

# Assignment 8

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Download all python codes from

<https://github.com/Adarsh541/AI1103-prob-and-ranvar/blob/main/Assignment8.1/codes/Assignment8.1.py>

and latex-tikz codes from

<https://github.com/Adarsh541/AI1103-prob-and-ranvar/blob/main/Assignment8.1/Assignment8.1.tex>

where

$$\mathbf{M} = \begin{pmatrix} \frac{4}{5} & -1/5 & -1/5 & -1/5 & -1/5 \\ -1/5 & \frac{4}{5} & -1/5 & -1/5 & -1/5 \\ -1/5 & -1/5 & \frac{4}{5} & -1/5 & -1/5 \\ -1/5 & -1/5 & -1/5 & \frac{4}{5} & -1/5 \\ -1/5 & -1/5 & -1/5 & -1/5 & \frac{4}{5} \end{pmatrix} \quad (2.0.5)$$

(2.0.6)

we also have

$$\mathbf{M}^2 = \mathbf{M} \quad (2.0.7)$$

## 1 PROBLEM

Let  $X_1, X_2, X_3, X_4, X_5$  be a random sample of size 5 from a population having standard normal distribution. If  $\bar{X} = \frac{1}{5} \sum_{i=1}^5 X_i$  and  $T = \sum_{i=1}^5 (X_i - \bar{X})^2$  then  $E[T^2 \bar{X}^2]$  is equal to

- 1) 3
- 2) 3.6
- 3) 4.8
- 4) 5.2

## 2 SOLUTION

Let  $\mathbf{x} \sim N(\mathbf{0}, \mathbf{I})$  be a standard normal random vector

$$\mathbf{x} = \begin{pmatrix} X_1 \\ X_2 \\ X_3 \\ X_4 \\ X_5 \end{pmatrix} \quad (2.0.1)$$

Then  $\bar{X}$  can be written as

$$\bar{X} = \frac{1}{5} \mathbf{u}^T \mathbf{x} \quad (2.0.2)$$

where

$$\mathbf{u} = \begin{pmatrix} 1 \\ 1 \\ 1 \\ 1 \\ 1 \end{pmatrix} \quad (2.0.3)$$

$$T = \mathbf{x}^T \mathbf{M} \mathbf{x} \quad (2.0.4)$$

## Lemma 2.1.

$$(\mathbf{u}^T \mathbf{x})(\mathbf{x}^T \mathbf{v}) = \mathbf{u}^T (\mathbf{x} \mathbf{x}^T) \mathbf{v} \quad (2.0.8)$$

*This lemma is verified using python simulation.*

**Lemma 2.2.** If  $\mathbf{A}, \mathbf{B}$  are two continuous independent random vectors then  $\mathbf{A}, \mathbf{B}^T \mathbf{B}$  are independent.

*Proof.* Since  $\mathbf{A}, \mathbf{B}$  are independent, for any  $\mathbf{a}, \mathbf{b}$  we can write

$$\Pr(\mathbf{A} = \mathbf{a}, \mathbf{B} = \mathbf{b}) = \Pr(\mathbf{A} = \mathbf{a}) \Pr(\mathbf{B} = \mathbf{b}) \quad (2.0.9)$$

without loss of generality we can assume the solution of the set

$$\{\mathbf{B} | \mathbf{B}^T \mathbf{B} = \mathbf{c}\} = \{\mathbf{b}_1, \mathbf{b}_2, \dots, \mathbf{b}_k\} \quad (2.0.10)$$

$$\begin{aligned} \Pr(\mathbf{A} = \mathbf{a}, \mathbf{B}^T \mathbf{B} = \mathbf{c}) &= \Pr(\mathbf{A} = \mathbf{a}, \mathbf{B} = \mathbf{b}_1) \\ &+ \Pr(\mathbf{A} = \mathbf{a}, \mathbf{B} = \mathbf{b}_2) + \dots + \Pr(\mathbf{A} = \mathbf{a}, \mathbf{B} = \mathbf{b}_k) \end{aligned} \quad (2.0.11)$$

$$= \Pr(\mathbf{A} = \mathbf{a}) (\Pr(\mathbf{B} = \mathbf{b}_1) + \dots + \Pr(\mathbf{B} = \mathbf{b}_k)) \quad (2.0.12)$$

$$= \Pr(\mathbf{A} = \mathbf{a}) \Pr(\mathbf{B}^T \mathbf{B} = \mathbf{c}) \quad (2.0.13)$$

□

**Definition 2.1** (cross-covariance).

$$\text{Cov}[\mathbf{x}, \mathbf{y}] = E[(\mathbf{x} - E[\mathbf{x}])(\mathbf{y} - E[\mathbf{y}])^T] \quad (2.0.14)$$

**Lemma 2.3.** Two jointly normal vectors are independent if and only if their cross-covariance is zero.

**Theorem 2.4.** Let  $\mathbf{x}$  be a  $5 \times 1$  standard multivariate normal random vector. Let  $\mathbf{B}$  be an  $l \times 5$  real matrix. Then the  $l \times 1$  random vector  $\mathbf{y}$  defined by  $\mathbf{y} = \mathbf{B}\mathbf{x}$  has multivariate normal distribution with mean  $E[\mathbf{y}] = \mathbf{0}$  and covariance matrix  $\text{Var}[\mathbf{y}] = \mathbf{B}\mathbf{B}^\top$

*Proof.* The joint moment generating function of  $\mathbf{x}$  is

$$M_{\mathbf{x}}(\mathbf{t}) = \exp\left(\mathbf{t}^\top \mu + \frac{1}{2} \mathbf{t}^\top \mathbf{V} \mathbf{t}\right) \quad (2.0.15)$$

since for standard normal distribution  $\mu = \mathbf{0}$  and  $\mathbf{V} = \mathbf{I}$ . So

$$M_{\mathbf{x}}(\mathbf{t}) = \exp\left(\frac{1}{2} \mathbf{t}^\top \mathbf{I} \mathbf{t}\right) \quad (2.0.16)$$

Therefore the joint moment generating function of  $\mathbf{y}$  is

$$M_{\mathbf{y}}(\mathbf{t}) = M_{\mathbf{x}}(\mathbf{B}^\top \mathbf{t}) \quad (2.0.17)$$

$$= \exp\left(\frac{1}{2} \mathbf{t}^\top \mathbf{B} \mathbf{B}^\top \mathbf{t}\right) \quad (2.0.18)$$

on comparing with (2.0.16) we can say  $\mathbf{y}$  has multivariate normal distribution.  $\square$

**Theorem 2.5.** Let  $\mathbf{x}$  be a  $5 \times 1$  standard multivariate normal random vector. Let  $\mathbf{A}, \mathbf{B}$  be two matrices. Define

$$\mathbf{T}_1 = \mathbf{A}\mathbf{x} \quad (2.0.19)$$

$$\mathbf{T}_2 = \mathbf{B}\mathbf{x} \quad (2.0.20)$$

Then  $\mathbf{T}_1$  and  $\mathbf{T}_2$  are two independent random vectors if and only if  $\mathbf{A}\mathbf{B}^\top = \mathbf{0}$

*Proof.* From theorem 2.4,  $\mathbf{T}_1$  and  $\mathbf{T}_2$  are jointly normal. Their cross-covariance is

$$\text{Cov}[\mathbf{T}_1, \mathbf{T}_2] = E[(\mathbf{T}_1 - E[\mathbf{T}_1])(\mathbf{T}_2 - E[\mathbf{T}_2])^\top] \quad (2.0.21)$$

$$= E[(\mathbf{A}\mathbf{x} - E[\mathbf{A}\mathbf{x}])(\mathbf{B}\mathbf{x} - E[\mathbf{B}\mathbf{x}])^\top] \quad (2.0.22)$$

$$= \mathbf{A}E[(\mathbf{x} - E[\mathbf{x}])(\mathbf{x} - E[\mathbf{x}])^\top]\mathbf{B}^\top \quad (2.0.23)$$

$$= \mathbf{A}\text{Var}[\mathbf{x}]\mathbf{B}^\top \quad (2.0.24)$$

$$= \mathbf{A}\mathbf{B}^\top \quad (2.0.25)$$

So  $\mathbf{T}_1$  and  $\mathbf{T}_2$  are independent if and only if  $\mathbf{A}\mathbf{B}^\top = \mathbf{0}$   $\square$

**Theorem 2.6.** Let  $\bar{X}$  be the sample mean of size 5 from a standard normal distribution. Then

$$1) \bar{X} \sim N(0, \frac{1}{5})$$

$$2) \bar{X} \text{ and } T \text{ are independent.}$$

$$3) T \sim \chi_4^2$$

where  $\chi_4^2$  is chi-square distribution with 4 degrees of freedom and  $T$  is defined as

$$T = \sum_{i=1}^5 (X_i - \bar{X})^2 \quad (2.0.26)$$

*Proof.* 1)

$$\bar{X} = \frac{1}{5} \mathbf{u}^\top \mathbf{x} \quad (2.0.27)$$

From theorem 2.4 we can say  $\bar{X}$  has normal distribution with mean  $E[\bar{X}] = \mathbf{0}$  and covariance matrix

$$\text{Var}[\bar{X}] = \frac{1}{25} \mathbf{u}^\top \mathbf{u} \quad (2.0.28)$$

$$= \frac{1}{5} \quad (2.0.29)$$

2) since  $\mathbf{M}$  is symmetric and idempotent we have

$$T = \mathbf{x}^\top \mathbf{M} \mathbf{x} \quad (2.0.30)$$

$$= \mathbf{x}^\top \mathbf{M} \mathbf{M} \mathbf{x} \quad (2.0.31)$$

$$= \mathbf{x}^\top \mathbf{M}^\top \mathbf{M} \mathbf{x} \quad (2.0.32)$$

$$= (\mathbf{M}\mathbf{x})^\top (\mathbf{M}\mathbf{x}) \quad (2.0.33)$$

$$\text{Cov}\left[\frac{1}{5} \mathbf{u}^\top \mathbf{x}, \mathbf{M}\mathbf{x}\right] = E\left[\left(\frac{1}{5} \mathbf{u}^\top \mathbf{x} - E\left[\frac{1}{5} \mathbf{u}^\top \mathbf{x}\right]\right)(\mathbf{M}\mathbf{x} - E[\mathbf{M}\mathbf{x}])^\top\right] \quad (2.0.34)$$

Using lemma 2.1 we get

$$\begin{aligned} \text{Cov}\left[\frac{1}{5} \mathbf{u}^\top \mathbf{x}, \mathbf{M}\mathbf{x}\right] &= \frac{1}{5} \mathbf{u}^\top E[(\mathbf{x} - E[\mathbf{x}])(\mathbf{x} - E[\mathbf{x}])^\top] \mathbf{M}^\top \quad (2.0.35) \\ &= \frac{1}{5} \mathbf{u}^\top \text{Var}[\mathbf{x}] \mathbf{M} \quad (2.0.36) \\ &= \frac{1}{5} \mathbf{u}^\top \mathbf{M} \quad (2.0.37) \\ &= 0 \quad (2.0.38) \end{aligned}$$

So  $\mathbf{M}\mathbf{x}$  and  $\frac{1}{5} \mathbf{u}^\top \mathbf{x}$  are independent. Using lemma 2.2 we get  $\bar{X}$  and  $T$  are independent.

3) Since  $\mathbf{M}$  is symmetric it can be expressed as

$$\mathbf{M} = \mathbf{P} \mathbf{D} \mathbf{P}^\top \quad (2.0.39)$$

where  $\mathbf{P}$  is orthogonal and  $\mathbf{D}$  is diagonal. Then

$$T = \mathbf{x}^\top \mathbf{M} \mathbf{x} \quad (2.0.40)$$

$$= \mathbf{x}^\top \mathbf{P} \mathbf{D} \mathbf{P}^\top \mathbf{x} \quad (2.0.41)$$

$$= (\mathbf{P}^\top \mathbf{x})^\top \mathbf{D} \mathbf{P}^\top \mathbf{x} \quad (2.0.42)$$

$$= \mathbf{y}^\top \mathbf{D} \mathbf{y} \quad (2.0.43)$$

where  $\mathbf{y} = \mathbf{P}^\top \mathbf{x}$ . From theorem 2.4 and orthogonality of  $\mathbf{P}$  we can say  $\mathbf{y} \sim N(\mathbf{0}, \mathbf{I})$

- a) The eigen values of  $\mathbf{M}$  are 1, 1, 1, 1, 0. So  $\mathbf{D}$  can be written as

$$\mathbf{D} = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{pmatrix} \quad (2.0.44)$$

Let  $\mathbf{v}_1, \dots, \mathbf{v}_5$  be corresponding eigen vectors. Then  $\mathbf{P} = (\mathbf{v}_1, \dots, \mathbf{v}_5)$ . Since  $\mathbf{P}$  is orthogonal the dot product of any two eigen vectors is zero. i.e

$$\mathbf{v}_i^\top \mathbf{v}_j = 0 \quad (2.0.45)$$

for any  $i \neq j$

$$\mathbf{y} = \begin{pmatrix} y_1 = \mathbf{v}_1^\top \mathbf{x} \\ y_2 = \mathbf{v}_2^\top \mathbf{x} \\ y_3 = \mathbf{v}_3^\top \mathbf{x} \\ y_4 = \mathbf{v}_4^\top \mathbf{x} \\ y_5 = \mathbf{v}_5^\top \mathbf{x} \end{pmatrix} \quad (2.0.46)$$

$$(2.0.47)$$

From (2.0.45) and theorem 2.5, it follows  $y_1, y_2, y_3, y_4, y_5$  are mutually independent.

$$T = y_1^2 + y_2^2 + y_3^2 + y_4^2 \quad (2.0.48)$$

So  $T$  is sum of squares of four independent standard normal variables which is chi-square distribution with 4 degrees of freedom.

□

$$E[T^2 \bar{X}^2] = E[T^2] E[\bar{X}^2] \quad (2.0.49)$$

$$E[\bar{X}^2] = \text{Var}[\bar{X}] + (E[\bar{X}])^2 \quad (2.0.50)$$

$$= \frac{1}{5} \quad (2.0.51)$$

since  $T$  is chi-squared distributed with 4 degrees of freedom

$$E[T] = 4 \quad (2.0.52)$$

$$\text{Var}[T] = 8 \quad (2.0.53)$$

$$\Rightarrow E[T^2] = \text{Var}[T] + (E[T])^2 \quad (2.0.54)$$

$$= 24 \quad (2.0.55)$$

From (2.0.49)

$$E[T^2 \bar{X}^2] = 4.8 \quad (2.0.56)$$