

Revision on Conditional Probability and intro to inferential
stats + probability distributions 21oct22

20 October 2022 07:11 PM ✓

Set theory elements:

1. Union ✓
2. Intersection ✓
3. Difference ✓
4. Complement ✓
5. P(AUB) formula ✓
6. And = x or = +

If A & B are 2 events from a sample space or
A & B are 2 events from an experiment

- $P[A \text{ or } B \text{ occurs}] = P[A \cup B]$
- $P[A \text{ and } B \text{ both occur}] = P[A \cap B]$

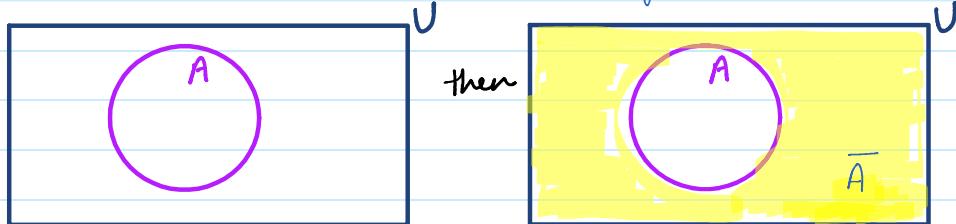
$A - B$, $B - A$ in sets are different
 $A \Delta B$ or symmetric difference $\rightarrow (A \cup B) - (A \cap B)$



- $P(A \cup B) = P(A) + P(B) - P(A \cap B)$
- complement of a set :

Universal set (U) : A set which contains all the elements under consideration for my experiment.

If $A \subset U$ then A^c or $\bar{A} = U - A$
↑
Subset (everything other than A)



1. $P(A)$ lies b/w : A - 1 to 1 C - 0 to 1
B - ∞ to ∞ D - 1 to 0

2. $P(U) = 1$

Extra information about what has already happened thereby reducing the sample space

$P(A|B)$: Prob. of event A occurring when event B has already occurred.

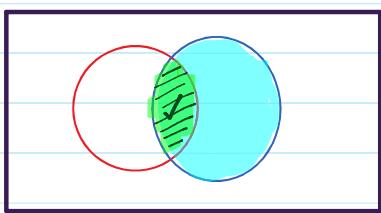
$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

For example, given that you drew a red card, what's the probability that it's a four ($P(\text{four}|\text{red}) = 2/26 = 1/13$).

$$P(\text{four}|\text{Red}) = P(\text{red} \cap \text{four}) = \frac{2/52}{1/2} = \frac{2/52}{1/2} = \frac{2}{52} \times 2 = \frac{4}{52} = \frac{2}{26} = \frac{1}{13}$$

Show on venn

Show on venn



$$P(A|B) = \frac{P(A \cap B)}{P(B)}$$

DEFINITION

For any two events A and B with $P(B) > 0$, the conditional probability of A given that B has occurred is defined by

$$P(A|B) = \frac{P(A \cap B)}{P(B)} \quad (2.3)$$

* Hint $P(A \cup B \cup C) = P(A) + P(B) + P(C) - P(A \cap B) - P(A \cap C) - P(B \cap C) + P(A \cap B \cap C)$

Example 2.26 A news magazine publishes three columns entitled "Art" (A), "Books" (B), and "Cinema" (C). Reading habits of a randomly selected reader with respect to these columns are

Read regularly	only A	only B	only C	$A \cap B$	$A \cap C$	$B \cap C$	$A \cap B \cap C$
Probability	.14	.23	.37	.08	.09	.13	.05

Figure 2.9 illustrates relevant probabilities.

calc. $P(A \cup B | C)$: 33

$$\begin{aligned} P(A \cup B) &= \frac{0.14 + 0.23}{0.37} \\ &= 0.29 \end{aligned}$$

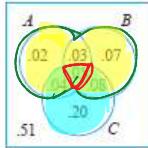


Figure 2.9 Venn diagram for Example 2.26

We thus have

$$P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{.08}{.23} = .348$$

$$P(A|B \cup C) = \frac{P(A \cap (B \cup C))}{P(B \cup C)} = \frac{.04 + .05 + .03}{.47} = \frac{.12}{.47} = .255$$

$$\begin{aligned} P(A|\text{reads at least one}) &= P(A|A \cup B \cup C) = \frac{P(A \cap (A \cup B \cup C))}{P(A \cup B \cup C)} \\ &= \frac{P(A)}{P(A \cup B \cup C)} = \frac{.14}{.49} = .286 \end{aligned}$$

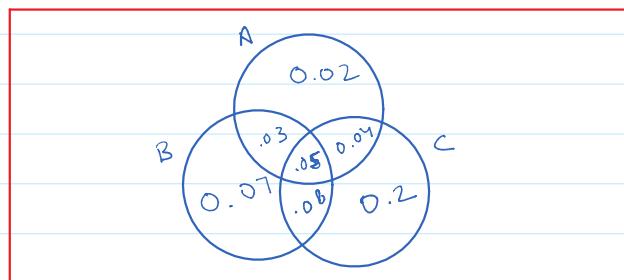
and

$$P(A \cup B | C) = \frac{P((A \cup B) \cap C)}{P(C)} = \frac{.04 + .05 + .08}{.37} = .459 \quad \blacksquare$$

$$P((A \cup B) \cap C) = P[(A \cap C) \cup (B \cap C)]$$

$$= P[A \cap C] + P[B \cap C] - P[(A \cap C) \cap (B \cap C)]$$

Read regularly	only A	only B	only C	$A \cap B$	$A \cap C$	$B \cap C$	$A \cap B \cap C$
Probability	.14	.23	.37	.08	.09	.13	.05



Example 1: The table below shows the occurrence of diabetes in 100 people. Let D and N be the events where a randomly selected person "has diabetes" and "not overweight". Then find $P(D | N)$.

	Diabetes (D)	No Diabetes (D')
Not overweight (N)	5 D ∩ N	45 N
Overweight (N')	17	33

Solution:

From the given table, $P(N) = (5+45) / 100 = 50/100$.

$$P(D \cap N) = 5/100.$$

By the conditional probability formula,

$$P(D | N) = P(D \cap N) / P(N)$$

$$= (5/100) / (50/100)$$

$$= 5/50$$

$$= 1/10$$

HW

Example 3: If a fair die is rolled twice, observe the numbers that face up. Find the conditional probability that the sum of the numbers is 7, given that the first number is 2.

Solution:

Let us determine the sample space of rolling a die twice. $S = \{(1,1), (1,2), (1,3), (1,4), (1,5), (1,6), (2,1), (2,2), (2,3), (2,4), (2,5), (2,6), (3,1), (3,2), (3,3), (3,4), (3,5), (3,6), (4,1), (4,2), (4,3), (4,4), (4,5), (4,6), (5,1), (5,2), (5,3), (5,4), (5,5), (5,6), (6,1), (6,2), (6,3), (6,4), (6,5), (6,6)\}$

Considering events A and B as given: we have

A : the sum of the numbers is 7. Thus set A = {(1,6), (2,5), (3,4), (4,3), (5,2), (6,1)}

B: the first number is 2. Thus set B = {(2,1), (2,2), (2,3), (2,4), (2,5), (2,6)}

$$A \cap B: \{(2,5)\}$$

By the conditional probability, we know that

$$P(A) = P(A \cap B) / P(B)$$

$$P(A) = \frac{1}{36}$$

$$\frac{6}{36}$$

$$P(A) = 1/6$$

Answer: The conditional probability that the sum of the numbers is 7, given that the first number is 2 is $1/6$

HW

Example: Two dies are thrown simultaneously and the sum of the numbers obtained is found to be 7. What is the probability that the number 3 has appeared at least once?

Solution: The sample space S would consist of all the numbers possible by the combination of two dies. Therefore S consists of 6×6 i.e. 36 events.

Event A indicates the combination in which 3 has appeared at least once.

Event B indicates the combination of the numbers which sum up to 7.

$$A = \{(3, 1), (3, 2), (3, 3), (3, 4), (3, 5), (3, 6), (1, 3), (2, 3), (4, 3), (5, 3), (6, 3)\}$$

$$B = \{(1, 6), (2, 5), (3, 4), (4, 3), (5, 2), (6, 1)\}$$

$$P(A) = 11/36$$

$$P(B) = 6/36$$

$$A \cap B = 2$$

$$P(A \cap B) = 2/36$$

Applying the conditional probability formula we get,

$D \rightarrow \text{diabetes}$

$N \rightarrow \text{not overweight}$

$$P(D | N) = \frac{P(D \cap N)}{P(N)} = \frac{5 / (5+45+17+33)}{(5+45) / (5+45+17+33)} = \frac{5 / 100}{50 / 100} = \frac{1}{10}$$

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$$P(A) = 11/36$$

$$P(B) = 6/36$$

$$A \cap B = 2$$

$$P(A \cap B) = 2/36$$

Applying the conditional probability formula we get,

$$P(A|B)$$

$$= \frac{P(A \cap B)}{P(B)}$$

$$= \frac{\frac{2}{36}}{\frac{6}{36}}$$

$$= \frac{1}{3}$$

H.W

Example 3 Ten cards numbered 1 to 10 are placed in a box, mixed up thoroughly and then one card is drawn randomly. If it is known that the number on the drawn card is more than 3, what is the probability that it is an even number?

Solution Let A be the event 'the number on the card drawn is even' and B be the event 'the number on the card drawn is greater than 3'. We have to find $P(A|B)$.

Now, the sample space of the experiment is $S = \{1, 2, 3, 4, 5, 6, 7, 8, 9, 10\}$

$$\text{Then } A = \{2, 4, 6, 8, 10\}, B = \{4, 5, 6, 7, 8, 9, 10\}$$

and $A \cap B = \{4, 6, 8, 10\}$

$$\text{Also } P(A) = \frac{5}{10}, P(B) = \frac{7}{10} \text{ and } P(A \cap B) = \frac{4}{10}$$

$$\text{Then } P(A|B) = \frac{P(A \cap B)}{P(B)} = \frac{\frac{4}{10}}{\frac{7}{10}} = \frac{4}{7}$$

Dependent and Independent events

Statistical definition: Two events A & B are said to be statistically independent if and only if $[P(A) \times P(B) = P(A \cap B)]$

Example 10 A die is thrown. If E is the event 'the number appearing is a multiple of 3' and F be the event 'the number appearing is even' then find whether E and F are independent?

Solution We know that the sample space is $S = \{1, 2, 3, 4, 5, 6\}$

$$\text{Now } E = \{3, 6\}, F = \{2, 4, 6\} \text{ and } E \cap F = \{6\}$$

$$\text{Then } P(E) = \frac{2}{6} = \frac{1}{3}, P(F) = \frac{3}{6} = \frac{1}{2} \text{ and } P(E \cap F) = \frac{1}{6}$$

$$\text{Clearly } P(E \cap F) = P(E) \cdot P(F)$$

Hence E and F are independent events.

$E \rightarrow$ mult of 3
 $F \rightarrow$ no. is even.

$$P(E) = 2/6$$

$$P(F) = 3/6$$

$$P(E \cap F) = 1/6$$

$$P(E) \cdot P(F) = \frac{2}{6} \times \frac{3}{6} = \frac{1}{6}$$

$$\therefore P(E) \cdot P(F) = P(E \cap F)$$

\Rightarrow independent ✓

Example 11 An unbiased die is thrown twice. Let the event A be 'odd number on the first throw' and B the event 'odd number on the second throw'. Check the independence of the events A and B.

Solution If all the 36 elementary events of the experiment are considered to be equally likely, we have

$$P(A) = \frac{18}{36} = \frac{1}{2} \text{ and } P(B) = \frac{18}{36} = \frac{1}{2}$$

$$\text{Also } P(A \cap B) = P(\text{odd number on both throws})$$

$$= \frac{9}{36} = \frac{1}{4}$$

$$\text{Now } P(A) \cdot P(B) = \frac{1}{2} \times \frac{1}{2} = \frac{1}{4}$$

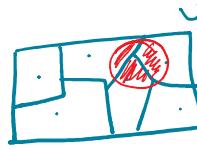
$$\text{Clearly } P(A \cap B) = P(A) \times P(B)$$

Thus, A and B are independent events

Bayes Theorem

The Bayes theorem establishes the likelihood of an event occurring given any condition. It is considered for the case of conditional probability. Also, this is known as the formula for the likelihood of "causes".

$$P(A|B) = \frac{P(B|A)P(A)}{P(B)}$$



- A, B = Events
- $P(A|B)$ = Probability of A given B is true
- $P(B|A)$ = Probability of B given A is true
- $P(A), P(B)$ = The independent probabilities of A and B

If A_1, A_2, \dots, A_n is the partition of U and B is any other event.

$$P(A_i|B) = \frac{P(B|A_i)P(A_i)}{P(B|A_1)P(A_1) + P(B|A_2)P(A_2) + \dots + P(B|A_n)P(A_n)}$$

Example 16 Bag I contains 3 red and 4 black balls while another Bag II contains 5 red and 6 black balls. One ball is drawn at random from one of the bags and it is found to be red. Find the probability that it was drawn from Bag II. *given info*

Solution Let E_1 be the event of choosing the bag I, E_2 the event of choosing the bag II and A be the event of drawing a red ball.

Then $P(E_1) = P(E_2) = \frac{1}{2}$

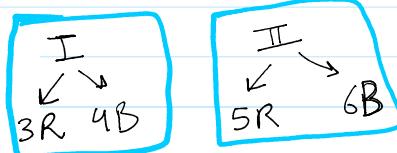
Also $P(A|E_1) = P(\text{drawing a red ball from Bag I}) = \frac{3}{7}$

and $P(A|E_2) = P(\text{drawing a red ball from Bag II}) = \frac{5}{11}$

Now, the probability of drawing a ball from Bag II, being given that it is red, is $P(E_2|A)$

By using Bayes' theorem, we have

$$P(E_2|A) = \frac{P(E_2)P(A|E_2)}{P(E_1)P(A|E_1) + P(E_2)P(A|E_2)} = \frac{\frac{1}{2} \times \frac{5}{11}}{\frac{1}{2} \times \frac{3}{7} + \frac{1}{2} \times \frac{5}{11}} = \frac{35}{68}$$



1 ball is drawn
→ ball is red

$$P(\text{II}|R)$$

$$\frac{P(\text{II} \cap R)}{P(R)}$$

using Bayes theorem here:

$$P(\text{II}|R) = \frac{P(R|\text{II})P(\text{II})}{P(R|\text{II})P(\text{II}) + P(R|\text{I})P(\text{I})}$$

$$\begin{aligned} P(\text{I}) &= \frac{1}{2} & P(R|\text{II}) &= \frac{5}{11} \\ P(\text{II}) &= \frac{1}{2} & P(R|\text{I}) &= \frac{3}{7} \end{aligned}$$

$$P(\text{II}|R) = \frac{\frac{5}{11} \times \frac{1}{2}}{\left(\frac{5}{11} \times \frac{1}{2}\right) + \left(\frac{3}{7} \times \frac{1}{2}\right)} = \frac{\frac{5}{11}}{\frac{5}{11} + \frac{3}{7}} = \frac{\frac{5}{11}}{\frac{68}{77}} = \frac{35}{68}$$

∴ $P(\text{II}|R) = \frac{35}{68}$

Example 18 Suppose that the reliability of a HIV test is specified as follows:

Of people having HIV, 90% of the test detect the disease but 10% go undetected. Of people free of HIV, 99% of the test are judged HIV-ive but 1% are diagnosed as showing HIV+ive. From a large population of which only 0.1% have HIV, one person is selected at random, given the HIV test, and the pathologist reports him/her as HIV+ive. What is the probability that the person actually has HIV?

Solution Let E denote the event that the person selected is actually having HIV and A the event that the person's HIV test is diagnosed as +ive. We need to find $P(E|A)$.

Also E' denotes the event that the person selected is actually not having HIV.

Clearly, (E, E') is a partition of the sample space of all people in the population. We are given that

$$P(E) = 0.1\% = \frac{0.1}{100} = 0.001$$

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$$\begin{aligned} P(E') &= 1 - P(E) = 0.99 \\ P(A|E) &= P(\text{Person tested as HIV+ive given that he/she is actually having HIV}) \\ &= 90\% = \frac{90}{100} = 0.9 \end{aligned}$$

and

$$\begin{aligned} P(A|E') &= P(\text{Person tested as HIV+ive given that he/she is actually not having HIV}) \\ &= 1\% = \frac{1}{100} = 0.01 \end{aligned}$$

Now, by Bayes' theorem

$$\begin{aligned} P(E|A) &= \frac{P(E)P(A|E)}{P(E)P(A|E) + P(E')P(A|E')} \\ &= \frac{0.001 \times 0.9}{0.001 \times 0.9 + 0.999 \times 0.01} = \frac{90}{1089} \\ &= 0.083 \text{ approx.} \end{aligned}$$

Thus, the probability that a person selected at random is actually having HIV given that he/she is tested HIV+ive is 0.083.

XIV

Example 21 A man is known to speak truth 3 out of 4 times. He throws a die and reports that it is a six. Find the probability that it is actually a six.

Solution Let E be the event that the man reports that six occurs in the throwing of the die and let S_1 be the event that six occurs and S_2 be the event that six does not occur.

Then $P(S_1) = \text{Probability that six occurs} = \frac{1}{6}$

$P(S_2) = \text{Probability that six does not occur} = \frac{5}{6}$

$P(E|S_1) = \text{Probability that the man reports that six occurs when six has actually occurred on the die}$

$= \text{Probability that the man speaks the truth} = \frac{3}{4}$

$P(E|S_2) = \text{Probability that the man reports that six occurs when six has not actually occurred on the die}$

$= \text{Probability that the man does not speak the truth} = 1 - \frac{3}{4} = \frac{1}{4}$

Thus, by Bayes' theorem, we get

$P(S_1|E) = \text{Probability that the report of the man that six has occurred is actually a six}$

$= \frac{P(S_1)P(E|S_1)}{P(S_1)P(E|S_1) + P(S_2)P(E|S_2)}$

$$= \frac{\frac{1}{6} \times \frac{3}{4}}{\frac{1}{6} \times \frac{3}{4} + \frac{5}{6} \times \frac{1}{4}} = \frac{1}{8} \times \frac{24}{8} = \frac{3}{8}$$

Hence, the required probability is $\frac{3}{8}$

For intro to inf stats see old class deck <theoretical concepts are present only>

- Sample & Population :
- Statistic & Parameter
- changes w/ samp · fixed
- function of samp · function of pop

sample ⊂ Pop^n

$$\rightarrow \text{sample mean } (\bar{x}) = \frac{\sum_{i=1}^n x_i}{n} = \frac{\text{pop}^n(\mu)}{\sum_{i=1}^n x_i/N}$$

$$\rightarrow \text{sample var. } (\sigma^2) = \frac{1}{n-1} \sum_{i=1}^n (x_i - \bar{x})^2 \quad \text{pop var. } \sigma^2 = \frac{1}{N} \sum (x_i - \mu)^2$$



Types of Sampling.

$X \rightarrow$ random variable : X - represents a population which follows a probability distribution. Since X can take various values at different trials, there is a probability attached to each value or a range of values that X can take, this combination of the values that X takes and the attached probability / probability density make up a probability distribution.

Probability distributions

We will be looking upon some useful p.d. today.

Revision of random variable and properties of probability distribution like pdf, range, definition, $E(X)$, $V(X)$

\rightarrow Definition, Range, $\underset{\text{unit.}}{\text{PDF/PMF}}$, $\underset{\text{discrete}}{E(X), V(X)}$

Discrete vs continuous

Probability Density Function and Probability Mass Function

- PDF is a statistical term that describes the probability distribution of the **Continuous** random variable

Probability Density Function

$$F(x) = P(a \leq x \leq b) = \int_a^b f(x) dx \geq 0$$

to calculate prob. we need to integrate pdf over the required range.

- PMF is a statistical term that describes the probability distribution of the **Discrete** random variable

$$p(x) = P(X = x)$$

The probability of x = The probability ($X = \text{one specific } x$)

- People often get confused between PDF and PMF. The PDF is applicable for continuous random variables while PMF is applicable for discrete random variables

Examples:

DISCRETE

1. uniform \rightarrow It has uniformity of probability at each value of X

7.8. Discrete Uniform Distribution. A random variable X is said to have uniform distribution on n points $[x_1, x_2, \dots, x_n]$ if its p.m.f. is given by :

$$P(X = x_i) = \frac{1}{n}, i = 1, 2, \dots, n \quad \dots(7.31)$$

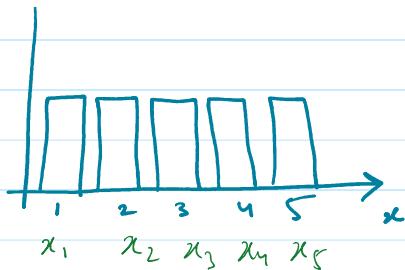


7-8. Discrete Uniform Distribution. A random variable X is said to have uniform distribution on n points $\{x_1, x_2, \dots, x_n\}$ if its p.m.f. is given by :

$$P(X = x_i) = \frac{1}{n}; i = 1, 2, \dots, n \quad \dots(7.31)$$

For example, if X has a uniform distribution on the points $\{0, 1, 2, \dots, n\}$, then $P(X = i) = \frac{1}{n+1}; i = 0, 1, 2, \dots, n$. $\dots(7.31a)$

Such distributions can be conceived in practice if under the given experimental conditions, the different values of the random variable become equally likely. Thus for a die experiment, and for an experiment with a deck of cards such distribution is appropriate.



If $X \sim U(5, 100)$ find $P[20 < X < 35]$

eg: $X \sim U(5, 100)$ find $P[20 < X < 35]$

$$= \frac{14}{96} = \frac{5+6+7+\dots+100}{96} = \frac{5040}{96} = 52.5$$

range: $5 \leq x \leq 100$

$$E(X) = \sum_{i=5}^{100} i \cdot \frac{1}{96} = \sum p_i x_i = \frac{\sum x_i}{n}$$

$$E(X) = \frac{a+b}{2} = \frac{5+100}{2} = \frac{105}{2} = 52.5$$

2. Bernoulli

7-1. Bernoulli Distribution. A random variable X which takes two values 0 and 1, with probabilities q and p respectively, i.e., $P(X = 1) = p$, $P(X = 0) = q$, $q = 1 - p$ is called a *Bernoulli variate* and is said to have a Bernoulli distribution.

X	$P(x)$
1	p
0	q

success
failure

$$N(X) = pq$$

$$E(X) = \sum p_i x_i = 1(p) + 0(q) \\ = p$$

$$\begin{aligned} V(X) &= E(X^2) - [E(X)]^2 \\ &= (1^2 \cdot p + 0^2 \cdot q) - p^2 \\ &= p - p^2 = p(1-p) = pq. \end{aligned}$$

3. Binomial

used a lot

If $X \sim \text{Bin}(n, p)$ then X measures

the number of successes in ' n ' trials of a bernoulli experiment w/ prob of success as ' p ' in each trial.

7-2

Fundamentals of Mathematical Statistics

being finite), in which the probability 'p' of success in any trial is constant for each trial. Then $q = 1 - p$, is the probability of failure in any trial.

$$p(x) = {}^n C_x \cdot p^x \cdot q^{n-x}$$

$$\text{eg. } n=5 \\ \dots x=3 \\ p = \left(\frac{1}{4}\right)$$

$$\begin{aligned} \frac{5}{4} \cdot \frac{5}{4} \cdot \frac{5}{4} \cdot \frac{1}{4} \cdot \frac{1}{4} &\rightarrow \left(\frac{1}{4}\right)^3 \left(\frac{3}{4}\right)^2 \\ \frac{1}{4} \cdot \frac{1}{4} \cdot \frac{1}{4} \cdot \frac{3}{4} \cdot \frac{1}{4} &\rightarrow \left(\frac{1}{4}\right)^3 \left(\frac{3}{4}\right)^2 \end{aligned}$$

eg. $n=3$ $P(x=3)$

$$E(X) = \sum p_i x_i = \sum_{x=0}^n {}^n C_x p^x q^{n-x} \cdot x = np$$

$$V(X) = npq$$

Example 7-11. Determine the binomial distribution for which the mean is 4 and variance 3. (Madurai Kamraj Univ B.Sc. 1993)

Solution, Let $X \sim B(n, p)$, then we are given that

$$E(X) = np = 4 \quad \dots (*)$$

and

$$\text{Var}(X) = npq = 3 \quad \dots (**)$$

Dividing (**) by (*), we get

$$q = \frac{3}{4} \Rightarrow p = 1 - q = \frac{1}{4}$$

Hence from (*), $n = \frac{4}{p} = 16$

Thus the given binomial distribution has parameters $n = 16$ and $p = 1/4$.

$$E(X) = 4$$

$$V(X) = 3$$

$$np = 4 \Rightarrow \frac{npq}{np} = \frac{3}{4}$$

$$npq = 3 \Rightarrow q = \frac{3}{4}$$

$$\Rightarrow p = \frac{1}{4}$$

$$np = 4 \Rightarrow n \left(\frac{1}{4} \right) = 4$$

$$\Rightarrow n = 16$$

hint: Range of $Y = \{0, 1, 2, 3, 4, 5\}$

eg. $Y \sim \text{Bin}(n=5, p=1/3)$ find $P[Y \leq 1]$
 $P[Y \leq 1] = P[Y=0] + P[Y=1]$

$$= {}^5 C_0 \left(\frac{1}{3}\right)^0 \left(\frac{2}{3}\right)^5 + {}^5 C_1 \left(\frac{1}{3}\right)^1 \left(\frac{2}{3}\right)^4$$

$$= 1 \times 1 \times \left(\frac{2}{3}\right)^5 + 5 \times \frac{1}{3} \times \left(\frac{2}{3}\right)^4$$

$$= 0.131 + \frac{5}{3} \times \frac{16}{81}$$

$$= 0.131 + 0.329$$

$$= 0.46$$

4. Poisson

mostly it is used when we are measuring occurrences of an event.

Definition. A random variable X is said to follow a Poisson distribution

if it assumes only non-negative values and its probability mass function is given by

$$P(x, \lambda) = P(X=x) = \frac{e^{-\lambda} \lambda^x}{x!}; x = 0, 1, 2, \dots; \lambda > 0$$

$$= 0, \text{ otherwise}$$

...(7-14)

$$p(x; \lambda) = P(X = x) = \frac{\lambda^x}{x!} e^{-\lambda}; x = 0, 1, 2, \dots; \lambda > 0$$

= 0, otherwise

... (7.14)

Here λ is known as the parameter of the distribution.

We shall use the notation $X \sim P(\lambda)$ to denote that X is a Poisson variate with parameter λ .

$$X \sim P(\lambda)$$

PMF :

$$p(x) = \frac{e^{-\lambda} \lambda^x}{x!}$$

$$\mathbb{E}(X) = \lambda \quad \text{V}(X) = \lambda$$

Range: 0, 1, 2, 3, 4, ...

mathematical constant
 $e \approx 2.7182$

Example 7.24. A car hire firm has two cars which it hires out day by day. The number of demands for a car on each day is distributed as Poisson variate with mean 1.5. Calculate the proportion of days on which (i) neither car is used, and (ii) some demand is refused. [Meerut Univ. B.Sc. 1993]

Solution. The proportion of days on which there are x demands for a car

$$= P\{|x \text{ demands in a day}\}$$

$$= \frac{e^{-1.5} (1.5)^x}{x!},$$

since the number of demands for a car on any day is a Poisson variate with mean 1.5. Thus

$$P(X = x) = \frac{e^{-1.5} (1.5)^x}{x!}; x = 0, 1, 2, \dots$$

(i) Proportion of days on which neither car is used is given by

$$P(X = 0) = e^{-1.5}$$

$$= \left[1 - 1.5 + \frac{(1.5)^2}{2!} - \frac{(1.5)^3}{3!} + \frac{(1.5)^4}{4!} - \dots \right]$$

$$= 0.2231$$

(ii) Proportion of days on which some demand is refused is

$$P(X > 2) = 1 - P(X \leq 2)$$

$$= 1 - [P(X = 0) + P(X = 1) + P(X = 2)]$$

$$= 1 - e^{-1.5} \left[1 + 1.5 + \frac{(1.5)^2}{2!} \right]$$

$$= 1 - 0.2231 \times 3.625 = 0.19126$$

$$\text{mean} = 1.5 \quad \lambda = 1.5 \quad X \sim P(1.5)$$

$$(i) P(X=0)$$

$$e^{-1.5} \frac{(1.5)^0}{0!} = e^{-1.5} = 0.2231$$

$$= (2.7182)^{-1.5} = 0.2231$$

$$(ii) P(X > 2)$$

$$0, 1, 2, \boxed{3, 4, 5, \dots}$$

needed

Since total prob. = 1

$$P(X > 2) = 1 - P(X \leq 2) = 1 - \left[\frac{e^{-1.5} (1.5)^0}{0!} + \frac{e^{-1.5} (1.5)^1}{1!} + \frac{e^{-1.5} (1.5)^2}{2!} \right]$$

complete as HW

CONTINUOUS

5. Normal

Definition. A random variable X is said to have a normal distribution with parameters μ (called "mean") and σ^2 (called "variance") if its density function is given by the probability law:

$$f(x; \mu, \sigma) = \frac{1}{\sigma \sqrt{2\pi}} \exp \left[-\frac{1}{2} \left\{ \frac{x-\mu}{\sigma} \right\}^2 \right]$$

$$\text{or } f(x; \mu, \sigma) = \frac{1}{\sigma \sqrt{2\pi}} e^{-(x-\mu)^2/2\sigma^2},$$

$-\infty < x < \infty, -\infty < \mu < \infty, \sigma > 0$... (8.3)

~~imp.~~ RANGE ALWAYS

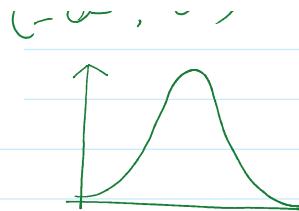
$(-\infty, \infty)$



or $f(x; \mu, \sigma) = \frac{1}{\sigma \sqrt{2\pi}} e^{-(x-\mu)^2/2\sigma^2}$
 $\quad -\infty < x < \infty, -\infty < \mu < \infty, \sigma > 0$... (8.3)

SNV

$E(X) = \mu \quad V(X) = \sigma^2$



Example 8.12. X is normally distributed and the mean of X is 12 and S.D.

is 4. (a) Find out the probability of the following :

- (i) $X \geq 20$, (ii) $X \leq 20$, and (iii) $0 \leq X \leq 12$

H.W

H.W

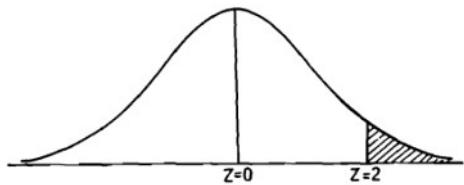
$x \sim N(12, 4^2)$

(i) $P(X \geq 20) = ?$

When $X = 20$, $Z = \frac{20-12}{4} = 2$

$\therefore P(X \geq 20) = P(Z \geq 2) = 0.5 - P(0 \leq Z \leq 2) = 0.5 - 0.4772 = 0.0228$

(ii) $P(X \leq 20) = 1 - P(X \geq 20)$
 $= 1 - 0.0228 = 0.9772$ (\because Total probability = 1)



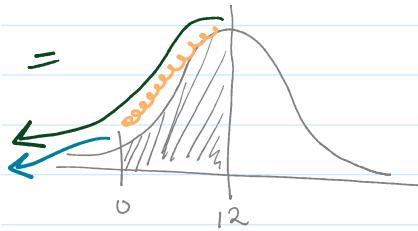
(iii) $P(0 \leq X \leq 12) = P(-3 \leq Z \leq 0)$

$= P(0 \leq Z \leq 3) = 0.49865$

$$\left(Z = \frac{X-12}{4} \right)$$

(From symmetry)

$P(0 \leq X \leq 12)$

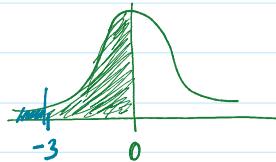


$P(X \leq 12) - P(X \leq 0)$

= reqd. prob.

Converting to std. normal using 2 easy steps.

$$\begin{aligned} P\left[\frac{X-12}{4} \leq \frac{12-12}{4}\right] - P\left[\frac{X-12}{4} \leq \frac{0-12}{4}\right] \\ = P[Z \leq 0] - P[Z \leq -3] \\ = \frac{1}{2} - 0.0013 \\ = \underline{\underline{0.4987}} \end{aligned}$$



6. Exponential

8.6. The Exponential Distribution. A continuous random variable X assuming non-negative values is said to have an exponential distribution with parameter $\theta > 0$, if its p.d.f. is given by

$$f(x) = \begin{cases} \theta e^{-\theta x}, & x \geq 0 \\ 0, & \text{otherwise} \end{cases} \quad \dots (8.24)$$

$x \sim \exp(\theta)$

$X > 0 \text{ or } x \in [0, \infty)$

p.d.f.: $\theta e^{-\theta x}$

$V(X) = \frac{1}{\theta^2}$

$E(X) = \frac{1}{\theta}$

Example: calculate $P[X < 5]$ if $X \sim \exp(0.25)$

$x \sim \exp(0.25) : P(X < s)$

in continuous
 $P(X \leq s)$ is
also same.

also same.

$$\begin{aligned}\int_0^5 \theta e^{-\theta x} dx &= \int_0^5 \frac{1}{4} e^{-x/4} dx \\&= \frac{1}{4} \left[\frac{e^{-x/4}}{-1/4} \right]_0^5 = \frac{1}{4} \times -\frac{1}{1} \times \left[e^{-5/4} - e^0 \right] \\&= -1 \{ 0.2865 - 1 \} \\&= -1 \{ -0.713 \} \\&= \boxed{0.713}\end{aligned}$$