QUESTION 3

ABHAY SHANKAR K: CS21BTECH11001 & KARTHEEK TAMMANA: CS21BTECH11028

(I) Question: Derive the expression of likelihood and prior for a heteroscedastic setting for a single data point with input $\mathbf{x_n}$ and output t_n .

Solution:

(a) Consider the following formula for the target variable:

$$t = \mathbf{w}^{\mathbf{T}} \phi(\mathbf{x}) + \epsilon(\mathbf{x})$$

where ϵ is a zero mean Gaussian random variable with precision $\beta(\mathbf{x})$ and ϕ is a deterministic function. Due to heteroscedasticity, the Gaussian noise is dependent on the input \mathbf{x} .

Due to the properties of Gaussian distribution, t is also normal, with its distribution i.e. the likelihood function given by:

$$p(t_n|\mathbf{x}_n, \mathbf{w}, \boldsymbol{\beta}) = \sqrt{\frac{\beta_n}{2\pi}} \exp\left\{-\frac{\beta_n}{2} \left(t_n - \mathbf{w}^{\mathbf{T}} \boldsymbol{\phi}(\mathbf{x}_n)\right)^2\right\}$$
(1)

So,

$$p(t_n|\mathbf{x_n}, \mathbf{w}, \boldsymbol{\beta}) = \mathcal{N}(t_n|\mathbf{w^T}\boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1})$$

(b) We may assume a Gaussian prior for \mathbf{w} , with arbitrary mean $\mathbf{m_0}$ and covariance matrix $\mathbf{S_0}$ in which case the prior is given by

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w}|\mathbf{m_0}, \mathbf{S_0})$$

(II) Question: Provide the expression for the objective function that you will consider for the ML and MAP estimation of the parameters considering a data set of size N.

Solution:

(a) To express this more succinctly, we define the design matrix

$$oldsymbol{\Phi} = egin{pmatrix} oldsymbol{\phi}(\mathbf{x}_1)^T \ dots \ oldsymbol{\phi}(\mathbf{x}_n)^T \end{pmatrix}$$

the data set $\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_n\}$ and the diagonal weighing matrix \mathbf{R} with $R_{ii} = \beta_i$.

Thus, the objective function for ML estimation is given by:

$$p(\mathbf{t}|\mathbf{X}, \mathbf{w}, \boldsymbol{\beta}) = \prod_{n=1}^{N} p(t_n | \mathbf{x_n}, \mathbf{w}, \boldsymbol{\beta})$$

$$= \prod_{n=1}^{N} \mathcal{N}(t_n | \mathbf{w^T} \boldsymbol{\phi}(\mathbf{x}_n), \beta_n^{-1})$$

$$= \frac{|\mathbf{R}|}{(2\pi)^{\frac{N}{2}}} \exp\left\{-\frac{1}{2} \sum_{n=1}^{N} \beta_n \left(t_n - \mathbf{w^T} \boldsymbol{\phi}(\mathbf{x}_n)\right)^2\right\}$$

$$= \frac{|\mathbf{R}|}{(2\pi)^{\frac{N}{2}}} \exp\left\{-\frac{1}{2} \left(\mathbf{t} - \boldsymbol{\Phi} \boldsymbol{w}\right)^T \mathbf{R} \left(\mathbf{t} - \boldsymbol{\Phi} \boldsymbol{w}\right)\right\}$$

$$= \mathcal{N}(\mathbf{t}|\boldsymbol{\Phi} \mathbf{w}, \mathbf{R}^{-1})$$
(2)

Date: October 9, 2023.

(b) Given

$$p(\mathbf{w}) = \mathcal{N}(w|\mathbf{m_0}, \mathbf{S_0})$$

and

$$p(\mathbf{t}|\mathbf{w}) = \mathcal{N}(\mathbf{\Phi}\mathbf{w}, \mathbf{R}^{-1})$$

Let $\mathbf{z} = \begin{pmatrix} \mathbf{w} \\ \mathbf{t} \end{pmatrix}$ as follows:

$$\ln p(\mathbf{z}) = \ln p(\mathbf{w}) + \ln p(\mathbf{t}|\mathbf{w})$$

$$= -\frac{1}{2} (\mathbf{w} - \mathbf{m}_0)^T \mathbf{S}_0^{-1} (\mathbf{w} - \mathbf{m}_0)$$

$$-\frac{1}{2} (\mathbf{t} - \mathbf{\Phi} \mathbf{w})^T \mathbf{R} (\mathbf{t} - \mathbf{\Phi} \mathbf{w}) + \text{constant}$$

$$= -\frac{1}{2} [\mathbf{w}^T \mathbf{S}_0^{-1} \mathbf{w} - 2\mathbf{w}^T \mathbf{S}_0^{-1} \mathbf{m}_0]$$

$$-\frac{1}{2} [\mathbf{t}^T \mathbf{R} \mathbf{t} - 2\mathbf{t}^T \mathbf{R} \mathbf{\Phi} \mathbf{w} + \mathbf{w}^T \mathbf{\Phi}^T \mathbf{R} \mathbf{\Phi} \mathbf{w}] + \text{constant}$$
(3)

where we have removed terms which are constant with respect to \mathbf{w} and \mathbf{t} .

Furthermore, due to linearity of expectation, we have

$$\mathbb{E}\left[\mathbf{z}
ight] = egin{pmatrix} \mathbb{E}\left[\mathbf{w}
ight] \\ \mathbb{E}\left[\mathbf{t}
ight] \end{pmatrix}$$

from which

$$cov[\mathbf{z}] = \begin{pmatrix} var[\mathbf{w}] & cov[\mathbf{w}, \mathbf{t}] \\ cov[\mathbf{t}, \mathbf{w}] & var[\mathbf{t}] \end{pmatrix}$$

From (3), it is clear that $p(\mathbf{z})$ is a Gaussian distribution. Now we complete the square.

To find the covariance of $\mathbf{w}|\mathbf{t}$, we consider the single term of second order in \mathbf{w} from (3):

$$\frac{1}{2}\mathbf{w^T}\boldsymbol{\Sigma}^{-1}\mathbf{w} = \frac{1}{2}\mathbf{w^T}(\mathbf{S_0}^{-1} + \boldsymbol{\Phi^T}\mathbf{R}\boldsymbol{\Phi})\mathbf{w}$$

We treat \mathbf{t} as a constant.

Thus, the covariance is given by

$$\mathbf{\Sigma} = (\mathbf{S_0}^{-1} + \mathbf{\Phi^T} \mathbf{R} \mathbf{\Phi})^{-1}$$

Similarly, we may obtain μ using the terms of (3) of first order in w. We have

$$\mathbf{w^T} \boldsymbol{\Sigma}^{-1} \boldsymbol{\mu} = \mathbf{w^T} \mathbf{S_0}^{-1} \mathbf{m_0} + \mathbf{w^T} \boldsymbol{\Phi^T} \mathbf{R} \mathbf{t}$$

which yields

$$\boldsymbol{\mu} = \boldsymbol{\Sigma} (\mathbf{S_0}^{-1} \mathbf{m_0} + \boldsymbol{\Phi^T} \mathbf{R} \mathbf{t})$$

Thus, the MAP objective function is

$$p(\mathbf{w}|\mathbf{t}) = \mathcal{N}(\mathbf{w}|\boldsymbol{\mu}, \boldsymbol{\Sigma})$$

(III) Question: Show

$$E_{\mathcal{D}} = \sum_{n=1}^{N} r_n \left\{ t_n - \mathbf{w}^{\mathbf{T}} \phi(\mathbf{x_n}) \right\}^2$$

and find w that minimizes $E_{\mathcal{D}}$. Solution:

QUESTION 3

(a) Taking logarithm of (1)

$$-\ln p(\mathcal{D}|\mathbf{w}) = \frac{N}{2} \ln \beta_n - \frac{N}{2} \ln(2\pi) + E_{\mathcal{D}}$$

with

$$E_{\mathcal{D}} = \frac{1}{2} \sum_{n=1}^{N} \beta_n \left\{ t_n - \mathbf{w}^{\mathbf{T}} \phi(\mathbf{x_n}) \right\}^2$$
(4)

Setting $r_n = \frac{\beta_n}{2}$, we obtain the desired equation.

(b) Now, we may obtain the **w** that minimizes $E_{\mathcal{D}}$ by differentiating (4) and setting the derivative to zero like so:

$$\sum_{n=1}^{N} \beta_n \left\{ t_n - \mathbf{w}^{\mathbf{T}} \phi(\mathbf{x_n}) \right\} \phi(\mathbf{x_n})^{\mathbf{T}} = 0$$

Whence

$$\sum_{n=1}^{N} t_n \beta_n \phi\left(\mathbf{x_n}\right)^T = \mathbf{w^T} \sum_{n=1}^{N} \beta_n \phi\left(\mathbf{x_n}\right) \phi\left(\mathbf{x_n}\right)^T$$

Using the matrix \mathbf{R} defined as before, taking the transpose of both sides yields

$$\Phi^{\mathbf{T}}\mathbf{R}\mathbf{t} = (\Phi^{\mathbf{T}}\mathbf{R}\Phi)\mathbf{w}$$

$$\implies \mathbf{w}_{\mathbf{ML}} = (\Phi^{\mathbf{T}}\mathbf{R}\Phi)^{-1}\Phi^{\mathbf{T}}\mathbf{R}\mathbf{t}$$
(5)