

My Presentation

Adepu Adarsh Sai

IITH(AI)

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chi-square distribution

Let X_1, X_2, \dots, X_k be i.i.d standard normal random variables. Define a random variable Y as

$$Y = X_1^2 + X_2^2 + \dots + X_k^2 \quad (1)$$

We say Y is chi-square distributed with k degrees of freedom. The mean and variance is given by

$$E[Y] = k \quad (2)$$

$$\text{Var}[Y] = 2k \quad (3)$$

cross-covariance

$$\text{Cov}[x, y] = E[(x - E[x])(y - E[y])^T] \quad (4)$$

Theorem-1

Let x be a $k \times 1$ standard multivariate normal random vector. Let B be an $l \times k$ real matrix. Then the $l \times 1$ random vector y defined by $y = Bx$ has multivariate normal distribution with mean $E[y] = 0$ and covariance matrix $\text{Var}[y] = BB^T$

PROOF:

The joint moment generating function of x is

$$M_x(t) = \exp\left(t^T \mu + \frac{1}{2} t^T V t\right) \quad (5)$$

since for standard normal distribution $\mu = 0$ and $V = I$. So

$$M_x(t) = \exp\left(\frac{1}{2} t^T I t\right) \quad (6)$$

Therefore the joint moment generating function of y is

$$M_y(t) = M_x(B^T t) \quad (7)$$

$$= \exp\left(\frac{1}{2} t^T B B^T t\right) \quad (8)$$

on comparing with (6) we can say y has multivariate normal distribution.

Theorem-2

Let x be a $k \times 1$ standard multivariate normal random vector. Let A, B be two matrices. Define

$$T_1 = Ax \quad (9)$$

$$T_2 = Bx \quad (10)$$

Then T_1 and T_2 are two independent random vectors if and only if $AB^T = 0$

PROOF

From theorem-1, T_1 and T_2 are jointly normal vectors. Their cross-covariance is

$$\text{Cov}[T_1, T_2] = E[(T_1 - E[T_1])(T_2 - E[T_2])^T] \quad (11)$$

$$= E[(Ax - E[Ax])(Bx - E[Bx])^T] \quad (12)$$

$$= AE[(x - E[x])(x - E[x])^T]B^T \quad (13)$$

$$= A \text{Var}[x] B^T \quad (14)$$

$$= AB^T \quad (15)$$

Two jointly normal vectors are independent if and only if their cross-covariance is zero. So T_1 and T_2 are independent if and only if $AB^T = 0$

Theorem-3

Let \bar{X} be the sample mean of size 5 from a standard normal distribution. Then

- 1 $\bar{X} \sim N(0, \frac{1}{5})$
- 2 \bar{X} and T are independent.
- 3 $T \sim \chi_4^2$

where χ_4^2 is chi-square distribution with 4 degrees of freedom and T (Sample variance) is given by

$$T = \sum_{i=1}^5 (X_i - \bar{X})^2 \quad (16)$$

PROOF

$$\overline{X} = \frac{1}{5} \mathbf{u}^\top \mathbf{x} \quad (17)$$

From theorem-1 we can say \overline{X} has normal distribution with mean $E[\overline{X}] = 0$ and covariance matrix

$$\text{Var} [\overline{X}] = \frac{1}{25} \mathbf{u}^\top \mathbf{u} \quad (18)$$

$$= \frac{1}{5} \quad (19)$$

$$T = x^T M x \quad (20)$$

where

$$M = \begin{pmatrix} \frac{4}{5} & -1/5 & -1/5 & -1/5 & -1/5 \\ -1/5 & \frac{4}{5} & -1/5 & -1/5 & -1/5 \\ -1/5 & -1/5 & \frac{4}{5} & -1/5 & -1/5 \\ -1/5 & -1/5 & -1/5 & \frac{4}{5} & -1/5 \\ -1/5 & -1/5 & -1/5 & -1/5 & \frac{4}{5} \end{pmatrix} \quad (21)$$

$$(22)$$

we also have

$$M^2 = M \quad (23)$$

$$T = x^T Mx \quad (24)$$

$$= x^T M Mx \quad (25)$$

$$= x^T M^T Mx \quad (26)$$

$$= (Mx)^T (Mx) \quad (27)$$

$$\text{Cov} \left[\frac{1}{5} u^T x, Mx \right] = E \left[\left(\frac{1}{5} u^T x - E \left[\frac{1}{5} u^T x \right] \right) (Mx - E[Mx])^T \right] \quad (28)$$

$$= \frac{1}{5} u^T E[(x - E[x]) (x - E[x])^T] M^T \quad (29)$$

$$= \frac{1}{5} u^T \text{Var}[x] M \quad (30)$$

$$= \frac{1}{5} u^T M \quad (31)$$

$$= 0 \quad (32)$$

Lemma

Functions of independent random variables are themselves independent.

proof that $y = x^\top x$ is a function

Assume that it is not a function. So there exist $x_1 \in \mathbb{R}^n$, $a, b \in \mathbb{R}$, $a \neq b$ such that

$$x_1^\top x_1 = a \quad (33)$$

$$x_1^\top x_1 = b \quad (34)$$

But from (33) and (34) it can clearly be seen that

$$a = b \quad (35)$$

which is a contradiction. So our assumption is wrong. Therefore $y : \mathbb{R}^n \rightarrow \mathbb{R}$ is a function.

So Mx and $\frac{1}{\sqrt{5}}u^T x$ are independent. Since functions of two independent variables are also independent, we can say \bar{X} and T (A function of Mx) are independent.

Since M is symmetric it can be expressed as

$$M = PDP^T \quad (36)$$

where P is orthogonal and D is diagonal. Then

$$T = x^T Mx \quad (37)$$

$$= x^T PDP^T x \quad (38)$$

$$= (P^T x)^T D P^T x \quad (39)$$

$$= y^T D y \quad (40)$$

where $y = P^T x$. Since x is standard normal, from theorem-1 we can say y is also jointly normal with

$$E[y] = 0 \quad (41)$$

$$\text{Var}[y] = P^T (P^T)^T \quad (42)$$

$$= P^T P \quad (43)$$

$$= I \quad (\text{Since } P \text{ is orthogonal}) \quad (44)$$

So $y \sim N(0, I)$ is standard normal

The eigen values of M are 1, 1, 1, 1, 0. So D can be written as

$$D = \begin{pmatrix} 1 & 0 & 0 & 0 & 0 \\ 0 & 1 & 0 & 0 & 0 \\ 0 & 0 & 1 & 0 & 0 \\ 0 & 0 & 0 & 1 & 0 \\ 0 & 0 & 0 & 0 & 0 \end{pmatrix} \quad (45)$$

Let v_1, \dots, v_5 be corresponding eigen vectors. Then $P = (v_1, \dots, v_5)$. Since P is orthogonal the dot product of any two eigen vectors is zero. i.e

$$v_i^T v_j = 0 \quad \text{for any } i \neq j \quad (46)$$

$$y = P^T x \quad (47)$$

$$\Rightarrow y = \begin{pmatrix} y_1 = v_1^T x \\ y_2 = v_2^T x \\ y_3 = v_3^T x \\ y_4 = v_4^T x \\ y_5 = v_5^T x \end{pmatrix} \quad (48)$$

$$(49)$$

From (46) and theorem-2, it follows y_1, y_2, y_3, y_4, y_5 are mutually independent.

$$T = y^T D y \quad (50)$$

$$\Rightarrow T = y_1^2 + y_2^2 + y_3^2 + y_4^2 \quad (51)$$

So T is sum of squares of four independent standard normal variables which is chi-square distribution with 4 degrees of freedom.

Question

gov/stats/2019/STATS-P1-IESISS,(Q.25)

Let X_1, X_2, X_3, X_4, X_5 be a random sample of size 5 from a population having standard normal distribution. If $\bar{X} = \frac{1}{5} \sum_{i=1}^5 X_i$ and

$T = \sum_{i=1}^5 (X_i - \bar{X})^2$ then $E[T^2 \bar{X}^2]$ is equal to

- ① 3
- ② 3.6
- ③ 4.8
- ④ 5.2

Solution

$$E[T^2\bar{X}^2] = E[T^2]E[\bar{X}^2] \quad (52)$$

$$E[\bar{X}^2] = \text{Var} [\bar{X}] + (E[\bar{X}])^2 \quad (53)$$

$$= \frac{1}{5} \quad (54)$$

since T is chi-squared distributed with 4 degrees of freedom

$$E[T] = 4 \quad (55)$$

$$\text{Var} [T] = 8 \quad (56)$$

$$\implies E[T^2] = \text{Var} [T] + (E[T])^2 \quad (57)$$

$$= 24 \quad (58)$$

From (52)

$$E[T^2\bar{X}^2] = 4.8 \quad (59)$$