

# Assignment 8

Adarsh Sai - AI20BTECH11001

Download latex-tikz codes from

<https://github.com/Adarsh541/AI1103-prob-and-ranvar/blob/main/Assignment8/Assignment8.tex>

## 1 PROBLEM(GOV/STATS/2015/STATISTICS-I(1),Q.3(A))

Let  $X_1, X_2, \dots, X_n$  be independent Poisson variates with  $E[X_i] = \mu_i$ . Find the conditional distribution of  $X_1, \dots, X_n \mid \sum_{i=1}^n X_i = y$

## 2 SOLUTION(GOV/STATS/2015/STATISTICS-I(1),Q.3(A))

$$F\left\{x_1, \dots, x_n \mid \sum_{i=1}^n X_i = y\right\} = \Pr\left(X_1 \leq x_1, \dots, X_n \leq x_n \mid \sum_{i=1}^n X_i = y\right) \quad (2.0.1)$$

$$= \frac{\Pr(X_1 \leq x_1, \dots, X_n \leq x_n, \sum_{i=1}^n X_i = y)}{\Pr(\sum_{i=1}^n X_i = y)} \quad (2.0.2)$$

1) To find numerator of (2.0.2), we solve for two random variables and extend it to n random variables.

a) Solving for two variables. Consider the sets

$$\{X_1 \leq x_1\} \quad (2.0.3)$$

$$\{X_2 \leq x_2\} \quad (2.0.4)$$

$$\{X_1 + X_2 = y\} \quad (2.0.5)$$

i) If  $x_1 + x_2 = y$ . Adding (2.0.3) and (2.0.4)

$$X_1 + X_2 \leq x_1 + x_2 = y \quad (2.0.6)$$

and the equality holds when

$$X_1 = x_1 \quad (2.0.7)$$

$$X_2 = x_2 \quad (2.0.8)$$

From (2.0.5) and (2.0.6)

$$X_1 + X_2 = y \quad (2.0.9)$$

So the intersection of the sets is

$$\begin{aligned} \{X_1 \leq x_1, X_2 \leq x_2, X_1 + X_2 = y\} \\ = \{X_1 = x_1, X_2 = x_2\} \end{aligned}$$

Since  $X_1$  and  $X_2$  are independent

$$\begin{aligned} \Rightarrow \Pr(X_1 \leq x_1, X_2 \leq x_2, X_1 + X_2 = y) \\ = \Pr(X_1 = x_1) \Pr(X_2 = x_2) \quad (2.0.10) \end{aligned}$$

$$= \prod_{i=1}^2 \left( \frac{\mu_i^{x_i} e^{-\mu_i}}{x_i!} \right) \quad (2.0.11)$$

$$= (e^{-\sum \mu_i}) \prod_{i=1}^2 \left( \frac{\mu_i^{x_i}}{x_i!} \right) \quad (2.0.12)$$

ii) if  $x_1 + x_2 < y$ . Adding (2.0.3) and (2.0.4)

$$X_1 + X_2 \leq x_1 + x_2 < y \quad (2.0.13)$$

but from (2.0.5)

$$X_1 + X_2 = y \quad (2.0.14)$$

So there exist no values for  $X_1$  and  $X_2$  Therefore

$$\begin{aligned} \Pr(X_1 \leq x_1, X_2 \leq x_2, X_1 + X_2 = y) \\ = \begin{cases} 0 & x_1 + x_2 < y \\ (e^{-\sum \mu_i}) \prod_{i=1}^2 \left( \frac{\mu_i^{x_i}}{x_i!} \right) & x_1 + x_2 = y \end{cases} \quad (2.0.15) \end{aligned}$$

Extending to n variables we get

$$\begin{aligned} \Pr\left(X_1 \leq x_1, \dots, X_n \leq x_n, \sum_{i=1}^n X_i = y\right) \\ = \begin{cases} (e^{-\sum \mu_i}) \prod_{i=1}^n \left( \frac{\mu_i^{x_i}}{x_i!} \right) & \sum_{i=1}^n x_i = y \\ 0 & \sum_{i=1}^n x_i < y \end{cases} \quad (2.0.16) \end{aligned}$$

2) **Theorem:** If the random variables  $X_1, X_2$  are independent and poisson distributed with parameters  $\lambda_1, \lambda_2$ , then their sum  $Z = X_1 + X_2$  is also poisson distributed with parameter  $\lambda_1 + \lambda_2$ . **proof:** The characteristic function of a pois-

son random variable is given by

$$\Phi_X(\omega) = e^{-\lambda(1-e^{j\omega})} \quad (2.0.17)$$

Using convolution

$$\Phi_Z(\omega) = \Phi_{X_1}(\omega)\Phi_{X_2}(\omega) \quad (2.0.18)$$

$$= e^{-\lambda_1(1-e^{j\omega})} e^{-\lambda_2(1-e^{j\omega})} \quad (2.0.19)$$

$$= e^{-(\lambda_1+\lambda_2)(1-e^{j\omega})} \quad (2.0.20)$$

This theorem can be extended to n variables. So

$$\Pr\left(\sum_{i=1}^n X_i = y\right) = \frac{(\sum \mu_i)^y e^{-\sum \mu_i}}{y!} \quad (2.0.21)$$

From (2.0.2),(2.0.16) and (2.0.21)

$$\begin{aligned} & F\left\{x_1, \dots, x_n \left| \sum_{i=1}^n X_i = y\right.\right\} \\ &= \begin{cases} \frac{y!}{(\sum \mu_i)^y} \prod_{i=1}^n \left(\frac{\mu_i^{x_i}}{x_i!}\right) & \sum_{i=1}^n x_i = y \\ 0 & \sum_{i=1}^n x_i < y \end{cases} \quad (2.0.22) \end{aligned}$$