

# Assignment 7

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Download all python codes from

<https://github.com/Adarsh541/AI1103-prob-and-ranvar/blob/%main/Assignment6/codes/Assignment6.py>

Download latex-tikz codes from

<https://github.com/Adarsh541/AI1103-prob-and-ranvar/blob/main/Assignment7/Assignment7.tex>

## 1 PROBLEM(CSIR UGC NET EXAM(JUNE 2012)Q.112)

Let  $X_1, X_2, X_3, \dots, X_n$  be i.i.d observations from a distribution with continuous probability density function  $f$  which is symmetric around  $\theta$  i.e

$$f(x - \theta) = f(\theta - x) \quad (1.0.1)$$

for all real  $x$ . Consider the test  $H_0 : \theta = 0$  vs  $H_1 : \theta > 0$  and the sign test statistic

$$S_n = \sum_{i=1}^n \text{sign}(X_i) \quad (1.0.2)$$

where

$$\text{sign}(x) = \begin{cases} 1 & x > 0 \\ 0 & x = 0 \\ -1 & x < 0 \end{cases} \quad (1.0.3)$$

Let  $z_\alpha$  be the upper  $100(1 - \alpha)^{\text{th}}$  percentile of the standard normal distribution where  $0 < \alpha < 1$ . Which of the following is/are correct?

- 1) If  $\theta = 0$  then  $\lim_{n \rightarrow \infty} P\{S_n > \sqrt{n}z_\alpha\} = 1$
- 2) If  $\theta = 0$  then  $\lim_{n \rightarrow \infty} P\{S_n > \sqrt{n}z_\alpha\} = \alpha$
- 3) If  $\theta > 0$  then  $\lim_{n \rightarrow \infty} P\{S_n > \sqrt{n}z_\alpha\} = 1$
- 4) If  $\theta > 0$  then  $\lim_{n \rightarrow \infty} P\{S_n > \sqrt{n}z_\alpha\} = \alpha$

## 2 SOLUTION(CSIR UGC NET EXAM(JUNE 2012)Q.112)

2.1  $H_0 : \theta = 0$

Assume hypothesis  $H_0 : \theta = 0$  is true.

1) Given  $X$  is symmetric around zero.

$$f_X(x) = f_X(-x) \quad (2.1.1)$$

$$\int_0^\infty f_X(x)dx = \int_0^\infty f_X(-x)dx \quad (2.1.2)$$

a) Solving LHS of (2.1.2)

$$\int_0^\infty f_X(x)dx = \Pr(X \geq 0) \quad (2.1.3)$$

b) Solving RHS of (2.1.2)

$$\int_0^\infty f_X(-x)dx \quad (2.1.4)$$

Changing  $-x \rightarrow x$  we get

$$\int_0^\infty f_X(-x)dx = \int_{-\infty}^0 f_X(x)dx \quad (2.1.5)$$

$$= \Pr(X \leq 0) \quad (2.1.6)$$

but

$$\int_{-\infty}^0 f_X(x)dx + \int_0^\infty f_X(x)dx = 1 \quad (2.1.7)$$

from (2.1.2), (2.1.5) and (2.1.7)

$$\int_{-\infty}^0 f_X(x)dx = \int_0^\infty f_X(x)dx = \frac{1}{2} \quad (2.1.8)$$

$$\implies \Pr(X \leq 0) = \Pr(X \geq 0) = \frac{1}{2} \quad (2.1.9)$$

2) Let  $Y$  be a random variable such that

$$Y = \text{sign}(X) \quad (2.1.10)$$

$$Y = \begin{cases} 1 & X > 0 \\ -1 & X < 0 \end{cases} \quad (2.1.11)$$

From (2.1.9) and (2.2.4) we have

$$\Pr(Y = -1) = \Pr(Y = 1) = \frac{1}{2} \quad (2.1.12)$$

So  $Y = \text{sign}(X)$  is also symmetric around zero.

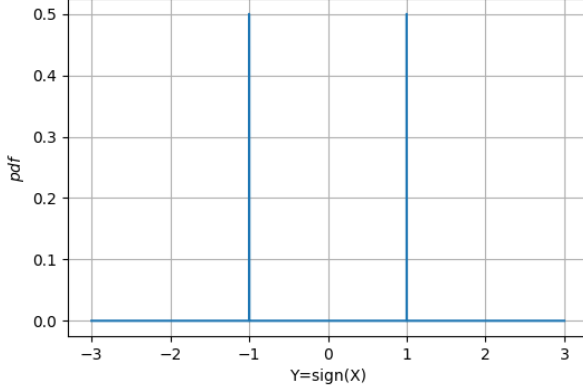


Fig. 2: pdf of  $Y = \text{sign}(X)$

$$\Rightarrow \mu_y = 0 \quad (2.1.13)$$

and variance is

$$\sigma_y^2 = (-1)^2 \left(\frac{1}{2}\right) + (1)^2 \left(\frac{1}{2}\right) \quad (2.1.14)$$

$$= 1 \quad (2.1.15)$$

3) Given

$$S_n = \sum_{i=1}^n \text{sign}(X_i) \quad (2.1.16)$$

$$S_n(\theta = 0) = \sum_{i=1}^n Y_i \quad (2.1.17)$$

From central limit theorem

$$Z = \lim_{n \rightarrow \infty} \sqrt{n} \left( \frac{\frac{S_n}{n} - \mu_y}{\sigma_y} \right) \quad (2.1.18)$$

$$= \lim_{n \rightarrow \infty} \sqrt{n} \left( \frac{S_n}{n} \right) \quad (2.1.19)$$

$$= \lim_{n \rightarrow \infty} \left( \frac{S_n}{\sqrt{n}} \right) \quad (2.1.20)$$

where  $Z$  is a standard normal variable  $N(0,1)$ .

a) Given

$$\alpha = P\{Z > z_\alpha\} \quad (2.1.21)$$

So from (2.1.20) and (2.1.21)

$$\lim_{n \rightarrow \infty} P\left\{ \frac{S_n}{\sqrt{n}} > z_\alpha \right\} = \alpha \quad (2.1.22)$$

$$\Rightarrow \lim_{n \rightarrow \infty} P\{S_n > \sqrt{n}z_\alpha\} = \alpha \quad (2.1.23)$$

2.2  $H_1 : \theta > 0$  is true

1) Given  $X$  is symmetric around  $\theta > 0$ . Let us assume  $\theta = \theta_0 > 0$ .

$$f_X(\theta_0 - x) = f_X(\theta_0 + x) \quad (2.2.1)$$

$$\int_{\theta_0}^{\infty} f_X(\theta_0 - x) dx = \int_{\theta_0}^{\infty} f_X(\theta_0 + x) dx \quad (2.2.2)$$

a) Solving LHS of (2.2.2). Changing  $(\theta_0 - x) \rightarrow t$

$$\int_{\theta_0}^{\infty} f_X(\theta_0 - x) dx = \int_{-\infty}^0 f_X(t) dt \quad (2.2.3)$$

$$= \Pr(X \leq 0) \quad (2.2.4)$$

b) Solving RHS of (2.2.2). Changing  $(\theta_0 + x) \rightarrow t$

$$\int_{\theta_0}^{\infty} f_X(\theta_0 + x) dx = \int_{2\theta_0}^{\infty} f_X(t) dt \quad (2.2.5)$$

$$= \int_0^{\infty} f_X(t) dt - \int_0^{2\theta_0} f_X(t) dt \quad (2.2.6)$$

$$= \Pr(X \geq 0) - k \quad (2.2.7)$$

where

$$k = \int_0^{2\theta_0} f_X(t) dt > 0 \quad (2.2.8)$$

From (2.2.2), (2.2.4) and (2.2.7)

$$\Pr(X \geq 0) > \Pr(X \leq 0) \quad (2.2.9)$$

2) So

$$\Pr(Y = 1) > \Pr(Y = -1) \quad (2.2.10)$$

Therefore, if we perform the experiment and find the value of  $\left(\frac{S_n}{\sqrt{n}}\right)$ , it is most likely to occur on the right side of the distribution of  $\left(\frac{S_n}{\sqrt{n}}\right)$ . In (2.1.20) it is shown that the distribution of the random variable  $\left(\frac{S_n}{\sqrt{n}}\right)$  is  $N(0, 1)$  when  $n$  is very large. So

$$\lim_{n \rightarrow \infty} P\left\{ \frac{S_n}{\sqrt{n}} > z_\alpha \right\} = 1 \quad (2.2.11)$$