```
\tfrac{\partial}{\partial \mathbf{x}}(\|\mathbf{x}-\mathbf{b}\|_2) = \tfrac{\mathbf{x}-\mathbf{b}}{\|\mathbf{x}-\mathbf{b}\|_2} \quad \  \tfrac{\partial}{\partial \mathbf{x}}(\|\mathbf{x}\|_2^2) = \tfrac{\partial}{\partial \mathbf{x}}(\mathbf{x}^\top\mathbf{x}) =
0 Essentials
Matrix/Vector
                                                                                                          \frac{\partial}{\partial \mathbf{X}}(\|\mathbf{X}\|_F^2) = 2\mathbf{X} \frac{\partial}{\partial \mathbf{X}}\log(x) = \frac{1}{x}
Vectors: Unit vector: u^{\dagger}u = 1 Orthogonal vec-
tors: u^{\dagger}v = 0 Range, Kernel, Nullity: range(\mathbf{A}) =
                                                                                                          Eigendecomposition
\{\mathbf{z}|\exists\mathbf{x}:\mathbf{z}=\mathbf{A}\mathbf{x}\}=span(\text{columns of A})
                                                                                                         \mathbf{A} \in \mathbb{R}^{N \times N} then \mathbf{A} = \mathbf{Q} \Lambda \mathbf{Q}^{-1} with \mathbf{Q} \in \mathbb{R}^{N \times N}.
rank(\mathbf{A}) = dim(range(\mathbf{A})) \ kernel(A) = \{\mathbf{x} : \mathbf{A}\mathbf{x} =
                                                                                                         if fullrank: \mathbf{A}^{-1} = \mathbf{Q} \mathbf{\Lambda}^{-1} \mathbf{Q}^{-1} and (\mathbf{\Lambda}^{-1})_{i,i} = \frac{1}{\lambda}.
\{0\} (spans nullspace) nullity(A) = dim(kernel(A))
                                                                                                         if A symmetric: A = \mathbf{Q} \Lambda \mathbf{Q}^{\top} (Q orthogonal). Eigen-
Ranks: rank(XY) \le rank(X) \forall X \in R^{mxn}, Y \in R^{nxk}
                                                                                                         value \lambda: solve det(A - \lambda I) = 0 Eigenvector \nu: solve
eq. if Y \in R^{n \times n}, rank(Y) = n Rank-nullity Theo-
                                                                                                          (A - \lambda I) * v = 0'
rem: dim(kernel(\mathbf{A})) + dim(range(\mathbf{A})) = n
                                                                                                         Probability / Statistics
Orthogonal mat. A^{-1} = A^{\top}, AA^{\top} = A^{\top}A = I,
                                                                                                         • P(x) := Pr[X = x] := \sum_{y \in Y} P(x, y) \cdot P(x|y) :=
\det(\mathbf{A}) \in \{+1, -1\}, \det(\mathbf{A}^{\top}\mathbf{A}) = 1, preserves inner
product, norm, distance, angle, rank, matrix ortho-
                                                                                                        Pr[X = x | Y = y] := \frac{P(x,y)}{P(y)}, \text{ if } P(y) > 0 \bullet \forall y \in
gonality Outer Product: \mathbf{u}\mathbf{v}^{\top}, (\mathbf{u}\mathbf{v}^{\top})_{i,j} = \mathbf{u}_{i}\mathbf{v}_{j}
                                                                                                          Y: \sum_{x \in X} P(x|y) = 1 (property for any fixed
Inner Product: \langle \mathbf{x}, \mathbf{y} \rangle = \mathbf{x}^{\top} \mathbf{y} = \sum_{i=1}^{N} \mathbf{x}_{i} \mathbf{y}_{i}. \langle \mathbf{x} \pm \mathbf{y} \rangle
                                                                                                         v) • P(x,y) = P(x|y)P(y) • posterior P(A|B) =
\mathbf{y}, \mathbf{x} \pm \mathbf{y} \rangle = \langle \mathbf{x}, \mathbf{x} \rangle \pm 2 \langle \mathbf{x}, \mathbf{y} \rangle + \langle \mathbf{y}, \mathbf{y} \rangle
                                                                                                          \frac{\operatorname{prior} P(A) \times \operatorname{likelihood} P(B|A)}{\operatorname{prior} P(A) \times \operatorname{likelihood} P(B|A)} \text{ (Bayes' rule)} \cdot P(x|y) =
(\mathbf{u}_i^T \mathbf{v}_i) \mathbf{v}_i = (\mathbf{v}_i \mathbf{v}_i^T) \mathbf{u}_i Cross product: \vec{a} \times \vec{b} =
                                                                                                          P(x) \Leftrightarrow P(y|x) = P(y) (iff X, Y independent)
(a_2b_3-a_3b_2,a_3b_1-a_1b_3,a_1b_2-a_2b_1)^{\top}
                                                                                                         • P(x_1,...,x_n) = \prod_{i=1}^n P(x_i) (iff IID) • Variance
Trace: trace(XYZ) = trace(ZXY)
                                                                                                         Var[X] := E[(X - \mu_x)^2] := \sum_{x \in X} (x - \mu_x)^2 P(x) =
Transpose: ({\bf A}^{\top})^{-1} = ({\bf A}^{-1})^{\top}, \ ({\bf A}{\bf B})^{\top} = {\bf B}^{\top}{\bf A}^{\top},
                                                                                                         E(X^2) - E(X)^2 \operatorname{Var}(aX) = a^2 \operatorname{Var}(X) \cdot \operatorname{expectati}
(\mathbf{A} + \mathbf{B})^{\top} = \mathbf{A}^{\top} + \mathbf{B}^{\top}
                                                                                                         on \mu_x := E[X] := \sum_{x \in X} x P(x) \cdot E[X + Y] = E[X] +
Cauchy-Schwarz inequality: |\langle \mathbf{u}, \mathbf{v} \rangle| \leq ||\mathbf{u}|| ||\mathbf{v}||
                                                                                                          E[Y] • standard deviation \sigma_x := \sqrt{Var[X]}
Jensen inequality: for convex function f, non
negative \lambda_i st. \sum_{i=1}^n \lambda_i = 1: f(\sum_{i=1}^n \lambda_i x_i) \leq
                                                                                                          Lagrangian Multipliers
                                                                                                          Minimize f(\mathbf{x}) s.t. g_i(\mathbf{x}) \leq 0, i = 1,...,m
\sum_{i=1}^{n} \lambda_i f(x_i) Note: for concave, inequality sign
switches Convexity: f(\theta x + (1 - \theta)y) \le \theta f(x) + \theta f(x)
                                                                                                          (inequality constr.) and h_i(\mathbf{x}) = \mathbf{a}_i^{\mathsf{T}} \mathbf{x} - b_i = 0 or
(1-\theta)f(y), \forall \theta \in [0,1] Least Squares equations:
                                                                                                         h_i(\mathbf{x}) = \sum_{w} x_{w,i} - b_i = 0, i = 1,..,p (equality cons-
\operatorname{arg\,min}_{\beta \in \mathbb{R}^{p+1}} \|\mathbf{y} - \mathbf{X}\boldsymbol{\beta}\|^2, \, \hat{\boldsymbol{\beta}} = (X^\top X)^{-1} X^\top y
                                                                                                         L(\mathbf{x}, \alpha, \beta) := f(\mathbf{x}) + \sum_{i=1}^{m} \alpha_i g_i(\mathbf{x}) + \sum_{i=1}^{p} \beta_i h_i(\mathbf{x})
Einstein matrix notation: (A \cdot B)_{ij} = \sum_{k=1}^{n} A_{ik} \cdot B_{kj}
                                                                                                          1 Principal Component Analysis
Kullback-Leibler: KL(P||Q) = \sum_{x \in X} P(x) \log \frac{P(x)}{Q(x)}
                                                                                                         \mathbf{X} \in \mathbb{R}^{D \times N}. N observations, K rank.
                                                                                                         1. Empirical Mean: \overline{\mathbf{x}} = \frac{1}{N} \sum_{n=1}^{N} \mathbf{x}_n.
• \|\mathbf{x}\|_0 = |\{i | x_i \neq 0\}|
                                                                                                         2. Center Data: \overline{\mathbf{X}} = \mathbf{X} - [\overline{\mathbf{x}}, \dots, \overline{\mathbf{x}}] = \mathbf{X} - \mathbf{M}.
• \|\mathbf{x}\|_2 = \sqrt{\sum_{i=1}^N \mathbf{x}_i^2} = \sqrt{\langle \mathbf{x}, \mathbf{x} \rangle}
                                                                                                         3. Cov.: \Sigma = \frac{1}{N} \sum_{n=1}^{N} (\mathbf{x}_n - \overline{\mathbf{x}}) (\mathbf{x}_n - \overline{\mathbf{x}})^{\top} = \frac{1}{N} \overline{\mathbf{X}} \overline{\mathbf{X}}^{\top}.
• \|\mathbf{u} - \mathbf{v}\|_2 = \sqrt{(\mathbf{u} - \mathbf{v})^\top (\mathbf{u} - \mathbf{v})}
                                                                                                         4. Eigenvalue Decomposition: \Sigma = \mathbf{U}\Lambda\mathbf{U}^{\top}.
                                                                                                          5. Select K < D, only keep U_K, \lambda_K.
• \|\mathbf{x}\|_p = (\sum_{i=1}^N |x_i|^p)^{\frac{1}{p}}; \|\mathbf{x}\|_{\infty} = \max_{i=1,...,n} |x_i|
                                                                                                         6. Transform data onto new Basis: \overline{\mathbf{Z}}_K = \mathbf{U}_K^{\top} \overline{\mathbf{X}}.
• \|\mathbf{M}\|_F = \sqrt{\sum_{i=1}^m \sum_{j=1}^n \mathbf{m}_{i,j}^2} = \sqrt{\sum_{i=1}^{\min\{m,n\}} \sigma_i^2} =
                                                                                                         7. Reconstruct to original Basis: \overline{\mathbf{X}} = \mathbf{U}_k \overline{\mathbf{Z}}_K.
\|\boldsymbol{\sigma}(\mathbf{A})\|_2 = \sqrt{trace(\mathbf{M}^T\mathbf{M})}
                                                                                                         8. Reverse centering: \tilde{\mathbf{X}} = \overline{\mathbf{X}} + \mathbf{M}.
                                                                                                         For compression save U_k, \overline{Z}_K, \overline{x}.
• \|\mathbf{M}\|_G = \sqrt{\sum_{ij} g_{ij} x_{ij}^2} (weighted Frobenius)
                                                                                                          \mathbf{U}_k \in \mathbb{R}^{D \times K}, \Sigma \in \mathbb{R}^{D \times D}, \overline{\mathbf{Z}}_K \in \mathbb{R}^{K \times N}, \overline{\mathbf{X}} \in \mathbb{R}^{D \times N}
• \|\mathbf{M}\|_1 = \sum_{i,j} |m_{i,j}|
                                                                                                         Calculation of: var(X) = \frac{1}{N} \sum_{n=1}^{N} (X_i - \overline{X})^2
• \|\mathbf{M}\|_2 = \sigma_{\max}(\mathbf{M}) = \|\sigma((M))\|_{\infty} (spectral)
                                                                                                          Iterative View
• \|\mathbf{M}\|_p = \max_{\mathbf{v} \neq 0} \frac{\|\mathbf{M}\mathbf{v}\|_p}{\|\mathbf{v}\|_p}
                                                                                                          Residual r_i: x_i - \tilde{x}_i = I - uu^T x_i
                                                                                                         Cov of r: \frac{1}{n} \sum_{i=1}^{n} (I - uu^{T}) x_{i} x_{i}^{T} (I - uu^{T})^{T} =
• \|\mathbf{M}\|_{\star} = \sum_{i=1}^{\min(m,n)} \sigma_i = \|\sigma(\mathbf{A})\|_1 (nuclear)
                                                                                                         (I - uu^T)\Sigma(I - uu^T)^T = \Sigma - 2\Sigma uu^T + uu^T\Sigma uu^T =
Derivatives
\frac{\partial}{\partial \mathbf{x}}(\mathbf{b}^{\top}\mathbf{x}) = \frac{\partial}{\partial \mathbf{x}}(\mathbf{x}^{\top}\mathbf{b}) = \mathbf{b} \frac{\partial}{\partial \mathbf{x}}(\mathbf{x}^{\top}\mathbf{x}) = 2\mathbf{x}
                                                                                                          1. Find principal eigenvector of (\Sigma - \lambda uu^T)
\frac{\partial}{\partial \mathbf{x}}(\mathbf{x}^{\top}\mathbf{A}\mathbf{x}) = (\mathbf{A}^{\top} + \mathbf{A})\mathbf{x} \frac{\partial}{\partial \mathbf{x}}(\mathbf{b}^{\top}\mathbf{A}\mathbf{x}) = \mathbf{A}^{\top}\mathbf{b}
                                                                                                          2. which is the second eigenvector of \Sigma
\frac{\partial}{\partial \mathbf{X}}(\mathbf{c}^{\top}\mathbf{X}\mathbf{b}) = \mathbf{c}\mathbf{b}^{\top} \frac{\partial}{\partial \mathbf{X}}(\mathbf{c}^{\top}\mathbf{X}^{\top}\mathbf{b}) = \mathbf{b}\mathbf{c}^{\top}
                                                                                                          3. iterating to get d principal eigenvector of \Sigma
```

```
1/N \| (U_K U_K^\top - I_d) \overline{X} \|_F^2 = 1/N * trace((U_K U_K^\top - I_d) \overline{X}) \|_F^2
I_d)\overline{XX}^{\top}(U_KU_K^{\top}-I_d)^{\top}=1/N*trace(([U_K;0]-I_d)^{\top})
U)\Lambda([U_K;0]-U)^{\perp})
2 Singular Value Decomposition
\begin{aligned} \mathbf{A} &= \mathbf{U}\mathbf{D}\mathbf{V}^{\top} = \sum_{k=1}^{\mathrm{rank}(\mathbf{A})} d_{k,k} u_k (v_k)^{\top} \\ \mathbf{A} &\in \mathbb{R}^{N \times P}, \mathbf{U} \in \mathbb{R}^{N \times N}, \mathbf{D} \in \mathbb{R}^{N \times P}, \mathbf{V} \in \mathbb{R}^{P \times P} \end{aligned}
\mathbf{U}^{\top}\mathbf{U} = \mathbf{I} = \mathbf{V}^{\top}\mathbf{V} (U, V orthonormal)
 U columns are eigvecs of AA^{\top}, V columns are eig-
vecs of \mathbf{A}^{\top}\mathbf{A}, \mathbf{D} diag. elements are singular values.
 (\mathbf{D}^{-1})_{i,i} = \frac{1}{\mathbf{D}_{i,i}} (don't forget to transpose)
1. calculate \mathbf{A}^{\top}\mathbf{A}.
 2. calculate eigvals of \mathbf{A}^{\top}\mathbf{A}, the square root of them,
 in descending order, are the diagonal elements of
 3. calc. eigvecs of \mathbf{A}^{\top}\mathbf{A} using eigvals resulting in
 the columns of V.
 4. calculate the missing matrix: \mathbf{U} = \mathbf{AVD}^{-1}.
 normalize each column of U and V.
 Low-Rank approximation
Use only K largest eigvals (and corresp. eigvecs).
 \tilde{\mathbf{A}}_{i,j} = \sum_{k=1}^K \mathbf{U}_{i,k} \mathbf{D}_{k,k} \mathbf{V}_{j,k} = \sum_{k=1}^K \mathbf{U}_{i,k} \mathbf{D}_{k,k} (\mathbf{V}^\top)_{k,j}.
 Echart-Young Theorem
\mathbf{A}_k = \operatorname{arg\,min}_{rank(B)=k} \|\mathbf{A} - \mathbf{B}\|_F^2 \quad (\text{not convex})
\min_{rank(B)=K} ||A - B||_F^2 = ||A - A_k||_F^2 = \sum_{r=k+1}^{rank(A)} \sigma_r^2
 \min_{rank(B)=K} ||A - B||_2 = ||A - A_k||_2 = \sigma_{k+1}
 3 Matrix Approximation & Reconstruction
 \min_{rank(B)=k} [\sum_{(i,j)\in I} (a_{ij} - b_{ij})^2], I = \{(i,j): ob.\}
 Alternating Least Squares
f(U, v_i) = \sum_{(i,j) \in I} (a_{i,j} - \langle u_j, v_i \rangle)^2
f(u_i, V) = \sum_{(i,j) \in I} (a_{i,j} - \langle u_j, v_i \rangle)^2
 Convex when fixed one.
 Convex Optimization
 Def.: \{(x,t)|x\in domf, f(x)\leq t\}, f:\mathbb{R}^D\to\mathbb{R} is
 convex, if dom f is a convex set, and if \forall x, y \in
 dom f, and \forall \alpha \in [0,1]: f(\alpha \mathbf{x} + (1-\alpha)\mathbf{y}) \leq
 \alpha f(\mathbf{x}) + (1 - \alpha) f(\mathbf{y}). Convex \iff Hessian p.s.d
  ⇔ local=global
 Positive semi-definite: all principal minors (same-
 indexed rows and columns) > 0
 Positive definite: leading principal minors > 0
 Convex Relaxation
 Replace non-convex rank constraints by convex
 norm constraints (superset). Then project optimum
 back (hopefully still optimal).
```

 $\min_{\mathbf{B}\in P_k} \|\mathbf{A} - \mathbf{B}\|_G^2, P_k = \{\mathbf{B} : \|\mathbf{B}\|_* \le k\} \supseteq Q_k =$

Power Method

Power iteration: $v_{t+1} = \frac{Av_t}{||Av_t||}$, $\lim_{t\to\infty} v_t = u_1$

Reconstruction Proof Sketch

Assuming $\langle u_1, v_0 \rangle \neq 0$ and $|\lambda_1| > |\lambda_i| (\forall i \geq 2)$

Given: $\tilde{X} = U_K U_K^{\top} \overline{X}$ To prove: squared reconstructi-

on error is the sum of the lowest D - K eigenvalues

of Σ . $err = 1/N \sum_{i=1}^{N} \|\tilde{x}_i - \overline{x}_i\|_2^2 = 1/N \|\tilde{X} - \overline{X}\|_F^2 = 1/N \|\tilde{X} - \overline{X}\|_F^2$

Iteration: $\mathbf{B}_{t+1} = \mathbf{B}_t + \eta_t \Pi(\mathbf{A} - shrink_{\tau}(\mathbf{B}_t))$ **4** Non-Negative Matrix Factorization $\mathbf{X} \in \mathbb{Z}_{>0}^{N \times M}$, NMF: $\mathbf{X} \approx \mathbf{U}^{\top} \mathbf{V}, x_{i,i} = \sum_{z} u_{zi} v_{z,i} =$ $\langle \mathbf{u}_i \mathbf{v}_i \rangle$ Decompose object into features: topics, face parts, etc.. u weights on parts, v parts (bases). More interpretable (PCA: holistic repre.). EM for MLE for pLSA (No global opt. guarantee) Context Model: $p(w|d) = \sum_{z=1}^{K} p(w|z)p(z|d)$ Conditional independence assumption (*): $p(w|d) = \sum_{z} p(w,z|d) = \sum_{z} p(w|d,z)p(z|d) \stackrel{*}{=}$ $\sum_{z} p(w|z)p(z|d)$ or p(w|d,z) = p(w|z)**Symmetric parameterization:** $p(w,d) = \sum_{z} p(z)p(w|z)p(d|z)$ Log-Likelihood: $L(\mathbf{U}, \mathbf{V}) = \sum_{i,j} x_{i,j} \log p(w_i|d_i)$ $= \sum_{(i,j)\in X} \log \sum_{z=1}^{K} p(w_j|z) p(z|d_i)$ $p(w_i|z) = v_{zi}, p(z|d_i) = u_{zi}, \sum_{i=1}^{N} v_{zi} = \sum_{z=1}^{K} u_{zi} = 1$ E-Step (optimal q: posterior of z over (d_i, w_i)): $q_{zij} = \frac{\bar{p}(w_j|z)p(z|d_i)}{\sum_{k=1}^{K} p(w_j|k)p(k|d_i)} := \frac{v_{zj}u_{zi}}{\sum_{k=1}^{K} v_{kj}u_{ki}}, \sum_{z} q_{zij} = 1$ M-Steps: $p(z|d_i) = \frac{\sum_j x_{ij} q_{zij}}{\sum_j x_{ij}}, p(w_j|z) = \frac{\sum_i x_{ij} q_{zij}}{\sum_{i,l} x_{il} q_{zil}}$ Lower Bound of $L(\mathbf{U}, \mathbf{V})$ Jensen ineq. : $\sum_{i,j\in X}\sum_{z=1}^{K}q_{zij}(log(v_{zj})+log(u_{zi})-log(q_{zij}))$ **Latent Dirichlet Allocation** To sample a new document, we need to extend X and U^T with a new row, s.t. $X = U^T V$. (While pLSA fixes both dimensions)

For each d_i sample topic weights $\mathbf{u}_i \sim \text{Dirichlet}(\alpha)$: $p(u_i|\alpha) = \prod_{z=1}^K u_{zi}^{\alpha_k - 1}$, then topic $z^t \sim \text{Multi}(u_i)$,

Multinom. obsv. model on we vec: $p(\mathbf{x}|V,u) =$

Bayesian averaging over **u**: $p(\mathbf{x}|\mathbf{V},\alpha) =$

 $\min_{\mathbf{U},\mathbf{V}} J(\mathbf{U},\mathbf{V}) = \frac{1}{2} \|\mathbf{X} - \mathbf{U}^{\top} \mathbf{V}\|_{F}^{2}$ (non-negativity)

1. sampling model: Gaussian vs multinomial 2. ob-

jective: quadratic vs KL divergence 3. constraints:

3. upd. $(\mathbf{V}\mathbf{V}^{\top})\mathbf{U} = \mathbf{V}\mathbf{X}^{\top}$, proj. $u_{zi} = \max\{0, u_{zi}\}$

NMF Algorithm for quadratic cost function

 $\frac{l!}{\prod_{j} \mathbf{x}_{j}!} \prod_{j} \pi_{j}^{\mathbf{x}_{j}}$ where $\pi_{j} = \sum_{z} v_{zj} u_{z}, \ l = \sum_{j} x_{j}$

word $w^t \sim \text{Multi}(v_{\tau^t})$

 $\int p(\mathbf{x}|\mathbf{V},\mathbf{u})p(\mathbf{u}|\alpha)d\mathbf{u}$

s.t. $\forall i, j, z : u_{7i}, v_{7i} \geq 0$

not normalized

Comparison with pLSA:

Alternating least squares:

2. repeat 3~4 for *maxIters*:

1. init: $\mathbf{U}, \mathbf{V} = rand()$

 $\{\mathbf{B} : rank(\mathbf{B}) \le k\}$ (in fact tightest convex lower-

 $\mathbf{B}^* = shrink_{\tau}(\mathbf{A}) = arg \min_{\mathbf{B}} \{ \|\mathbf{A} - \mathbf{B}\|_F^2 + \tau \|\mathbf{B}\|_* \}$

Then with SVD $\mathbf{A} = \mathbf{UDV_T}, \mathbf{D} = diag(\sigma_i)$, holds

bound $rank(\mathbf{B}) \ge \|\mathbf{B}\|_*, for \|\mathbf{B}\|_2 \le 1$

 $\mathbf{B}^* = \mathbf{U}\mathbf{D}_{\tau}\mathbf{V}^{\mathsf{T}}, \mathbf{D}_{\tau} = diag(\max\{0, \sigma_i - \tau\})$

SVD Thresholding

 \rightarrow 2-sided loss. cutoff n_{max} : limit influence of high $\frac{\pi_k^{(t-1)} \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_k^{(t-1)}, \boldsymbol{\Sigma}_k^{(t-1)})}{\boldsymbol{\Sigma}_{j=1}^K \pi_i^{(t-1)} \mathcal{N}(\mathbf{x}_n | \boldsymbol{\mu}_i^{(t-1)}, \boldsymbol{\Sigma}_i^{(t-1)})}$ freq. $f(n) \stackrel{n \to 0}{\to} 0$: as small counts very noisy 1. sample $(i, j)u.a.r, s.t.n_{ij} > 0$ M-Step: $\mu_k^{(t)} := \frac{\sum_{n=1}^N q_{k,n} \mathbf{x}_n}{\sum_{n=1}^N q_{k,n}}, \pi_k^{(t)} := \frac{1}{N} \sum_{n=1}^N q_{k,n}$ 2. $\mathbf{x}_i^{new} \leftarrow \mathbf{x}_i + 2\eta f(n_{ij})(\log n_{ij} - \langle \mathbf{x}_i, \mathbf{y}_j \rangle)\mathbf{y}_i$ $\Sigma_{k}^{(t)} = \frac{\sum_{n=1}^{N} q_{k,n} (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t)}) (\mathbf{x}_{n} - \boldsymbol{\mu}_{k}^{(t)})^{\top}}{\sum_{n=1}^{N} q_{k,n}}$ 3. $\mathbf{y}_{i}^{new} \leftarrow \mathbf{y}_{j} + 2\eta f(n_{ij})(\log n_{ij} - \langle \mathbf{x}_{i}, \mathbf{y}_{j} \rangle)\mathbf{x}_{i}$ embeds can model analogies and relatedness, but K-means vs. EM hard vs soft; spherical clusters vs antonyms are usually not well captured. covariance matrix; fast and cheap vs slow and more 6 Data Clustering & Mixture Models iteration; K-means can be used as init. for EM. K-K-means Target: $\min_{\mathbf{U},\mathbf{Z}} J(\mathbf{U},\mathbf{Z}) = \|\mathbf{X} - \mathbf{U}\mathbf{Z}\|_F^2$ means as a special case of GMM with covariances $=\sum_{n=1}^{N}\sum_{k=1}^{K}\mathbf{z}_{k,n}\|\mathbf{x}_{n}-\mathbf{u}_{k}\|_{2}^{2}$ $\Sigma_i = \sigma^2 I$. limit of $\sigma \to 0$ is K-means (hard asgmts). 1. Initiate: choose K centroids $\mathbf{U} = [\mathbf{u}_1, \dots, \mathbf{u}_K]$ Model Order Selection (AIC / BIC for GMM) 2. Cluster Assign: data points to clusters. $k^*(\mathbf{x}_n) =$ Trade-off between data fit (i.e. likelihood $p(\mathbf{X}|\theta)$) $\underset{k}{\operatorname{arg\,min}} \{ \|\mathbf{x}_n - \mathbf{u}_k\|_2 \}$ returns cluster k^* , whose cenand complexity (i.e. # of free parameters $\kappa(\cdot)$). For troid \mathbf{u}_{k^*} is closest to data point \mathbf{x}_n . Set $\mathbf{z}_{k^*,n} = 1$, choosing *K*: AIC($\theta | \mathbf{X}$) = $-\log p_{\theta}(\mathbf{X}) + \kappa(\theta)$ and for $l \neq k^* \mathbf{z}_{l,n} = 0$. $BIC(\theta|\mathbf{X}) = -\log p_{\theta}(\mathbf{X}) + \frac{1}{2}\kappa(\theta)\log N$ 3. Update centroids: $\mathbf{u}_k = \frac{\sum_{n=1}^N z_{k,n} \mathbf{x}_n}{\sum_{n=1}^N z_{k,n}}$. # of free params, fixed covariance matrix: $\kappa(\theta) =$ $K \cdot D + (K-1)$ (K: # clusters, D: dim(data) = 4. Repeat until $\|\mathbf{Z} - \mathbf{Z}^{\text{new}}\|_0 = \|\mathbf{Z} - \mathbf{Z}^{\text{new}}\|_F^2 = 0$. $\dim(\mu_i)$, K-1: π of # free clusters), full covariance matrix: $\kappa(\theta) = K(D + \frac{D(D+1)}{2}) + (K-1)$. Computational cost: $O(k \cdot n \cdot d)$ Prior: p(z) = 1/K**K-Means++:** 1. Choose centroid \mathbf{u}_1 randomly from Compare AIC/BIC for different K – the smaller the datapoints S 2. For $x \in S$, calculate min. squared dibetter. BIC penalizes complexity more. stance $d_m(x)$ to existing centroids $c_1, ..., c_m$ 3. Add 7 Sparse Coding new centroid c_{m+1} , choosen randomly from S with prob. $p(x) = d_m(x) / \sum_{z \in S} d_m(z)$ 4. Repeat until K Orthogonal Basis Pros: fast inverse; preserves energy. For x and orcentroids choosen \rightarrow proceed with K-means Gaussian Mixture Models (GMM) thog. mat. U compute $\mathbf{z} = \mathbf{U}^{\top} \mathbf{x}$. Approx $\hat{\mathbf{x}} = \mathbf{U}\hat{\mathbf{z}}$, $\hat{z}_i = z_i$ if $|z_i| > \varepsilon$ else 0. Reconstruction Error Gaussian $p(x) = \frac{1}{\sqrt{2\pi}\sigma} exp(-\frac{(x-\mu)^2}{2\sigma^2})$ Multivariate $p(x;\mu;\Sigma) = \frac{1}{|\Sigma|^{\frac{1}{2}}(2\pi)^{\frac{D}{2}}} exp[-\frac{1}{2}(x-\mu)^T \Sigma^{-1}(x-\mu)]$ $\|\mathbf{x} - \hat{\mathbf{x}}\|^2 = \sum_{d \in \sigma} \langle \mathbf{x}, \mathbf{u}_d \rangle^2$. Choice of base depends on signal. Fourier: global support, good for sine like waves; wavelet: local support, poor for non-For GMM let $\theta_k = (\mu_k, \Sigma_k); p_{\theta_k}(\mathbf{x}) =$ vanishing signal; PCA basis optimal for given Σ . $\mathcal{N}(\mathbf{x}|\boldsymbol{\mu}_k, \boldsymbol{\Sigma}_k)$ Stripes & check patterns: hi-freq in Fourier. Fourier: $O(D \cdot log D)$, Wavelet: O(D) or $O(D \cdot log D)$ Mixture Models: $p_{\theta}(\mathbf{x}) = \sum_{k=1}^{K} \pi_k p_{\theta_k}(\mathbf{x})$ Haar Wavelets (form orthogonal basis) Assignment variable (generative model): scaling fnc. $\phi(x) = [1, 1, 1, 1]$, mother W(x) = $z_{ij} \in \{0,1\}, \sum_{i=1}^{k} z_{ij} = 1$ [1,1,-1,-1], dilated W(2x) = [1,-1,0,0], trans-Prior: $\Pr(z_k = 1) = \pi_k \Leftrightarrow p(\mathbf{z}) = \prod_{k=1}^K \pi_k^{z_k}$ lated W(2x-1) = [0,0,1,-1] Must be normalized **Complete data distribution:**

 $p_{\theta}(\mathbf{x}, \mathbf{z}) = \prod_{k=1}^{K} (\pi_k p_{\theta_k}(\mathbf{x}))^{z_k}$

 $\Pr(z_k = 1 | \mathbf{x}) = \frac{\Pr(z_k = 1) p(\mathbf{x} | z_k = 1)}{\sum_{l=1}^K \Pr(z_l = 1) p(\mathbf{x} | z_l = 1)} = \frac{\pi_k p_{\theta_k}(\mathbf{x})}{\sum_{l=1}^K \pi_l p_{\theta_l}(\mathbf{x})}$

 $p_{\theta}(\mathbf{X}) = \prod_{n=1}^{N} p_{\theta}(\mathbf{x}_n) = \prod_{n=1}^{N} \left(\sum_{k=1}^{K} \pi_k p_{\theta_k}(\mathbf{x}_n) \right)$

 $\geq \sum_{n=1^N} \sum_{k=1}^K q_k [\log p_{\theta_k}(\mathbf{x}_n) + \log \pi_k - \log q_k]$

1. sample cluster index $j \sim Categorical(\pi)$

2. given j, sample data $x \sim \text{Normal}(\mu_i, \Sigma_i)$

Expectation-Maximization (EM) for GMM

 $\Pr[z_{k,n} = 1|\mathbf{x}_n] =$

evidenceP(B)

 $posterior P(A|B) = \frac{prior P(A) \times likelihood P(B|A)}{r}$

Max. Likelihood Estimation (MLE):

with $\sum_{k=1}^{K} q_k = 1$ by Jensen Inequality.

 $\arg \max_{\theta} \sum_{n=1}^{N} \log \left(\sum_{k=1}^{K} \pi_{k} p_{\theta_{k}}(\mathbf{x}_{n}) \right)$

Likelihood of observed data X:

Posterior Probabilities:

Generative Model

4. update $(\mathbf{U}\mathbf{U}^{\top})\mathbf{V} = \mathbf{U}\mathbf{X}$, proj. $v_{zi} = \max\{0, v_{zi}\}$

Distr. Model: $p_{\theta}(w|w') = \Pr[w \text{ in context of } w']$

 $\log p_{\theta}(w|w') = \langle y_w, x_{w'} \rangle + b_w$, word y_w , c'txt $x_{w'}$ use GloVe obj., negative sampling (logistic class.)

Co-occ.: $\mathbf{N} = (n_{ij}) \in \mathbb{R}^{|V| \times |C|} = \#w_i \text{ in context } w_i$

Objective: $H(\theta; \mathbf{N}) = \sum_{n_{ij}>0} f(n_{ij}) (\log n_{ij} - \sum_{n_{ij}>0} f(n_{ij})) (\log n_{ij} - \sum_{n_{ij}>0} f(n_{ij})) (\log n_{ij} - \sum_{n_{ij}>0} f(n_{ij})) (\log n_{ij})$

 $\log \exp[\langle \mathbf{x}_i, \mathbf{y}_i \rangle + b_i + d_i])^2, \quad f(n) =$

 $\min\{1, (\frac{n}{n_{max}})^{\alpha}\}, \alpha \in (0; 1] (= 3/4)$ unnorm. distr. E-Step:

 $L(\theta; \mathbf{w}) = \sum_{t=1}^{T} \sum_{\Lambda \in I} \log p_{\theta}(w^{(t+\Delta)} | w^{(t)})$

 $p_{\theta}(w|w') = \frac{\exp[\langle \mathbf{x}_{w}, \mathbf{x}_{w'} \rangle + b_{w}]}{\sum_{v \in V} \exp[\langle \mathbf{x}_{v}, \mathbf{x}_{u'} \rangle + b_{v}]} \text{ (soft-max)}.$

GloVe (Weighted Square Loss)

Latent Vector Model: $w \to (\mathbf{x}_w, b_w) \in \mathbb{R}^{D+1}$

5 Word Embeddings

Log-likelihood:

Modifications:

```
s.t. \mathbf{x} = \mathbf{U}\mathbf{z}. NP-hard \rightarrow approximate with 1-norm
  (convex) or with MP.
 Coherence • m(\mathbf{U}) = \max_{i,j:i\neq j} |\mathbf{u}_i^\top \mathbf{u}_j| \bullet m(\mathbf{B}) = 0
 if B orthog. matrix • m([\mathbf{B}, \mathbf{u}]) \ge \frac{1}{\sqrt{D}} if atom u is
  added to orthog. basis \mathbf{B} (o.n.b. = orthonormal base)
   Matching Pursuit (MP) approximation of x onto
   U, using K entries. Objective: \mathbf{z}^* \in \arg\min_{\mathbf{z}} ||\mathbf{x} - \mathbf{z}||
   \mathbf{U}\mathbf{z}\|_{2}, s.t. \|\mathbf{z}\|_{0} \leq K 1. init: z \leftarrow 0, r \leftarrow x 2. whi-
   le \|\mathbf{z}\|_0 < K do 3. select atom index with smal-
  lest angle i^* = \operatorname{arg\,max}_i |\langle \mathbf{u}_i, \mathbf{r} \rangle| 4. update coeffi-
  cients: z_{i^*} \leftarrow z_{i^*} + \langle \mathbf{u}_{i^*}, \mathbf{r} \rangle 5. update residual: \mathbf{r} \leftarrow
  \mathbf{r} - \langle \mathbf{u}_{i^*}, \mathbf{r} \rangle \mathbf{u}_{i^*}.
  Exact recovery when: K < 1/2(1 + 1/m(\mathbf{U}))
  Compressive Sensing: Compress data while ga-
 thering: • \mathbf{x} \in \mathbb{R}^D, K-sparse in o.n.b. U. \mathbf{y} \in \mathbb{R}^M
 with y_i = \langle \mathbf{w}_i, \mathbf{x} \rangle: M lin. combinations of signal;
 \mathbf{y} = \mathbf{W}\mathbf{x} = \mathbf{W}\mathbf{U}\mathbf{z} = \Theta\mathbf{z}, \, \Theta \in \mathbb{R}^{M \times D} \bullet \text{Reconstruct}
 \mathbf{x} \in \mathbb{R}^D from \mathbf{y}; find \mathbf{z}^* \in \arg\min_{\mathbf{z}} ||\mathbf{z}||_0, s.t. \mathbf{y} = \Theta \mathbf{z}
  (e.g. with MP, or convex it with 1-norm: can be
  eq.!). Given z, reconstruct x = Uz
 Any orthogonal U sufficient if: \bullet W = Gaussi-
 an random projection, i.e. w_{ij} \sim \mathcal{N}(0, \frac{1}{D}) \cdot M
   \geq cKlog(\frac{D}{K}), where c is some constant
   8 Dictionary Learning
  Adapt dict. to signal characteristics. Obj: (\mathbf{U}^{\star}, \mathbf{Z}^{\star}) \in
 \underset{\mathbf{U},\mathbf{Z}}{\operatorname{arg\,min}}_{\mathbf{U},\mathbf{Z}} \|\mathbf{X} - \mathbf{U} \cdot \mathbf{Z}\|_F^2 not jointly convex but
convex in either. Matrix Fact. by Iter Greedy
 Min. 1. Coding step: \mathbf{Z}^{t+1} \in \operatorname{arg\,min}_{\mathbf{Z}} \|\mathbf{X} - \mathbf{U}^t \mathbf{Z}\|_F^2
 subject to Z being sparse (\mathbf{z}_n^{t+1} \in \operatorname{arg\,min}_{\mathbf{z}} \|\mathbf{z}\|_0
s.t.\|\mathbf{x}_n - \mathbf{U}^t \mathbf{z}\|_2 \le \sigma \|\mathbf{x}_n\|_2 2. Dict update step:
 \mathbf{U}^{t+1} \in \arg\min_{\mathbf{U}} \|\mathbf{X} - \mathbf{U}\mathbf{Z}^{t+1}\|_F^2, subj to \forall l \in [L]: \ \zeta \sim \mathcal{N}(\mathbf{0}, \mathbf{I}), \mathbf{z} = \mu + U\zeta then z \sim \mathcal{N}(\mu, \mathbf{U}^\top \mathbf{U})
 \|\mathbf{u}_l\|_2 = 1. (set \mathbf{U} = [\mathbf{u}_1^t \cdots \mathbf{u}_l \cdots \mathbf{u}_l^t], \min_{u_l} \|\mathbf{X} - \mathbf{u}_l\|_2
 \|\mathbf{U}\mathbf{Z}^{t+1}\|_F^2 = \min_{u_l} \|\mathbf{R}_l^t - \mathbf{u}_l(\mathbf{z}_l^{t+1})^\top\|_F^2 \text{ with } \mathbf{R}_l^t = \mathbf{u}_l^T \mathbf{v}_l^T \mathbf{v}_l^T
\tilde{\mathbf{U}}\Sigma\tilde{\mathbf{V}}^{\top} by \mathbf{u}_{i}^{*}=\tilde{\mathbf{u}}_{1})
  9 Neural Networks
 Activation: scalar, non-linear tanh(x) = \frac{e^x - e^{-x}}{e^x + e^{-x}}
sigmoid: \sigma(x) = \frac{1}{1+\sigma^{-x}}, \sigma'(x) = \sigma(x)(1-\sigma(x))
Neurons: F_{\sigma}(\mathbf{x}; \mathbf{w}) = \sigma(w_0 + \sum_{i=1}^{M} x_i w_i) = \sigma(w^{\top} x)
 Output: linear regression \mathbf{y} = \mathbf{W}^{L} \mathbf{x}^{L-1}, binary (lo-
  gistic) y_1 = P[Y = 1 | \mathbf{x}] = \frac{1}{1 + \exp(-\mathbf{w}^T \mathbf{x}^{L-1})}, multiclass
(soft-max) y_k = P[Y = k | \mathbf{x}] = \frac{\exp(\mathbf{w}_k^T \mathbf{x}^{\hat{L}-1})}{\sum_{m=1}^K \exp(\mathbf{w}_k^T \mathbf{x}^{\hat{L}-1})}. Loss
l(y,\hat{y}): squared loss \frac{1}{2}(y-\hat{y})^2, cross-entropy loss
  -y\log \hat{y} - (1-y)\log(1-\hat{y}) \ 0 \le \hat{y} \le 1, y \in \{0,1\}
 or y \in [0,1]. layer-wise: \mathbf{x}^l = \sigma^l \left( \mathbb{W}^{(l)} \mathbf{x}^{(l-1)} \right)
```

Overcomplete Basis

 $\mathbf{U} \in \mathbb{R}^{D \times L}$ for # atoms = $L > D = \dim(\text{data})$. Deco-

ding involved \rightarrow add constraint $\mathbf{z}^* \in \arg\min_{\mathbf{z}} ||\mathbf{z}||_0$

Backpropagation Layer-to-layer Jacobian: $\mathbf{x} = \text{prev.}$ layer activation, \mathbf{x}^+ = next layer activation. Jacobian matrix $\mathbf{J} = J_{ij}$ of mapping $\mathbf{x} \to \mathbf{x}^+$, $\mathbf{x}_i^+ = \sigma(\mathbf{w}_i^\top \mathbf{x})$, $J_{ij} = \frac{\partial \mathbf{x}_i^{\top}}{\partial \mathbf{x}_i} = w_{ij} \cdot \mathbf{\sigma}'(\mathbf{w}_i^{\top} \mathbf{x})$. Across multiple layers: $\frac{\partial \mathbf{x}^{(l)}}{\partial \mathbf{x}^{(l-n)}} = \mathbf{J}^{(l)} \cdot \frac{\partial \mathbf{x}^{(l-1)}}{\partial \mathbf{x}^{(l-n)}} = \mathbf{J}^{(l)} \cdot \mathbf{J}^{(l-1)} \cdots \mathbf{J}^{(l-n+1)}$ and then back prop. $\nabla_{\mathbf{y}^{(l)}}^{\top} \ell = \nabla_{\mathbf{y}}^{\top} \ell \cdot \mathbf{J}^{(L)} \cdots \mathbf{J}^{(l+1)}$ Weights: $\frac{\partial l}{\partial w_{ij}^{(l)}} = \frac{\partial l}{\partial x_i^{(l)}} \frac{\partial x_i^{(l)}}{\partial w_{ij}^{(l)}}, \frac{\partial x_i^l}{\partial w_{ij}^l} =$ $\sigma'([\mathbf{w}_{i}^{(l)}]^T\mathbf{x}^{(l-1)}) \cdot x_{i}^{(l-1)}$ (sensitivity of downstream unit · activation of up-stream unit)< **Gradient Descent (or Deepest Descent) Gradient:** $\nabla f(\mathbf{x}) := \left(\frac{\partial f(\mathbf{x})}{\partial \mathbf{x}_1}, \dots, \frac{\partial f(\mathbf{x})}{\partial \mathbf{x}_D}\right)$ $\mathbf{x}^{(t+1)} = \mathbf{x}^{(t)} - \gamma \nabla f(\mathbf{x}^{(t)})$, usually $\gamma \approx \frac{1}{4}$ **SGD** Assume additive obj. $f(x) = \frac{1}{N} \sum_{n=1}^{N} f_n(x)$ sample $n \in_{u,a,r} \{1,\ldots,N\}$, then $\mathbf{x}^{(t+1)} = \mathbf{x}^{(t)} - \gamma \nabla f_n(\mathbf{x}^{(t)})$, typically $\gamma \approx \frac{1}{2}$. **Neural Networks for Images (CNN)** $F_{n,m}(\mathbf{x};\mathbf{w}) = \sigma(b + \sum_{k=-2}^{2} \sum_{l=-2}^{2} w_{k,l} x_{n+k,m+l}).$ 10 Deep Unsupervised Learning **AR:** Image $p(\mathbf{x}) = \prod_{i=1}^{n^2} p(x_i|x_1,\dots,x_{i-1})$ **ELBO:** $\mathbb{E}_{x \sim P_{\mathbf{x}}} \left[\mathbb{E}_{z \sim O} \log P_{g}(x|z) - D_{KL}(Q(z|x)||P(z)) \right]$ Q enc. posterior distr., P(z) prior distr. on latent var z, P_{ρ} likelihood of dec. generated x. Jointly trained: enc. optimize regularizer term, sample $\mathbf{z} \sim Q$, feed to dec., produce \hat{x} to max. reconstruction quality. Both terms diff'able, can use SGD to train endto-end. **Reparam. trick:** use variational distr.s s.t. $q_{\phi}(\mathbf{z}; \mathbf{x}) = g_{\phi}(\zeta; \mathbf{x})$ with eg. $\zeta \sim \mathcal{N}(\mathbf{0}, \mathbf{I})$ Example: