

# 1 Question 1

## 1.1 Part a)

Let  $\Omega$  be the sample space. Therefore  $P(\{\Omega\}) = 1$ . Adding all the joint pmf values must sum to 1:

$$\begin{aligned}\{\Omega\} &= \bigcup_x \bigcup_y \{X = x\} \cap \{Y = y\} \\ P(\{\Omega\}) &= 1 \\ \implies 1 &= P((\{X = -1\} \cap \{Y = -1\}) \cup \dots \cup (\{X = 1\} \cap \{Y = 1\})) \\ &= P(\{X = -1\} \cap \{Y = -1\}) + \dots + P(\{X = 1\} \cap \{Y = 1\}) \\ &= (p - \frac{1}{16}) + (\frac{1}{4} - p) + (0) + (\frac{1}{8}) + (\frac{3}{16}) + (\frac{1}{8}) + (p + \frac{1}{16}) + (\frac{1}{16}) + (\frac{1}{4} - p) \\ 1 &= -\frac{1}{16} + \frac{4}{16} + \frac{7}{16} + \frac{1}{16} + \frac{1}{16} + \frac{4}{16} \\ 1 &= 1\end{aligned}$$

Unfortunately, this tells us no information about  $p$ . From the definition of probability,  $P(\{c\})$  for  $c \in \Omega$  must be greater or equal to 0,  $P(\{c \in \Omega\}) \geq 0$ . This can be used to restrict the possible values of  $p$ :

$$\begin{aligned}P(A \subseteq \Omega) &\geq 0 \\ \implies P(\{X = -1\} \cap \{Y = -1\}) &\geq 0 \\ p - \frac{1}{16} &\geq 0 \\ p &\geq \frac{1}{16} \\ \implies P(\{X = 0\} \cap \{Y = -1\}) &\geq 0 \\ \frac{1}{4} - p &\geq 0 \\ p &\leq \frac{1}{4} \\ \implies P(\{X = -1\} \cap \{Y = 1\}) &\geq 0 \\ p + \frac{1}{16} &\geq 0 \\ p &\leq \frac{1}{16} \\ \implies p &\in [\frac{1}{16}, \frac{1}{4}] \tag{1}\end{aligned}$$

Therefore,  $\frac{1}{16} \leq p \leq \frac{1}{4}$ , and can be any value within this range.

## 1.2 Part b)

Aim is to find  $P(\{X = Y\})$ :

$$\begin{aligned}
P(\{X = Y\}) &= \sum_a P(\{X = a\} \cap \{Y = a\}) \\
&= P(\{X = -1\} \cap \{Y = -1\}) + P(\{X = 0\} \cap \{Y = 0\}) + P(\{X = 1\} \cap \{Y = 1\}) \\
&= (p - \frac{1}{16}) + (\frac{3}{16}) + (\frac{1}{4} - p) \\
&= \frac{6}{16} = \frac{3}{8}
\end{aligned}$$

### 1.3 Part c)

The marginal pdf of X is  $f_X(x)$ , which is equal to  $P(\{X = x\})$  and can be manually evaluated:

$$\begin{aligned}
P(\{X = -1\}) &= \sum_y P(\{X = -1\} \cap \{Y = y\}) \\
&= (p - \frac{1}{16}) + (\frac{1}{8}) + (p + \frac{1}{16}) \\
&= 2p + \frac{1}{8} \\
P(\{X = 0\}) &= \sum_y P(\{X = 0\} \cap \{Y = y\}) \\
&= (\frac{1}{4} - p) + (\frac{3}{16}) + (\frac{1}{16}) \\
&= -p + \frac{1}{2} \\
P(\{X = 1\}) &= \sum_y P(\{X = 1\} \cap \{Y = y\}) \\
&= (0) + (\frac{1}{8}) + (\frac{1}{4} - p) \\
&= -p + \frac{3}{8} \\
\Rightarrow f_X(x) = P(\{X = x\}) &= \begin{cases} 2p + \frac{1}{8} & x = -1 \\ -p + \frac{1}{2} & x = 0 \\ -p + \frac{3}{8} & x = 1 \\ 0 & \text{otherwise} \end{cases}
\end{aligned}$$

$$\begin{aligned}
P(\{Y = -1\}) &= \sum_x P(\{X = x\} \cap \{Y = -1\}) \\
&= (p - \frac{1}{16}) + (\frac{1}{4} - p) + (0) \\
&= \frac{3}{16} \\
P(\{Y = 0\}) &= \sum_x P(\{X = x\} \cap \{Y = 0\}) \\
&= (\frac{1}{8}) + (\frac{3}{16}) + (\frac{1}{8}) \\
&= \frac{7}{16} \\
P(\{Y = 1\}) &= \sum_x P(\{X = x\} \cap \{Y = 1\}) \\
&= (p + \frac{1}{16}) + (\frac{1}{16}) + (\frac{1}{4} - p) \\
&= \frac{6}{18} = \frac{3}{8} \\
\Rightarrow f_Y(x) = P(\{Y = y\}) &= \begin{cases} \frac{3}{16} & y = -1 \\ \frac{7}{16} & y = 0 \\ \frac{3}{8} & y = 1 \\ 0 & \text{otherwise} \end{cases}
\end{aligned}$$

## 1.4 Part d)

X and Y are independant if

$$P(\{X = x\} \cap \{Y = y\}) = P(\{X = x\}) \cdot P(\{Y = y\}) = f_X(x) \cdot f_Y(y) \quad (2)$$

for all possible values  $x$  and  $y$ . Therefore, this must be true for  $x = -1$  and  $y = 1$ :

$$\begin{aligned}
\text{LHS} &= P(\{X = -1\} \cap \{Y = 1\}) \\
&= p + \frac{1}{16} \\
\text{RHS} &= P(\{X = -1\})P(\{Y = 1\}) \\
&= (2p + \frac{1}{8})(\frac{3}{8}) \\
&= \frac{3}{4}p + \frac{3}{64}
\end{aligned}$$

As shown above, LHS and RHS are only equal for zero or one values of  $p$ . Letting LHS = RHS, we can find this exact value (or lack thereof):

$$\begin{aligned}
p + \frac{1}{16} &= \frac{3}{4}p + \frac{3}{64} \\
\frac{1}{4}p &= \frac{3}{64} - \frac{1}{16} \\
p &= -\frac{1}{64} \cdot 4 = -\frac{1}{16}
\end{aligned}$$

Therefore  $\text{LHS} = \text{RHS}$  only when  $p = -\frac{1}{16}$ , however from (1) this is not within the potential domain of  $p$ . Therefore  $\text{LHS} \neq \text{RHS}$ , showing one counterexample to (2), hence  $X$  and  $Y$  are not independent.

## 1.5 Part e)

$$\begin{aligned}
E(X) &= \sum_x xP(\{X = x\}) \\
&= -1(2p + \frac{1}{8}) + 0(-p + \frac{1}{2}) + 1(-p + \frac{3}{8}) \\
&= -2p - \frac{1}{8} - p + \frac{3}{8} \\
\therefore E(X) &= -3p + \frac{1}{4} \\
E(Y) &= \sum_y yP(\{Y = y\}) \\
&= -1(\frac{3}{16}) + 0(\frac{7}{16}) + 1(\frac{3}{8}) \\
\therefore E(Y) &= -\frac{3}{16} + \frac{3}{8} = \frac{3}{16} \\
\text{Cov}(X, Y) &= E[(X - E(X))(Y - E(Y))] \\
\text{Cov}(X, Y) &= \sum_{c \in \Omega} (X(c) - E(X))(Y(c) - E(Y))P(\{c\}) \\
&= \sum_{x,y} (x - (-3p + \frac{1}{4}))(y - (\frac{3}{16}))P(\{X = x\} \cap \{Y = y\})
\end{aligned}$$

Expanding this sum is tedious and results in nine trinomials. The following sum for the  $\text{Cov}(X, Y)$  expansion significantly reduces the algebra necessary by computing  $E(XY)$  instead:

$$\begin{aligned}
\text{Cov}(X, Y) &= E(XY) - E(X)E(Y) \\
E(XY) &= \sum_{c \in \Omega} X(c)Y(c)P(\{c\}) \\
&= \sum_{x,y} xyP(\{X = x\} \cap \{Y = y\}) \\
&= (-1)(-1)(p - \frac{1}{16}) + (-1)(0)(\frac{1}{4} - p) + (-1)(1)(0) \\
&\quad + (0)(-1)(\frac{1}{8}) + (0)(0)(\frac{3}{16}) + (0)(1)(\frac{1}{8}) \\
&\quad + (1)(-1)(p + \frac{1}{16}) + (1)(0)(\frac{1}{16}) + (1)(1)(\frac{1}{4} - p) \\
&= (p - \frac{1}{16}) - (p + \frac{1}{16}) + (\frac{1}{4} - p) \\
&\therefore E(XY) = -p + \frac{1}{8} \\
\implies \text{Cov}(X, Y) &= (-p + \frac{1}{8}) - (-3p + \frac{1}{4})(\frac{3}{16}) \\
&= -p + \frac{1}{8} + \frac{9}{16}p - \frac{3}{64} \\
&\therefore \text{Cov}(X, Y) = -\frac{7}{16}p - \frac{5}{64}
\end{aligned}$$

## 2 Question 2

$\Omega$  is continuous, which implies  $X$  is continuous and  $Y$  is continuous. Let  $D$  be the set of all  $\langle x, y \rangle$  that is inside (or on the boundary) of the triangle given.

### 2.1 Part a)

We are told that the joint pdf of  $X$  and  $Y$  is uniform over  $D$ , and assume it is 0 everywhere else. The area of the triangle  $D$  on a cartesian plane is  $A = \frac{1}{2}bh = 1$ , and we know

$$\int_{d \in D} f_{X,Y}(d) dd = 1$$

Since  $f_{X,Y}$  is uniform, this integral can be interpreted as the geometric volume of a triangular prism, extruded from  $D$  by  $f_{X,Y}$

$$\begin{aligned} A \cdot f_{X,Y} &= 1 \\ f_{X,Y} &= 1 \end{aligned}$$

Therefore the joint pdf of  $(X, Y)$  is  $f_{X,Y} = 1$ .

### 2.2 Part b)

The set of vectors  $D_2$  containing all vectors  $\langle x, y \rangle$  satisfying  $x > y$  in  $D$  forms a triangle on a cartesian plane with vertices at  $(0, 0)$ ,  $(0.5, 0.5)$  and  $(1, 0)$ . The area of this triangle is exactly  $\frac{1}{4}$  the total area of  $D$ , since  $A = \frac{1}{2}bh = \frac{1}{2} \cdot 1 \cdot 0.5 = \frac{1}{4}$ . Therefore:

$$\begin{aligned} P(\{X \geq Y\}) &= \int_{d \in D_2} f_{X,Y}(d) dd \\ &= f_{X,Y} \cdot A \\ &= \frac{1}{4} \\ \therefore P(\{X \geq Y\}) &= \frac{1}{4} \end{aligned}$$

### 2.3 Part c)

Since

$$\begin{aligned} F_X &= \int_{-\infty}^x f_X(x) dx \\ f_X(x) &= \frac{d}{dx} F_X(x) \end{aligned}$$

Figure 1 illustrates the geometric cases involved with evaluating  $f_X$ :

We know

$$\begin{aligned} F_X(-1) &= 0 \\ F_X(1) &= 1 \end{aligned}$$

$$\begin{aligned} F_X(x < -1) &= 0 \\ F_X(x > 1) &= 1 \end{aligned}$$

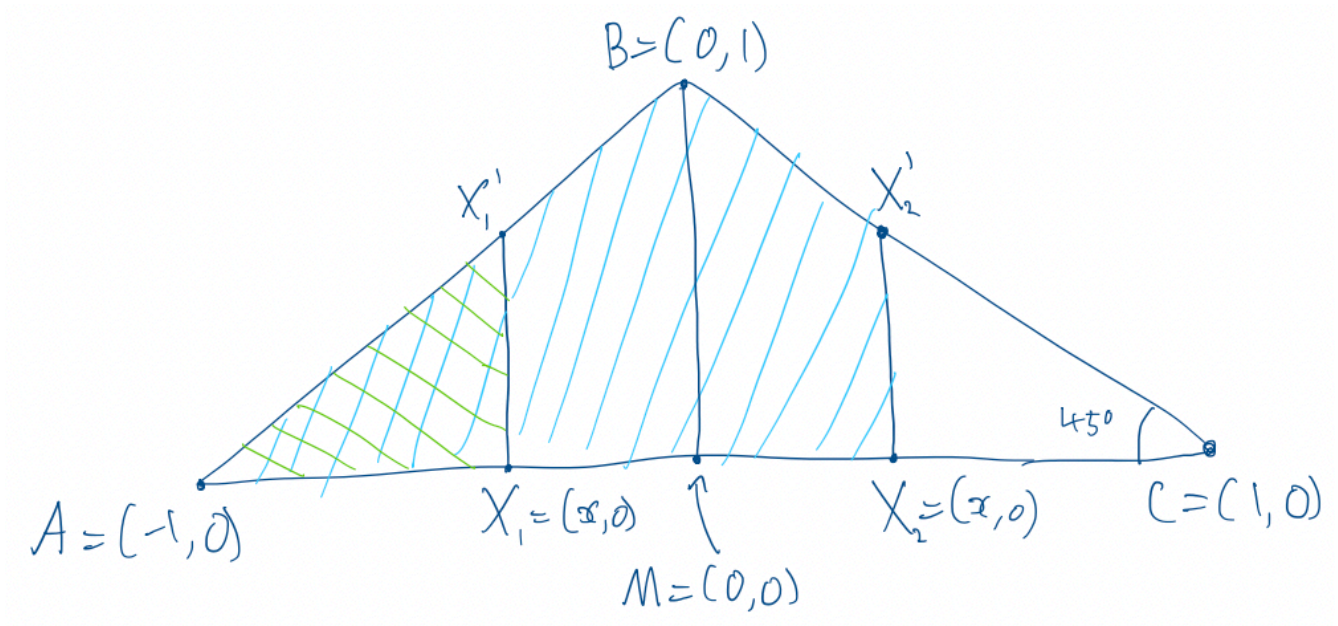


Figure 1: Geometric rendering of  $D$  showing the cases of  $F_X$

Case 1:  $-1 \leq x \leq 0$ . This implies  $F_X$  is the area of  $\triangle AX_1'X_1$ . Let  $AX_1 = x_1 = X_1X_1'$ , which implies  $x_1 = x + 1$ :

$$\begin{aligned} \text{Area}(AX_1'X_1) &= \frac{1}{2}bh \\ &= \frac{1}{2}x_1^2 \end{aligned}$$

Therefore

$$F_X(-1 \leq x \leq 0) = \frac{1}{2}(x+1)^2$$

Case 2:  $0 < x \leq 1$ . This implies  $F_X(x)$  is the area of  $ABX_2'X_2$ . Note  $x = MX_2$ , and that  $X_2X_2' = X_2C$

$$\begin{aligned}
\text{Area}(\text{ABX}_2'\text{X}_2) &= \text{Area}(\triangle \text{ABM}) + \text{Area}(\text{MBX}_2'\text{X}_2) \\
\text{Area}(\triangle \text{ABM}) &= \frac{1}{2} \\
\text{Area}(\text{MBX}_2'\text{X}_2) &= \text{Area}(\text{MBC}) - \text{Area}(\text{X}_2\text{X}_2'\text{C}) \\
\text{Area}(\text{MBC}) &= \text{Area}(\text{ABM}) = \frac{1}{2} \\
\text{Area}(\text{X}_2\text{X}_2'\text{C}) &= \frac{1}{2}bh \\
&= \frac{1}{2} \cdot (1-x) \cdot (1-x) \\
&= \frac{1}{2}(1-x)^2
\end{aligned}$$

This is enough information to express  $F_X$ :

$$\begin{aligned}
\Rightarrow F_X(0 < x \leq 1) &= \left(\frac{1}{2}\right) + \left(\frac{1}{2} - \frac{1}{2}(1-x)^2\right) \\
&= 1 - \frac{1}{2}(1-2x+x^2) \\
&= 1 - \frac{1}{2} + x - \frac{1}{2}x^2 \\
&= -\frac{1}{2}x^2 + x + \frac{1}{2}
\end{aligned}$$

Combining cases:

$$\begin{aligned}
\Rightarrow F_X &= \begin{cases} 0 & : x < -1 \\ \frac{1}{2}(1+x)^2 & : -1 \leq x \leq 0 \\ -\frac{1}{2}x^2 + x + \frac{1}{2} & : 0 < x \leq 1 \\ 1 & : x > 1 \end{cases} \\
\therefore f_X(x) &= \begin{cases} x+1 & : -1 \leq x \leq 0 \\ -x+1 & : 0 < x \leq 1 \\ 0 & : \text{otherwise} \end{cases}
\end{aligned}$$

This so happens to be the geometric shape of  $D$  on a cartesian plane, ABC.

$$F_Y(y \leq 0) = 0$$

$$F_Y(y \geq 1) = 1$$

Case  $0 < y < 1$



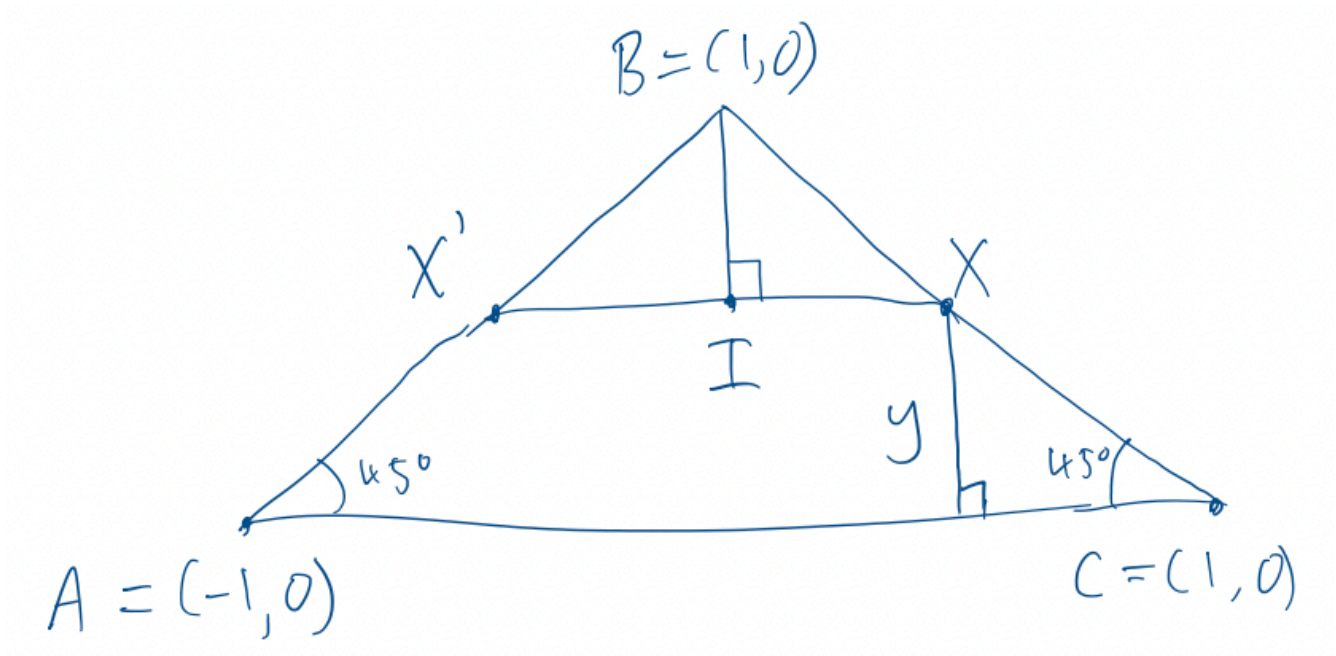


Figure 2: Geometric rendering of  $D$  showing the cases of  $F_Y$

$$\begin{aligned}
 F_Y(y) &= \text{Area}(AX'XC) \\
 &= \text{Area}(\triangle ABC) - \text{Area}(\triangle XBX') \\
 &= 1 - 2 \cdot \text{Area}(\triangle XIB)
 \end{aligned}$$

Note  $IB = 1 - y = IX$

$$\begin{aligned}
 \text{Area}(\triangle XIB) &= \frac{1}{2}bh \\
 &= \frac{1}{2}(1 - y)^2 \\
 &= \frac{1}{2}(y^2 - 2y + 1) \implies F_Y &= 1 - (y^2 - 2y + 1) \\
 &= -y^2 + 2y
 \end{aligned}$$

Combining cases

$$\begin{aligned}
 F_Y(y) &= \begin{cases} 1 & : y \geq 1 \\ -y^2 + 2y & : 0 < y < 1 \\ 0 & : y < 0 \end{cases} \\
 f_Y(y) &= \begin{cases} -2y + 2 & : 0 < y < 1 \\ 0 & : \text{otherwise} \end{cases}
 \end{aligned}$$

The geometric interpretation of this is not as intuitive to realise. Morph A to  $(0, 2)$ , B to  $(0, 1)$  and C to  $(0, 0)$ , and the initial value of  $f_X(0) = AC$  is now placed on the y-axis.

## 2.4 Part d)

The continuous independence rule can be stated like so:

$$f_{X,Y}(< x, y >) = f_X(x) \cdot f_Y(y) \quad \forall x, y \in \mathbb{R}$$

Suppose  $< x, y > \in D$ .

$$\begin{aligned} \text{LHS} &= f_{X,Y}(< x, y >) \\ &= 1\text{RHS} &= f_Y(y) \cdot f_X(x) \\ &= (-2y + 2) \cdot \begin{cases} x + 1 & : -1 \leq x \leq 0 \\ -x + 1 & ; 0 < x \leq 1 \end{cases} \end{aligned}$$

Suppose further  $< x, y > = < 0, 0 >$

$$\begin{aligned} \text{RHS} &= (-2 \cdot 0 + 2)(0 + 1) \\ &= 2 \end{aligned}$$

Therefore there exists an  $< x, y > \in D$  such that the independence rule fails,  $\text{LHS} \neq \text{RHS}$ . Therefore  $X$  and  $Y$  are not independent.

## 2.5 Part e)

It can be immediately noted that, due to the symmetry across the  $x = 0$  "line" and the fact that  $\text{Cov}(X, Y) = \int_{x,y} xyP(\{X = x\} \cap \{Y = y\})$  is negative (or zero) for  $x < 0$  and positive (or zero) for  $x > 0$  demonstrates intuitively that  $\text{Cov}(X, Y) = 0$ .

Somehow this seems dissatisfying to me.  $\text{Cov}(X, Y) = E(XY) - E(X)E(Y)$  is another way to evaluate  $\text{Cov}(X, Y)$ . This method requires evaluating a double integral or equivalent.

$$\begin{aligned} E(X) &= \int_{-\infty}^{+\infty} x f_X(x) dx \\ &= \int_{-1}^0 x(x+1) dx + \int_0^1 x(-x+1) dx \\ &= \int_{-1}^0 x^2 + x dx + \int_0^1 -x^2 + x dx \\ &= \left. \frac{1}{3}x^3 + \frac{1}{2}x^2 \right|_{x=-1}^{x=0} + \left. -\frac{1}{3}x^3 + \frac{1}{2}x^2 \right|_{x=0}^{x=1} \\ &= (0+0) - \left( \frac{1}{3}(-1)^3 + \frac{1}{2}(-1)^2 \right) + \left( -\frac{1}{3}(1)^3 + \frac{1}{2}(1)^2 \right) - (0+0) \\ &= +\frac{1}{3} + \frac{1}{2} - \frac{1}{3} + \frac{1}{2} \\ &= 1 \end{aligned}$$

$$\begin{aligned}
E(Y) &= \int_{-\infty}^{+\infty} y f_Y(y) dy \\
&= \int_0^1 -2y + 2 dy \\
&= -y^2 + 2y \Big|_{y=0}^{y=1} \\
&= (-(1)^2 + 2(1)) - (0 + 0) \\
&= -1 + 2 \\
&= 1
\end{aligned}$$

$$\begin{aligned}
E(XY) &= \int_{-\infty}^{+\infty} \int_{-\infty}^{+\infty} xy f_X(x) f_Y(y) dx dy \\
&= \int_{-\infty}^{+\infty} x \left( \int_0^1 y f_Y(y) dy \right) f_X(x) dx \\
&= E(Y) \int_{-1}^{+1} x f_X(x) dx \\
&= E(Y) \cdot E(X) \\
&= 1
\end{aligned}$$

Therefore  $\text{Cov}(X, Y) = 1 - 1 \cdot 1 = 0$

### 3 Question 3

The sample space  $\Omega$  is parameterized by the number of orders in the month.

$$\Omega_n = \{(x_1, x_2, \dots, x_n) : \forall n \in \mathbb{Z} \text{ and } x_n \in \{1, 2, 3, 4\}\}$$

Let  $X_i$  be the random variable for the number of items in the  $i$ 'th order.

$$X_i((x_1, \dots, x_i, \dots, x_n) \in \Omega_n) = x_i$$

From the question statement, for all  $i \in \mathbb{Z}$ :

$$P(\{X_i = 1\}) = 0.54$$

$$P(\{X_i = 2\}) = 0.22$$

$$P(\{X_i = 3\}) = 0.15$$

$$P(\{X_i = 4\}) = 0.09$$

Let  $T$  be the random variable for the total number of items:

$$T((x_1, \dots, x_n) \in \Omega_n) = \sum_{i=1}^n x_i$$

### 3.1 Part a)

Note  $n = 500$ . The question is asking to evaluate  $P(\{T > 900\})$   $P(\{T > 900\}) < 0.05$