

# Forecasting: principles and practice

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3.1 Dynamic regression

#### **Outline**

- 1 Regression with ARIMA errors
- 2 Lab session 19
- 3 Some useful predictors for linear models
- 4 Dynamic harmonic regression
- 5 Lab session 20
- **6** Lagged predictors

#### **Regression models**

$$y_t = \beta_0 + \beta_1 x_{1,t} + \cdots + \beta_k x_{k,t} + \varepsilon_t,$$

- y<sub>t</sub> modeled as function of k explanatory variables  $x_{1,t}, \ldots, x_{k,t}$ .
- In regression, we assume that  $\varepsilon_t$  was WN.
- Now we want to allow  $\varepsilon_t$  to be autocorrelated.

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- In regression, we assume that  $\varepsilon_t$  was WN.
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#### **Example: ARIMA(1,1,1) errors**

$$y_{t} = \beta_{0} + \beta_{1} x_{1,t} + \dots + \beta_{k} x_{k,t} + \eta_{t},$$
  
$$(1 - \phi_{1} B)(1 - B)\eta_{t} = (1 + \theta_{1} B)\varepsilon_{t},$$

where  $\varepsilon_t$  is white noise.

#### **Residuals and errors**

#### Example: $\eta_t$ = ARIMA(1,1,1)

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#### Example: $\eta_t$ = ARIMA(1,1,1)

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 $(1 - \phi_1 B)(1 - B)\eta_t = (1 + \theta_1 B)\varepsilon_t,$ 

- Be careful in distinguishing  $\eta_t$  from  $\varepsilon_t$ .
- Only the errors  $\eta_t$  are assumed to be white noise.
- In ordinary regression,  $\eta_t$  is assumed to be white noise and so  $\eta_t = \varepsilon_t$ .

4

#### **Estimation**

If we minimize  $\sum \eta_t^2$  (by using ordinary regression):

- Estimated coefficients  $\hat{\beta}_0, \dots, \hat{\beta}_k$  are no longer optimal as some information ignored;
- 2 Statistical tests associated with the model (e.g., t-tests on the coefficients) are incorrect.
- *p*-values for coefficients usually too small ("spurious regression").
- 4 AIC of fitted models misleading.

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- Statistical tests associated with the model (e.g., t-tests on the coefficients) are incorrect.
- *p*-values for coefficients usually too small ("spurious regression").
- 4 AIC of fitted models misleading.
  - Minimizing  $\sum \varepsilon_t^2$  avoids these problems.
  - Maximizing likelihood is similar to minimizing  $\sum \varepsilon_t^2$ .

# **Stationarity**

#### **Regression with ARMA errors**

$$y_t = \beta_0 + \beta_1 x_{1,t} + \cdots + \beta_k x_{k,t} + \eta_t,$$
 where  $\eta_t$  is an ARMA process.

- If we estimate the model while any variable is non-stationary, the estimated coefficients can be incorrect.
- Difference variables until all stationary.
- If necessary, apply same differencing to all variables to preserve interpretability.

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#### **Original data**

$$\begin{aligned} \mathbf{y}_t &= \beta_0 + \beta_1 \mathbf{x}_{1,t} + \dots + \beta_k \mathbf{x}_{k,t} + \eta_t \\ \text{where} \quad \phi(\mathbf{B}) (1 - \mathbf{B})^d \eta_t &= \theta(\mathbf{B}) \varepsilon_t \end{aligned}$$

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#### After differencing all variables

$$\begin{aligned} \mathbf{y}_t' &= \beta_1 \mathbf{x}_{1,t}' + \dots + \beta_k \mathbf{x}_{k,t}' + \eta_t'. \\ \text{where} \quad \phi(\mathbf{B}) \eta_t &= \theta(\mathbf{B}) \varepsilon_t \\ \text{and} \quad \mathbf{y}_t' &= (\mathbf{1} - \mathbf{B})^d \mathbf{y}_t \end{aligned}$$

#### **Model selection**

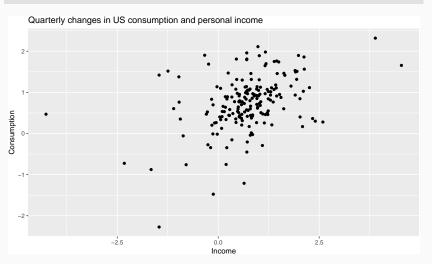
- Fit regression model with automatically selected ARIMA errors.
- Check that  $\varepsilon_t$  series looks like white noise.

#### **Selecting predictors**

- AICc can be calculated for final model.
- Repeat procedure for all subsets of predictors to be considered, and select model with lowest AICc value.

```
autoplot(uschange[,1:2], facets=TRUE) +
  xlab("Year") + ylab("") +
  ggtitle("Quarterly changes in US consumption and personal income")
   Quarterly changes in US consumption and personal income
 -1 -
 -2-
-2.5 -
     1970
                   1980
                                                 2000
                                                               2010
```

```
qplot(Income,Consumption, data=as.data.frame(uschange)) +
   ggtitle("Quarterly changes in US consumption and personal income")
```

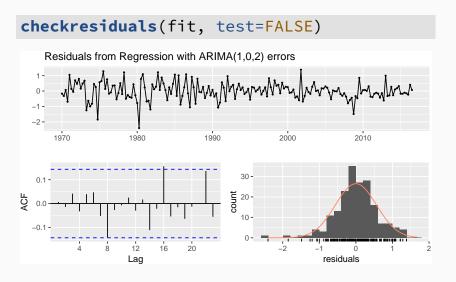


- No need for transformations or further differencing.
- Increase in income does not necessarily translate into instant increase in consumption (e.g., after the loss of a job, it may take a few months for expenses to be reduced to allow for the new circumstances). We will ignore this for now.

```
(fit <- auto.arima(uschange[,1], xreg=uschange[,2]))</pre>
## Series: uschange[, 1]
## Regression with ARIMA(1,0,2) errors
##
## Coefficients:
##
          ar1
                 ma1
                        ma2 intercept
                                        xreg
##
        0.692 - 0.576 0.198
                                 0.599
                                       0.203
## s.e. 0.116 0.130 0.076
                                 0.088
                                       0.046
##
## sigma^2 estimated as 0.322: log likelihood=-156.9
## AIC=325.9 AICc=326.4 BIC=345.3
```

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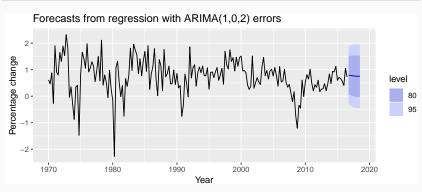
Write down the equations for the fitted model.



#### checkresiduals(fit, plot=FALSE)

```
##
## Ljung-Box test
##
## data: Residuals from Regression with ARIMA(1,0,2) errors
## Q* = 5.9, df = 3, p-value = 0.1
##
## Model df: 5. Total lags used: 8
```

```
fcast <- forecast(fit,
    xreg=rep(mean(uschange[,2]),8), h=8)
autoplot(fcast) + xlab("Year") +
    ylab("Percentage change") +
    ggtitle("Forecasts from regression with ARIMA(1,0,2) errors")</pre>
```

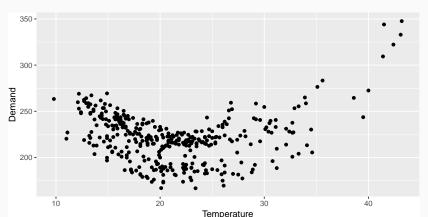


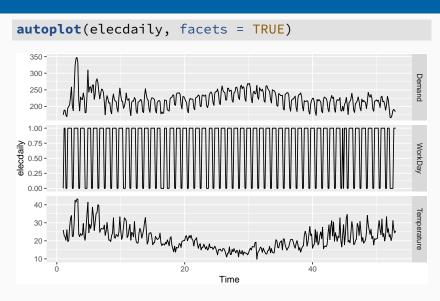
# **Forecasting**

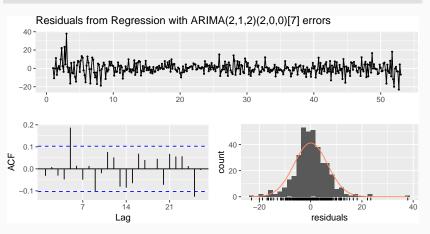
- To forecast a regression model with ARIMA errors, we need to forecast the regression part of the model and the ARIMA part of the model and combine the results.
- Some predictors are known into the future (e.g., time, dummies).
- Separate forecasting models may be needed for other predictors.
- Forecast intervals ignore the uncertainty in forecasting the predictors.

Model daily electricity demand as a function of temperature using quadratic regression with ARMA errors.

```
qplot(elecdaily[,"Temperature"], elecdaily[,"Demand"]) +
    xlab("Temperature") + ylab("Demand")
```

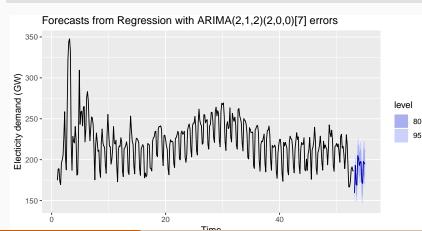






```
##
## Ljung-Box test
##
## data: Residuals from Regression with ARIMA(2,1,2)(2,0,0)[7] errors
## Q* = 28, df = 4, p-value = 1e-05
##
## Model df: 10. Total lags used: 14
```

```
fcast <- forecast(fit,
    xreg = cbind(rep(26,14), rep(26^2,14),
        c(0,1,0,0,1,1,1,1,1,0,0,1,1,1)))
autoplot(fcast) + ylab("Electicity demand (GW)")</pre>
```



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# **Lab Session 19**

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#### **Trend**

#### Linear trend

$$x_t = t$$

- t = 1, 2, ..., T
- Strong assumption that trend will continue.

# **Dummy variables**

If a categorical variable takes only two values (e.g., 'Yes' or 'No'), then an equivalent numerical variable can be constructed taking value 1 if yes and 0 if no. This is called a dummy variable.

	A	В
1	Yes	
2	Yes	
3	No	
4	Yes	
5	No	
6	No	
7	Yes	
8	Yes	
9	No	
10	No	
11	No	
12	No	
13	Yes	
14	No	

# **Dummy variables**

If there are more than two categories, then the variable can be coded using several dummy variables (one fewer than the total number of categories).

		Α	В	С	D	Е
,	1	Monday	1	0	0	0
	2	Tuesday	0	1	0	0
	3	Wednesday	0	0	1	0
	4	Thursday	0	0	0	1
	5	Friday	0	0	0	0
	6	Monday	1	0	0	0
	7	Tuesday	0	1	0	0
	8	Wednesday	0	0	1	0
	9	Thursday	0	0	0	1
	10	Friday	0	0	0	0
	11	Monday	1	0	0	0
	12	Tuesday	0	1	0	0
	13	Wednesday	0	0	1	0
	14	Thursday	0	0	0	1
	15	Friday	0	0	0	0

# Beware of the dummy variable trap!

- Using one dummy for each category gives too many dummy variables!
- The regression will then be singular and inestimable.
- Either omit the constant, or omit the dummy for one category.
- The coefficients of the dummies are relative to the omitted category.

## **Uses of dummy variables**

#### **Seasonal dummies**

- For quarterly data: use 3 dummies
- For monthly data: use 11 dummies
- For daily data: use 6 dummies
- What to do with weekly data?

#### **Uses of dummy variables**

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#### **Outliers**

If there is an outlier, you can use a dummy variable (taking value 1 for that observation and 0 elsewhere) to remove its effect.

## **Uses of dummy variables**

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#### **Outliers**

If there is an outlier, you can use a dummy variable (taking value 1 for that observation and 0 elsewhere) to remove its effect.

#### **Public holidays**

■ For daily data: if it is a public holiday, dummy=1, otherwise dummy=0.

#### **Fourier series**

Periodic seasonality can be handled using pairs of Fourier terms:

$$s_k(t) = \sin\left(\frac{2\pi kt}{m}\right) \qquad c_k(t) = \cos\left(\frac{2\pi kt}{m}\right)$$
$$y_t = a + bt + \sum_{k=1}^{K} \left[\alpha_k s_k(t) + \beta_k c_k(t)\right] + \varepsilon_t$$

- Every periodic function can be approximated by sums of sin and cos terms for large enough K.
- Choose K by minimizing AICc.
- Called "harmonic regression"
- fourier() function generates these.

#### **Intervention variables**

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Equivalent to a dummy variable for handling an outlier.

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Variable takes value 0 before the intervention and 1 afterwards.

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Equivalent to a dummy variable for handling an outlier.

#### Steps

Variable takes value 0 before the intervention and 1 afterwards.

# Change of slope

■ Variables take values 0 before the intervention and values  $\{1, 2, 3, \dots\}$  afterwards.

## **Holidays**

#### For monthly data

- Christmas: always in December so part of monthly seasonal effect
- Easter: use a dummy variable  $v_t = 1$  if any part of Easter is in that month,  $v_t = 0$  otherwise.
- Ramadan and Chinese new year similar.

# **Trading days**

With monthly data, if the observations vary depending on how many different types of days in the month, then trading day predictors can be useful.

```
z<sub>1</sub> = # Mondays in month;
z<sub>2</sub> = # Tuesdays in month;
:
z<sub>7</sub> = # Sundays in month.
```

## **Distributed lags**

Lagged values of a predictor.

Example: x is advertising which has a delayed effect

```
    x<sub>1</sub> = advertising for previous month;
    x<sub>2</sub> = advertising for two months previously;
    :
    x<sub>m</sub> = advertising for m months previously.
```

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## Dynamic harmonic regression

#### **Combine Fourier terms with ARIMA errors**

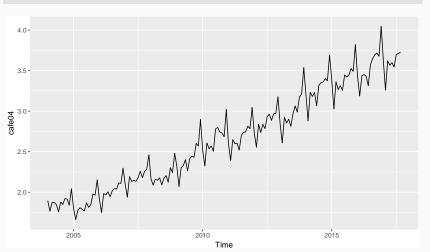
#### **Advantages**

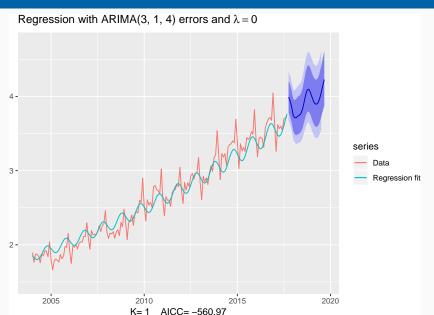
- it allows any length seasonality;
- for data with more than one seasonal period, you can include Fourier terms of different frequencies;
- the seasonal pattern is smooth for small values of K (but more wiggly seasonality can be handled by increasing K);
- the short-term dynamics are easily handled with a simple ARMA error.

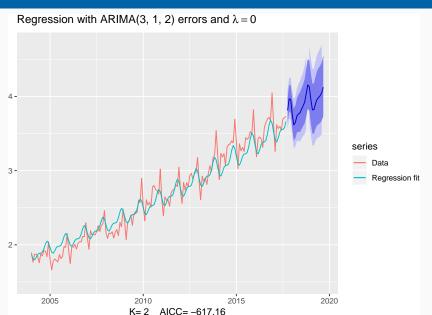
#### **Disadvantages**

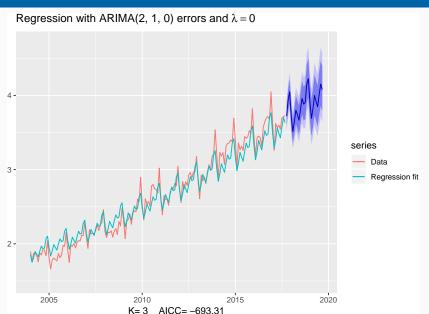
seasonality is assumed to be fixed

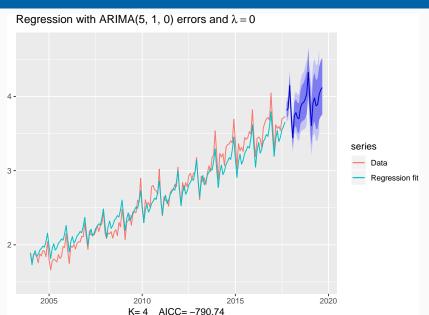
```
cafe04 <- window(auscafe, start=2004)
autoplot(cafe04)</pre>
```

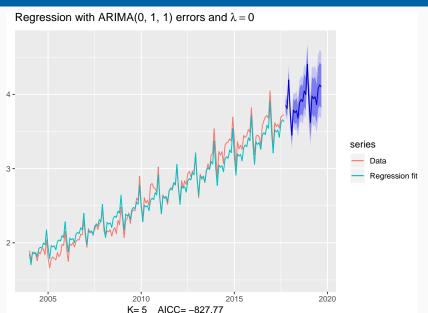


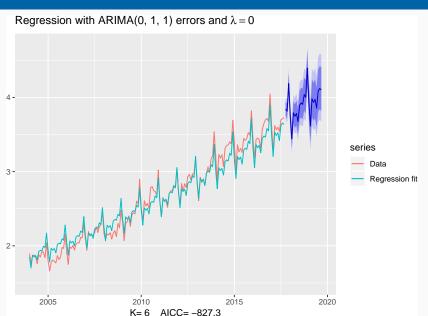


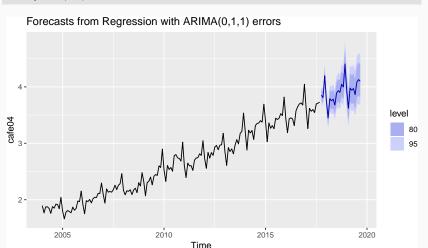






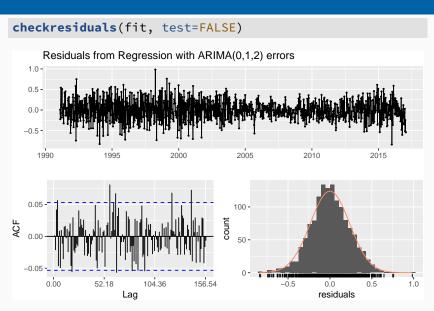






harmonics <- fourier(gasoline, K = 13)

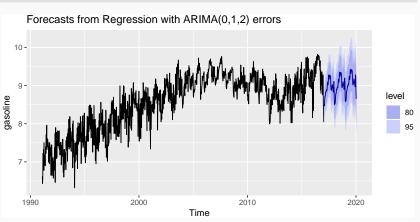
```
(fit <- auto.arima(gasoline, xreg = harmonics, seasonal = FALSE))
## Series: gasoline
  Regression with ARIMA(0,1,2) errors
##
  Coefficients:
##
          ma1
                ma2
                     drift S1-52 C1-52
                                         S2-52
##
      -0.961
              0.094 0.001 0.031 -0.255 -0.052
## s.e. 0.027
              0.029 0.001 0.012 0.012 0.009
##
       C2-52 S3-52 C3-52 S4-52 C4-52 S5-52
##
     -0.017
              0.024 -0.099 0.032 -0.026 -0.001
## s.e. 0.009 0.008 0.008 0.008 0.008 0.008
      C5-52 S6-52 C6-52 S7-52 C7-52 S8-52
##
##
      -0.047
              0.058 -0.032 0.028
                                 0.037
                                        0.024
## s.e. 0.008
              0.008 0.008 0.008
                                  0.008 0.008
##
      C8-52 S9-52 C9-52 S10-52
                                 C10-52 S11-52
     0.014 -0.017 0.012 -0.024 0.023 0.000
##
## s.e. 0.008
              0.008
                     0.008 0.008 0.008 0.008
##
       C11-52 S12-52 C12-52 S13-52 C13-52
    -0.019 -0.029 -0.018 0.001 -0.018
##
## s.e. 0.008 0.008 0.008 0.008 0.008
##
  sigma^2 estimated as 0.056: log likelihood=43.66
## AIC=-27.33
            AICc=-25.92
                          BIC=129
```



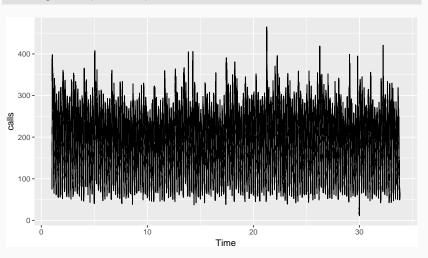
#### checkresiduals(fit, plot=FALSE)

```
##
## Ljung-Box test
##
## data: Residuals from Regression with ARIMA(0,1,2) errors
## Q* = 130, df = 75, p-value = 6e-05
##
## Model df: 29. Total lags used: 104.357142857143
```

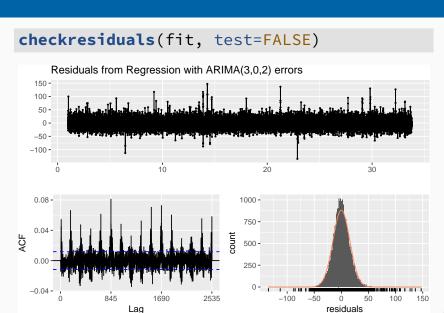
```
newharmonics <- fourier(gasoline, K = 13, h = 156)
fc <- forecast(fit, xreg = newharmonics)
autoplot(fc)</pre>
```



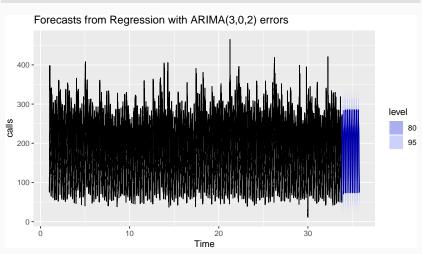
#### autoplot(calls)



```
xreg <- fourier(calls, K = c(10,0))
(fit <- auto.arima(calls, xreg=xreg, seasonal=FALSE, stationary=TRUE))</pre>
## Series: calls
## Regression with ARIMA(3,0,2) errors
##
## Coefficients:
##
         ar1
               ar2
                      ar3
                             ma1
                                    ma2 intercept
     0.841 0.192 -0.044 -0.590 -0.189 192.070
##
## s.e. 0.169 0.178 0.013 0.169 0.137
                                            1.764
##
   S1-169 C1-169 S2-169 C2-169 S3-169
##
   55.245 -79.087 13.674 -32.375 -13.693
## s.e. 0.701 0.701 0.379 0.379 0.273
    C3-169 S4-169 C4-169 S5-169 C5-169 S6-169
##
##
      -9.327 -9.532 -2.797 -2.239 2.893 0.173
## s.e. 0.273 0.223 0.196 0.196 0.179
##
     C6-169 S7-169 C7-169 S8-169 C8-169 S9-169
      3.305 0.855 0.294 0.857 -1.391 -0.986
##
## s.e. 0.179 0.168 0.168
                             0.160 0.160
                                           0.155
##
       C9-169 S10-169 C10-169
   -0.345 -1.196 0.801
##
## s.e. 0.155 0.150 0.150
##
  sigma^2 estimated as 243: log likelihood=-115412
## ATC=230877
            AICc=230877
                          BIC=231099
```



```
fc <- forecast(fit, xreg = fourier(calls, c(10,0), 1690))
autoplot(fc)</pre>
```



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# **Lab Session 20**

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# Sometimes a change in $x_t$ does not affect $y_t$ instantaneously

- $y_t$  = sales,  $x_t$  = advertising.
- $y_t$  = stream flow,  $x_t$  = rainfall.
- $y_t$  = size of herd,  $x_t$  = breeding stock.

# Sometimes a change in $x_t$ does not affect $y_t$ instantaneously

- $y_t$  = sales,  $x_t$  = advertising.
- $y_t$  = stream flow,  $x_t$  = rainfall.
- $y_t$  = size of herd,  $x_t$  = breeding stock.
- These are dynamic systems with input  $(x_t)$  and output  $(y_t)$ .
- $\blacksquare$   $x_t$  is often a leading indicator.
- There can be multiple predictors.

The model include present and past values of predictor:  $x_t, x_{t-1}, x_{t-2}, \ldots$ 

$$y_t = a + \nu_0 x_t + \nu_1 x_{t-1} + \dots + \nu_k x_{t-k} + \eta_t$$

where  $\eta_t$  is an ARIMA process.

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$$y_t = a + \nu_0 x_t + \nu_1 x_{t-1} + \cdots + \nu_k x_{t-k} + \eta_t$$

where  $\eta_t$  is an ARIMA process.

#### Rewrite model as

$$y_t = a + (\nu_0 + \nu_1 B + \nu_2 B^2 + \dots + \nu_k B^k) x_t + \eta_t$$
  
=  $a + \nu(B) x_t + \eta_t$ .

The model include present and past values of predictor:  $x_t, x_{t-1}, x_{t-2}, \ldots$ 

$$y_t = a + \nu_0 x_t + \nu_1 x_{t-1} + \dots + \nu_k x_{t-k} + \eta_t$$

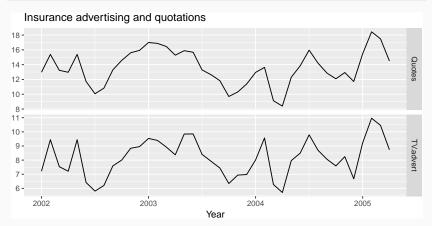
where  $\eta_t$  is an ARIMA process.

#### Rewrite model as

$$y_t = a + (\nu_0 + \nu_1 B + \nu_2 B^2 + \dots + \nu_k B^k) x_t + \eta_t$$
  
=  $a + \nu(B) x_t + \eta_t$ .

- $\nu$ (B) is called a *transfer function* since it describes how change in  $x_t$  is transferred to  $y_t$ .
- x can influence y, but y is not allowed to influence x.

```
autoplot(insurance, facets=TRUE) +
   xlab("Year") + ylab("") +
   ggtitle("Insurance advertising and quotations")
```



```
Advert <- cbind(
    AdLag0 = insurance[,"TV.advert"],
    AdLag1 = lag(insurance[,"TV.advert"],-1),
    AdLag2 = lag(insurance[,"TV.advert"],-2),
    AdLag3 = lag(insurance[,"TV.advert"],-3)) %>%
  head(NROW(insurance))
# Restrict data so models use same fitting period
fit1 <- auto.arima(insurance[4:40,1], xreg=Advert[4:40,1],</pre>
  stationary=TRUE)
fit2 <- auto.arima(insurance[4:40,1], xreg=Advert[4:40,1:2],</pre>
  stationary=TRUE)
fit3 <- auto.arima(insurance[4:40,1], xreg=Advert[4:40,1:3],</pre>
  stationary=TRUE)
fit4 <- auto.arima(insurance[4:40,1], xreg=Advert[4:40,1:4],</pre>
  stationary=TRUE)
c(fit1$aicc, fit2$aicc, fit3$aicc, fit4$aicc)
```

```
(fit <- auto.arima(insurance[,1], xreg=Advert[,1:2],
    stationary=TRUE))</pre>
```

```
## Series: insurance[, 1]
## Regression with ARIMA(3,0,0) errors
##
## Coefficients:
##
         ar1
                 ar2 ar3 intercept AdLag0
                                             AdLag1
   1.412 -0.932 0.359
                               2.039 1.256
                                              0.162
##
## s.e. 0.170 0.255 0.159
                               0.993 0.067 0.059
##
## sigma^2 estimated as 0.217: log likelihood=-23.89
## AIC=61.78 AICc=65.28 BIC=73.6
```

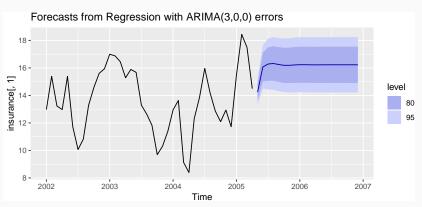
```
(fit <- auto.arima(insurance[,1], xreg=Advert[,1:2],
    stationary=TRUE))</pre>
```

```
## Series: insurance[, 1]
## Regression with ARIMA(3,0,0) errors
##
## Coefficients:
## ar1 ar2 ar3 intercept AdLag0 AdLag1
## 1.412 -0.932 0.359 2.039 1.256 0.162
## s.e. 0.170 0.255 0.159 0.993 0.067 0.059
##
## sigma^2 estimated as 0.217: log likelihood=-23.89
## AIC=61.78 AICc=65.28 BIC=73.6
```

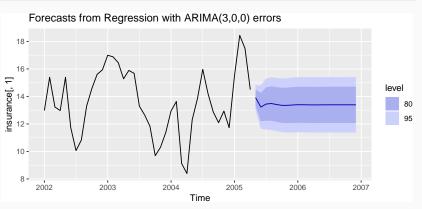
$$y_t = 2.04 + 1.26x_t + 0.16x_{t-1} + \eta_t,$$
  

$$\eta_t = 1.41\eta_{t-1} - 0.93\eta_{t-2} + 0.36\eta_{t-3} + \varepsilon_t,$$

```
fc <- forecast(fit, h=20,
    xreg=cbind(c(Advert[40,1],rep(10,19)), rep(10,20)))
autoplot(fc)</pre>
```



```
fc <- forecast(fit, h=20,
    xreg=cbind(c(Advert[40,1],rep(8,19)), rep(8,20)))
autoplot(fc)</pre>
```



```
fc <- forecast(fit, h=20,
    xreg=cbind(c(Advert[40,1],rep(6,19)), rep(6,20)))
autoplot(fc)</pre>
```

