Machine Learning Homework 2

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1 Optimising Likelihood: Monotonic Transforms

Problem 1

$$\begin{split} f &= \theta^t (1-\theta)^h \\ \text{First Derivative:} \\ \frac{df}{d\theta} &= t(1-\theta)^h \theta^{t-1} - h \theta^t (1-\theta)^{h-1} \\ \text{Second Derivative:} \\ \frac{d}{d\theta} \frac{df}{d\theta} &= h t \theta^{t-1} (1-\theta)^{h-1} + t (t-1) \theta^{t-2} (1-\theta)^h - \{h t \theta^{t-1} (1-\theta)^{h-1} + h (h-1) \theta^t (1-\theta)^{h-2} \} \\ f &= \log(\theta^t (1-\theta)^h) \\ f &= t \log(\theta) + h \log(1-\theta) \\ \text{First Derivative:} \\ \frac{df}{d\theta} &= \frac{t}{\theta} + \frac{h}{1-\theta} \\ \text{Second Derivative:} \\ \frac{d}{d\theta} \frac{df}{d\theta} &= \frac{-t}{\theta^2} + \frac{h}{(1-\theta)^2} \end{split}$$

Problem 2

Let local maxima for any positive differentiable function $f(\theta)$ be θ^* . Then θ^* satisfies the condition:

1.
$$\frac{d}{d\theta}f(\theta^*)=0$$

$$2. \ \frac{d^2}{d\theta^2} f(\theta^*) < 0$$

For the function $log(f(\theta))$, we find the first derivative and equate to zero,

$$\frac{d}{d\theta}log(f(\theta)) = \frac{1}{f(\theta)}\frac{d}{d\theta}f(\theta)$$

For a positive function $f(\theta)$ is positive and from 1, $\frac{d}{d\theta}f(\theta) = 0$ at θ^* . Also, the second derivative of $log(f(\theta))$ is

$$\frac{d^2}{d\theta^2}log(f(\theta)) = -\frac{1}{f(\theta)^2}\frac{d}{d\theta}f(\theta) + \frac{1}{f(\theta)}\frac{d^2}{d\theta^2}f(\theta)$$

In the above equation we can notice that the first term is negative and from 2 second term is also negative hence proving that θ^* is a local maximum of both $f(\theta)$ and $log(f(\theta))$

From the two problem we understand that Monotonic functions preserve the critical points. Conversion of the problem to a log scale makes the computation easy and is preferred.

2 Properties of MLE and MAP

Problem 3

from Bayes rule we know that

$$p(\theta|D) = \frac{p(D|\theta)p(\theta)}{p(D)} \tag{1}$$

now we know that $p(\theta|D)$ is the prior estimate θ_{MAP} and $p(D|\theta)$ is the Maximum Likelihood Estimate θ_{MLE} so we get the relation

 $posterior \propto likelihood \times prior$

for a uniform prior $p(\theta)$ we can say that θ_{MLE} is a special case of θ_{MAP}

Problem 4

 $posterior \propto likelihood \propto prior$

The prior is beta distribution which is conjugate prior for the binomial distribution of the likelihood. Therefore, from lectures the distribution of the posterior is same as the prior distribution.

posterior
$$\propto {}^{N}C_{m}\theta^{m}(1-\theta)^{N-m} * \frac{\Gamma(a+b)}{\Gamma(a)+\Gamma(b)}\theta^{a-1}(1-\theta)^{N-m}$$

by simple pattern matching with Beta distribution we get posterior mean as

$$\mathbb{E}[\theta|\mathcal{D}] = \frac{m+a}{m+l+a+b} \tag{2}$$

From the mean of the beta distribution we get

$$\mathbb{E}[p(\theta)] = \frac{a}{a+b} \tag{3}$$

Also θ_{MLE} From the Eq. 2

$$\frac{m+a}{m+l+a+b} = \frac{m}{m+l+a+b} + \frac{a}{m+l+a+b}$$

$$= \frac{m}{m+l} \frac{m+l}{m+l+a+b} + \frac{a}{a+b} \frac{a+b}{m+l+a+b}$$

$$= (1-\lambda) \frac{m}{m+l} + \lambda \frac{a}{a+b}$$
(4)
(5)

$$= \frac{m}{m+l} \frac{m+l}{m+l+a+b} + \frac{a}{a+b} \frac{a+b}{m+l+a+b}$$

$$= (1-\lambda) \frac{m}{m+l} + \lambda \frac{a}{a+b}$$
(6)

Hence the posterior mean value $\mathbb{E}[\theta|\mathcal{D}]$ lies between prior mean of θ and the θ_{MLE} .

3 Poisson Distribution

Problem 5

Part a

The $\mathbb{E}[\lambda]$ is the mean of the Poisson distribution which is λ , which gives the maximum likelihood estimate.

Part b

It is known that the conjugate prior of the Poisson distribution is $Gamma(\alpha, \beta)$. Therefore our posterior is also $Gamma(\alpha, \beta)$.

So,

$$p(\lambda|\mathcal{D}) \qquad \propto p(\mathcal{D}|\lambda) * p(\lambda)$$

$$= e^{(-n\lambda)} \prod_{i=1}^{n} \frac{\lambda^{k_i}}{k_i!} e^{-\beta\lambda}$$

$$\propto \lambda^{\sum_{i=1}^{n} k_i + \alpha - 1} e^{-n\lambda - \beta\lambda}$$

By simple pattern matching the mode gamma distribution, we get

$$\lambda = \frac{\sum_{i=1}^{n} k_i + \alpha - 1}{n + \beta} \tag{7}$$