1 Question 1

1.1 Part a)

Let Ω be the sample space. Therefore $P(\{\Omega\}) = 1$. Adding all the joint pmf values must sum to 1:

$$\{\Omega\} = \bigcup_x \bigcup_y \{X = x\} \cap \{Y = y\}$$

$$P(\{\Omega\}) = 1$$

$$\implies 1 = P((\{X = -1\} \cap \{Y = -1\}) \cup \ldots \cup (\{X = 1\} \cap \{Y = 1\}))$$

$$= P(\{X = -1\} \cap \{Y = -1\}) + \ldots + P(\{X = 1\} \cap \{Y = 1\}))$$

$$= (p - \frac{1}{16}) + (\frac{1}{4} - p) + (0) + (\frac{1}{8}) + (\frac{3}{16}) + (\frac{1}{8}) + (p + \frac{1}{16}) + (\frac{1}{16}) + (\frac{1}{4} - p)$$

$$1 = -\frac{1}{16} + \frac{4}{16} + \frac{7}{16} + \frac{1}{16} + \frac{1}{16} + \frac{4}{16}$$

$$1 = 1$$

Unfortunately, this tells us no information about p. From the definition of probability, $P(\{c\})$ for $c \in \Omega$ must be greater or equal to 0, $P(\{c \in \Omega\}) \geq 0$. This can be used to restrict the possible values of p:

$$P(A \subseteq \Omega) \geq 0$$

$$\implies P(\{X = -1\} \cap \{Y = -1\}) \geq 0$$

$$p - \frac{1}{16} \geq 0$$

$$p \geq \frac{1}{16}$$

$$\implies P(\{X = 0\} \cap \{Y = -1\}) \geq 0$$

$$\frac{1}{4} - p \geq 0$$

$$p \leq \frac{1}{4}$$

$$\implies P(\{X = -1\} \cap \{Y = 1\}) \geq 0$$

$$p + \frac{1}{16} \geq 0$$

$$p \leq \frac{1}{16}$$

$$\implies p \in [\frac{1}{16}, \frac{1}{4}] \tag{1}$$

Therefore, $\frac{1}{16} \leq p \leq \frac{1}{4}$, and can be any value within this range.

1.2 Part b)

Aim is to find $P(\{X = Y\})$:

$$P(\{X = Y\}) = \sum_a P(\{X = a\} \cap \{Y = a\})$$
$$= P(\{X = -1\} \cap \{Y = -1\}) + P(\{X = 0\} \cap \{Y = 0\}) + P(\{X = 1\} \cap \{Y = 1\})$$
$$= (p - \frac{1}{16}) + (\frac{3}{16}) + (\frac{1}{4} - p)$$
$$= \frac{6}{16} = \frac{3}{8}$$

1.3 Part c)

The marginal pdf of X is $f_X(x)$, which is equal to $P(\{X = x\})$ and can be manually evaluated:

$$P(\{X = -1\}) = \sum_y P(\{X = -1\} \cap \{Y = y\})$$
$$= (p - \frac{1}{16}) + (\frac{1}{8}) + (p + \frac{1}{16})$$
$$= 2p + \frac{1}{8}$$
$$P(\{X = 0\}) = \sum_y P(\{X = 0\} \cap \{Y = y\})$$
$$= (\frac{1}{4} - p) + (\frac{3}{16}) + (\frac{1}{16})$$
$$= -p + \frac{1}{2}$$
$$P(\{X = 1\}) = \sum_y P(\{X = 1\} \cap \{Y = y\})$$
$$= (0) + (\frac{1}{8}) + (\frac{1}{4} - p)$$
$$= -p + \frac{3}{8}$$
$$\implies f_X(x) = P(\{X = x\}) = \begin{cases} 2p + \frac{1}{8} & x = -1 \\ -p + \frac{1}{2} & x = 0 \\ -p + \frac{3}{8} & x = 1 \\ 0 & \text{otherwise} \end{cases}$$

$$P(\{Y = -1\}) = \sum_x P(\{X = x\} \cap \{Y = -1\})$$

$$= (p - \frac{1}{16}) + (\frac{1}{4} - p) + (0)$$

$$= \frac{3}{16}$$

$$P(\{Y = 0\}) = \sum_x P(\{X = x\} \cap \{Y = 0\})$$

$$= (\frac{1}{8}) + (\frac{3}{16}) + (\frac{1}{8})$$

$$= \frac{7}{16}$$

$$P(\{Y = 1\}) = \sum_x P(\{X = x\} \cap \{Y = 1\})$$

$$= (p + \frac{1}{16}) + (\frac{1}{16}) + (\frac{1}{4} - p)$$

$$= \frac{6}{18} = \frac{3}{8}$$

$$\implies f_Y(x) = P(\{Y = y\}) = \begin{cases} \frac{3}{16} & y = -1 \\ \frac{7}{16} & y = 0 \\ \frac{3}{8} & y = 1 \\ 0 & \text{otherwise} \end{cases}$$

1.4 Part d)

X and Y are independant if

$$P(\{X = x\} \cap \{Y = y\}) = P(\{X = x\}) \cdot P(\{Y = y\}) = f_X(x) \cdot f_Y(y) \qquad (2)$$

for all possible values x and y. Therefore, this must be true for $x = -1$ and $y = 1$:

$$\text{RHS} = P(\{X = -1\})P(\{Y = 1\})$$

$$\text{LHS} = P(\{X = -1\} \cap \{Y = 1\}) \qquad = (2p + \frac{1}{8})(\frac{3}{8})$$

$$= p + \frac{1}{16} \qquad = \frac{3}{4}p + \frac{3}{64}$$

As shown above, LHS and RHS are only equal for zero or one values of p. Letting LHS = RHS, we can find this exact value (or lack thereof):

$$p + \frac{1}{16} = \frac{3}{4}p + \frac{3}{64}$$

$$\frac{1}{4}p = \frac{3}{64} - \frac{1}{16}$$

$$p = -\frac{1}{64} \cdot 4 = -\frac{1}{16}$$

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Therefore LHS = RHS only when $p = -\frac{1}{16}$, however from (1) this is not within the potential domain of p. Therefore LHS $\neq$ RHS, showing one counterexample to (2), hence X and Y are not independent.

1.5 Part e)

$$\mathrm{E(X)} = \sum_x x\mathrm{P}(\{X = x\})$$

$$= -1(2p + \frac{1}{8}) + 0(-p + \frac{1}{2}) + 1(-p + \frac{3}{8})$$

$$= -2p - \frac{1}{8} - p + \frac{3}{8}$$

$$\therefore \mathrm{E(X)} = -3p + \frac{1}{4}$$

$$\mathrm{E(Y)} = \sum_y y\mathrm{P}(\{Y = y\})$$

$$= -1(\frac{3}{16}) + 0(\frac{7}{16}) + 1(\frac{3}{8})$$

$$\therefore \mathrm{E(Y)} = -\frac{3}{16} + \frac{3}{8} = \frac{3}{16}$$

$$\mathrm{Cov(X, Y)} = \mathrm{E}[(X - \mathrm{E(X)})(Y - \mathrm{E(Y)}]$$

$$\mathrm{Cov(X, Y)} = \sum_{c \in \Omega}(X(c) - \mathrm{E(X)})(Y(c) - \mathrm{E(Y)})\mathrm{P}(\{c\})$$

$$= \sum_{x,y}(x - (-3p + \frac{1}{4}))(y - (\frac{3}{16}))\mathrm{P}(\{X = x\} \cap \{Y = y\})$$

Expanding this sum is tedious and results in nine trinomials. The following sum for the $\mathrm{Cov(X, Y)}$ expansion significantly reduces the algebra necessary by computing $\mathrm{E(XY)}$ instead:

$$\text{Cov(X, Y)} = \text{E(XY)} - \text{E(X)E(Y)}$$

$$\text{E(XY)} = \sum_{c \in \Omega} \text{X}(c)\text{Y}(c)\text{P}(\{c\})$$

$$= \sum_{x,y} xy\text{P}(\{X = x\} \cap \{Y = y\})$$

$$= (-1)(-1)(p - \frac{1}{16}) + (-1)(0)(\frac{1}{4} - p) + (-1)(1)(0)$$

$$+ (0)(-1)(\frac{1}{8}) + (0)(0)(\frac{3}{16}) + (0)(1)(\frac{1}{8})$$

$$+ (1)(-1)(p + \frac{1}{16}) + (1)(0)(\frac{1}{16}) + (1)(1)(\frac{1}{4} - p)$$

$$= (p - \frac{1}{16}) - (p + \frac{1}{16}) + (\frac{1}{4} - p)$$

$$\therefore \text{E(XY)} = -p + \frac{1}{8}$$

$$\implies \text{Cov(X, Y)} = (-p + \frac{1}{8}) - (-3p + \frac{1}{4})(\frac{3}{16})$$

$$= -p + \frac{1}{8} + \frac{9}{16}p - \frac{3}{64}$$

$$\therefore \text{Cov(X, Y)} = -\frac{7}{16}p - \frac{5}{64}$$

2   Question 2

Ω is continuous, which implies X is continuous and Y is continuous. Let D be the set of all $< x, y >$ that is inside (or on the boundary) of the triangle given.

2.1   Part a)

We are told that the joint pdf of X and Y is uniform over D, and assume it is 0 everywhere else. The area of the triangle D on a cartisian plane is $A = \frac{1}{2}bh = 1$, and we know

$$\int_{d \in D} f_{X,Y}(d)\mathrm{d}d = 1$$

Since $f_{X,Y}$ is uniform, this integral can be interpreted as the geometric volume of a triangular prism, extruded from D by $f_{X,Y}$

$$A \cdot f_{X,Y} = 1$$
$$f_{X,Y} = 1$$

Therefore the joint pdf of (X, Y) is $f_{X,Y} = 1$.

2.2   Part b)

The set of vectors D_2 containing all vectors $< x, y >$ satisfying $x > y$ in D forms a triangle on a cartisian plane with verticies at $(0, 0)$ $(0.5, 0.5)$ and $(1, 0)$. The area of this triangle is exactly $\frac{1}{4}$ the total area of D, since $A = \frac{1}{2}bh = \frac{1}{2} \cdot 1 \cdot 0.5 = \frac{1}{4}$. Therefore:

$$P(\{X \geq Y\}) = \int_{d \in D_2} f_{X,Y}(d)\mathrm{d}d$$
$$= f_{X,Y} \cdot A$$
$$= \frac{1}{4}$$
$$\therefore P(\{X \geq Y\}) = \frac{1}{4}$$

2.3   Part c)

Since

$$F_X = \int_{-\infty}^{x} f_X(x)\mathrm{d}x$$

$$f_X(x) = \frac{\mathrm{d}}{\mathrm{d}x}F_X(x)$$

Figure 1 illustrates the geometric cases involved with evaluating f_X:

We know

$$F_X(-1) = 0 \qquad\qquad\qquad F_X(x < -1) = 0$$
$$F_X(1) = 1 \qquad\qquad\qquad F_X(x > 1) = 1$$
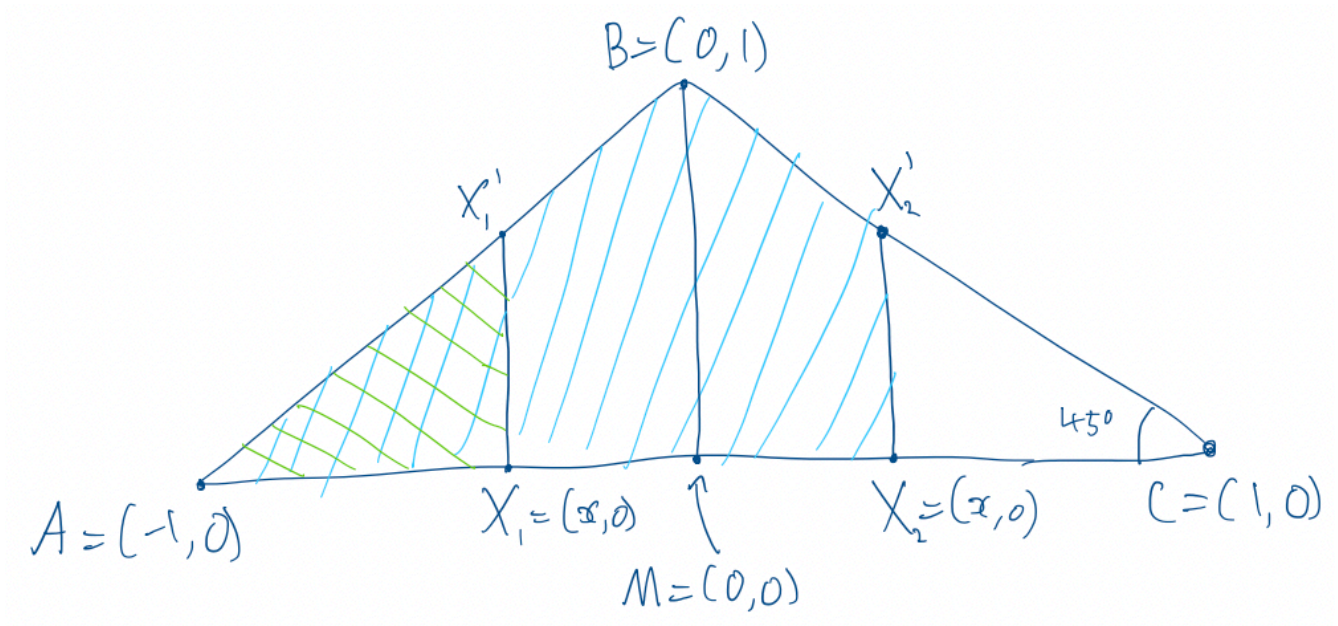
Figure 1: Geometric rendering of D showing the cases of F_X

Case 1: $-1 \leq x \leq 0$. This implies F_X is the area of $\triangle AX_1'X_1$. Let $AX_1 = x_1 = X_1X_1'$, which implies $x_1 = x + 1$:

$$\begin{aligned} \text{Area}(AX_1'X_1) &= \frac{1}{2}bh \\ &= \frac{1}{2}x_1^2 \end{aligned}$$

Therefore

$$F_X(-1 \leq x \leq 0) = \frac{1}{2}(x+1)^2$$

Case 2: $0 < x \leq 1$. This implies $F_X(x)$ is the area of $ABX_2'X_2$. Note $x = MX_2$, and that $X_2X_2' = X_2C$

$$\text{Area}(\text{ABX}'_2\text{X}_2) = \text{Area}(\triangle \text{ABM}) + \text{Area}(\text{MBX}'_2\text{X}_2)$$

$$\text{Area}(\triangle ABM) = \frac{1}{2}$$

$$\text{Area}(MBX'_2X_2) = \text{Area}(MBC) - \text{Area}(X_2X'_2C)$$

$$\text{Area}(MBC) = \text{Area}(ABM) = \frac{1}{2}$$

$$\text{Area}(X_2X'_2C) = \frac{1}{2}bh$$

$$= \frac{1}{2} \cdot (1-x) \cdot (1-x)$$

$$= \frac{1}{2}(1-x)^2$$

This is enough information to express F_X:

$$\implies F_\text{X}(0 < x \le 1) = (\frac{1}{2}) + (\frac{1}{2} - \frac{1}{2}(1-x)^2)$$

$$= 1 - \frac{1}{2}(1 - 2x + x^2)$$

$$= 1 - \frac{1}{2} + x - \frac{1}{2}x^2$$

$$= -\frac{1}{2}x^2 + x + \frac{1}{2}$$

Combining cases:

$$\implies F_\text{X} = \begin{cases} 0 & : x < -1 \\ \frac{1}{2}(1+x)^2 & : -1 \le x \le 0 \\ -\frac{1}{2}x^2 + x + \frac{1}{2} & : 0 < x \le 1 \\ 1 & : x > 1 \end{cases}$$

$$\therefore f_\text{X}(x) = \begin{cases} x + 1 & : -1 \le x \le 0 \\ -x + 1 & : 0 < x \le 1 \\ 0 & : \text{otherwise} \end{cases}$$

This so happens to be the geometric shape of D on a cartesian plane, ABC.

$$F_\text{Y}(y \le 0) = 0$$

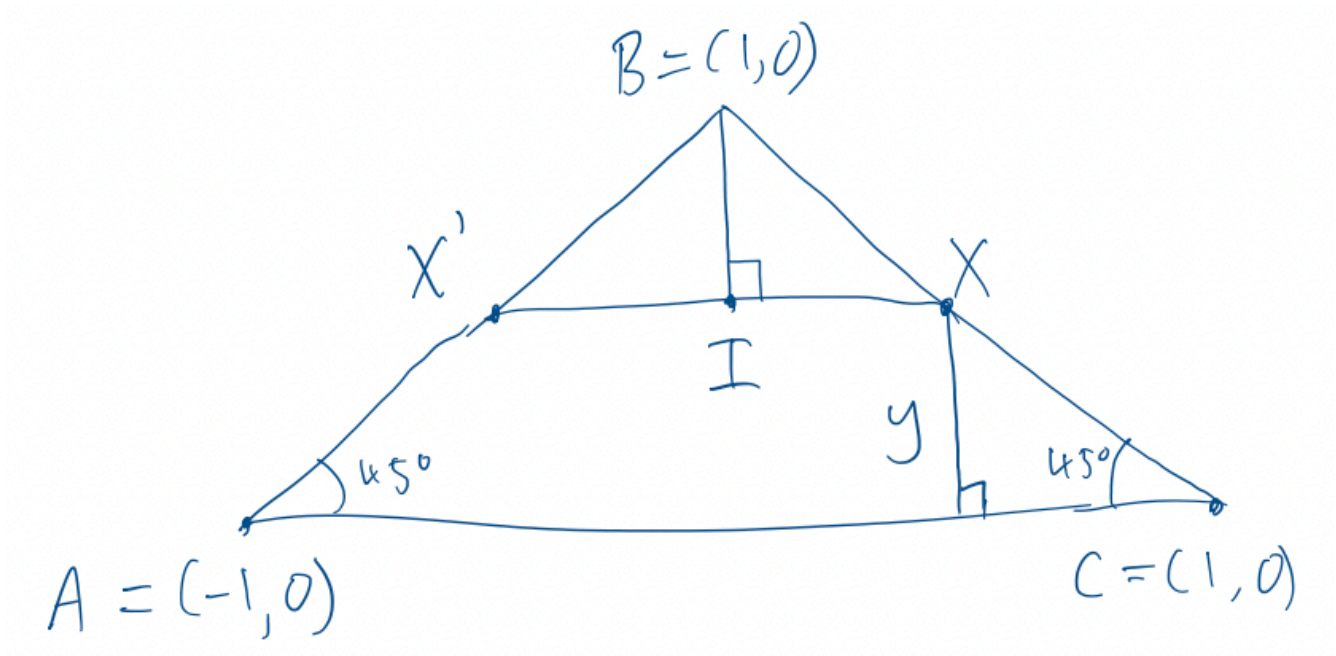$$F_\text{Y}(y \ge 1) = 1$$

Case $0 < y < 1$

Figure 2: Geometric rendering of D showing the cases of F_Y

$$
\begin{aligned}
F_Y(y) &= \text{Area}(AX'XC) \\
&= \text{Area}(\triangle ABC) - \text{Area}(\triangle XBX') \\
&= 1 - 2 \cdot \text{Area}(\triangle XIB)
\end{aligned}
$$

Note $IB = 1 - y = IX$

$$
\begin{aligned}
\text{Area}(\triangle XIB) &= \frac{1}{2}bh \\
&= \frac{1}{2}(1-y)^2 \\
&= \frac{1}{2}(y^2 - 2y + 1) \implies F_Y \qquad\qquad = 1 - (y^2 - 2y + 1) \\
&= -y^2 + 2y
\end{aligned}
$$

Combining cases

$$
F_Y(y) = \begin{cases} 1 & : y \geq 1 \\ -y^2 + 2y & : 0 < y < 1 \\ 0 & : y < 0 \end{cases}
$$

$$
f_Y(y) = \begin{cases} -2y + 2 & : 0 < y < 1 \\ 0 & : \text{otherwise} \end{cases}
$$

The geometric interpretation of this is not as intuitive ro realise. Morph A to $(0, 2)$, B to $(0, 1)$ and C to $(0, 0)$, and the initial value of $f_X(0) = AC$ is now placed on the y-axis.

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2.4   Part d)

The continuous independance rule can be stated like so:

$$f_{X,Y}(<x,y>) = f_X(x) \cdot f_Y(y) \quad \forall x, y \in \mathbb{R}$$

Suppose $<x, y> \in D$.

$$
\begin{aligned}
\text{LHS} &= f_{X,Y}(<x,y>) \\
&= 1\text{RHS} \qquad\qquad\qquad\qquad\qquad\qquad\qquad\qquad = f_Y(y) \cdot f_X(x) \\
&= (-2y+2) \cdot
\begin{cases}
x+1 & : -1 \le x \le 0 \\
-x+1 & ; 0 < x \le 1
\end{cases}
\end{aligned}
$$

Suppose further $<x, y> = <0, 0>$

$$
\begin{aligned}
\text{RHS} &= (-2 \cdot 0 + 2)(0 + 1) \\
&= 2
\end{aligned}
$$

Therefore there exists an $<x, y> \in D$ such that the independance rule fails, LHS $\neq$ RHS. Therefore X dn Y are not independant.

2.5   Part e)

It can be immediately noted that, due to the symmetry across the $x = 0$ "line" and the fact that $\text{Cov}(X, Y) = \int_{x,y} xy P(\{X = x\} \cap \{Y = y\})$ is negative (or zero) for $x < 0$ and positive (or zero) for $x > 0$ demonstrates intuitively that $\text{Cov}(X, Y) = 0$.

Somehow this seems dissatisfying to me. $\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$ is another way to evaluate $\text{Cov}(X, Y)$. This method requires evaluating a double integral or equivalent.

$$
\begin{aligned}
E(X) &= \int_{-\infty}^{+\infty} x f_X(x) \, dx \\
&= \int_{-1}^{0} x(x+1) \, dx + \int_{0}^{1} x(-x+1) \, dx \\
&= \int_{-1}^{0} x^2 + x \, dx + \int_{0}^{1} -x^2 + x \, dx \\
&= \frac{1}{3}x^3 + \frac{1}{2}x^2 \Big|_{x=-1}^{x=0} + -\frac{1}{3}x^3 + \frac{1}{2}x^2 \Big|_{x=0}^{x=1} \\
&= (0 + 0) - (\frac{1}{3}(-1)^3 + \frac{1}{2}(-1)^2) + (-\frac{1}{3}(1)^3 + \frac{1}{2}(1)^2) - (0 + 0) \\
&= +\frac{1}{3} + \frac{1}{2} - \frac{1}{3} + \frac{1}{2} \\
&= 1
\end{aligned}
$$

$$E(Y) = \int_{-\infty}^{+\infty} y f_Y(y) \, dy$$

$$= \int_0^1 -2y + 2 \, dy$$

$$= -y^2 + 2y \big|_{y=0}^{y=1}$$

$$= (-(1)^2 + 2(1)) - (0 + 0)$$

$$= -1 + 2$$

$$= 1$$

$$E(XY) = \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f_X(x) f_Y(y) \, dx \, dy$$

$$= \int_{-\infty}^{+\infty} x \left(\int_0^1 y f_Y(y) \, dy \right) f_X(x) \, dx$$

$$= E(Y) \int_{-1}^{+1} x f_X(x) \, dx$$

$$= E(Y) \cdot E(X)$$

$$= 1$$

Therefore $\text{Cov}(X, Y) = 1 - 1 \cdot 1 = 0$

3 Question 3

The sample space Ω is parameterized by the number of orders in the month.

$$\Omega_n = \{(x_1, x_2, \ldots x_n) : \forall n \in \mathbb{Z} \text{ and } x_n \in \{1, 2, 3, 4\}\}$$

Let X_i be the random variable for the number of items in the i'th order. Notice X_i are all independant and identically distributed (iid).

$$X_i((x_1, \ldots x_i, \ldots x_n) \in \Omega_n) = x_i$$

From the question statement, for all $i \in \mathbb{Z}$:

$$P(\{X_i = 1\}) = 0.54$$
$$P(\{X_i = 2\}) = 0.22$$
$$P(\{X_i = 3\}) = 0.15$$
$$P(\{X_i = 4\}) = 0.09$$

Let T be the random variable for the total number of items:

$$T((x_1, \ldots x_n) \in \Omega_n) = \sum_{i=1}^n x_i$$

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3.1 Part a)

Note $n = 500$. The question is asking to evaluate $P(\{T < 900\}) \geq 0.95$, equivalently $P(\{T > 900\}) < 0.05$.

$$
\begin{aligned}
E(X_i) &= \sum_x x P(\{X_i = x\}) \\
&= 1 \cdot 0.54 + 2 \cdot 0.22 + 3 \cdot 0.15 + 4 \cdot 0.09 \\
&= 1.79 \\
Var(X_i) &= E[(X_i - E(X_i)^2] \\
&= \sum_{c \in \Omega} (X_i(c) - 1.79)^2 P(\{c\}) \\
&= \sum_x (x - 1.79)^2 P(\{X_i = x\}) \\
&= (1 - 1.79)^2 \cdot 0.55 + (2 - 1.79)^2 \cdot 0.22 + (3 - 1.79)^2 \cdot 0.15 + (4 - 1.79)^2 \cdot 0.09 \\
&= 1.0059 \\
E(T) &= \sum_{i=1}^{500} E(X_i) \\
&= 500 \cdot 1.79 \\
&= 895 \\
\implies \mu_T &= 895 \\
Var(T) &= \sum_{i=1}^{500} Var(X_i) \\
&= 500 \cdot 1.0059 \\
&= 502.95 \\
\implies \sigma_T &= \sqrt{502.95} \approx 22.4265
\end{aligned}
$$

From the Central Limit Theorem (CLT), we can assume T follow a normal distribution. $n = 500$ is a reasonable large n for this to be a good approximation.

$$
T \sim \mathcal{N}(\mu = 895, \sigma = 502.95)
$$

We will ignore the inequality until the end:

$$
\begin{aligned}
P(\{T > 900\}) &= P(\{\frac{T - \mu}{\sigma} > \frac{900 - 895}{22.4265}\}) \\
&= P(\{Z > 0.2230\}) \\
&= 1 - P(\{Z < 0.2230\}) \text{Using stats table} \qquad = 1 - 0.5871 \\
&= 0.4129
\end{aligned}
$$

$\therefore$ No, there is approximatelly a 41% change of exceeding the 900-item limit, which is significantly higher than the threshold 5%, therefore the company cannot process $n = 500$ orders.