STAT1301 Advanced Analysis of Scientific Data Semester 2, 2025

Assignment 2

1 Introduction

Throughout the report, the following syntactical shortcuts and notation will be used.

If X is a random variable of the sample space Ω , an abbreviation of set notation is as follows:

Abbreviate
$$\{d : \forall d \in \Omega \text{ and } X(d) = x\}$$
 (1)

$$As \{X = x\}$$
 (2)

Additionally, when thinking in terms of sets becomes obselete,

Abbreviate
$$P(\{d : \forall d \in \Omega \text{ and } X(d) = x\})$$
 (3)

$$As P(X = x) (4)$$

The abbriviation will be used with inequalities as well.

Given some random variable X, there must exist a function mapping from the sample space Ω to the domain of X, which can be at most \mathbb{R} . This function is (intuitively) named X. This function incidentally defines the random variable, which is the motivating reason for using its letter to represent its mapping. The notation Domain[X] will be used throughout this report to indicate the domain of the function mapping X and hence the random variable X itself by definition.

2 Question 1

To begin, lets define the sample space

$$\Omega = \{(a, b) \in \{1, 2, 3, 4, 5, 6\}\}$$
$$|\Omega| = 36$$

Notice this is uniform, and hence that a and b are independent

$$P(A) = \frac{|A|}{|\Omega|} = \frac{|A|}{36} \tag{5}$$

2.1 Part a)

Let X be a random variable representing the payout of a given dice roll $(a, b) \in \Omega$:

$$X((a,b) \in \Omega) = a \cdot b$$

Let f_X be the PMF of X. Note $f_X(x \in \Omega) = P(\{X = x\})$. By cases, the probability distribution of X can be deduced:

$$f_{X}(1) = \frac{1}{36}$$

$$f_{X}(2) = \frac{2}{36}$$

$$f_{X}(2) = \frac{2}{36}$$

$$f_{X}(3) = \frac{2}{36}$$

$$f_{X}(4) = \frac{3}{36}$$

$$f_{X}(4) = \frac{3}{36}$$

$$f_{X}(5) = \frac{2}{36}$$

$$f_{X}(10) = \frac{2}{36}$$

For all other values x, $f_X(x) = 0$

2.2 Part b)

This makes determining the expected value of X trivial:

$$E(X) = \sum_{c \in \Omega} X(c)P(c)$$
 (6)

$$= \sum_{x \in \text{Domain}[X]} x P(\{X = x\}) \tag{7}$$

$$= 1 \cdot f_{X}(1) + 2 \cdot f_{X}(2) + \dots \cdot 30 \cdot f_{X}(30) + 36 \cdot f_{X}(36)$$
(8)

$$=\frac{1}{36} + \frac{4}{36} + \dots + \frac{60}{36} + \frac{36}{36} \tag{9}$$

$$=\frac{441}{36} = \frac{49}{4} = 12.25\tag{10}$$

2.3 Part c)

Evaluating Var(X) is similarly trivial

$$Var(X) = E[(X - E(X))^2]$$
(11)

$$= \sum_{c \in \Omega} (X(c) - \frac{49}{4})^2 P(\{c\})$$
 (12)

$$= \sum_{x \in \text{Domain}[X]} (x - \frac{49}{4})^2 P(\{X = x\})$$
 (13)

$$= (1 - \frac{49}{4})^2 \cdot \frac{1}{36} + (2 - \frac{49}{4})^2 \cdot \frac{2}{36} + \dots + (30 - \frac{49}{4})^2 \cdot \frac{2}{36} + (36 - \frac{49}{4})^2 \cdot \frac{1}{36}$$
 (14)

$$=\frac{11515}{144}\approx 79.97\tag{15}$$

$$\implies \sigma_{\mathcal{X}} = \sqrt{\operatorname{Var}(\mathcal{X})} = \sqrt{\frac{11515}{144}} \approx 8.942 \tag{16}$$

3 Question 2

Understanding this question in terms of a sample space isn't very fruitful. Ω is completely unspecified, we can only deduce that $|\Omega| \geq (0, 20)$, which implies it is continuous. $P(A) : \exists A \in \Omega$ is also completely unknown.

3.1 Part a)

Let X be the continuous random variable of algae growth as measured in grams of biomass produced. Note Domain[X] = (0, 20).

Since X is a random variable, its PDF f_X must sum to 1:

$$1 = \int_{c \in \Omega} P(\{c\}) \tag{17}$$

$$= \int_{x \in \text{Domain}[X]} P(\{X = x\})$$
 (18)

$$= \int_0^{20} c(x^2 - 60x + 800) dx \tag{19}$$

$$= c\left[\frac{1}{3}x^3 - 30x^2 + 800x\right]_{x=0}^{x=20} \tag{20}$$

$$1/c = \left[\frac{1}{3}(20)^3 - 30(20)^2 + 800(20)\right] - \left[0 - 0 + 0\right] \tag{21}$$

$$1/c = \frac{20000}{3} \tag{22}$$

$$c = \frac{3}{20000} \tag{23}$$

3.2 Part b)

Let F_X be the CDF of X:

$$F_{X} = \int_{-\infty}^{x} f_{X}(x) dx \tag{24}$$

$$= \int_0^x c(x^2 - 60x + 800) dx \tag{25}$$

$$= c\left[\frac{1}{3}x^3 - 30x^2 + 800x\right]_{x=0}^{x=x}$$
 (26)

$$\frac{F_{\rm X}}{c} = \left[\frac{1}{3}x^3 - 30x^2 + 800x\right] - \left[\frac{1}{3}0^3 - 30 \cdot 0^2 + 800 \cdot 0\right] \tag{27}$$

$$\implies F_{\mathcal{X}} = c(\frac{1}{3}x^3 - 30x^2 + 800x) \text{ for } 0 \le x \le 20$$
 (28)

$$=\frac{1}{20000}x^3 - \frac{9}{2000}x^2 + \frac{3}{25}x\tag{29}$$

3.3 Part c)

$$E(X) = \int_{x \in Domain[X]} x f_X dx$$
(30)

$$= \int_0^{20} x \cdot c(x^2 - 60x + 800) dx \tag{31}$$

$$\frac{E(X)}{c} = \int_0^{20} x^3 - 60x^2 + 800x dx \tag{32}$$

$$= \left[\frac{1}{4}x^4 - 20x^3 + 400x^2\right]_{x=0}^{x=20} \tag{33}$$

$$= \left[\frac{1}{4}(20)^4 - 20(20)^3 + 400(20)^2\right] - \left[0 - 0 + 0\right] \tag{34}$$

$$= 40000 - 160000 + 160000 \tag{35}$$

$$E(X) = c \cdot 40000 \tag{36}$$

$$E(X) = 6 \text{ grams} \tag{37}$$

3.4 Part d)

$$P(\{X > 10\} | \{X > 2\}) \tag{38}$$

$$= \frac{P(\{X > 10\} \cap \{X > 2\})}{\{X > 2\}} \tag{39}$$

$$= \frac{P(\{X > 10\})}{P(\{X > 2\})} \tag{40}$$

From the CDF definition of X, $P({X < x}) = f_X(x)$

$$\implies P(\{X > 10\}) = 1 - P(\{X < 10\}) \tag{41}$$

$$= 1 - F_{\rm X}(10) \tag{42}$$

$$=1-\frac{4}{5} (43)$$

$$=\frac{1}{5}\tag{44}$$

$$\implies P(\{X > 2\}) = 1 - P(\{X < 2\})$$
 (45)

$$=1-F_{\rm X}(2)$$
 (46)

$$=1-\frac{139}{625}\tag{47}$$

$$=\frac{486}{625}\tag{48}$$

$$\implies \frac{P(\{X > 10\})}{P(\{X > 2\})} = \frac{\frac{1}{5}}{\frac{486}{625}} = \frac{125}{486} \approx 0.2572 \tag{49}$$

Therefore, the probability that the biomass exceeds 10 grams, given that it is detectable, is above in $(49) = \frac{125}{486}$.

4 Question 3

Assume that p = 0.25 for all the products, not just the 25 that were sampled.

The sample space for this is again completely unspecified, and the P probability function is practically useless for this question. For convenience, the sample space Ω is therefore defined as the domain of X, representing the number of products passing the specific inspection.

$$\Omega = \{1, 2, 3...24, 25\}$$

This makes the definition of X trivial, and its domain incidentally the entire sample space:

$$X(a \in \Omega) = a$$

4.1 Part a)

Since each product has a p = 0.25 probability of passing inspection, and there are 25 products, and it is assumed each inspection and product is independent of each other, X is a binomial distribution:

$$X \sim Bin(n = 25, p = 0.25)$$

Notes the following theorems about binomial distributions and X:

$$P(\{X = x\}) = \binom{n}{x} p^x (1-p)^{n-x} = \binom{25}{x} 0.25^x \cdot 0.75^{25-x}$$
$$E(X) = np = \frac{25}{4}$$
$$Var(X) = np(1-p) = \frac{75}{16}$$

4.2 Part b)

Let X_2 be the random variable representing the probability distribution of X with an n parameter such that the probability of finding a defect-free product eexceeds 99%:

$$X_2 \sim Bin(n, p = 0.25)$$

$$P(\{X_2 \ge 1\}) > 0.99 \tag{50}$$

$$0.99 < P(\{X_2 \ge 1\}) \tag{51}$$

$$0.99 < 1 - P(\{X_2 = 0\}) \tag{52}$$

$$0.99 - 1 < -\binom{n}{0}(0.25)^0(0.75)^n \tag{53}$$

$$0.01 > 1 \cdot 1 \cdot 0.75^n \tag{54}$$

$$\log_{0.75} 0.01 > n \tag{55}$$

$$\implies n < \log_{0.75} 0.01 \approx 16.008$$
 (56)

Therefore the minimum (integer) sample size is n = 16.

4.3 Part c)

The random variable Y is dependent on X. Given a possibility $a \in \Omega$ from the sample space, Y(a) explicitly depends upon X(a) such that it exactly equals:

$$Y(a \in \Omega) = 3X(a) - (25 - X(a))$$
 (57)

$$=4X(a)-25\tag{58}$$

This allows us to calculate E(X) and Var(X) relatively easily using probability theorems:

$$E(Y) = E(4X - 25)$$
 (59)

$$= 4E(X) - 25$$
 (60)

$$=4 \cdot \frac{25}{4} - 25 \tag{61}$$

$$=0 (62)$$

$$Var(Y) = Var(4X - 25) \tag{63}$$

$$=4^{2}Var(X) \tag{64}$$

$$= 16 \cdot \frac{75}{16} \tag{65}$$

$$=75\tag{66}$$

4.4 Part d)

Since Y is defined in terms of X, this isn't too difficult to evaluate:

$$P(\{Y \ge 27\}) = P(\{4X - 25 \ge 27\}) \tag{67}$$

$$= P(\{4X \ge 52\}) \tag{68}$$

$$= P(\{X \ge 13\}) \tag{69}$$

$$\approx 0.00337\tag{70}$$

This can be calculated by running 1 - pbinom(12, 25, 0.25) in R

5 Question 4

Let $\Omega = (-\infty, +\infty)$ in units °C, representing the continuous range of possible temperatures in the storeroom. An argument could be made to limit this to $(-\infty, 8)$.

Let X be a random variable for the temperature inside the storeroom.

5.1 Part a)

$$X \sim N(\mu = 7.5^{\circ}C, \sigma = 0.3^{\circ}C)$$

$$P({7.2 < X < 8}) = P({\frac{7.2 - 7.5}{0.3} < \frac{X - \mu}{\sigma} < \frac{8 - 7.5}{0.3}})$$
(71)

$$= P(\{-\frac{2}{3} < Z < \frac{5}{3}\}) \tag{72}$$

$$= P(\{Z < \frac{5}{3}\}) - P(\{-\frac{2}{3} < Z\})$$
 (73)

$$= P(\{Z < \frac{5}{3}\}) - P(\{Z > -\frac{2}{3}\})$$
 (74)

$$= P(\{Z < \frac{5}{3}\}) - (1 - P(\{Z < \frac{2}{3}\}))$$
 (75)

$$= P(\{Z < \frac{5}{3}\}) + P(\{Z < \frac{2}{3}\}) - 1$$
 (76)

Using stats tables this equals 0.9515 + 0.7454 - 1 = 0.6969. Using R running pnorm $(\frac{5}{3})$ - pnorm $(-\frac{2}{3}) = 0.6997$.

5.2 Part b)

$$X \sim N(\mu, \sigma = 0.3^{\circ}C)$$

$$P({X > 8^{\circ}C}) = 1\%$$
 (77)

$$0.01 = P(\{X > 8\}) \tag{78}$$

$$= 1 - P(\{X < 8\}) \tag{79}$$

$$0.99 = P(\{X < 8\}) \tag{80}$$

$$= P\left(\left\{\frac{X - \mu}{\sigma} = \frac{8 - \mu}{\sigma}\right\}\right) \tag{81}$$

$$0.99 = P(\{Z = \frac{8 - \mu}{3}\}) \tag{82}$$

(83)

Let z be the value which satisfies $P({Z < z}) = 0.99$.

$$\implies \frac{8-\mu}{3} = z \tag{84}$$

$$8 - \mu = 3z \tag{85}$$

$$-\mu = 3x - 8 \tag{86}$$

$$\mu = 8 - 3z \tag{87}$$

Using the stats table, $z\approx 2.33$ which implies $\mu\approx 8-3\cdot 2.33=1.01$ °C. Using R, z=qnorm(0.99) ≈ 2.326348 , which implies $\mu\approx 8-3\cdot 2.326348\approx 1.021$ °C.