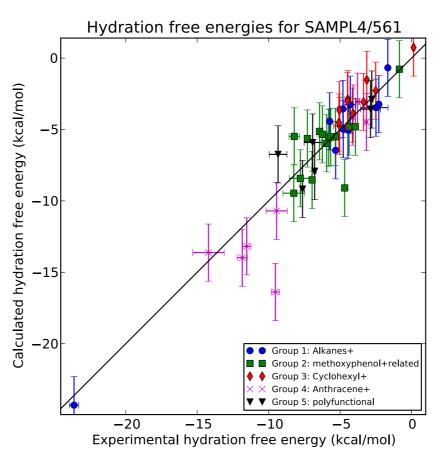
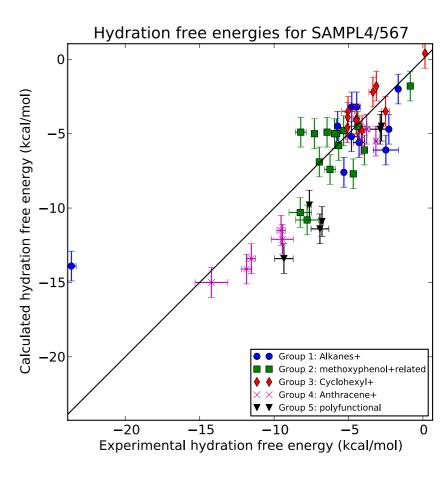
Error analysis may not seem that important for drug discovery -- but poor understanding of error is a huge reason modeling doesn't see more use in discovery settings

### A method's usefulness is partly determined by how well it can estimate error





- Average error
- RMS error
- Average unsigned error
- Kendall tau
- Pearson R

$$-0.12 + / -0.24$$

$$1.13 + / - 0.17$$

$$0.73 + / -0.05$$

$$0.93 + / - 0.04$$

$$-0.55 + / -0.33$$

$$2.38 + / - 0.40$$

$$1.78 + / -0.23$$

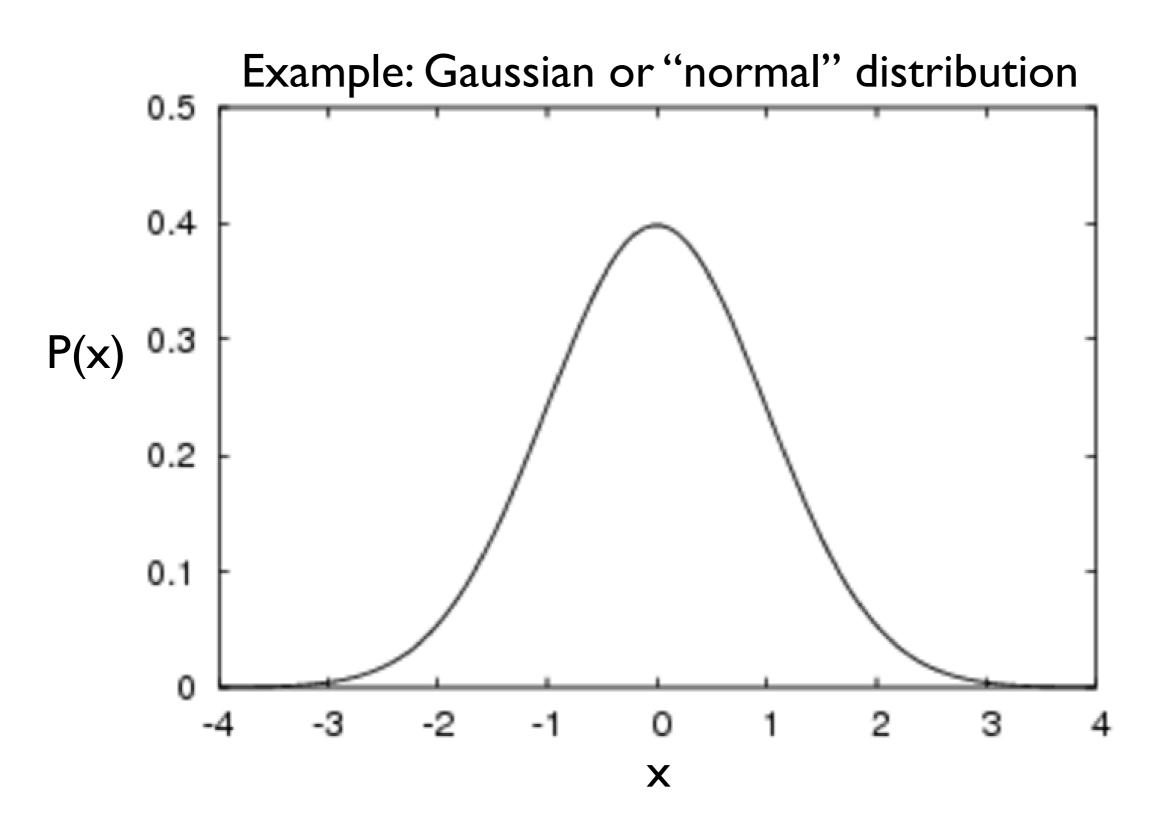
$$0.58 + / - 0.08$$

$$0.81 + / -0.05$$

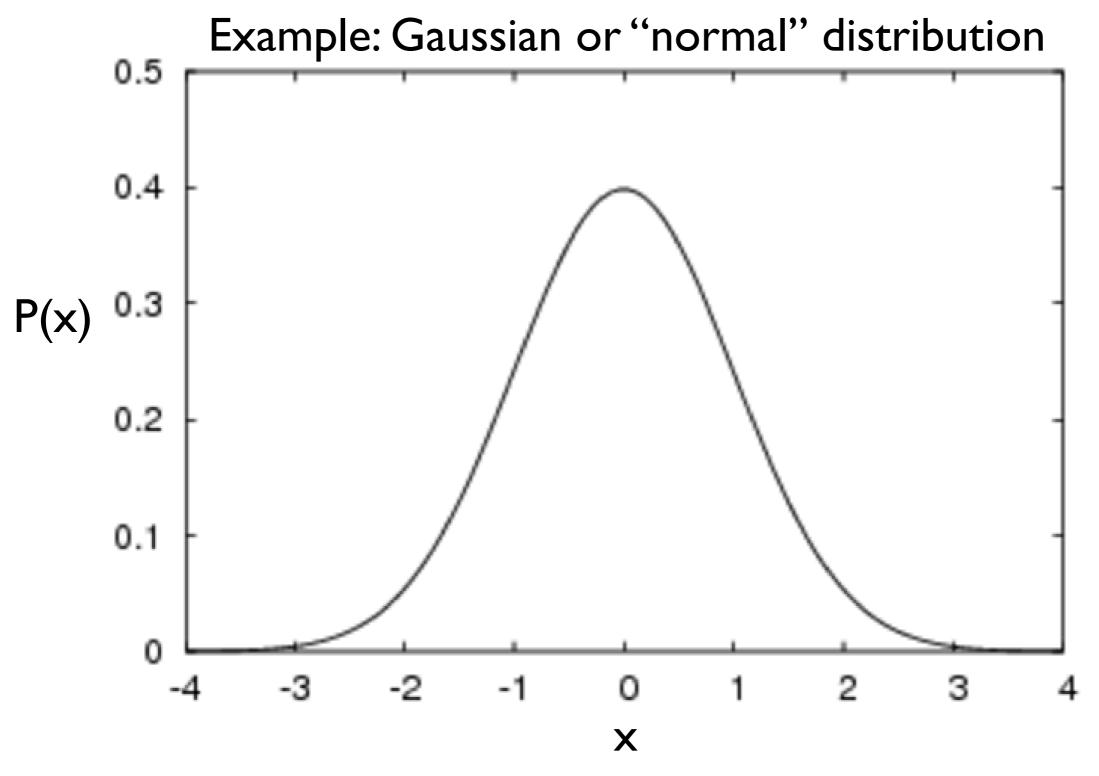
#### A couple of key references to save for your records

- Simulation-oriented: "Best practices for quantification of uncertainty and sampling quality in molecular simulations": <a href="https://github.com/dmzuckerman/Sampling-Uncertainty">https://github.com/dmzuckerman/Sampling-Uncertainty</a>
- General/screening oriented: "Confidence limits, error bars and method comparison in molecular modeling by A. Nicholls. Part 1 (<a href="http://dx.doi.org//10.1007/s10822-014-9753-z">http://dx.doi.org//10.1007/s10822-016-9904-5</a>).

## We also have to deal with distributions of continuous variables

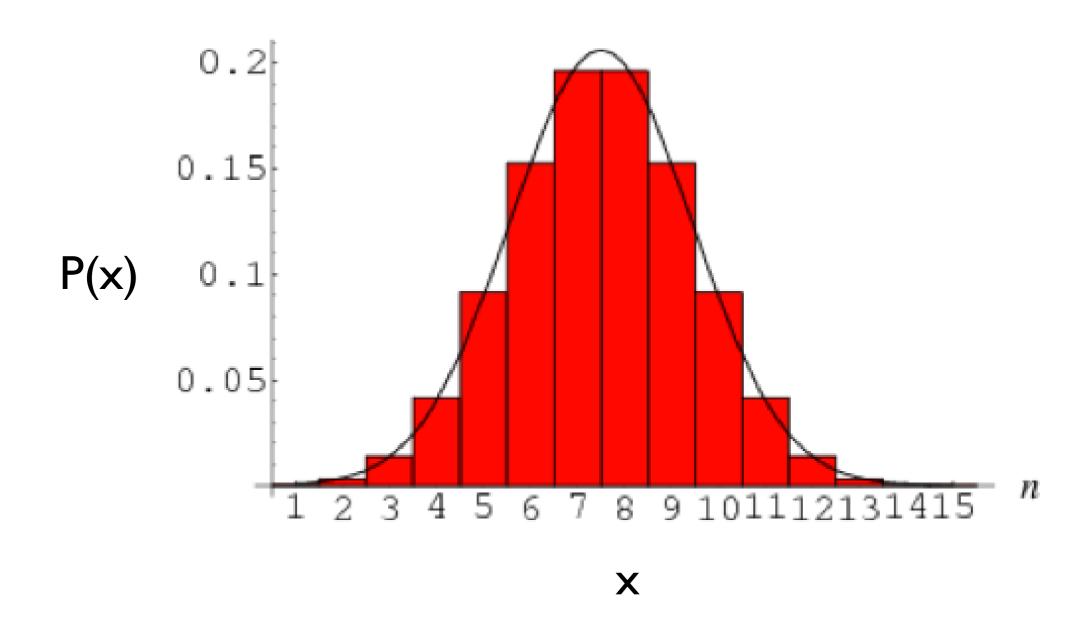


## We also have to deal with distributions of continuous variables

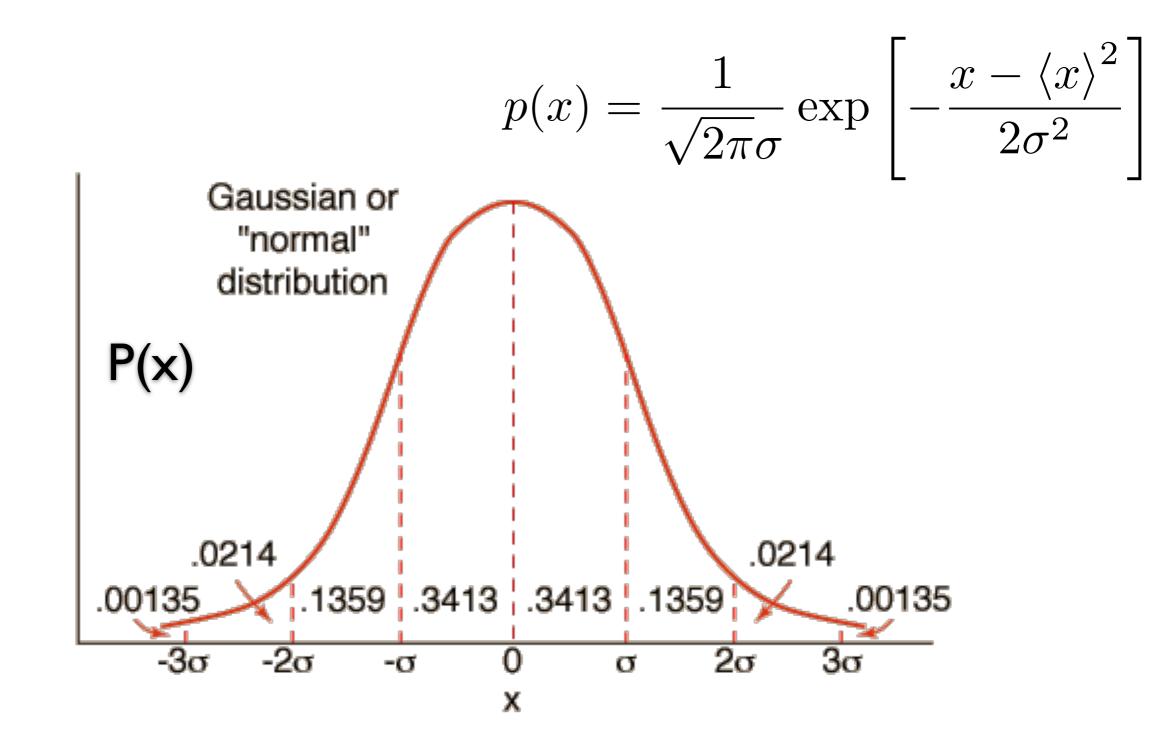


Probability density function

#### One can also "discretize" continuous distributions



#### Distributions have a characteristic width, o



# Averages are computed by integrating over the probability density function

$$\langle f(x) \rangle = \langle f \rangle = \int dx f(x) p(x)$$

Example, average x. For uniform distribution from 0 to 1 -that is, p(x) = 1:

$$\langle x \rangle = \int dx \cdot x = 1/2$$

For distribution p(x) = 2x on range 0 to 1:

$$\langle x \rangle = \int dx \cdot x \cdot 2x = 2/3$$

# Many experiments amount to averaging over probability distributions

- For example we make many observations
- Or we effectively average over many observations:
  - Observe the same property averaged over many molecules (~Avogadro's number)
  - Average over "long" times (seconds or more)

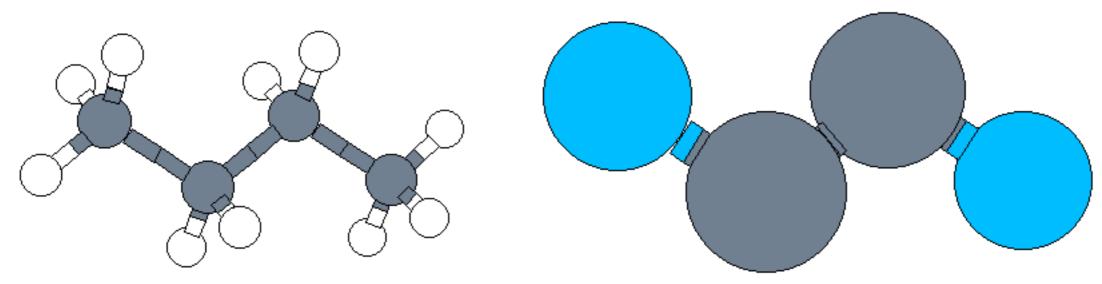
# Often, we perform repeated observations of some property to estimate its average

$$\langle f \rangle \doteq \frac{1}{N} \sum_{i=1}^{N} f(x_i)$$

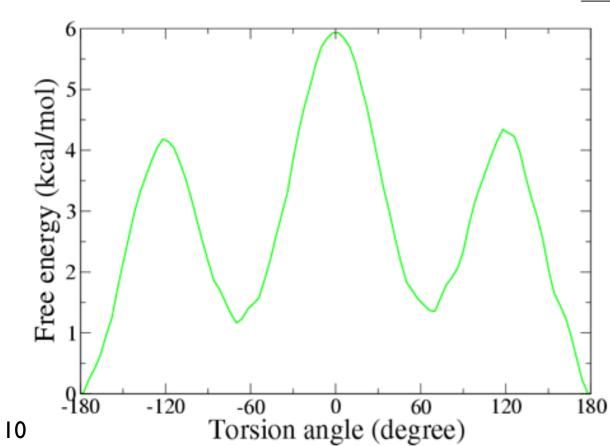
Many observations i=1 to N of some property f of our observed values  $x_i$ 

The dot indicates we approximate the true average -- only equal as  $N \rightarrow \infty$ 

### Example: Butane

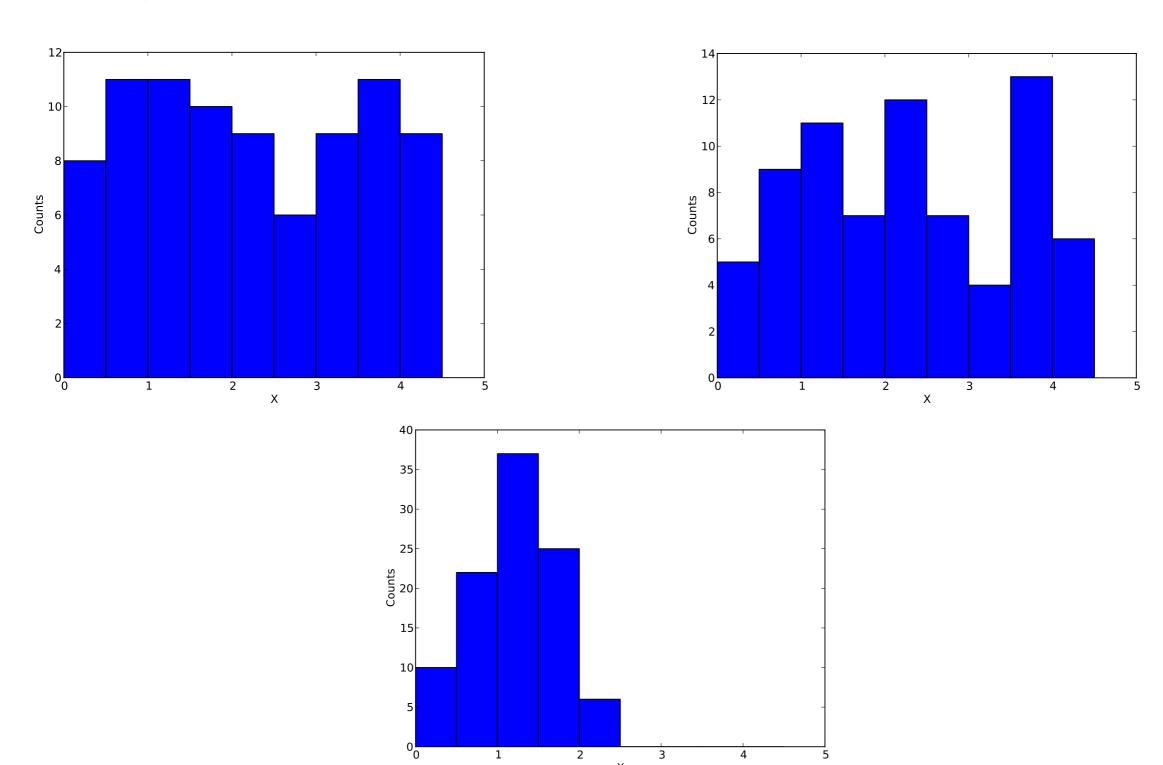


http://www.theo.chemie.tu-darmstadt.de/group/services/yaspdoc/ kurs/butan/index.html



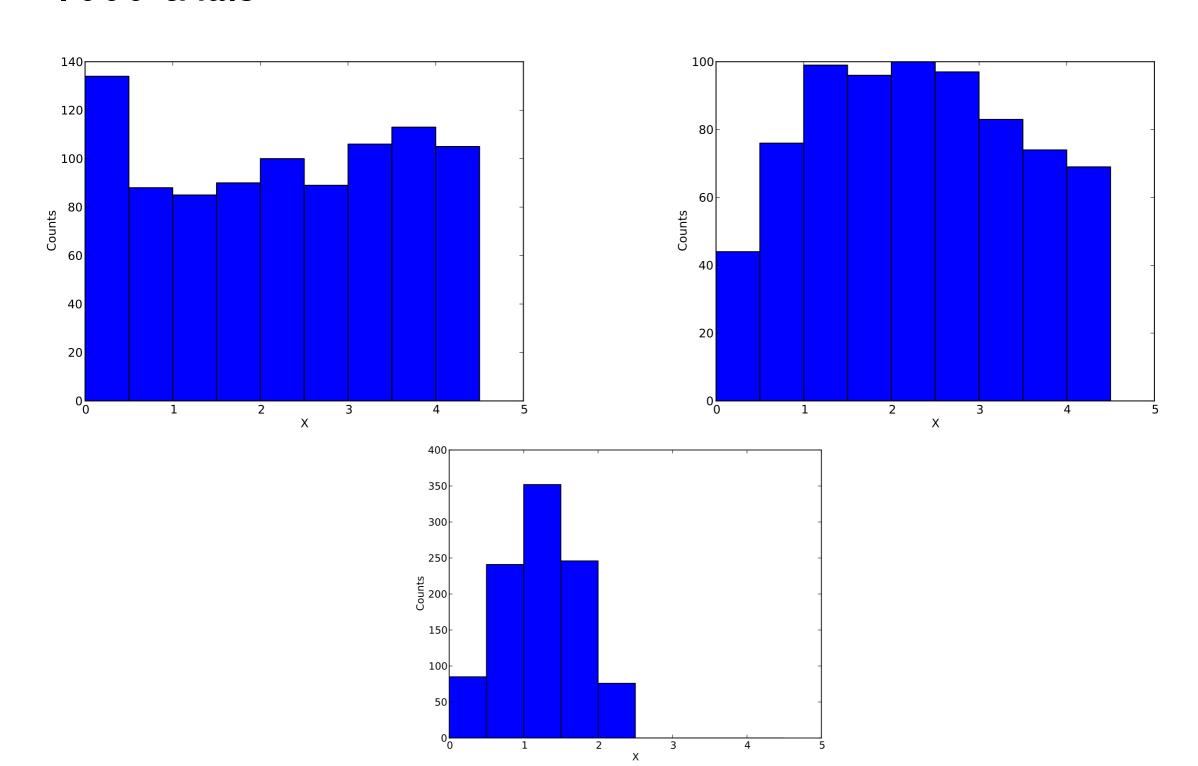
# Observations are taken from distributions; distinguishing distributions is nontrivial

#### 100 trials

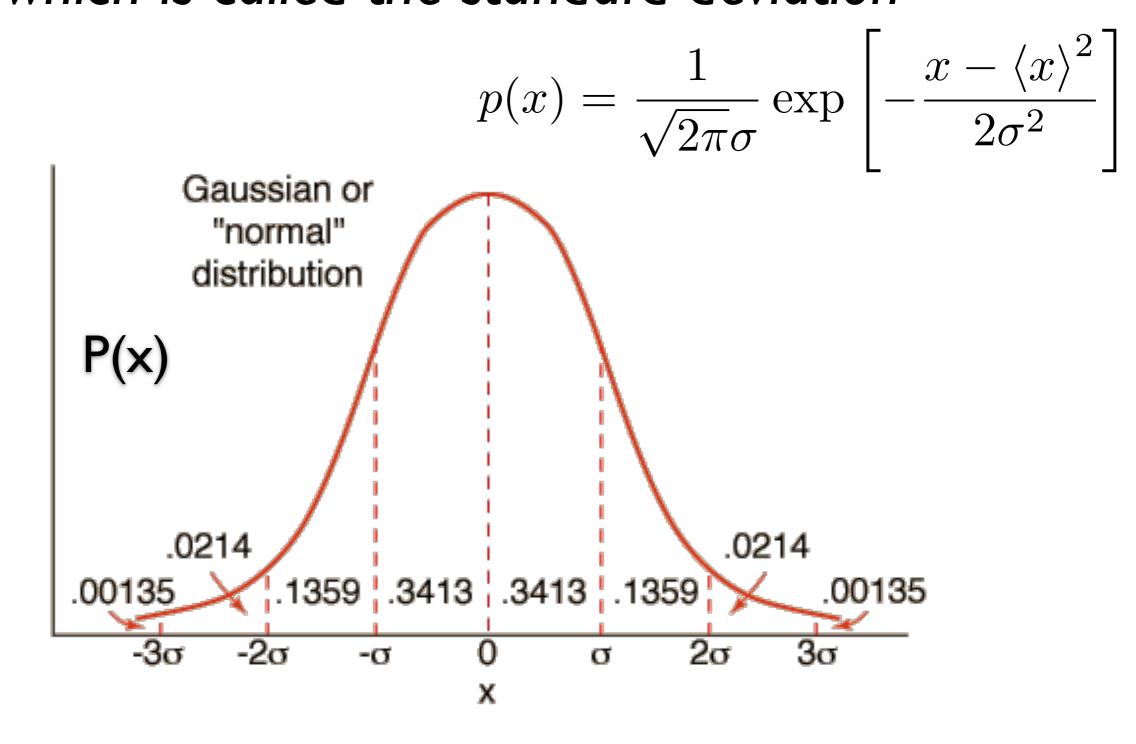


# Observations are taken from distributions; distinguishing distributions is nontrivial

#### 1000 trials



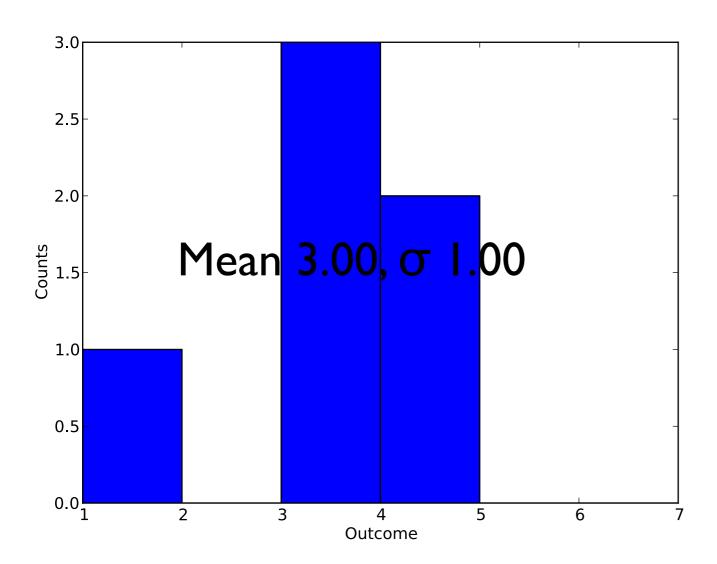
### Distributions have a characteristic width, $\sigma$ which is called the standard deviation

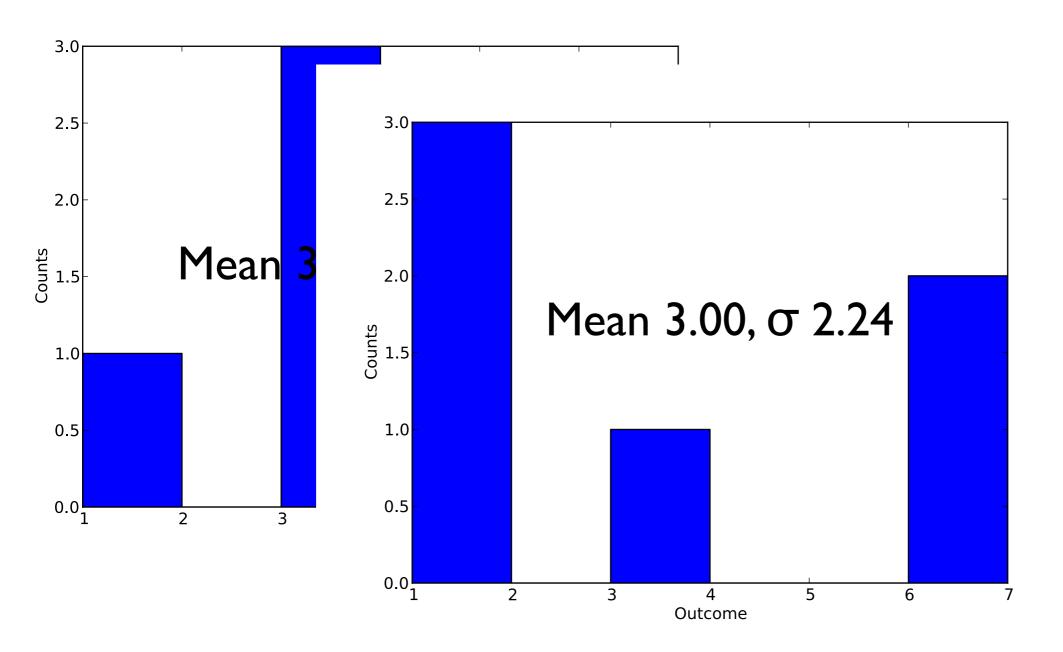


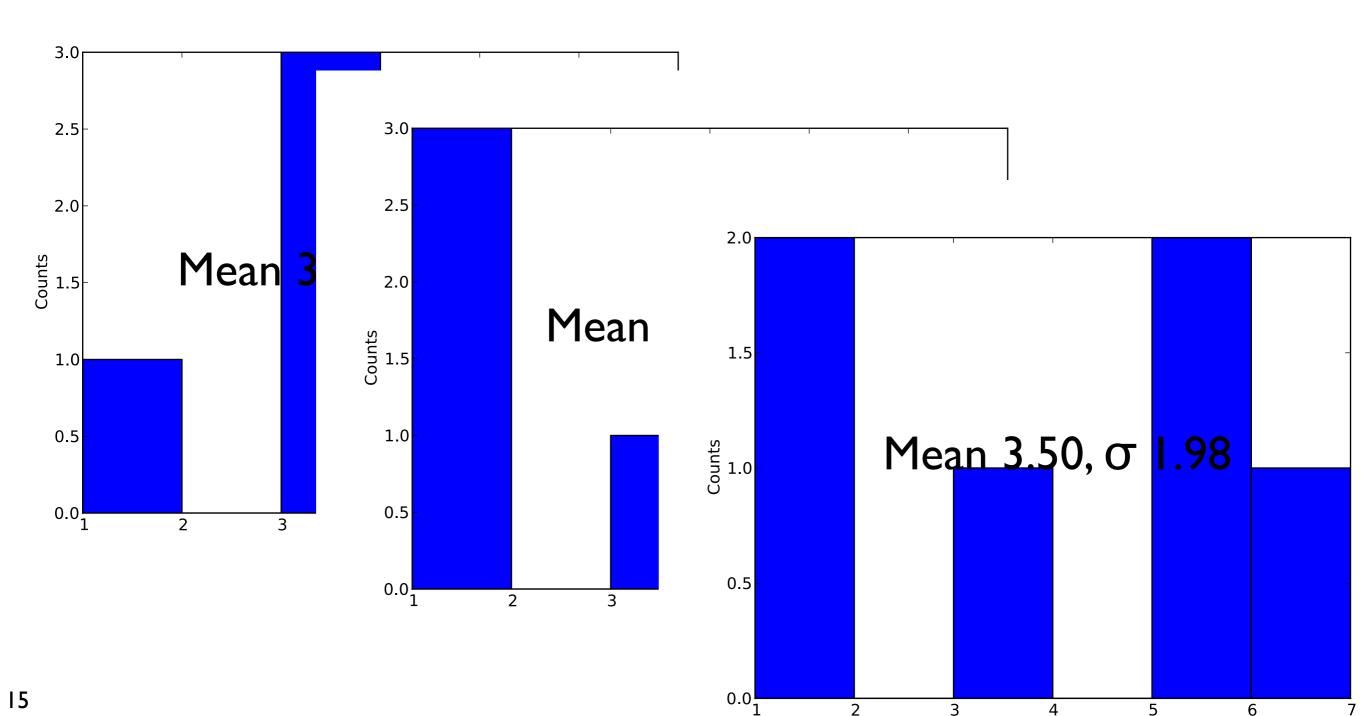
# The standard deviation is computed the same way regardless of the shape of the distribution

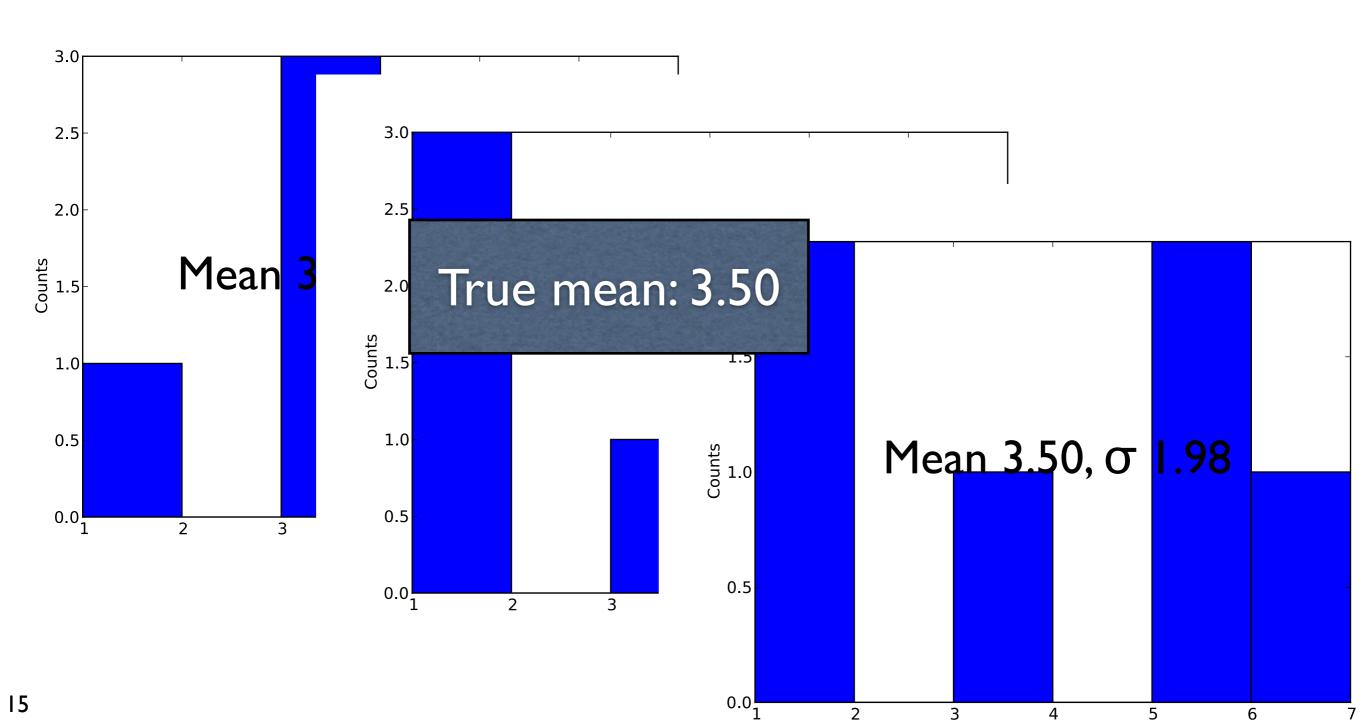
$$\sigma = \sqrt{\sum_i (x_i - \bar{x})^2}$$
 with measurements  $x_i$ 

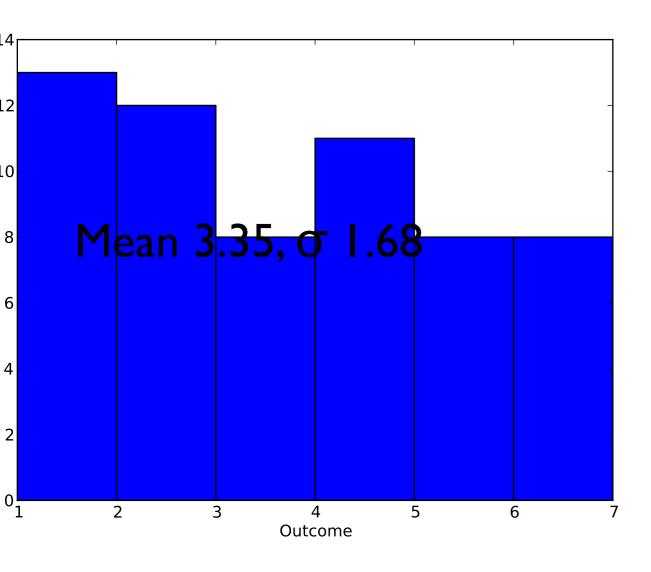
and the mean x is: 
$$\bar{x} = \frac{1}{N} \sum_{i=1}^{n} x_i$$

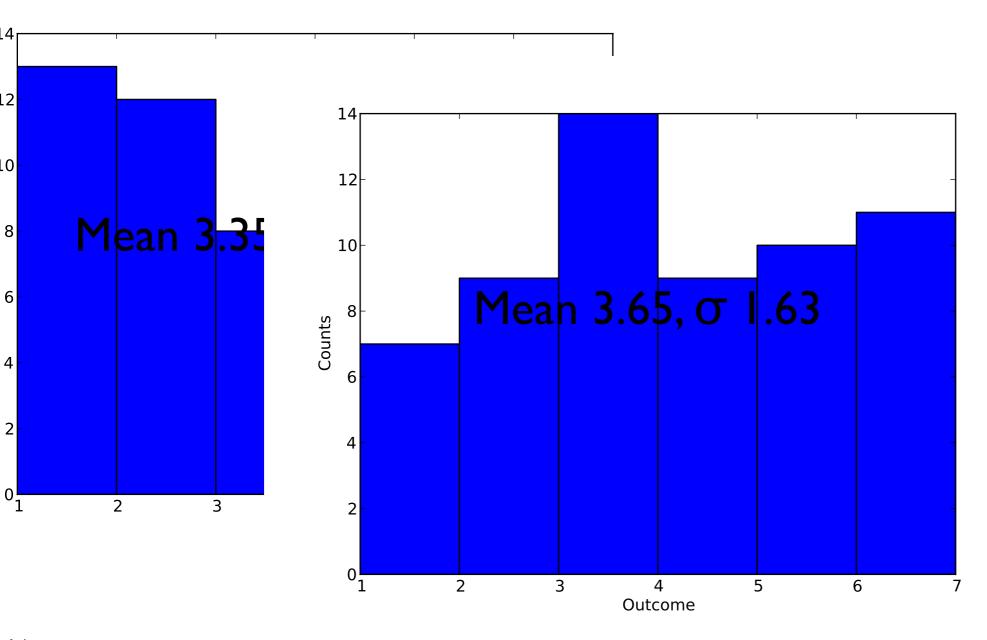


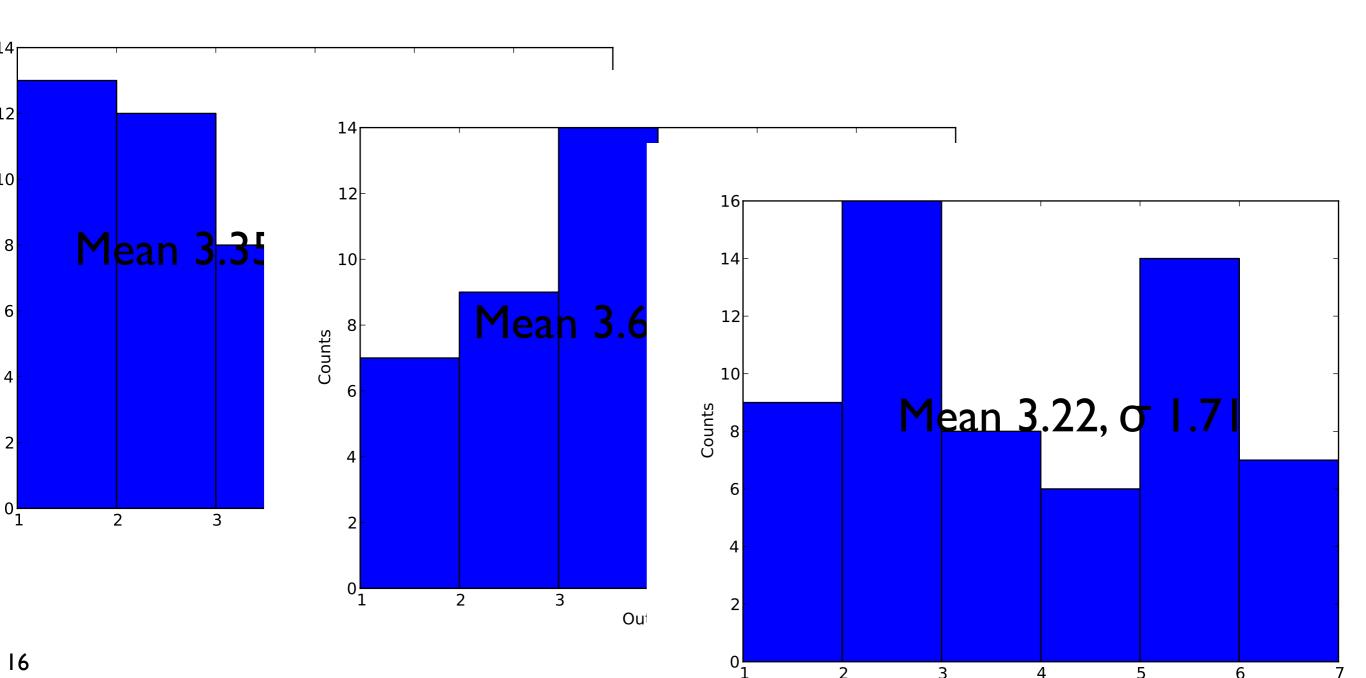


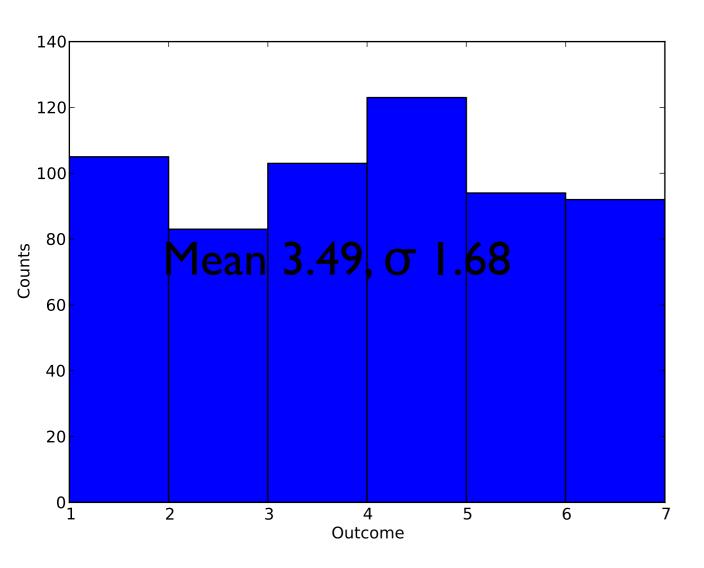


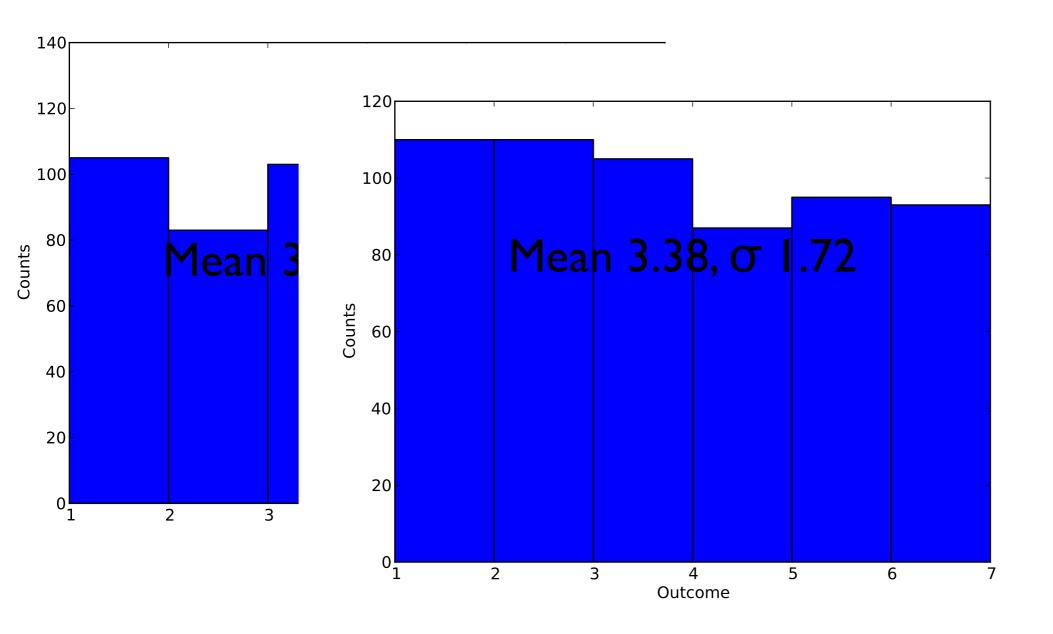


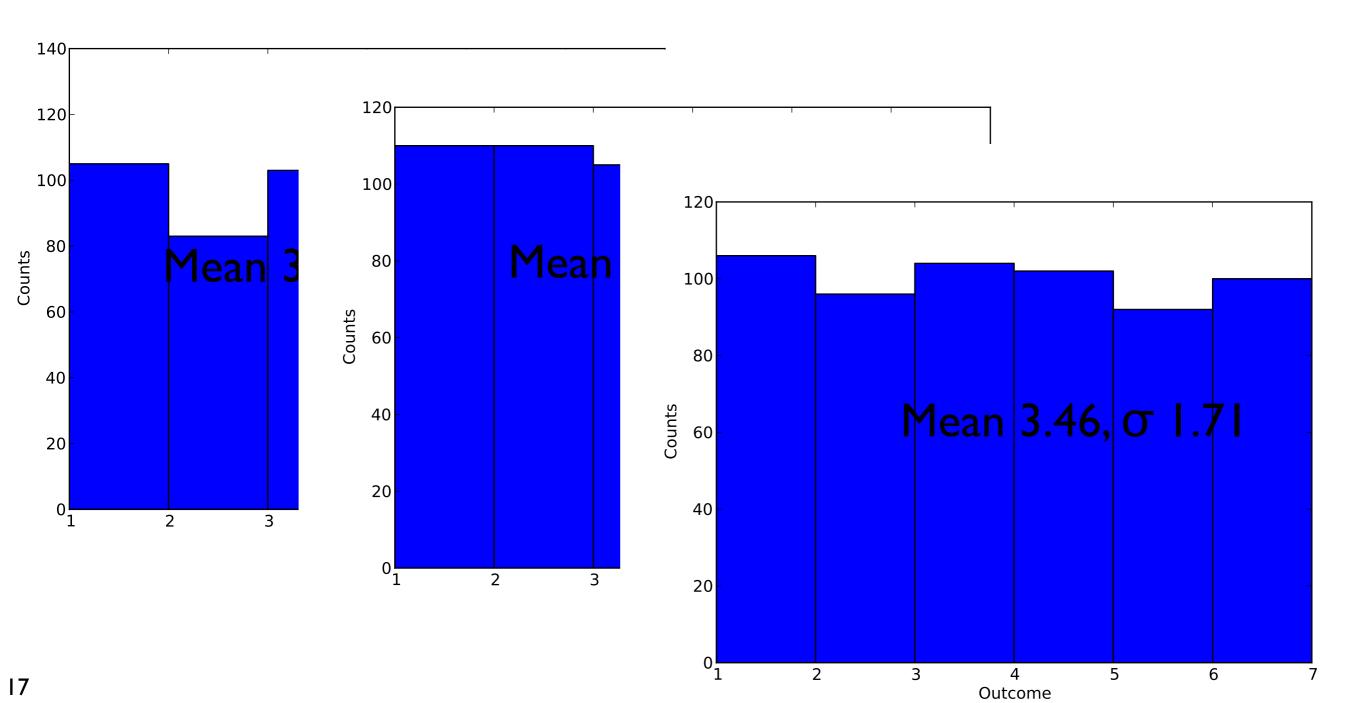












# Many problems involve multiplying multiple probability distributions

Simple example: Flipping a coin twice

```
P(two heads) = P(heads) * P(heads) = I/4
P(two tails) = P(tails) * P(tails) = I/4
P(heads then tails) = P(heads) * P(tails) = I/4
P(tails then heads) = P(tails) * P(heads) = I/4
```

#### But:

P(one tails, one heads) = P(tails) \* P(heads) + P(heads)\*P(tails)

More ways -- higher entropy

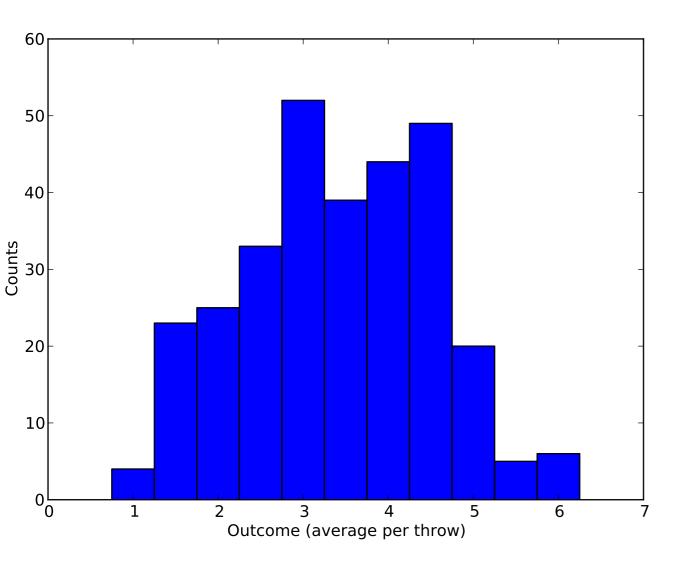
# Many problems involve multiplying multiple probability distributions

#### Rolling two dice:

Outcome	Probability	Ways	Outcome	Probability	Ways
			7	6/36	1+6, 2+5, 3+4
2	1/36	1+1	8	5/36	2+6, 3+5, 4+4
3	2/36	I+2	9	4/36	3+6, 4+5
4	3/36	I+3, 2+2	10	3/36	4+6, 5+5
5	4/36	I+4, 2+3		2/36	5+6
6	5/36	I+5, 2+4, 3+3	12	1/36	6+6

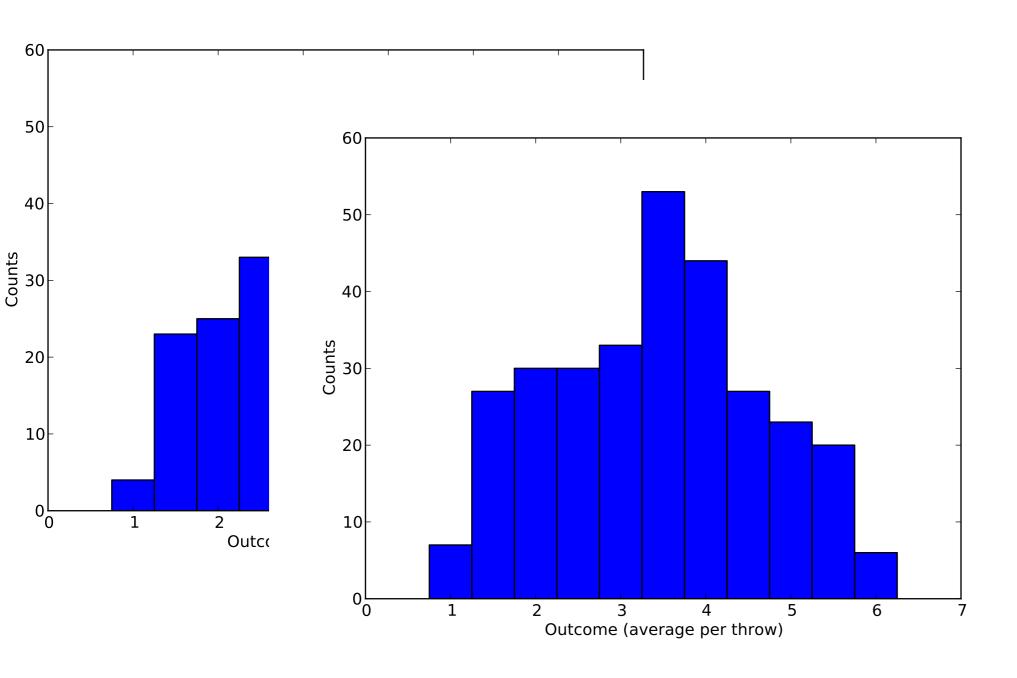
### Let's test this numerically -- two dice

#### 300 throws of two dice



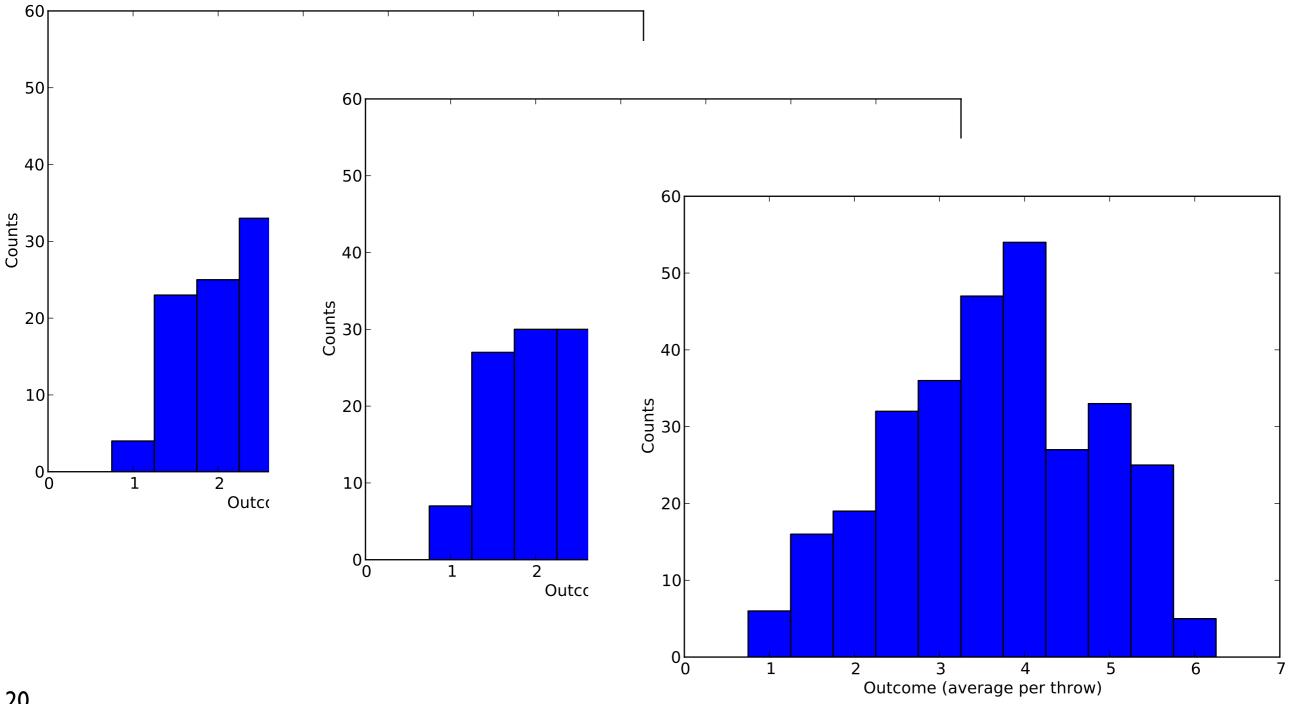
### Let's test this numerically -- two dice

#### 300 throws of two dice



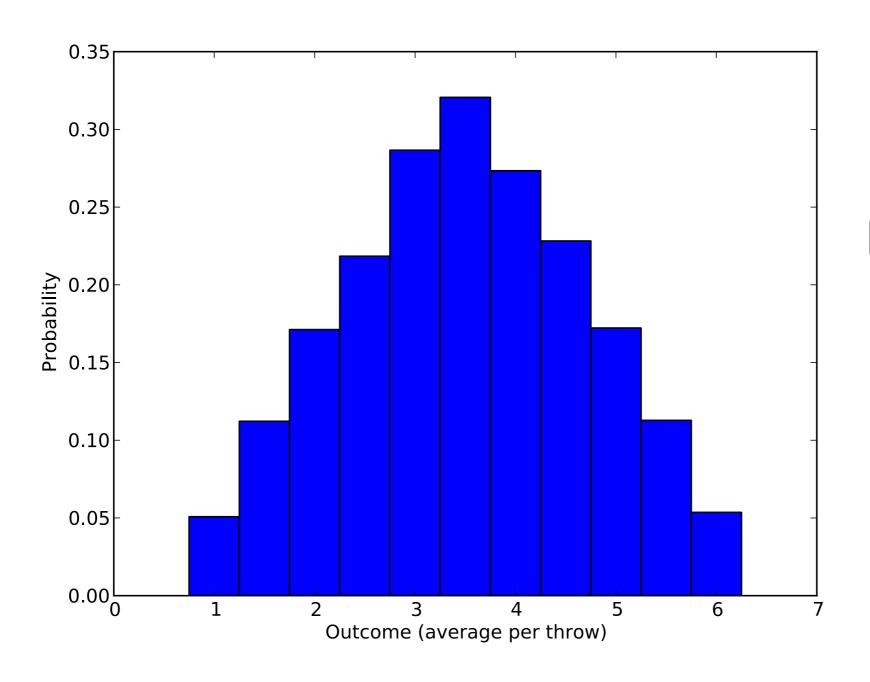
### Let's test this numerically -- two dice

#### 300 throws of two dice



# It's hard to verify our predicted probabilities without doing more trials

#### 10000 throws of two dice



Mean 3.5,  $\sigma$  1.2

### Take-aways:

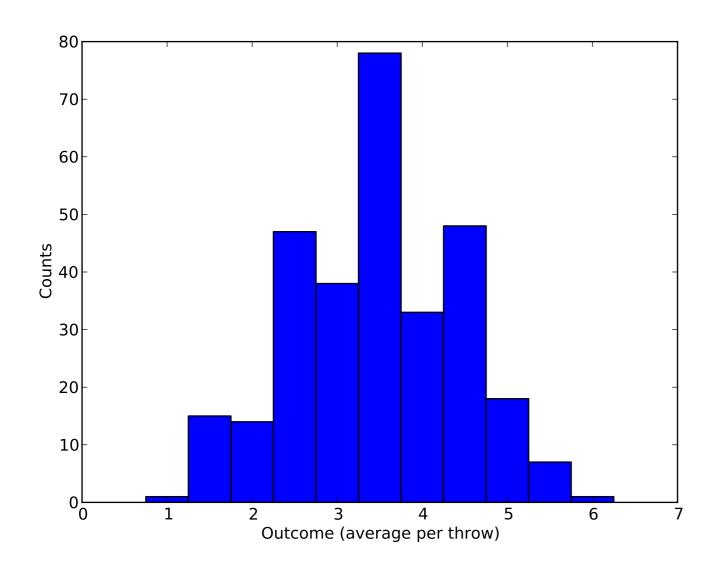
Combining uniform distributions can give nonuniform things

To really get the correct average takes a lot of trials

What about more dice throws?

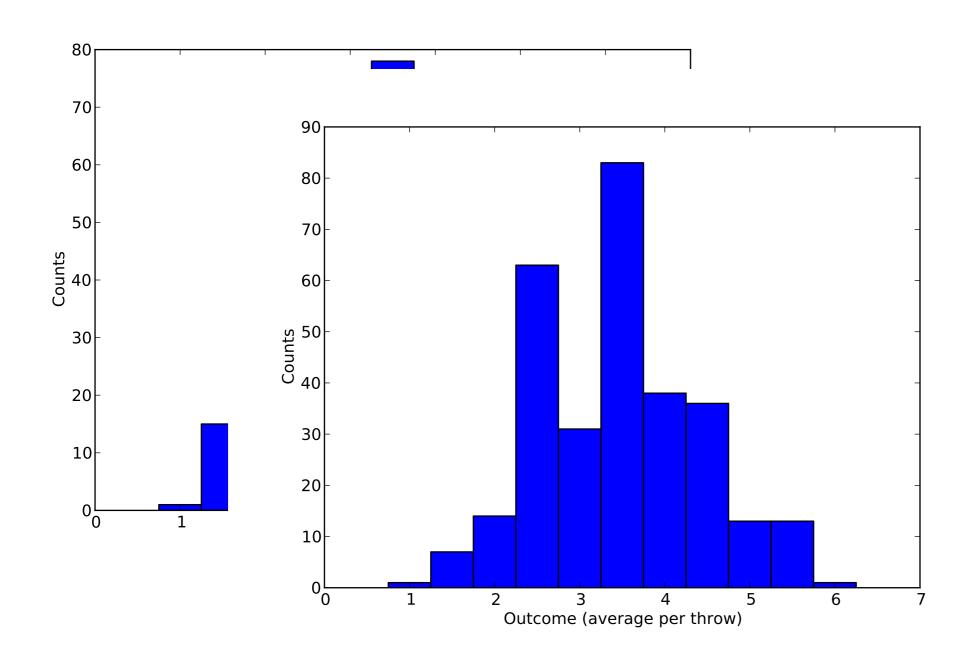
### Three dice -- more sharply peaked

#### 300 throws of three dice



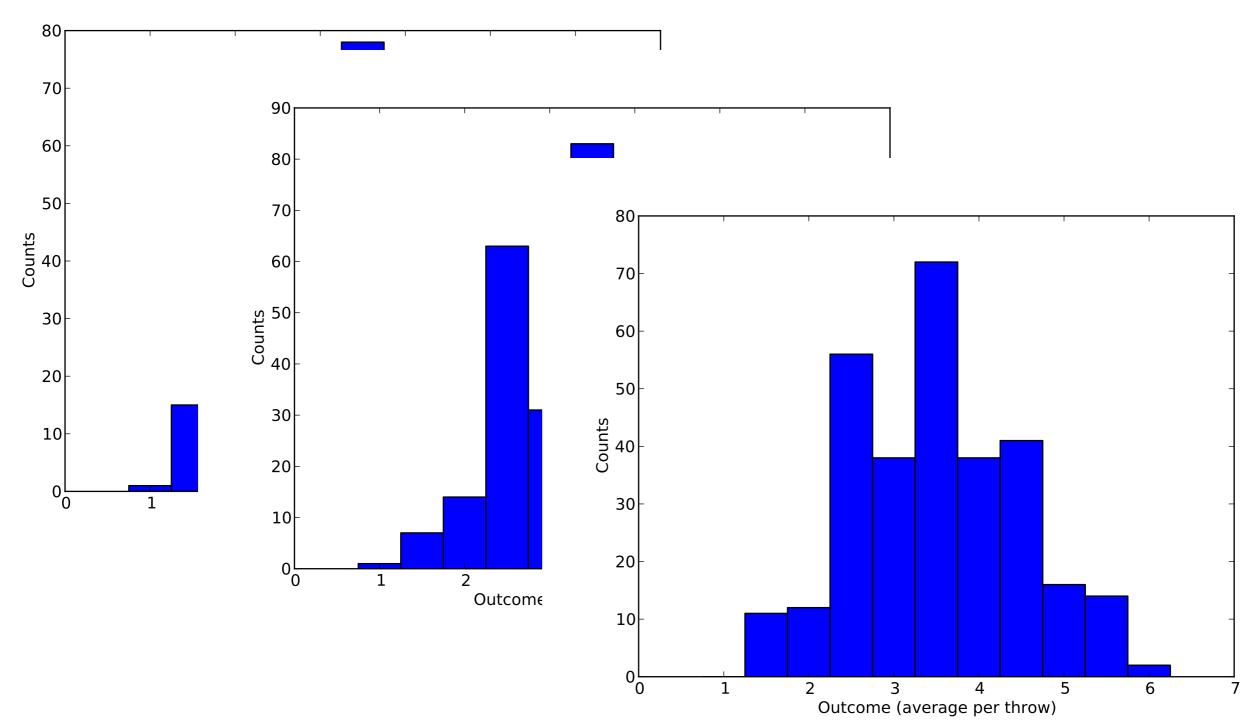
### Three dice -- more sharply peaked

#### 300 throws of three dice



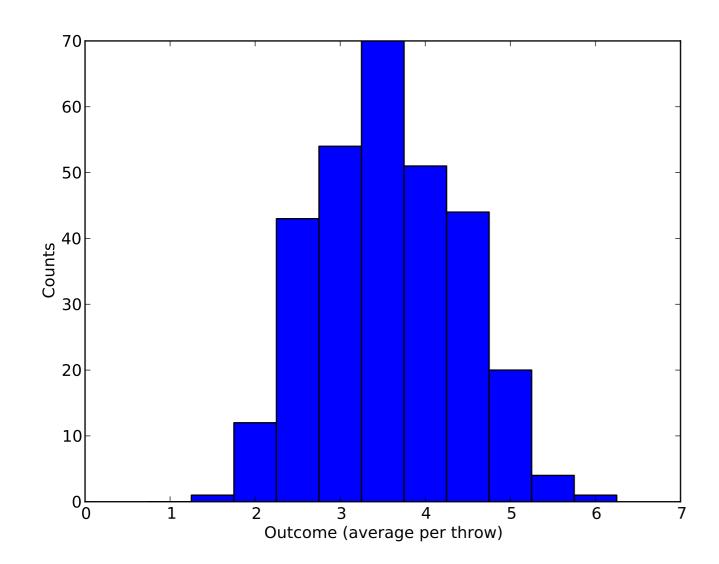
### Three dice -- more sharply peaked

#### 300 throws of three dice



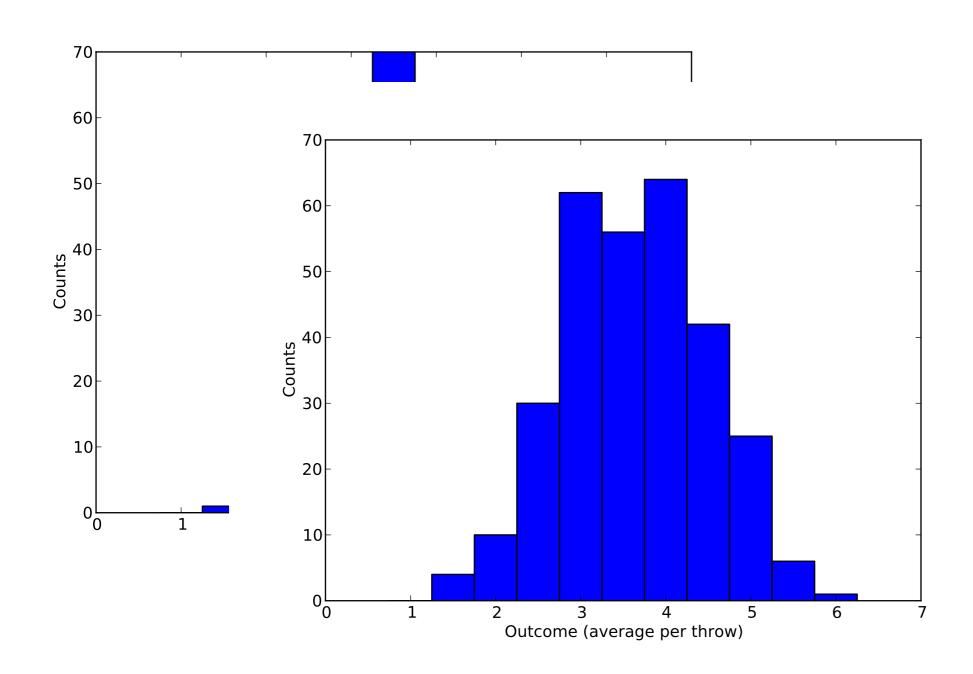
## Four dice -- even more sharply peaked

#### 300 throws of four dice



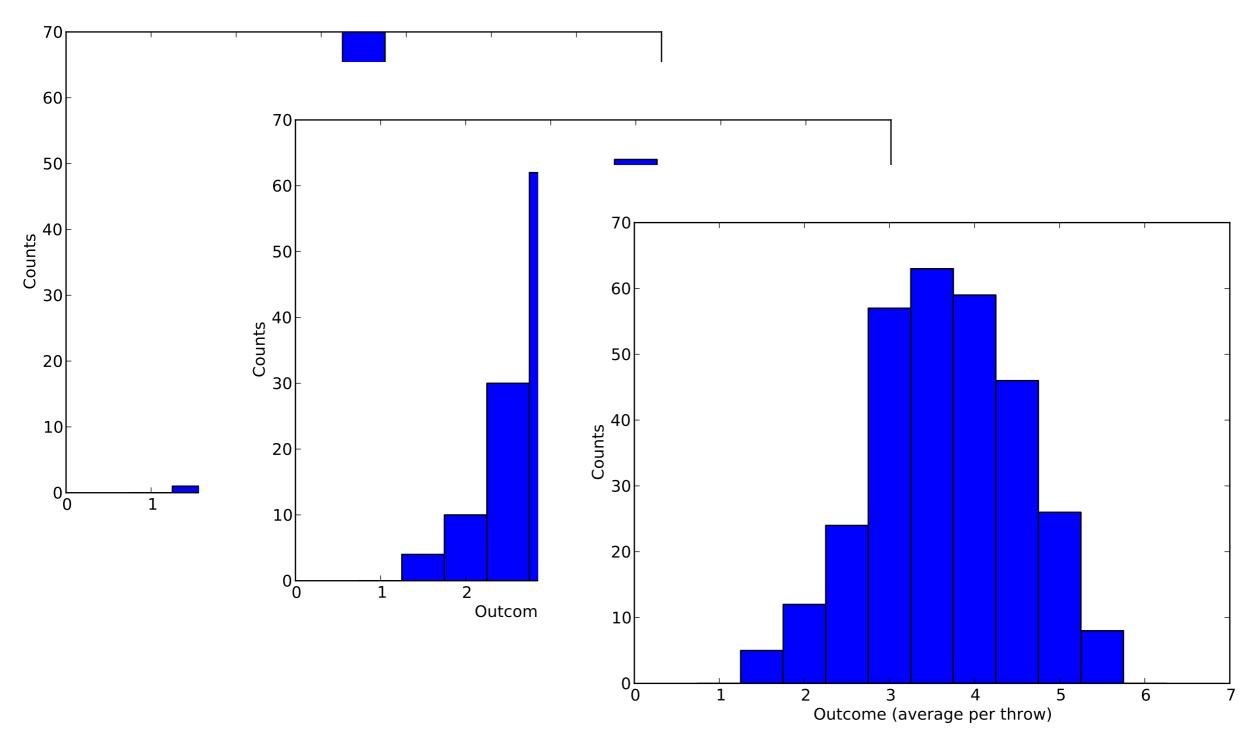
## Four dice -- even more sharply peaked

#### 300 throws of four dice



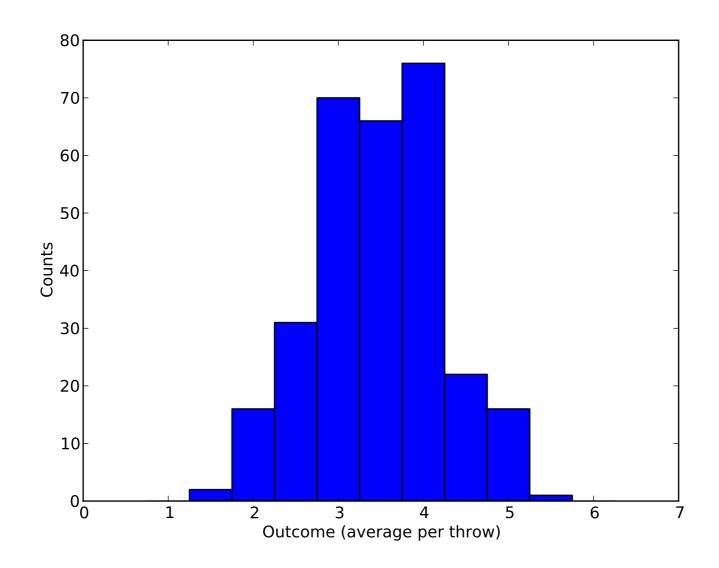
## Four dice -- even more sharply peaked

#### 300 throws of four dice



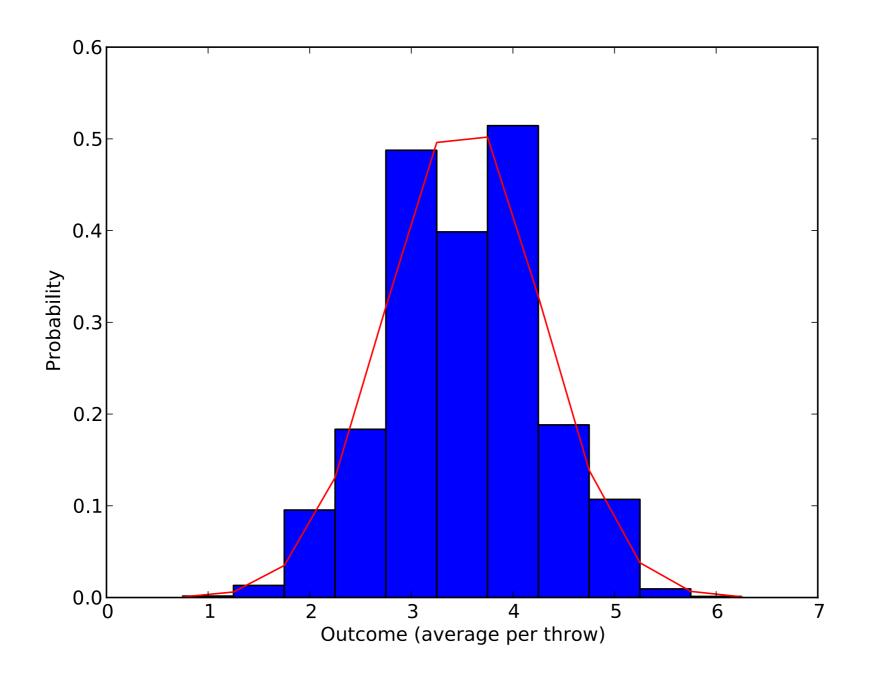
### Five dice -- is this starting to look familiar?

#### 300 throws of five dice

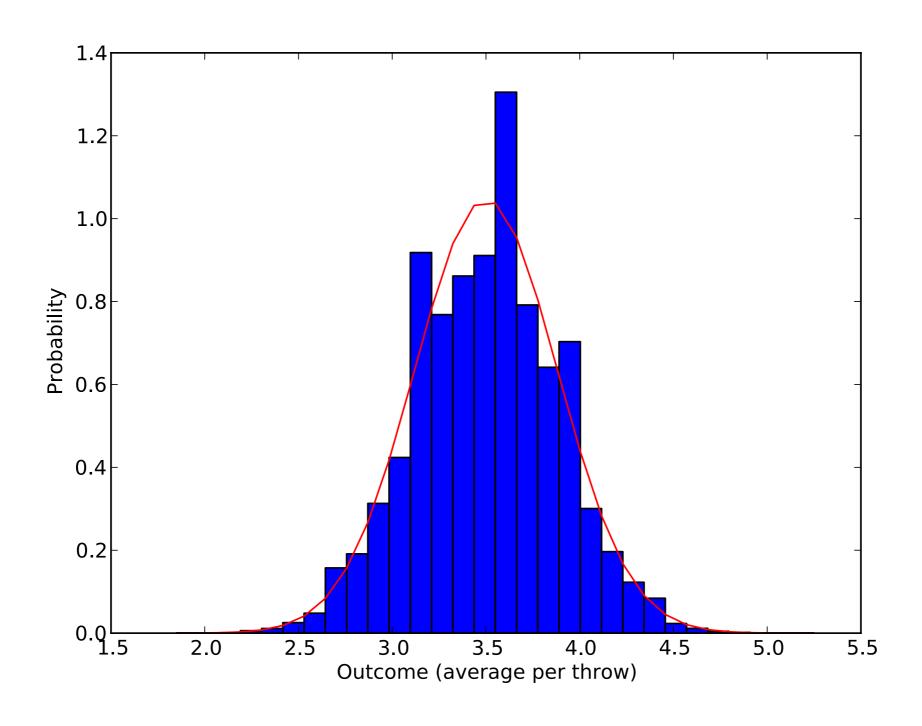


## Five dice -- is this starting to look familiar?

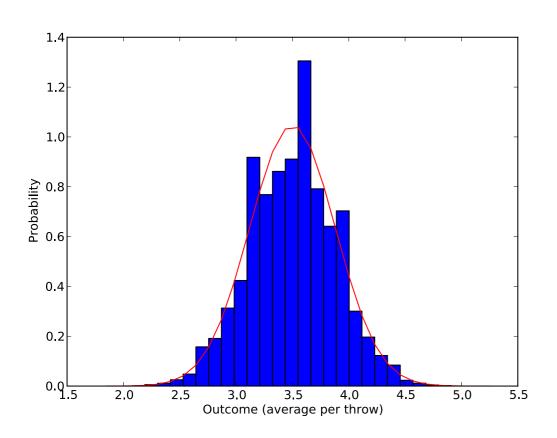
#### 10000 throws of five dice



### What about 20 dice?

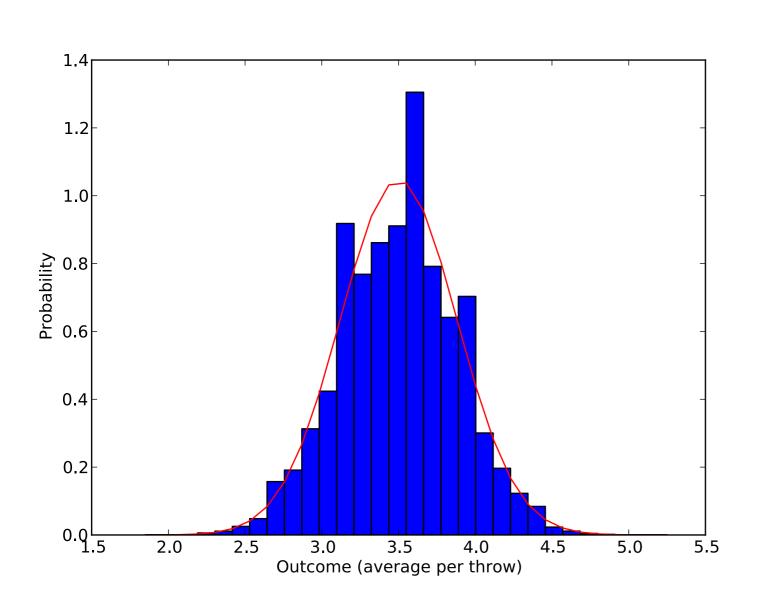


## It turns out there is something pretty profound behind this -- the "central limit theorem"



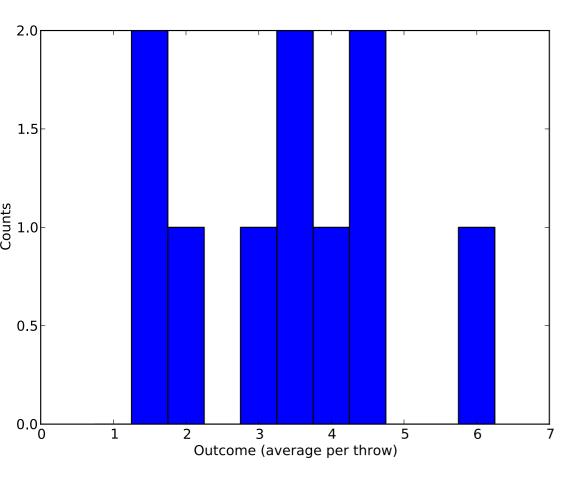
The distribution of a sum of many small values will follow a Gaussian (normal) distribution, regardless of the starting distribution

# When we compute observables from simulations, we usually expect property estimates to follow the central limit theorem

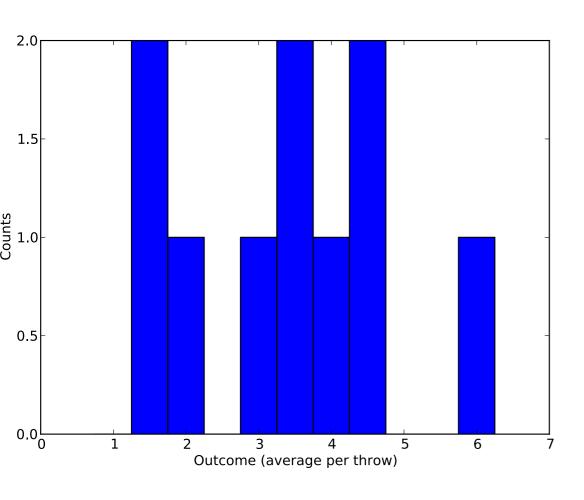


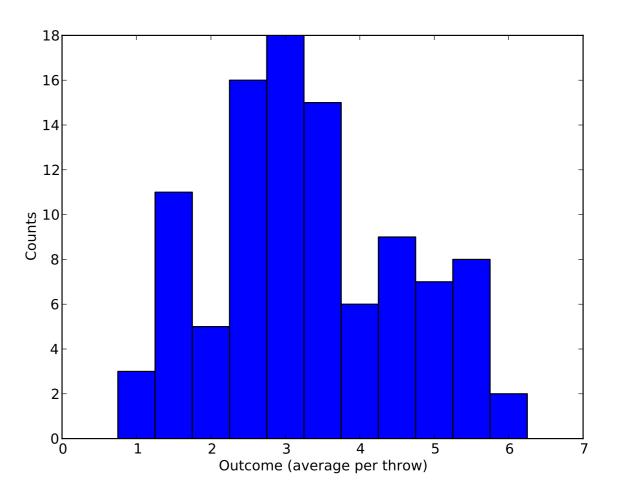
$$\langle f \rangle \doteq \frac{1}{N} \sum_{i=1}^{N} f(x_i)$$

# If we make a measurement from some distribution of possible results, how do we assess reliability?

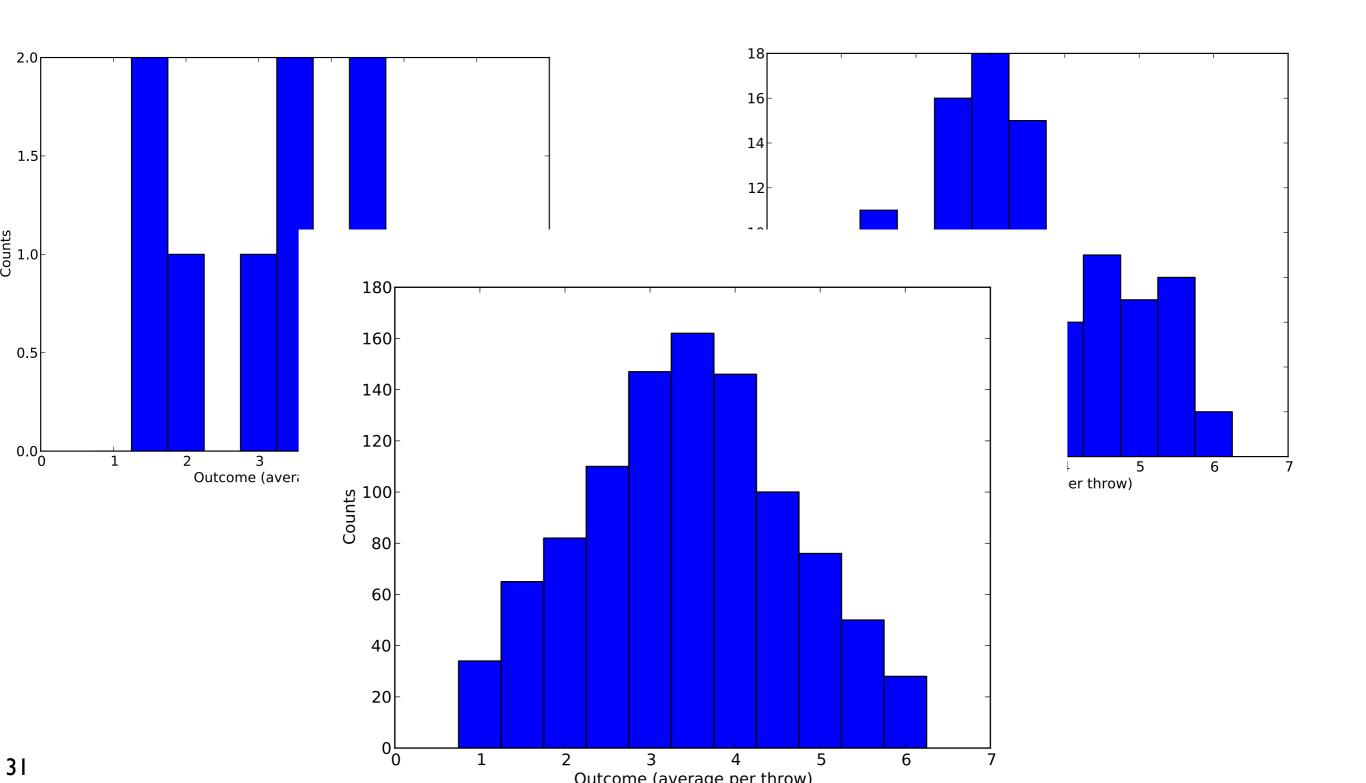


# If we make a measurement from some distribution of possible results, how do we assess reliability?





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# We need more information than $\sigma$ since collecting more data doesn't change $\sigma$ (just our estimate of it)

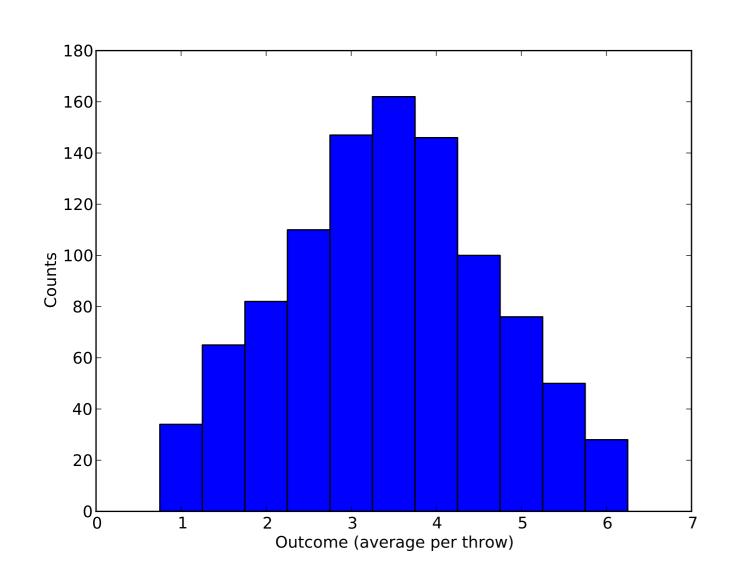
Particularly, we want the "standard error of the mean" -- "std-err"  $\sigma_{\theta} = \frac{\sigma}{\sqrt{N}}$ 

where N is the number of samples

sometimes called the uncertainty

(Assumes a symmetric distribution

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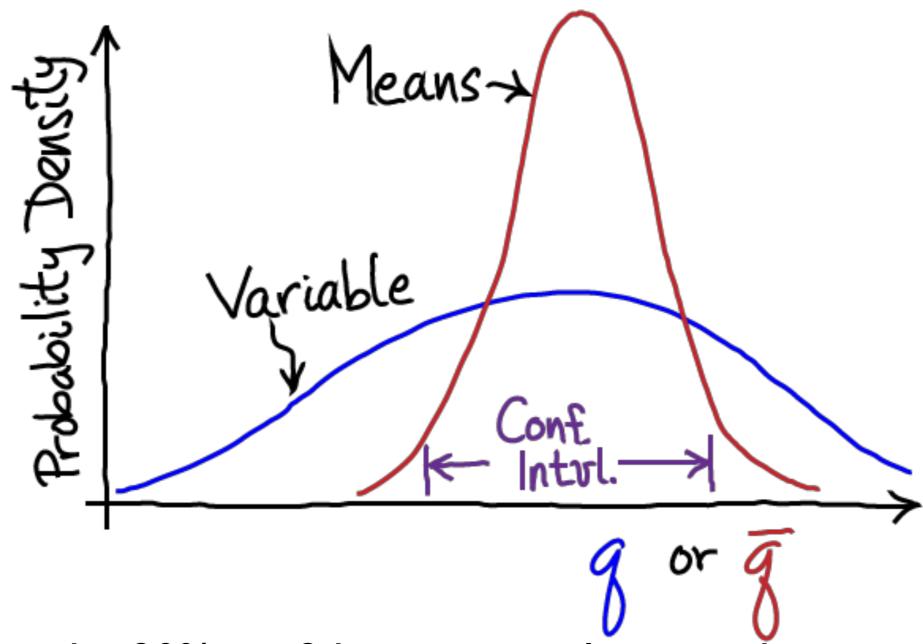
## Confidence intervals are another way of thinking about uncertainty

Such as, "90% chance that the true value of our observable falls between X and Y", i.e. "A 90% chance that the mean age of graduate students is between 25 and 26".

Turns out for a unimodal Gaussian distribution,  $2\sigma(err)$  is essentially a 90% confidence interval

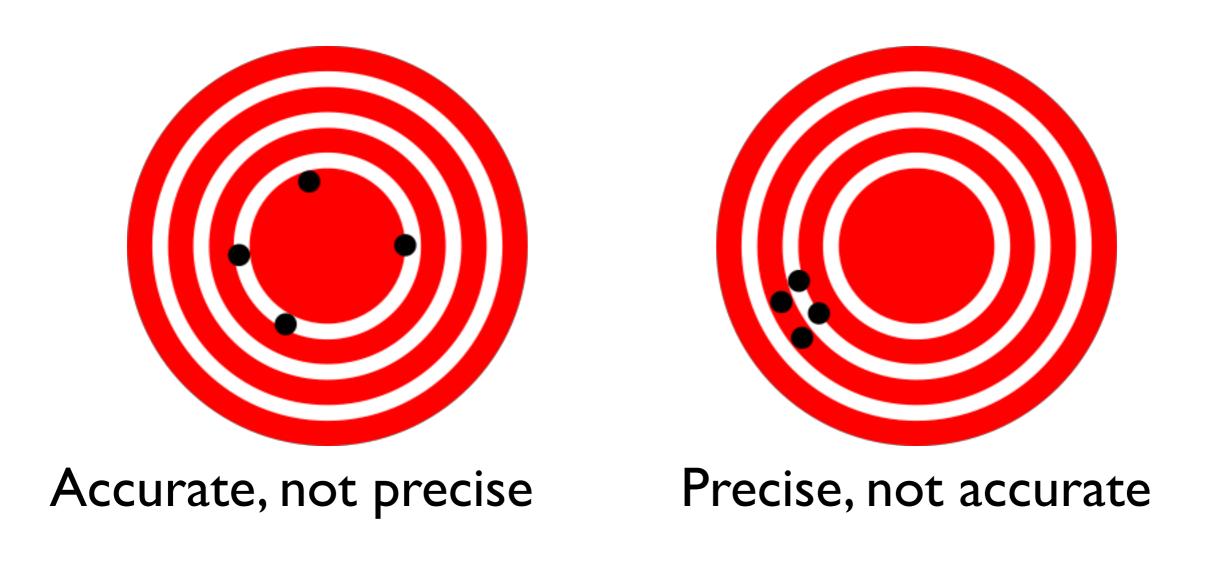
But confidence intervals work for asymmetric distributions

### Confidence intervals are very relevant for us



In this case the 90% confidence interval means there is a 90% chance any future (equivalent) simulation would fall within that range

## Confidence intervals and standard error assess precision of results



To assess accuracy, you have to calibrate/compare to other methods or experiment

## It's important to remember the difference between accuracy and precision

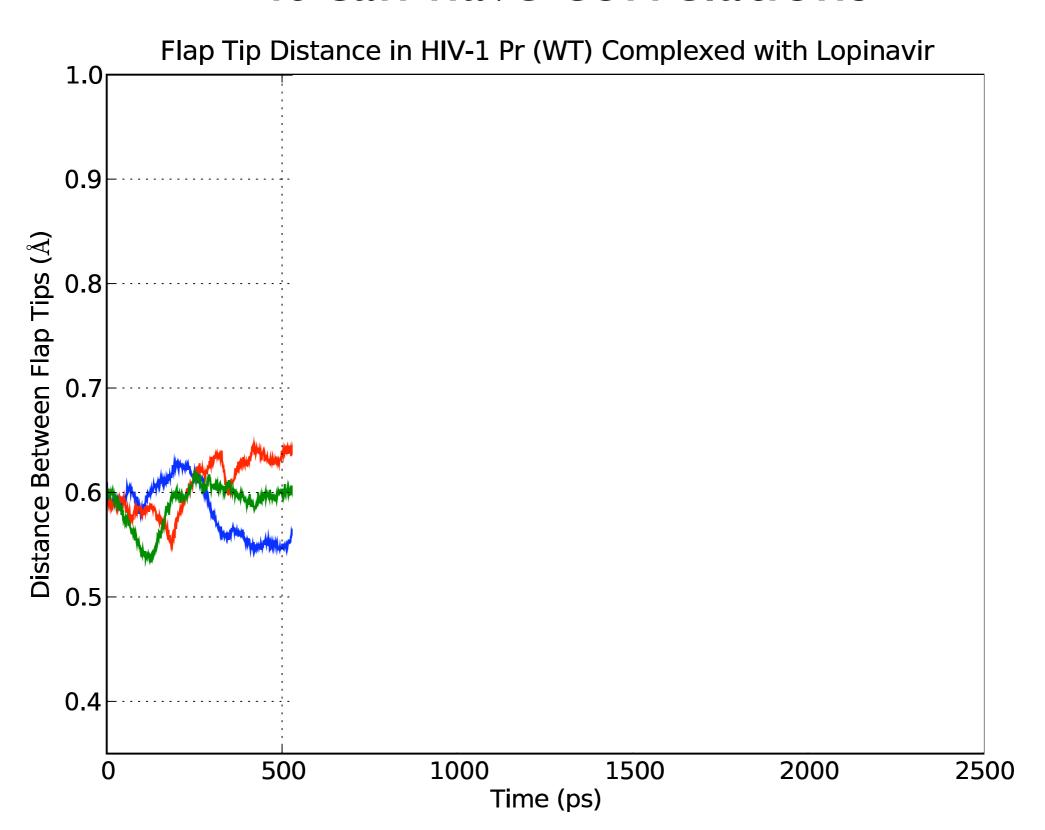
- Precision a computed observable (usually a mean over observations) is very precise if
  - We are very certain of the value we are reporting, that is:
  - Our error bars/uncertainties are very small
- A value is accurate if it agrees well with the true value
- A result may be accurate and not precise, or precise and not accurate, both, or neither

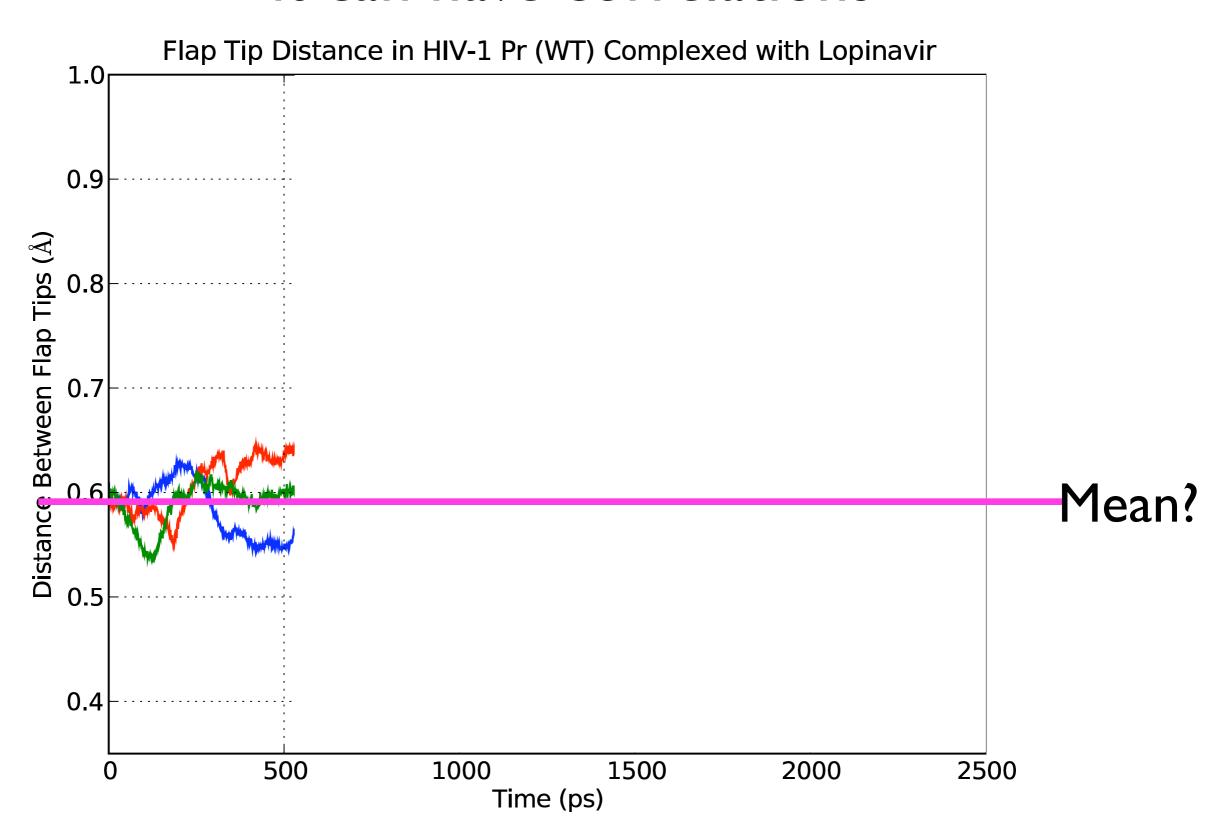
## Keeping standard errors/Cls in mind can avoid wild goose chases

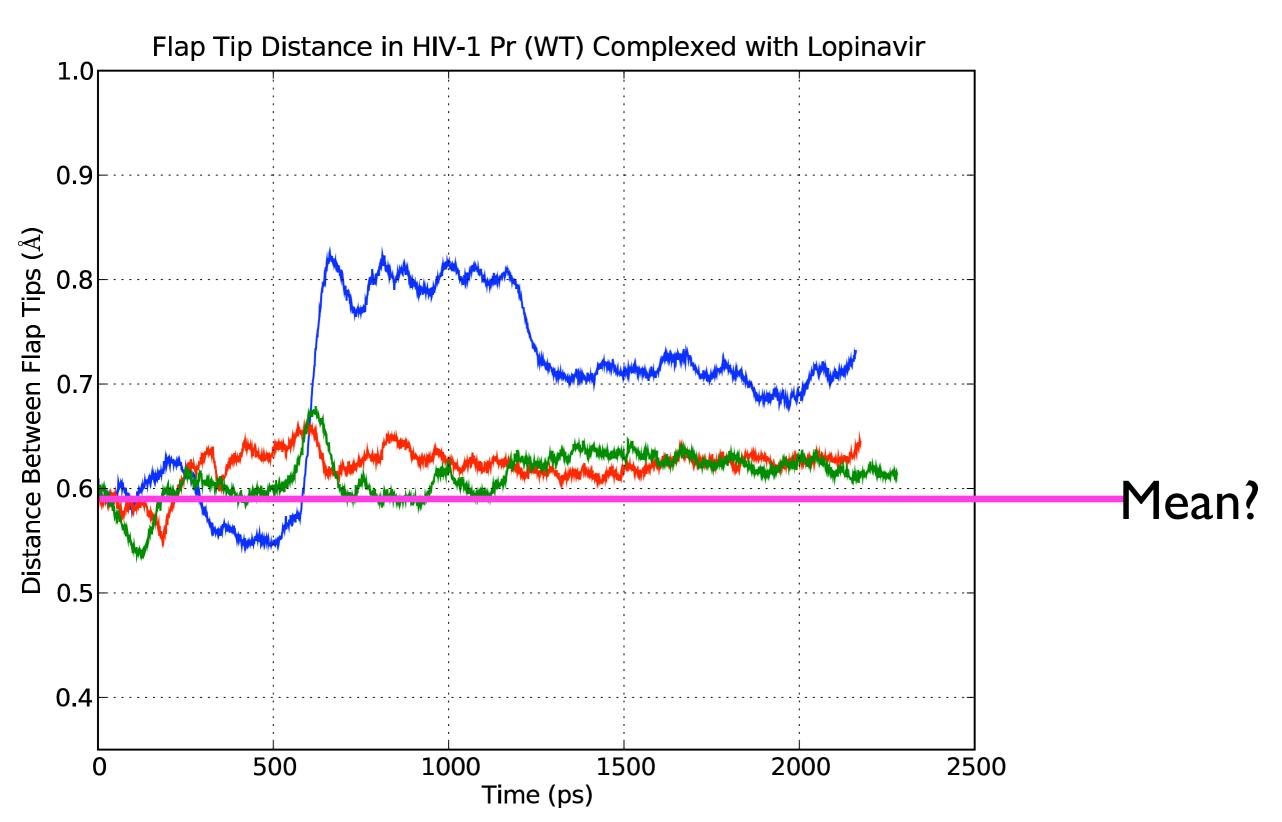
- i.e. experimental value for some observable is 3.5, and we compute 4.3. Difference is 0.7.
  - How precise are the values?
  - 3.5+/-1.5, 4.3+/-1.0; averaging these gives 3.9+/-1.8; both expt. and calculated values fall within uncertainty of this, so the results **agree** 
    - (Note that uncertainties combine, for a sum, by  $sqrt(\sigma_1^2 + \sigma_2^2)$ )
  - But, if the values were 3.5+/-0.1 and 4.3+/-0.2, they definitely disagree

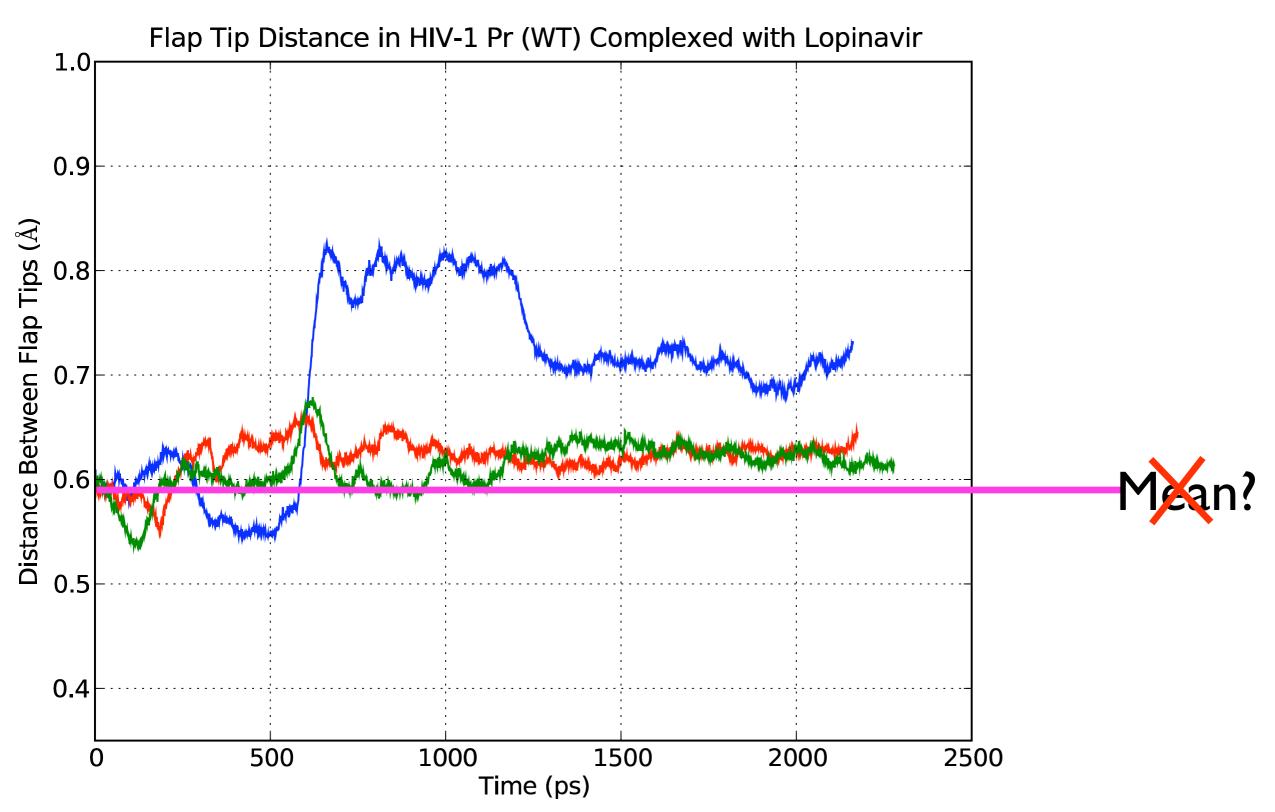
#### • Warnings:

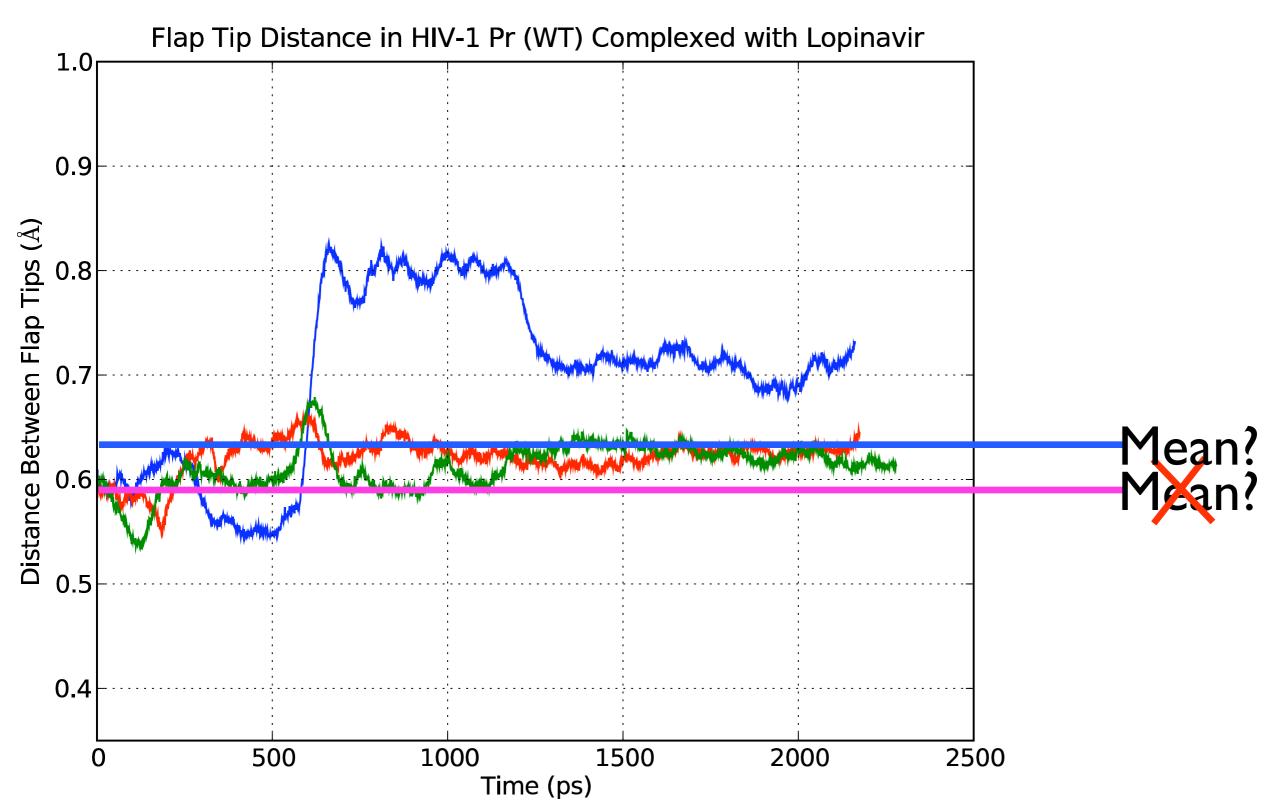
- Don't waste time tracking down differences between numbers when your precision is too low!
- Don't come up with explanations for differences that aren't significant
- ALWAYS know your precision before making comparisons

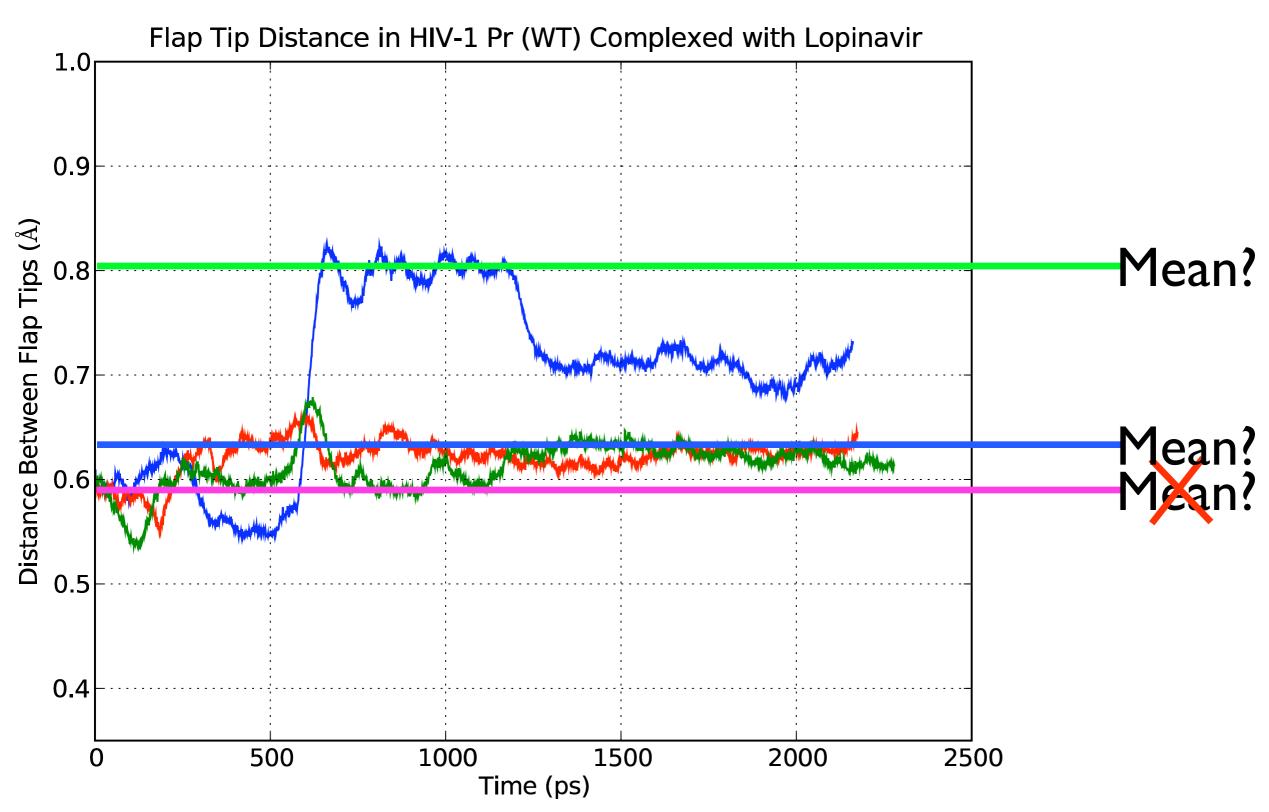


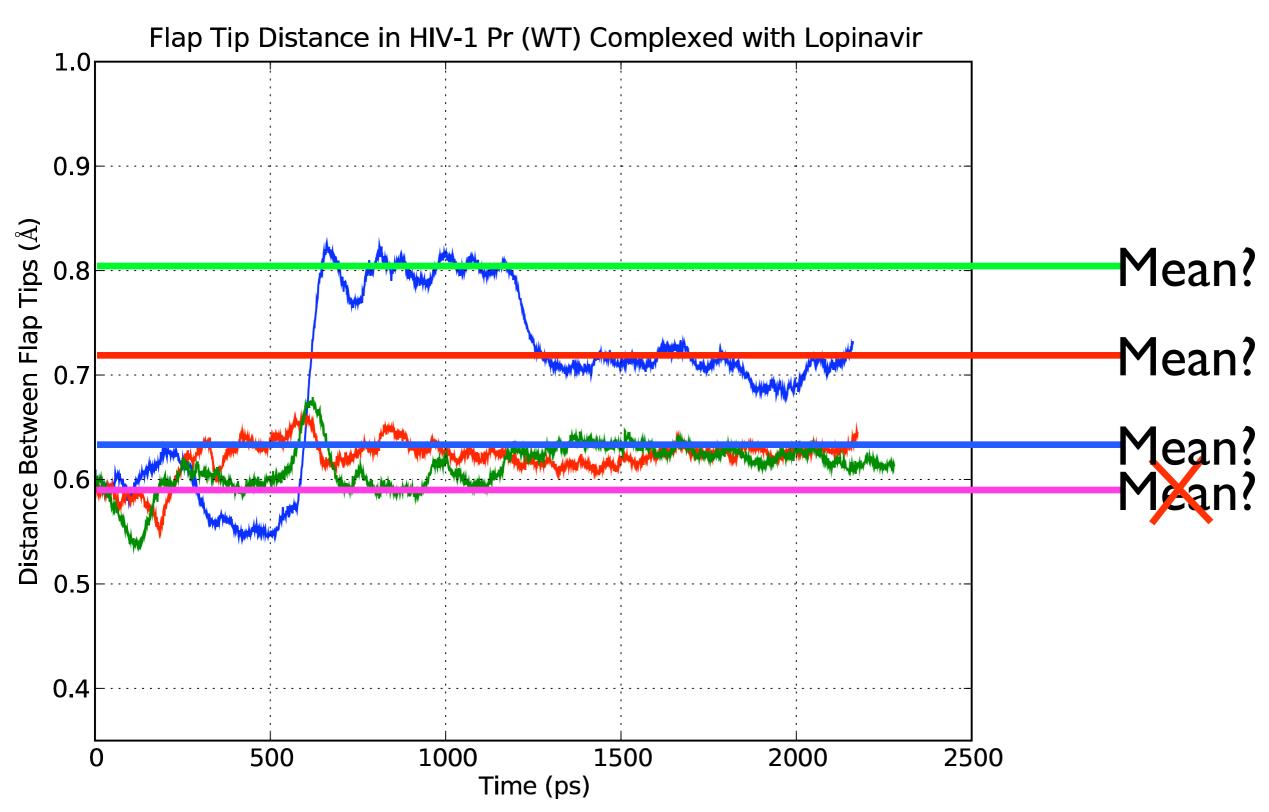




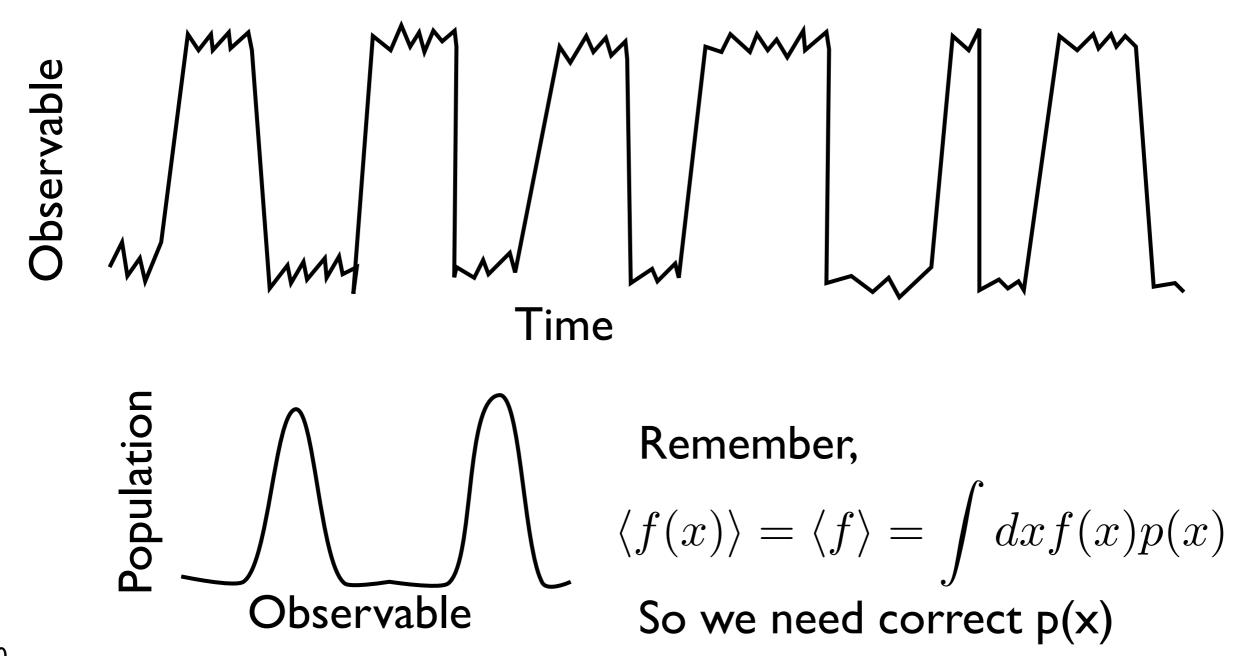




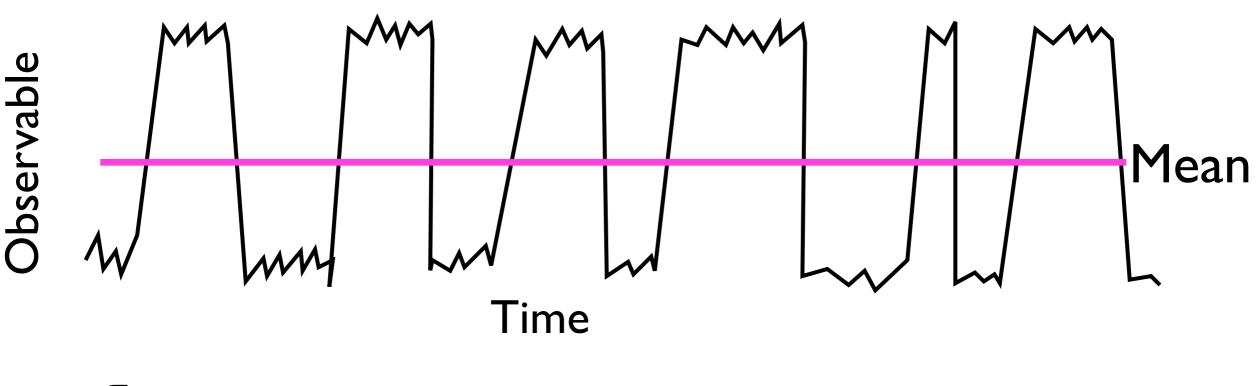


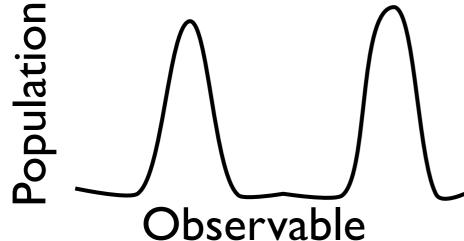


# To get an accurate value of an observable with transitions, we need to see enough transitions to get the populations correct



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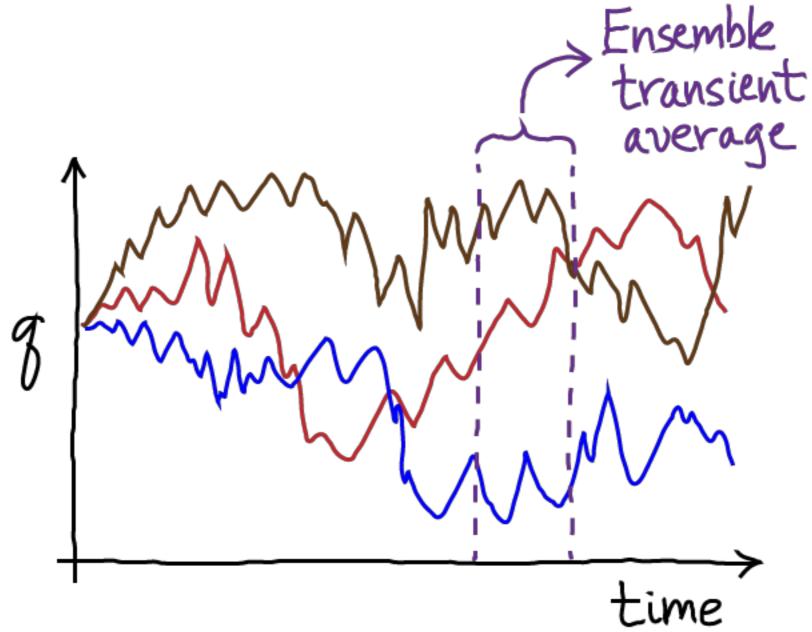


Remember,

$$\langle f(x) \rangle = \langle f \rangle = \int dx f(x) p(x)$$

So we need correct p(x)

## For uncertainty with "too short trajectories, perhaps use an ensemble transient average

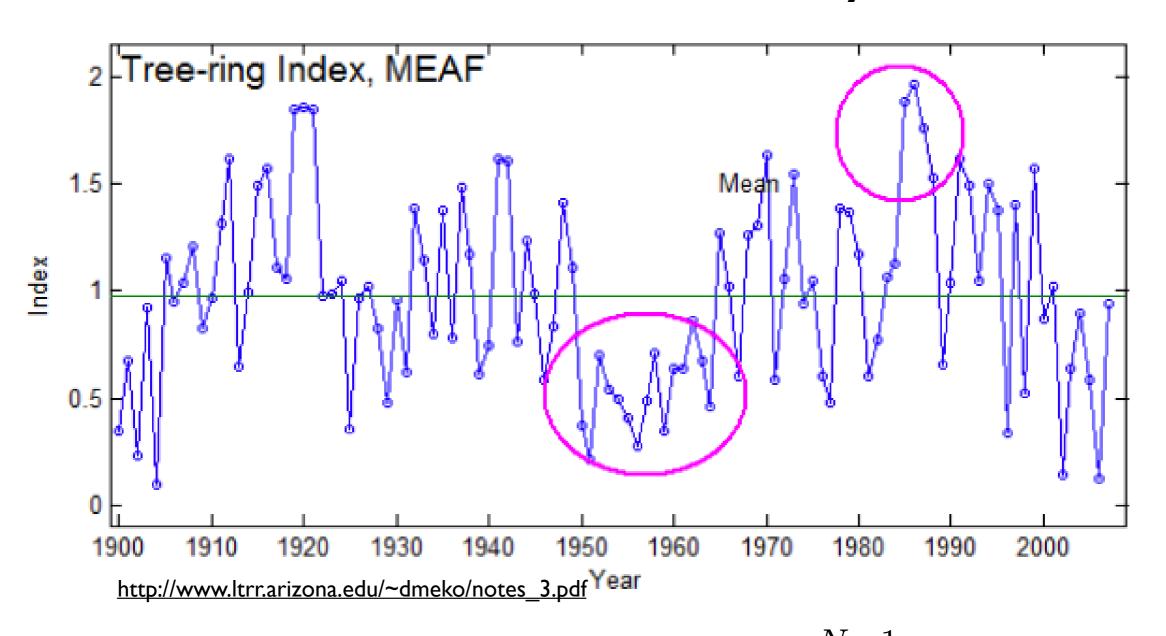


Having several "too short" trajectories generated independently, we can at least estimate a conditional uncertainty (given starting state and time elapsed), Zuckerman points out

## For correlated data, we have to deal with the number of effective samples

- Compute the standard error of the mean, using error analysis
  - Standard error:  $\sigma_{\theta}^2 = \frac{\sigma_x^2}{n}$
  - ullet But if we take frequent samples, many may be correlated  $N_{eff}=n/g$   $\sigma_{ heta}^2=rac{\sigma_x^2}{N_{eff}}$
  - For correlated samples, we have fewer "effective" samples:
    - g is the "statistical inefficiency" and is related to how correlated the measurements are; we will come back to this
- Error analysis should always be done, but also useful to test that separate trials agree
  - Some use standard error over trials as estimate of uncertainty, but probably not adequate

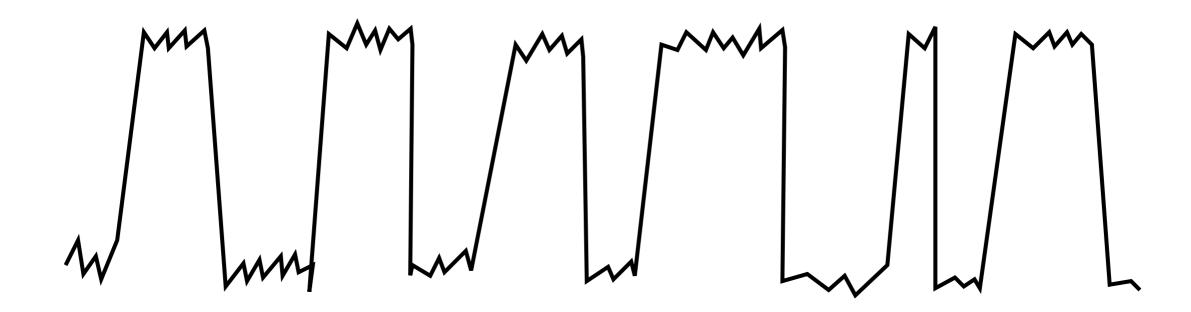
## Correlation detection can be done using autocorrelation analysis



Autocorrelation function

$$\rho_i = \sum_{j=0}^{N-1} a_j a_{j+i}$$

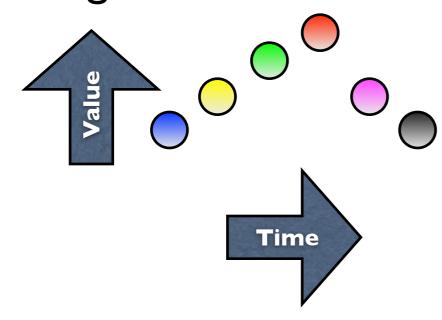
## The autocorrelation time measures the timescale for correlations



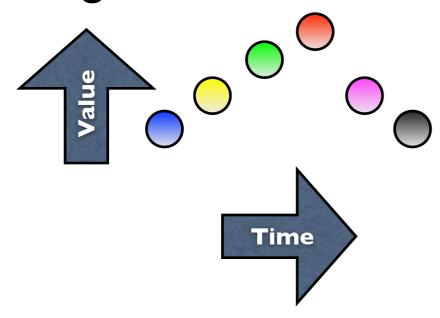
- An easy way to factor this in:
  - conda install pymbar
  - pymbar.statisticalInefficiency( timeseries ) gives the statistical inefficiency g

• 
$$N_{eff} = N/g$$

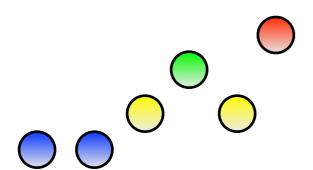
#### Original data



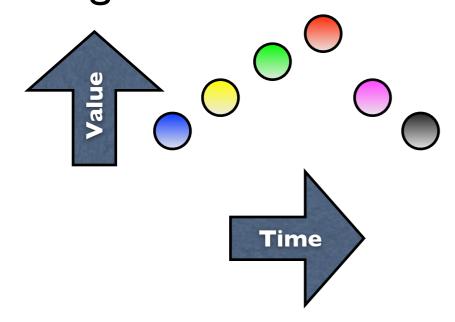
Original data



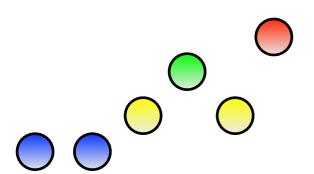
Bootstrap set I



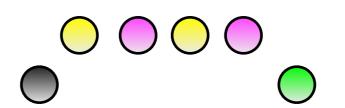
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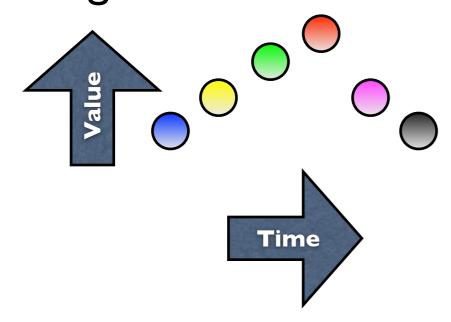
Bootstrap set I



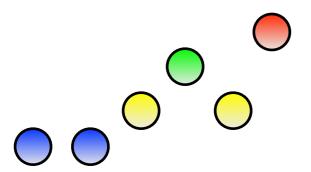
Bootstrap set 2



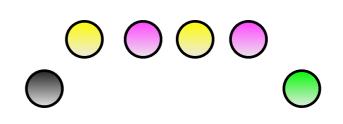
Original data



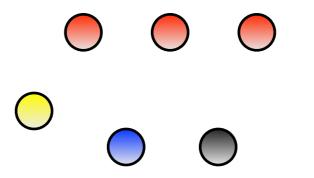
Bootstrap set I



Bootstrap set 2



Bootstrap set 3



# Time correlations may not be that important in much experimental data -- but know when to check

- What is the timescale of your measurement compared the timescales of the system?
- Are there conditions which could be changing with time (hr, day) in your lab?
- etc.

## Error propagates when doing math -- propagation is easy for addition and subtraction

$$\sigma_{\theta,z}^2 = \sigma_{\theta,x}^2 + \sigma_{\theta,y}^2$$

for 
$$z = x + y$$
 or  $z = x - y$ 

that is

$$\sigma_{\theta,z} = \sqrt{\sigma_{\theta,x}^2 + \sigma_{\theta,y}^2}$$

## Error propagation is harder when mathematical expressions are more involved

In general for f(x,y):

$$\sigma_{\theta,f}^2 = \left(\frac{\partial f(x,y)}{\partial x}\right)^2 \sigma_{\theta,x}^2 + \left(\frac{\partial f(x,y)}{\partial y}\right)^2 \sigma_{\theta,y}^2$$

• For example for multiplication,  $z = x^*y$ :

$$\frac{\partial f(x,y)}{\partial x} = y \qquad \frac{\partial f(x,y)}{\partial y} = x$$

$$\sigma_{\theta,f}^2 = y^2 \sigma_{\theta,x}^2 + x^2 \sigma_{\theta,y}^2 \quad \text{or} \quad \frac{\sigma_{\theta,f}}{xy} = \sqrt{\left[\frac{\sigma_{\theta,x}}{x}\right]^2 + \left[\frac{\sigma_{\theta,y}}{y}\right]^2}$$

## The "null hypothesis" is vitally important in experimental design and analysis

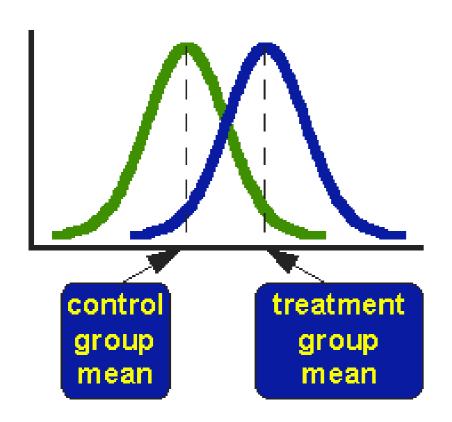
- Null hypothesis: An uninteresting outcome for your experiment
  - i.e. what you are changing makes no difference in the outcome
  - or you are observing only noise/random fluctuations
  - or you observe no binding of your compound
  - or...

### The null hypothesis can never be proven, but it can be rejected or not rejected

- An experiment can reject or fail to reject the null hypothesis
- Example: Testing new treatment on control vs test groups
  - With enough data you may be able to reject the null hypothesis
  - Without enough data, there MAY still be a difference, but not one you can see

# The t-test is a commonly used approach to assess significance of results (rejection of the null hypothesis)

 The t-value relates the difference in measured means to the standard error in the mean



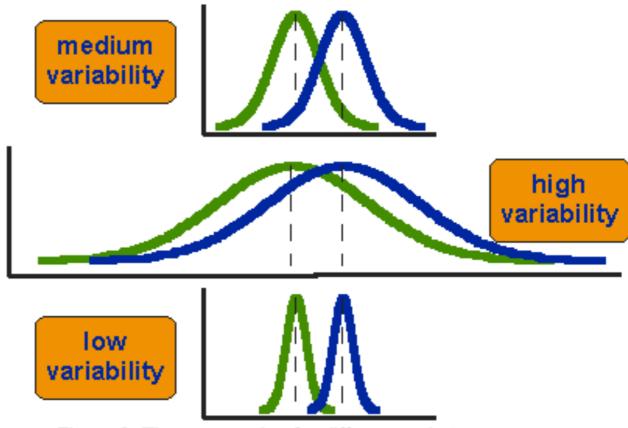


Figure 2. Three scenarios for differences between means.

## The t-value is easy to calculate (though there are different forms for different cases)

 The t-value relates the difference in measured means to the standard error in the mean

$$t = \frac{\mu_c - \mu_t}{\sqrt{\sigma_{\theta,c}^2 + \sigma_{\theta,t}^2}}$$

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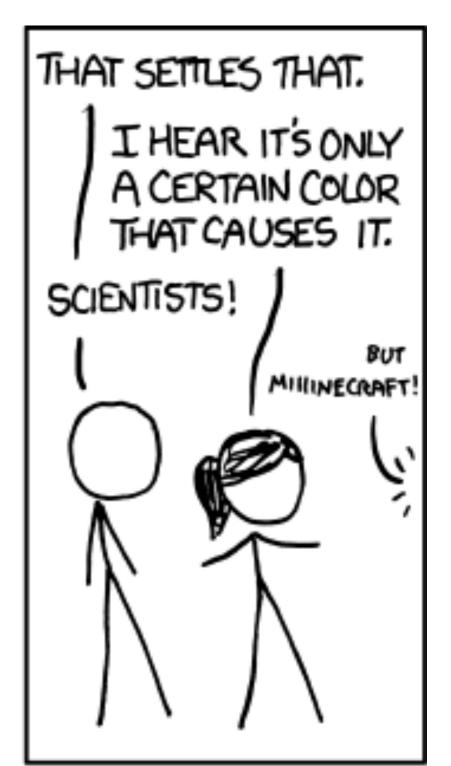
$$t = \frac{\mu_c - \mu_t}{\sqrt{\sigma_{\theta,c}^2 + \sigma_{\theta,t}^2}}$$

Degrees of Freedom	Probability, p			
	0.1	0.05	0.01	0.001
1	6.31	12.71	63.66	636.62
2	2.92	4.30	9.93	31.60
3	2.35	3.18	5.84	12.92
4	2.13	2.78	4.60	8.61
Tabulated t valu	ies for the	confidence	with which	VOU <sup>6.87</sup> 2n
rabaraced c var	1.94	2.45	3.71	1 ) Ou Carr

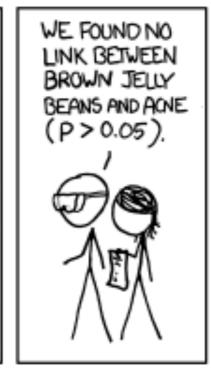
reject the null hypothesis

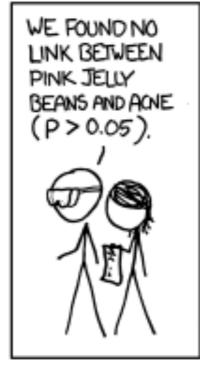


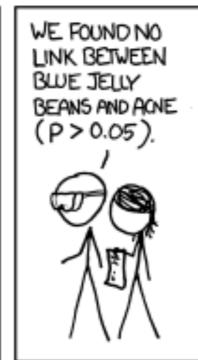


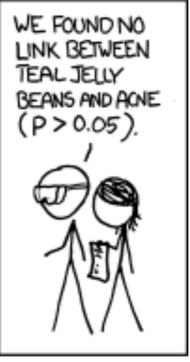




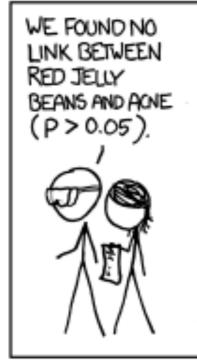


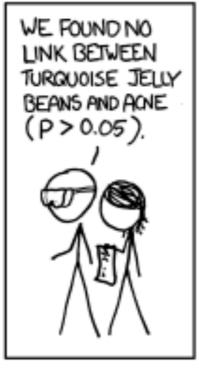


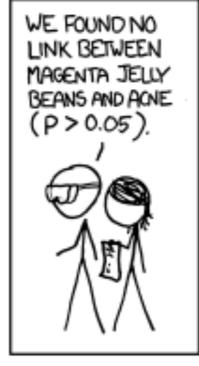


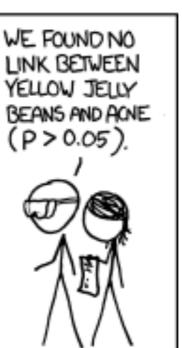












WE FOUND NO LINK BETWEEN PURPLE JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN BROWN JELLY BEANS AND ACNE (P > 0.05).



RED JELLY

BEANS AND A

(P>0.05

WE FOUND NO LINK BETWEEN PINK JELLY BEANS AND ACNE (P>0.05).



WE FOUND NO LINK BETWEEN BLUE JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN TEAL JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN SALMON JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO
LINK BETWEEN
GREY JELLY
BEANS AND ACNE
WE FOUND N
(P > 0.05).



WE FOUND NO LINK BETWEEN TAN JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN CYAN JELLY BEANS AND ACNE (P>0.05).



WE FOUND A LINK BETWEEN GREEN JELLY BEANS AND ACNE (P < 0.05).



WE FOUND NO LINK BETWEEN MAUVE JELLY BEANS AND ACNE (P > 0.05)



WE FOUND NO LINK BETWEEN BEIGE JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN LILAC JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN BLACK JELLY BEANS AND ACNE (P > 0.05).

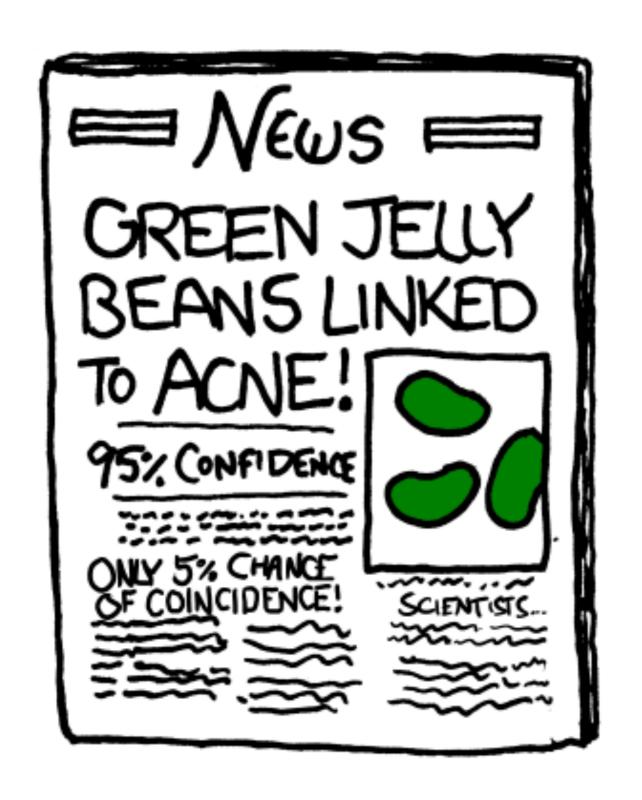


WE FOUND NO LINK BETWEEN PEACH JELLY BEANS AND ACNE (P > 0.05).



WE FOUND NO LINK BETWEEN ORANGE JELLY BEANS AND ACNE (P > 0.05).







### Who wants to spend their career analyzing statistical fluctuations?

- Tests for significance prevent wasted time
- Examples:
  - My lab: HIV protease, hydration work, ...
  - Presidential polling
  - From the class?

Note, though, there are other kinds of uncertainty such as model uncertainty