected.
 - That means the non-zero linear association between X and Y is statistically significant.

$$\mathbf{H_0}$$
: $\beta_1 = 0$ \longrightarrow X and Y are unrelated.

$$\mathbf{H_A}: \beta_1 \neq 0$$
 \longrightarrow X and Y are related.

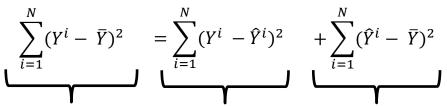
In plain English, using $\hat{\beta}_1$ and SE($\hat{\beta}_1$), we can figure out how sure we are that there is an actual linear relation between X and Y.

- $\hat{\beta}_1$ is still an estimation of that linear relation.
- How good is $\hat{\beta}_1$? That is determined by RMSE/MAE and another metric R² (up next)

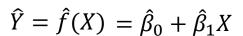


Question: How much does \hat{f} explain the variation in the data?

- The responses (Y's) in the data has a certain variation. → Quantified by Var(Y)
- Using X's and the optimum parameters, $(\hat{\beta}_0 \text{ and } \hat{\beta}_1)$ how much of this variance can the model explain/capture?
- Let's start with the formula for the variation of Y, what we try to explain:



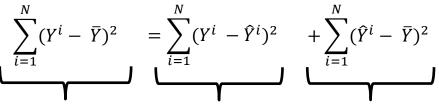
- Total sum of squares (TSS)
- Total variation
- Residual sum of squares (RSS)
- · Unexplained variation
- Explained sum of squares (RSS)
- Explained variation



Question: How much does \hat{f} explain the variation in the data?

R² is the metric that quantifies this question by the ratio of the explained variation to the total variation.

$$R^2 = \frac{Explained\ variation}{Total\ variation} = \frac{Total\ variation - Unexplained\ variation}{Total\ variation} = 1 - \frac{Unexplained\ variation}{Total\ variation}$$



- Total sum of squares (TSS)
- Total variation
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 (RSS)
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$$R^{2} = 1 - \frac{\sum_{i=1}^{N} [Y^{i} - \hat{Y}^{i}]^{2}}{\sum_{i=1}^{N} [Y^{i} - \bar{Y}]^{2}}$$

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- TSS can be calculated before the training/prediction.
- RSS is calculated after the prediction we need the \widehat{Y} 's.
- So, after a model and trained and the results are returned,
 - RMSE/MAE the numeric assessment of the prediction accuracy Prediction Task
 - R^2 the numeric assessment of how well the function is fit \longrightarrow Inference Task
 - A ratio between 0 and 1
 - Also, R^2 is the square of Pearson correlation coefficient.

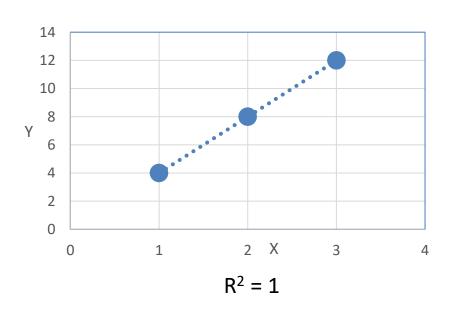
statsmodels table

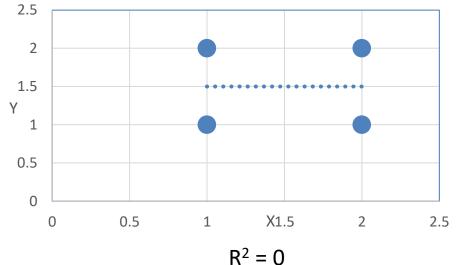
$$R^{2} = 1 - \frac{\sum_{i=1}^{N} [Y^{i} - \hat{Y}^{i}]^{2}}{\sum_{i=1}^{N} [Y^{i} - \bar{Y}]^{2}} = 1 - \frac{RSS}{TSS}$$

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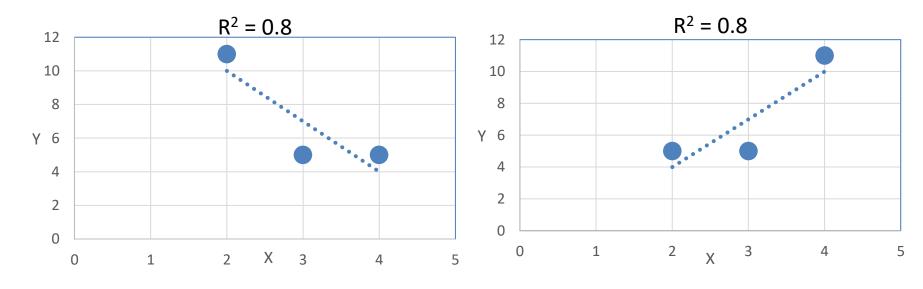
- Ideal scenario: RMSE/MAE as low as possible and R² as close to 1 as possible.
- Note that RMSE/MAE is a continuous number
 - Its assessment is subjective
 - It can be too high, tolerable or good, depending on the prediction task.
- On the other hand, R^2 is a ratio a more objective metric
 - A low RMSE/MAE can still end up with a relatively low R² if TSS is low
- If you are after accurate predictions: RMSE/MAE
- If you are after a good statistical explanation between X and Y: R^2
- Usually, a low RMSE/MAE and a high R² go hand-in-hand but not always

Visualization with two toy models: Extreme cases





Two more toy models: Same R^2 with different slope



To wrap up:

Prediction:

A Simple Linear Regression (SLR) model: $Y = f(X) = \beta_0 + \beta_1 X$ One-line Python implementations:

olf (SLK) model.
$$I = f(X) = p_0 + p_1 X$$
 of the line rythorn implementations:

ols_object = smf.ols(formula = 'price~engineSize', data = train)

Training: • Find
$$\hat{\beta}_0$$
 and $\hat{\beta}_1$ (using training data and the equations)

Using $\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 X$, find the predictions for test data

- Find training and test RMSE/MAE
 - Compare test RMSE/MAE with training RMSE/MAE/MSE for overfitting/underfitting

np.sqrt(((testp.price - pred_price)**2).mean())

pred price = model.predict(testf)

model = ols_object.fit()

- **Inference:** Find **95% confidence intervals** for $\hat{\beta}_0$ and $\hat{\beta}_1$
 - Find the prediction intervals
 - Find the t-statistic and p-value for statistical
 - Find the R² value.

significance

intervals = model.get_prediction(testf)

model.summary()

intervals.summary frame(alpha=0.05)