# Dynamic Programming and Structural Estimation

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## Outline

Introduction

# Stochastic Dynamic Programming

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  - Grids
- Interpolation
- Today:
  - Uncertainty:
    - Future income is uncertain
    - + Another state variable: Permanent income
    - + "Normalization" of one state variable.

## Stochastic Dynamic Programming

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- Estimation:

Simulated Method of Moments (SMM/SMD) Relate to "reduced-form" Can we combine approaches?

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Simulated Method of Moments (SMM/SMD)

Relate to "reduced-form"

Can we combine approaches?

- Example: Buffer Stock model of Deaton (1991); Carroll (1992)
  - Estimated in Gourinchas and Parker (2002)

# Gourinchas and Parker (2002)

Introduction 000

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Structural Estimation

- Approach:
  - 1. **Estimate model** with 2+ motives:

Buffer-stock motive: Income risk while working.

Life cycle motive: Consumption in retirement.

# Gourinchas and Parker (2002)

Research question: "Which savings motives dominate across life?"

- Approach:
  - 1. **Estimate model** with 2+ motives: Buffer-stock motive: Income risk while working. Life cycle motive: Consumption in retirement.
  - 2. Quantify importance of these motives over life Counterfactual simulations

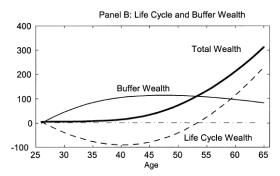


FIGURE 7.—The role of risk in saving and wealth accumulation.

## Outline

Stochastic DP

# Buffer-stock model (Deaton-Carroll) Bellman equation

• Simplest version of the buffer-stock model is

$$V_{t}(M_{t}, P_{t}) = \max_{C_{t}} \frac{C_{t}^{1-\rho}}{1-\rho} + \beta \mathbb{E}_{t} \left[ V_{t+1}(M_{t+1}, P_{t+1}) \right]$$
s.t.
$$A_{t} = M_{t} - C_{t}$$

$$M_{t+1} = RA_{t} + Y_{t+1}$$

$$Y_{t+1} = \xi_{t+1}P_{t+1}$$

$$P_{t+1} = GP_{t}\psi_{t+1}$$

$$A_{t} \geq 0, \forall t$$

Structural Estimation

• where  $\mathbb{E}_{t}[V_{t+1}(M_{t+1}, P_{t+1})] = \mathbb{E}[V_{t+1}(M_{t+1}, P_{t+1})|M_{t}, P_{t}, C_{t}]$ 

• Last period: Everything is consumed,

$$C_T^{\star}(M_T, P_T) = M_T$$

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• Gourinchas and Parker (2002): Retirement-periods Assumes a linear post-retirement value (w.  $P_{T+1} = P_T$ )

$$V_{T+1}(M_{T+1}, P_{T+1}) = \kappa \cdot (M_{T+1} + h \cdot P_{T+1})$$

Motivated by a deterministic perfect credit market solution (estimate  $\kappa$  and h, through  $\gamma_0$  and  $\gamma_1$ )

• They also allow for time-varying taste-shifters,  $v_t(Z_t)$ .

• **Defining**  $c_t \equiv C_t/P_t$ ,  $m_t \equiv M_t/P_t$  etc. implies

$$A_t = M_t - C_t$$

$$A_t/P_t = M_t/P_t - C_t/P_t$$

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### Normalization I

• **Defining**  $c_t \equiv C_t/P_t$ ,  $m_t \equiv M_t/P_t$  etc. implies

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and state transition

$$M_{t+1} = RA_t + Y_{t+1}$$
 $M_{t+1}/P_{t+1} = RA_t/P_{t+1} + Y_{t+1}/P_{t+1}$ 
 $m_{t+1} = Ra_tP_t/P_{t+1} + \xi_{t+1}$ 
 $m_{t+1} = \frac{R}{G\psi_{t+1}}a_t + \xi_{t+1}$ 

The **adjustment factor**  $\frac{1}{G\psi_{t+1}}$  is due to changes in permanent income

• **Defining**  $v_t(m_t) = V_t(M_t, P_t) / P_t^{1-\rho}$  implies

$$\begin{split} V_t(M_t, P_t) &= \max_{C_t} \frac{C_t^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ V_{t+1}(M_{t+1}, P_{t+1}) \right] \\ &= \max_{c_t} \frac{(c_t P_t)^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ V_{t+1}(M_{t+1}, P_{t+1}) \right] \Leftrightarrow \\ V_t(M_t, P_t) / P_t^{1-\rho} &= \max_{c_t} \frac{(c_t P_t)^{1-\rho} / P_t^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ V_{t+1}(M_{t+1}, P_{t+1}) / P_t^{1-\rho} \right] \Leftrightarrow \\ v_t(m_t) &= \max_{c_t} \frac{c_t^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ \underbrace{V_{t+1}(M_{t+1}, P_{t+1}) / P_{t+1}^{1-\rho}}_{t-1} \cdot \underbrace{P_{t+1}^{1-\rho} / P_t^{1-\rho}}_{=(G\psi_{t+1})^{1-\rho}} \right] \\ &= \max_{c_t} \frac{c_t^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ (G\psi_{t+1})^{1-\rho} v_{t+1}(m_{t+1}) \right] \end{split}$$

$$v_{t}(m_{t}) = \max_{c_{t}} \frac{c_{t}^{1-\rho}}{1-\rho} + \beta \mathbb{E}_{t} \left[ (G\psi_{t+1})^{1-\rho} v_{t+1}(m_{t+1}) \right]$$
s.t.
$$a_{t} = m_{t} - c_{t}$$

$$m_{t+1} = \frac{1}{G\psi_{t+1}} Ra_{t} + \xi_{t+1}$$

$$a_{t} \geq 0$$

• **Benefit:** Dimensionality of state space reduced,  $2 \rightarrow 1$ . Can this always be done?

## Bellman equation in ratio form

$$v_{t}(m_{t}) = \max_{c_{t}} \frac{c_{t}^{1-\rho}}{1-\rho} + \beta \mathbb{E}_{t} \left[ (G\psi_{t+1})^{1-\rho} v_{t+1}(m_{t+1}) \right]$$
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Structural Estimation

- **Benefit:** Dimensionality of state space reduced,  $2 \rightarrow 1$ . Can this always be done?
- No... Uses that utility is homothetic (budget constraint also important)

$$V_T(M_T, P_T) = \frac{M_T^{1-\rho}}{1-\rho} = \frac{(m_T P_T)^{1-\rho}}{1-\rho} = \frac{m_T^{1-\rho}}{1-\rho} P_T^{1-\rho}$$

such that  $v_T(m_T) = V_T(M_T, P_T)/P_T^{1-\rho}$  holds!

## Solving the model: Numerical Integration

Solved by backwards induction

Terminal period:

$$v_T(m_T) = \frac{m_T^{1-\rho}}{1-\rho}$$

Structural Estimation

For t < T:

$$v_t(m_t) = \max_{c_t} \frac{c_t^{1-\rho}}{1-\rho} + \beta \mathbb{E}_t \left[ (G\psi_{t+1})^{1-\rho} v_{t+1}(m_{t+1}) \right]$$

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Structural Estimation

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• How to evaluate expectations?

$$\mathbb{E}_{t}\left[\bullet\right] = \int_{\psi_{t+1}} \int_{\xi_{t+1}} \left[ (G\psi_{t+1})^{1-\rho} v_{t+1}(m_{t+1}) \right] f(d\psi_{t+1}, d\xi_{t+1})$$

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• How to evaluate expectations?

$$\mathbb{E}_t\left[\bullet\right] = \int_{\psi_{t+1}} \int_{\xi_{t+1}} \left[ (G\psi_{t+1})^{1-\rho} \mathsf{v}_{t+1}(m_{t+1}) \right] f(d\psi_{t+1}, d\xi_{t+1})$$

Numerical Integration: Discretize into sum (Gauss-Hermite)

$$\mathbb{E}_{t}\left[\bullet\right] \approx \sum_{j=1}^{J} \sum_{k=1}^{K} \left[ \left( G \psi^{(j)} \right)^{1-\rho} v_{t+1}(m^{(j,k)}) \right] \omega_{j} \omega_{k}$$

and interpolate  $v_{t+1}(m^{(j,k)}) = rac{1}{G\psi^{(j)}} Ra_t + \xi^{(k)})$  for values of the  $\overrightarrow{m}$  grid.

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## Outline

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Structural Estimation

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- Alternative: Estimate reduced-form equations "derived" from model
- My (and others) claim: To turn reduced form parameter estimates into policy advice requires a lot of assumptions
  - "All econometric work relies heavily on a priori assumptions. The main difference between structural and experimental (or "atheoretic") approaches is not in the number of assumptions but the extent to which they are made explicit." (Keane, 2010)

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Structural Estimation

#### Benefit of models:

- 1. Ensure *consistent* world view
- 2. Better models are well defined: Clear how to progress.
- 3. Hopefully "deep" policy-invariant parameters (Lucas critique).

- Critique of structural estimation: Requires many assumptions
- Alternative: Estimate reduced-form equations "derived" from model
- My (and others) claim: To turn reduced form parameter estimates into policy advice requires a lot of assumptions

"All econometric work relies heavily on a priori assumptions. The main difference between structural and experimental (or "atheoretic") approaches is not in the number of assumptions but the extent to which they are made explicit." (Keane, 2010)

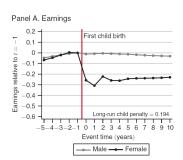
Structural Estimation

#### Benefit of models:

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- **Frontier:** Use exogenous variation to estimate structural model.

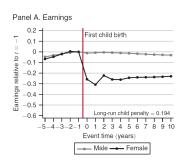
• Example: Event-studies (child-birth, Kleven, Landais and Søgaard, 2019)

- Reduced-form to be casual: "statistical" assumptions
  - No self-selection (timing)
  - No anticipation effects.
  - Parallel trends.



• Example: Event-studies (child-birth, Kleven, Landais and Søgaard, 2019)

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  - No self-selection (timing)
  - No anticipation effects.
  - Parallel trends.



- A model can allow for these assumptions to be violated But only through the chosen functional forms and mechanisms
  - "Economic" assumptions
  - Easier to debate and approve upon (?)

We know how to solve dynamic programming models

Structural Estimation

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- How can we estimate them? We need
  - 1. Data on (some) *states*
  - 2. Data on (some) choices

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Structural Estimation

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- Simulated versions:
  - Maximum Simulated Likelihood (MSL, SML)
  - 2. Method of Simulated Moments (MSM, SMM)

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- Simulated versions:
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- Example model: Life-cycle buffer-stock model
  - States:  $M_{it}$ ,  $P_{it}$
  - Choice: Cit
- **Parameters** to estimate:  $\theta = \{\beta, \rho\}$ 
  - Calibration: G,  $\sigma_{\psi}$ ,  $\sigma_{\xi}$ , R, and  $\lambda$  ("known")

# Method of Simulated Moments (MSM)

•  $\Lambda^d = \frac{1}{N} \sum_{i=1}^{N} \Lambda_i^d$  are some moments in the data Could be avg., var, cov, regression-coefs, etc.

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- $\Lambda^d = \frac{1}{N} \sum_{i=1}^N \Lambda_i^d$  are some moments in the data Could be avg., var, cov, regression-coefs, etc.
- $\Lambda^m(\theta)$  are the same moments calculated on data **simulated** from the model solved with parameter values  $\theta$
- MSM then is

$$\hat{\theta} = \arg\min_{\theta} \left( \Lambda^d - \Lambda^m(\theta) \right)' W \left( \Lambda^d - \Lambda^m(\theta) \right)$$

where W is a positive-definite weighting matrix.

# Weighting Matrix

- Common weighting matrices, W, are
  - 1. Theoretically optimal Inverse of covariance matrix of empirical moments Can cause problems in finite samples
  - 2. Identity, / Equal weighting. Does not take level-differences out of moments
  - 3. **Diagonal matrix** with *inverse* of empirical moment *variances* Removes "level" differences. Scales with uncertainty about empirical moments

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4. Freely chosen Focus on fitting some specific dimensions of the data

- 1. **Solve** the buffer-stock model and **simulate** a full panel
- 2. Construct a data set from the simulated data
- 3. Try to **estimate**  $\theta = \{\beta, \rho\}$ using as moments the average wealth for each age between 40 and 55  $\Lambda^d = (A_{40}, A_{41}, \dots, A_{55})$

### Estimation experiment

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I will now describe how to calculate the objective function

$$Q(\theta) = \left(\Lambda^d - \Lambda^m(\theta)\right)' W \left(\Lambda^d - \Lambda^m(\theta)\right)$$

for a given value of  $\theta$ .

This function should then be minimized to get

$$\hat{\theta} = \arg\min_{\theta} \, Q(\theta)$$

1. Solve model to get  $c_t^*(m;\theta)$  for all t on a grid of m (2-dim array)

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- 2. For s = 1, ..., S:
  - 2.1 Simulate N agents for T periods to get

$$C_{it}^{(s)}(\theta) = P_{it}^{(s)} \cdot \check{c}_{t}^{\star}(M_{it}^{(s)}(\theta)/P_{it}^{(s)};\theta)$$

$$M_{it}^{(s)}(\theta) = RA_{it-1}^{(s)}(\theta) + Y_{it}^{(s)}$$

$$A_{it-1}^{(s)}(\theta) = M_{it-1}^{(s)}(\theta) - C_{it-1}^{(s)}(\theta)$$

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for some initial  $A_{i0}$  and  $P_{i0}$  and draws of ?

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for some initial  $A_{i0}$  and  $P_{i0}$  and draws of  $\xi_{i*}^{(s)}$  and  $\psi_{i*}^{(s)}$ .

- 1. Solve model to get  $c_t^*(m;\theta)$  for all t on a grid of m (2-dim array)
- 2. For s = 1, .... S:
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$$\begin{split} C_{it}^{(s)}(\theta) &= P_{it}^{(s)} \cdot \check{c}_{t}^{\star}(M_{it}^{(s)}(\theta)/P_{it}^{(s)};\theta) \\ M_{it}^{(s)}(\theta) &= RA_{it-1}^{(s)}(\theta) + Y_{it}^{(s)} \\ A_{it-1}^{(s)}(\theta) &= M_{it-1}^{(s)}(\theta) - C_{it-1}^{(s)}(\theta) \\ Y_{it}^{(s)} &= P_{it}^{(s)} \xi_{it}^{(s)} \\ P_{it}^{(s)} &= GP_{it-1}^{(s)} \psi_{it}^{(s)} \end{split}$$

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for some initial  $A_{i0}$  and  $P_{i0}$  and draws of  $\xi_{it}^{(s)}$  and  $\psi_{it}^{(s)}$ .

2.2 Calculate moments using simulated data,  $\Lambda_s(\theta) = \{\frac{1}{N} \sum_{i=1}^{N} A_{i*}^{(s)}(\theta)\}_{t=40}^{55}$ 

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Structural Estimation

for some initial  $A_{i0}$  and  $P_{i0}$  and draws of  $\xi_{it}^{(s)}$  and  $\psi_{it}^{(s)}$ .

- 2.2 Calculate moments using simulated data,  $\Lambda_s(\theta) = \{\frac{1}{N} \sum_{i=1}^{N} A_{it}^{(s)}(\theta)\}_{t=40}^{55}$
- 3. Calculate the objective function with  $\Lambda^m(\theta) = \frac{1}{5} \sum_{s=1}^{5} \Lambda_s(\theta)$

$$Q(\theta) = \left(\Lambda^d - \Lambda^m(\theta)\right)' W\left(\Lambda^d - \Lambda^m(\theta)\right)$$

1. Solve model to get  $c_t^*(m; \theta)$  for all t on a grid of m (2-dim array)

- 1. Solve model to get  $c_t^*(m;\theta)$  for all t on a grid of m (2-dim array)
- 2. Simulate  $\tilde{S} = SN$  agents for T periods to get

$$C_{t}^{(s)}(\theta) = P_{t}^{(s)} \cdot \check{c}_{t}^{\star}(M_{i}^{(s)}(\theta)/P_{t}^{(s)};\theta)$$

$$M_{t}^{(s)}(\theta) = RA_{t-1}^{(s)}(\theta) + Y_{t}^{(s)}$$

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Structural Estimation

for some initial  $A_0$  and  $P_0$  and draws of  $\mathcal{E}_{+}^{(s)}$  and  $\psi_{+}^{(s)}$ .

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Structural Estimation

for some initial  $A_0$  and  $P_0$  and draws of  $\xi_{\star}^{(s)}$  and  $\psi_{\star}^{(s)}$ 

3. Calculate simulated moments,  $\Lambda^m(\theta) = \{\frac{1}{5} \sum_{s=1}^{\tilde{S}} A_t^{(s)}(\theta)\}_{t=40}^{55}$  now

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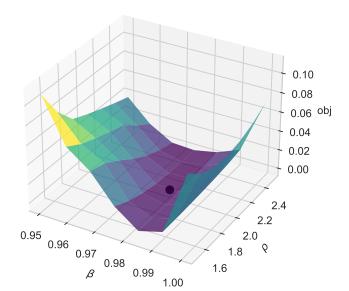
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for some initial  $A_0$  and  $P_0$  and draws of  $\zeta_{+}^{(s)}$  and  $\psi_{+}^{(s)}$ 

- 3. Calculate simulated moments,  $\Lambda^m(\theta) = \{\frac{1}{\xi} \sum_{s=1}^{\tilde{\xi}} A_t^{(s)}(\theta)\}_{t=40}^{55}$  now
- 4. Calculate the objective function

$$Q(\theta) = \left(\Lambda^d - \Lambda^m(\theta)\right)' W\left(\Lambda^d - \Lambda^m(\theta)\right)$$

### Buffer-stock: MSM



## Indirect inference / minimum distance

- Many different names for very similar approaches
  - McFadden (1989): Method of Simulated Moments (MSM)
  - Duffie and Singleton (1993): Simulated Minimum Distance (SMD)

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• Gouriéroux, Monfort and Renault (1993) + Smith (1993): Indirect Inference (II)

### Indirect inference / minimum distance

- Many different names for very similar approaches
  - McFadden (1989): Method of Simulated Moments (MSM)
  - Duffie and Singleton (1993): Simulated Minimum Distance (SMD)

- Gouriéroux, Monfort and Renault (1993) + Smith (1993): Indirect Inference (II)
- SMD/II rely on an auxillary statistical model
  - ullet Let  $\Lambda^d$  be the parameters of the auxillary model when estimated on the actual data
  - Let  $\Lambda_s(\theta)$  be the parameters of the auxiliary model when estimated on simulated data
- Note: The auxiliary statistical model is misspecified and its parameters are thus typically not interpretable

### Simulation Pitfalls

- FIX the seed (or draws!)
- Flat objective function!
  - Discrete choices: Taking a mean of an indicator function

- Gradient based numerical optimization will likely FAIL!
  - Use, e.g., scipy.optimize.minimize(fun , method='Nelder-Mead') (Nelder-Mead)
  - Or some smoothing device (e.g. Logit)
- As  $N, S \rightarrow \infty$  this problem vanishes
- The problem is also less severe around  $\theta_0$
- Continuous outcomes do not have this problem

 MSM is consistent and asymptotically normal under standard assumptions

$$\sqrt{\textit{N}}(\hat{ heta}- heta_0) 
ightarrow \mathcal{N}( exttt{0,} (1+\textit{S}^{-1})\textit{V})$$

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where  $\theta_0$  are the true parameters

Standard formulas for V:

$$V = (G'WG)^{-1}G'W\Omega W'G(G'WG)^{-1}$$

where  $G = -\frac{\partial \Lambda^m(\theta)}{\partial \theta}$  is the Jacobian of the objective function.  $\Omega = Var(\Lambda_i^d)$  is the variance of the (individual) moments in the data. **Remember:** Standard errors are large if large changes in  $\theta$  imply small changes in the objective function

### Identification

• Is there enough variation in the data to identify  $\theta$ ? Very hard to prove anything because the model is typically strongly non-linear

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- Problems:
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- **Graphical inspection is useful:** Plot the objective function in the neighborhood of the found optimum
  - "Informativeness of moments": Honoré, Jørgensen and de Paula (2020)

#### Curse of dimensionality and lack of identification

⇒ we cannot estimate all the parameters of the model

- ⇒ first step calibration is necessary
  - 1. Calculations on own data (e.g. exogenous processes)
  - 2. References to previous estimates
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- Or the opposite: When does the result break down?
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- Calibration is also important for
  - 1. Gaining intuition for how the model works
  - 2. Initial guesses for estimation algorithm

#### Next time:

Static and dynamic labor supply Recap for some + new stuff for most.

#### Literature:

Keane (2011, sections 1–5): "Labor Supply and Taxes: A Survey"

- Read before lecture
- Reading guide:
  - Section 1: short Introduction
  - Section 2: Optimal Taxation, Motivation. Skim fast.
  - Section 3: Basic model. Key, focus here.
  - Section 4: Econometric issues. Skim.
  - Section 5: Roadmap of empirical literature. *Short, read.*
  - (Remaining: empirical literature.)

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