

*Student Name:* Abhishek Pardhi

*Roll Number:* 200026

*Date:* May 2, 2023

It has been given that  $\alpha$ -divergence is

$$D_{\alpha}(p||q) = \frac{4}{1-\alpha^2} \left( 1 - \int p(z)^{\frac{(1+\alpha)}{2}} q(z)^{\frac{(1-\alpha)}{2}} dz \right)$$

Let's find  $\lim_{\alpha \rightarrow 1} D_{\alpha}(p||q)$ , but before that notice that the numerator as well as the denominator tends to 0, hence we can use L'hospital rule to find this kind of limit:

{using the result  $(a^x)' = a^x \ln a$ }

$$\Rightarrow \lim_{\alpha \rightarrow 1} \frac{-4}{1-\alpha^2} \int \left( \frac{\ln p(z)}{2} p(z)^{\frac{(1+\alpha)}{2}} q(z)^{\frac{(1-\alpha)}{2}} - \frac{\ln q(z)}{2} p(z)^{\frac{(1+\alpha)}{2}} q(z)^{\frac{(1-\alpha)}{2}} \right) dz$$

$$\Rightarrow \lim_{\alpha \rightarrow 1} \frac{1}{\alpha} \int \ln \frac{p(z)}{q(z)} p(z)^{\frac{(1+\alpha)}{2}} q(z)^{\frac{(1-\alpha)}{2}} dz$$

$$\Rightarrow \int p(z) \ln \frac{p(z)}{q(z)} dz = - \int p(z) \ln \frac{q(z)}{p(z)} dz = KL(p||q)$$

Hence,  $KL(p||q)$  corresponds to  $\alpha$ -divergence as  $\alpha \rightarrow 1$ .

Now similarly, we can find  $\lim_{\alpha \rightarrow -1} D_{\alpha}(q||p)$ :

$$\Rightarrow \lim_{\alpha \rightarrow -1} \frac{1}{\alpha} \int \ln \frac{p(z)}{q(z)} p(z)^{\frac{(1+\alpha)}{2}} q(z)^{\frac{(1-\alpha)}{2}} dz$$

$$\Rightarrow - \int q(z) \ln \frac{p(z)}{q(z)} dz = KL(q||p)$$

Hence,  $KL(q||p)$  corresponds to  $\alpha$ -divergence as  $\alpha \rightarrow -1$ .

Student Name: Abhishek Pardhi

Roll Number: 200026

Date: May 2, 2023

Given:

- Likelihood  $p(x_n|\mu, \tau) = \mathcal{N}(\mu, \tau^{-1})$
- Priors  $p(\mu) = \frac{1}{\sigma_\mu}, p(\tau) = \frac{1}{\tau}$
- let  $\mathbf{X} = [x_1, x_2, \dots, x_N]$

The log-joint will be:

$$\Rightarrow \log p(\mathbf{X}, \mu, \tau) = \log p(\mathbf{X}|\mu, \tau) + \log p(\mu|\tau) + \log p(\tau)$$

Since  $\mu$  doesn't depend on  $\tau$  we can write the above expression as:

$$\Rightarrow \log p(\mathbf{X}, \mu, \tau) = \log p(\mathbf{X}|\mu, \tau) + \log p(\mu) + \log p(\tau)$$

Let's find  $q_\mu^*(\mu)$  by using the above expression (only keeping terms that involve  $\mu$ )

$$\log q_\mu^*(\mu) = \mathbb{E}_{q_\tau}[\log p(\mathbf{X}|\mu, \tau) + \log p(\mu)] + \text{const}$$

$$\Rightarrow -\frac{\mathbb{E}_{q_\tau}[\tau]}{2} \sum_{n=1}^N (x_n - \mu)^2 + \text{const}$$

Therefore, using [1] and putting  $\lambda_0 = 0$ , we get:

$$q_\mu^* = \mathcal{N}(\mu|\mu_N, \lambda_N) \text{ where } \mu_N = \bar{x} \text{ and } \lambda_N = N\mathbb{E}_{q_\tau}[\tau]$$

Let's find  $q_\tau^*(\tau)$  by using the log-joint expression (only keeping terms that involve  $\tau$ )

$$\log q_\tau^*(\tau) = \mathbb{E}_{q_\mu}[\log p(\mathbf{X}|\mu, \tau) + \log p(\tau)] + \text{const}$$

$$\Rightarrow -\frac{\tau}{2} \sum_{n=1}^N \mathbb{E}_{q_\mu}[(x_n - \mu)^2] + \frac{N}{2} \log \tau - \log \tau + \text{const}$$

$$q_\tau^*(\tau) \propto \tau^{\frac{N}{2}-1} \exp(-\frac{\tau}{2} \sum_{n=1}^N \mathbb{E}_{q_\mu}[(x_n - \mu)^2])$$

This is a Gamma function, therefore:

$$q_\tau^* = \text{Gamma}(\tau|a_N, b_N) \text{ where } a_N = \frac{N}{2} \text{ and } b_N = \frac{\sum_{n=1}^N \mathbb{E}_{q_\mu}[(x_n - \mu)^2]}{2}$$

Student Name: Abhishek Pardhi

Roll Number: 200026

Date: May 2, 2023

The conditional posterior of  $z_{dn}$ :

$$p(z_{dn} = k | \mathbf{W}, \mathbf{Z}_{-dn}) \propto p(z_{dn} = k | \mathbf{Z}_{-dn}) \times p(w_{dn} | z_{dn} = k, \mathbf{W}_{-dn}, \mathbf{Z}_{-dn})$$

The prior on  $z_{dn}$  is:

$$p(z_{dn} = k | \mathbf{Z}_{-dn}) = \int p(z_{dn} = k | \mathbf{Z}_{-dn}, \theta_d) p(\theta_d | \mathbf{Z}_{-dn}) d\theta_d$$

$$\Rightarrow \mathbb{E}_{p(\theta_d | \mathbf{Z}_{-dn})} [p(z_{dn} = k | \mathbf{Z}_{-dn}, \theta_d)]$$

$$\Rightarrow \mathbb{E}_{p(\theta_d | \mathbf{Z}_{-dn})} [\theta_{dk}]$$

let's find  $p(\theta_d | \mathbf{Z}_{-dn})$ :

$$p(\theta_d | \mathbf{Z}_{-dn}) \propto p(\theta_d) p(\mathbf{Z}_{-dn} | \theta_d) \Rightarrow \text{Dirichlet}(\alpha, \dots, \alpha) \prod_{i \neq n} \text{multinoulli}(\theta_d)$$

$$\Rightarrow \theta_{dk}^{(\alpha - 1 + \sum_{i \neq 1} \mathbb{I}[z_{di} = k])}$$

Or we can say:

$$p(\theta_d | \mathbf{Z}_{-dn}) = \text{Dirichlet}([\alpha + \sum_{i \neq n} \mathbb{I}[z_{di} = k]]_{k=1}^K)$$

Hence, we have the prior on  $z_{dn}$ :

$$p(z_{dn} | \mathbf{Z}_{-dn}) = \frac{\alpha + \sum_{i \neq n} \mathbb{I}[z_{di} = k]}{K\alpha - 1 + N_d}$$

Now let's find  $p(w_{dn} | z_{dn} = k, \mathbf{W}_{-dn}, \mathbf{Z}_{-dn})$ :

$$p(w_{dn} | z_{dn} = k, \mathbf{W}_{-dn}, \mathbf{Z}_{-dn}) = \int p(w_{dn} = v | \phi_k) p(\phi_k | \mathbf{Z}_{dn}, \mathbf{W}_{-dn}) d\phi_k$$

$$\Rightarrow \mathbb{E}_{p(\phi_k | \mathbf{Z}_{dn}, \mathbf{W}_{-dn})} [p(w_{dn} = v | \phi_k)]$$

$$\Rightarrow \mathbb{E}_{p(\phi_k | \mathbf{Z}_{dn}, \mathbf{W}_{-dn})} [\phi_{kv}]$$

let's find  $p(\phi_k | \mathbf{Z}_{dn}, \mathbf{W}_{-dn}) \propto p(\mathbf{W}_{-dn} | \phi_k, \mathbf{Z}_{dn}) p(\phi_k)$

$$\phi_k^\eta \prod_{i \neq n} \prod_{j \neq d} p(w_{ij} | \phi_k, z_{ij})$$

$$\Rightarrow \phi_k^{\eta + \sum_{i \neq n} \sum_{j \neq d} \mathbb{I}[w_{ij} = v] \mathbb{I}[z_{ij} = k]}$$

Or we can say:

$$p(\phi_k | \mathbf{Z}_{dn}, \mathbf{W}_{-dn}) = \text{Dirichlet}([\eta + \sum_{i \neq n} \sum_{j \neq d} \mathbb{I}[w_{ij} = v] \mathbb{I}[z_{ij} = k]]_{v=1}^V)$$

Hence, we have:

$$p(w_{dn} | z_{dn} = k, \mathbf{W}_{-dn}, \mathbf{Z}_{-dn}) = \frac{\eta + \sum_{i \neq n} \sum_{j \neq d} \mathbb{I}[w_{ij} = v] \mathbb{I}[z_{ij} = k]}{V\eta + \sum_{i \neq n} \sum_{j \neq d} \mathbb{I}[z_{ij} = k]}$$

Putting all of these together, we will get the posterior of  $z_{dn}$  as follows:

$$p(z_{dn} = k | \mathbf{W}, \mathbf{Z}_{-dn}) \propto \frac{\alpha + \sum_{i \neq n} \mathbb{I}[z_{di} = k]}{K\alpha - 1 + N_d} \times \frac{\eta + \sum_{i \neq n} \sum_{j \neq d} \mathbb{I}[w_{ij} = v] \mathbb{I}[z_{ij} = k]}{V\eta + \sum_{i \neq n} \sum_{j \neq d} \mathbb{I}[z_{ij} = k]}$$

#### Intuitive sense behind CP of $z_{dn}$

The probability of a word belonging to a particular topic is calculated by looking at the frequency of that word in the entire corpus, as well as in the current document (excluding current instance). The frequency of a word in the corpus helps to determine the topic vectors for the entire corpus, while the frequency in the document helps to determine the topic distribution for that particular document. Therefore, both the corpus and the document are taken into account to calculate the probability of a word belonging to a certain topic.

### Gibbs Sampler

- Initialize  $\mathbf{Z}^{(0)}$  and for  $t = 1..T$  do:
- $z_{dn}^{(t)} \sim p(z_{dn} = k | \mathbf{W}, \mathbf{Z}_{-dn})$

We can use *Monte Carlo* approximation to compute an approximation of  $\theta_{dk}$  and  $\phi_{kv}$ :  
First we take  $L$  samples  $\{Z^{(l)}\}_{l=1}^L$  from the *CP*  
then we find the approximation:

$$\mathbb{E}[\theta_{dk}] = \sum_{l=1}^L \frac{\alpha + \sum_{i \neq n} \mathbb{I}[z_{di}=k]}{K\alpha - 1 + N_d}$$
$$\mathbb{E}[\phi_{kv}] = \sum_{l=1}^L \frac{\eta + \sum_{i \neq n} \sum_{j \neq d} \mathbb{I}[w_{ij}=v] \mathbb{I}[z_{ij}=k]}{V\eta + \sum_{i \neq n} \sum_{j \neq d} \mathbb{I}[z_{ij}=k]}$$

### Intuitive meaning of the expectations

The expected value of the topic distribution  $\mathbb{E}[\theta_{dk}]$  for a given document  $d$  is determined by the number of words in the document assigned to the topic  $k$  based on the samples  $\mathbf{Z}(s)$ . Similarly,  $\mathbb{E}[\phi_{kv}]$  is determined by the number of times the word  $v$  appears in the corpus assigned to topic  $k$ , as well as the number of words assigned to topic  $k$  across the entire corpus.

*Student Name:* Abhishek Pardhi

*Roll Number:* 200026

*Date:* May 2, 2023

Given:  $p(r_{ij}|\mathbf{u}_i, \mathbf{v}_j) = \mathcal{N}(r_{ij}|\mathbf{u}_i^T \mathbf{v}_j, \beta^{-1})$  and  $p(r_{ij}|\mathbf{R}) = \int p(r_{ij}|\mathbf{u}_i, \mathbf{v}_j)p(\mathbf{u}_i, \mathbf{v}_j|\mathbf{R})d\mathbf{u}_i d\mathbf{v}_j$   
 Using Monte Carlo sampling, we can find approximation of mean and variance of any entry  $r_{ij}$  as follows:

$p(r_{i,j}|\mathbf{R}) = \frac{1}{S} \sum_{s=1}^S p(r_{ij}|\mathbf{u}_i^{(s)}, \mathbf{v}_j^{(s)})$  where  $\mathbf{u}_i^{(s)}, \mathbf{v}_j^{(s)}$  are known since we know  $\mathbf{U}^{(s)} = \{\mathbf{u}_i^{(s)}\}_{i=1}^N$  and  $\mathbf{V}^{(s)} = \{\mathbf{v}_i^{(s)}\}_{i=1}^N$

Let's prove the following two identities:

Suppose we need to find  $\mathbb{E}[f] = \int f(z)p(z)dz$ , and we approximated this expectation using monte-carlo approximation  $\hat{f} = \frac{1}{S} \sum_{s=1}^S f(z^{(s)})$

Claim 1:  $\mathbb{E}[\hat{f}] = \mathbb{E}[f]$

Proof:

$$\begin{aligned} \mathbb{E}[\hat{f}] &= \mathbb{E}\left[\frac{1}{S} \sum_{s=1}^S f(z^{(s)})\right] \\ &\Rightarrow \frac{1}{S} \sum_{s=1}^S \mathbb{E}[f(z^{(s)})] \\ &\Rightarrow \frac{1}{S} \sum_{s=1}^S \mathbb{E}[f] = \mathbb{E}[f] \end{aligned}$$

Hence, proved

Claim 2:  $\text{Var}[\hat{f}] = \frac{1}{S} \mathbb{E}[(f - \mathbb{E}[f])^2]$

Proof:

$$\begin{aligned} \text{Var}[\hat{f}] &= \mathbb{E}[\hat{f}^2] - \mathbb{E}[\hat{f}]^2 \\ &\Rightarrow \mathbb{E}\left[\left(\frac{1}{S} \sum_{s=1}^S f(z^{(s)})\right)^2\right] - \mathbb{E}\left[\left(\frac{1}{S} \sum_{s=1}^S f(z^{(s)})\right)\right]^2 \\ &\Rightarrow \mathbb{E}\left[\frac{1}{S^2} \sum_{s=1}^S f^2(z^{(s)}) + \frac{1}{S^2} \sum_{s=1}^S \sum_{l=1}^L f(z^{(s)})f(z^{(l)})\right] - \mathbb{E}\left[\frac{1}{S} \sum_{s=1}^S f(z^{(s)})\right]^2 \\ &\Rightarrow \frac{1}{S^2} \sum_{s=1}^S \mathbb{E}[f^2(z)] + \frac{1}{S^2} \sum_{s=1}^S \sum_{l=1}^L \mathbb{E}[f(z)]^2 - \mathbb{E}\left[\frac{1}{S} \sum_{s=1}^S f(z)\right]^2 \\ &\Rightarrow \frac{1}{S} \mathbb{E}[f^2(z)] - \frac{1}{S} \mathbb{E}[f(z)]^2 \\ &= \frac{1}{S} \text{Var}[f] \end{aligned}$$

Hence, proved

Now using above claims we get  $E[p(r_{i,j}|\mathbf{R})] = \mathbf{u}_i^T \mathbf{v}_j$

$$\begin{aligned} &\Rightarrow E[\hat{r}_{ij}] = E[r_{ij}] = \frac{1}{S} \sum_{s=1}^S \mathbf{u}_i^{(s)T} \mathbf{v}_j^{(s)} \\ &\Rightarrow \text{Var}[\hat{r}_{ij}] = \frac{1}{S} \text{Var}[r_{ij}] = \frac{1}{\beta S} \end{aligned}$$

Hence, the required values are  $E[\hat{r}_{ij}]$  and  $\text{Var}[\hat{r}_{ij}]$ .

**Part 1:**

First we tune the value of  $\sigma$  such that we get as smaller value of  $M$  as possible which almost envelopes the distribution. We find the value of  $M$  by finding the largest possible value of  $\frac{\tilde{p}(x)}{q(x)}$  and assigning it to  $M$ . This is done in the python notebook. Below is the graph for Rejection Sampler where it can be seen that the samples from the Rejection Sampler(blue) resembles the actual distribution(green). The black dotted line is  $Mq(x)$ . The actual distribution is plotted by dividing the pdf values by area under the curve of the distribution.

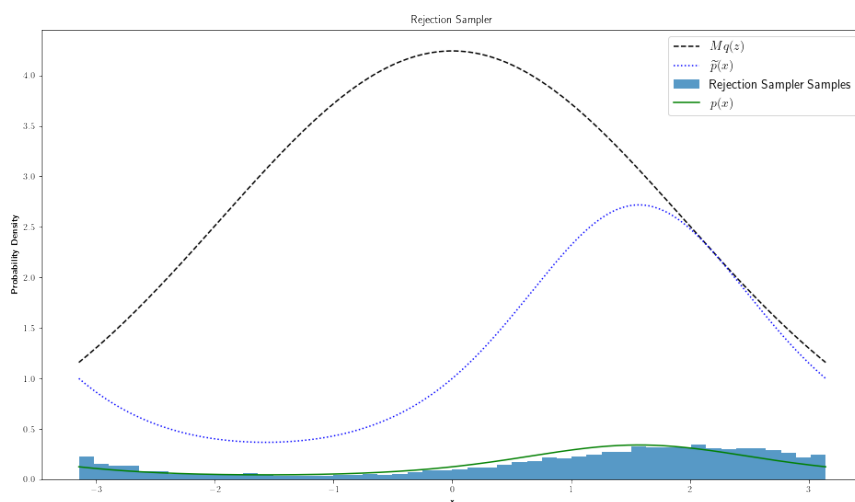


Figure 1: Rejection Sampler

**Part 2:**

- $\sigma^2 = 1$  is giving good approximation of the 2-D Gaussian where as  $\sigma^2 = 0.01$  is performing badly.  $\sigma^2 = 100$  is giving almost same approximation as  $\sigma^2 = 1$  but took a long time to build due to high rejection rate.
- For  $\sigma^2 = 0.01$ , Rejection rate = 9.7%
- For  $\sigma^2 = 1$ , Rejection rate = 59.7%
- For  $\sigma^2 = 100$ , Rejection rate = 98.8%

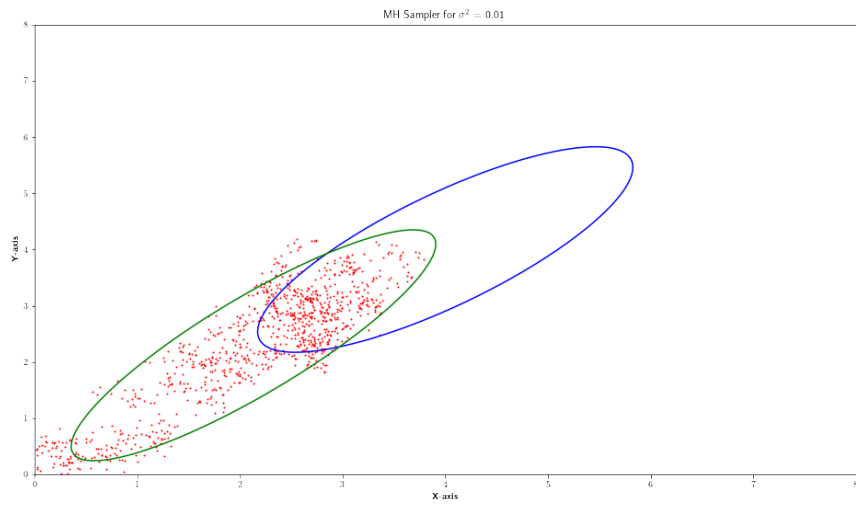


Figure 2:  $\sigma^2 = 0.01$

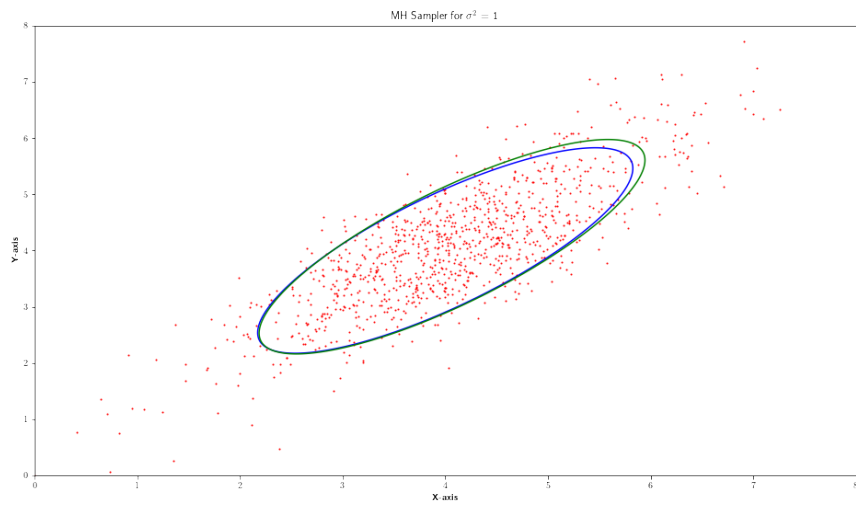


Figure 3:  $\sigma^2 = 1$

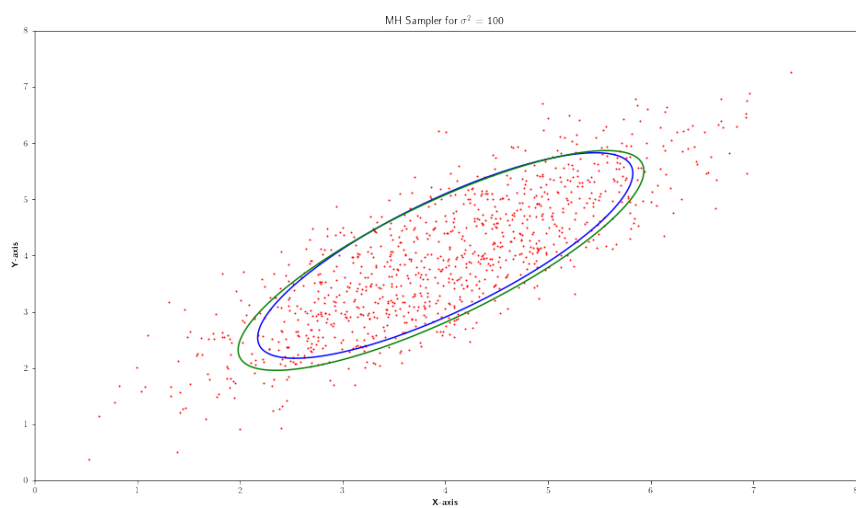


Figure 4:  $\sigma^2 = 100$



## References

- [1] Piyush Rai. *CS772 Lecture*. Slide 13, Page 29.